Grow with the flow: a latitudinal cline in physiology is associated with more variable precipitation in Erythranthe cardinalis

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Running Head: Latitudinal cline and climate in Erythranthe

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Abstract

Local adaptation is commonly observed in nature: organisms perform well in their natal environment, but poorly outside it. Correlation between traits and latitude, or latitudinal clines, are among the most common pieces of evidence for local adaptation, but identifying the traits under selection and the selective agents are challenging. Here, we investigated a latitudinal cline in growth and photosynthesis across 16 populations of the perennial herb *Erythranthe cardinalis* (Phrymaceae). Using machine learning methods, we identify interannual variation in precipitation as a likely selective agent: Southern populations from more variable environments had higher photosynthetic rates and grew faster. We hypothesize that selection may favor a more annualized life history – grow now rather than save for next year – in environments where severe droughts occur more often. Thus our study provides insight into how species may adapt if Mediterranean climates become more variable due to climate change.

Introduction

- Local adaptation within species has been documented in numerous species; populations
- 2 generally have higher fitness in their native environment, but perform poorly outside it
- ³ (Schluter, 2000; Leimu and Fischer, 2008; Hereford, 2009). However, the prevalance of local
- 4 adaptation remains difficult to assess because researchers rarely test for local adaptation
- 5 unless there are obvious phenotypic or environmental differences (but see Hereford and
- 6 Winn 2008). When local adaptation occurs, it frequently leads to clines in both phenotypes

and allele frequencies when selection varies over environmental gradients (Huxley, 1938; Endler, 1977; Barton, 1999). Phenotypic differences between populations along a cline most often have a genetic basis and can be studied in a common garden (Turesson, 1922; Clausen et al., 1940; Hiesey et al., 1942). Despite a long history of studying local adaptation and clines, it remains challenging to identify exactly which traits are under selection and which differ for nonadaptive reasons. In particular, the role that physiological differences play in local adaptation is poorly understood, despite the fact that physiology is frequently assumed to explain adaptation to the abiotic environment. A related problem is identifying which of the myriad and often covarying aspects of the environment causes spatially varying selective pressures.

When populations are locally adapted, reaction norms for fitness will cross, such that local genotypes have higher fitness than foreign genotypes and rank orders change across envi-18 ronments (Kawecki and Ebert, 2004). The traits that underlie local adaptation, however, 19 need not mirror this pattern. Populations can have fixed genetic differences conferring trait values that are adaptive at home but neutral or maladaptive away. Alternatively, 21 the ability to plastically respond to a particular environment or the magnitude of response to an environment could be adaptive. We distinguish between these patterns of adaptive trait differences by referring to 'intrinsic' and 'plastic' trait variation, respectively. Both 24 intrinsic and plastic trait variation can be explained by genetic differences and both are 25 involved in adaptation. For example, intrinsic differences in photoperiod responses (Blackman et al., 2011) and developmental rate (Stinchcombe et al., 2004) allow organisms to properly time their life history with the local environment. Conversely, sun and shade 28 plants do not have intrinsically higher or lower rates of carbon assimilation, but rather, 29 genotype-by-environment interactions cause sun plants to assimilate more under high light and shade plants under low light (Givnish, 1988). In plants especially, we know little about 31 the prevalence and adaptive significance of variation in fundamental physiological traits like 32 photosynthesis and their impact on plant performance.

A basic approach to identify candidate traits underlying local adaptation is to find associations between traits and environments. Either intrinsic and/or plastic variation should vary clinally along environmental gradients. Indeed, clines in ecologically important traits 36 are widespread in nature (Endler, 1977) and often adaptive, but in most cases the selective 37 agent is unknown. For example, in *Drosophila* numerous latitudinal clines exist for traits 38 like thermal tolerance (Hoffmann et al., 2002), body size (Coyne and Beecham (1987) and references therein), and life history (Schmidt et al., 2005). Some Drosophila clines have 40 evolved multiple times (Oakeshott et al. (1982); Huey et al. (2000), see also Bradshaw and 41 Holzapfel (2001)) or shifted in response to climate change (Umina et al., 2005), evincing 42 climatic adaptation. Similarly, plant species exhibit latitudinal clines in traits like flowering 43 time (Stinchcombe et al., 2004), cyanogenesis (Koovers and Olsen, 2012), leaf morphology (Hopkins et al., 2008; Stock et al., 2014), and drought response (Kooyers et al., 2015) that likely relate to climatic variation.

Despite the fact that latitudinal clines have been studied for a long time, latitude per se cannot be a selective agent. Latitude may be strongly correlated with one or two key climatic variables, such as temperature, precipitation, or growing degree-days. Latitude 49 may also correlate with the strength of biotic interactions (Schemske et al., 2009) or other nonclimatic aspects of the environment, though as we explain below, we do not yet have 51 compelling data that these are important in our study system. Hence, we focus on whether 52 latitude could be an effective proxy for an underlying climatic driver, in which case we would expect a yet stronger relationship between traits and the key climatic variable(s) driving selection. Alternatively, latitude may be more strongly related to traits than any single climatic variable for at least two reasons. First, latitude may be correlated with several climatic agents of selection that are individually weak, but add up to a strong latitudinal cline. Alternatively, gene flow among neighbouring populations could smooth out local climatic effects, since alleles will experience selection across populations linked by migration (Slatkin, 1978; Paul et al., 2011; Hadfield, 2016). We refer to this as the

'climatic neighborhood'. For example, in mountainous regions average temperature at a given latitude varies widely, but in aggregate, a lower latitude set of populations will experience warmer climate than a higher latitude one. Thus, any particular low latitude population would be warm-adapted, even if it was located in a cooler (e.g. high elevation) site. Because many climatic factors vary latitudinally, and which climatic factors vary latitudinally changes over the earth's surface (e.g. coastal vs. continental), dissecting the evolution of latitudinal clines across many species will help identify generalities, such as whether thermal tolerance maxima or seasonal timing is more important (Bradshaw and Holzapfel, 2008), or whether local versus regional climate shape selective pressures.

In this study, we investigated two major questions: 1) whether intrinsic or plastic physiological trait variation corresponds with latitude; and 2) what climatic factor(s) could plausibly be responsible for latitudinal clines. Within question 2, we tested three hypotheses outlined in the previous paragraph: latitudinal clines are explained by a single dominant climatic 73 factor, multiple climatic factors, or the climatic neighborhood experienced by nearby population connected through gene flow. These hypotheses are not mutually exclusive since, for example, single or multiple factors in a climatic neighborhood may lead to latitudinal clines. We focused on climate because climate often determines and where species are found and also can exert strong selection on populations within species [CITATION?]. There is 78 also a compelling need to know how populations are (or are not) locally adapted to climate so as to predict how theyâĂŹll respond to climate change [CITATION?]. We did not test for local adaptation to other environmental characteristicss, such as edaphic or biotic (e.g. competitors, pollinators) factors, because we did not have compelling observations from which to make testable hypotheses about the latter in our focal species (described) below).

We examined these questions in *Erythranthe cardinalis* (formerly *Mimulus cardinalis* [Neson 2014]) because linking physiological traits to potentially complex patterns of local adaptation requires integrating multiple lines of evidence from comparative, experimental,

and genomic studies under both lab and field conditions. Many classic and contemporary studies of local adaptation use Mimulus sensu lato species because of its natural history, 89 easy propagation, and genetic/genomic resources (Clausen et al., 1940; Hiesey et al., 1971; 90 Bradshaw and Schemske, 2003; Wu et al., 2008; Lowry and Willis, 2010; Wright et al., 91 2013). Yet, there is a conspicuous deficiency of links between local adaptation and physiological mechanisms (Angert (2006); Angert et al. (2008); Wu et al. (2010), but see Wright et al. (2013) on mechanisms of copper tolerance). We measured genetic and genotypeby-environment variation in response to temperature and drought among 16 populations distributed over 10.7° of latitude. We found a latitudinal cline of intrinsic variation in photosynthesis and growth, but little evidence for variation in plasticity. Interannual variation 97 in precipitation and temperature are associated with this axis of variation, suggesting that 98 climatic variance rather than mean may be an important driver of local adaptation in E. cardinalis. The climatic neighborhoods around populations explained trait variation 100 better than local climate, indicating that latitudinal clines may be common because lati-101 tude integrates effects of selection on populations connected through gene flow. We place 102 these findings in the context of life history theory and consider future directions in the 103 Discussion. 104

105 Material and Methods

106 Population Selection

We used 16 populations from throughout the range of E. cardinalis (Table 1). These populations were intentionally chosen to span much of the climatic range of the species based on all known occurrences (see below). Seeds were collected in the field from mature, undehisced fruit left open for 2-4 weeks to dry, then stored at room temperature. We used seeds from 154 families, 4–12 (mean = 9.6, median = 12) families per population.

