

Wage Contracts and Financial Frictions

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Abstract

Financial crises often lead to drastic reductions in firms' access to credit, impacting significantly their ability to finance their operations. This paper shows that firms can partly offset the effects of these shocks by optimally adjusting their wage bills. We augment a baseline financial frictions model to account for two well-documented features of the labor market: wages are set at the firm level and within long-term employment relationships. Because of these features, wage dynamics depend on the financial conditions of firms, reflecting a trade-off between smoothing wages of risk-averse workers and investing in capital. We validate the model predictions on wage dynamics using matched employer-employee data from Italy. We find that more constrained firms adjust wages more in response to idiosyncratic shocks. In addition, firms that suffer the most during recessions back-load wages by offering steeper wage-tenure profiles to their workers. When matching these statistics with our general equilibrium model, we find that these wage adjustments reduce the sensitivity of output to financial shocks by 20%: wage back-loading enhances investment and job creation while improving allocative efficiency. We conclude by studying policies aimed at reducing inputs cost during recessions. Our findings show that these wage adjustments diminish the effectiveness of temporary payroll subsidies while enhancing the effectiveness of temporary investment subsidies in stimulating output.

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1 Introduction

After the Great Recession there has been a wide consensus that financial shocks played an important role in recent cycles. The central framework for understanding the macroeconomic effects of financial shocks hinges on the role of firms' credit, building on the work of [Kiyotaki and Moore \(1997\)](#) and [Bernanke, Gertler, and Gilchrist \(1999\)](#). The main mechanism is based on the idea that a reduction in the availability of credit forces employers to cut investment and hiring because of the shortage of funds, leading to large macroeconomic effects of financial shocks.

The standard mechanism implicitly relies on firms purchasing inputs in spot markets. In other words, whenever firms experience a financial shock the burden of adjustment must fall on quantities. This assumption is quantitatively important in explaining the large macroeconomic effects of financial shocks, especially when applied to the labor market, as the wage bill constitutes a significant share of firms' costs.

However, this assumption conflicts with two well-documented characteristics of the labor market: a significant portion of wages is determined at the firm level ([Card, Heining, and Kline, 2013](#); [Card et al., 2018](#)), and wages are set within long-term employment relationships. We argue these features are crucial as the effects of financial shocks are heterogeneous and persistent across firms. Moreover, these features are consistent with evidence that firms adjust wages over time and in response to shocks ([Guiso, Pistaferri, and Schivardi, 2005, 2012](#)), suggesting that firms can adjust wages rather than quantities when access to credit is limited.¹ How do firms adjust wages during periods of credit tightening? What are the macroeconomic implications of these wage adjustments? And how effective are stabilization policies aimed at reducing the cost of inputs during financial crises?

This paper proposes answers to these questions with four contributions. First, we propose a general equilibrium model of frictional financial and labor markets, where wages are set at the firm-level as part of long-term employment relationships, and firms face occasionally binding financial constraints. Second, we provide novel empirical evidence using matched employer-employee data that supports model predictions on wage dynamics: firms adjust wages based on their financial conditions and to ease the effects of credit constraints. Third, we show that these wage adjustments are quantitatively important during business cycle, as they significantly mitigate the macroeconomic impacts of a credit tightening. Fourth, we study the effects of temporary payroll and investment subsidies, commonly used stabilization policy to reduce the cost of inputs. We find that optimal wage adjustments over long-term employment relationships act as substitute to

¹[Guiso, Pistaferri, and Schivardi \(2005\)](#) finds that firms adjust wages in response to shocks in a way to partially insure workers' earnings from idiosyncratic fluctuations in firms' productivity. [Guiso, Pistaferri, and Schivardi \(2012\)](#) finds that firms in less financially developed regions pay steeper wage-tenure profile, thus adjusting wage payments over time.

payroll subsidies and as complement to investment subsidies.

We build on a canonical model of firms’ dynamics under financial frictions (Moll, 2014; Khan and Thomas, 2013). In our economy, firms face idiosyncratic productivity shocks and issue debt in order to finance their operations subject to a potentially binding financial constraint. As in Jermann and Quadrini (2012), we model an aggregate financial shock as a tightening of this financial constraint. Importantly, we introduce search frictions in the labor market and long-term wage contracts between firms and their workers in the tradition of Thomas and Worrall (1988).

Credit market frictions generates a trade-off between providing insurance to risk-averse workers and investing in capital. To understand the mechanism, consider a firm that currently operates below “optimal scale” due to a binding borrowing constraints. Over time, as this constraint is gradually eased, the firm’s output will increase.² Risk-averse workers would like to receive a constant wage throughout this transition, but the firm would like to pay lower wages and invest more in capital when borrowing constraints are binding. More generally, wages adjust over time and in response to shocks depending on firm-specific financial conditions.

In the model we illustrate two key findings. First, firms significantly affected by financial frictions adjust wages more in response to idiosyncratic shocks, as it is more costly for them to hedge workers against these shocks. Second, firms back-load wage payments when they require more credit, resulting in a steeper wage-tenure profile. This means that during recessions, firms more impacted by a credit tightening offer lower wages so to invest more in productive capital and ease the credit constraint, with the implicit promise of future wage increases.

We propose new empirical evidence supporting these two key predictions of the model on the dynamics of wages. We use matched employer-employee data from Italy, including administrative data on workers’ compensation and firms’ balance sheets. First, using a model-consistent indicator alongside several others, we show that in the cross-section firms that are more financially constrained adjust wages significantly more in response to shocks. Specifically, we estimate the “pass-through” of value-added per worker to wages, a commonly used statistic that measures the extent to which workers are subject to firm-specific shock. We find that this pass-through coefficient is 1.5 times larger for more financially constrained firms. Second, we show that in recession more financially constrained firms back-load wages by paying their workers according to a steeper wage tenure profile. We compare the wage-tenure profile of workers hired during the Great Recession by constrained and unconstrained firms, and we find that wages grow approximately 1.6 and 3.4 percentage points more after one and four years of tenure at constrained firms.

²This dynamics is standard and common to several models of financial frictions, as in Moll (2014) and Midrigan and Xu (2014).

The estimated model is consistent both with the empirical evidence on the heterogeneous wage dynamics across firms and stylized facts on aggregate wage dynamics. Despite firms adjust wages to offset the effects of a credit tightening, the average wage remains relatively stable in line with evidence that it moves little during recessions (Grigsby, 2022). However, the modest cyclicalities of the average wage masks substantial heterogeneity in the cross-section, where financially constrained firms contract and substantially reduce wage payments, while unconstrained firms expand and moderately increase wages. In the aggregate, this firm-level heterogeneity in wage adjustments implies that the skewness of the wage adjustment distribution is lower in recessions than in booms, consistent with evidence documented by Adamopoulou et al. (2016).

We use our model to understand how these wage adjustments affect the transmission of financial shocks. When firms back-load wage payments during recessions, this frees resources for current *and* future investment. A temporary lower wages boosts current investment *and* leads to more output next period. Because the effects of financial frictions are persistent, as in Moll (2014), an increase in output and retained earnings next period further enhances future investment. This virtuous cycle propagates forward up to a time when the effects of borrowing constraints are eased, and an increase in retained earnings has no effects on investment. The effect of wage back-loading on current *and* future investment enhances job creation, by substantially increasing the present discounted value of output over the length of the employment relationship. Additionally, the ability to make state-contingent wage adjustments lowers the expected cost of hiring a worker by minimizing the expected present discounted value of all future wage payments, discounted using the firm-specific stochastic discount factor.

In the quantitative analysis we show that firm-specific wage adjustments substantially mitigate the effects of an aggregate financial shock. Toward that purpose, we construct an alternative economy in which firms cannot commit to future wages, a restriction that prevents firms from adjusting wages over time and in response to shocks. This second economy works as the canonical model of financial frictions, so the comparison with the baseline will help understand how far dynamic wage contracts go into smoothing the effects of shocks. By comparing the impulse response functions to an aggregate financial shock in the two economies, we find that output drops by 20% less in our model.

The differential response of output to an aggregate financial shock is primarily driven by differences in aggregate employment and productivity. Employment falls less in the baseline economy, as dynamic wage contracts enhance job creation. Investment also falls less in the baseline economy, since the newly matched entrepreneurs invest in capital and wage back-loading frees resources for current investment. However, general equilibrium effects render the difference in aggregate investment between the two economies small. This is due to investment in newly created matches with high marginal product of capital “crowding-out” investment in existing matches with low marginal product of

capital. In other words, despite aggregate investment follows a similar path in the two economies, in the baseline model capital is re-allocated towards more productive firms. Consequently, the drop in endogenous total factor productivity, a measure of allocative efficiency, is less pronounced in the economy with dynamic wage contracts.

Finally, we illustrate that incorporating the dynamic structure of wage contracts during financial crises has important policy implications. We study the impact of policies aimed at reducing inputs cost during recessions, such as temporary investment and payroll subsidies.³ We find that a payroll subsidy on new hires is not as effective as a standard model with a spot labor market would suggest because firms' optimal wage adjustments and payroll subsidies act as subsidies, meaning they are both aimed at reducing the cost of labor. As a result, payroll subsidies are less effective at stimulating output during recessions when firms optimally backload wages to reduce the cost of labor. On the other hand, we find that an investment subsidy is more effective than in a standard model with a spot labor market because firms' optimal wage adjustments and investment subsidies act as complements. Indeed, when an investment subsidies transfer resources to firms and make investment opportunities more attractive, financially constrained entrepreneurs will backload wages even more during recessions as to free additional resources for investment, thus amplifying the stimulative effect of the policy. We illustrate these results quantitatively by simulating in our model investment and payroll subsidies similar to those implemented in the United States after the Great Recession.

Related Literature

This paper relates to a large literature studying the role of financial frictions during recessions. Early research showed that financial imperfections can amplify shocks originated outside the financial sectors, as in the work of [Kiyotaki and Moore \(1997\)](#), [Bernanke, Gertler, and Gilchrist \(1999\)](#), [Gertler and Karadi \(2011\)](#), [Arellano, Bai, and Kehoe \(2019\)](#). Subsequent studies showed that shocks originated in the financial sector can propagate to the rest of the economy leading to financial recessions, as in the work of [Jermann and Quadrini \(2012\)](#), [Khan and Thomas \(2013\)](#), [Buera and Moll \(2015\)](#).

Assessing the quantitative importance of these mechanisms has been a large and an ongoing area of research, as several forces can either dampen or amplify their effects. For instance, [Chari \(2012\)](#) and [Moll \(2014\)](#) pointed out that non-financial firms might be able to self-finance themselves. More closely related to our paper, [Di Tella \(2017\)](#), [Carlstrom, Fuerst, and Paustian \(2016\)](#), [Dmitriev and Hoddenbagh \(2017\)](#) show that financial contracts that are state contingent with respect to aggregate shocks can mitigate financial amplification, while [Bocola and Bornstein \(2023\)](#) study study how trade credit

³These policies have been implemented by several OECD countries, often as a temporarily accelerated tax depreciation and temporary payroll tax cuts on new hires.

within long-term supplier relationships amplifies the effects of financial shocks.⁴

Our paper differs from these studies by highlighting a new channel that mitigate the effects of financial shocks. In practice, we focus on wage contracts rather than financial contracts, and we provide rich empirical evidence supporting the main mechanism at the micro level. Conceptually, we emphasize the importance of contracts that are long-term in nature and state-contingent on idiosyncratic characteristics –and not only on aggregates.

This paper also links to the literature on optimal wage contracts within long-term employment relationships. Building on the seminal work of [Thomas and Worrall \(1988\)](#) and [Harris and Holmstrom \(1982\)](#), optimal wage contracts have been studied in the context of rich search models of the labor market by [Burdett and Coles \(2003\)](#), [Shi \(2009\)](#), [Menzio and Shi \(2010\)](#), [Fukui \(2020\)](#), [Balke and Lamadon \(2022\)](#), [Souchier \(2023\)](#), with an emphasis on the role of on-the-job search. However, this literature abstracts from firms’ investment decisions and financial market imperfections limiting firms’ access to credit. We see our paper as complementary to their work, as we study the role of financial frictions and firms’ investment decision in affecting optimal wage contracts. Related to the seminal studies by [Michelacci and Quadrini \(2009\)](#) and [Berk, Stanton, and Zechner \(2010\)](#), which explore properties of optimal wage contracts with firms’ financial constraints, we develop a business cycle model where wages respond to both idiosyncratic and aggregate shocks, provide empirical evidence supporting the model’s predictions, and assess quantitatively their macroeconomic implications for business cycle and stabilization policies.

This paper also addresses the recent macroeconomic literature exploring the link between wage rigidity and financial frictions. The work of [Favilukis, Lin, and Zhao \(2020\)](#), [Schoefer \(2021\)](#), [Wang \(2022\)](#), [Donangelo et al. \(2019\)](#), and [Acabbi, Panetti, and Sforza \(2020\)](#), is based on the idea of “labor as leverage”, where wage rigidity increases the leverage of firms, as poorly flexible payrolls act like predetermined debt obligations. Our research adds to this by considering wage rigidity arising endogenously from the optimal contract between firms and risk-averse workers. While nesting the idea of “labor as leverage” in our framework, we demonstrate through both modeling and data that financially constrained firms tend to adjust wages more after a shock. Crucially, we highlight the role of wage backloading arising within a long-term employment relationship in easing financial constraints. Thus, we show that the view of long-term employment relationship as simply a source of wage rigidity misses important patterns in the data that have large quantitative implications over the business cycle.

Finally, this paper relates to several studies documenting the effectiveness of payroll and investment subsidies. [Cahuc, Carcillo, and Le Barbanchon \(2018\)](#) and [House and Shapiro \(2008\)](#) found that the stimulus of these policies is substantial, while [Neumark and Grijalva \(2017\)](#) and [Zwick and Mahon \(2017\)](#) highlighted their increased effectiveness

⁴Other related papers introducing optimal contracts in business cycle models are [Boldrin and Horvath \(1995\)](#), [Kehoe and Perri \(2002\)](#), [Cooley, Marimon, and Quadrini \(2004\)](#).

during the Great Recession.⁵ Using a general equilibrium model, we document rich interactions between these policies and how firms set wages over long-term employment relationships during a financial crisis. We show both conceptually and quantitatively that optimal wage adjustments implied by the dynamic nature of wage contracts make temporary payroll subsidies less effective at stimulating output while making temporary investment subsidies more effective.

The paper is organized as follows. Section 2 describes the model. Section 3 illustrates the model mechanism: it characterizes properties of dynamic wage contracts, it shows how wages vary with firms' financial conditions, and illustrates how hiring and investment decisions depend on the structure of wage contracts. Section 4 illustrates novel empirical evidence on wage dynamics that validate predictions of the model. Section 5 presents the main quantitative results, discussing the macroeconomics implications of dynamic wage contracts for business cycle and stabilization policies. Section 6 concludes.

2 Model

We consider an economy populated by a continuum of entrepreneurs and workers. Entrepreneurs are heterogeneous in their productivity and produce output using capital and labor. There are financial frictions in the form of a collateral constraint, that is the borrowing capacity of entrepreneurs is limited by a fraction of the value of their capital stock, that serves as collateral. We model aggregate financial shocks as a decrease in the collateral value of capital. These features of our environment are common to several business cycle models with financial frictions and heterogeneity, as [Khan and Thomas \(2013\)](#), [Buera and Moll \(2015\)](#), [Kiyotaki and Moore \(2019\)](#).

Workers and entrepreneurs meet in a frictional labor market and engage in long-term employment relationships. As a result workers can be employed or non-employed, and entrepreneurs can be either matched with workers or vacant. Before matching with a worker, entrepreneurs offer wage contracts that specify the path of wages for any possible history of future shocks, as in [Thomas and Worrall \(1988\)](#). We describe the environment in detail in Section 2.1, we illustrate the decision problems and value functions of entrepreneurs and workers in Section 2.2, we define macroeconomic aggregates in Section 2.3, we define the equilibrium in Section 2.4, and we conclude by discussing some of the model assumptions in 2.5

⁵[Saez, Schoefer, and Seim \(2019\)](#) also documented larger impacts of payroll subsidies on financially constrained firms

2.1 Environment

Time is discrete and indexed by $t = 0, 1, \dots$. There is a continuum of entrepreneurs with measure 1, indexed by $j \in [0, 1]$, and a continuum of workers with measure M , indexed by $i \in [0, M]$. All agents in the economy have time-separable preferences with discount factor β , but entrepreneurs and workers differ in their utility functions. Entrepreneurs have utility

$$E_0 \sum_{t=0}^{\infty} \beta^t v(c_{jt}), \quad v(c) = \frac{c^{1-\sigma_E}}{1-\sigma_E}$$

and workers have utility

$$E_0 \sum_{t=0}^{\infty} \beta^t u(c_{it}), \quad u(c) = \frac{c^{1-\sigma_W}}{1-\sigma_W}$$

Production technology

If entrepreneurs are not matched with a worker, they produce the value of home production \bar{b} . Similarly, non-employed workers also produce \bar{b} . Matched entrepreneurs produce output using a standard constant returns to scale production function $f(k, \ell)$, and they are heterogeneous in their idiosyncratic productivity z , so that output y_{jt} is given by

$$y_{jt} = z_{jt} \times f(k_{jt}, \ell_{jt})$$

Idiosyncratic productivity z follows a discrete Markov process, taking values $z \in \{z_1, \dots, z_{N_z}\}$, with transition matrix Π_z . We assume that realizations of productivity shocks are independent across entrepreneurs and also independent over time. These assumptions imply a law of large numbers so the share of entrepreneurs experiencing any particular sequence of shocks is deterministic. As we assume that entrepreneurs can hire only one worker, we write the production function more compactly as $f(k) = f(k, 1)$. Entrepreneurs have access to a technology that can convert final good into physical capital one for one. In what follows we drop the j subscript whenever it is not needed for clarity.

Financial markets

Entrepreneurs are the only agents in the economy that have access to financial markets. They can borrow or save using uncontingent risk-free bonds that are in zero net-supply. Borrowing is subject to a standard collateral constraint, as entrepreneurs can borrow an amount b_{t+1} that must be less or equal to a share ξ_t of their capital stock k_{t+1} , that is

$$b_{t+1} \leq \xi_t k_{t+1}$$

While the collateral constraint is assumed to be exogenous here, it can be obtained as an endogenous outcome in an environment with limited enforcement of debt contract, where entrepreneurs can decide to not repay their debt and steal profits and share $(1 - \xi_t)$ of their capital stock.⁶

The collateral value of capital ξ_t is common to all entrepreneurs and it follows a discrete Markov Process. It can take two values $\xi \in \{\xi_L, \xi_H\}$ with transition matrix Π_ξ . We interpret ξ_H as the collateral value of capital in normal times, and ξ_L as the collateral value in recession. The stochastic behaviour of the collateral value of capital allows us to study the macroeconomics implications of aggregate financial shocks, as in [Jermann and Quadrini \(2012\)](#), [Khan and Thomas \(2013\)](#), [Buera and Moll \(2015\)](#), [Bocola and Bornstein \(2023\)](#). We refer to a change from ξ_H to ξ_L as a financial shock, as when $\xi = \xi_L$ entrepreneurs face more limited access to credit.

We assume that workers do not have access to financial markets, that is they are hand-to-mouth. This is a common assumption in models of dynamic wage contracts as it simplifies the contracting problem, and is also in line with the view that firms have better access to financial markets than workers. We see this as a conservative assumption, as it greatly limits the ability of firms to adjust wages.

Labor market

Entrepreneurs can be matched to a worker or they can be vacant. Workers can be employed or not employed. Matched entrepreneurs and employed workers separate with probability ϕ . Non-employed workers and vacant entrepreneurs meet in a frictional labor market with directed search, as in [Moen \(1997\)](#). In the economy there is a continuum of sub-markets with a constant returns to scale matching function $m(v, s)$, where v is the measure of vacant entrepreneurs and s is the measure of workers searching in a given sub-market.

Each vacant entrepreneur can open a vacancy in one sub-market. When an entrepreneur opens a vacancy he commits to a wage contract $\mathcal{C} = \{w_\tau(z^\tau, \xi^\tau)\}_{\tau=t}^\infty$ that specifies wages contingent on all future histories of idiosyncratic shocks z^τ and aggregate shocks ξ^τ . We assume that workers can also commit to a wage contract upon matching with an entrepreneur. Each non employed worker can search for a job in one sub-market. We use $\theta = v/s$ to denote the labor market tightness of each sub-market. Given the matching function, one can define the job finding probability $\lambda_w(\theta)$ and the probability of filling a vacancy $\lambda_f(\theta)$ as

$$\lambda_w(\theta) = \frac{m(v, s)}{s}, \quad \lambda_f(\theta) = \frac{m(v, s)}{v}$$

⁶See [Bocola and Lorenzoni \(2023\)](#) for an example of the limited enforcement problem with a collateral constraint.

Each sub-market is indexed by the tuple (θ, W) , where W is the expected utility of a worker conditional on finding a job in that sub-market. When entrepreneurs open a vacancy in a sub-market indexed by (θ, W) , they commit to a wage contract that will deliver the worker an expected utility equal to W . As each entrepreneur can open only one vacancy, we assume that there are no vacancy posting costs. A non employed worker can search for a job, but search is costly and it implies a disutility cost, as in models of non-participation similar to [Krusell et al. \(2017\)](#). We assume that non-employed workers who search for a job have to forgo a share x of the value of home production, in line with empirical evidence that non employed workers have to spend time searching for a job. Therefore, the flow utility of a non employed worker earning the flow output of home production \bar{b} and searching for a job is equal to $u((1-x)\bar{b})$.

Timing

At the beginning of each period idiosyncratic and aggregate shocks z_{jt}, ξ_t are realized. Each period can be divided into two stages, that we label as morning (or *before* matching and separation) and afternoon (or *after* matching and separation).

