Selection on early survival does not explain germination rate clines in *Mimulus cardinalis*

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Abstract

- Premise Many traits covary with environmental gradients to form phenotypic clines. While
- local adaptation to the environment can generate phenotypic clines, other nonadaptive processes
- may also. If local adaptation causes phenotypic clines, then the direction of genotypic selection
- on traits should shift from one end of the cline to the other. Traditionally genotypic selection on
- non-Gaussian traits like germination rate have been hampered because it is challenging to
- measure their genetic variance.
- Methods Here we used quantitative genetics and reciprocal transplants to test whether a
- ₂₀ previously discovered cline in germination rate showed additional signatures of adaptation in the
- scarlet monkeyflower (*Mimulus cardinalis*). We measured genotypic and population level

- 22 covariation between germination rate and early survival, a component of fitness. We developed a
- novel discrete log-normal model to estimate genetic variance in germination rate.
- Results Contrary to our adaptive hypothesis, we found no evidence that genetic variation in
- ₂₅ germination rate contributed to variation in early survival. Across populations, southern
- populations in both gardens germinated earlier and survived more.
- ²⁷ Conclusions Southern populations have higher early survival but this is not caused by faster
- germination. This pattern is consistent with nonadaptive forces driving the phenotypic cline in
- 29 germination rate, but future work will need to assess whether there is selection at other life
- ₃₀ stages. This statistical framework should help expand quantitative genetic analyses for other
- waiting-time traits.
- Keywords: cline; genotypic selection; germination; Mimulus; reciprocal transplant

Introduction

Populations within species differ in ecologically important traits that often result from adaptation to different local environments (Turesson, 1922; Clausen et al., 1948). However, it is still rare that we understand the traits and selective agents underlying local adaptation (Wadgymar, Lowry, et al., 2017). The classic signature of local adaptation is crossing reaction norms for fitness measured in a reciprocal transplant experiment (Kawecki and Ebert, 2004; Johnson et al., 2021), such that populations have higher relative fitness in their local environment but lower relative fitness in a foreign environment. For example, annual inland ecotypes of the yellow Monkeyflower Mimulus guttatus (Phrymaceae) have higher relative fitness in seasonally dry sites compared to perennial 41 coastal ecotypes that are adapted to year-round water availability. The relative fitness of coastal and inland ecotypes is reversed on the coast because of life history and other genetic differences between ecotypes (Lowry and Willis, 2010). The prevalence of local adaptation implies that selection varies spatially, populations have heritable variation in fitness, and selection is stronger than migration or drift. Exceptional cases of local adaptation over short temporal or spatial scales (e.g. Wright et al., 2013; Grant and Grant, 2014; Richardson et al., 2014; Lescak et al., 2015; Barrett et al., 2019; DiVittorio et al., 2020) likely contribute to an availability bias that leads many to perceive that local adaptation is ubiquitous. However, systematic meta-analyses reveal that local adaptation is often weak or nonexistent (Leimu and Fischer, 2008; Hereford, 2009; Brady et al., 2019) and may be getting weaker because of anthropogenic climate change (Bontrager et al., 2020). Populations may not be locally adapted if differential selection is weak relative to migration or drift or they lack heritable variation in traits under selection. Alternatively, recent anthropogenic climate change may have erased the signature of local adaptation by causing some foreign populations to have higher fitness than local populations (Bontrager et al., 2020). For example, in the alpine plant Boechera strica, simulated climate change (mimicking early snow melt) favors populations from lower elevations and snow addition restores the relative fitness advantage of high-elevation populations (Anderson and Wadgymar, 2020). Understanding what causes trait variation among populations within species, whether adaptive or not, will help biologists better predict how populations will respond to environmental change.

Phenotypic clines are commonly interpreted as evidence for local adaptation, but nonadaptive