Table 1: Geographic region, latitude, longitude, and elevation (mas = meters above seal level) of 16 focal populations used in this study.

Name	Region	Latitude	Longtiude	Elevation (mas)
Hauser Creek	South Margin	32.657	-116.532	799
Cottonwood Creek	South Margin	32.609	-116.7	267
Sweetwater River	South Margin	32.9	-116.585	1180
Grade Road Palomar	South Margin	33.314	-116.871	1577
Whitewater Canyon	Transverse	33.994	-116.665	705
Mill Creek	Transverse	34.077	-116.873	2050
West Fork Mojave River	Transverse	34.284	-117.378	1120
North Fork Middle Tule River	South Sierras	36.201	-118.651	1314
Paradise Creek	South Sierras	36.518	-118.759	926
Redwood Creek	South Sierras	36.691	-118.91	1727
Wawona	Central Sierras	37.541	-119.649	1224
Rainbow Creek	Central Sierras	37.819	-120.007	876
Middle Yuba River	North Sierras	39.397	-121.082	455
Little Jamison Creek	North Sierras	39.743	-120.704	1603
Deep Creek	North Coast	41.668	-123.11	707
Rock Creek	North Margin	43.374	-122.957	326

Plant propagation

On 14 April, 2014, 3-5 seeds per family were sown directly on sand (Quikrete Play Sand, 113 Georgia, USA) watered to field capacity in RLC4 Ray Leach cone-tainers placed in RL98 114 98-well trays (Stuewe & Sons, Inc., Oregon, USA). We used pure sand because E. cardinalis 115 typically grows in sandy, riparian soils (A. Angert, pers. obs.). Two jumbo-sized cotton 116 balls at the bottom of cone-tainers prevented sand from washing out. Cone-tainers sat in 117 medium-sized flow trays (FLOWTMD, Stuewe & Sons, Inc., Oregon, USA) to continuously 118 bottom-water plants during germination in greenhouses at the University British Columbia 119 campus in Vancouver, Canada (49°15' N, 123°15' W). Misters thoroughly wetted the top of 120 the sand every two hours during the day. Most seeds germinated between 1 and 2 weeks, 121 but we allowed 3 weeks before transferring seedlings to growth chambers. We recorded 122 germination daily between one to two weeks after sowing, and every 2-3 days thereafter. 123

On 5 May (21 days after sowing), we transferred seedlings to one of two growth chambers (Conviron, Manitoba, Canada). We thinned seedlings to one plant per cone-tainer, leaving the center-most plant. 702 of 768 (91.4%) had plants that could be used in the experiment. We allowed one week at constant, non stressful conditions (day: 20°C, night: 16°C) for plants to acclimate to growth chambers before starting treatments. The initial size of seedlings, measured as the length of the first true leaves, did not differ between populations, families, or treatments (Table S1).

131 Temperature and drought treatments

We imposed four treatments, a fully-factorial cross of two temperature levels and two 132 watering levels. The temperature levels closely simulated an average growing season at the 133 thermal extremes of the species range, which we designate as Hot and Cool treatments. 134 Watering levels contrasted a perennial and seasonal stream, which we refer to as Well-135 watered and Drought treatments. A detailed description of treatments is provided in the 136 Supplemental Materials and Methods and summarized in Fig 1. Because growth chambers 137 cannot be subdivided, one chamber was assigned to the Hot treatment level and another 138 to the Cool treatment level. Within each chamber, there were two Well-watered blocks 139 and two Drought blocks. The photosynthetically active radiation in both chambers was 140 approximately 400 μ mol quanta m⁻² s⁻¹. The growth chambers did not control humidity, 141 but because of watering and high plant transpiration rates, the relative humidity was quite 142 high in both temperature levels (data not shown). Lower humidity would have made the 143 drought more severe, but low soil moisture is stressful in and of itself. The total number of 144 plants in each treatment was: $n_{\text{cool,dry}} = 169$; $n_{\text{cool,ww}} = 174$; $n_{\text{hot,dry}} = 176$; $n_{\text{hot,ww}} = 183$. 145 Each population had 8-12 individuals per treatment level (mean = 11, median = 11). 146

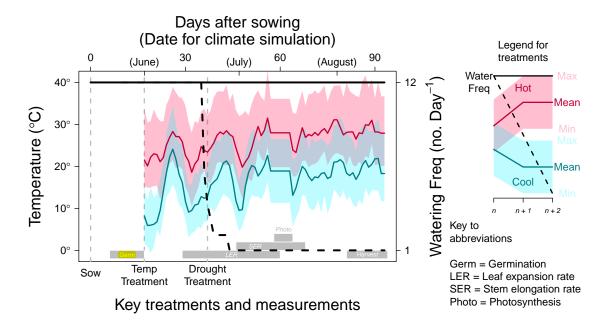


Figure 1: Overview of experimental treatments and timing of key trait measurements. All plants germinated within 21 days of sowing. At that time, we began temperature treatments (left axis), simulating a typical June-August weather pattern at Hot (red) and Cool (blue) sites. The bold lines track the average daily temperatures. Within each day, there was a maximum daytime temperature (top of translucent polygons) and minimum nighttime temperature (bottom of translucent polygons). The drought treatment commenced later by ramping down the frequency of bottom-watering episodes (dashed black line; right axis), while watering frequency was maintained in the control treatment (solid black line). Grey boxes on the bottom of the plot outline the period of key measurements described in the Material and Methods.

147 Trait measurements

We measured five traits in response to temperature and watering treatments (Table 2).

Days to germination We tested for population variation in germination rate, measured as Days to Germination, using a lognormal survival model fit using the survreg function in the R package survival version 2.38 (Therneau, 2015). We treated Population as a fixed effect and Family as random effect using a Γ frailty function. Statistical significance of the

Table 2: Key traits measured in this study.

Trait	Units
Days to germination	day
Leaf expansion rate	$\mathrm{mm}\ \mathrm{day}^{-1}$
Stem elongation rate	${\rm cm~day^{-1}}$
Photosynthetic rate	μ mol CO ₂ m ⁻² s ⁻¹
Mortality	probability of death

Population effect was determined using analysis of deviance. Note that, unlike other traits discussed below, we did not include Block, Treatment, or Population × Treatment interactions because during germination plants had not been placed into blocks and treatments had not yet been applied.

Growth rate: leaf expansion and stem elongation We measured growth rate dur-157 ing two phases: leaf expansion and stem elongation. Growth measurements were taken 158 during the early vegetative stage. We censused leaf length twice per week shortly after 159 the emergence of true leaves from 12 May – 12 June (28–59 days after sowing), resulting 160 in 10 measurements. We ceased measuring leaf length once it appeared to asymptote and 161 growth shifted to stem elongation. We also censused plant height on 7 occasions (twice 162 per week) between 29 May and 20 June (45 to 67 days after sowing) until plants began to initiate floral buds. Thus all growth measurements occured during the vegetative, pre-164 reproductive phase. Both leaf expansion and stem elongation were modelled separately 165 as second-order polynomials. We used empirical Bayes' estimates of growth for each indi-166 vidual plant from linear mixed-effects models fit with the R package lme4 version 1.1-12 167 (Bates et al., 2015). 168

Photosynthesis During the week of 10 to 16 June (57 to 63 days after sowing), we measured daytime photosynthetic rate on a subset of 329 plants evenly spread between

treatments and families within populations. The youngest, fully-expanded leaf acclimated for 3 minutes to reach steady state in a 6-cm² chamber of a LI-COR 6400XT Portable Photosynthesis System (LI-COR Biosciences, Lincoln, Nebraska). We made all measurements at ambient light (400 μ mol m⁻² s⁻¹ of photosynthetically active radiation), atmospheric CO₂ (400 ppm), temperature, and moderate relative humidity. During this period, we suspended normal day-to-day temperature fluctuations and set daytime temperatures to the average for that period (Cool: 26.5°; Hot: 36.1°) so that all plants within a temperature level could be measured under the same conditions.

