Exercise2_cc4515_cy2550_ml4404

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```
[26]: import numpy as np
from scipy.misc import derivative
from scipy.stats import norm
from sympy import diff, symbols
import random
```

Exercise 2

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(1)

$$\begin{split} E(\bar{X}^2) &= E(\frac{\sum_{i=1}^n X_i^2 + 2\sum_{i=1}^{n-1} \sum_{j=i+1}^n X_i X_j}{n^2}) \\ &= \frac{1}{n} EX^2 + (1 - \frac{1}{n})(EX)^2 \\ EX^2 &= VarX + (EX)^2 = \lambda + \lambda^2 \\ EX &= \lambda \end{split}$$

So
$$E(\bar{X}^2) = \lambda^2 + \frac{\lambda}{n}$$

(2)

$$\begin{split} Es^2 &= \frac{1}{n-1} E(\sum_{i=1}^n X_i^2 - n\bar{X}^2) \\ &= \frac{1}{n-1} (n\lambda + n\lambda^2 - n\lambda^2 - \lambda) = \lambda \end{split}$$

(3)

$$\begin{split} EY_i &= EX_i^2 - 2\lambda EX_i + \lambda^2 - EX_i \\ &= \lambda^2 + \lambda - 2\lambda^2 + \lambda^2 - \lambda = 0 \\ VarY_i &= EY_i^2 - (EY_i)^2 = EY_i^2 = E(X_i^2 - (2\lambda + 1)X_i + \lambda^2)^2 \end{split}$$

Using moment generating function, we can get

$$EX = \lambda$$

$$EX^{2} = \lambda^{2} + \lambda$$

$$EX^{3} = \lambda^{3} + 3\lambda^{2} + \lambda$$

$$EX^{4} = \lambda^{4} + 6\lambda^{3} + 7\lambda^{2} + \lambda$$
So $VarY_{i} = 2\lambda^{2}$

(4)

$$s^{2} - \bar{X} = \frac{1}{n-1} \left(\sum_{i=1}^{n} X_{i}^{2} - n\bar{X}^{2} - (n-1)\bar{X} \right)$$

$$Y_{i} = (X_{i} - \bar{X})^{2} + (\bar{X} - \lambda)^{2} + 2(X_{i} - \bar{X})(\bar{X} - \lambda) - X_{i}$$

$$\sum_{i=1}^{n} Y_{i} = \sum_{i=1}^{n} (X_{i} - \bar{X})^{2} + n(\bar{X} - \lambda)^{2} - \sum_{i=1}^{n} X_{i}$$

$$= (n-1)s^{2} + n(\bar{X} - \lambda)^{2} - n\bar{X}$$
So
$$s^{2} - \bar{X} = \frac{\sum_{i=1}^{n} Y_{i} - n(\bar{X} - \lambda)^{2} + \bar{X}}{n-1}$$
(5)
By CLT, $\sqrt{n\bar{Y}} \Rightarrow_{D} N(0, 2\lambda^{2})$

$$\frac{n}{n-1} \Rightarrow_{P} 1$$
By LLN, $\bar{X} - \lambda \Rightarrow_{P} 0$
By CLT, $\sqrt{n}(\bar{X} - \lambda) \Rightarrow_{D} N(0, \lambda)$

Then Slutsky's Theorem tells that

$$\sqrt{n}(\bar{X} - \lambda)(\bar{X} - \lambda) = \sqrt{n}(\bar{X} - \lambda)^2 \Rightarrow_D 0 * N(0, \lambda) = 0$$
$$\frac{\bar{X}}{\sqrt{n}} \Rightarrow_P 0$$

So
$$(\sqrt{n}\bar{Y} - \sqrt{n}(\bar{X} - \lambda)^2 + \frac{\bar{X}}{\sqrt{n}}) \Rightarrow_D N(0, 2\lambda^2)$$

Because $\frac{n}{n-1} \Rightarrow_P 1$

$$\frac{n}{n-1}(\sqrt{n}\bar{Y}-\sqrt{n}(\bar{X}-\lambda)^2+\frac{\bar{X}}{\sqrt{n}})\Rightarrow_D N(0,2\lambda^2)$$

So
$$\sqrt{n}(s^2 - \bar{X}) \Rightarrow_D N(0, 2\lambda^2)$$

$$\bar{X} \Rightarrow_P \lambda$$

So
$$\sqrt{\frac{n}{2}} \frac{(s^2 - \bar{X})}{\bar{X}} \Rightarrow_D N(0, 1)$$

(6)

For sample from Poisson distribution, when $E(\bar{X}) = E(s^2)$, $P(|\sqrt{\frac{n}{2}} \frac{(s^2 - \bar{X}) - E(s^2 - \bar{X})}{\bar{X}}| \le Z_{1-\alpha/2}) = \alpha$ So if $|\sqrt{\frac{n}{2}} \frac{(s^2 - \bar{X})}{\bar{X}}| \ge Z_{1-\alpha/2}$, we can reject H0.

(7)

```
[33]: model_2 = model2_gen(500)
m2,v2 = np.mean(model_2), np.var(model_2,ddof = 1)
```

```
[34]: print("Model A: ", np.sqrt(500/2)*(v1-m1)/m1) print("Model B: ",np.sqrt(500/2)*(v2-m2)/m2)
```

Model A: 0.2692375458087828 Model B: 33.74859256105128

```
[35]: norm.ppf(0.975)
```

[35]: 1.959963984540054

As we can see, model A 0.27 < 1.96, model B 33.75 > 1.96.

So for model A we can accept H0 at significant level 0.05.

For model B we should reject H0 at significant level 0.05.

(8)

```
[36]: f = [1,4,15,31,39,55,54,49,47,31,16,9,8,4,3]
    data = []
    for i,fre in enumerate(f):
        for j in range(fre):
            data.append(i)
        data = np.array(data)
```

```
[37]: m3,v3 = np.mean(data), np.var(data,ddof = 1)
np.sqrt(len(data)/2)*(v3-m3)/m3
```

[37]: 0.9786745788901424

0.98<1.96, so we can accept H0 at significant level 0.05.

The test doesn't detect any overdispersions.

```
[60]: #e = np.exp(1)
#t = symbols('x', real=True)
#l = symbols('l', real = True)
#diff(e**(l*(e**t-1)), t, 4).subs(t,0)
```