In the morning, matched entrepreneurs produce output and pay wages to the employed workers. Vacant entrepreneurs and non-employed workers produce \bar{b} . Then, vacant entrepreneurs post vacancies and non-employed workers search for jobs. Vacant entrepreneurs decide whether they want to open a vacancy or not, and if they do so they decide in which market indexed by (θ, W) . Some entrepreneurs may not find it profitable to open a vacancy, and if so they will stay vacant until the beginning of the subsequent period. Similarly, non employed workers decide whether to search for a job or not. Conditional on searching, non employed workers choose a market indexed by (θ, W) where to locate.

At the end of each morning matching and separation take place. Matched entrepreneurs can become vacant with probability ϕ , while vacant entrepreneurs who opened a vacancy in a market with tightness θ will be matched to a worker with probability $\lambda_f(\theta)$. Similarly, non-employed workers who search for a job in a market with tightness θ will be matched to an entrepreneur with probability $\lambda_w(\theta)$.

In the afternoon all agents consume. Workers consume the income earned in the morning, before matching and separation: if they were employed they consume the wage they earned, if they were not employed they consume \bar{b} or $(1-x)\bar{b}$, depending on whether they searched for a job. All entrepreneurs solve a consumption/saving problem. We assume that vacant entrepreneurs cannot hold capital, so they save using risk-free bonds.⁷ Matched entrepreneurs decide on how much to borrow or save in the risk-free bonds, and

⁷Note that vacant entrepreneurs would never choose to hold physical capital as long as the interest rate is not less than minus the depreciation rate, as capital depreciates without producing any output when entrepreneurs are not matched. As a result, vacant entrepreneurs cannot borrow and they save using the risk-free bonds.

how much capital stock to hold next period.⁸

2.2 Value functions and wage determination

We describe the problem of each agent recursively. First we discuss the recursive state space, and then we describe in details the maximization problem of each agent.

Recursive state space

We characterize the optimal contract recursively. We define a recursive contract as wages and promised utility $w'(z', \xi')$, $W'(z', \xi')$ that depends on current state variables and are contingent only on the realizations of shocks next period (z', ξ') . This formulation requires to include the utility promised to the worker W as a state variable in the problem of matched entrepreneurs. In other words, in each period, matched entrepreneurs choose state-contingent wages and promised utility for the next period, subject to a promise-keeping constraint where the expected utility of the worker must equal the utility W promised in the previous period.

Heterogeneity across matched entrepreneurs can be summarized by the exogenous state variable z and two endogenous state variables (m, W) , where W is the utility promised to the worker and m is net worth, or cash-on hand, that is equal to the sum of output and undepreciated capital stock, minus wage payments and the repayment of outstanding debt, according to the law of motion:

$$m'(z', \xi') \leq z'f(k') + (1 - \delta)k' - w'(z', \xi') - b' \quad (1)$$

Similarly, heterogeneity across vacant entrepreneurs can be summarized by the exogenous state variable z and the endogenous state variable m . Heterogeneity across employed workers is fully summarized by their expected utility W , and there is no heterogeneity across non-employed workers as they are hand-to-mouth and there is no ex-ante heterogeneity. The aggregate state of the economy, that is denoted by S is summarized by the realization of the aggregate shock ξ , and distribution of matched and vacant entrepreneurs over their states, that we denote by $\Lambda^m(m, W, z)$ and $\Lambda^v(m, z)$.

Matched entrepreneurs

At the core of our model there is the decision problem of matched entrepreneurs, whose solution characterizes the optimal wage contract. We denote by $J(m, W, z, S)$ their value function after matching and separation, according to equation (2). This depends on net worth m , idiosyncratic productivity z , utility promised to the worker W and the aggregate

⁸Capital is predetermined, as in standard business cycle models. Note that this implies the investment decision of entrepreneurs is risky, as they choose the capital stock before observing realizations of idiosyncratic productivity shocks, as in [Angeletos \(2007\)](#) and [David, Schmid, and Zeke \(2022\)](#).

state of the economy S .

These entrepreneurs choose how much to consume this period, how much to borrow or save b' —where $b' > 0$ means they borrow—, and the capital stock that will be productive next period k' . They also choose how to fulfill their promise to the worker, meaning they decide how to deliver the utility W with state contingent wages and continuation values $w'(z', \xi'), W'(z', \xi')$.

The budget constraint of matched entrepreneurs at the end of period implies that the sum of consumption and physical capital has to be equal to the sum of net worth m and net borrowing qb' . The law of motion of net worth is given by the sum of output and the undepreciated capital stock minus wages paid to the worker and the repayment of outstanding debt. Borrowing is limited by a collateral constraint, so that firms can borrow up to a share ξ of their future capital stock. The promise keeping constraint makes sure the expected utility of the worker is at least equal to the promised utility W .

The value of being matched with a worker at the end of period is equal to the flow utility of consumption plus the expected continuation values of the entrepreneur. With probability $(1 - \phi)$ the match will survive until the end of next period, while with probability ϕ the match will separate and the entrepreneur will get the continuation value V of being vacant at the end of next period. We define the value V in (4).

$$J(m, W, z, S) = \max_{\substack{c^e, b', k', m'(z', \xi'), \\ w'(z', \xi'), W'(z', \xi')}} \left\{ v(c^e) + \beta(1 - \phi) \underbrace{\mathbb{E}[J(m'(z', \xi'), W'(z', \xi'), z', S') | z, S]}_{\text{not separate}} \right. \\ \left. + \beta\phi \underbrace{\mathbb{E}[V(m'(z', \xi'), z', S') | z, S]}_{\text{separate}} \right\} \quad (2)$$

$$(\text{Budget constraint : } \lambda^e) \quad c^e + k' \leq m + qb'$$

$$(\text{Net worth : } \eta(z', \xi')) \quad m'(z', \xi') \leq z'f(k') + (1 - \delta)k' - w'(z', \xi') - b'$$

$$(\text{Collateral constraint : } \mu) \quad b' \leq \xi k'$$

$$(\text{Promise keeping : } \gamma) \quad W \leq \mathbb{E}\left[u(w'(z', \xi')) + \beta(1 - \phi)W'(z', \xi') + \beta\phi\mathcal{U}(S'') | z, S\right]$$

Vacant entrepreneurs

Vacant entrepreneurs face two decision problems: before matching and separation they can choose to open a vacancy to become matched by the end of the period, and after matching and separation they face a consumption/savings problem.

First, there is a discrete choice problem between posting a vacancy or not. En-

entrepreneurs that decide to open a vacancy have to choose a sub-market (θ, W) where to open it. With probability $\lambda_f(\theta)$ the entrepreneur is matched to a worker, where $J(m, W, z, S)$ denotes the value of a matched entrepreneur with utility W promised to the worker. With probability $1 - \lambda_f(\theta)$ the entrepreneur remains vacant, where $V(m, z, S)$ denotes the value of a vacant entrepreneur in the afternoon, after matching and separation. The entrepreneur opens a vacancy only if the expected continuation value from doing so is greater than the value $V(m, z, S)$ of being vacant at the end of period.

$$\widehat{V}(m, z, S) = \max \left(\max_{(\theta, W)} \left\{ [\lambda_f(\theta)J(m, W, z, S) + (1 - \lambda_f(\theta))V(m, z, S)] \right\}, V(m, z, S) \right) \quad (3)$$

After matching and separation, vacant entrepreneurs decide how much to consume and how much to save, according to (4). The value of being vacant at the end of period is equal to the flow utility of consumption plus the expected continuation value of being vacant next period, before matching and separation. At that stage, net worth will be equal to the returns on savings plus the flow value of home production \bar{b} .

$$V(m, z, S) = \max_{a', c^e, m'} \left\{ v(c^e) + \beta \mathbb{E} \left[\widehat{V}(m', z', S') \mid z, S \right] \right\} \quad (4)$$

$$(\text{Budget constraint}) : \quad c^e + qa' \leq m$$

$$(\text{Net worth}) : \quad m' \leq a' + \bar{b}$$

Workers

Workers decide whether to search for a job, and if they search they choose a sub-market (θ, W) where to locate.

The value of a non employed worker before matching and separation, that we denote by \mathcal{U} , is defined in equation (5). First, they face a discrete choice problem between searching and not searching.

$$\mathcal{U}(S) = \max \left(\underbrace{u(\bar{b}) + \beta \mathbb{E} [\mathcal{U}(S') \mid S]}_{\text{if not search}}, \underbrace{u(\bar{b}(1-x)) + \beta \mathcal{W}(S)}_{\text{if search}} \right) \quad (5)$$

The value of a non employed worker who does not search is equal to the flow utility of home production and the expected continuation value of being not employed next period. The value of a non employed worker who does search is given by the flow utility of home production, adjusted for the disutility cost of searching, and the expected continuation

value of a worker who search, denoted by $\mathcal{W}(S)$ and defined in equation (6).

$$\mathcal{W}(S) = \max_{(\theta, W)} \{ \lambda_w(\theta)W + [1 - \lambda_w(\theta)] \mathbb{E} [\mathcal{U}(S') | S] \} \quad (6)$$

A non employed worker who searches in sub-market (θ, W) finds a job with probability $\lambda_w(\theta)$, and receives expected utility W next period conditional on finding a job. We say that a sub-market is active if there are at last some workers and some entrepreneurs searching in that sub-market. The problem of a worker who searches, as defined by equation (6), implies that workers search in a given sub-market $(\tilde{\theta}, \tilde{W})$ if and only if it is weakly better than searching in any other sub-market, that is:

$$\lambda_w(\tilde{\theta})\tilde{W} + [1 - \lambda_w(\tilde{\theta})] \mathbb{E} [\mathcal{U}(S') | S] \geq \max_{(\theta, W)} \{ \lambda_w(\theta)W + [1 - \lambda_w(\theta)] \mathbb{E} [\mathcal{U}(S') | S] \}$$

As all non-employed workers are homogeneous, the expected continuation value of a worker who searches $\mathcal{W}(S)$ must be equalized across all the active sub-markets.

2.3 Aggregation

We define aggregate output Y_t , capital K_t , employment N_t , and investment I_t as follows.⁹

$$Y_t = \int_0^1 y_{jt} \quad (7)$$

$$K_t = \int_0^1 k_{jt} \quad (8)$$

$$N_t = \int d\Lambda_{t-1}^m(m, W, z) \quad (9)$$

$$I_t = K_t - (1 - \delta)K_{t-1} \quad (10)$$

We define aggregate debt B_t as the sum of gross debt of matched entrepreneurs:

$$B_t = \int \max(b(m, W, z, S_{t-1}), 0) d\Lambda_{t-1}^m(m, W, z)$$

Finally, we define aggregate productivity A_t such that:

$$Y_t = A_t K_t^\alpha L_t^{1-\alpha}$$

We can characterize aggregate productivity as a function of three terms, according to equation (11), where the expectation and covariance operators are taken with respect to the distribution of matched entrepreneurs Λ_{t-1}^m -normalized so that it adds up to one. The

⁹Recall that $\Lambda_t^m(m, W, z)$ and $\Lambda_t^v(m, z)$ are the distributions of matched and vacant entrepreneurs at the end of period t , after matching and separation occurred.

first term shows that A_t is proportional to the average productivity of active entrepreneurs $E[z_t]$. The second term, that is the ratio between $E[k^\alpha]$ and $E[k]^\alpha$, shows that aggregate productivity is lower when the cross-sectional dispersion in capital is higher. Finally, the covariance term implies that A_t is larger when more productive entrepreneurs produce a larger share of aggregate output.

$$A_t = E[z_t] \left[\frac{E[k^\alpha]}{E[k]^\alpha} \left(1 + \frac{\text{Cov}(z, k^\alpha)}{E[z]E[k^\alpha]} \right) \right] \quad (11)$$

More precisely, let us denote by $\tilde{\Lambda}_{t-1}^m$ the normalized distribution of matched entrepreneurs. We have that:

$$A_t = \int \sum_{z'} z' \Pi(z'|z) d\tilde{\Lambda}_{t-1}^m(m, W, z) \left[\frac{\left(\int k(m, W, z)^\alpha d\tilde{\Lambda}_{t-1}^m(m, W, z) \right)}{\left(\int k(m, W, z) d\tilde{\Lambda}_{t-1}^m(m, W, z) \right)^\alpha} \left(1 + \frac{\text{Cov}(z, k^\alpha)}{E[z]E[k^\alpha]} \right) \right] \quad (12)$$

2.4 Competitive equilibrium

Definition 1. *A recursive competitive equilibrium is defined as : i) a law of motion Γ for the aggregate state S , ii) entrepreneurs' policy functions and value functions, iii) workers' policy functions and value functions, iv) distributions of matched and vacant entrepreneurs $\Lambda^m(m, W, z), \Lambda^v(m, z)$ v) price q and market tightness $\{\theta\}$ in active sub-markets, such that:*

- *non-employed workers solve their problem given $\Gamma, \{(\theta, W)\}$*
- *entrepreneurs solve their problem given \mathcal{W}, q, Γ*
- *law of motion Γ is consistent with the policy functions and the value functions*
- *the measure of workers who search is consistent with $\{\theta\}$*
- *q is such that the bond market clears*

$$\int b'(m, W, z, S) d\Lambda^m(m, W, z) = \int a'(m, z, S) d\Lambda^v(m, z)$$

In Appendix A we show that the definition of equilibrium implies the resource constraint from Walras' Law, and we define formally the law of motion Γ for the aggregate state.

Solving this model poses some challenges, as the decision problems of matched and vacant entrepreneurs depend on $q, \mathcal{W}(S), \mathcal{U}(S)$, that are all endogenous objects, that depend on the exogenous aggregate state ξ as well as the endogenous state S . In Proposition

1, we show that the values \mathcal{W}, \mathcal{U} do not depend on the aggregate state of the economy S .¹⁰ This result substantially simplifies the analysis we conduct in Section 5, as the problem of entrepreneurs depends on S only through (ξ, q) .

Proposition 1. *If the measure of workers M is large, in equilibrium these properties hold: (i) a positive measure of workers does not search (ii) non employed workers are indifferent between searching and not searching (iii) the values $\mathcal{W}(S)$ and $\mathcal{U}(S)$ do not depend on the aggregate state S , (iv) the value functions of entrepreneurs depend on S only through (ξ, q) .*

Proof: See Appendix A

2.5 Discussion

Before moving on, let us discuss some of the assumptions we made.

First, we assume that financial frictions takes the form of a collateral constraint, consistent with a large literature that built on Kiyotaki and Moore (1997). We could have alternatively modeled financial frictions as a working capital constraint, in the spirit of Jermann and Quadrini (2012) or Bocla and Lorenzoni (2023), where firms raise funds with intraperiod loans to purchase inputs. This alternative modeling assumption would amplify the effects of wage back-loading on investment: lower wages would not only free up resources for investment by increasing net worth but also enhance borrowing capacity, as entrepreneurs would need to borrow less to pay the wage bill.

Second, we assume that entrepreneurs -firms' owners- have a concave utility function and cannot issue equity. While this assumption is fairly common in models of financial frictions (e.g. Moll (2014), Kiyotaki and Moore (2019)), our mechanism would extend to an economy with risk-neutral entrepreneurs -e.g. Khan and Thomas (2013)- that can issue equity subject to some adjustment costs, as in Jermann and Quadrini (2012). Wage contracts would still solve a risk-sharing problem, where wages co-move with the marginal value of a dollar for the entrepreneur.¹¹

Third, we assume that workers and firms can commit to a wage contract. Firms' commitment is a common assumption in the literature studying dynamic wage contracts, as in Harris and Holmstrom (1982) and Balke and Lamadon (2022), and it is often motivated by firms' reputational concerns. Here we assume that workers can also commit to

¹⁰For this step it is crucial to show that the values \mathcal{W}, \mathcal{U} do not depend on any endogenous aggregate state. Indeed, if these values were functions of exogenous aggregate states only, they would be easy to forecast and they would not pose any computational challenge to solving the model.

¹¹In this class firms don't distribute dividends and don't consume, as long as there is a positive probability of facing a binding borrowing constraint at any time in the future.

a contract, as in [Boldrin and Horvath \(1995\)](#), because we abstract from on-the-job search in our model. Indeed, the reason why workers usually cannot commit to a wage contract is they keep searching for better opportunities while currently employed. Interestingly, workers' commitment is not key to one of the main mechanism of the model, that is firms back-load wage payments when they need more credit. Indeed, if workers could not commit to a wage contract firms would *further* tilt the wage-tenure profile of their workers using wage back-loading as a tool to retain workers (e.g. [Balke and Lamadon \(2022\)](#), [Souchier \(2023\)](#)).

Fourth, we assume that each entrepreneur can hire only one worker. Despite this assumption is common to several search models of the labor market, it is not without loss of generality in a model with investment and search frictions. For instance, firms' ability to adjust wages of incumbent workers might affect the decision of hiring new employees within the same multi-worker firms. While this would be an interesting extension, the implied contracting problem would be intractable with an almost infinite dimensional state variable, since one would have to keep track of the promised utility offered to each worker.¹²

Fifth, we assume that workers are hand-to-mouth. This assumption is common to several search models of the labor market and models of dynamic wage contracts, and consistent with the view that firms have better access to financial markets than workers. Note that if workers had unrestricted access to financial markets, then entrepreneurs would implicitly borrow from their employees as to completely offset the effects of financial frictions. It would be interesting to consider an intermediate case, where workers could save and borrow subject to some borrowing constraint ([Souchier, 2024](#)). In this setting firms could implicitly borrow from their employees even more than they do under our stark assumption of hand-to-mouth workers, potentially giving firms more room to back-load wages after a credit tightening.

3 Model mechanism

This section characterizes properties of the optimal wage contract and illustrates the main model mechanism.

In [Section 3.1](#) we consider a special case of our model that is analytically tractable. We show that the optimal wage contract takes a simple form: wages are a constant share of entrepreneurs' net worth over the length of an employment relationship. This tractability allows to explicitly characterize how wages adjust over time and in response to shocks as

¹²Workers hired in different years will have different promised utilities, so one would have at least one state variable for each cohort of incumbent workers. The problem can be made tractable with some specific assumptions, as in [Michelacci and Quadrini \(2009\)](#), that consider a deterministic environment, but the general case is not.

a function of entrepreneurs' leverage, that is b/k . We focus on leverage as in our model firms with higher leverage are more likely to be financially constrained. We show that entrepreneurs with high leverage adjust wages more in response to idiosyncratic shocks, and increase wages more over time.

In Section 3.2 we turn to the general case. We use optimality conditions to illustrate the determinants of the investment, consumption, and savings decision. Then, we explain the main trade-off in wage setting, highlighting similarities with the special case. Both cases share the same underlying logic, which accounts for the heterogeneity in wage dynamics across firms. We illustrate that in the general case entrepreneurs that are more financially constrained *i*) back-load wage payments (i.e. pay steeper wage-tenure profile), *ii*) back-load wage payments even more during recessions, and *iii*) adjust wages more in response to idiosyncratic shocks.

Finally, Section 3.3 highlights the link between the structure of wage contracts, investment, and job creation, by considering the problem of vacant entrepreneurs.

3.1 Special case: analytical results

In this section we consider a special case of the problem of matched entrepreneurs that is analytically tractable. This special case allows us to solve analytically for the policy functions of consumption and wages, and to characterize how wage dynamics vary with entrepreneurs' financial conditions. We consider an economy where both entrepreneurs and workers have log utility, the production technology is linear in capital, and there is no separation.

$$v(c_t^e) = \log(c_t^e), \quad u(w_t) = \log(w_t), \quad f(k_t) = k_t, \quad \phi = 0$$

Combining the optimality conditions for state-contingent wages and workers' continuation values we obtain a risk-sharing condition, according to equation (13). Indeed, because both entrepreneurs and workers can commit to future wages, a wage contract offers a full set of state-contingent claims within the employment relationship that implies perfect risk-sharing. Since we assume that entrepreneurs and workers have the same preferences, this risk-sharing problem implies that wages move one-for-one with the entrepreneur's consumption.

$$\frac{c^e(z^{t+s}, \xi^{t+s})}{w(z^{t+s}, \xi^{t+s})} = \frac{c^e(z^{t+k}, \xi^{t+k})}{w(z^{t+k}, \xi^{t+k})} \quad \forall z^{t+s}, z^{t+k}, \xi^{t+s}, \xi^{t+k} \quad (13)$$

The next proposition characterizes policy functions for consumption and wages.

Proposition 2. *Entrepreneur's consumption and worker's wage are linear in net worth:*

$$c_t = (1 - x)m_t, \quad w_t = \gamma(1 - x)m_t$$

$$\text{with } 1 - x = \frac{1 - \beta}{1 + \beta\gamma}$$

where γ is the multiplier on the promise keeping constraint, that is constant over time.

Proof. See Appendix A

Proposition 2 shows that wages are a constant share of entrepreneurs' net worth over the length of an employment relationship. This constant share depends on the discount factor β and the multiplier on the promise keeping constraint γ . If entrepreneurs initially promised a higher utility W to the worker, the multiplier γ will be higher. Intuitively, this result implies that entrepreneurs with low net worth pay lower wages, and wages are expected to grow and adjust with net worth.

An immediate corollary of this result is that wages increase when idiosyncratic productivity increases, as that leads to an increase in future net worth of the entrepreneur.

Corollary 1. *State contingent wages $w(z_{t+s}|z^{t+s-1})$ are increasing in productivity z_{t+s} .*

We now turn to study how wage dynamics depend on entrepreneurs' financial conditions. To this end, we define leverage as the ratio between entrepreneurs' debt and capital, that is b/k . Because of the collateral constraint, leverage must be below the collateral value of capital ξ , and entrepreneurs with higher leverage are more financially constrained. The next proposition characterizes how wages change both in response to idiosyncratic shocks and over time depending on the leverage of the entrepreneur.

Proposition 3. *For each time t , the pass-through of idiosyncratic productivity shocks to wages, defined in (14), and the expected growth rate of wages, defined in (15), are increasing in entrepreneurs' leverage b_{t+1}/k_{t+1} .*

$$\frac{\partial[\log(w_{t+1}) - \log(w_t)]}{\partial[\log(z_{t+1}) - \log(z_t)]} \quad (14)$$

$$\mathbb{E}[\log(w_{t+1}) - \log(w_t)] \quad (15)$$

Proof. See Appendix A

Proposition 3 implies that highly levered entrepreneurs adjust wages more in response to idiosyncratic shocks, as for the same change in productivity they adjust wages more. Intuitively, this result is implied by the law of motion of net worth and Proposition 2. Because entrepreneurs borrow using uncollateralized debt, a higher leverage makes their net

worth in $t + 1$ more sensitive to fluctuations in future productivity at z_{t+1} . As their net worth fluctuates more, so must do the wage according to Proposition 2.