Mortality We assayed mortality during twice-weekly growth measurements. We analyzed the probability of surviving until the end of the experiment as a function of population, treatment, and their interactions using a Generalized Linear Mixed Model (GLMM) assuming binomially distributed errors. We included Family and Block as random effects. We assessed significance of fixed effects using Type-II Analysis of Deviance with Wald χ^2 tests in the R package car (Fox and Weisberg, 2011).

185 Intrinsic variation and plasticity

For all traits (Table 2) except germination (see above), we tested for Population, Treat-186 ment (Temperature, Water, and Temperature × Water), and Population × Treatment 187 interactions (Population \times Temperature, Population \times Water, and Population \times Temperature, 188 ature × Water). We interpreted significant Population effects to indicate intrinsic variation 189 and Population × Treatment interactions to indicate variation in plasticity. As mentioned 190 above, we used survival and GLMM models for germination rate and mortality, respec-191 tively. For all other traits, we used mixed model ANOVAs with Family and Block included 192 as random factors. We fit models using restricted maximum likelihood in lmer, a function 193 in the R package lme4 (Bates et al., 2015). We determined significant fixed effect terms us-194

ing a step-wise backward elimination procedure implemented with the step function in the
R package lmerTest version 2.0-32 (Kuznetsova et al., 2016). This package uses Satterthwaite's approximation to calculate denominator degrees of freedom for F-tests. We also
included days to germination as a covariate in growth analyses. To ensure that Population
and Treatment effects were specific to a particular growth phase, we included germination
day as a covariate in leaf expansion and stem elongation analyses.

201 Principal components of germination, growth, and photosynthesis

For each single-trait model above, we extracted the Population coefficient (factoring out
Treatment and other effects). The multivariate distribution of these coefficients was then
summarized using principal components analysis. The first principal component of these
traits (TraitPC1) loaded positively with germination, growth, and photostynthetic rate,
therefore we define this as a phenotypic axis delineating fast to slow growth.

107 Identifying putative selective agents

Latitudinal clines are common, but it is often difficult to ascribe this variation to a par-208 ticular selective agent. To reiterate, we tested three non-mutually exclusive hypotheses 209 about how such latitudinal clines emerge: 1) one or two climatic variable explains latitudi-210 nal trait variation; 2) latitude is a proxy for multiple climatic factors that together shape 211 trait variation; and 3) latitude integrates selection in a broader climatic neighborhood. We 212 found that a population's position along TraitPC1 correlated strongly with the latitude of 213 origin (see Results) and next used Random Forest regression (Liaw and Wiener, 2002) to 214 identify putative climatic factors underlying trait-latitude associations in E. cardinalis. We reasoned that if we identified a single climatic factor that explained more trait variation 216 than latitude, then this would suggest that factor is a key selective agent underlying the 217 latitudinal cline (Hypothesis 1). On the other hand, if multiple climatic factors together

together have imposed selection for the latitudinal cline (Hypothesis 2). We hereafter refer 220 to factors identified in this analysis as 'Climate-TraitPC1' variables. To test Hypothesis 221 3 about climatic neighborhoods driving selection, we directly competed local with neigh-222 borhood climate. We used the immediate collection location for local climate. For climate 223 neighborhoods, we sampled climate at 1000 random points (at 90-m resolution) within a 224 62-km radius buffer around the collection and took the average. We chose this buffer size 225 because neutral genetic differentiation increases slowly with geographic distance, indicating significant gene flow between nearby populations (Paul et al., In review). Significant spatial autocorrelation persisted for approximately 62 km. Since E. cardinalis is found exclusively in riparian areas, we only selected points along streams using the National Hydrogeoraphy 229 Dataset (United States Geological Survey, 2015). Climatic means and variances (see below) 230 were weighted by their climatic suitability as determined using a multimodel ensemble av-231 erage of ecological niche models (Angert et al., 2016). In addition to competing local and 232 neighborhood climate, we compared the univariate correlation between local and neigh-233 borhood climate with TraitPC1 and Latitude using paired t-tests. We adjusted degrees 234 of freedom to account for the fact that many climatic factors are highly correlated and 235 not independent. Specifically, we calculated the effective number of independent climatic 236 factors (M_{eff}) using the formula $M_{\text{eff}} = 1 + (M-1)(1 - \text{Var}(\lambda)/M)$ (Chevrud, 2001), where 237 M is the original number of climatic factors and λ are the eigenvalues of the correlation 238 matrix of all climatic factors. 239 To help eliminate potentially spurious correlations between TraitPC1 and climate, we tested

are necessary to explain trait variation, then this would suggest that many climatic factors

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To help eliminate potentially spurious correlations between TraitPC1 and climate, we tested for overlap between climatic variables that best predict latitude of all *E. cardinalis* occurrence records (see detail below), not just the 16 focal populations. We refer to these climatic factors as 'Climate-Latitude' variables. The logic is that climatic factors associated with both TraitPC1 and latitude for all populations are more likely to be important selective agents than climatic factors that happen to correlate with TraitPC1 but do not covary

with latitude throughout the E. cardinalis range. Therefore, we did not consider Climate-TraitPC1 variables to be candidate selective agents unless the same or very similar variable was found in the Climate-Latitude analysis.

We selected Climate-Latitude and Climate-TraitPC1 variables independently using Vari-249 able Selection Using Random Forest (VSURF) algorithm in the R package VSURF version 250 1.0.3 (Genuer et al., 2016). Random Forest regression is useful for cases like ours when 251 the number of potential predictors is similar to or greater than the number of observations 252 ('high p, low n' problem). VSURF is a multistip algorithm that progressively retains or 253 eliminates variables based on their importance over regression trees in the forest. Variable 254 importance is defined as the average amount a climate variable reduces mean-squared er-255 ror in the predicted response (TraitPC1 or Latitude), compared to a randomly permuted 256 dataset, across all trees in the random forest (see Genuer et al. [2015] for further detail). 257 Hence, VSURF automatically eliminates unimportant and redundant variables based on 258 the data without having to arbitrarily choose among colinear climate variables before the 259 analysis. We kept only variables selected for prediction, the most stringent criterion. A 260 visual overview of how we selected climatic variables is depicted in Fig 2. 261

For Climate-Latitude analyses, we compiled a representative set of 356 recent (since 2000) 262 known E. cardinalis occurrences. These occurrences were thinned by 50% to correct for 263 uneven sampling from a comprehensive set of herbarium records and an exhaustive field survey in 2010-11 (Angert et al., 2016). For both Climate-TraitPC1 analyses (16 focal populations) and Climate-Latitude (many populations), we used a 90-m digital elevation 266 model from HydroSHEDS (Lehner et al., 2006) to extract elevation. Monthly interpolated 267 climate layers were calculated using ClimateWNA version 5.30 (Wang et al., 2012), which 268 accurately downscales climate data specifically for the rugged topography of western North 269 America. For each occurrence, we calculated bioclimatic variables using the biovars func-270 tion in the R package dismo version 1.1-1 (Hijmans et al., 2016). We included 24 climatic 271 factors, 9 from ClimateWNA and 15 bioclimatic variables (Table S2). The bioclimatic

variables included all permutations of two climatic factors, temperature and precipitation, and six temporal scales (annual average, coldest quarter, warmest quarter, wettest quarter, 274 driest quarter, or seasonality) as well as mean diurnal range, isothermality, annual temper-275 ature range. For each variable, we calculated both a 30-year normal by averaging annual 276 values between 1981 and 2010 and 30-year coefficient of variation, a standardized metric 277 of interannual climatic variation. Temperatures were converted to Kelvin to be on a ratio 278 scale appropriate for calculating the coefficient of variation (CV). In total, the VSURF al-279 gorithm selected among 96 climate variables: 24 climatic factors × 2 types (30-year average 280 and CV) \times 2 spatial scales (local and neighborhood). 281

2 Results

A coordinated latitudinal cline in germination, growth, and photosynthesis

There are strong genetically-based trait differences in time to germination, growth, and 285 photosynthetic rate among populations of E. cardinalis, as evidenced by large and signif-286 icant population effects for these traits (Table 3). A single principal component captured 287 71.6 % of the trait variation among populations, defining an axis of variation from fast to 288 slow growth. A population's position along this axis strongly covaried with its latitude of 289 origin; southern populations grew faster than northern populations (Fig 3). As we explain 290 below, intrinsic differences between populations in terms of plant function (photosynthesis) 201 and performance (growth) contrasted with the low amount of variation in plasticity. There 292 were similar latitudinal clines for individual traits underlying PC1 (Figures S1 to S4).

