# HW<sub>1</sub>

## Exercise 1.

### Question 1.

 $Q_D(p)$  is  $p^{th}$  population quantile such that  $P(D \le Q_D(p)) = p$ .

In order word:

$$\int_{0}^{Q_{D}(p)} \lambda e^{-\lambda D} dD = 1 - e^{-\lambda Q_{D}(p)} = p$$

$$Q_{D}(p) = -\frac{1}{\lambda} \ln(1 - p)$$

#### Question 2

First empirical moment of the exponential distribution:

$$\hat{\mu} = D_n$$

Population moment of the exponential distribution:

$$E(D_1) = \frac{1}{\lambda}$$

The MOM estimator of  $\lambda$  is :  $\hat{\lambda}^{MOM} = {}^{D_n}$ 

Therefore the method of moments-based estimator of  $Q_D(p)$ :

$$Q_D(p)^{MOM} = -\frac{1}{\hat{\lambda}^{MOM}} - \ln(1-p) = -\frac{D_n \ln(1-p)}{\ln(1-p)}$$

#### Question3

From the CLT

$$\begin{array}{ccc}
 & \xrightarrow{D} & \xrightarrow{1} \\
\sqrt{n}(D_n - \lambda) & \xrightarrow{N \to \infty} \mathcal{N}(0, \lambda^2)
\end{array}$$

Hence, by Delta Method we can get, let g(t) = t \* ln(1-p) so g'(t) = ln(1-p)

$$\begin{array}{ccc}
 & & \mathcal{D} & \frac{(\ln(1-p))^2}{\sqrt{n}(\ln(1-p)^{D_n} + Q_D(p))} & \frac{\lambda^2}{\sqrt{n}} & \frac{\lambda^2}{\sqrt{n}}
\end{array}$$

Since 
$$\hat{\lambda}^{MOM} = {}^{D_n}$$

$$\sqrt{n}(\ln(1-p)^{\frac{1}{\lambda}} + Q_D(p)) \xrightarrow{D} \frac{(\ln(1-p))^2}{\lambda^2}$$

So, the *approximate*  $(1-\alpha)$ -confidence interval for  $Q_D(p)$  is

$$\begin{bmatrix} - & \frac{z_{1-\alpha/2} \times \ln(1-p)}{\lambda \sqrt{n}} & - & \frac{z_{1-\alpha/2} \times \ln(1-p)}{\lambda \sqrt{n}} \\ [-D_n \ln(1-p) - & \frac{\lambda \sqrt{n}}{\lambda \sqrt{n}} & - \end{bmatrix}$$

#### Question 4.

We know that if  $D_1, \dots, D_n$  are independent exponential random variables with parameter  $\lambda$ , then

$$\sum_{i=1}^{n} D_{i} \sim \Gamma(n, \lambda)$$

Therefore

$$\lambda^{D_n} = \frac{\lambda}{n} \sum_{i=1}^{n} D_i \sim \Gamma(n, n)$$

So,  $\lambda^{D_n}$  is independent of the parameter  $\lambda$ , which means it is an exact pivot.

From previous question

$$Q_D(p) = -\frac{1}{\lambda} \ln(1-p)$$

$$Q_D(0.5) = \frac{1}{\lambda} \ln(2)$$

To construct 95% confidence interval, let a and b be the 0.025 and 0.975 quantile of  $\Gamma(n,n)$ 

Therefore

$$P(a < \lambda^{D_n} < b) = 0.95$$

$$- - - \frac{D_{nln}(2)}{b} < \frac{1}{\lambda} \ln(2) < \frac{D_{nln}(2)}{a} = 0.95$$

Processing math: 100% rval is 
$$\begin{bmatrix} \frac{D_{nln(2)}}{b}, \frac{D_{nln(2)}}{a} \end{bmatrix}$$