Proposition 3 also implies that the expected growth rate of wages is higher at highly levered entrepreneurs. The fact that leverage is endogenous means that highly levered entrepreneurs are choosing to borrow more and invest more in capital –relative to their net worth–. This means that highly levered entrepreneurs must expect higher returns from their investment, and thus they promise higher returns to their workers in the form of higher wage growth.

Note that the results from Proposition 3 apply to a cross-section of firm at any given point in time. We want to stress this feature of our results with respect to growth rate of wages. This means that in normal times wages grow faster at more levered entrepreneurs, but also that during recessions wages grow faster at more levered entrepreneurs.

3.2 General case

We now turn to the problem of a matched entrepreneur in the general case. First, we discuss the optimal choice of investment, debt and consumption using the optimality conditions. Then, we illustrate how wage dynamics depends on entrepreneurs' financial conditions, emphasizing how the general case shares the same underlying logic underneath the analytical result from Section 3.1.

3.2.1 Investment, debt, and consumption

The recursive problem of matched entrepreneurs defined in (2) implies fairly common optimality conditions for debt and capital:

$$v'(c^e) = \frac{\mu}{q} + \frac{1}{q} \mathbb{E} [\eta(z', \xi') \mid z, \xi] \quad (16)$$

$$v'(c^e) = \mu\xi + \mathbb{E} [\eta(z', \xi') [z' f'(k') + 1 - \delta] \mid z, \xi] \quad (17)$$

where μ denotes the multiplier on the collateral constraint, $\eta(z', \xi')$ denotes the multiplier on the law of motion of net worth that is the marginal value of a dollar for the entrepreneur in state (z', ξ') .

In the general case we have $f(k) = k^\alpha$. Compared to the special case, the marginal product of capital is decreasing in k . This implies there is always a positive and fined

first-best level of capital, according to equation (18), that we denote by k^{FB} .¹³

$$\frac{1}{q} - (1 - \delta) = f'(k^{FB})\mathbb{E}[z'|z] \quad (18)$$

Because of financial frictions, capital is always below the first best level, that is $k' < k^{FB}$. Indeed, distortions appear in the optimality condition for capital, that can be rearranged to obtain (19) by combining (16) and (17). The presence of borrowing constraint on the left hand-side of (19) reduces entrepreneurs' capital whenever the constraint is binding, that is $\mu > 0$. The presence of uninsurable idiosyncratic risk lowers the returns to capital on the right hand-side of (19), which are equal to the product of the risk-neutral marginal product of capital and a risk adjustment term. This is because, as in the seminal work of Guiso and Parigi (1999), investing in capital is risky as entrepreneurs' consumption is correlated with their idiosyncratic productivity. This risk premium term reduces returns to capital because the marginal value of a dollar for the entrepreneur, $\eta(z', \xi')$, is negatively correlated with future productivity z' .

$$\underbrace{\frac{\mu}{\mathbb{E}[\eta(z', \xi')|z, \xi]} \left(\frac{1}{q} - \xi \right) + \frac{1}{q} - [1 - \delta]}_{\text{borrowing constraint}} = f'(k') \mathbb{E}[z'|z] \underbrace{\left[1 - \frac{\text{Cov}(\mathbb{E}[\eta(z', \xi')|z', \xi], z'|z)}{\mathbb{E}[\eta(z', \xi')|z, S] \mathbb{E}[z'|z]} \right]}_{\text{idiosyncratic risk}} \quad (19)$$

Entrepreneurs with low values of net worth face a binding collateral constraint. Indeed, they are less able to self-finance investment in capital and *ceteris paribus* they need to borrow more. For higher level of net worth the collateral constraint will not be binding, and entrepreneurs will approach the first-best level of capital. As the collateral constraint is eased, leverage of entrepreneurs decreases, and entrepreneurs with higher net worth build a sufficiently high buffer stock of savings to insure against idiosyncratic shocks. As a result, the covariance term on the right-hand side of (19) decreases with net worth, making capital further close to k^{FB} .

3.2.2 Risk sharing

Equation (20) is the optimality conditions for state-contingent wages $w'(z', \xi')$. On the left-hand side of (20), the marginal cost of increasing $w'(z', \xi')$ is equal to the marginal value of a dollar for the entrepreneur in that state of the world. Intuitively, the marginal value of a dollar for the entrepreneur is high when net worth is low, that is when the entrepreneurs' level of capital is likely below the first best level. On the right-hand side,

¹³The first-best level of capital is the optimal choice of entrepreneurs in an economy with complete markets. This means there are no borrowing constraints, entrepreneurs can issue state-contingent claims to perfectly insure against idiosyncratic shocks. Note that in the first-best case there are no aggregate financial shocks, as borrowing is unconstrained.

the benefit of increasing wages is large in states when workers' marginal utility is low, and proportional to the multiplier γ on the promise keeping constraint.

$$\eta(z', \xi') = \gamma u'(w'(z', \xi')) \quad (20)$$

$$\gamma = \gamma(z', \xi') \quad (21)$$

Crucially, (20) implies a trade-off underneath wage setting between smoothing earnings of risk-averse workers –from the right hand-side– and smoothing output to relax the effects of financial frictions –from the left hand-side–. If the collateral constraint binds, smoothing output means invest more in capital in the short run to ease the constraint. If the collateral constraint does not bind, smoothing output means building a buffer stock of savings to avoid binding constraints in the future.

As in the special case, the optimal wage contracts between entrepreneurs and workers can be described as the solution to a risk-sharing problem. The ratio between the marginal utility of the worker and the marginal value of a dollar for the entrepreneur is equal to γ , which does not depend on future states (z', ξ') . The optimality conditions for state-contingent continuation values $W'(z', \xi')$ implies that γ has to be constant over time, according to equation (21). Therefore, as long as the entrepreneur and the worker do not separate, the optimal wage contract implies perfect risk-sharing with respect to idiosyncratic productivity shocks and aggregate financial shocks.¹⁴ As consumption of entrepreneurs depends on their net worth, wages also depend on net worth because of optimal risk sharing, in a way that reminds the special case.

Differently from the special case, the degree of risk-sharing underneath wage contracts depend on the relative risk-aversion coefficients of entrepreneurs and workers, σ_E and σ_W . Under the assumption of CRRA preferences, σ_W captures both the relative risk-aversion coefficient and the inverse elasticity of inter-temporal substitution. In Appendix D we characterize the solution of the risk-sharing problem under the assumption that workers have Epstein-Zin preferences (Epstein and Zin, 1989), and we highlight the role played by risk-aversion and inter-temporal substitution. Interestingly, the ability of firms to adjust wages over time depends both on the elasticity of inter-temporal substitution and the degree of risk-aversion of workers. Indeed, future earnings are uncertain due to the risk of separation, and wage dynamics depend on workers' preference towards persistence risk, that is a function of both risk-aversion and inter-temporal substitution.

3.2.3 Wage dynamics and firms' financial conditions

In the special case presented in Section 3.1 wage dynamics depend on financial frictions: more constrained entrepreneurs adjust wages more in response to idiosyncratic shocks and

¹⁴Since wages cannot be paid after separation, entrepreneurs and workers are still exposed to the idiosyncratic risk of separation.

pay steeper wage-tenure profiles. We now illustrate the mechanism in more details for the general case, where entrepreneurs with low net worth are *ceteris paribus* financially constrained, as they are not able to self-finance investment.¹⁵

To illustrate the mechanism consider two entrepreneurs that start at time $t = 0$ with different net worth m_0 , but same promised utility to the worker and same productivity (W_0, z_0) . Figure 1 plots whether the constraint binds, wages, and capital as a function of time for these two entrepreneurs.

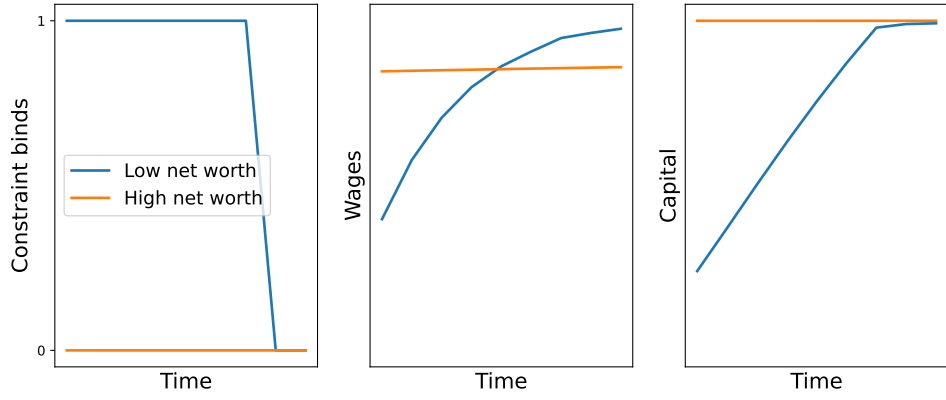
Panel (a) considers a case where no shocks are realized. In this case, the entrepreneur that starts with high net worth is unconstrained, and both capital and wages are almost flat over time. On the other hand, the entrepreneur that starts with lower net worth is financially constrained and offers a steep wage tenure profile to the worker. This means wages are lower initially, but they increase more over time as the entrepreneur accumulates net worth and approaches the first best level of capital. As in this example both workers receive the same expected utility W_0 , but the timing of wage payments differ, we say that the entrepreneur with low net worth back-loads wages. This property of wage contracts aligns with the findings of Michelacci and Quadrini (2009), that studies optimal wage contract when firms are subject to financial constraints in a deterministic environment.

Consider now an aggregate financial shock. Panel (b) plots whether the constraint binds, wages and capital as a function of time, when there is a one-time financial shock at $t = 0$ (dashed line). The entrepreneur with high net worth is only mildly affected by the shock, that leads to small changes in wages and capital. The entrepreneurs with low net worth is greatly affected by the shock, leading to a substantial contraction in current and future investment. In response to the shock, the entrepreneur with low net worth makes the wage-tenure profile of the worker steeper, thus back-loading wages more after an aggregate financial shock.

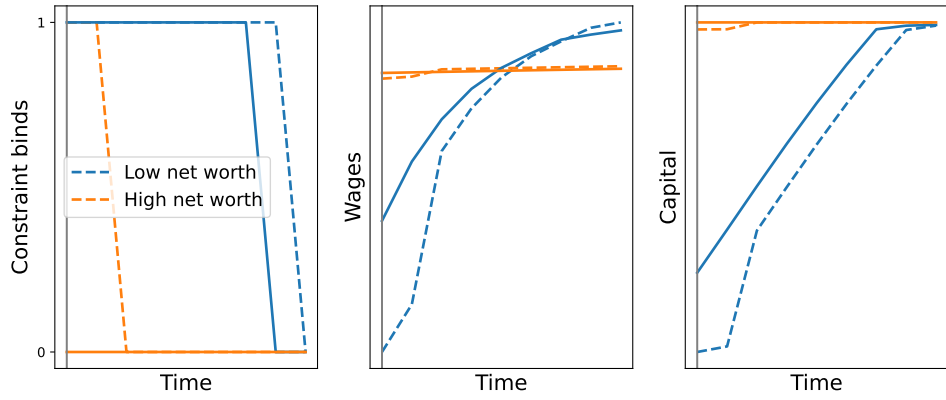
The dashed lines in Panel (c) plot whether the constraint binds, wages and capital as a function of time when there is a positive idiosyncratic productivity shock at time $t = 0$.¹⁶ The mechanism is the same underneath Corollary 1 and Proposition 3. An increase in productivity leads to higher net worth and higher wages, because of optimal risk-sharing. The entrepreneur with low net worth is more sensitive to changes in idiosyncratic productivity, as higher leverage makes his returns more correlated with fluctuations in output. Consequently, entrepreneurs with low net worth adjust wages more in response to idiosyncratic shocks.

¹⁵This means that in the entrepreneurs with low net worth have higher leverage, so that by emphasizing the link between wage dynamics and net worth we also illustrate how the findings from the special case extend to the general case.

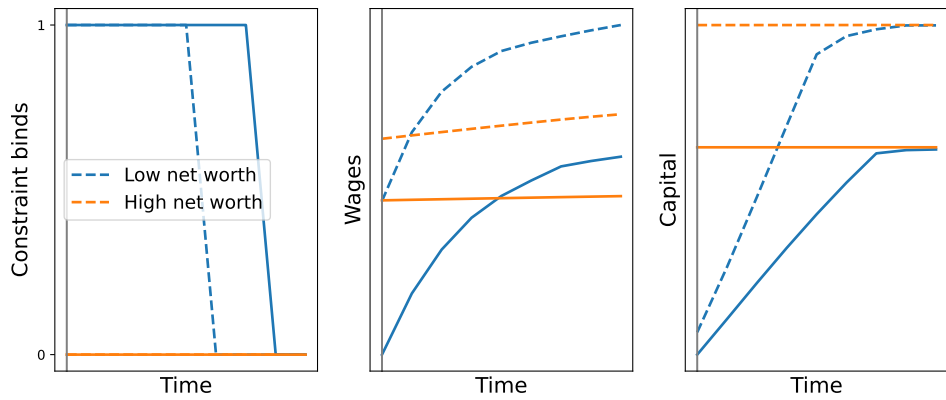
¹⁶We consider a permanent increase in productivity that occurs at period $t = 0$, so that productivity increases and stays higher in future periods.



(a) No shocks



(b) Aggregate financial shock: $\xi \downarrow$



(c) Idiosyncratic productivity shock: $z \uparrow$

Figure 1: Wage-tenure profile, capital and net worth dynamics for entrepreneurs with same (W, z) but different net worth m in the first period. Wages, capital and net worth are plotted an entrepreneur with low net worth (blue line) and high net worth (orange line).

3.3 Job creation, investment, and wages

The dynamic structure of wage contract, namely the fact that entrepreneurs can adjust wages both over time and in response to shocks, affect both investment and job creation. First, we discuss the implications for investment, relating to the problem of matched entrepreneurs. Then, we characterize the optimal job creation decision of vacant entrepreneurs and we explain how it depends on the dynamic structure of wage contracts.

3.3.1 Investment and wages

The ability of firms to adjust wages both over time and in response to shocks enhances investment. First, the ability to back-load wages over time allow entrepreneurs that are currently financially constrained to pay lower wages in the short term so to have more resources to invest in capital. Additionally, as illustrated in (20), wage contracts determine the degree of risk-sharing between the entrepreneur and the worker. When an entrepreneur shares some idiosyncratic risk with the worker, all else equal, it decreases his exposure to idiosyncratic risk, thus making investment less risky and more attractive.

To understand these effects, consider an alternative economy where wages are determined on a spot labor market. Spot market wages would not be contingent on firms' idiosyncratic productivity simply because of the law of one price. As a result, the entrepreneurs would bear entirely the investment risk, making capital less attractive. Similarly, if wages would be determined on a spot labor market, entrepreneurs would not adjust the wage-tenure profile of workers according to their financial conditions.

Importantly, even a small change in the current wage can have large effects on the entire dynamics of investment. Consider an entrepreneur with a binding constraint that offer a lower wage at time t . Current investment mechanically increases as more resources become available. If the collateral constraint binds also in future periods, future investment also increases. Indeed, higher investment in t increases net worth in $t + 1$, which further enhances investment in $t + 1$, thus increasing net worth in $t + 2$, and so on.¹⁷ As highlighted in Khan and Thomas (2013) and Moll (2014), the effects of financial frictions are persistent both in recession and in the cross-section, suggesting that a small change in wages can have persistent effects on investment. This dynamic effect of current wages on future investment has implications for job creation, that we illustrate in the next section.

3.3.2 Job creation

Job creation is determined by the solution of the problem of vacant entrepreneurs, according to (3), who decide whether to open a vacancy and a sub-market in which to post. The latter decision implies the job creation equation (22) that mimics a standard

¹⁷Note that if an entrepreneur becomes unconstrained in $t + 1$, higher earnings in $t + 1$ will have little effects on investment in $t + 1$ as the entrepreneur would be able to self-finance himself.

surplus sharing rule common to a large class of search models.¹⁸ On the left-hand side of (22) there is the ratio between the surplus of the entrepreneur, that is $J - V$, and the surplus of the worker, that is $W - \mathcal{U}$. On the right-hand side there is the product between a standard term in search models that reminds of the Hosios' condition¹⁹ and the multiplier on the promise keeping constraint γ , that intuitively captures how costly it is for the entrepreneur to deliver the promised utility W .

In the standard competitive search model the path of wages is not uniquely pinned down, as standard assumptions on risk-sharing and complete financial markets make it irrelevant in terms of the equilibrium allocation. In this model, because of incomplete markets, the specific path of wages is allocative for job creation and thus affects the equilibrium allocation. We emphasize two ways in which the dynamic structure of wage contracts affect job creation.

$$\frac{J(m, W, z, S) - V(m, z, S)}{W - \mathcal{U}} = \underbrace{\left(-\frac{\partial J(m, W, z, S)}{\partial W} \right)}_{\gamma(m, W, z, S)} \left(\frac{1 - \eta}{\eta} \right) \quad (22)$$

First, the structure of wage contracts affects job creation through the surplus ($J - V$). For instance, when an entrepreneur with low net worth back-load wages to enhance investment, he will make the worker more productive over the length of a match increasing the surplus ($J - V$). In the previous section we emphasized that even a small change in current wage can have large effects on the dynamics of investment when credit constraints bind for several periods. Crucially, job creation that depends on the present discounted value of a match and thus internalizes these dynamics effects of wages on investment.

Second, the entrepreneur optimally set wages in order to minimize the present discounted value of all future wage payments made to the worker. Intuitively, the ability of making state-dependent wage adjustments in the future decreases the cost of hiring a worker in present discounted value terms today, thus fostering job creation. The following lemma formalizes this result.

¹⁸Appendix A derives equation (22) from the problem of vacant entrepreneurs using Proposition 1.

¹⁹In the efficient allocation of a standard search model, the ratio between firms' surplus and workers' surplus is equal to the ratio $(1 - \eta)/\eta$. See Hosios (1990).

Lemma 1. *The optimal wage contract that solves (2) is also a solution to:*

$$\min_{\{w_{s+1}(z^{s+1}, \xi^{s+1})\}_{s=t}^{\infty}} \mathbb{E} \left[\sum_{s=t}^{\infty} w_{s+1}(z^{s+1}, \xi^{s+1}) \times [\beta(1 - \phi)]^{(s-t)} \frac{\eta_{s+1}(z^{s+1}, \xi^{s+1})}{v'(c_t^e)} \right] \quad (23)$$

$$\text{subject to: } W \leq \mathbb{E} \left[\sum_{s=t}^{\infty} [\beta(1 - \phi)]^{(s-t)} u(w_{s+1}(z^{s+1}, \xi^{s+1})) + [\beta(1 - \phi)]^{(s-t)} \beta \phi \mathcal{U} \right]$$

Proof: See Appendix A

4 Empirical evidence

We use matched employer-employee data from Italy to provide empirical evidence supporting model's prediction on wage dynamics.

In Section 3.1 we showed that financially constrained firms – those with higher leverage according to the model – adjust wages more in response to idiosyncratic shocks and offer wages that are expected to grow faster. In Section 3.2 we illustrated the broad implications of these results for how wages adjust with tenure at different firms. Financially constrained firms offer their workers a steeper wage-tenure profile, and thus increasing wages over time as they accumulate more capital and net worth. This heterogeneity in wage-tenure profiles across firms is particularly pronounced after an aggregate financial shock, where firms with high leverage make the wage-tenure profiles of their workers even steeper in response to the shock.

In the empirical analysis, we validate two key predictions of the model: financially constrained firms adjust wages more in response to idiosyncratic shocks and offer steeper wage-tenure profiles during recessions. Testing these predictions is crucial for two reasons. First, the differential wage response to idiosyncratic shocks provides evidence for the *risk-sharing* mechanism underneath wage-setting. In practice, we estimate that financially constrained firms display higher pass-through of value-added per worker to wages than unconstrained firms. We focus on pass-through coefficients as a measure of risk-sharing between firms and workers, building on the work of (Guiso, Pistaferri, and Schivardi, 2005; Guiso and Pistaferri, 2020). Second, the steeper wage-tenure profiles offered by constrained firms during recessions highlight the *dynamic* nature of wage adjustments over the length of employment relationships, at the same time emphasizing how firms optimally adjust wages to ease the effects of an aggregate tightening. In the data, we show that during the Great Recession financially constrained firms offered steeper wage-tenure profiles to their workers than the unconstrained ones.

Since identifying financially constrained firms is notoriously difficult, we show that

our results are robust to using several methods to classify firms as more or less financially constrained. In the baseline specifications we consider firms with higher leverage to be more financially constrained. This is consistent with both the model and several studies that use leverage as a proxy for the strength of financial frictions across firms (Gopinath et al., 2017; Ottonello and Winberry, 2020). Moreover, leverage is a particularly suitable proxy in countries like Italy, where firms rely heavily on bank financing and where bank credit is largely collateralized (Garrido, Kopp, and Weber, 2016; Affinito, Sabatini, and Stacchini, 2021). Furthermore, evidence shows that firms with higher leverage were more significantly impacted by the Great Recession and the EU Sovereign Debt crisis, as they rely more on external financing and were therefore more affected by a credit tightening (Arellano, Bai, and Bocola, 2017; Kalemli-Özcan, Laeven, and Moreno, 2022). Leverage also correlates with other commonly used proxies for financial constraints, such as credit scores and distance to default (Rodano, Serrano-Velarde, and Tarantino, 2018; Ottonello and Winberry, 2020).