Table 3: Summary of Population, Treatment, and Population \times Treatment effects. We used different statistical modeling for the diverse traits assayed – glmer: generalized linear mixed model using the R package **Ime4** (Bates et al., 2015); Imer: linear mixed model using the R package **Ime4** (Bates et al., 2015); survreg: survival regression using the R package **survival** (Therneau, 2015). Note that temperature and water treatments were imposed after germination, hence are not applicable to this trait. Complete analysis of variance/deviance tables for each trait are available in the Supporting Information. Key to statistical significance: *P < 0.05; ** *P < 0.01; *** *P < 0.001

Trait Statistical model	Germination survreg	Leaf expansion lmer	Stem elongation lmer	Photosynthesis lmer	Mortality glmer
Population	***	***	***	***	
Temperature	NA	***	***	**	***
Water	NA	*			***
$Pop \times Temp$	NA			*	
$Pop \times Water$	NA	*			
$Temp \times Water$	NA				***
$\underline{\text{Pop} \times \text{Temp} \times \text{Water}}$	NA				

4 Little evidence for variation in plasticity

Genotype × environment (G×E) interactions are also a common signature of local adaptation. We found little evidence of $G \times E$ in E. cardinalis. There were only two statisti-296 cally significant Population × Treatment interactions (Table 3, Fig. S5), but these were 297 not strong compared to Population and Temperature effects. Otherwise, populations re-298 sponded similarly to treatments: faster growth in the hot treatment, slower growth in the 299 dry treatment, and high mortality in the hot, dry treatment (Table 3). Note that inter-300 actions were calculated after factoring out intrinsic trait differences, necessarily reducing 301 statistical power to detect significant interactions relative to main effects. However, the fact that the Population and Temperature effects were often highly significant $(P \ll 0.001)$ in most cases) suggests that statistical power alone cannot explain why we failed to detect Population X Treatment interactions. Complete ANOVA tables are available in the Supporting Information (Tables S3 to S6)

Neighborhood climatic variability best explains latitudinal cline

Interannual variation in climate averaged over 62-km radius climatic neighborhoods around 308 populations correlated most strongly with trait variation and latitude of E. cardinalis 309 occurrences (Fig. 4, Table S7). All 16 Climate-Latitude and 3 Climate-TraitPC1 vari-310 ables were neighborhood rather than local variables (Fig. 4). In fact, neighborhood cli-311 mate almost always correlated better with TraitPC1 and Latitude than local climate 312 (Fig. S6). On average, neighborhood Climate-TraitPC1 correlation coefficients were 0.16 313 higher than correlations with local-scale climate variables (paired t-test, t = 7.87, d.f. = 314 33.6, $P = 3.94 \times 10^{-9}$). Likewise, neighborhood Climate-Latitude correlation coefficients 315 were 0.13 higher than those for local-scale climate (paired t-test, t = 6.71, d.f. = 36.8, 316 $P = 7.22 \times 10^{-8}$). 317

The most important climatic variables related to the interannual variation rather than 318 average climate. Among Climate-Latitude and Climate-TraitPC1 variables, neighborhood 310 climatic variability over 30 years (1981–2010) in either winter precipitation (bio16 $_{\sigma}$) and/or 320 temperature (bio11 $_{\sigma}$) are the strongest candidates driving the latitudinal cline in E. cardi-321 nalis (see Table S2 for a key to climate variable abbreviations). Note that the coefficient 322 of variation of a climatic factor is subscripted with σ whereas the mean is subscripted with 323 μ . More specifically, greater winter precipitation variability and lower winter temperature 324 variability are associated with Southern latitudes and higher TraitPC1 values (Fig. 5A,B). 325 Neighborhood interannual variation in winter precipitation (bio16 $_{\sigma}$) was the most impor-326 tant Climate-Latitude variable (Fig. 4A). However, neighborhood bio 16_{σ} did not overlap 327 with Climate-TraitPC1 variables (Fig. 4B). We nevertheless consider it a plausible can-328 didate for two reasons. First, neighborhood bio16 $_{\sigma}$ correlated strongly with TraitPC1 329 (Fig. 5A). Second, one of the most important Climate-TraitPC1 variables (neighborhood 330 $bio15_{\sigma}$; Fig. 5B,C) is very similar to $bio16_{\sigma}$. In Mediterranean climates like California, most 331 precipitation occurs in the wettest quarter (winter), so years with low winter precipitation 332

also have low precipitation seasonality. Hence, highly variable year-to-year winter precipitation at lower latitude (Fig. 5D) is closely associated with large swings in precipitation seasonality (Fig. 5C).

Interannual variation in temperature of the coldest quarter (neighborhood bio11 $_{\sigma}$) is an-336 other plausible candidate because it was only variable in both Climate-Latitude and Climate-337 TraitPC1 analyses (Fig. 4). Neighborhood bio11 $_{\sigma}$ explained more variation in TraitPC1 338 than latitude (latitude $r^2 = 0.55$ vs. bio 11_{σ} $r^2 = 0.6$; Fig. S7), whereas neighborhood 339 bio16 $_{\sigma}$ did slightly worse (bio16 $_{\sigma}$ $r^2=0.49$). Models using bio15 $_{\sigma}$ or bio11 $_{\sigma}$ to predict 340 TraitPC1 also had significantly lower Akaike Information Criteria (AIC) than the latitude 341 model (AIC of different models – bio15 $_{\sigma}$: 48.5; bio11 $_{\sigma}$: 52.4; latitude: 54.5). The best two-342 factor model including both neighborhood bio 15_{σ} and bio 11_{σ} did not significantly improve 343 explanatory power ($r^2 = 0.71$, AIC= 49.2). In summary, either variation in precipita-344 tion or temperature seasonality may be important selective agents, but there is no strong 345 evidence that they are both important. The most important Climate-TraitPC1 variable, neighborhood variation in mean diurnal range ($bio2_{\sigma}$; Fig. 4B) did not have any obvious 347 similarity to Climate-Latitude variables. Given the large number of potential associations, 348 we therefore think this may be a spuriously strong relationship and do not consider it 349 further. 350

Discussion

We found evidence for one of two common signatures of local adaptation in the perennial
herb *Erythranthe cardinalis*. Latitudinal clines in germination rate, photosynthesis, and
growth suggest adaptive differentiation in important physiological traits of the species.
However, we found little evidence that populations respond differently to temperature or
drought. Due to low replication within families, we did not have power to assess within
population genotype-by-environment interactions, which may be present. As we discuss

below, low variation in plasticity among populations result may indicate that some dimensions of the fundamental abiotic niche are relatively conserved. Finally, our results suggest that neighborhood-scale climate and interannual variation are more important selective agents than local averages. In the paragraphs that follow, we tie these results into the broader threads of evolutionary theory that might help explain why intrinsic variation in physiology changes clinally, whereas plastic responses to temperature and drought are relatively similar. One caveat to bear in mind is that we are limited by the size of the climate grid ($\approx 90 \text{ m}^2$) and therefore unable to detect very fine-scale local adaptation.

Evolutionary theory indicates that the shape of fitness tradeoffs, demography, and gene 366 flow can constrain adaptation (Levins, 1968; Ronce and Kirkpatrick, 2001; Lenormand, 367 2002) and hence the type of variation maintained within species. Specifically, adaptive 368 variation cannot be maintained by spatially varying selection if tradeoffs are too strong, 369 demography is strongly asymmetric, and/or maladaptive gene flow is too high. To clarify, 370 strong tradeoffs can prevent local adaptation in variable environments because selection favors habitat specialists that track a specific habitat regardless of its frequency in the environment (Levins, 1968). For example, a riparian specialist may experience similar 373 selection in rivers of highfall regions and deserts, even though the habitat is much rarer in the latter. In E. cardinalis we found substantial genetically based variation among populations along a phenotypic axis from fast to slow growth that varied over a large 376 spatial scale (Fig. 3). If this variation is adaptive, it suggests one of several possibilities 377 to investigate in the future: fitness tradeoff between low versus high latitude environments 378 is not too strong nor swamped by demographic asymmetry or maladaptive gene flow. 379 That is, alleles favoured at one latitude are not strongly selected against when they flow 380 to another population, allowing locally adaptive genetic variation to be maintained by 381 spatially heterogenous selection. We also know from previous work that population size 382 does not vary strongly with latitude. Gene flow appears to be high, but attenuates at 383 broad spatial scales, especially between Southern (< 35°N) and Northern portions of the

range (Paul et al., In review).

