Assignment 4

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Problem 1

Consider M = 20,000 hypotheses and their corresponding p-values $\{P_1, \ldots, P_M\}$. Suppose the p-values are from the following mixture distribution:

$$P_i = \begin{cases} \text{Uniform}[0,1], & \text{if } y_i = 0, \\ Beta(0.5,1), & \text{if } y_i = 1, \end{cases}$$

where y_i is a variable indicate whether the *i*-th hypothesis is null $(y_i = 0)$ or nonnull $(y_i = 1)$. Suppose we also know $\Pr(y_i = 0) = 0.9$ and $\Pr(y_i = 1) = 0.1$.

- (a) What is analytic expression of the local false discovery rate, i.e., fdr(P) = Pr(null|P), where P can be any values in (0,1]. Plot function fdr(P) using R or Python based on your analytic calculation.
- (b) Implement Bonferroni correction (BC) and Benjamini-Hochberg Procedure (BHP). Use computer program to check, on average, how many null hypothesis would be rejected by BC (family-wise error rate ≤ 0.1) and BHP (global false discovery rate ≤ 0.1). (Hint: you need to generate p-values $\{P_1,\ldots,P_M\}$ from the mixture distribution and apply BC and BHP to this p-value set. Then repeat this process L times, e.g., L=1,000, to check the average numbers.)
- (c) Do you know how to check where BHP controls global false discovery rate at a desired level, e.g., ≤ 0.1 ? (Hint: Compare the hypothesis rejected by BHP to the ground truth given by the indicator variable $\{y_1, \ldots, y_M\}$.)

Problem 2

Consider M = 20,000 hypotheses and their z-statistics $z_i | \mu_i \sim \mathcal{N}(\mu_i, 1)$,

$$\mu_i = \begin{cases} \mathcal{N}(0.0, \sigma_0^2), & \text{if } y_i = 0, \\ \mathcal{N}(2.5, \sigma_1^2), & \text{if } y_i = 1, \end{cases}$$

where $\sigma_0^2 = \sigma_1^2 = 0.5$ and y_i is a variable indicate whether the *i*-th hypothesis is null $(y_i = 0)$ or nonnull $(y_i = 1)$. Suppose we also know $\Pr(y_i = 0) = 0.95$ and $\Pr(y_i = 1) = 0.05$.

- (a) Derive the analytic solution for the marginal density function f(z) of z.
- (b) Write down the expression for local false discovery rate fdr(z) = Pr(null|z) and plot fdr(z) using R or Python.
- (c) Obtain $\mathbb{E}(\mu_i|z_i)$.

Problem 3

Consider the problem motivating the James-Stein Estimator: $z_i|\mu_i \sim \mathcal{N}(\mu_i, 1)$, i = 1, ..., N., how to obtain a good estimate of $\boldsymbol{\mu} = [\mu_1, ..., \mu_N]^T$ from $\mathbf{z} = [z_1, ..., z_N]^T$?

(a) Assign prior distribution $g(\mu)$: $\mu_i \sim \mathcal{N}(0, \alpha^{-1}), i = 1, ..., N$. Now consider μ as the latent variable and α^{-1} as the model parameter. Derive an EM algorithm for parameter estimation and then use the estimated parameter $\hat{\alpha}$ to obtain an estimate of μ .

(b) Implement the above EM algorithm and compare the estimate with the result obtained from the James-Stein estimator.

Problem 4

Let $\mathbf{y} = [y_1, \dots, y_K]^T$ be a vector of random variables from the multinomial distribution

$$\operatorname{Mult}(\mathbf{y}|\boldsymbol{\mu}, N) = {N \choose y_1 y_2 \dots y_K} \prod_{k=1}^K \mu_k^{y_k},$$

where $0 \le \mu_k \le 1$, $\sum_{k=1}^K \mu_k = 1$ and $\sum_{k=1}^K y_k = N$. Assume that the piror distribution of μ is the Dirichlet distribution

$$Dir(\boldsymbol{\mu}|\boldsymbol{\alpha}) = \frac{\Gamma(\sum_{k=1}^{K} \alpha_k)}{\Gamma(\alpha_1) \cdots \Gamma(\alpha_K)} \prod_{k=1}^{K} \mu_k^{\alpha_k - 1}$$

where $\alpha_k > 0$.

(a) Suppose we have collected a data set $\mathcal{D} = \{\mathbf{y}_1, \dots, \mathbf{y}_n\}$, where

$$\mathbf{y}_i \sim \mathrm{Mult}(\mathbf{y}_i|\boldsymbol{\mu}_i, N), \quad \boldsymbol{\mu}_i \sim \mathrm{Dir}(\boldsymbol{\mu}_i|\boldsymbol{\alpha})$$

Derive an EM algorithm to estimate $\alpha \mathbb{R}^K$, where μ is viewed as the latent variable. Let $\hat{\alpha} = \arg \max_{\alpha} \log p(\mathcal{D}|\alpha) = \arg \max_{\alpha} \sum_{i=1}^{n} \log p(\mathbf{y}_i|\alpha)$ be the MLE, where

$$p(\mathbf{y}_i|\boldsymbol{\alpha}) = \int p(\mathbf{y}_i|\boldsymbol{\mu}_i) p(\boldsymbol{\mu}_i|\boldsymbol{\alpha}) d\boldsymbol{\mu}_i.$$

Can you guarantee that your EM algorithm converges to MLE $\hat{\alpha}$? Explain your reason.

(b) Consider the generative model is as follows:

$$\mathbf{y}_i \sim \text{Mult}(\mathbf{y}_i | \boldsymbol{\mu}_i, N), \quad \boldsymbol{\mu}_i \sim \text{Dir}(\boldsymbol{\mu}_i | \boldsymbol{\alpha}_i), \quad \boldsymbol{\alpha}_i = \exp(\boldsymbol{\beta}^T \mathbf{x}_i).$$

where the parameter α_i in the Dirichlet prior is modulated by side information encoded in $\mathbf{x} \in \mathbb{R}^p$ and $\boldsymbol{\beta}$ is a $p \times K$ matrix. Please derive an EM algorithm to estimate $\boldsymbol{\beta}$.

(c) Suppose the data set $\mathcal{D} = \{\mathbf{y}_i, \mathbf{x}_i\}_{i=1,\dots,n}$ is generated via the probabilistic model in (b), where $\mathbf{y}_i = [y_{i,1}, \dots, y_{i,4}]^T$ is a vector of K = 4 random variables and $\sum_{k=1}^K y_{i,k} = 50$, \mathbf{x}_i is a vector of p = 5 random variables. Apply your algorithm derived in (b) to the given data set \mathcal{D} to estimate $\boldsymbol{\beta} \in \mathbb{R}^{5\times 4}$. The data set is given in data.txt.

Requirement

- You need to submit a report, in which you should clearly describe your method and explain your idea.
 The code should also be included.
- You can use R or Python for coding.
- Your report should be in the **pdf** or **html** format, which is automatically generated by either R markdown or Jupyter notebook.
- The report is due to May, 7, 23:59 pm, 2020 (HK time).