We find similar results using several alternative methods to classify firms as financially constrained. In a first set of robustness exercises, we classify firms as financially constrained using other common indicators constructed from balance sheet data. More precisely, we identify financially constrained firms depending on their debt-to-output ratio and interest coverage ratio. This approach is motivated by evidence that these ratios are central to commonly used loan covenants, which restrict borrowing for firms with low earnings relative to their stock of debt or interest payments (Greenwald, 2019; Drechsel, 2023). In a second set of robustness exercises, we classify firms as financially constrained based on heterogeneous fluctuations in local credit supply during the Great Recession (Greenstone, Mas, and Nguyen, 2020; Barone, de Blasio, and Mocetti, 2018). More precisely, a shift-share style research design is implemented to isolate local lending shocks using variation in preexisting bank market shares and bank supply shifts. In this way, we identify as more financially constrained those firms operating in local credit markets that experienced a larger credit supply shock in 2008.

The section is organized as follows: first we describe the data and the institutional background. Section 4.3 presents the empirical strategy and the main results on heterogeneous wage-tenure profiles during the Great Recession, and Section 4.4 present results on heterogeneous pass-through coefficients.

4.1 Matched employer-employee data

The analysis relies on matched employer-employee data sourced from official social security records maintained by the Italian Social Security Institute (Istituto Nazionale Previdenza Sociale, INPS). This comprehensive dataset, spanning from 2005 to 2019, includes the entire employment history of all private-sector employees who were at any time

employed by firms participating in the Survey of Industrial and Service Firms (INVIND survey) conducted by the Bank of Italy. The relevant population for the INVIND survey includes firms in manufacturing and services, with at least 20 employees. Known as the INPS-INVIND dataset, it integrates employer and employee data, covering approximately 10 million workers and 10 thousand firms over the 2005-2019 period.²⁰

From INPS records, for each job spell in every year, we observe worker and firm identifiers, along with gross earnings, the number of weeks worked in full time equivalent units, part-time status and a coarse occupational code (apprentice, blue collar, high-skilled blue collar, white collar, middle manager or manager). For each worker we also observe a series of basic demographic characteristics such as gender and year of birth. We complement the INPS-INVIND data by matching it to balance sheet and income statement information from CERVED.²¹ For each firm in the sample we retrieve firms' total assets, value added and various measures of firm debt, including debt with banks, suppliers, or other intermediaries. These variables are particularly important as they allow us to construct different measures of leverage. We describe in more details the data, the cleaning procedure, and the sample construction in Appendix B.

In some robustness exercises, we use granular data on bank lending from the Bank of Italy Supervisory Report database to construct local credit supply indexes, and data from INVIND survey questions, such as the total share of overtime hours at the firm level.

4.2 Institutional background

The relevant tiers for wage formation in the Italian labor market are at the industry and company levels (Guiso, Pistaferri, and Schivardi, 2005). The first pillar consists of sectoral bargaining agreements at the industry level that establish “minimum wages” (contractual minimums, or *minimi tabellari*) for different job ladder levels. The second pillar consists of company-specific components of the compensation package, allowing the firm to unilaterally determine certain elements or negotiate them in a company contract with unions. The most significant portion is the company-level wage increment (*super-minimum*), which permanently raises the contractual minimum wage in nominal terms and includes both firm- and worker-level components. Additional forms of firm-specific compensation include transitory production bonuses (*premi di produzione*) and a variable pay component (*retribuzione variabile*).

A substantial share of workers' compensation package is determined at the firm level. Using data from 1975 to 2000, Card, Devicienti, and Maida (2014) show that nearly

²⁰The INPS-INVIND data have been used in a number of recent studies, including (Macis and Schivardi, 2016; Daruich, Di Addario, and Saggio, 2023; Di Addario et al., 2023) among others.

²¹CERVED is a leading Italian data provider, offering detailed balance sheet data and comprehensive business information.

all employees earn some premium above the contractual minimum, with the median premium being 24%. Guiso, Pistaferri, and Schivardi (2005) reports that in 1994 the average wage component due to firm-specific pay policies was around 23%. The latter grew to around 30% in 2009 according to the same data source (Darulich, Di Addario, and Saggio, 2023). Since then, the Italian labor market has witnessed a gradual erosion of centralized bargaining agreements (D’Amuri and Giorgiantonio, 2015). This process culminated in the Inter-sectoral Agreement of June 2011 which broadened the scope of decentralized wage-bargaining and defined the procedures for its activation. Meanwhile Law 148/2011 introduced the possibility of signing firm and local-level agreements in derogation of the law and of the national collective agreements.

4.3 Wage backloading in the Great Recession

We now discuss the estimation of wage-tenure profiles for workers employed by firms with different levels of leverage, to provide empirical evidence that firms with high leverage offered steeper wage-tenure profiles during the Great Recession. In Section 4.3.3 we discuss other methods to identify more financially constrained firms.

We consider a log wage equation according to (24). In equation (24) $w_{ij(t_0)t}$ and $T_{ij(t_0)t}$ are earnings and tenure at time t of worker i , who was hired by firm j in t_0 . We allow wages to be a non-parametric function of tenure, defined by tenure-specific coefficients β_s for $s = 0, 1, \dots$. In equation (24) lev_{jt_0-1} is leverage of firm j at time $t_0 - 1$, and X_{ijt_0} is a vector of firms’ and workers characteristics at time t_0 . We allow different wage-tenure profiles depending on firms leverage at the time of hire –whether it was above or below the median at $t_0 - 1$ –, and on workers’ and firms’ observable characteristics at t_0 . Wages also depend on a match-specific time invariant term, that we denote by μ_{ij} .

$$\begin{aligned} \log w_{ij(t_0)t} = & \mu_{ij} + \sum_{s=0}^S \beta_s 1(T_{ij(t_0)t} = s) + \sum_{s=0}^S \gamma_s 1(T_{ij(t_0)t} = s) 1(lev_{jt_0-1} > \text{median}) \\ & + \sum_{s=0}^S \delta_s X'_{ijt_0} 1(T_{ij(t_0)t} = s) + u_{ij(t_0)t}, \end{aligned} \quad (24)$$

Before discussing threats to identification, we take differences of (24) between t and t_0 . On the left-hand side of (25) there is the cumulative wage growth between t and t_0 of worker i , who was hired by firm j in t_0 . Since we are interested in differences in wage growth between firms with different leverage, our analysis focuses on the coefficients $\tilde{\gamma}_s = (\gamma_s - \gamma_0)$.²²

²²Note that permanent match-specific quality, as captured by the term μ_{ij} , is not a concern as it does not affect estimates of $\tilde{\gamma}_s$.

$$\begin{aligned}
(\log w_{ij(t_0)t} - \log w_{ij(t_0)t_0}) &= \sum_{s=1}^S \tilde{\beta}_s 1(T_{ij(t_0)t} = s) \\
&+ \sum_{s=1}^S \tilde{\gamma}_s 1(T_{ij(t_0)t} = s) 1(lev_{jt_0-1} > \text{median}) \\
&+ \sum_{s=1}^S \tilde{\delta}_s 1(T_{ij(t_0)t} = s) X'_{ijt_0} + (u_{ij(t_0)t} - u_{ij(t_0)t_0}) \quad (25)
\end{aligned}$$

We restrict our attention to workers hired in January and whose first year of year-round employment at the firm was in $t_0 = 2009$. This way, we compare wage-tenure profiles of one cohort of workers hired during the Great Recession by different firms.²³ We focus on heterogeneous wage-tenure profiles within a relatively medium-term horizon of five years, where we expect to see more action, thus setting $S = 4$.

4.3.1 Identification

There are two main concerns associated with estimating $\{\tilde{\gamma}_s\}$ from equation (25). First, leverage is an endogenous variable that can be correlated with other firms' and workers' characteristics that influence future wage growth independently from leverage. Second, tenure is the outcome of endogenous decisions of workers, who may decide to change jobs, and firms, which can terminate a job. We now discuss how we address these concerns in detail.

Leverage at $t_0 - 1$ may be correlated with firms' and workers' observable characteristics that have an effect on the wage-tenure profile independently from leverage (e.g. leverage may be correlated with firms' size). To address this concern, we include firms' sector, total assets and value-added per worker at $t_0 - 1$ in $X_{ij(t_0)t_0}$, thus controlling for sector, size and value-added per worker specific tenure profiles. Similarly, we include dummy variables for workers' occupation, age and gender in $X_{ij(t_0)t_0}$, thus controlling for occupation, age and gender specific tenure profiles.

A similar concern is that leverage at $t_0 - 1$ may be correlated with firms' permanent unobservable characteristics that affect wage growth. For instance, some firms may have better access to credit and persistently sustain high leverage without being financially constrained.²⁴ To address this concern, we consider a specification similar to (25) where

²³Pooling different cohorts of workers to compare post-Great Recession wage growth requires separately accounting for wage changes during the recession based on pre-existing tenure and experience. This is conceptually challenging and requires extensive data. Since data on workers' careers are available starting from 2005, tenure for all incumbent workers cannot be observed.

²⁴A model assumption is that the collateral value of capital ξ is common to all firms. In practice different firms may have different values of ξ . Moreover, in the model financial constraints are occasionally binding, meaning that firms expect to ease the effects of the constraints over time. In practice some firms may always have high values of leverage, for instance due to "zombie lending".

we classify firms as financially constrained using within-firm variation in leverage, rather than pure cross-sectional variation (Ottonello and Winberry, 2020). More precisely, we allow the wage tenure profile to depend on whether leverage at t_0 was greater than the firms' own average, that is if $lev_{jt_0-1} > E[lev]_j$, where we calculate the average over the entire sample.

To further corroborate our results and address potential endogeneity issues, we perform a series of robustness exercises using alternative methods to identify more financially constrained firms, which will be discussed in more detail in the next section. Specifically, we employ a shift-share style research design to isolate local lending shocks based on variations in preexisting bank market shares and bank supply shifts (Greenstone, Mas, and Nguyen, 2020). This approach enables us to identify financially constrained firms without relying on firm-specific characteristics, thereby mitigating the endogeneity problems associated with using balance sheet data to derive proxies for financial constraints. Additionally, our results remain robust even when using different measures of financial constraints derived from balance sheet data.

Comparing workers with same tenure levels at different firms can induce bias in the coefficients $\{\tilde{\gamma}_s\}$ as workers' mobility and thus tenure is endogenous (e.g. better workers or higher-quality matches may correlate positively with lower quit rates, and workers' mobility decision may be correlated with firms' leverage). We address this issue by implementing two corrections that combine insights from Abraham and Farber (1987); Dustmann and Meghir (2005).

First, exploiting the nature of our exercise that aims to compare wage growth for workers in the first $S = 4$ years of tenure, we focus on a sample of workers hired at t_0 who then remain with the same firm for at least $S = 4$ years. This way, we compare workers with same *completed tenure*.²⁵ While comparing workers with same completed tenure controls for different quit rates, this step introduces sample selection. For instance, highly levered firms may fire workers more often over time, keeping only the most productive ones.²⁶

Second, we correct for sample selection using a common exclusion restriction (Dustmann and Meghir, 2005; Guiso, Pistaferri, and Schivardi, 2012), based on whether the *current* job is found following firm closure or a mass layoff at the *previous* job.²⁷ The idea

²⁵Since we are interested in tenure profiles over the first five years of tenure, the relevant completed tenure is within five years horizon, which we observe. If one was interested in estimating tenure profiles over an unbounded horizon, then a proxy for estimated completed tenure should be used as in Abraham and Farber (1987).

²⁶A similar concern has to do with learning: highly levered firms may keep only workers that, over time, discovered to be the best fit for the firm. These effects are often ruled out in the empirical literature (Dustmann and Meghir, 2005), as the effects of tenure are hard to interpret with learning about match quality.

²⁷We define job displacement exactly as in Jacobson, LaLonde, and Sullivan (1993): we identify the workers who left their employer at the same time that the employer experienced a 30% or larger decline in employment.

is that displaced workers must start searching for a new job sampling from the unconditional distribution of match values, while those who moved voluntarily to the current firm did it because they improved their match value, (i.e. they sampled from the conditional distribution). Hence, the probability of being a mover out of the *current* job must be higher for the displaced worker than for the average worker.

4.3.2 Results

Table 1 reports the estimates of equation (25) using only the sample of job stayers. Since workers' mobility decision is endogenous, we present estimates with and without controlling for selection into the current job (Columns 2 and 3, respectively).

We report the probit estimates for completed tenure in Column 1, where the dependent variable is the probability that workers' completed tenure is greater or equal than S . These are used to construct an estimate of the inverse Mills ratio. The exclusion restriction in the mobility probit (a variable that affects the decision to move but does not affect wage growth) have the expected effect: being a displaced worker has a negative effect on completed tenure implying a higher probability of moving out of the current job. The instrument is statistically significant, suggesting that the estimates do not suffer from weak exclusion restriction problems. Further details about the mobility probit and the correction for sample selection are discussed in Appendix B.3.

We report estimates of the coefficients $\{\tilde{\gamma}_s\}$ in Columns 2 and 3. The estimated coefficients are positive, statistically significant and increasing in tenure, meaning that highly levered firms offered steeper wage tenure profiles. This result aligns with our model prediction, that during a credit tightening firms with high leverage offer a steeper wage-tenure profile to their workers. Our baseline measure of leverage is the ratio between firms' debt, measured as the sum of all financial debt and debt towards suppliers, divided by the firm's total assets. Adjusting for selection in Column 3 results in lower estimates of the coefficients $\{\tilde{\gamma}_s\}$, as expected. Moreover, the coefficients on the inverse Mills ratio become larger and statistically significant as tenure increases, suggesting that selection becomes more and more relevant over time.

4.3.3 Robustness

We perform several robustness exercises using alternative methods to classify firms as financially constrained. Results are reported in Table 2. In a first set of exercises, we explore alternative indicators of financial frictions derived from balance sheet data, using both cross-sectional and within-firm variation. We present additional evidence for these results by classifying firms as financially constrained based on heterogeneous fluctuations in local credit supply during the Great Recession.

First, we propose an alternative way to construct the leverage ratio b/k , where the

| Wage growth and firms' leverage | | | |
|--|------------------------|------------------------|-----------------------|
| | Probit (1) | Cumulative wage growth | |
| | | (2) | (3) |
| 1 ($lev_{jt_0-1} > \text{median}$) | 0.0486*** (0.0140) | | |
| Tenure=1: 1 ($lev_{jt_0-1} > \text{median}$) | | 0.0237*** (0.0023) | 0.0166*** (0.0019) |
| Tenure=2: 1 ($lev_{jt_0-1} > \text{median}$) | | 0.0418*** (0.0028) | 0.0236*** (0.0023) |
| Tenure=3: 1 ($lev_{jt_0-1} > \text{median}$) | | 0.0480*** (0.0029) | 0.0326*** (0.0024) |
| Tenure=4: 1 ($lev_{jt_0-1} > \text{median}$) | | 0.0494*** (0.0031) | 0.0344*** (0.0026) |
| Displaced | -0.1553*** (.0372) | | |
| Tenure=1: Inverse Mills ratio | | | -0.0005 (0.0098) |
| Tenure=2: Inverse Mills ratio | | | 0.0068 (0.0118) |
| Tenure=3: Inverse Mills ratio | | | 0.0293** (0.0123) |
| Tenure=4: Inverse Mills ratio | | | 0.0334** (0.0132) |
| Workers' controls | Yes | Yes | Yes |
| Firms' controls | Yes | Yes | Yes |
| N. of workers | 130,775 | 100,680 | 100,680 |
| N | 130,775 | 402,720 | 402,720 |

Table 1: Column (1) reports the results of the probit estimates of whether completed tenure of each worker-firm match is above or below S , using the entire sample of workers hired at t_0 . Columns (2) and (3) report the estimates of equation (25) using the sample of workers who stayed at least 4 years. Robust standard errors in parenthesis. The superscripts ***, **, and * denote statistical significance at the 1, 5, and 10 percent level, respectively.

denominator includes only financial debt, excluding debt to suppliers. We report results using cross-sectional variation in leverage and within-firm variation in leverage in Columns 1 and 2 respectively. As explained above, when using within-firm variation (Ottonello and Winberry, 2020) we replace the indicator function on whether firms’ leverage at $t_0 - 1$ is above median with an indicator function for whether firms’ leverage at $t_0 - 1$ is above that firms’ own average. In practice, the indicator function is replaced by $1(lev_{jt01} > E[lev]_j)$, where $E[lev]_j$ is the sample average of firm j . We find similar results using cross-sectional and within-firm variation in leverage. We interpret these results as evidence that our findings are not driven by permanent unobservable heterogeneity across firms.

Second, we consider firms with a high debt-to-output ratio to be more financially constrained, consistent with a view of earnings-based constraints as opposed to collateral constraints (Drechsel, 2023). In practice, firms’ borrowing capacity might be limited by their earnings and not by the collateral value of their assets. Despite the Italian credit market relies more on collateralized borrowing than on this type of loan covenants, we find similar results. We measure firms’ debt-to-output ratio as firms’ financial debt divided by firms’ value added. In this case, we also report results using cross-sectional variation and within-firm variation in debt-to-output ratio. In Appendix B.4 we present similar findings by classifying firms as more or less financially constrained based on their interest coverage ratio (Greenwald, 2019). These findings show that results are robust to alternative methods to classify firms as more or less financially constrained derived from balance sheet data.

Finally, we identify as more financially constrained those firms exposed to a more pronounced tightening of local credit conditions. To this end, we rely on a shift-share style research design to isolate local lending shocks using variation in preexisting bank market shares and bank supply shifts (Greenstone, Mas, and Nguyen, 2020). This approach allows us to identify more financially constrained firms without relying on firm-specific characteristics, thus addressing potential endogeneity problems associated with using balance sheet information to derive proxies for financial constraints.

We use granular data on bank lending from the Bank of Italy Supervisory Report database to construct a province-time index of credit supply as in Barone, de Blasio, and Mocetti (2018).²⁸ We denote by S_{pt} a measure of change in credit supply for province p in year t . This measure is the weighted average of bank- b -specific changes in outstanding loans δ_{bt} between year t and $t - 1$, weighted by the market share W_{bp} of each bank b in

²⁸Provinces are administrative units roughly comparable to a US county. According to the Italian Antitrust authority the “relevant market” in banking for antitrust purposes is the province (Guiso, Pistaferri, and Schivardi, 2012). Italy is divided into 110 provinces.

| | Wage growth and financial frictions | | | | |
|-----------------------|---|-----------------------|-------------------------|-----------------------|------------------------------|
| | Leverage ratio (excluding suppliers) | | Debt-to-output ratio | | Local credit supply shock |
| | (1) | (2) | (3) | (4) | (5) |
| Tenure=1: constrained | 0.0206*** (0.0017) | 0.0197*** (0.0018) | 0.0065*** (0.0019) | 0.0149*** (0.0019) | 0.0205*** (0.0022) |
| Tenure=2: constrained | 0.0320*** (0.0021) | 0.0402*** (0.0022) | 0.0156*** (0.0022) | 0.0259*** (0.0023) | 0.0346*** (0.0027) |
| Tenure=3: constrained | 0.0351*** (0.0022) | 0.0385*** (0.0023) | 0.0200*** (0.0024) | 0.0262*** (0.0028) | 0.0303*** (0.0028) |
| Tenure=4: constrained | 0.0323*** (0.0024) | 0.0345*** (0.0025) | 0.0204*** (0.0026) | 0.0284*** (0.0026) | 0.0310*** (0.0030) |
| Within firm variation | No | Yes | No | Yes | |
| Inverse Mills ratio | Yes | Yes | Yes | Yes | Yes |
| Workers' controls | Yes | Yes | Yes | Yes | Yes |
| Firms' controls | Yes | Yes | Yes | Yes | Yes |
| N | 402,720 | 402,720 | 402,720 | 402,720 | 395,000 |

Table 2: Estimates of the coefficients $\tilde{\gamma}_s$ from equation (25) obtained using different methods to classify firms as more or less financially constrained. Robust standard errors in parenthesis. The superscripts ***, **, and * denote statistical significance at the 1, 5, and 10 percent level, respectively.

province p before the crises.²⁹

$$S_{pt} = \sum_b W_{bp} \hat{\delta}_{bt} \quad (26)$$

We consider firms that operate in province p with an index of credit supply conditions S_{pt} below the mean as more financially constrained. The estimates in Column 5 of Table 2 indicate that wages grow faster in the first two years of tenure at firms experiencing a stronger deterioration in local credit conditions. These results, based on an exogenous measure of tight credit conditions, further corroborate the mechanism and address remaining endogeneity issues associated with using balance sheet data to identify constrained firms. More details on how we construct this credit supply index and additional results are included in Appendix B.6.

One concern associated with measurements of earnings is that estimates of equation (25) could be influenced by adjustments in hours worked. Since we focus on full-time workers employed year-round in all empirical exercises, we expect hours to remain stable. However, firms may still adjust overtime hours for these workers. In Appendix B.5, we

²⁹In the baseline specification we use market shares in 2006. In Appendix B.6 we show that results are not sensitive to this choice.

use data on average overtime hours at the firm level from the INVIND survey to show that differences in overtime hours between constrained and unconstrained firms are small and not statistically significant during and after the Great Recession and do not exhibit a similar dynamic pattern as that of wages.

In Appendix B.8 we exploit differences in wage flexibility across firms to propose descriptive evidence that links these wage adjustments to firms' investment decision: firms with greater wage flexibility back-loaded wages more and experienced a smaller drop in investment during the Great Recession.

4.4 Estimates of pass-through coefficients

We estimate the pass-through coefficients of value-added per worker to wages, following a well established methodology that built on [Guiso, Pistaferri, and Schivardi \(2005\)](#). Estimates of pass-through coefficients in the range between 5% and 15% have often been interpreted as evidence of some form of risk-sharing, or partial insurance, within long-term employment relationships.

In the empirical analysis we isolate idiosyncratic changes in firms' value added per worker from fixed heterogeneity and changes common to all firms in the same sector. To this end, we construct residuals of log of value added per worker against firm-specific fixed effects and 2-digit NACE sector-year fixed effects. Then, we take first differences of the residuals, that we denote by ε_{jt} . For each firm-worker match we measure average monthly earnings in any given year. We focus on a sample of *stayers*, who work full-time and for 52 weeks in the same firm for at least two consecutive years.