Nevertheless, local gene flow from similar environments may shape how selection varies 386 with latitude. Theory predicts that populations will not be perfectly adapted to their 387 immediate habitat when there is gene flow from surrounding populations with different 388 optima (Lenormand, 2002). With spatial heterogeneity and gene flow, traits will not covary 389 perfectly with the local optimum (Slatkin, 1978; Paul et al., 2011; Hadfield, 2016), but 390 should instead better match the average environment experienced by nearby populations 391 connected through gene flow, which we refer to as the climatic neighborhood. Gene flow 392 and spatial heterogeneity may therefore be important in maintaining genetic variation 393 (Yeaman and Jarvis, 2006). As this hypothesis predicts, climatic neighborhoods (62-km 394 buffer around populations) correlated with traits and latitude of occurrences better than 395 local climate (Fig. 4). We interpret this as suggestive evidence that gene flow between 396 neighboring E. cardinalis populations shapes selection – populations are locally adapted to 397 prevailing climate in their neighborhood, but perhaps not perfectly adapted to their local 398 climate. This may not greatly constrain local adaptation because local and neighborhood 399 climate values at the resolution of ClimateWNA (90 m^2) were generally similar in E. 400 cardinalis populations (Fig. S6). Therefore, we would predict in reciprocal transplants 401 that populations whose local climate is farther from their neighborhood average would be 402 less well adapted than those close to their neighborhood average. 403

It is reasonable to predict that southern populations, which appear to experience more frequent drought years (see below), might have physiological adaptation to survive and grow in drier soil. We found no evidence for this type of drought tolerance; all populations responded to drought and temperature similarly (Table 3). Plants grew faster in the Hot treatment, but there was little effect of drought on growth. Rather, the effects of drought took longer to materialize but resulted in high mortality, especially in the Hot treatment. However, there was no differential mortality among populations in this treatment. Although our results indicate that this axis of the species niche may be constrained,

plants have multiple ways to resist drought through both tolerance and escape (Ludlow, 1989; Kooyers, 2015). Next, we consider why drought tolerance may less important in local adaptation than a form of escape for this species.

We hypothesize that tolerance to dry soil may be constrained by a combination of strong 415 fitness tradeoffs, demographic asymmetry, and gene flow. Soil moisture in riparian habitats 416 where E. cardinalis live are highly heterogeneous at very small spatial scales (several me-417 ters). Plants in the stream never have to tolerate drought whereas plants only a few meters 418 away may experience extreme drought since there is little direct precipitation during the 419 growing season in Mediterranean climates of western North America. We hypothesize alle-420 les that confer greater drought tolerance may be quite costly in well-watered soils, and vice 421 versa, leading to strong fitness tradeoffs. Such tradeoffs would promote specialization to 422 one soil moisture or another, thereby inhibiting the evolution of broad environmental tol-423 erance within a population. Demography and gene flow may reinforce niche conservatism. 424 A new mutant with increased drought tolerance that could survive at the resource-poor 425 margin of a population would likely be demographically overwhelmed by the larger census 426 populations that can be maintained in higher-resource environments. Infrequent wet years 427 may also produce most seeds, so selection is weighted towards alleles that have high fitness 428 in the wet environment, even if dry years are more frequent (Templeton and Levin, 1979; 429 Brown and Venable, 1986). Finally, gene flow, which is generally high among E. cardinalis 430 populations within the same ecoregion (Paul et al., In review), will thwart local adapta-431 tion and reinforce specialization. Thus, the spatial grain of the environment, demographic 432 asymmetry, and gene flow may conspire to constrain local adaptation along this environ-433 mental axis. Consistent with this hypothesis, recent record-setting droughts have caused 434 the decline or even local extinction of some natural populations of E. cardinalis (Sheth and 435 Angert, 2017). 436

In sum, these results indicate that differences in physiology and growth, but not responses to temperature and drought, mediate local adaptation to climate in *E. cardinalis*. Next,

we would like to understand why variation in these particular traits may be adaptive. We 439 argue that temporally more variable environments, as experienced by southern populations, select for a more 'annualized' life-history strategy, a form of drought escape. Demographic 441 observations in natural populations of E. cardinalis reveal that southern populations tend 442 to flower earlier at a smaller size, while northern populations invest more in vegetative 443 growth (Sheth and Angert, 2017). The association between position along the 'fast-slow' 444 continuum and associated traits in E. cardinalis is similar to interspecific relationships 445 between growth, functional traits, and life history (Adler et al., 2014; Salguero-Gómez 446 et al., 2016). However, we cannot exclude unexplored factors (e.g. edaphic conditions, 447 competitors, pollinators, etc.) which may drive the latitudinal cline. 448

Greater investment in aboveground growth, as opposed to belowground storage for future 449 seasons, may be favoured in climates with more frequent drought years, but maladaptive in 450 climates with more consistent precipitation. This is a form of drought escape in that plants 451 are investing more reproduction in the present to avoid possible drought in subsequent 452 years. Suppose plants that grow quickly and allocate new resources to continued growth 453 rather than storage have higher fitness over a single growing season. However, by not 454 allocating resources to storage, these fast-growing plants begin future seasons at a deficit. 455 Therefore, in a stable environment where winter survivorship is assured in most years, 456 failure to store resources may reduce lifetime fitness. But the environment is not constant. 457 For perennial herbs in Mediterranean climates, a dry winter (rainy season) can kill the 458 rhizomes (underground stems that store nutrients for future growth) before emergence or aboveground stems before flowering. If drought years occur frequently enough, selection 460 may favour the fast-growing strategy because there is no advantage to storage if drought 461 kills plants before flowering. Considering life-history strategy as a continuum from no 462 storage (annual) to lots of storage (perennial), we hypothesize that the optimal allocation to 463 aboveground growth is more 'annualized' in southern climates that have greater interannual 464 variation in precipitation. This scenario differs from classic drought escape syndromes in which plants speed up development early in the season before the onset of drought.

The hypothesis that greater precipitation variability selects for an annualized life history 467 is tentative, but consistent with theory and data from other species. Life history theory 468 shows that less variable environments are one factor that favours the evolution of perenniality (Stearns, 1976; Iwasa and Cohen, 1989; Friedman and Rubin, 2015). Populations of 470 the perennial Plantago asiatica show a similar latitudinal cline in growth and allocation to 471 storage (Sawada et al., 1994), though these authors attribute the cline to variation in grow-472 ing season length. There are also life history clines in the closely related species E. guttata, 473 but the underlying traits and climatic drivers are quite different. Annual E. quttata flower 474 sooner and produce fewer stolons in response to climates with shorter seasons and more 475 intense summer drought (Lowry and Willis, 2010; Friedman et al., 2015; Kooyers et al., 476 2015). In contrast, there are no truly annual (monocarpic and semelparous) populations of 477 E. cardinalis. Rather, our hypothesis states that climatic variability selects on quantitative 478 variation in allocation to growth versus storage. This hypothesis makes several indepen-479 dent, testable predictions. The allocation tradeoff predicts that northern populations will 480 provision more assimilate to rhizomes compared with southern populations. If southern 481 populations are indeed more 'annualized' because more frequent droughts cause mortality, 482 then we predict that species distribution models using recent climate would best predict 483 occurrences in the south, whereas longer term climate would be a better predictor in the 484 north. Finally, we predict that southern populations would show greater variation in the 485 size of recruits and higher maximum population growth rates. Exploratory analysis also 486 pointed to variation in winter temperature as a potential selective agent. Specifically, win-487 ter temperature is more variable in the north, which may lead to less predictability in the 488 start of the growing season or late frost. This could affect selection on phenology, which we did not study here, but it is not obviously connected to the latitudinal cline in physiological 490 traits. 491

492 In summary, we found evidence for a coordinated latitudinal cline in germination rate,

photosynthesis, and growth, suggesting local adaptation. We therefore predict to find 493 different optima for these traits in different climates. We did not find evidence that the 494 relative performance of populations shifts with temperature or watering regime, suggesting 495 relatively little variation in plasticity. Exploratory analysis implicate that more variable 496 precipitation regimes at lower latitude drive much of the latitudinal cline, though other 497 climatic factors could also contribute. Interestingly, the climatic neighborhood may shape 498 selective pressures more than local climate. In the future, we will use field experiments to 490 test whether greater variation in precipitation selects for faster growth and that selection 500 on temperature/drought responses does not vary among populations. By doing so, we aim 501 to understand why certain physiological and developmental mechanisms, but not others, 502 contribute to local adaptation. 503

