### Problem Set: Non Parametric Statistics: Group 18

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3/30/2021

### Question 1

### Question 2

Compare the MISE and AMISE criteria in three densities in nor1mix of your choice.

- 1. Code (2.33) and the AMISE expression for the normal kernel, and compare the two error curves.
- 2. Compare them for n = 100, 200, 500, adding a vertical line to represent the  $h_{MISE}$  and  $h_{AMISE}$  bandwidths. Describe in detail the results and the major takeaways.

For this exercise we require some mathematical ideas that we will developed briefly.

We start with the KDE estimator  $\hat{f}(x;h) = \sum_{i=1}^{n} K_h(x-X_i)$ . The expectation and variance for this estimator are given bit he following expressions:

(1) 
$$E[\hat{f}(x;h)] = (K_h * f)(x)$$
.

(2) 
$$\operatorname{Var}[\hat{f}(x;h)] = \frac{1}{n}((K_h^2 * f)(x) - (K_h * f)^2(x))$$

Then, we develop some assymptotic expression for (1) and (2):

(3) 
$$E[\hat{f}(x;h)] - f(x) = Bias[\hat{f}(x;h)] = \frac{1}{2}\mu_2(K)f''(x)h^2 + o(h^2)$$

(4) 
$$\operatorname{Var}[\hat{f}(x;h)] = \frac{R(K)}{nh} f(x) + o((nh)^{-1})$$

Then, from whe equations (3) and (4) we obtain the following expression for the MSE:

(5) 
$$MSE[\hat{f}(x;h)] = \frac{\mu_2^2(K)}{4} (f''(x))^2 h^4 + \frac{R(K)}{nh} f(x) + o(h^4 + (nh)^{-1})$$

In equations (3), (4) and (5) we have:

(6) 
$$\mu_2(K) := \int z^2 K(z) dz$$

(7) 
$$R(K) := \int (K(x))^2 dx$$

Then, we use  $\text{MISE}[\hat{f}(\cdot;h)]$  as global error criteria for measuring the performance of  $\hat{f}$  in relation to the target density f.

(8) 
$$MISE[\hat{f}(\cdot;h)] = \int MSE[\hat{f}(x;h)]$$

Then, we obtain the following asymptotic expansion for the MISE:

(9) 
$$\text{MISE}[\hat{f}(\cdot;h)] = \frac{1}{4}\mu_2^2(K)R(f^{"})h^4 + \frac{R(K)}{nh} + o(h^4 + (nh)^{-1})$$

We define the dominating part of equation (9) as  $AMISE[\hat{f}(\cdot;h)]$  defined as:

(10) AMISE
$$[\hat{f}(\cdot;h)] = \frac{1}{4}\mu_2^2(K)R(f'')h^4 + \frac{R(K)}{nh}$$

with the expression R(f'') given by:

(11) 
$$R(f'') = \int (f''(x))^2 dx$$

Finally, the bandwidth the minimizes the AMISE is:

(12) 
$$h_{AMISE} = \left[ \frac{R(K)}{\mu_2^2(K)R(f'')n} \right]^{1/5}$$

Now, we consider our particular case of study. In this case, we reduce our analysis considering:

- a) A normal kernel  $K_h(\cdot)$  with distribution  $\mathcal{N}(0,1)$
- b) The density function f is based on the family of normal r-mixtures:

(13) 
$$f(x; \mu, \sigma, \mathbf{w}) = \sum_{j=1}^{r} w_j \phi_{\sigma_j}(x - \mu_j)$$

where 
$$w_j \ge 0, j = 1, ..., r$$
 and  $\sum_{j=1}^{r} w_j = 1$ .

With this two expression, we can obtain a specific value for the AMISE in equation (10). With assumption a), we obtain the following expressions for equations (6) and (7)

(6.1) 
$$\mu_2(K) = 1$$

(7.1) 
$$R(K) = \frac{1}{2\sqrt{\pi}}$$

Expression for equation (11) can be obtained from the following adaption of expression given in *Theorem 4.1* given by *Marron and Wand (1992)* 

(11.1) 
$$R(f'') = \int (f''(x))^2 dx =$$

With this expression, we obtain the reduced form of the AMISE:

(10.1) AMISE
$$[\hat{f}(\cdot;h)] = \frac{1}{4}R(f'')h^4 + \frac{1}{2nh\sqrt{\pi}}$$

With optimal bandwidth  $h_{AMISE}$ 

(12.1) 
$$h_{AMISE} = \left[ \frac{(2\sqrt{\pi})^{-1}}{R(f'')n} \right]^{1/5}$$

Finally, under this assumptions, we obtain a explicit and exact MISE expression of equation (8):