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processes also generate similar patterns (Endler, 1977). A phenotypic cline is a correlation
   between a heritable character and the environment or a proxy for environmental gradients such
   as latitude or elevation (Huxley, 1938). Clines may be caused by either abrupt or gradual
   environmental change. For example, heavy metal contamination of soil around mines is an
   abrupt change from the surrounding pasture soil. A similarly abrupt change in the heavy metal
   tolerance of Anthoxanthum odoratum occurs across the mine boundary, but other traits under
   weaker or correlated selection vary gradually over space (Antonovics and Bradshaw, 1970).
   Other clines track gradual environmental change. Latitudinal clines in the size of Drosophila
   subobscura flies have evolved multiple times and most likely track gradual variation in
   temperature and/or phenology (Huey et al., 2000). But clines are not necessarily adaptive. When
   there is genetic isolation by distance (Wright, 1943), it is possible for phenotypic differences to
   correlate with genetic differences throughout the species range (Vasemägi, 2006). In practice,
   nonadaptive clines are difficult to demonstrate because it is hard to reject adaptive explanations.
   The proportion of cyanogenesis varies clinally in white clover (Trifolium repens) but field
   experiments have not observed spatially varying selection that could explain this cline (Wright
   et al., 2021). Hence, spatial variation in cyanogenesis may be nonadaptive or the design of the
   experiment may have missed selection occurring during germination and early development.
   When species expand their range, such as during biological invasions, nonadaptive clines can
   readily evolve because of multiple introductions and serial founder events (Colautti and Lau,
   2015). Even parallel clines, which are usually considered strong evidence of natural selection, can
   result from nonadaptive processes when there is epistasis and consistent spatial variation in the
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   strength of genetic drift (Santangelo et al., 2018). Evolutionary ecological experiments such as
   reciprocal transplants are needed to provide additional evidence to distinguish between adaptive
   and nonadaptive hypotheses for clines (Wadgymar, Daws, et al., 2017).
   There are several experimental methods for measuring the agents and strength of selection to
   different environments under natural conditions (Wadgymar, Lowry, et al., 2017), but here we
   focus on differential selection. One key prediction is that the direction of selection and position
   of the phenotypic optimum will differ among environments. If genetic variation in a trait causes
   fitness to vary, then genotypic selection analyses (Rausher, 1992) should find that genotypes with
   high trait expression are favored in one environment but disfavored in another. For example,
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selection favors higher seed dormancy in southern environments of Arabidopsis thaliana and lower dormancy in northern environments, presumably to optimally time seedling establishment 93 with the onset of favorable summer conditions (Postma and Agren, 2016). Populations can also simultaneously differ as a result of both adaptive and non-adaptive processes. When locally adapted populations differ in traits that do not themselves confer local adaptation, we expect population-level covariation between divergent traits and fitness but do not expect within-population variation to correlate with fitness. Hence, comparing the direction of among-genotype and among-population selection can help distinguish between phenotypic clines that are important for local adaptation and those that may have evolved nonadaptively. In this study, we used a reciprocal transplant to test whether a cline in Mimulus cardinalis germination rate could be locally adaptive in terms of early survival between seedling establishment and flowering. Because seeds of M. cardinalis do not require cold or warm 103 stratification to break dormancy, they will germinate whenever conditions are sufficiently warm 104 and moist. Within this germination niche, there is genetic variation for germination time that covaries with latitude and other traits associated with a fast-slow growth continuum (Muir and Angert, 2017; Sheth and Angert, 2018; Nelson et al., 2021). In greenhouse conditions, seeds from 107 the southern end of the range (San Diego County, CA) germinate in 6–7 days; seeds from the 108 northern end of the range (Lane County, OR) germinate in 9–10 days. The growing season is 100 typically longer in the southern end of the range because of reduced or nonexistent snowpack, 110 but for the same reason southern populations may be more susceptible to late-season drought 11 (Sheth and Angert, 2018). The adaptive hypothesis is that faster germination increases early survival in the southern end of the range and slower germination increases early survival at the 113 northern end. We would predict either stabilizing genotypic selection for different optimal 114 germination rates or opposing signs of directional genotypic selection. If the cline is maintained by nonadaptive forces then germination rate should be uncorrelated with genotypic fitness (even if it is correlated with population mean fitness) or the direction of selection should be the same 117 in both parts of the range. A limitation of this study is that we cannot measure selection on 118 seedling establishment since that occurred in the greenhouse (see Materials and Methods) or 110 later survival and fecundity (these fitness components will be analyzed in a follow-up paper).

Germination and early survival are potentially important for local adaptation but often 121 overlooked compared to later life stages. Failure to measure seedling establishment and early 122 survival can bias estimates of selection on traits that only express at later life stages, such as flower size (Mojica and Kelly, 2010). In the herb Arabidopsis thaliana, quantitative differences in germination timing contributed to local adaptation and are maintained by a balance of selection on seedling establishment, survival, and fecundity, but the optimal timing varies in northern and 126 southern environments (Postma and Ågren, 2016, 2018). A challenge of measuring genotypic 127 selection on germination rate is that quantitative genetic theory and analysis are built around Gaussian distributions (Lynch and Walsh, 1998; but see Villemereuil et al., 2016), but the distribution of times until germination are most likely not Gaussian. Although it is convenient to 130 express germination as a rate (events per time), this and other "waiting time" traits are measured 131 as the time until an event, the inverse of germination rate. Throughout this paper we distinguish 132 between germination rate as the trait of interest which we estimate by measuring the time to 133 germination. The time until germination (or another event) can be modeled as a waiting time, which is bounded at zero and usually right skewed. Waiting times are distributed exponentially 135 if the probability of germination is constant through time, but this unlikely to be true. Rather, the probability of germination almost certainly increases through time once conditions are 137 conducive. Hence, more complex waiting time distributions should occur in nature. A second challenge is that waiting time is usually modeled as a continuous trait, but is observed at discrete intervals (e.g. once per day). If the average time to germination is sufficiently long relative to the 140 interval between observations, then a continuous-time approximation may be sufficient. 141 However, when time to germination is several days, as with M. cardinalis, then accurately modeling the trait may need to account for the discrete observation process. Here we show that non-Gaussian distributions and measurement process can be incorporated into and improve 144 analysis of time to germination. This approach should be useful, with some customization, to 145 many other traits and biological systems. Hence, investigating spatially variable selection on 146 germination time is important for expanding our knowledge of local adaptation and phenotypic 147 clines for poorly studied life stages with challenging statistical properties.

Materials and Methods

550 Study system

Mimulus cardinalis (=Erythranthe cardinalis Lowry et al., 2019) is a perennial forb native to the 15 Western US (California and Oregon). It is self-compatible, predominantly outcrossing, and 152 hummingbird pollinated. We used five focal populations from throughout the geographic range of *M. cardinalis* (Table 1; Appendix S1, Fig. S1; see the Supplementary Data with this article). 154 These five focal populations are a subset of those we used previously to identify a cline in 155 germination and other traits (Muir and Angert, 2017). Seeds were collected in the field from 156 mature, undehisced fruits left in open coin envelopes for 2–4 weeks to dry, then stored at room 157 temperature. To minimize maternal effects, we grew a large number of field-collected seeds in the greenhouse and generated seed families for this experiment by hand-pollinating individuals 159 using the breeding design described in the next section. 160