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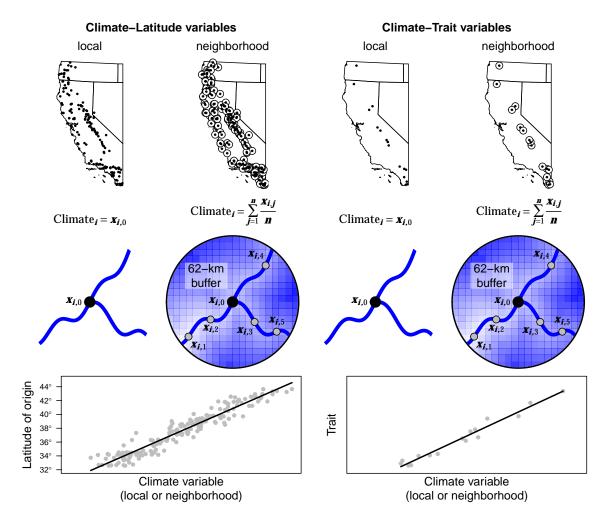


Figure 2: Overview of method for identifying putative climatic selective agents underlying latitudinal cline. We looked for climate variables that explained both the latitude of 356 E. cardinalis occurrences ('Climate-Latitude variables') and with traits ('Climate-Trait variables'). For Climate-Latitude variables we extracted climate data from recent occurrences located throughout California and Oregon, USA (shown in map). For Climate-Trait variables, we extracted climatic data for the 16 focal populations. For both analyses, we extracted local and neighborhood climate. Local climate refers to climate only from where a population was collected $(x_{i,0})$. Neighborhood climate was calculated as the average over 1000 points in a 62-km radius climatic neighborhood $(x_{i,1}, x_{i,2}, \ldots)$, but only along stream habitats as E. cardinalis is riparian. We identified climatic factors that most strongly predicted latitude of occurrences (Climate-Latitude variables) and traits (Climate-Trait variables), as shown for hypothetical data in plots at the bottom of the figure.

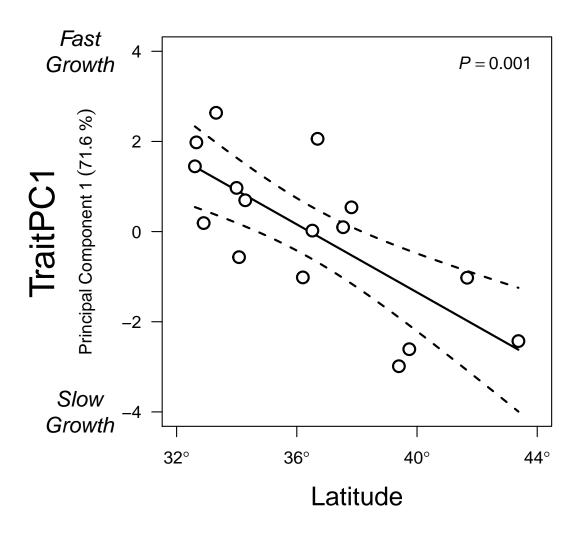


Figure 3: Trait variation, from fast to slow growth, is closely associated with latitude. Each point is a population's latitude of origin (x-axis) and position along the slow to fast growth axis (y-axis), defined as Principal Component 1 of four traits (see Material and Methods). The line and 95% confidence intervals were estimated using linear regression.

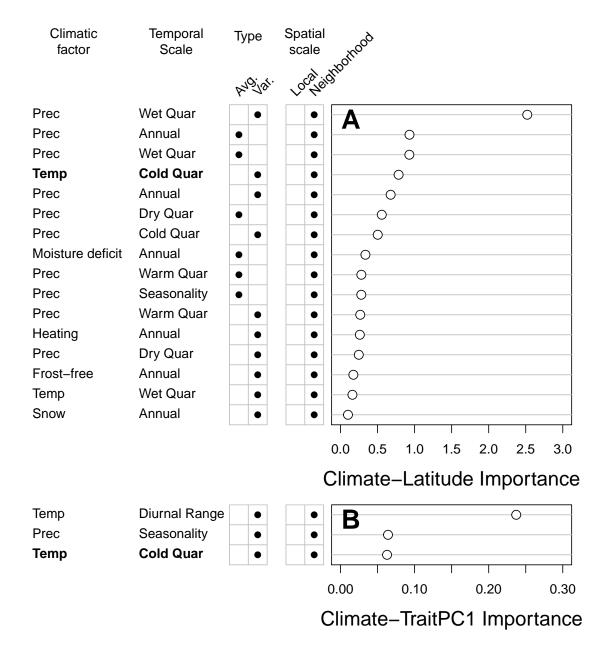


Figure 4: Climatic variation integrated over climatic neighborhood is closely correlated with latitude of *E. cardinalis* and trait variation. A. Using Random Forest regression, we identified 16 climatic variables significantly (high importance) associated with latitude of *E. cardinalis* occurrences. B. Only one of of the most important Climate-Latitude variables (in bold) was among the most important Climate-TraitPC1 variables. Variable importance is defined as the average amount a climate variable reduces mean-squared error in the predicted response (TraitPC1 or Latitude), compared to a randomly permuted dataset, across all trees in the random forest (see Genuer et al. [2015] for further detail). Note that the Importance values in A and B are not comparable because the dependent variables (Latitude and Trait PC1, respectively) are on different scales. Climatic variables (left of A; right of B) are defined by four qualities: Climatic factor – Temperature (Temp), Precipitation (Prec), Heating degreedays (Heating), Snow (precipitation as snow); Temporal scale – Annual, Coldest quarter (Cold Quar), Warmest Quarter (Warm Quar), Wettest quarter (Wet Quar), Driest Quarter (Dry Quar), or Seasonality; Type – 30-year average (Avg.) or coefficient of variation (Var.); Spatial scale – local or 62-km radius climatic neighborhood.

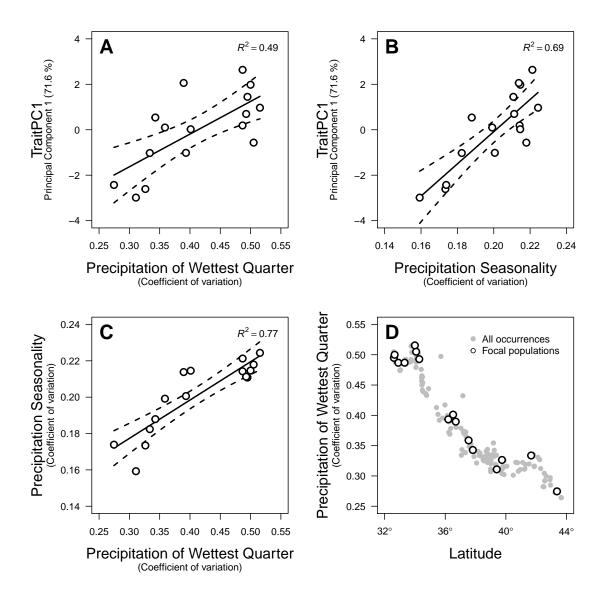


Figure 5: Variation in precipitation is correlated with TraitPC1 and latitude. A. Greater values of TraitPC1 are associated with greater interannual variation in precipitation of the wettest quarter. This was the most important Climate-Latitude variable, but not among the most important Climate-TraitPC1 variables. B. However, a closely related parameter, interannaul variation in precipitation seasonality, was among the most important Climate-TraitPC1 variables. C. Across focal populations, variation in precipitation of the wettest quarter and seasonality are closely correlated. D. Southern populations of *E. cardinalis* experience much greater interannual variationi in precipitation. In all panels, we report climatic neighborhood values (see Material and Methods). Regression lines, 95% confidence intervals, and coefficients of determination (\mathbb{R}^2) were calculated using linear regression.

- Supporting Information
- Supporting Tables

Table S1: Initial size of seedlings did not vary among Populations, Families, or Treatments. We used a censored Gaussian model of initial size at the outset of the experiment (longest leaf length of the first true leaves). The model was censored because we could not accurately measure leaves less than 0.25 mm with digital callipers (217 of 702, 30.9%, were too small). We fit models using a Bayesian MCMC method implemented using the MCMCglmm function with default priors in the R package **MCMCglmm** version 2.17 (Hadfield, 2010). We estimated the posterior distribution from 1000 samples of an MCMC chain run for 10^5 steps after a 10^4 step burn-in. We used step-wise backward elimination procedure to find the best-supported model according to Deviance Information Criterion (DIC).

