# Exercise2\_cc4515\_cy2550\_ml4404

February 9, 2020

```
[18]: import numpy as np
from scipy.misc import derivative
from scipy.stats import norm
from sympy import diff, symbols
```

## 1 Exercise 2

Chutian Chen cc4515; Congcheng Yan cy2550; Mingrui Liu ml4404

## 1.1 (1)

$$\begin{split} E(\bar{X}^2) &= E(\frac{\sum_{i=1}^n X_i^2 + 2\sum_{i=1}^{n-1} \sum_{j=i+1}^n X_i X_j}{n^2}) \\ &= \frac{1}{n} EX^2 + (1 - \frac{1}{n}) E(X)^2 \\ EX^2 &= VarX + (EX)^2 = \lambda + \lambda^2 \\ EX &= \lambda \\ \text{So } E(\bar{X}^2) &= \lambda^2 + \frac{\lambda}{n} \end{split}$$

#### 1.2(2)

$$Es^{2} = \frac{1}{n-1}E(\sum_{i=1}^{n} X_{i}^{2} - n\bar{X}^{2})$$
$$= \frac{1}{n-1}(n\lambda + n\lambda^{2} - n\lambda^{2} - \lambda) = \lambda$$

#### 1.3 (3)

$$EY_{i} = EX_{i}^{2} - 2\lambda EX_{i} + \lambda^{2} - EX_{i}$$

$$= \lambda^{2} + \lambda - 2\lambda^{2} + \lambda^{2} - \lambda = 0$$

$$VarY_{i} = EY_{i}^{2} - (EY_{i})^{2} = EY_{i}^{2} = E(X_{i}^{2} - (2\lambda + 1)X_{i} + \lambda^{2})^{2}$$

Using moment generating function, we can get

$$EX = \lambda$$

$$EX^2 = \lambda^2 + \lambda$$

$$EX^3 = \lambda^3 + 3\lambda^2 + \lambda$$

$$EX^4 = \lambda^4 + 6\lambda^3 + 7\lambda^2 + \lambda$$

So 
$$VarY_i = 2\lambda^2$$

## 1.4(4)

$$s^{2} - \bar{X} = \frac{1}{n-1} \left( \sum_{i=1}^{n} X_{i}^{2} - n\bar{X}^{2} - (n-1)\bar{X} \right)$$

$$Y_{i} = (X_{i} - \bar{X})^{2} + (\bar{X} - \lambda)^{2} + 2(X_{i} - \bar{X})(\bar{X} - \lambda) - X_{i}$$

$$\sum_{i=1}^{n} Y_{i} = \sum_{i=1}^{n} (X_{i} - \bar{X})^{2} + n(\bar{X} - \lambda)^{2} - \sum_{i=1}^{n} X_{i}$$

$$= (n-1)s^{2} + n(\bar{X} - \lambda)^{2} - n\bar{X}$$
So 
$$s^{2} - \bar{X} = \frac{\sum_{i=1}^{n} Y_{i} - n(\bar{X} - \lambda)^{2} + \bar{X}}{n-1}$$

### 1.5 (5)

By CLT, 
$$\sqrt{n}\bar{Y} \Rightarrow_D N(0, 2\lambda^2)$$

$$\frac{n}{n-1} \Rightarrow_P 1$$

By LLN, 
$$(\bar{X} - \lambda \Rightarrow_P 0)$$

By CLT, 
$$\sqrt{n}(\bar{X} - \lambda) \Rightarrow_D N(0, \lambda)$$

Then Slutsky's Theorem tells that

$$\sqrt{n}(\bar{X} - \lambda)(\bar{X} - \lambda) = \sqrt{n}(\bar{X} - \lambda)^2 \Rightarrow_D 0 * N(0, \lambda) = 0$$

$$\frac{\bar{X}}{\sqrt{n}} \Rightarrow_P 0$$

So 
$$(\sqrt{n}\bar{Y} - \sqrt{n}(\bar{X} - \lambda)^2 + \frac{\bar{X}}{\sqrt{n}}) \Rightarrow_D N(0, 2\lambda^2)$$

Because  $\frac{n}{n-1} \Rightarrow_P 1$ 

$$\frac{n}{n-1}(\sqrt{n}\bar{Y} - \sqrt{n}(\bar{X} - \lambda)^2 + \frac{\bar{X}}{\sqrt{n}}) \Rightarrow_D N(0, 2\lambda^2)$$

So 
$$\sqrt{n}(s^2 - \bar{X}) \Rightarrow_D N(0, 2\lambda^2)$$

$$\bar{X} \Rightarrow_P \lambda$$

So 
$$\sqrt{\frac{n}{2}} \frac{(s^2 - \bar{X})}{\bar{X}} \Rightarrow_D N(0, 1)$$

#### 1.6 (6)

$$P(|\sqrt{\frac{n}{2}} \frac{(s^2 - \bar{X})}{\bar{X}}| \le Z_{1 - \alpha/2}) = \alpha$$

So if  $|\sqrt{\frac{n}{2}} \frac{(s^2 - \bar{X})}{\bar{X}}| \ge Z_{1-\alpha/2}$ , we can reject H0.

## 1.7 (7)

```
for i in range(n):
    a.append(np.random.poisson(np.random.gamma(2.5,2,1),1)[0])
return np.array(a)
```

```
[53]: model_2 = model2_gen(500)
m2,v2 = np.mean(model_2), np.var(model_2)
```

```
[55]: print("Model A: ", np.sqrt(500/2)*(v1-m1)/m1) print("Model B: ",np.sqrt(500/2)*(v2-m2)/m2)
```

Model A: -0.23346323869833238 Model B: 29.64822055088575

```
[57]: norm.ppf(0.975)
```

#### [57]: 1.959963984540054

As we can see, model A 0.23 < 1.96, model B 29.65 > 1.96.

So for model A we can accept H0 at significant level 0.05.

For model B we should reject H0 at significant level 0.05.

#### 1.8 (8)

```
[58]: f = [1,4,15,31,39,55,54,49,47,31,16,9,8,4,3]
    data = []
    for i,fre in enumerate(f):
        for j in range(fre):
            data.append(i)
    data = np.array(data)
```

```
[59]: m3,v3 = np.mean(data), np.var(data)
np.sqrt(len(data)/2)*(v3-m3)/m3
```

#### [59]: 0.9390395410831514

0.93<1.96, so we can accept H0 at significant level 0.05.

The test doesn't detect any overdispersions.

```
[60]: #e = np.exp(1)

#t = symbols('x', real=True)

#l = symbols('l', real = True)

#diff(e**(l*(e**t-1)), t, 4).subs(t,0)
```