161 Genetic variance and heritability

To estimate selection among genotypes and populations, we first need to quantify genetic variance in germination rate and early survival. We estimated genetic variance and heritability of germination rate and winter survival using a quantitative genetic breeding design. For each 164 population, we crossed 15 parental individuals using a partial diallel design (Lynch and Walsh, 165 1998) with three dams per sire and three sires per dam for a total of 45 full-sib families per 166 population. The number of parents is low for quantitative genetic analyses, but we were restricted garden space and by the availability of wild-collected seeds in some source 168 populations. We opted for fewer parental individuals and more populations so that we can better 169 understand range-wide patterns of local adaptation. Individual plants of this hermaphroditic 170 species were used as both sires (pollen parent) and dams (ovule parent). We did not make crosses 171 between populations. One family from the Little Jamison Creek population did not produce enough seeds, resulting in 44 usable families. In total, we used $5 \times 45 - 1 = 224$ families. 173 We estimated genetic variances and heritabilities for each of the five populations (Table 1). The 174 total phenotypic variance within a population V_P can be partitioned into genetic (V_G) , maternal 175 (V_M) , and environmental (V_E) components:

$$V_P = V_G + V_M + V_E$$

Genetic and environmental factors contribute to maternal affects (Lynch and Walsh, 1998).

Although we attempted to minimize maternal effects by growing a refresher generation in a

common greenhouse environment and M. cardinalis seeds are tiny and not well provisioned, seed

traits may be particularly sensitive to maternal effects. We did not have statistical power to

estimate nonadditive genetic contributions, therefore we estimated broad-sense (H^2) heritability,

the fraction of phenotypic variance contributed by all sources of (including dominance and

epistasis) genetic variance:

$$H^2 = \frac{V_G}{V_P}$$

When separate individuals are used as dams and sires, additive genetic variance of the base population is often estimated from the variance among sires: $V_G=4~\sigma_{\rm sire}$ (Lynch and Walsh, 1998). Since our breeding design used hermaphroditic individuals as both dams and sires, we estimated parental breeding values from their contributions as both sires and dams. Hence, we estimated V_G from the variance among parents:

$$V_G = 4 \sigma_{\rm parent}$$

Using $\sigma_{\rm parent}$ rather than $\sigma_{\rm sire}$ is more powerful because it uses all of the data. We estimated σ_{parent} and other parameters using a Bayesian mixed effects model. We describe the general approach here and provide detail specific to germination rate and survival below. We fit the model in Stan version 2.29.2 (Stan Development Team, 2022) using cmdstanr version 0.5.2 192 (Gabry and Češnovar, 2022). Stan calculates the posterior distribution of the model using the 193 Hamiltonian Monte Carlo algorithm, which is similar to the more widely used Markov Chain 194 Monte Carlo, but is faster and more efficient at sampling for many applications (Monnahan et al., 2017). We used weakly informative normally distributed priors for parameters that affect the 196 trait mean (intercepts and coefficients) and half-Cauchy priors for variance parameters. Weakly 197 informative priors are strongly recommended for complex mixed models (McElreath, 2016). We 198

ran 4 parallel chains for 4000 warmup iterations, 4000 sampling iterations, and a thinning interval of 4. This configuration allowed parameters to converge, which we defined as the convergence diagnostic $\hat{R} < 1.05$ (Vehtari et al., 2021). We inspected posterior predictive plots to assess model adequacy using the pp_check function from the R package **bayesplot** version 1.8.1 (Gabry et al., 2019). We used the posterior median for point estimates and calculated uncertainty with the 95% quantile intervals from the posterior distribution.

Germination rate

We estimated germination rate under benign greenhouse conditions by measuring time to germination, the number of days to visible emergence from the soil. There were 48 randomized blocks split evenly between two time cohorts (North cohort: August 3–4 and South cohort: 208 August 22–24, 2015) in the University of British Columbia greenhouse (Vancouver, Canada). The North seed cohort went into the North transplant garden (gardens described below) and the South seed cohort went into the South transplant garden. Seed cohorts were staggered so that seedlings were transplanted at the same ontogenetic stage, as described below. We distinguish 212 between cohorts and gardens because germination for each transplant garden was assayed in the 213 greenhouse prior to transplanting. Emergence is a reliable proxy for germination since seeds 214 were sown directly on top of the soil, resulting in little delay between germination and visible cotyledons. We sowed 3–5 seeds per plug on moist perlite-peat moss potting medium (Sunshine #1, Sungro Horticulture, Agawam, Massachusetts, USA) and recorded the first day on which a plug had at least one germinated seedling. We later thinned each plug to one seedling by 218 selecting the central-most individual, so the first germinated seedling recorded is generally not 210 the same individual transplanted into the field. Hence, individual correlations between germination and survival are inappropriate and we analyze relationships between breeding values or population averages. We censused germination daily for most of the experiment, but 222 census dates were more spread out later in the experiment when only a few plants remained. We 223 accounted for census interval in our statistical model (see below). Gentle misting kept soil moist 224 for the duration of the experiment. Most plugs (9320 of 10650, 87.5%) had at least one germinant 225 by the end of the experiment (North cohort: August 27, 2015; South cohort: September 20, 2015); we treated plants that did not germinate as missing data. 5 (0.05%) plants emerged before day 5.

We removed data on these individuals because they were likely not *M. cardinalis*, but rather contamination from another species' seed in the potting medium. Species can be difficult to tell apart as the cotyledons first emerge, but in plugs where a single germinant emerged before day 5 it was always a contaminant species. It therefore seems most likely that all germinants before day 5 were contamination, but observations during thinning showed that contamination was rare and it unlikely to have any substantial impact on our estimate of germination timing.

Estimating genetic variance in germination rate

In this section we describe the custom probability distribution we used to estimate quantitative genetic parameters of germination rate in this study. However, this approach could be extended to other ecologically important non-Gaussian traits censused at discrete temporal or spatial intervals. To modify the approach for other systems requires specifying an appropriate statistical distribution for the trait of interest and using integration to calculate the probability of an event occurring within a definite time or spatial interval. Probabilistic programming languages like

Stan (Stan Development Team, 2022) have made it easier for biologists to specify statistical models most appropriate for their data.