Model	Random	DIC
Population + Water + Temperature +	Family	1638
Population:Water +		
Population:Temperature +		
Water:Temperature +		
Population:Water:Temperature		
Population + Water + Temperature +	Family	1605.2
Population:Water +		
Population:Temperature +		
Water:Temperature		
Population + Water + Temperature +	Family	1603.4
Population:Water +		
Population:Temperature		
Population + Water + Temperature +	Family	1577.5
Population:Water +		
Water:Temperature		
Population + Water + Temperature +	Family	1579.9
Population:Temperature +		
Water:Temperature		
Population + Water + Temperature +	Family	1577.3
Population:Water	- "	4550.5
Population + Water + Temperature +	Family	1550.5
Water:Temperature	- "	4540.0
Population + Water + Temperature	Family	1549.3
Population + Water	Family	1541.7
Population + Temperature	Family	1546.8
Water + Temperature	Family	1551.1
Population	Family	1541.9
Water	Family	1543.9
-	Family	1541.7
-	-	1538.3

Table S2: Climatic variables used

Abbreviation	Climate variable
$\overline{\mathrm{DD}}_{-0}$	degree-days below 0°C(chilling degree-days)
DD5	degree-days above 5°C(growing degree-days)
DD_18	degree-days below 18°C(heating degree-days)
DD18	degree-days above 18°C(cooling degree-days)
NFFD	number of frost-free days
PAS	precipitation as snow (mm) between August in previous year and July
	in current
Eref	Hargreaves reference evaporation (mm)
CMD	Hargreaves climatic moisture deficit (mm)
RH	mean annual relative humidity
bio1	annual mean temperature
bio2	mean diurnal range (mean of monthly (max temp - min temp))
bio3	isothermality (bio2/bio7) (* 100)
bio4	temperature seasonality (standard deviation *100)
bio5	max temperature of warmest month
bio6	min temperature of coldest month
bio7	temperature annual range (bio5-bio6)
bio8	mean temperature of wettest quarter
bio9	mean temperature of driest quarter
bio10	mean temperature of warmest quarter
bio11	mean temperature of coldest quarter
bio12	annual precipitation
bio15	precipitation seasonality (coefficient of variation)
bio16	precipitation of wettest quarter
bio17	precipitation of driest quarter
bio18	precipitation of warmest quarter
bio19	precipitation of coldest quarter

Table S3: Analysis of varianace (ANOVA) table on leaf expansion rate (LER) using **ImerTest** (Kuznetsova et al., 2016). Family and Block were included as random effects. Abbreviations: SS = sum of squares; MS = mean sum of squares (SS / df1); df1 = numerator degrees of freedom; df2 = denominator degrees of freedom.

	SS	MS	df1	df2	F-value	P-value
Day to Germination	12.12	12.12	1	637	35.21	4.9 ×10 ⁻⁹
Population	22.22	1.48	15	118	4.3	2.5×10^{-6}
Temperature	80.42	80.42	1	5	233.61	2.6×10^{-5}
Water	4.1	4.1	1	5	11.92	0.019
Temperature × Water	0.03	0.03	1	4	0.07	0.801
Population × Temperature	2.76	0.18	15	547	0.53	0.925
Population \times Water	9.66	0.64	15	562	1.87	0.024
	4.11	0.27	15	530	0.78	0.700

Table S4: Analysis of varianace (ANOVA) table on stem elongation rate (SER) using **ImerTest** (Kuznetsova et al., 2016). Family and Block were included as random effects. Abbreviations: SS = SUM = SUM

	SS	MS	df1	df2	F-value	P-value
Day to Germination	3.6	3.6	1	662	21.1	5.1×10^{-6}
Population	12	8.0	15	113	4.7	5.8×10^{-7}
Temperature	12.4	12.4	1	6	72.8	1.5×10^{-4}
Water	0.6	0.6	1	5	3.7	0.113
Temperature \times Water	0.9	0.9	1	4	5.2	0.093
Population \times Temperature	3.6	0.2	15	549	1.4	0.126
Population \times Water	2.8	0.2	15	536	1.1	0.330
Population \times Temperature \times Water	1.5	0.1	15	518	0.6	0.874

Table S5: Analysis of varianace (ANOVA) table on photosynthetic rate using **ImerTest** (Kuznetsova et al., 2016). Family and Block were included as random effects. Abbreviations: SS = sum of squares; MS = mean sum of squares (SS / df1); df1 = numerator degrees of freedom; df2 = denominator degrees of freedom.

	SS	MS	df1	df2	F-value	P-value
Population	347.7	23.2	15	78	3.02	7.5×10^{-4}
Temperature	134.1	134.1	1	6	17.46	6.4×10^{-3}
Water	51	51	1	4	6.64	0.066
Temperature × Water	0.7	0.7	1	3	0.09	0.781
Population × Temperature	218.6	14.6	15	263	1.9	0.024
Population × Water	87.7	5.8	15	233	0.76	0.724
Population \times Temperature \times Water	91.4	6.1	15	208	0.79	0.686

Table S6: Analysis of deviance table on the probability of mortality by the end of the experiment using Type-II Wald χ^2 tests in the R package **car** (Fox and Weisberg, 2011). Family and Block were included as random effects. Abbreviations: df = degrees of freedom

	χ^2	df	P-value
Population	32	31	0.419
Temperature	31.8	6	1.8×10^{-5}
Water	69.2	12	4.6×10^{-10}
Temperature \times Water	20.7	1	5.3×10^{-6}
Population \times Temperature	5.6	15	0.985
Population \times Water	8.6	15	0.897
Population \times Temperature \times Water	0.2	15	1.000

Table S7: Important climatic variables predicting latitude of E. cardinalis populations ('Climate-Latitude') and the first principal component of traits measured in a common garden ('Climate-TraitPC1'). Local climatic variables were measured from the exact location of collection; neighborhood climatic variables were averaged from a 62-km neighborhood around population (see Material and Methods). Importance and significance were determined using the variable selection using random forests (VSURF) algorithm (see Material and Methods). Climatic variables are described in Table S2. μ signifies the mean of the climate variables from 1981–2010; σ indicates coefficient of variation among years.

Climate-Latitude variables	Climate-TraitPC1 variables
Precipitation of wettest quarter $(\sigma, \text{neighborhood})$ Annual precipitation $(\mu, \text{neighborhood})$ Precipitation of wettest quarter $(\mu, \text{neighborhood})$ Mean temperature of coldest quarter $(\sigma, \text{neighborhood})$ Annual precipitation $(\sigma, \text{neighborhood})$ Precipitation of driest quarter $(\mu, \text{neighborhood})$ Precipitation of coldest quarter $(\sigma, \text{neighborhood})$ Precipitation of warmest quarter $(\mu, \text{neighborhood})$ Precipitation of warmest quarter $(\mu, \text{neighborhood})$ Precipitation seasonality $(\mu, \text{neighborhood})$ Precipitation of warmest quarter $(\sigma, \text{neighborhood})$ Precipitation of driest quarter $(\sigma, \text{neighborhood})$ Precipitation of driest quarter $(\sigma, \text{neighborhood})$ Number of frost-free days $(\sigma, \text{neighborhood})$ Mean temperature of wettest quarter $(\sigma, \text{neighborhood})$ Precipitation as snow $(\sigma, \text{neighborhood})$	Mean diurnal range (σ , neighborhood) Precipitation seasonality (σ , neighborhood) Mean temperature of coldest quarter (σ , neighborhood)

682 Supporting Figures

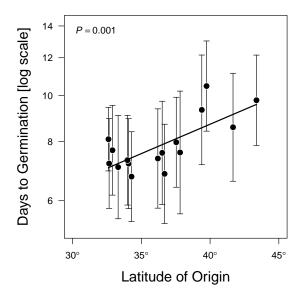


Figure S1: Southern populations germinate faster. Each point is a population of *E. cardinalis* showing its latitude of origin (x-axis) and model-predicted days to germination in days under growth chamber conditions (see Material and Methods). Bars around each point are 95% confidence intervals. Predicted time to germination and confidence intervals are based on survival regression (see Materials and Materials). The line is the linear regression of log(model-predicated days to germination) \sim latitude. The P-value of the regression is given in the upper left corner.