(14) 
$$\mathrm{MISE}_r[\hat{f}(\cdot;h)] = (2\sqrt{\pi}nh)^{-1} + \boldsymbol{w}'\{(1-n^{-1})\Omega_2 - \Omega_1 + \Omega_0\}\boldsymbol{w}$$

with 
$$(\Omega_a)_{i,j} = \phi_{(ah^2 + \sigma_i^2 + \sigma_i^2)^{1/2}}(\mu_i - \mu_j)$$
 for  $i, j = 1, ..., r$ 

Finally, we can proceed with a numeric approach over equation (14) and obtain:

(15) 
$$\operatorname{arg\ min}_{h>0} \operatorname{MISE}[\hat{f}(\cdot;h)]$$

We start with a easy example using the dataset nor1mix::MW.nm1 that is distributed according to  $\mathcal{N}(0,1)$ 

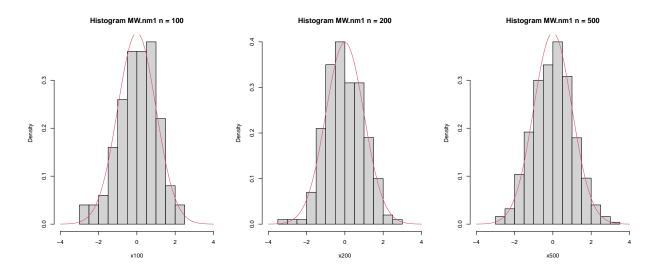


Figure 1: Histograms from sample obtained from n = 100, 200 and 500 for object MW.nm1

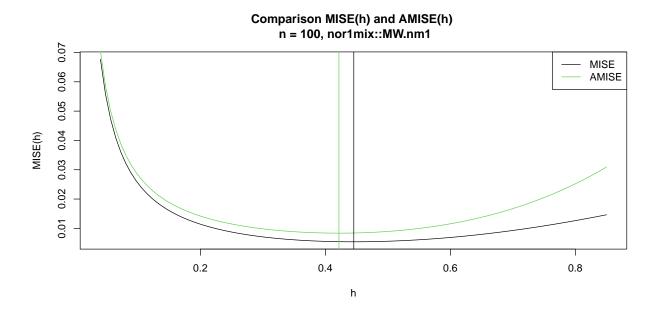


Figure 2: MISE and AMISE for range bandwith between 0.04 and 0.85 and n=100

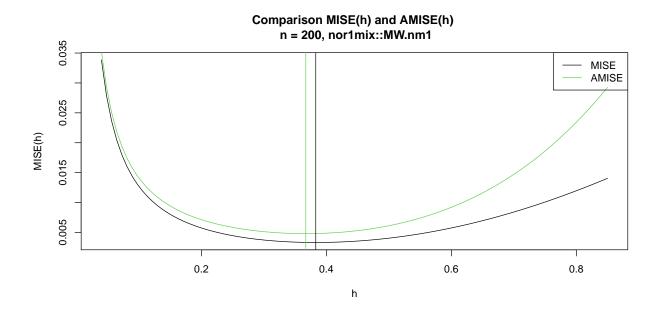


Figure 3: MISE and AMISE for range bandwith between 0.04 and 0.85 and  $\rm n=200$ 

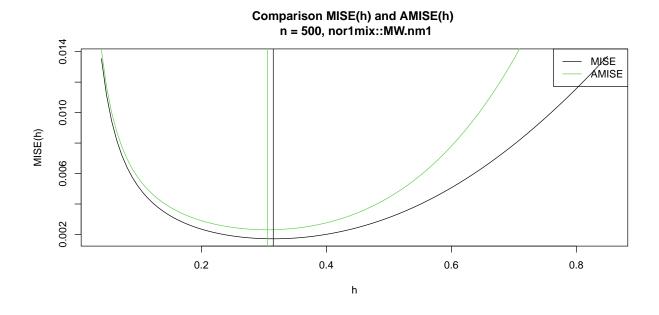


Figure 4: MISE and AMISE for range bandwith between 0.04 and 0.85 and n=500

### Question 3

Adapt the np\_pred\_CI function to include the argument type\_boot, which can take either the value "naive" or "wild".

- 1) If type\_boot = "wild", then the function must perform the wild bootstrap algorithm described above, implemented from scratch following substeps i—iv.
- 2) Compare and validate the correct behavior of the confidence intervals, for the two specifications of type\_boot, in the model considered in Exercise 5.8 (without performing the full simulation study).

For this question we describe briefly the necessary steps related with the implementation of the wild-bootstrap:

Fitted values:  $\hat{Y}_i := \hat{m}(X_i; q, h)$  with i = 1, ..., n.