The full germination-rate model included a fixed effect of population, fixed effect of cohort, a 243 random effect of greenhouse block, a random maternal effect of dam, and a random effect of parent as either dam or sire to estimate genetic variance. Most quantitative traits are modeled with a Gaussian (i.e. normal) distribution, but this is inappropriate for germination which has a lower bound at 0 and is censused at discrete time points. A Gaussian distribution might be an 247 adequate approximation if the time to germination was longer and more spread out so that the 248 effects of a zero lower bound and discrete sampling interval were inconsequential. However, 249 given the rapid germination rate and small spread in timing for *M. cardinalis*, we used a different 250 approach. Instead, we modeled genetic variation in a continuously distributed latent rate parameter that is measured as a discrete number of days to germination. Time to germination 252 can be modeled similar to other waiting time or survival processes. There are several common 253 waiting-time distributions with a lower bound at 0, but we used the log-normal distribution 254 because unlike some other common distributions (e.g. exponential), the mean and variance of

log-normal distribution are separate parameters. Hence, we can model the variance components separate from differences in the mean germination rate. We discretized the distribution using the definite integral of the probability density to calculate a probability mass. Let the probability density with parameters $\vec{\theta}$ be $f(x|\vec{\theta})$. If a seed was censused to have germinated on day t_i , it could have germinated at any time between the previous census on day t_{i-1} . The associated probability mass $g(x|\theta)$ is:

$$g(x=t_i|\vec{\theta}) = \int_{t_{i-1}}^{t_i} f(x|\vec{\theta}) dx$$

Preliminary analyses revealed a germination threshold of four days since we did not observe any germination before the fifth day after sowing. The model did not fit as well without this threshold because models without a threshold predict a significant amount of germination between days 0 and 5 (results not shown). The variances and heritability are reported on the log-transformed scale.

₂₆₇ Transplant gardens and winter survival

After germination, we transplanted 20 out of 24 blocks of seedlings each to two experimental gardens located at the northern and southern portions of the species' range. We refer to these as 260 the "North" and "South" gardens, respectively. We used some plants from the remaining 4 blocks 270 to replace those that did not germinate or died during transport. We transplanted seedlings at similar ontogenetic stages but temporally staggered. We planted the North and South cohorts 5–6 weeks after germination on September 9–17, 2015 (North seed cohort into North garden) 273 and October 2–9, 2015 (South seed cohort into South garden), respectively. We chose times when we observed natural seedlings were of similar size. The North garden was located near natural populations along the Middle Fork of the Feather River (39° 46' 59.7" N, 120° 38' 31.1" W, 1314 mas, Plumas County, California, USA); the South garden was located along King Creek (32° 54' 23.4" N, 116° 37' 26.0" W, 1011 mas, San Diego County, California, USA). These gardens are 278 located close to two of the source populations (Appendix S₁, Fig. S₁). The North garden is 7.1 km 270 from the Little Jamison Creek source population; the South garden is 3.8 km from Sweetwater 280 River source population. Since M. cardinalis is a riparian specialist, gardens were located within

50 m of natural waterways where the microclimate closely resembled that of nearby natural populations. We irrigated plots with microperforated drip tape (Toro, Bloomington, Minnesota, 283 USA) placed about each seedling to help establish transplants and mimic riparian soil where M. 284 cardinalis naturally germinates. Seedlings were initially irrigated once per day. The duration of 285 watering was adjusted at each garden to give each plant approximately the same volume of water. Once plants were established, we gradually reduced irrigation as evaporation declined while preventing drought stress. Irrigation ended when winter precipitation began. We 288 discarded data on plants that died within the first month after transplanting, as this probably indicates transplant shock rather than natural mortality. We monitored reemergence in the spring using monthly censuses beginning on March 14, 2016 (South) and April 15, 2016 (North). 29 If a plant was recorded in the last census of 2015 and emerged in spring 2016, we counted it as 292 alive. If a plant did not reemerge, we recorded winter mortality.

Estimating genetic variance in winter survival

We treated winter survival as a binomially distributed trait determined by a latent probability of survival $p_{\rm surv}$. We therefore estimated genetic variance in $p_{\rm surv}$ and then used methods described by Villemereuil et al. (2016) for non-Gaussian traits to calculate heritability. We analyzed Garden as a fixed effect and garden Block as a random effect. The primary difference from the germination model is that we estimated genotype-by-environment ($G \times E$) interactions to test whether populations are locally adapted. If there is local adaptation, southern populations should have higher survival in the South garden and lower survival in the North garden, and *vice versa*.

Selection on germination rate

We estimated genotypic and population-level selection on winter survival as a function of germination rate. Long-term demographic studies of natural *M. cardinalis* populations use sizerather age-based models, so we do not know with certainty how important a fitness component winter survival is. However, population growth rate is moderately sensitive to the survival of prereproductive individuals of uncertain age; early survival is less important than recruitment, but more important than fecundity (Angert, 2006). If genetic variation in germination rate within

and among populations causes the probability of survival to change, then there should be significant genotypic selection on germination rate. Alternatively, populations with different 310 mean germination rates may have different probabilities of survival as a consequence of other, confounding trait differences. If this is true, we predict a population-level correlation between germination rate and survival, but no evidence of genotypic selection on variation within 313 populations. For each sample of breeding values and population averages from the posterior distribution, we used ordinary linear regression to calculate the relationship between germination rate and winter survival in both gardens at both genotypic and populational levels. We also calculated the variance among populations $(V_{
m pop})$ in germination rate and survival from the posterior distribution. We used the median and quantile intervals to estimate the slopes and 318 quantify uncertainty in our estimates. We used the same approach to estimate quadratic 319 selection coefficients as well.