We isolate idiosyncratic changes in workers' earnings from changes that can be attributed to demographics characteristics (gender-age-occupation fixed effects) or aggregate trends (year fixed effects). Then, we take first differences of the residuals, that we denote by ω_{jt} .³⁰

Once we have constructed measures of the idiosyncratic components of earnings growth and growth in value added per worker, we estimate the pass-through of value added per worker to wages using standard techniques from [Guiso, Pistaferri, and Schivardi \(2005\)](#); [Guiso and Pistaferri \(2020\)](#). We define the pass-through coefficient as the following moment of the data:

$$\mathcal{P} = \frac{Cov\left(\Delta w_{ijt}, \sum_{s=-1}^1 \Delta \varepsilon_{jt+s}\right)}{Cov\left(\Delta \varepsilon_{jt}, \sum_{s=-1}^1 \Delta \varepsilon_{jt+s}\right)} \quad (27)$$

³⁰We winsorize ε_{jt} and ω_{jt} at the 1st and 99th percentile to remove for possibly extraordinary events that are not present in our model.

| Estimates of pass-through coefficients | | | |
|---|-----------------------|-----------------------|-----------------------|
| | (1) | (2) | (3) |
| $\Delta\varepsilon_{jt}$ | 0.0612*** (0.0004) | | |
| $\Delta\varepsilon_{jt} \times 1(\text{Below median leverage})$ | | 0.0474*** (0.0004) | 0.0414*** (0.0005) |
| $\Delta\varepsilon_{jt} \times 1(\text{Above median leverage})$ | | 0.0680*** (0.0006) | 0.0760*** (0.0005) |
| Firms' controls | No | No | Yes |
| N | 11,052,040 | 11,052,040 | 11,052,040 |

Table 3: Estimates of the pass-through coefficient of firms' value-added per worker to wages, obtained using the instrumental variable estimator defined in (27). The first column reports the average pass-through. In the other columns we split the sample and report estimates for firms with leverage above and below median. Robust standard errors in parentheses. The superscripts ***, **, and * denote statistical significance at the 1, 5, and 10 percent level, respectively.

which corresponds to the instrumental variable regression of Δw_{ijt} on $\Delta\varepsilon_{jt}$, using $(\varepsilon_{jt+1} - \varepsilon_{jt-2})$ as an instrument. The instrumental variable strategy filters out the effect of purely transitory shocks, which we interpret as measurement error. In a model with permanent productivity shocks and static pass-through as Guiso, Pistaferri, and Schivardi (2005), this estimator recovers the true pass-through of shocks to the persistent component of firms' idiosyncratic productivity to workers' earnings.

In our model the pass-through coefficient is not static –as wages adjust dynamically in response to shocks– and productivity shocks are not permanent. Consequently, in our model the pass-through coefficient of value-added per worker to wages is different from the pass-through of idiosyncratic productivity shocks to wages. We regard these pass-through coefficients of value-added per worker to wages as crucial moments of the data that measure the degree of risk-sharing between firms and workers, which we use to discipline and validate our model. We estimate an average pass-through coefficient of 6.1%, as reported in Column 1 of Table 3, in line with the existing literature. Using Italian data from 1982 to 1994 Guiso, Pistaferri, and Schivardi (2005) finds an average pass-through coefficient of 6.8%.

We proceed by estimating how the pass-through coefficient varies with firms' leverage. We split firms according to whether at $t - 1$ their leverage was above or below the median in the leverage distribution of that year, and estimate the pass-through coefficient for each sub-sample. Our baseline measure of leverage is the ratio between firms' debt, measured as the sum of all financial debt and debt towards suppliers, divided by the firm's total

assets. Results are reported in Column 2 of Table 3.³¹ Firms with leverage above median have a pass-through coefficient of 6.8%, that is almost 1.5 times larger than the estimated pass-through for firms having leverage below median.

As we discussed in Section 4.3, leverage can be correlated with other firms' observable characteristics that affects the pass-through coefficient independently from leverage.³² In Column 3 we present results obtained by first residualizing leverage at $t - 1$ against log assets and log value-added per worker at $t - 1$, and then categorizing firms based on whether this residualized measure of leverage is below or above the median. Controlling for firms' observable characteristics, firms with leverage above median have a pass-through coefficient of 7.6%, that is almost 1.8 times larger than the estimated pass-through for firms having leverage below median.

In Appendix B.7 we illustrate how the results presented in Table 3 are robust to using alternative methods to classify firms as financially constrained.

5 Quantitative analysis

We present the quantitative model starting from the calibration of model's parameters. Then, we assess the ability of the quantitative model to reproduce salient features of the data in Section 5.2. Sections 5.3 and 5.4 present the main quantitative results aimed at quantifying the importance of dynamic wage contracts for the propagation of financial shocks and for the effects of stabilization policies. We solve the model using techniques from Krusell and Smith (1997), including the price q as a state variable in the entrepreneurs' problems and approximating a forecasting rule for future prices. Details are described in Appendix C.

5.1 Calibration

We begin by describing how we choose parameters for our quantitative analysis. We interpret one period in the model as one quarter. The firm level production function and the matching technology are both Cobb-Douglas and described in equations (28), (29), where B is a constant that measures matching efficiency, η is the matching function elasticity and α is the production function elasticity.

$$f(k, \ell) = k^\alpha \ell^{1-\alpha} \tag{28}$$

$$m(v, s) = Bv^{1-\eta}s^\eta \tag{29}$$

³¹Results are robust to considering alternative measures of leverage.

³²Correlation of firms' leverage with workers' observable characteristics is not a source of concern here, as ω_{ijt} is the idiosyncratic component of earnings that is orthogonal to workers' observable characteristics.

| Assigned parameters | | | | |
|----------------------------------|---|------------------------------------|-------|--------|
| Parameter | Intuition | Value | | |
| β | Discount factor | 0.99 | | |
| δ | Depreciation rate | 0.025 | | |
| α | Capital share | 0.3 | | |
| η | Matching function elasticity | 0.5 | | |
| σ_E | RA coefficient of entrepreneurs | 1 | | |
| $P(\xi_L \xi_H)$ | Probability of financial recession | 0.01 | | |
| $P(\xi_L \xi_L)$ | Persistence of financial recession | 0.8 | | |
| ϕ | Separation probability | 0.024 | | |
| x | Share of time spent searching for a job | 0.05 | | |
| Externally calibrated parameters | | | | |
| Parameter | Intuition | Estimated value | | |
| ρ | Persistence of idiosyncratic productivity z | 0.96 | | |
| $\sigma(\varepsilon)$ | Std. deviation of innovations in z | 0.07 | | |
| Internally calibrated parameters | | | | |
| Parameter | Value | Moment | Model | Target |
| ξ_L | 0.33 | Drop of corporate debt in 2008 | -14% | -14% |
| ξ_H | 0.87 | 99p of leverage distribution | 0.87 | 0.87 |
| σ_W | 11 | Pass-through of VA/worker to wages | 0.06 | 0.06 |
| B | 0.5 | Job finding rate | 0.30 | 0.30 |
| \bar{b} | 0.15 | Unemployment rate | 8.1% | 8.1% |

Table 4: Values for all model’s parameters.

Nine parameters are assigned and listed in Table 4. We set the common discount factor β equal to 0.99, the depreciation of capital equal to 0.025, and the elasticity of the Cobb-Douglas production function equal to 0.3, as these are standard values in macroeconomics. We follow [Ljungqvist and Sargent \(2017\)](#) in setting the matching function elasticity η equal to 0.5, based on estimate for several EU countries reviewed in [Petrongolo and Pissarides \(2001\)](#). We assume that entrepreneurs have log-utility, that is $\sigma_E = 1$, following a large body of work that studied the effects of financial frictions in models with heterogeneity, as [Midrigan and Xu \(2014\)](#), [Moll \(2014\)](#), [Kiyotaki and Moore \(2019\)](#).³³ We set $P(\xi_H|\xi_H) = 0.99$ to be consistent with the notion that financial crises are rare events in advanced economies. We set $P(\xi_L|\xi_L) = 0.8$, meaning the average duration of a recession in the model is five quarters, that corresponds to the length of the Great Recession in Italy according to the OECD based recession indicators. Finally, we set the probability of separation in the model equal to 0.024, as to match the separation rate measured in [D’Amuri et al. \(2022\)](#) for Italy.³⁴ We set the disutility cost of searching for a job in line

³³This assumption is not far from empirical evidence. Estimates from [Herranz, Krassa, and Villamil \(2015\)](#) imply a median relative risk aversion coefficient of entrepreneurs of 1.5, obtained using a sample of small firms in the US.

³⁴Using microdata from the Labour Force Survey, [D’Amuri et al. \(2022\)](#) find a quarterly EU rate (employment to unemployment) of 1.2% and a quarterly EN rate (employment to out of labor force) of 1.2% for men aged 35-55, that implies a quarterly separation rate equal to 2.4%. We focus on males

with evidence on the amount of time spent searching for a job (Manning, 2011; Krueger and Mueller, 2010). In practice, we set x equal to 0.05, meaning that agents have to forego 5% of the value of home production when they search for a job.

The remaining parameters are estimated in two separate exercises. We recover the parameters that discipline the stochastic process for firms' idiosyncratic productivity, the persistence ρ and the standard deviation $\sigma(\varepsilon)$, by estimating a stochastic process for firms' productivity from balance sheet data. We use a GMM estimator that filters out fixed heterogeneity across firms as well as purely transitory productivity shocks, as none of them are present in the model. Details are discussed in Appendix B.9. The remaining five parameters $(\xi_H, \xi_L, B, \bar{b}, \sigma_W)$ are chosen simultaneously so that the model matches a set of moments of interest. Below, we describe the targeted moments and discuss heuristically which model parameters they help us discipline.

We pin down the collateral value of capital ξ_L during financial crises by targeting the observed 14% drop in the aggregate debt of the non-financial sector in 2008³⁵, following the same calibration strategy proposed in Khan and Thomas (2013). We set the collateral value of capital during normal times ξ_H , that is the maximum leverage b/k that a firm can have, as to match the 99-th percentile of firms' leverage distribution. Motivated by the analytical results illustrated in Section 3, that is wages co-move with entrepreneurs' consumption according to the ratio of the two relative risk aversion coefficients, we set the relative risk aversion coefficient of workers in order to match the average pass-through of value-added per worker to wages that we estimated in Section 4. We estimate the value of home production \bar{b} and the efficiency of the matching function B following a calibration strategy similar to Shimer (2005). We target an average unemployment rate of 8.1%, that was the average unemployment rate in Italy for the years before 2008, and an average job finding rate equal to 30%, consistent with empirical evidence from D'Amuri et al. (2022), Cingano and Rosolia (2012).

5.2 Model vs Data

The quantitative model can match several untargeted moments describing wage dynamics, leverage distribution and investment dynamics. The first panel of Table 5 shows that the model can quantitatively replicates the heterogeneity in wage dynamics across leverage that we documented in Section 4.4. As in the data, the pass-through of value-added worker to wages is higher for firms with leverage above median, and the magnitudes are comparable. Also, we estimate the wage tenure-profile for workers hired in recession by firms with leverage above and below median. We find that the wages of workers hired by

between 35-55 as the impact of fertility, schooling, and retirement decision on the EN rate is negligible, as these features are not present in the model.

³⁵The outstanding amount of total debt securities in non-financial corporations sector was 100 billions in Q1 of 2008 and 87 billions in Q4 of 2008.

highly leveraged firms during a recession increase by approximately 2 percentage points after one year of tenure and 3 percentage points after four years of tenure, consistent with the empirical evidence presented in Section 4.3.

The model generates a cross-sectional distribution of firms' leverage similar to the data. Panel B in Table 5 reports the three quartiles of the leverage distribution, that are remarkably similar between the model and the data, despite targeting only the 99th percentile. Note that in normal times only few firms face a binding collateral constraint, as the 75-th percentile of the leverage distribution is substantially below ξ_H . We measure the average standard deviation of investment rates (i/k) in the data using a balanced panel of firms, finding the same value documented in Cooper and Haltiwanger (2006), which is very close to the value implied by the model.³⁶

We inspect the ability of our model to reproduce the dynamics of macroeconomic aggregates during the 2008 recession in Italy. Panel C in Table 5 reports the drop in aggregate output, employment, total factor productivity, and investment from peak to trough during the first year of recession. The model has been calibrated to match a drop in aggregate debt of 14%, and it accounts for a substantial share of the observed drop in aggregate output, consistent with the view that financial shocks played a key role in the Great Recession.

The model also implies a substantial drop in aggregate employment, investment and total factor productivity. The dynamics of employment and capital depend on a direct effect and a general equilibrium effect.

A drop in the collateral value of capital ξ leads to a contraction in debt and investment for entrepreneurs with low net-worth that are not able to self-finance themselves. The contraction in aggregate debt generates excess savings in the economy that increases the price of bonds q . At the same time, an increase in q induces wealthy unconstrained entrepreneurs to substitute away from risk free bonds and to invest more in physical capital.³⁷

The response of aggregate employment is disciplined by the job creation decision of vacant entrepreneurs. When ξ drops, the surplus of a match $J - V$ falls for entrepreneurs with low net worth. As a result, they either post a lower promised value W – with a drop in the vacancy filling rate $\lambda_f(\theta)$ – or they decide not to open a vacancy at all, leading to a drop in aggregate employment. In general equilibrium the price of bonds increases fostering job creation of unconstrained entrepreneurs. Indeed, entrepreneurs with high net worth experience an increase in the match surplus $J - V$, as investment in capital

³⁶An alternative calibration strategy for the productivity process would have been to estimate the standard deviation of innovation to productivity by targeting the standard deviation of investment rates, as in Khan and Thomas (2013). We rather estimate a productivity process externally and we keep this moment for the validation.

³⁷This heterogeneous effects of an aggregate financial shock on firms' investment decision is similar to the one illustrated by Khan and Thomas (2013).

| Panel A: Wage dynamics | | | | |
|--|--|--------------------|---|-----------------------|
| | Pass-through coefficients $\times 100$ | | Additional wage growth at levered firms | |
| | Leverage < median | Leverage > median | 1y horizon | 4y horizon |
| Model | 4.1 | 7.1 | 1.9pp | 3.0pp |
| Data | 4.1 | 7.6 | 1.7pp | 3.4pp |
| Panel B: Investment and leverage | | | | |
| | p25 leverage | p50 leverage | p75 leverage | Investment volatility |
| Model | 0.20 | 0.46 | 0.63 | 0.30 |
| Data | 0.25 | 0.43 | 0.70 | 0.33 |
| Panel C: Macroeconomic effects of a financial shock | | | | |
| | Drop in aggregate variables, first year of recession | | | |
| | $\Delta \log(Y_t)$ | $\Delta \log(N_t)$ | $\Delta \log(I_t)$ | $\Delta \log(A_t)$ |
| Model | -4.7 % | -2.9% | -24 % | -3.6 % |
| Data | -7.1 % | -3.4% | -12 % | . |
| Panel D: Distribution of wage adjustments | | | | |
| | Skewness | | Kelley's skewness | |
| | Recession | Normal times | Recession | Normal times |
| Model | 1.1 | 2.5 | -0.47 | - 0.02 |
| Data | 0.20 | 0.71 | 0.01 | 0.20 |

Table 5: Panel A reports moments associated to wage dynamics in the model and empirical findings from Section 4. Panel B reports moments for the distribution of leverage and investment dynamics in the model and in the data. Panel C reports the drop of macroeconomics aggregates from peak to trough in the first year of the recession in the data and according to the impulse response functions of the model. Panel D reports the skewness of the wage adjustment distribution in the model and in the data ([Adamopoulou et al., 2016](#)).

becomes relatively more attractive than saving in risk-free bonds.

While the direct effect dominates, the coexistence of these two countervailing channels affects allocative efficiency. Because of the direct effect, constrained entrepreneurs with high marginal product of capital are forced to reduce their investment, while wealthy entrepreneurs with lower marginal product of capital invest more. This reallocation of capital towards unconstrained entrepreneurs leads to a drop in aggregate productivity A , as defined in equation (11).

The model is not only consistent with evidence on the heterogeneous wage dynamics presented in Section 4, but also with stylized facts on aggregate wage dynamics. The model implies a modest cyclicalities of the average wage in recession, that drops by 0.3%. The modest decline in the average wage masks substantial heterogeneity in the cross-section: firms strongly impacted by the credit tightening cut their wages substantially, while unconstrained firms expand and pay higher wages. In the aggregate, this firm-level heterogeneity in wage adjustments implies that the skewness of the wage adjustment distribution is lower in recessions than in booms. This feature of our model is consistent with recent evidence that the skewness of the wage adjustment distribution changes during recession (Adamopoulou et al., 2016), as reported in Panel D of Table 5.³⁸ The model is also consistent with evidence from Gertler, Huckfeldt, and Trigari (2020), Grigsby, Hurst, and Yildirmaz (2021) that wages of new hires are as cyclical as wages of incumbents, and with evidence from Kudlyak (2014) and Basu and House (2016) that the average user cost of labor is substantially more cyclical than the average wage. We provide more details in Appendix C.4.

5.3 The macroeconomic implications of dynamic wage contracts

Dynamic wage contracts, meaning the ability of firms to adjust wages over time and in response to shocks, play an important role when the economy is hit by an aggregate financial shock. In order to quantitatively evaluate this channel we compare our model to a counterfactual economy where firms cannot optimally adjust wages.

There are two fundamental assumptions in the model that allow firms to adjust wages over time and in response to shocks: employment relationships are long-term in nature and firms can commit to future wages. Indeed, for firms to be able to back-load wage payments they must stay matched with the same workers for more than one period, and they must be able to implicitly promise higher future wages to their employees. To quantitatively assess the importance of dynamic wage contracts during financial crises, we relax the former assumption on firms' commitment, as this exercise can be done without

³⁸The model is consistent with evidence that the skewness of the wage adjustment distribution dropped substantially during recessions. The values of such skewness differ between model and data because several other features that affects wage adjustments, such as human capital, are not present in our model.

changing the economic environment.³⁹

We propose a counterfactual economy that has two features of a spot labor market: firms cannot adjust timing of wage payments over the length of an employment relationship, and the allocative wage for job creation is only the current wage w_t . These features are obtained by assuming entrepreneurs can commit only to a wage for the first period after matching with a worker. More precisely, if an entrepreneur hires a worker in period t , we assume that the entrepreneur can commit only to a wage w_{t+1} , that is for the first period in which the match is productive. Consistently with the assumption of no commitment on the firms' side, we also assume that workers cannot commit to an employment relationship with a given entrepreneur. As a result, entrepreneurs would pay workers their outside option \bar{b} for any period $s > t + 1$, as this is a dominant strategy compared to any wages above \bar{b} . As in the baseline model, firms compete for workers in the expected utility W of a match, but they do so by choosing only w_t . Therefore, the main difference from the baseline model is that firms cannot choose how to deliver the utility W over time and in response to shocks by adjusting wages.

Figure 2 plots the impulse response function of macroeconomic aggregates to an aggregate financial shock in the baseline model and in the counterfactual economy with no commitment. The two economies display the same drop in aggregate debt.⁴⁰ However, in the counterfactual economy output is more than one percentage point lower compared to the baseline economy, meaning that dynamic wage contracts substantially mitigate the effects of an aggregate financial shock on output.

The differential response of output is primarily driven by differences in aggregate employment and productivity. In the baseline economy firms optimally adjust wages to ease the effects of the credit tightening. As a result, aggregate employment and investment fall less in the baseline economy, as illustrated in Figure 2. In other words, in the baseline economy financially constrained entrepreneurs invest more than in the counterfactual economy. This implies that in general equilibrium unconstrained entrepreneurs invest *less* in the baseline economy, as the opportunity cost of capital q is greater. As a result, the impulse response functions of aggregate investment do not differ substantially between the two economies. However, this general equilibrium effect implies that in the baseline economy aggregate productivity is higher, as capital is reallocated from unconstrained entrepreneurs with low marginal product of capital to constrained entrepreneurs with high marginal product of capital. This reallocation implies that the drop in aggregate productivity is less pronounced in the baseline economy, as shown in Figure 2.

³⁹On the other hand, relaxing the assumption of long-term employment relationships would substantially change the economic environment, thus limiting the extent of the quantitative exercise to capture only the role of dynamic wage contracts, which is the main channel we aim to isolate.

⁴⁰We re-calibrated the value of ξ_L in the counterfactual economy as to obtain the same drop in aggregate debt, as we interpret a drop in aggregate debt, rather than a lower value of ξ_L *per se*, is the primitive shock propelling a financial recession.