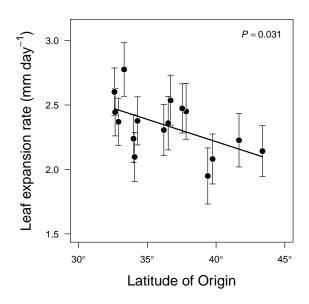


Figure S2: Southern populations grow faster. Each point is a population of E. cardinalis showing its latitude of origin (x-axis) and model-predicted leaf expansion rate during the rosette phase. Bars around each point are 95% confidence intervals. Predicted leaf expansion rate based least-square mean estimates and confidence intervals were calculated from linear mixed-effects models (see Materials and Materials). The line is the linear regression of model-predicated leaf expansion rate \sim latitude. The P-value of the regression is given in the upper right corner.

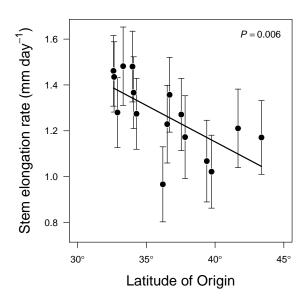


Figure S3: Southern populations grow faster. Each point is a population of E. cardinalis showing its latitude of origin (x-axis) and model-predicted stem elongation rate. Bars around each point are 95% confidence intervals. Predicted stem elongation rate based least-square mean estimates and confidence intervals were calculated from linear mixed-effects models (see Materials and Materials). The line is the linear regression of model-predicated stem elongation rate \sim latitude. The P-value of the regression is given in the upper right corner.

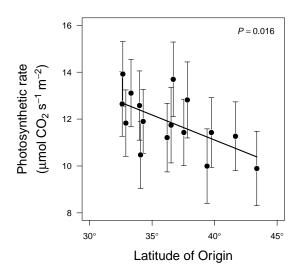


Figure S4: Southern populations photosynthesize faster. Each point is a population of E. cardinalis showing its latitude of origin (x-axis) and model-predicted instantaneous photosynthetic rate. Bars around each point are 95% confidence intervals. Predicted photosynthetic rates based least-square mean estimates and confidence intervals were calculated from linear mixed-effects models (see Materials and Materials). The line is the linear regression of model-predicated photosynthetic rate \sim latitude. The P-value of the regression is given in the upper right corner.

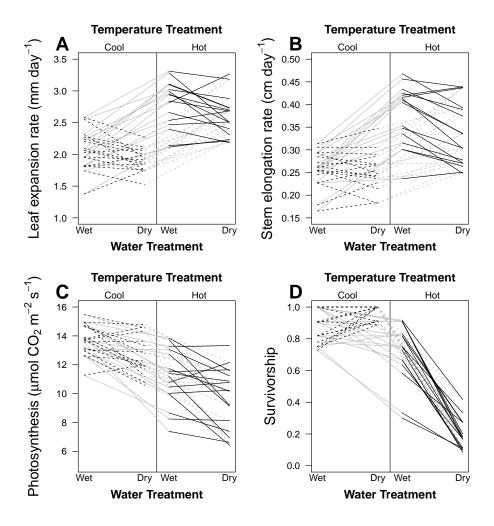


Figure S5: Reaction norms signify little Population \times Treatment interactions. For all panels, black lines represent population-level reaction norms from Wet to Dry in the Cool temperature treatment (dashed black lines) and Hot temperature treatment (solid black lines); gray lines represent reaction norms from Cool to Hot in the Wet treatment (solid gray lines) and Dry treatment (dashed gray lines). The responses shown are (A) leaf expansion rate, (B) stem elongation rate, (C) photosynthesis, and (D) survivorship (= 1 - mortality).



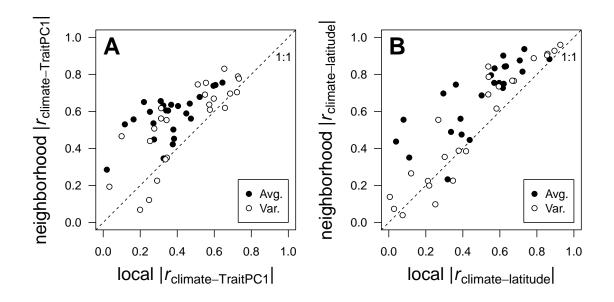


Figure S6: Neighborhood climate predicts TraitPC1 ('Climate-trait', panel A) and Latitude of occurences ('Climate-latitude', panel B) better than local climate. Each point is the absolute value of the Pearson correlation coefficient (|r|) between TraitPC1 (A) or latitude (B) for 24 climatic factors, for which we used both the 30-year mean (closed circles) and coefficient of variation (open circles). Most points lie above the 1:1 line, indicating stronger correlations with neighborhood compared to local climate. Neighborhood climate was integrated over a 62-km radius around focal populations (see text for further detail).

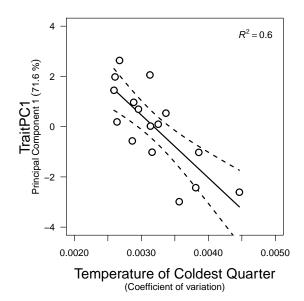


Figure S7: Trait variation, from fast to slow growth, is closely associated with neighborhood variation in temperature of the coldest quarter ($bio11_{\sigma}$) Each point is a population coefficient of variation in bio11 averaged over a 62-km climatic neighborhood (x-axis) and position along the slow to fast growth axis (y-axis), defined as Principal Component 1 of four traits (see Material and Methods). The line and 95% confidence intervals were estimated using linear regression.

683 Supporting Material and Methods

684 Temperature treatments

We simulated typical growing season (June 1 - August 15) air temperatures at the two most 685 thermally divergent focal sites in our study, Whitewater Canyon (WWC, Hot) and Little 686 Jameson (LIJ, Cool). We downloaded daily interpolated mean, minimum, and maximum 687 air temperature from 13 years (2000-2012) at both sites from ClimateWNA (Wang et al., 688 2012). This range was chosen because seeds used in the experiment were collected around 689 2012, thus their presence in that location at that time suggests that populations were able 690 to persist there for at least some years before collection. Monthly temperatures from Cli-691 mateWNA are highly correlated with the air temperature recorded from data loggers in 692 the field at these sites (A. Angert, unpub. data). Hence, the ClimateWNA temperature 693 profiles are similar to actual thermal regimes experienced by E. cardinalis in nature. We 694 simulated realistic temperature regimes by calculating the mean temperature trend from 695 June to August using LOESS (Cleveland et al., 1992). The residuals were highly autocor-696 related at both sites (warmer than average days are typically followed by more warm days) 697 and there was strong correlation (r = 0.65) between sites (warm days in WWC were also 698 warm in LIJ). The 'VARselect' function in the vars package for R (Pfaff, 2008) indicated 699 that a lag two Vector Autoregression (VAR(2)) model best captured the within-site auto-700 correlation as well as between-site correlation in residuals. We fit and simulated from the 701 VAR(2) model using the package dse (Gilbert, 2014) in R. Simulated data closely resem-702 bled the autocorrelation and between-site correlation of the actual data. From simulated 703 mean temperature, we next selected minimum and maximum daily temperatures. Mean, 704 min, and max temperature were highly correlated at both sites. We chose min and max 705 temperatures using site-specific fitted linear models between mean, max, and min tem-706 perature, with additional variation given by normally distributed random deviates with 707 variance equal to the residual variance of the linear models. For each day, the nighttime (22:00 - 6:00) chamber temperature was set to the simulated minimum temperature. During the middle of the day, temperature was set to the simulated maximum temperature, with a variable period of transition between min and max so that the average temperature was equal the simulated mean temperature.

3 Watering treatments

For watering treatments, we simulated two extreme types of streams where E. cardinalis 714 grows. In the well-watered treatment, we simulated a large stream that never goes dry 715 during the summer growing season. In the drought treatment, we simulated a small stream 716 that has ample flow at the beginning of the season due to rain and snow melt, but gradually 717 dries down through the summer. In both treatments, plants were bottom-watered using 718 water chilled to 7.5°C. Plants in the well-watered treatment were fully saturated every two 719 hours during the day. Watering in the drought treatment gradually declined from every 720 two hours to every day between May 20 (36 days after sowing) and 10 June (57 days after 721 sowing). Simultaneously, the amount of bottom-watering per flood decreased, such that only the bottom of the cone-tainers were wetted by the end of the experiment.