Climate

We compared the climate during the experiment (September 2015 to May 2016) to a standardized climate normal. Following Sheth and Angert (2018), we downloaded climate variables for 1961–1990 and 2015–16 derived from ClimateNA version 7.10 (Wang et al., 2016). We compared the seasonal temperature average (autumn: September-November; winter: December-February; spring: March-May) for the experiment with the 1961–1990 normal for each population.

Results

Southern source populations germinate faster and have a higher probability of winter survival

The vast majority of seeds germinated in 1-2 weeks, but northern source populations took a few days longer on average. 4206 out of 5321 plugs (79.0%) had germinants in the North cohort;

5114 out of 5328 plugs (96.0%) had germinants in the South cohort. South cohort seedlings also germinated about one day faster (mean = 7.7, median = 7) than the North cohort (mean = 8.9, median = 8). The estimated log-mean difference and 95% CIs in time to germinate in the North

cohort was 0.395 (0.284, 0.505). After accounting for cohort and block effects, southern source populations consistently germinate faster (Fig. 1a). For example, seeds from the slowest 336 population (Rock Creek) germinated 2.9 (95% CI [1.8, 4.2]) days slower than the fastest population (West Fork Mojave River). The discretized log-normal distribution was an adequate model of germination rate based on the similarity of the posterior predictions to the observed 339 data (Appendix S1, Fig. S2). 340 Fewer plants survived over winter in the North garden than in the South garden (65.5% versus 341 94.2%). In both gardens, southern source populations survived more often than northern source 342 populations even after accounting for block effects (Fig. 1b). For example, in the North garden, we estimated that plants originating from the most southern source population, Sweetwater River, survived $80.5\% \ (95\% \ \text{CI} \ [71.5\%, 88.0\%])$ of the time compared to 345 52.3%~(95%~CI~[40.1%,65.8%]) for the most northern source population Rock Creek. Similarly, 346 in the South garden the local population from Sweetwater River survived $95.2\%~(95\%~{\rm CI}~[91.1\%, 97.2\%])$ of the time compared to $85.5\%~(95\%~{\rm CI}~[76.4\%, 90.8\%])$ for Rock Creek. For most populations, the average temperature was warmer than the historical normal from 1961–1990, especially for the winter season in the South garden (Appendix S1, Fig. 350 S₃). However, temperatures for the North garden were close to their historical norm. 35

Variance in germination rate and survival within and among populations

The genetic variation in germination rate within populations of *M. cardinalis* is similar to the genetic differences among them (Fig. 2a). The difference between $V_{\rm pop}$ and V_G for the log-mean 354 germination rate is -0.017 (95% CI [-0.118, 0.080]), but the confidence intervals include 0. The 355 environmental variance V_E is greater than V_G , resulting in a moderate heritability H^2 (Fig. 2b). 356 Maternal and Block effects contributed little variance (Fig. 2a). In contrast, there is virtually no genetic variance in winter survival (p_{surv}) in either garden, resulting in heritabilities that are 358 effectively 0 (Fig. 3). There is some variation among populations, but it is substantially less than 359 the unexplained environmental variance in both gardens (Fig. 3). Parameter estimates and confidence intervals for all variance components and heritabilities are given in Appendix S1, 36 Table S₁ for germination rate and Appendix S₁, Table S₂ for winter survival.

Directional selection favors faster germination among but not within populations

Since there was almost no genetic variance in winter survival in either garden, but some
variance among populations, the only evidence for selection is among populations. Source
populations originating from farther south in the species' range germinated faster and had
higher winter survival than those originating for farther north (Fig. 4). However, within
populations, genotypes that germinated faster did not have a different probability of survival.
Selection coefficients and confidence intervals are given in Appendix S1, Table S3 for
among-population selection and Appendix S1, Table S4 for genotypic selection. There was no
evidence for stabilizing selection using quadratic regression (results not shown).

Discussion

Phenotypic clines are often interpreted as evidence for local adaptation but evolutionary forces besides selection can also generate associations between phenotypes and environment (Endler, 375 1977). If genotypic selection measured under natural conditions favors local phenotypes, this 376 supports the hypothesis that phenotypic clines are adaptive. Alternatively, failure to detect 377 selection could indicate that nonadaptive evolutionary forces maintain phenotypic variation. We measured the effect of germination rate in the greenhouse, which varies clinally with latitude in 379 Mimulus cardinalis (Muir and Angert, 2017), on early winter survival. Although there was 380 genetic variation in germination rate within and among the five source populations (Table 1; 38 Appendix S1, Table S1; Fig. 1a), there was no evidence selection consistently favored local genotypes. We predicted that earlier germination should be favored in the South garden and later germination favored in the North garden. Instead, in both North and South gardens, source populations that germinated faster on average also had higher survival, though survival in the 385 North garden was lower than the South overall (Fig. 4). Apparent selection for faster germination in both gardens is likely indirect rather than causal. Populations originating from the southern range limit survived more often and germinated faster. Within populations there was no genetic variance in winter survival despite genetic variance in germination rate, indicating that differences in germination rate are probably not the main cause of differences in

winter survival. Our results are not consistent with selection on survival maintaining phenotypic clines in germination rate. Rather, the results may indicate a role for nonadaptive evolution or that selection operates on germination rate during other life stages.

Why are plants from southern source populations more likely to survive winter?