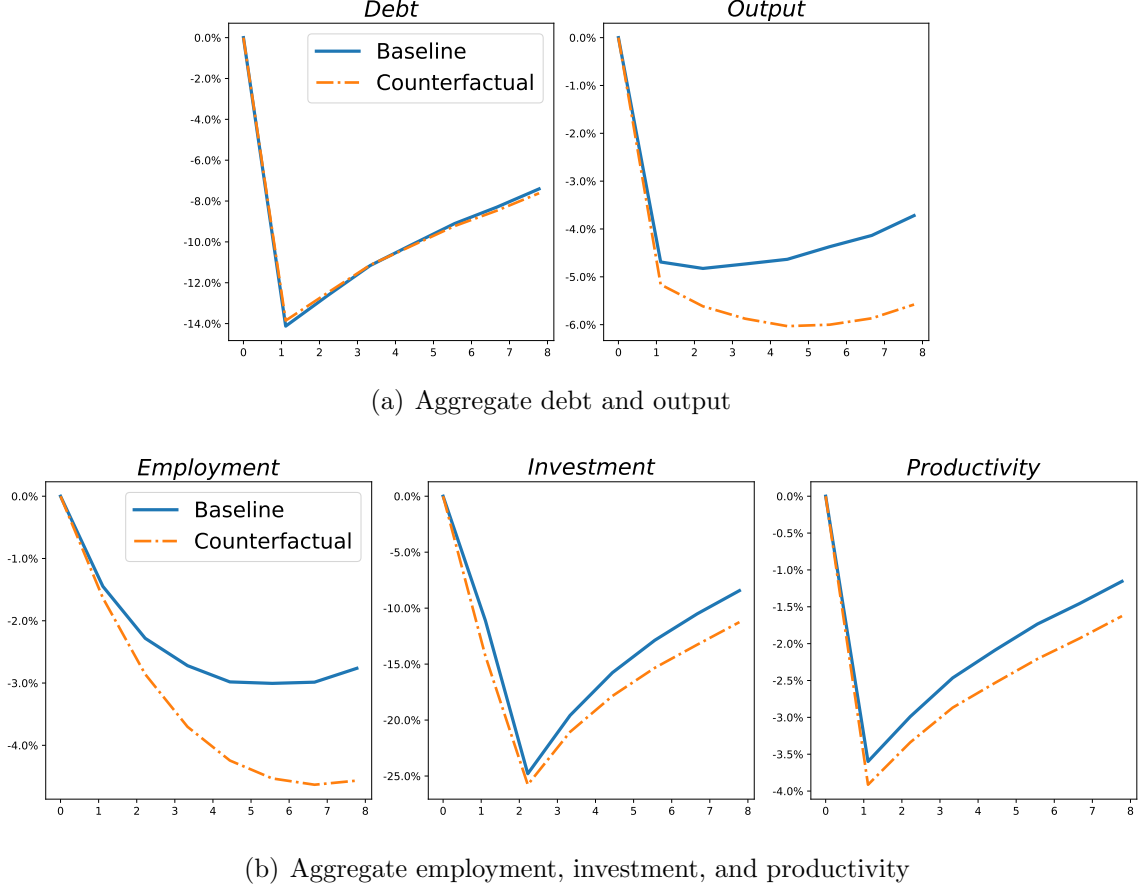


Figure 2: Impulse response functions for aggregate debt (D), output (Y), employment (N), investment (I) and productivity (A) in response to an aggregate financial shock. The solid line plots impulse response functions in the baseline economy, and the dashed line plots impulse response functions in the counterfactual economy with no commitment. We compute $2 \times M$ simulations of length T . We draw M sequences of uniform random numbers that we use to simulate realizations of ξ . In the first M simulations we set $\xi = \xi_L$ at $T - 10$. The IRFs are computed taking the difference in logs between the first and second set of simulations from $T - 10$ to T , averaging across M .

In Appendix C we consider a comparative static exercise where we increase the relative risk aversion coefficient of workers, thus decreasing firms' ability to adjust wages, and we find results similar to Figure 2.

5.4 Wage adjustments and stabilization policies

We now turn to study how effective are stabilization policies aimed at reducing input costs, in light of our findings that firms can optimally adjust the timing of wage payments over long-term employment relationships to reduce the cost of labor. We consider two broad types of policies: payroll subsidies and investment subsidies.

Payroll subsidies have been implemented by several OECD countries after the Great Recession in the form of payroll tax cuts. In some cases the tax cut was small and applied

to all employees, but more often it has been generous and it applied only to new hires. In the US the social security contribution for workers hired from unemployment has been set to zero as part of the HIRE Act, leading to a 6.2 percentage points cut. The contribution has been even more generous in some European countries, such as Ireland and Portugal, where the cost of employment contribution has been set to zero in 2010, leading to a 10 percentage points cut.⁴¹ These policies are often motivated by the presence of wage rigidity (Bils and Klenow, 2009) and targeted towards new hires based on the idea that the cost of incumbent workers is infra-marginal and thus not allocative. We focus on payroll subsidies for new hires, as these policies have been more used in practice and there is a large consensus on their potential positive effects.⁴²

Investment subsidies have often been implemented to foster recoveries in downturns. In practice, these subsidies are often implemented using accelerated depreciation schemes, that allow firms to deduct a large share of their investment from taxes immediately. Both in 2003 and in 2008 the United States introduced a 50 percent bonus depreciation, giving firms the possibility to immediately deduct 50 percent of investment purchases and then depreciate the remaining 50 percent under standard depreciation schedules. According to House and Shapiro (2008) this policy was equivalent to an investment subsidy ranging between 0.5% and 4.5%, depending on the recovery period of the investment good and the nominal interest rate. To facilitate the comparison between payroll subsidies and investment subsidies, we focus on investment subsidies targeted to newly created matches.

How effective are these policies when firms can optimally adjust the timing of wage payments over long-term employment relationships? To answer these questions we compare the effects of these policies between our model and the counterfactual economy described in Section 5.3 where firms cannot adjust wages over time. We show that a payroll subsidy on new hires is not as effective as it is the counterfactual economy because it crowds out the firms incentives to backload wages and its effectiveness is lower when firms already pay lower wages during recessions. On the other hand, an investment subsidy makes firms want to invest more, which spur firms' incentives to backload wages as the current marginal value of a dollar in the firm increases relative to the future. As a results, firms make wage tenure profiles steeper in response to the policy making more resources available for investment and amplifying the initial stimulus.

⁴¹See OECD (2010) for a detailed discussion of payroll tax cuts and hiring credit measures after 2008.

⁴²As illustrated by Schoefer (2021), in a model with financial constraints broad-based subsidies applied to incumbent workers too would increase hiring and investment because the cost of incumbent workers is not infra-marginal. Another rational for broad-based subsidies is to prevent firms from firing workers during downturns and prevent a spike in unemployment. As in our model separation is exogenous and matches are one-to-one, we focus on policies targeted to newly created matches.

5.4.1 Modeling payroll and investment subsidies

We introduce payroll and investment subsidies in the baseline model. The government sets payroll subsidies, investment subsidies, lump sum taxes, and government debt. The subsidies $\tau_N(\xi), \tau_I(\xi)$ depend on the realization of the aggregate shock, with $\tau_N(\xi_L) > 0, \tau_I(\xi_L) > 0$ and $\tau_N(\xi_H) = 0, \tau_I(\xi_H) = 0$, meaning that there is a temporary subsidy in recession and no taxes or subsidies in normal times.⁴³ While the government is restricted to running a balanced-budget fiscal policy in the long run, we allow for short-run debt-financed government expenditure. This means the government can raise funds using public debt and lump-sum taxes. Denote by G government expenditure on payroll and investment subsidies. The budget constraint of the government reads according to (30).⁴⁴ We parameterize the persistence of government debt by ρ_B . Given the law of motion of government debt (31), lump-sum taxes are set period by period to satisfy the budget constraint (30). We assume that lump-sum taxes are levied on entrepreneurs and employed workers.

$$qB' = G + T + B \quad (30)$$

$$B' = \rho_B(B + G) \quad (31)$$

The problem of matched entrepreneurs in an economy with subsidies is defined in equation (32). Because of investment subsidies the value of a matched entrepreneur now depends explicitly on the current capital stock k . Since we focus on policies targeted to newly created matches, the value of a matched entrepreneur depends on whether he is eligible for the subsidy ($e = 1$) or not ($e = 0$). Newly created matches are always eligible for the subsidy, and an entrepreneur remains eligible if $\xi' = \xi_L$. The budget constraint and the law of motion of net worth are modified as to account for the subsidies and for lump-sum transfers T . Both policies transfer resources to possibly constrained entrepreneurs by reducing the cost of inputs, and at the same time provide incentives for job creation and investment. More details on the economy with subsidies are provided in Appendix C.2.

⁴³An alternative way to perform this exercise is to assume there are payroll taxes in normal times, that is $\tau_N(\xi_H) < 0$, and a payroll tax cut, rather than a payroll subsidy, in recession. This exercise would be country-specific. We choose a more stylized approach to make a broad argument on the effects of such policies. A logic applies to investment subsidies.

⁴⁴We allow the government to issue public debt for two reasons. First, an increase in government expenditure during economic downturns is often funded with government debt and not by a contemporaneous increase in taxes. Second, the use of government debt to foster an economic recovery can play an important role during a credit tightening, when private credit markets are disrupted

$$J(m, W, z, k, e, S) = \max_{\substack{c^e, b', k', i, m'(z', \xi') \\ w'(z', \xi'), W'(z', \xi')}} \left\{ v(c^e) + \beta(1 - \phi) \mathbb{E} [J(m'(z', \xi'), W'(z', \xi'), z', k', e', S') | z, S] \right. \\ \left. + \beta \phi \mathbb{E} [V(m'(z', \xi'), z', S') | z, S] \right\} \quad (32)$$

$$(\text{Budget constraint : } \lambda^e) \quad c^e + i[1 - e\tau_I] \leq m - k(1 - \delta) + q(s_0)b'$$

$$(\text{Capital : } I) \quad k' \leq i + (1 - \delta)k$$

$$(\text{Net worth : } \eta(z', \xi')) \quad m'(z', \xi') \leq z'f(k') + (1 - \delta)k' - w'(z', \xi')[1 - e\tau_N] - b' + T'(\xi', S')$$

$$(\text{Collateral constraint : } \mu) \quad b' \leq \xi k'$$

$$(\text{Promise keeping : } \gamma) \quad W \leq \mathbb{E} \left[u(w'(z', \xi') + T'(\xi', S')) + \beta(1 - \phi)W'(z', \xi') + \beta \phi \mathcal{U}(S'') | z, S \right]$$

5.4.2 The effects of payroll subsidies

We consider first the effects of payroll subsidies. When the HIRE act was passed in the United States, the former Chief Economist at the Treasury Alan Krueger said that the HIRE act “provides an incentive for private-sector employers to hire new workers sooner than they otherwise would”. The ability of these policies to stimulate employment depends crucially on how they lower the cost of labor. The cost of labor in present discounted value terms, accounting for payroll subsidies is:

$$PDV_t = \mathbb{E} \left[\sum_{s=t}^{\infty} \underbrace{[1 - \tau(\xi_s)] \times w_{s+1}(z_{s+1}, \xi_{s+1})}_{\text{flow cost of labor}} \times \underbrace{[\beta(1 - \phi)]^{(s-t)} \frac{\eta_{s+1}(z_{s+1}, \xi_{s+1})}{v'(c_t^e)}}_{\text{SDF of entrepr.}} \right] \quad (33)$$

Crucially, temporary payroll subsidies lower the cost of labor for new hires according to the share of PDV_t that is paid during recession, that is as long as the subsidy is in place. As the stochastic discount factor of entrepreneurs increases substantially during a credit tightening, the share of PDV_t paid in recession is large. This implies that temporary payroll subsidies can have large effects on employment during a financial crises.⁴⁵ On the other hand, the share of the flow cost of labor that is paid in recession depends on the solution of the dynamic contracting problem. To highlight this channel we study the

⁴⁵This is in line with evidence from [Saez, Schoefer, and Seim \(2019\)](#), finding that payroll subsidies are more effective on financially constrained firms.

effects of payroll subsidies in the baseline economy and in the counterfactual economy with no commitment that we presented in Section 5.3.

In our model firms back-load wages of new hires after an aggregate financial shock. When wages are back-loaded the share of PDV_t that is paid in recession, that is as long as the subsidy is in place, is lower because the flow cost of labor is lower. Consequently, one should expect temporary payroll subsidies to be less effective in lowering the present discounted value of wage payments in the baseline model when compared to a counterfactual economy with no dynamic wage contracts.⁴⁶

Temporary payroll subsidies have also an effect on wages, as subsidies distort the risk-sharing condition between entrepreneurs and workers, as illustrated in equation (34). Indeed, the policy provides incentives to pay higher wages in recession because it is less costly to deliver utility to the workers when wages are subsidized. In this sense, in our model entrepreneurs adjust wages in response to a payroll subsidy to increase the flow cost of labor and reducing resources available for investment.

$$\eta(z', \xi')[1 - \tau(\xi')] = \gamma u'(w'(z', \xi')) \quad (34)$$

Table 6 reports the effects of a payroll subsidy on new hires equal to 6 percentage points, as the one introduced in 2010 in the United States as part of the HIRE Act. The effect on aggregate output is substantially lower in the economy with dynamic wage contracts, both on impact and cumulatively one year after the beginning of the recession. Payroll subsidies on new hires are not as effective as they would be in a model where firms cannot adjust wages over long-term employment relationships. Intuitively firms' optimal wage adjustments and payroll subsidies are substitute to each other, as they both aim to reduce the cost of labor.

5.4.3 The effects of investment subsidies

We now turn to studying the effects of investment subsidies. Evidence from House and Shapiro (2008), Zwick and Mahon (2017) show that firms respond strongly to these policies and that financial frictions can amplify the investment response. Intuitively, constrained firms value future cash flows with high effective discount rates, which amplify the perceived value of bonus incentives because the difference in today's tax benefits dwarfs the present value comparison that matters in frictionless models. We illustrate the differential effects of these policies on constrained and unconstrained firms using the optimality conditions implied by (32). To understand this heterogeneous response we focus on whether it affects the intertemporal consumption margin, and thus wages through the risk sharing condition.

⁴⁶Similarly, a consequence of wage back-loading is also that the increase of the stochastic discount factor of entrepreneurs during recession is less pronounced.

| Panel A: Payroll subsidies | | | | | | |
|-------------------------------|-------------------------------------|-------|--------|---|-------|---------|
| | Effect on impact: $\Delta \log X_t$ | | | Cumulative effect: $\sum_{s=0}^3 \Delta \log X_{t+s}$ | | |
| | Y | N | I | Y | N | I |
| Baseline model | +0.1 % | +0.1% | +1.6 % | +1.1 % | +1.3% | +5.9 % |
| Counterfactual | +0.2 % | +0.2% | +3.3 % | +1.4 % | +1.3% | +11.2 % |
| Panel B: Investment subsidies | | | | | | |
| | Effect on impact: $\Delta \log X_t$ | | | Cumulative effect: $\sum_{s=0}^3 \Delta \log X_{t+s}$ | | |
| | Y | N | I | Y | N | I |
| Baseline model | +1.3 % | +1.5% | +10 % | +6.3 % | +7.4% | +32 % |
| Counterfactual | +0.8 % | +0.8% | +4.2 % | +5.1 % | +5.7% | +17 % |

Table 6: Macroeconomic effects of a temporary payroll and investment subsidy on newly created matches. Panel A reports results for an economy with only payroll subsidies $\tau_L(\xi_L) = 0.06$, while Panel B reports results for an economy with only investment subsidies $\tau_I(\xi_L) = 0.016$. The table reports differences in output, employment and investment between the economy with subsidies and the economy without. Differences are reported on impact and cumulatively at one year horizon. The first line of each panel reports results for the baseline model, while the second line for the counterfactual model. We set $\rho_B = 0.9$.

When the constraint does not bind the intertemporal consumption decision is unaffected by the investment subsidy, and is characterized by the Euler equation for bonds in equation (35). An investment subsidy makes capital more attractive, and this affects the optimal allocation of resources between debt and capital for unconstrained entrepreneurs. However, as long as entrepreneurs are not financially constrained the subsidy does not affect the intertemporal consumption decision and the Euler equation for bonds (35) is identical to the in the model without taxes. Through the risk-sharing condition this means that also the intertemporal path of wages is not affected by the subsidy.

$$qv'(c^e) = \mathbb{E}[\eta(z', \xi')] \quad (35)$$

However, when the constraint binds the investment subsidy affects the intertemporal consumption decision. By combining the optimality conditions for bonds and capital we obtain equation (36). An investment subsidy makes capital more attractive, but since constrained entrepreneurs cannot substitute away from bonds to invest more in capital, as the choice of bonds is not interior, the only way they can invest more in capital is by reducing their current consumption. As a results, the investment subsidy affects the intertemporal consumption decision for constrained firms by increasing the current marginal value of a dollar relative to the future. As entrepreneurs make their own consumption profile steeper in response to the subsidy, the risk sharing condition implies that they also make the wage tenure profile of their workers steeper. In this

sense, in our model entrepreneurs adjust wages in response to an investment subsidy to free more resources for investment thus amplifying the initial stimulus.

$$v'(c^e)[1 - q\xi] < \mathbb{E}\left[\eta(z', \xi')(z'f(k') + 1 - \delta - \xi)\right] + \underbrace{\left(\tau v'(c^e) - (1 - \delta)\mathbb{E}[\eta(z', \xi')\tau(\xi')]\right)}_{>0} \quad (36)$$

Table 6 reports the effects of an investment subsidy equal to 1.6%, that according to House and Shapiro (2008) mimics the effects of a bonus depreciation of 50%.⁴⁷ The effect on aggregate output is substantially larger in the economy with dynamic wage contracts. Investment subsidies are more effective than they would be in a model where firms cannot adjust wages over long-term employment relationships. Intuitively firms' optimal wage adjustments and investment subsidies are complement to each other: when the policy reduce the cost of investment by making capital more attractive, firms respond to the policy by backloading wages in a way to free more resources to invest in capital thus amplifying the effects of the stimulus.⁴⁸

6 Conclusion

This paper explored the macroeconomic effects of financial shocks through a novel lens, highlighting the role of dynamic wage contracts within firms facing credit constraints. By integrating a simple dynamic contracting problem in a general equilibrium model with financial frictions and aggregate shocks, we illustrated how firms adjust wages over time and in response to shocks, depending on their financial conditions. We illustrated the mechanism within a theoretical framework and we provided empirical evidence on wage dynamics using matched employer-employee data from Italy.

Our findings show that financially constrained firms tend to back-load wages and exhibit higher wage adjustments in response to shocks, thereby enhancing liquidity for investment and job creation. The quantitative analysis revealed that dynamic wage contracts can substantially mitigate the adverse effects of financial shocks on aggregate output, employment, and allocative efficiency. Quantitatively, our model replicates the observed cross-sectional heterogeneity in wage dynamics, and it is at the same time consistent with evidence that the average wage is relatively stable over the cycle. We highlighted the policy implications of our findings, particularly regarding the effectiveness of payroll and investment subsidies during financial crises. Our analysis suggests

⁴⁷Calculations from House and Shapiro (2008) show that a 50% bonus depreciation implies a 1.6% investment subsidy for investment goods with a recovery period of 10 years given a nominal interest rate of 3%.

⁴⁸This mechanism is consistent with evidence from Garrett, Ohn, and Surez Serrato (2020) that did not find a positive effect of investment subsidies on earnings per-worker, while documenting positive effects on employment.

that incorporating the dynamic structure of wage contracts dampen the effects of payroll subsidies while enhancing the effects of investment subsidies.

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Appendix

A Model Appendix and Proofs

A.1 Proof of Proposition 1

First, we show that the measure of workers who search for a job is bounded from above, and that this bound does not depend on the measure of workers M . Then, we will take M such that there must be a positive measure of workers who do not search. Once we show that there is positive measure of workers who do not search, the other results will follow.

The problem of a vacant entrepreneur before matching and separation, defined in (3), can be written as

$$\hat{V}(m, z, S) = \max \left(\max_{\theta, W} \left\{ [\lambda_f(\theta)J(m, W, z, S) + (1 - \lambda_f(\theta))V(m, z, S)] \right\}, V(m, z, S) \right) \quad (37)$$

$$\text{s.t.} \quad \mathcal{W}(S) \leq \lambda_w(\theta)W + [1 - \lambda_w(\theta)] \mathbb{E}[\mathcal{U}(S') \mid S]$$

where the constraint can be re-arranged as

$$\frac{\mathcal{W}(S) - \mathbb{E}[\mathcal{U}(S') \mid S]}{\lambda_w(\theta)} \leq W - \mathbb{E}[\mathcal{U}(S') \mid S] \quad (38)$$

The optimality conditions of this problem imply

$$\theta) \quad \lambda'_f(\theta)(J(m, W, z, S) - V(m, z, S)) + \nu \frac{\mathcal{W}(S) - \mathbb{E}[\mathcal{U}(S') \mid S]}{\lambda_w^2(\theta)} \lambda'_w(\theta) = 0$$

$$W) \quad \lambda_f(\theta)J'_W(m, W, z, S) + \nu = 0$$

Combining the two FOCs, one obtains

$$\theta) \quad J(m, W, z, S) - V(m, z, S) = -J'_W(m, W, z, S) \frac{\mathcal{W}(S) - \mathbb{E}[\mathcal{U}(S') \mid S]}{\lambda_w} \left(\frac{1 - \eta}{\eta} \right)$$

where we have used properties of a Cobb-Douglas matching function⁴⁹.

This optimality condition and the constraint jointly determine (W, θ) given workers' values. Market tightness in sub-market (W, θ) is implicitly characterized by

⁴⁹If $\lambda_f = B\theta^{-\eta}$, then $\lambda'_f = -B\eta\theta^{-\eta-1} = -\eta B\lambda_f/\theta < 0$, and $\lambda_w = B\theta^{1-\eta}$ so $\lambda'_w = B(1 - \eta)\theta^{-\eta} = B(1 - \eta)\lambda_w/\theta$

$$\lambda_w(\theta) = \left(\frac{1-\eta}{\eta} \right) \frac{\mathcal{W}(S) - \mathbb{E}[\mathcal{U}(S') | S]}{J(m, W, z, S) - V(m, z, S)} \gamma(m, W, z, S)$$

Then, it must be that the total measure of workers who search, denoted by $s(S)$, is

$$\begin{aligned} s(S) &= \int s(\theta, W) \\ s(S) &= \int \underbrace{\left(\frac{1}{B} \right)^{-\frac{1}{1-\eta}} \left[\left(\frac{1-\eta}{\eta} \right) \frac{\mathcal{W}(S) - \mathbb{E}[\mathcal{U}(S') | S]}{J(m, W, z, S) - V(m, z, S)} \gamma(m, W, z, S) \right]^{-\frac{1}{1-\eta}}}_{\theta^{-1}} \mathbb{1}(m, W, z, S) d\Lambda^m(m, z) \end{aligned}$$

where B denotes matching efficiency, and $\mathbb{1}(m, W, z, S)$ is an indicator function equal to one if the entrepreneur opens a vacancy and zero otherwise.