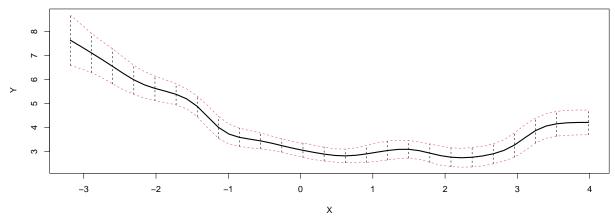
Prediction (Under a new grid):  $\hat{m}(x,q,h)$ . What about the uncertainty of  $\hat{m}(x;q,h)$ ?

- 1. Asymptotic approach:  $\hat{m}(x;q,h) \pm \hat{se}(\hat{m}(x;q,h))$
- 2. Bootstrap (we will focus on will bootstrap)
- 3. Compute  $\hat{m}(x;q,h) = \sum_{i=1}^{n} W_i^q(x) Y_i$  from the original sample  $(X_1,Y_1),...,(X_n,Y_n)$
- 4. Enter the wild bootstrap. For b = 1, ..., B.
- i. Simulate  $V_1^{*b}, ..., V_n^{*b}$  to be iid copies of V such that E[V] = 0 and Var[V] = 1,
- ii. Compute the pertubed residuals  $e_i^{*b} = \hat{e}_i V_i^{*b}$  where  $Y_i^{*b} := \hat{m}(X_i; q, h) + e_i^{*b}$  for i = 1, ..., n.
- iii. Obtain the bootstrap sample:  $(X_1, Y_1^{*b}), ..., (X_n, Y_n^{*b})$  where  $Y_i^{*b} := \hat{m}(X_i; q, h) + e_i^{*b}$  with i = 1, ..., n.
- iv. Compute  $\hat{m}^{*b}(x;q,h) = \sum_{i=1}^n W_i^q(x) Y_i^{*b}$  from  $(X_1^*,Y_1^{*b}),...,(X_n^8,Y_n^{*b})$

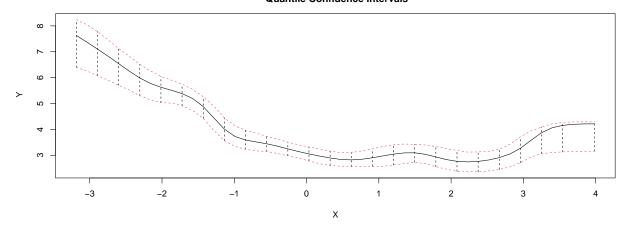
Our first simulation scenario is the following:

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- ## Warning in par(mforw = c(1, 2)): "mforw" is not a graphical parameter

## Bootstrap (npplot) Normal Approximation Confidence Intervals



### Bootstrap (npplot) Quantile Confidence Intervals



### Confidence Intervals quantile Naive Bootstrap

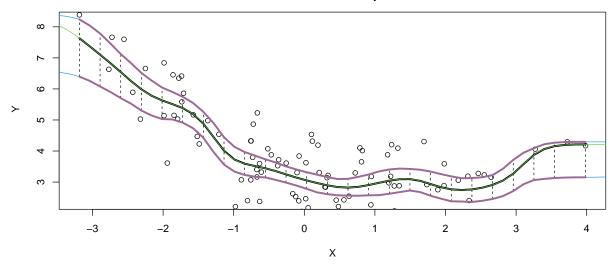


Figure 5: Function np\_pred\_CI: Naive Bootstrap

# Confidence Intervals quantile Wild Bootstrap – Normal perturbation

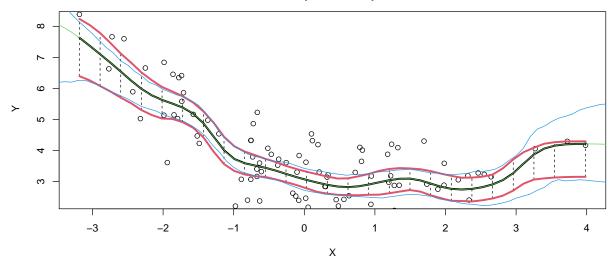


Figure 6: Function np\_pred\_CI: Wild Bootstrap and normal perturbation

# Confidence Intervals quantile Wild Bootstrap – Golden ratio perturbation

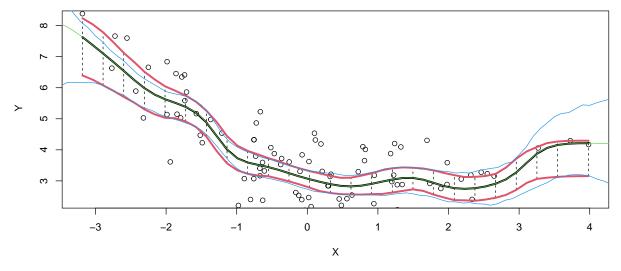


Figure 7: Function np\_pred\_CI: Wild Bootstrap and golden error perturbation