An unexpected finding is that southern source populations had higher survival in both South and North gardens. If populations were locally adapted, we would expect to see local populations 397 have the highest fitness. We consider three explanations. One explanation is that artificial conditions of the experiment may have favored genotypes from southern populations for an unknown reason. Although we attempted to situate gardens in natural conditions and at realistic conspecific densities, southern genotypes may be better adapted to artificial manipulations such 401 as removing heterospecific competitors, ground cloth, and/or irrigation watering. M. cardinalis 402 seeds require moderate temperatures similar to other co-occurring *Mimulus* species (Vickery, 1967) and water, but we do not know the minimum soil water potential required for germination. We also know little about light requirements or soil microbes and experimental coinditions might 405 have differed from natural conditions. A second possibility is that climatic anomalies favor 406 southern, warm-adapted populations in all gardens. Temperature in the North garden was 407 warmer than average for the local population (Little Jamison Creek) but not substantially warmer than the Rock Creek population, which had even lower survival. If we had planted in a season closer to the historical temperature and snowpack, then we might have observed local adaptation 410 in northern source populations. Third, early survival may tradeoff with other fitness components we have not measured or analyzed here, including seedling establishment, later viability, and/or fecundity. We cannot address selection on seedling establishment because we germinated seedlings under controlled greenhouse conditions. In Arabidopsis thaliana, faster germination 414 increased survival but decreased seedling establishment (Postma and Ågren, 2018). Southern source population growth rate could be more sensitive to germination timing, leading to stronger 416 response to selection, whereas in the north it matters less and/or conflicting selection at other life stages is stronger. We will address these hypotheses further in a separate paper including multiple seasons of viability and fecundity data from these experiments. We are not including

those data in the current analysis because we are using them to address a conceptually distinct question. However, we believe the current analysis advances the field by addressing selection on germination at one life stage and presents a new statistical method for the quantitative genetics of waiting-time traits.

What maintains genetic variation in germination rate?

One possibility is that germination rate measured in the greenhouse is not representative of 425 what occurs in nature. Species often have particular germination requirements, such as thermal time or water potential (Huang et al., 2016), to escape seed morality, as a bet-hedging strategy in unpredictable environments, or to synchronize establishment with favorable conditions (Donohue et al., 2010). In the field, seeds may use additional cues such as day length or 429 temperature that did not vary in our experiment. We may not be able to measure selection on variation in germination rate unless it occurs in the field. For example, seedlings in the north may require a longer duration of favorable conditions to ensure they do not germinate in late fall 432 or during a brief warm period in early spring. Either scenario might cause them to be exposed to 433 a hard, damaging frost. We may not be able to observe this type of selection without germinating 434 plants in the field at different times of the season. Alternatively, variation may be the result of 435 nonadaptive evolution. Mimulus cardinalis demography varies latitudinally in a pattern consistent with an ongoing northward range expansion (Sheth and Angert, 2018). Population 437 expansion can lead to the accumulation of deleterious alleles because of stronger genetic drift 438 (Peischl et al., 2013), serial founder events (Slatkin and Excoffier, 2012), and maladaptation 430 (Polechová, 2018). If selection actually favors faster germination, the slower germination and lower survival of northern populations may both be caused by the accumulation of deleterious genetic variation at the expanding range margins. The current cline could also be a legacy of 442 selection in the recent past which is now less strong or at least not observed during our study 443 (Saccheri et al., 2008). For example, snow addition resembling historical climate increases local 444 advantage in germination and establishment of Boechera stricta (Anderson and Wadgymar, 2020). If warming climate is advancing the germination window for M. cardinalis, especially in northern Sierra Nevada populations limited by winter snowmelt, this have eliminated any 447 advantage of slower germination rate.

Statistical approaches for non-Gaussian traits like germination time

Statistical challenges of analyzing non-Gaussian traits may direct research away from "difficult" 450 traits, while applying Gaussian methods to non-Gaussian traits may lead to poor estimation. We 45 show that the discretized log-normal distribution may be useful for analyzing genetic variation in germination time and other "waiting-time" traits with means close to zero. Discretized models 453 take into account that a continuous process is measured at discrete time points (e.g. daily, weekly, 454 etc.). For example, when we saw a seedling had not emerged on day t but had emerged by day 455 t+1, our estimation procedure takes into account that germination could have occurred at any time within that interval. Discretization becomes less important when the observation interval is small relative to the total duration, in which case a continuous approximation is sufficient. 458 Waiting-time distributions are naturally bounded at 0 and often skewed, both of which are 459 violated with a Gaussian distribution. We used the log-normal distribution, but other 460 distributions (e.g. exponential or Weibull) might be more appropriate in other situations. A challenge with our approach is that custom probability distributions need to be specified in a programming language. Fortunately, probabilistic programming languages like Stan (Stan 463 Development Team, 2022) are enabling evolutionary ecologists and quantitative geneticists to 464 specify custom probability distributions and estimate their parameters from data more easily and 465 reliably (see also Hadfield, 2010).

Conclusions

The reciprocal transplant experiment did not support our hypothesis that local adaptation
maintains clinal variation in germination rate in *M. cardinalis*. We cannot exclude the possibility
that artificial aspects of our experimental design may have prevented us from measuring
selection on germination rate realistically or that selection operates via different components of
fitness. It is also possible that climate anomalies during our experiment prevented us from
observing historical patterns of selection that led to present-day phenotypic variation. It is
plausible that the cline in germination rate and other phenotypes may be caused by nonadaptive
forces that increase the frequency of deleterious alleles during range expansion. Studies like ours
that combine clines with reciprocal transplants are necessary to determine what evolutionary

forces shape range-wide phenotypic variation within species.

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Author Contributions

- ⁴⁸⁶ CDM and ALA designed the experiment. CDM and CLV collected germination data. CDM and
- 487 ALA collected data on survival in the field. CDM analyzed the data and wrote the manuscript
- with input from ALA and CLV.

Data Availability Statement

- Data are available on Dryad [URL TBA]. Source code for all analyses is archived on Zenodo [URL
- TBA]. A repository with data and code is available on GitHub
- https://github.com/cdmuir/mimulus-germination.