Alternatively, one can write $s(S)$ using the participation constraint to obtain

$$s(S) = \int \underbrace{\left(\frac{1}{B} \right)^{-\frac{1}{1-\eta}} \left[\frac{\mathcal{W}(S) - \mathbb{E}[\mathcal{U}(S') | S]}{W - \mathbb{E}[\mathcal{U}(S') | S]} \right]^{-\frac{1}{1-\eta}}}_{\theta^{-1}} \mathbb{1}(m, W, z, S) d\Lambda^m(m, z) \quad (39)$$

First, note that if the ratio

$$\frac{W - \mathbb{E}[\mathcal{U}(S') | S]}{\mathcal{W}(S) - \mathbb{E}[\mathcal{U}(S') | S]} \quad (40)$$

is bounded in all sub-markets, than $s(S)$ must be bounded, as the measure of workers is 1.

We now show that the ratio (40) is bounded. If there is a positive measure of workers who search, we must have

$$\mathcal{W}(S) - \mathbb{E}[\mathcal{U}(S') | S] \geq u(\bar{b}) - u(\bar{b}(1-x))$$

so that the denominator of (40) is positive and bounded from below. Note that if there is not a positive measure of workers who search we simply have that $s(S) = 0$, and therefore $s(S)$ is still bounded. Moreover, since the productivity process for productivity follows a discrete Markov process with upper bound \bar{z} , it must be that there exists \bar{W} such that no entrepreneurs would ever find it profitable to open a vacancy with $W > \bar{W}$. Therefore, the numerator of (40) is bounded from above:

$$W - \mathbb{E}[\mathcal{U}(S') | S] \leq \bar{W} - \mathbb{E}[\mathcal{U}(S') | S] \leq \bar{W} - \frac{u(\bar{b})}{1-\beta}$$

where the last inequality follows from the fact that at least a positive measure of workers

is searching.

As we assumed that the production function is bounded from above by \bar{y} , for large values of capital, we have that \bar{W} is bounded from above

$$\bar{W} \leq \frac{u(\bar{y})}{1 - \beta}$$

as no entrepreneurs will be willing to offer a promised utility greater than \bar{W} . Note that in one could prove that the ergodic distribution of entrepreneurs' net worth is bounded from above, then W would be bounded without assuming that the production function is bounded.

As a result, the ratio in equation (40) is bounded, as the numerator is bounded from above and the denominator is bounded from below, and both must be positive. Note that the measure of employed workers is bounded above by one, that is the measure of entrepreneurs. Therefore, the measure of workers who search s is bounded as

$$s \leq \underbrace{\left(\frac{1}{B}\right)^{-\frac{1}{1-\eta}} \left[\frac{u(\bar{y})}{(1-\beta)h}\right]^{-\frac{1}{1-\eta}}}$$

which is a function of primitives \bar{y}, β, h, B . Then there exists a finite measure of workers M that satisfies

$$M > 1 + \underbrace{\left(\frac{1}{B}\right)^{-\frac{1}{1-\eta}} \left[\frac{u(\bar{y})}{(1-\beta)h}\right]^{-\frac{1}{1-\eta}}}_s$$

such that a positive measure of workers is not searching. If a positive measure of workers is not searching, this means that in equilibrium workers must be indifferent between searching and not searching. This implies

$$\mathcal{U}(S) = u(\bar{b}) + \beta \mathbb{E}[\mathcal{U}(S') \mid \mathcal{S}]$$

from which it follows that the value \mathcal{U} does not depend on S and it solves

$$\mathcal{U} = \frac{u(\bar{b})}{1 - \beta} \tag{41}$$

Moreover, workers being indifferent between searching and not searching also implies

$$u(\bar{b}) + \beta \mathbb{E}[\mathcal{U}(S') \mid \mathcal{S}] = u(\bar{b}(1-x)) + \beta \mathcal{W}(S)$$

and by combining it with equation (41) we get

$$\mathcal{W} = \mathcal{U} + \frac{u(\bar{b}) - u(\bar{b}(1-x))}{\beta}$$

Finally, note that the constraint in (47) does not depend on S anymore, and it simplifies to

$$\frac{u(\bar{b}) - u(\bar{b}(1-x))}{\lambda_w(\theta)\beta} \leq W - \mathcal{U}$$

A.2 Proof of Proposition 2

We guess that the consumption function takes the form

$$c_t = (1-x)m_t \tag{42}$$

with

$$x = \frac{\beta + \beta\gamma}{1 + \beta\gamma}$$

Combining (42) with the risk-sharing condition we obtain

$$w_t = \gamma(1-x)m_t \tag{43}$$

Combining (42), (43) with the law of motion of net worth we obtain

$$m_{t+1} = \frac{1}{1 + \gamma(1-x)} [k_{t+1}[z_{t+1} + 1 - \delta] - b_{t+1}]$$

We consider two cases: the collateral constraint does not bind, or the collateral constraint binds. We will show that in both cases the guess (42) is verified.

- Case 1: the collateral constraint does not bind.

Here we first guess and verify that the policy function for capital and debt take the form

$$k_{t+1} = \phi(z_t, \xi_t, q_t)xm_t, \quad b_{t+1} = -(1 - \phi(z_t, \xi_t, q_t))xm_t$$

where the function $\phi(z_t, \xi_t, q_t)$ crucially does not depend on net worth m_t .

Combining this guess with the law of motion of net worth we obtain

$$m_{t+1} = m_t \frac{x}{1 + \gamma(1-x)} [\phi(z_t, \xi_t, q_t)[z_{t+1} + 1 - \delta] + (1 + r_t)[1 - \phi(z_t, \xi_t, q_t)]] \tag{44}$$

where we denote $1 + r_t = \frac{1}{q_t}$.

We combine (44) with the Euler equation for bonds and capital and obtain

$$1 = \frac{\beta(1 + r_t)}{x} \mathbb{E} \left[\frac{1 + \gamma(1 - x)}{[\phi(z_t, \xi_t, q_t)[z_{t+1} + 1 - \delta] + (1 + r_t)[1 - \phi(z_t, \xi_t, q_t)]]} \right]$$

$$1 = \frac{\beta}{x} \mathbb{E} \left[\frac{[1 + \gamma(1 - x)](z_{t+1} + 1 - \delta)}{[\phi(z_t, \xi_t, q_t)[z_{t+1} + 1 - \delta] + (1 + r_t)[1 - \phi(z_t, \xi_t, q_t)]]} \right]$$

Combining the two Euler equation we are left with

$$\mathbb{E} \left[\frac{z_{t+1} - \delta - r_t}{\phi(z_t, \xi_t, q_t)[z_{t+1} - \delta - r_t] + 1 + r_t} \right] = 0$$

that provides an equation that implicitly characterizes $\phi(z_t, \xi_t, q_t)$ and also verifies the guess for the functional form of k_{t+1}, b_{t+1} .

Then, take a weighted average of the Euler equations for bonds and capital, with weights $\phi(z_t, \xi_t, q_t)$ and $1 - \phi(z_t, \xi_t, q_t)$, to obtain, after some manipulation:

$$1 = \beta \mathbb{E} \left[\frac{1 + \gamma(1 - x)}{x} \right]$$

which implies an equation for x :

$$x = \frac{\beta + \beta\gamma}{1 + \beta\gamma}$$

that does not depend on net worth.

Next we move to the case when the collateral constraint is binding, and we show that we obtain the same equation for x , that verifies our initial guess (42).

- Case 2: the collateral constraint binds.

In this case, the law of motion for net worth is

$$m_{t+1} = m_t \frac{1}{1 - q_t \xi_t} \frac{x}{1 + \gamma(1 - x)} [z_{t+1} + 1 - \delta - \xi_t] \quad (45)$$

Combining the Euler equation for bonds and capital to substitute for the multiplier on the collateral constraint we obtain

$$v'(c_t^e)[1 - q_t \xi_t] = \beta \mathbb{E} [v'(c_{t+1}^e)[z_{t+1} + 1 - \delta \xi]] \quad (46)$$

Combining (45) with (46) we obtain:

$$1 = \beta \mathbb{E} \left[\frac{1 + \gamma(1 - x)}{x} \right]$$

which implies an equation for x :

$$x = \frac{\beta + \beta\gamma}{1 + \beta\gamma}$$

which verifies the initial guess from (42)

A.3 Proof of Proposition 3

Using Proposition 2, entrepreneurs leverage can be expressed as:

$$\frac{b_t}{k_t} = 1 - \frac{1}{\phi(z_t, \xi_t, q_t)}$$

Using Proposition 2 the pass-through of idiosyncratic productivity shocks to wages can be expressed as

$$\frac{\partial [\log(w_{t+1}) - \log(w_t)]}{\partial [\log(z_{t+1}) - \log(z_t)]} = \frac{\phi(z_t, \xi_t, q_t) z_{t+1}}{\phi(z_t, \xi_t, q_t) [z_{t+1} - \delta] + (1 - \phi(z_t, \xi_t, q_t z_{t+1})) \frac{1}{q_t}}$$

from which we find that in the cross-section the pass-through of idiosyncratic shocks to wages is increasing in ϕ . As leverage is increasing in ϕ , then the pass-through is increasing in leverage.

Using Proposition 2 the average growth rate of wages can be expressed as:

$$\mathbb{E}[\log(w_{t+1}) - \log(w_t)] = \mathbb{E} \left[\frac{x}{1 + \gamma(1 - x)} \phi(z_t, \xi_t, q_t) [z_{t+1} + 1 - \delta] + [1 - \phi(z_t, \xi_t, q_t)] \frac{1}{q_t} \right]$$

From which we obtain, after some manipulation, that

$$\frac{\partial \mathbb{E}[\log(w_{t+1}) - \log(w_t)]}{\partial \phi} = \log \left(\frac{1}{\beta} \right) > 0$$

that implies the average growth rate of wages is increasing in entrepreneurs' leverage.

A.4 Proof of Lemma 1

Let denote by γ_0 the multiplier on the constraint defined in (23). The first order conditions of the problem defined in (23) imply:

$$\eta_{s+1}(z^{s+1}, \xi^{s+1}) = \frac{\gamma_0}{v'(c_t^e)} u'(w(z^{s+1}, \xi^{s+1})), \quad \forall s, z^{s+1}, \xi^{s+1}$$

Taking the ratio of the optimality conditions for two different histories we obtain a risk-sharing condition similar to the optimality condition of the recursive problem defined in

(2).

$$\frac{\eta_{s+1}(z^{s+1}, \xi^{s+1})}{\eta_{p+1}(z^{p+1}, \xi^{p+1})} = \frac{u'(w(z^{s+1}, \xi^{s+1}))}{u'(w(z^{p+1}, \xi^{p+1}))}, \quad \forall s, p, z^{s+1}, \xi^{s+1}, z^{p+1}, \xi^{p+1}$$

Given the same promised utility W , this implies that the optimal contract that solves (2) is also a solution to the problem defined in (23). To see that, start from $s = t$: as we have $z^{t+1} = z_{t+1}$, $\xi^{t+1} = \xi_{t+1}$ given z_t, ξ_t , the optimal wage contract implied by (2) trivially satisfies the first order conditions of (23). Similarly, one can use the intertemporal dimension of the risk-sharing condition to check that the optimal wage contract implied by (2) satisfies the optimality condition of (23) for $s > t$.

A.5 Optimal job creation of vacant entrepreneurs

Using Proposition 1, the problem of a vacant entrepreneur before matching and separation, defined in (3), can be written as

$$\widehat{V}(m, z, S) = \max \left(\max_{\theta, W} \left\{ [\lambda_f(\theta)J(m, W, z, S) + (1 - \lambda_f(\theta))V(m, z, S)] \right\}, V(m, z, S) \right) \quad (47)$$

$$\text{s.t.} \quad \frac{u(\bar{b}) - u(\bar{b}(1 - x))}{\beta \lambda_w(\theta)} = W - \mathcal{U}$$

The first order conditions for an interior solution to this problem imply

$$\theta) \quad \lambda'_f(\theta)(J(m, W, z, S) - V(m, z, S)) + \nu \frac{u(\bar{b}) - u(\bar{b}(1 - x))}{\beta \lambda_w^2(\theta)} \lambda'_w(\theta) = 0$$

$$W) \quad \lambda_f(\theta)J'_W(m, W, z, S) + \nu = 0$$

where ν denotes the multiplier on the constraint. Combining the two first order conditions, and replacing the constraint in the first order condition for θ , we obtain:

$$\theta) \quad J(m, W, z, S) - V(m, z, S) = -J'_W(m, W, z, S) [W - \mathcal{U}] \left(\frac{1 - \eta}{\eta} \right)$$

where we have used properties of a Cobb-Douglas matching function⁵⁰.

⁵⁰If $\lambda_f = B\theta^{-\eta}$, then $\lambda'_f = -B\eta\theta^{-\eta-1} = -\eta B\lambda_f/\theta < 0$, and $\lambda_w = B\theta^{1-\eta}$ so $\lambda'_w = B(1 - \eta)\theta^{-\eta} = B(1 - \eta)\lambda_w/\theta$

A.6 Walras' Law

If the market for risk-free bonds clears, then the resource constraint holds. First, note that the budget constraint of each matched entrepreneur j can be written as:

$$c_j^e + k_j' \leq y_j + (1 - \delta)k_j - w_j - b_j + qb_j'$$

The budget constraint of each vacant entrepreneur i can be written as:

$$c_i^e \leq -b_i + qb_i'$$

We can integrate the two budget constraints over the measure of matched entrepreneurs j and vacant entrepreneurs i ; and sum the integrated budget constraint to obtain:

$$C^e + K' + W \leq Y + (1 - \delta)K - B + qB'$$

where W denotes the sum of all wages paid to workers. Using the market clearing condition for risk-free bonds, that is $B = 0, B' = 0$, and the identity $W = C^w$ as workers are hand-to-mouth, to obtain

$$C^e + K' + C^w \leq Y + (1 - \delta)K$$

that is, the resource constraint holds

A.7 Law of motion Γ

The law of motion Γ for the aggregate state S is made of

- An exogenous law of motion for the aggregate shock ξ
- An endogenous law of motion H^m for the distribution $\Lambda^m(m, W, z)$:

$$H^m(\Lambda^m)(\mathcal{M}, \mathcal{W}, \Xi) = \int Q_{\Lambda^m}((m, W, z), \mathcal{M}, \mathcal{W}, \Xi) d\Lambda^m(m, W, z)$$

$$Q_{\Lambda^m}((m, W, z), \mathcal{M}, \mathcal{W}, \Xi) = \sum_{z' \in \Xi} \begin{cases} \pi(z' | z) & \text{if } m'(m, W, z; S) \in \mathcal{M}, W'(m, W, z; S) \in \mathcal{W} \\ 0 & \text{otherwise} \end{cases}$$

- An endogenous law of motion H^v for the distribution $\Lambda^v(m, z)$:

$$H^v(\Lambda^v)(\mathcal{M}, \Xi) = \int Q_{\Lambda^v}((m, z), \mathcal{M}, \Xi) d\Lambda^v(m, z)$$

$$Q_{\Lambda^v}((m, z), \mathcal{M}, \Xi) = \sum_{z' \in \Xi} \begin{cases} \pi(z' | z) & \text{if } m'(m, z; S) \in \mathcal{M} \\ 0 & \text{otherwise} \end{cases}$$

A.8 Average Job Finding Rate

The average job finding rate in the economy is

$$\begin{aligned}
 Average(\lambda_w(\theta)) &= \frac{\int \lambda_w(\theta_j) s_j dj}{\int s_j dj} \\
 Average(\lambda_w(\theta)) &= \frac{\int \lambda_w(\theta) \theta(m, W, z, S)^{-1} \underbrace{\mathbb{1}(m, W, z, S) d\Lambda^v(m, z)}_v}{\int \theta(m, W, z, S)^{-1} \underbrace{\mathbb{1}(m, W, z, S) d\Lambda^v(m, z)}_v}
 \end{aligned}$$

The model solution algorithm easily returns $\lambda_f(\theta) = B\theta^{-\eta}$, from which we can get

$$\theta = \left(\frac{\lambda_f(\theta)}{B} \right)^{-\frac{1}{\eta}}, \quad \lambda_w(\theta) = B\theta^{1-\eta}$$

$$Average(\lambda_w(\theta)) = \frac{\int \lambda_f(\theta) \underbrace{\mathbb{1}(m, W, z, S) d\Lambda^v(m, z)}_v}{\int \theta(m, W, z, S)^{-1} \underbrace{\mathbb{1}(m, W, z, S) d\Lambda^v(m, z)}_v}$$

B Data Appendix

B.1 Matched employer-employee data

In this section, we provide further details on the matched employer-employee data used for our analysis. We restrict our analysis to firms that participated in the Bank of Italy’s annual Survey of Industrial and Service Firms (INVIND). Each year the survey gathers information on investments, gross sales, workforce and other economic variables relating to Italian industrial and service firms with 20 or more employees. More precisely, the survey cover firms operating in the following industries: “Food and beverages”, “Textiles and apparel”, “Chemical, pharmaceutical, rubber”, “Non-metallic minerals”, “Metalworking industry”, “Wood, paper, furniture”, “Water and waste”, “Wholesale/retail trade”, “Hotels and restaurants”, “Transportation and telecommunication”, “Other (real estate etc.)”.

The National Social Security Institute (*Istituto Nazionale di Previdenza Sociale*, INPS) was asked to provide the complete works histories of all workers that ever transited in an INVIND firm. While the data included spells of employment in which workers were employed at firms not listed in the INVIND survey (e.g. they were employed at an INVIND firm in 2010, but then they changed job in 2012), we restrict our attention to INVIND firms for which we have information on the entire population of employees.

B.1.1 Cleaning procedure and sample construction

We start from the balance sheets data of all firms available in Cerved from 2005 to 2019. These are 1,526,216 firms in total. We restrict our sample to these firms that took part of the INVIND survey between 2005 and 2019 (i.e. that have been surveyed by the Bank of Italy and for which we have access to the entire work histories of all their employees). This step restricts our sample to 9,698 firms.

Then, we perform some cleaning of balance sheet data, as to remove extreme values and clearly implausible entries. We drop firms that have negative entries for value added in at least one year between 2005 and 2019. These are 1,949. We are left with 7,749 firms. We exclude 484 firms with intermittent participation in Cerved and drop 233 firms that appear in the data for less than 3 years (notice that the IV estimation of pass-through coefficients requires at least 3 years of data). We are left with 7032 firms.

For the estimation of pass-through coefficient we run a simple regression to remove persistent heterogeneity across firms, as well as aggregate and sector-specific trends:

$$\log(VA_{it}/L_{it}) = \alpha_i + 2\text{digsector-timeFE} + \varepsilon_{it} \quad (48)$$

and get the first difference of residuals. We further exclude 601 firms with residual changes

in VA per worker greater than 1 and 661 with changes smaller than -1. This leaves us with 6129 firms. Then, we drop firms for which Cerved declares that debt-related information is not reliable. Cerved classify some firms with non-reliable debt information whenever these firms are not required by the law to report detailed information on their debt in their balance sheets. This step leads us to remove 756 firms. Finally, we also clean from outliers in leverage, as there are firms with values greater than 100 or negative leverage. This leaves us with 5107 firms.

Then, we move to worker-level data. For each worker that has ever transited by an INVIND firm, we have access to their entire work history. We select a 25% random sample of workers from the sample. These are 2,521,206 workers. By merging the balance sheet data and the worker-level data, we find that basically all firms in our sample have at least one worker in the 25% sample. Once we focus on the sub-sample of firms that we obtained from the cleaning procedure we are left with 1,153,746 workers. We exclude 208 workers who have some duplicate record in the dataset. A duplicate record is defined by the combination of earnings \times type of contract \times number of weeks \times which months she worked at the same firm in the same year. This leaves us with 1,153,538. As our empirical analysis is focused on stayers, we only keep workers who in a given year have only one employer, and work 52 weeks. By considering workers who only have one employer in one year, we drop 97,782 workers. This leaves us with 1,055,756 workers. We drop workers who only have part-time contracts during our observation window. This is because this workers may be more affected by an hours response during a value added shock. This leads us to drop 252,719 workers. This leaves us with 803,037 workers.

B.2 Aggregate data

In Section 5 we use time-series data for several macro-economic aggregates. The data come mainly from OECD (variable codes in parenthesis), and we retrieved them from FRED, Federal Reserve Bank of St. Louis.

- Corporate debt: we measure corporate debt as the amount outstanding of total debt securities issued by corporations in the non-financial sector, including all maturities, such that the residence of the issuer is in Italy ('TDSAMRIAONCIT'). Data are available at quarterly frequency.
- Output: we measure aggregate output as real gross domestic product ('CLVMNAC-SCAB1GQIT') at quarterly frequency.
- Unemployment: we measure it as the unemployment rate ('LRUN64TTITQ156S') for people aged 15-64. Data are available at quarterly frequency.
- Employment: we measure the number of employed people aged 15-64 ('LFEM64TTITQ647S'). Data are available at quarterly frequency.

- Hours: we measure hours as the average annual hours worked by employed persons ('AVHWPEITA065NRUG'). Data are available at annual frequency, and we use linear interpolation to construct a measure of hours at quarterly frequency.
- Labor: we measure total labor inputs as the product of employment and hours.
- Investment: we measure investment as grossed fixed capital formation ('ITAGFCFQD-SNAQ'). Data are available at quarterly frequency.

B.3 Sample selection: mobility probit

We provide additional details for the sample selection correction that we discussed in Section 4.3. In practice, we estimate a probit regression where on the left-hand side there is a dummy equal to one if the completed tenure $CT_{ij(t_0)}$ of worker i at firm j hired in t_0 is greater or equal than S (i.e. the worker is included in the sub-sample of stayers). The variable $Displaced_{ij(t_0)}$ is equal to 1 if the current job started after a mass layoff or a firm closure at the previous employer. We include in $X_{ij(t_0)}$ the same set of controls that we include in our main specification.

$$Pr(CT_{ij(t_0)} \geq S) = \Phi\left(\beta_0 + \beta_1 Displaced_{ij(t_0)} + \beta_2 1(lev_{jt_0-1} > \text{median}) + \beta_3 X_{ij(t_0)} + \varepsilon_{ij(t_0)}\right) \quad (49)$$

We report detailed estimates of equation (49) in Table (7). We discussed the estimates for the coefficient on displaced workers in Section 4.3. Other coefficients also have the expected sign. Workers employed by more productive firms –higher VA/worker– are more likely to stay longer in the firm. Similarly, younger workers, female and managers are more like to have shorter tenure.

