

# Homework 1 Exercise 3

Chutian Chen cc4515; Congcheng Yan cy2550; Mingrui Liu ml4404

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## Exercise 3

1. Let  $Y$  has a standard normal distribution. Let  $X$  has a normal distribution  $N(\mu, \sigma^2)$ . We have

$$\sigma^2 = E(X^2) - (E(X))^2,$$

$$E(X^2) = \sigma^2 + \mu^2.$$

Then we have some conclusions about  $Y$ :

$$E(Y) = 0,$$

$$E(Y^2) = 0 + 1 = 1,$$

$$E(Y^3) = \int_{-\infty}^{\infty} x^3 \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} dx = 0.$$

Go back to  $\gamma$ ,

$$\begin{aligned}\gamma &= \mathbb{E}[R_1^3] \\ &= E[(Y * \sigma + \mu)^3] \\ &= E[Y^3 * \sigma^3 + 3 * Y^2 * \sigma^2 * \mu + 3 * Y * \sigma * \mu^2 + \mu^3] \\ &= E[3 * Y^2 * \mu * \sigma^2 + \mu^3] \\ &= \mu^3 + 3 * \mu * \sigma^2\end{aligned}$$

2.

(a)  $\frac{1}{n} \sum_{i=1}^n R_i$  has a normal distribution  $\mathbf{N}(\mu, \sigma^2/n)$ , so

$$\begin{aligned} E[\hat{\gamma}] - \gamma &= \mu^3 + 3 * \frac{\mu}{n} * \sigma^2 - \mu^3 - 3 * \mu * \sigma^2 \\ &= 3 * \frac{(1-n)\mu}{n} * \sigma^2 \end{aligned}$$

(b) Yes, it's consistent.

$$\begin{aligned} E[\hat{\gamma}^2] - (E[\hat{\gamma}])^2 &= \frac{O(n^6) + 6! * C_6^6 * \sum_{i=1}^6 R_i}{n^6} - \mu^6 - O\left(\frac{1}{n}\right) \\ &\xrightarrow{n \rightarrow \infty} 0 \end{aligned}$$

3. The bias of  $(\frac{1}{n} \sum_{i=1}^n R_i)^3$  is  $3 * \frac{(1-n)\mu}{n} * \sigma^2$ . From previous work, we also have

$$\begin{aligned} \gamma &= \mu^3 + 3 * \mu * \sigma^2 \\ E[\hat{\gamma}] &= \mu^3 + 3 * \frac{\mu}{n} * \sigma^2 \\ E[R_1 R_2 R_3] &= \mu^3 \end{aligned}$$

We can easily have an unbiased estimator  $n(\frac{1}{n} \sum_{i=1}^n R_i)^3 - (n-1)R_1 R_2 R_3$ .

4.

(a)

$$\begin{aligned} E[\tilde{\gamma}] - \gamma &= \frac{n}{n} E[R_1^3] - E[R_1^3] \\ &= 0 \end{aligned}$$

(b) Yes, it's consistent.

$$\begin{aligned}
E[\tilde{\gamma}^2] - (E[\tilde{\gamma}])^2 &= \frac{(\sum_{i=1}^n Y_i)^2}{n^2} - (\mu^3 + 3 * \mu * \sigma^2)^2 \\
&= \frac{O(n^2) + n * (n-1)E[R_1^3 R_2^3]}{n^2} - (\mu^3 + 3 * \mu * \sigma^2)^2 \\
&= (\mu^3 + 3 * \mu * \sigma^2)^2 - (\mu^3 + 3 * \mu * \sigma^2)^2 \\
&\xrightarrow{n \rightarrow \infty} 0
\end{aligned}$$

5. Based on previous work, we have two sufficient statistics  $\bar{R} = \frac{1}{n} \sum_{i=1}^n R_i$  and  $\bar{R}^2 = \frac{1}{n} \sum_{i=1}^n R_i^2$ . We also have unbiased and consistent estimator  $\tilde{\gamma} = \frac{1}{n} \sum_{i=1}^n R_i^3$ . The  $MVUE = E[\tilde{\gamma}|T]$ .

$$\begin{aligned}
E[\tilde{\gamma}|T] &= E\left[\frac{1}{n} \sum_{i=1}^n R_i^3 | \bar{R}, \bar{R}^2\right] \\
&= E\left[\frac{1}{n} \sum_{i=1}^n R_i - \bar{R} + \bar{R}^3 | \bar{R}, \bar{R}^2\right] \\
&= E\left[\frac{1}{n} \sum_{i=1}^n [(R_i - \bar{R})^3 + 3(R_i - \bar{R})^2 \bar{R} + 3(R_i - \bar{R}) \bar{R}^2 + \bar{R}^3] | \bar{R}, \bar{R}^2\right]
\end{aligned}$$

Similar with the derivation of  $E(Y) = 0, E(Y^3) = 0$ , in question 1, we get

$$E(R_i - \bar{R}) = 0, E((R_i - \bar{R})^3) = 0. \text{ So,}$$

$$\begin{aligned}
E[\tilde{\gamma}|T] &= E\left[\frac{1}{n} \sum_{i=1}^n [3(R_i - \bar{R})^2 \bar{R} + \bar{R}^3] | \bar{R}, \bar{R}^2\right] \\
&= E[3\bar{R}_i^2 \bar{R} - 6R_i \bar{R}^2 + 4\bar{R}^3 | \bar{R}, \bar{R}^2] \\
&= 3\bar{R}_i^2 \bar{R} - 2\bar{R}^3
\end{aligned}$$