Supporting Information

- 494 Additional Supporting Information may be found online in the supporting information section at
- the end of the article
 - Appendix S1: Supporting figures and tables (Fig. S1-S3 and Tables S1-S4)

₁₉₇ Literature Cited

526

```
Anderson, J. T., and S. M. Wadgymar. 2020. Climate change disrupts local adaptation and favours
498
       upslope migration. Ecology Letters 23: 181–192.
499
   Angert, A. L. 2006. Demography of central and marginal populations of monkeyflowers (
       Mimulus cardinalis and M. Lewisii ). Ecology 87: 2014-2025.
   Antonovics, J., and A. Bradshaw. 1970. Evolution in closely adjacent plant populations. 8. Clinal
       patterns at a mine boundary. Heredity 25: 349–362.
503
   Barrett, R. D. H., S. Laurent, R. Mallarino, S. P. Pfeifer, C. C. Y. Xu, M. Foll, K. Wakamatsu, et al.
       2019. Linking a mutation to survival in wild mice. Science 363: 499–504.
505
   Bontrager, M., C. D. Muir, C. Mahony, D. E. Gamble, R. M. Germain, A. L. Hargreaves, E. J.
       Kleynhans, et al. 2020. Climate warming weakens local adaptation. Evolutionary Biology.
507
   Brady, S. P., D. I. Bolnick, R. D. H. Barrett, L. Chapman, E. Crispo, A. M. Derry, C. G. Eckert, et al.
       2019. Understanding Maladaptation by Uniting Ecological and Evolutionary Perspectives.
500
       The American Naturalist 194: 495-515.
   Clausen, J., D. Keck, and W. Hiesey. 1948. Experimental studies on the nature of species. III:
       Environmental Responses of Climate Races of Achillea. Carnegie Institution of Washington,
512
       Washington, D.C.
513
   Colautti, R. I., and J. A. Lau. 2015. Contemporary evolution during invasion: Evidence for
514
       differentiation, natural selection, and local adaptation. Molecular Ecology 24: 1999–2017.
   DiVittorio, C. T., S. Singhal, A. B. Roddy, F. Zapata, D. D. Ackerly, B. G. Baldwin, C. R. Brodersen,
       et al. 2020. Natural selection maintains species despite frequent hybridization in the desert
517
       shrub encelia. Proceedings of the National Academy of Sciences 117: 33373-33383.
518
   Donohue, K., R. Rubio de Casas, L. Burghardt, K. Kovach, and C. G. Willis. 2010. Germination,
519
       Postgermination Adaptation, and Species Ecological Ranges. Annual Review of Ecology,
520
       Evolution, and Systematics 41: 293–319.
   Endler, J. A. 1977. Geographic variation, speciation, and clines. Princeton University Press,
522
       Princeton, N.J.
523
   Gabry, J., and R. Češnovar. 2022. Cmdstanr: R Interface to 'CmdStan'.
   Gabry, J., D. Simpson, A. Vehtari, M. Betancourt, and A. Gelman. 2019. Visualization in Bayesian
```

workflow. Journal of the Royal Statistical Society: Series A (Statistics in Society) 182: 389-402.

- Grant, P. R., and B. R. Grant. 2014. 40 years of evolution: Darwin's Finches on Daphne Major
 Island. Princeton University Press, Princeton; Oxford.
- Hadfield, J. D. 2010. MCMC Methods for Multi-Response Generalized Linear Mixed Models: The

 MCMCglmm r Package. Journal of Statistical Software 33.
- Hereford, J. 2009. A Quantitative Survey of Local Adaptation and Fitness Trade-Offs. *The*American Naturalist 173: 579–588.
- Huang, Z., S. Liu, K. J. Bradford, T. E. Huxman, and D. L. Venable. 2016. The contribution of germination functional traits to population dynamics of a desert plant community. *Ecology* 97: 250–261.
- Huey, R. B., G. W. Gilchrist, M. L. Carlson, D. Berrigan, and L. Serra. 2000. Rapid evolution of a geographic cline in size in an introduced fly. *Science* 287: 308–309.
- Huxley, J. 1938. Clines: An Auxiliary Taxonomic Principle. Nature 142: 219-220.
- Johnson, L. C., M. B. Galliart, J. D. Alsdurf, B. R. Maricle, S. G. Baer, N. M. Bello, D. J. Gibson, and
- A. B. Smith. 2021. Reciprocal transplant gardens as gold standard to detect local adaptation in
- grassland species: New opportunities moving into the 21st century. Journal of Ecology:
- 1365-2745.13695.
- Kawecki, T. J., and D. Ebert. 2004. Conceptual issues in local adaptation. *Ecology Letters* 7:

 1225–1241.
- Leimu, R., and M. Fischer. 2008. A meta-analysis of local adaptation in plants. *PLoS One* 3: e4010–e4010.
- Lescak, E. A., S. L. Bassham, J. Catchen, O. Gelmond, M. L. Sherbick, F. A. von Hippel, and W. A.
- Cresko. 2015. Evolution of stickleback in 50 years on earthquake-uplifted islands. *Proceedings*of the National Academy of Sciences 112: E7204–E7212.
- Lowry, D. B., J. M. Sobel, A. L. Angert, T. Ashman, R. L. Baker, B. K. Blackman, Y. Brandvain, et al.
- 2019. The case for the continued use of the genus name *mimulus* for all monkeyflowers.
- Taxon 68: 617-623.
- Lowry, D. B., and J. H. Willis. 2010. A widespread chromosomal inversion polymorphism
- contributes to a major life-history transition, local adaptation, and reproductive isolation.
- 555 PLoS biology 8: 2227.
- Lynch, M., and B. Walsh. 1998. Genetics and analysis of quantitative traits. Sinauer, Sunderland,

- Mass.
- McElreath, R. 2016. Statistical Rethinking: A Bayesian Course with Examples in R and Stan. 1st
- ed. Chapman; Hall/CRC.
- Mojica, J. P., and J. K. Kelly. 2010. Viability selection prior to trait expression is an essential
- component of natural selection. *Proceedings of the Royal Society B: Biological Sciences* 277:
- ⁵⁶² 2945-2950.
- Monnahan, C. C., J. T. Thorson, and T. A. Branch. 2017. Faster estimation of Bayesian models in
- ecology using Hamiltonian Monte Carlo. *Methods in Ecology and Evolution* 8: 339–348.
- Muir, C. D., and A. L. Angert. 2017. Grow with the flow: Faster growth is associated with more
- variable precipitation in a perennial herb. *Journal of Evolutionary Biology* 30: 2189–2203.
- Nelson, T. C., C. D. Muir, A. M. Stathos, D. D. Vanderpool, K. Anderson, A. L. Angert, and L.
- Fishman. 2021. Quantitative trait locus mapping reveals an independent genetic basis for
- joint divergence in leaf function, life-history, and floral traits between scarlet monkeyflower (
- Mimulus cardinalis) populations. American Journal of Botany 108: 844-856.
- Peischl, S., I. Dupanloup, M. Kirkpatrick, and L. Excoffier. 2013. On the accumulation of
- deleterious mutations during range expansions. *Molecular Ecology* 22: 5972–5982.
- Polechová, J. 2018. Is the sky the limit? On the expansion threshold of a species' range. PLOS
- Biology 16: e2005372.
- Postma, F. M., and J. Ågren. 2018. Among-year variation in selection during early life stages and
- the genetic basis of fitness in *Arabidopsis thaliana*. *Molecular Ecology* 27: 2498–2511.
- Postma, F. M., and J. Ågren. 2016. Early life stages contribute strongly to local adaptation in
- Arabidopsis thaliana. Proceedings of the National Academy of Sciences 113: 7590-7595.
- Rausher, M. D. 1992. The measurement of selection on quantitative traits: Biases due to
- environmental covariances between traits and fitness. Evolution 46: 616–626.
- 581 Richardson, J. L., M. C. Urban, D. I. Bolnick, and D. K. Skelly. 2014. Microgeographic adaptation
- and the spatial scale of evolution. *Trends in Ecology & Evolution* 29: 165–176.
- Saccheri, I. J., F. Rousset, P. C. Watts, P. M. Brakefield, and L. M. Cook. 2008. Selection and gene
- flow on a diminishing cline of melanic peppered moths. Proceedings of the National Academy
- of Sciences 105: 16212-16217.
- Santangelo, J. S., M. T. J. Johnson, and R. W. Ness. 2018. Modern spandrels: The roles of genetic

- drift, gene flow and natural selection in the evolution of parallel clines. Proceedings of the 587 Royal Society B: Biological Sciences 285: 20180230. 588 Sheth, S. N., and A. L. Angert. 2018. Demographic compensation does not rescue populations at a trailing range edge. Proceedings of the National Academy of Sciences 115: 2413–2418. Slatkin, M., and L. Excoffier. 2012. Serial Founder Effects During Range Expansion: A Spatial 59 Analog of Genetic Drift. Genetics 191: 171–181. 592 Stan Development Team. 2022. Stan Modeling Language Users Guide and Reference Manual. 593 Turesson, G. 1922. The genotypic response of the plant species to habitat. *Hereditas* 3: 211–350. Vasemägi, A. 2006. The Adaptive Hypothesis of Clinal Variation Revisited: Single-Locus Clines as a Result of Spatially Restricted Gene Flow. Genetics 173: 2411-2414. 596 Vehtari, A., A. Gelman, D. Simpson, B. Carpenter, and P.-C. Bürkner. 2021. Rank-Normalization, 597 Folding, and Localization: An Improved r for Assessing Convergence of MCMC (with 598 Discussion). Bayesian Analysis 16. 599 Vickery, R. K. 1967. Ranges of Temperature Tolerance For Germination of Mimulus Seeds From
- Diverse Populations. Ecology 48: 647–651.

 Villemereuil, P. de, H. Schielzeth, S. Nakagawa, and M. Morrissey. 2016. General Methods for

 Evolutionary Quantitative Genetic Inference from Generalized Mixed Models. Genetics 204:

604 1281-1294.

608

- Wadgymar, S. M., S. C. Daws, and J. T. Anderson. 2017. Integrating viability and fecundity
 selection to illuminate the adaptive nature of genetic clines. Evolution Letters 1: 26–39.
- Wadgymar, S. M., D. B. Lowry, B. A. Gould, C. N. Byron, R. M. Mactavish, and J. T. Anderson.
- processes that contribute to local adaptation. *Methods in Ecology and Evolution* 8: 738–749.

2017. Identifying targets and agents of selection: Innovative methods to evaluate the

- Wang, T., A. Hamann, D. Spittlehouse, and C. Carroll. 2016. Locally Downscaled and Spatially
 Customizable Climate Data for Historical and Future Periods for North America. PLOS ONE
 11: e0156720.
- Wright, K. M., D. Lloyd, D. B. Lowry, M. R. Macnair, and J. H. Willis. 2013. Indirect Evolution of

 Hybrid Lethality Due to Linkage with Selected Locus in Mimulus guttatus. *PLoS Biology* 11:

 e1001497.
- Wright, S. 1943. Isolation by distance. *Genetics* 28: 114–138.

- Wright, S. J., D. M. Goad, B. L. Gross, P. R. Muñoz, and K. M. Olsen. 2021. Genetic trade-offs
- underlie divergent life history strategies for local adaptation in white clover. *Molecular*
- Ecology: mec.16180.