To control for sample selection we include the inverse Mills ratio implied by equation (49) as a control in the main regression by interacting it with tenure.

$$\begin{aligned} (\log w_{ij(t_0)t} - \log w_{ij(t_0)t_0}) = & \sum_{s=1}^S \tilde{\beta}_s 1(T_{ij(t_0)t} = s) \\ & + \sum_{s=1}^S \tilde{\gamma}_s 1(T_{ij(t_0)t} = s) 1(lev_{jt_0-1} > \text{median}) \\ & + \sum_{s=1}^S \tilde{\delta}_s 1(T_{ij(t_0)t} = s) X'_{ijt_0} \\ & + \sum_{s=1}^S \tilde{\iota}_s 1(T_{ij(t_0)t} = s) \frac{\phi_{ij(t_0)}}{\Phi_{ij(t_0)}} + (u_{ij(t_0)t} - u_{ij(t_0)t_0}) \end{aligned} \quad (50)$$

| Variable | Coefficient | Std. Err. |
|--------------------------------------|-------------|-----------|
| Displaced | -0.1987*** | 0.0320 |
| 1 ($lev_{jt_0-1} > \text{median}$) | 0.0874*** | 0.0087 |
| log Assets | 0.0104*** | 0.0023 |
| log VA/worker | 0.3526*** | 0.0047 |
| <i>Year of birth</i> | | |
| 1960 | 0.0209 | 0.0233 |
| 1961 | -0.0112 | 0.0234 |
| 1962 | 0.0055 | 0.0239 |
| 1963 | -0.0053 | 0.0250 |
| 1964 | -0.0468* | 0.0244 |
| 1965 | 0.0688** | 0.0248 |
| 1966 | -0.0477* | 0.0248 |
| 1967 | -0.0393 | 0.0247 |
| 1968 | -0.0751*** | 0.0243 |
| 1969 | -0.0874*** | 0.0239 |
| 1970 | -0.1164*** | 0.0236 |
| 1971 | -0.1171*** | 0.0233 |
| 1972 | -0.2058*** | 0.0228 |
| 1973 | -0.2271*** | 0.0227 |
| 1974 | -0.2720*** | 0.0222 |
| 1975 | -0.3394*** | 0.0224 |
| 1976 | -0.3463*** | 0.0228 |
| 1977 | -0.4444*** | 0.0227 |
| 1978 | -0.4180*** | 0.0228 |
| 1979 | -0.4352*** | 0.0231 |
| 1980 | -0.4540*** | 0.0238 |
| <i>Occupation</i> | | |
| Blue Collar | 1.1960*** | 0.0565 |
| High Skilled Blue Collar | 1.4174*** | 0.1019 |
| White Collar | 1.2094*** | 0.0563 |
| Middle Manager | 1.2667*** | 0.0590 |
| Manager | 0.2413*** | 0.0702 |
| Others | 0.0413** | 0.0921 |
| Female | -1.2361*** | 0.0748 |
| Sectors FE | Yes | |
| <i>N</i> | 168,184 | |

Table 7: Estimated coefficients and standard errors for the mobility probit.

B.4 Robustness: wage backloading and interest coverage

Most of the empirical analysis presented in Section 4 focuses on heterogeneity in wage dynamics across firms with different leverage ratios. This approach is motivated both by our model presented in Section 2, where firms borrowing capacity is restricted by the collateral value of their capital, and by evidence that bank credit is largely collateralized in the Italian market (Garrido, Kopp, and Weber, 2016).

In practice, there might be other mechanisms that make firms more or less financially constrained independently of the collateral value of their assets. A common type of debt covenants –provisions in debt contracts that constrain future lending – is related to interest coverage (Greenwald, 2019). In practice these covenants impose a cap on the ratio of a firms interest payments to its earnings or EBITDA.

We show that the empirical results presented in Section 4.3 are robust to classifying firms as more or less financially constrained based on the ratio of their interests payments to their earnings. We propose a measure of interest coverage (“IC”) consistent with the structure of these debt covenants, that is defined as the ratio of firms’ financial expenses to EBITDA. Our measure of financial expenses is the sum of: interest paid on bank loans, bonds, and other types of financing, interest paid on overdraft accounts, costs of issuing debt instruments, bank fees and service charges related to loans.

$$\begin{aligned}
 (\log w_{ij(t_0)t} - \log w_{ij(t_0)t_0}) = & \sum_{s=1}^S \tilde{\beta}_s 1(T_{ij(t_0)t} = s) \\
 & + \sum_{s=1}^S \tilde{\gamma}_s 1(T_{ij(t_0)t} = s) 1(IC_{jt_0-1} > \text{median}) \\
 & + \sum_{s=1}^S \tilde{\delta}_s 1(T_{ij(t_0)t} = s) X'_{ijt_0} + (u_{ij(t_0)t} - u_{ij(t_0)t_0}) \quad (51)
 \end{aligned}$$

We estimate equation (51), that is we modify equation (25) from Section 4.3 in order to estimate different wage-tenure profiles for workers employed by firms whose IC ratio at $t_0 - 1$ is above or below the median. We report estimates in Table 8. Estimates for the mobility probit are reported in Column 1. Estimates for the coefficients $\{\tilde{\gamma}_s\}$ are reported in Columns 2 –without controlling for endogenous mobility– and 3 –controlling for endogenous mobility.

The estimated coefficients are qualitatively similar to those presented in Section 4.4: wages grow faster over time at more financially constrained firms. The magnitudes of the estimated coefficients are also similar to those reported in Table 1 for the baseline specification over the first two years of tenure.

| Wage growth and interest coverage | | | |
|--|-----------------------|------------------------|-----------------------|
| | Probit (1) | Cumulative wage growth | |
| | | (2) | (3) |
| $1(IC_{jt_0-1} > \text{median})$ | 0.0364*** (0.0123) | | |
| Tenure = 1 : 1 ($IC_{jt_0-1} > \text{median}$) | | 0.0184*** (0.0019) | 0.0158*** (0.0018) |
| Tenure = 2 : 1 ($IC_{jt_0-1} > \text{median}$) | | 0.0261*** (0.0023) | 0.0221*** (0.0022) |
| Tenure = 3 : 1 ($IC_{jt_0-1} > \text{median}$) | | 0.0296*** (0.0024) | 0.0240*** (0.0023) |
| Tenure = 4 : 1 ($IC_{jt_0-1} > \text{median}$) | | 0.0258*** (0.0026) | 0.0186*** (0.0025) |
| Displaced | -0.1550*** (.0676) | | |
| Tenure=1: Inverse Mills ratio | | | -0.0047 (0.0096) |
| Tenure=2: Inverse Mills ratio | | | 0.0051 (0.0116) |
| Tenure=3: Inverse Mills ratio | | | 0.0185 (0.0121) |
| Tenure=4: Inverse Mills ratio | | | 0.0165 (0.0129) |
| Workers' controls | Yes | Yes | Yes |
| Firms' controls | Yes | Yes | Yes |
| N. of workers | 130,775 | 100,680 | 100,680 |
| N | 130,775 | 402,720 | 402,720 |

Table 8: Column (1) reports the results of the probit estimates of whether completed tenure of each worker-firm match is above or below S , using the entire sample of workers hired at t_0 . Columns (2) and (3) report the estimates of equation (51) using the sample of workers who stayed at least 4 years. Robust standard errors in parenthesis. The superscripts ***, **, and * denote statistical significance at the 1, 5, and 10 percent level, respectively.

B.5 Robustness: wage backloading and overtime hours

One might be concerned that the estimates of equation (25) presented in Section 4.3 could be influenced by adjustments in hours worked. Indeed, we only observe average monthly earnings and not the base wage. Even if we focus on full-time workers employed year-round in all the empirical exercises, firms may adjust over-time hours for these workers. We exploit the nature of our data that includes information from the INVIND survey on average over-time hours at the firm level. For each firm in the sample we have information on the ratio between over-time hours and total hours in any given year.

We estimate a firm-level regression according to equation (52) similar to the main specification for wage growth in equation (25). On the left-hand side H_{jt} is the share of over-time hours at firm j in year t . On the left-hand side the coefficients $\{\tilde{\gamma}_s\}$ measures the difference in the share of over-time hours in year $t_0 + s$ between firms with high leverage in $t_0 - 1$ and firms with low leverage in $t_0 - 1$. We control firms' size and firms' value-added per worker by including them in the vector of controls X_{jt_0} , thus using the same firm-level controls that we used to estimate equation (25).

$$H_{jt} = \sum_{s=1}^S \tilde{\beta}_s 1(t = s) + \sum_{s=1}^S \tilde{\gamma}_s 1(t = s) 1(lev_{jt_0-1} > \text{median}) + \sum_{s=1}^S \tilde{\delta}_s 1(t = s) X_{jt_0} + (u_{ij(t_0)t} - u_{ij(t_0)t_0}) \quad (52)$$

We test whether there are any significant differences in the share of over-time hours between firms with different leverage starting from $t_0 = 2009$. We report point estimates and confidence intervals for the coefficients $\{\tilde{\gamma}_s\}$ in Figure 3. We find that differences in over-time hours over time are small and not statistically significant. To interpret the magnitudes, note that a coefficient of 0.2 means that on average firms with leverage above median have a share of over-time hours 0.2% larger than firms with leverage below the median. Moreover, while differences in wages increase over time we find that differences in over-time hours don't seem to increase over time. To corroborate these findings, in Figure 4 we report the estimates of $\{\tilde{\gamma}_s\}$ using debt-to-output ratio rather than leverage to construct the indicator function on the right-hand side of equation (52). Also in this case we do not find substantial differences in the share of over-time hours between firms with debt-to-output ratio above and below median.

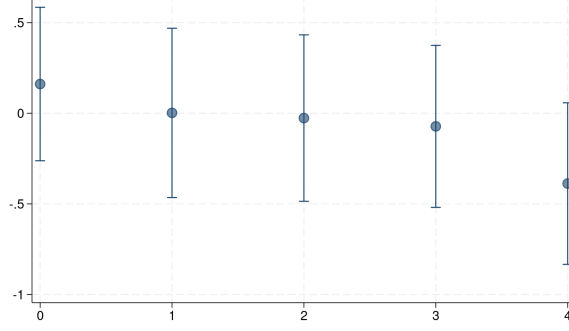


Figure 3: We report estimates of the coefficients $\{\tilde{\gamma}_s\}$ from equation (52) and the 95% confidence intervals.

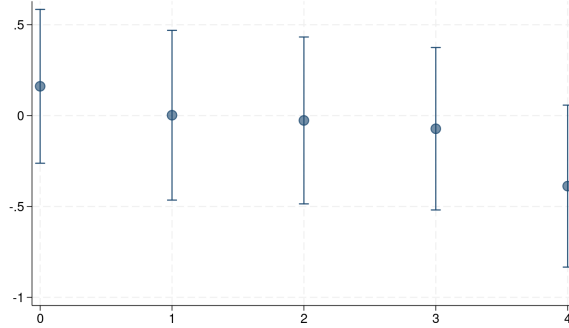


Figure 4: We report estimates of the coefficients $\{\tilde{\gamma}_s\}$ from equation (52) and the 95% confidence intervals using the debt-to-output ratio rather than the leverage ratio to construct the indicator function on the right-hand side of equation (52).

B.6 Robustness: indexes of local credit supply

We construct province-time index of credit supply following the methodology proposed in [Greenstone, Mas, and Nguyen \(2020\)](#), using the same data and procedure as in [Barone, de Blasio, and Mocetti \(2018\)](#). Provinces are the natural geographical unit for this exercise, as According to the Italian Antitrust authority the “relevant market” in banking for antitrust purposes is the province ([Guiso, Pistaferri, and Schivardi, 2012](#)).

To construct this province-time index we use confidential data drawn from the Bank of Italy Supervisory Report database. The data provides information on total outstanding loans –including credit lines, credit receivable and fixed-term loans– extended by Italian banks to the private sector (firms and households) in each province. The data allows us to distinguish between outstanding loans extended to firms or households. Using this database we describe the loan growth rate at the local level and we build a data-driven indicator of credit supply. Since we are interested in constructing these indexes to identify more financially constrained firms, we focus on firms’ loans.

Specifically, for each province p we estimate the following equation that separates the contribution of bank-specific fluctuations and province-specific fluctuations to bank lending.

$$\Delta \ln(L_{bkt}) = \alpha + \delta_{bt} + \gamma_{kt} + \varepsilon_{bkt}$$

On the left-hand side, the outcome variable is the percentage change in outstanding loans by bank b in province p between the years t and $t - 1$. On the right-hand side, γ_{kt} is a set of province-year fixed effects that aim to capture the variation in the change of lending due to purely local economic factors that are common to all banks. The coefficients δ_{bt} capture bank-year fixed effects and represent our parameters of interest, since they capture nationwide bank lending policies that are purged of province-specific loan demand (and of any other province-year level idiosyncratic shock).

$$S_{pt} = \sum_b W_{bp} \hat{\delta}_{bt} \quad (53)$$

We denote by S_{pt} our measure of credit supply in province p in year t . This measure is the weighted average of bank- b -specific changes in outstanding loans δ_{bt} between year t and $t - 1$, with the market share W_{bp} of each bank b in province p used as weights. We construct different indexes of credit supply using market shares of bank b in province p for different pre-crises years, namely 1999, 2002, and 2006.

We estimate equation (54), that is we modify equation (25) from Section 4.3 to estimate different wage-tenure profiles for workers employed by firms operating in provinces that experienced a more pronounced drop in credit supply. In this sense, we identify as more financially constrained firm j that operates in province p is the index of credit supply at $t_0 - 1$ is below the average of that same year.

$$\begin{aligned} (\log w_{ijp(t_0)t} - \log w_{ijp(t_0)t_0}) = & \sum_{s=1}^S \tilde{\beta}_s 1(T_{ijp(t_0)t} = s) \\ & + \sum_{s=1}^S \tilde{\gamma}_s 1(T_{ijp(t_0)t} = s) 1(S_{pt_0-1} < \text{mean}) \\ & + \sum_{s=1}^S \tilde{\delta}_s 1(T_{ijp(t_0)t} = s) X'_{ijpt_0} + (u_{ijp(t_0)t} - u_{ijp(t_0)t_0}) \end{aligned} \quad (54)$$

We report estimates of the mobility probit and equation (54) in Table 9, using banks' market shares in 2006 to construct the weights W_{bp} . Note that the estimated coefficients $\{\tilde{\gamma}_s\}$ reported in Column 3 are the same we reported in Table 2.

| Wage growth and local credit supply shocks | | | |
|--|------------------------------|----------------------------|----------------------------|
| | Probit (1) | Cumulative wage growth | |
| | | (2) | (3) |
| $1(S_{pt_0-1} < \text{mean})$ | -0.0842^{***} (0.0141) | | |
| Tenure=1: $1(S_{pt_0-1} < \text{mean})$ | | 0.0247^{***} (0.0021) | 0.0205^{***} (0.0022) |
| Tenure=2: $1(S_{pt_0-1} < \text{mean})$ | | 0.0384^{***} (0.0025) | 0.0346^{***} (0.0027) |
| Tenure=3: $1(S_{pt_0-1} < \text{mean})$ | | 0.0337^{***} (0.0027) | 0.0303^{***} (0.0028) |
| Tenure=4: $1(S_{pt_0-1} < \text{mean})$ | | 0.0312^{***} (0.0029) | 0.0310^{***} (0.0030) |
| Displaced | -0.1435^{***} (.0376) | | |
| Tenure=1: Inverse Mills ratio | | | 0.0975^{***} (0.0162) |
| Tenure=2: Inverse Mills ratio | | | 0.0883^{***} (0.0191) |
| Tenure=3: Inverse Mills ratio | | | 0.0753^{**} (0.0201) |
| Tenure=4: Inverse Mills ratio | | | 0.0179 (0.0217) |
| Workers' controls | Yes | Yes | Yes |
| Firms' controls | Yes | Yes | Yes |
| N. of workers | 127,229 | 98,750 | 98,750 |
| N | 127,229 | 395,000 | 395,000 |

Table 9: Column (1) reports the results of the probit estimates of whether completed tenure of each worker-firm match is above or below S , using the entire sample of workers hired at t_0 . Columns (2) and (3) report the estimates of equation (54) using the sample of workers who stayed at least 4 years. Robust standard errors in parenthesis. The superscripts ***, **, and * denote statistical significance at the 1, 5, and 10 percent level, respectively.

| Robustness: wage growth and local credit supply shocks | | | | |
|--|------------------------|-----------------------|------------------------|-----------------------|
| | Including province p | | Excluding province p | |
| | 1999 | 2002 | 1999 | 2002 |
| | (1) | (2) | (3) | (4) |
| Tenure=1: $1(S_{pt_0-1} < \text{mean})$ | 0.0175*** (0.0022) | 0.0204*** (0.0021) | 0.0161*** (0.0021) | 0.0141*** (0.0021) |
| Tenure=2: $1(S_{pt_0-1} < \text{mean})$ | 0.0300*** (0.0027) | 0.0350*** (0.0026) | 0.0290*** (0.0025) | 0.0287*** (0.0026) |
| Tenure=3: $1(S_{pt_0-1} < \text{mean})$ | 0.0285*** (0.0028) | 0.0302*** (0.0027) | 0.0234*** (0.0027) | 0.0199*** (0.0027) |
| Tenure=4: $1(S_{pt_0-1} < \text{mean})$ | 0.0279*** (0.0030) | 0.0301*** (0.0029) | 0.0254*** (0.0029) | 0.0203*** (0.0029) |
| Inverse Mills ratio | Yes | Yes | Yes | Yes |
| Workers' controls | Yes | Yes | Yes | Yes |
| Firms' controls | Yes | Yes | Yes | Yes |
| N | 395,000 | 395,000 | 395,000 | 395,000 |

Table 10: Estimates of the coefficients $\tilde{\gamma}_s$ from equation (54) obtained using alternative methods to construct indexes of credit supply S_{pt} . Robust standard errors in parenthesis. The superscripts ***, **, and * denote statistical significance at the 1, 5, and 10 percent level, respectively.

One concern associated with this approach has to do with large provinces having an effect on nationwide lending policies of certain banks. For instance, this may occur when a province market is sufficiently large with respect to the national credit market of a certain bank (e.g. small banks are typically geographically concentrated in few provinces) and, therefore, it may affect its lending policy. As a result, it would be hard to plausibly isolate changes in credit supply from S_{pt} . We address this concern as in [Barone, de Blasio, and Mocetti \(2018\)](#) by estimating bank-fixed effects δ_{bt} using all provinces but p . In this way, we isolate the component of bank-specific fluctuations in credit that is not driven by province-specific effects. In [Table 10](#) we report estimates of the coefficients $\{\tilde{\gamma}_s\}$ from equation (54) using alternative methods to construct indexes of credit supply S_{pt} . In Columns 1 and 2 we report estimates obtained using different banks' market shares in different years –1999 and 2002– to construct indexes of local credit supply. In Columns 3 and 4 we repeat the same exercise while estimating bank-fixed effects δ_{bt} using all provinces but p . We find that results are robust to these alternative methods to construct indexes of credit supply.

We consider firms that operate in provinces p with an index of credit supply conditions S_{pt} below the median as more financially constrained.

B.7 Robustness: pass-through coefficients

We present several robustness exercises to further corroborate the results presented in Section 4.4 on how the pass-through coefficient of value-added per worker varies with firms' financial conditions.

In Section 4.4 we illustrated how firms with higher leverage pass-through more of changes in value-added per worker to wages. This approach is motivated both by our model presented in Section 2, where firms borrowing capacity is restricted by the collateral value of their capital, and by evidence that bank credit is largely collateralized in the Italian market (Garrido, Kopp, and Weber, 2016).

Here we consider alternative ways to classify firms as more or less financially constrained. In practice, there might be other mechanisms that make firms more or less financially constrained independently of the collateral value of their assets. For instance, firms' borrowing capacity can be limited by their earnings as opposed to their assets (Drechsel, 2023; Lian and Ma, 2020). Also, a common type of debt covenants –provisions in debt contracts that constrain future lending – is related to interest coverage (Greenwald, 2019). In practice these covenants impose a cap on the ratio of a firms interest payments to its earnings or EBITDA. Therefore, firms' borrowing capacity can be limited by the ratio between interest payments and earnings.

In this Section we illustrate that estimates presented in Section 4.4 are robust to using alternative methods to classify firms as financially constrained. First, we consider an alternative way to construct the leverage ratio b/k , where the denominator includes only financial debt, excluding debt to suppliers. Results are reported in Column 1 of Table 11. Second, we consider firms with a high debt-to-output ratio to be more financially constrained, consistent with a view of earnings-based constraints as opposed to collateral constraints (Drechsel, 2023). Results are reported in Column 2 of Table 11. Third, we consider firms with a interest coverage ratio to be more financially constrained, consistent with a view of interest coverage covenants (Greenwald, 2019). Results are reported in Column 3 of Table 11. Fourth, as we explained in Section 4.3, one might be concerned that leverage at $t_0 - 1$ may be correlated with firms permanent unobservable characteristics that affect wage growth. For instance, some firms may have better access to credit and persistently sustain high leverage without being financially constrained. To address this concern we consider firms with leverage above their own average to be more financially constrained. Results are reported in Column 3 of Table 11.

B.8 Wage back-loading and investment

This section provides descriptive evidence supporting the model mechanism by exploiting differences in wage flexibility across firms. We show that firms with more flexibility in wage setting back-loaded wages more during the Great Recession while experiencing a

| Robustness: pass-through coefficients | | | | |
|--|--------------------------|--------------------------------|-----------------------------|--------------------------|
| | Leverage ratio (1) | Debt-to-output ratio (2) | Interest coverage (3) | Leverage ratio (4) |
| $\Delta\varepsilon_{jt} \times$ less constrained | 0.045*** (0.001) | 0.041*** (0.001) | 0.049*** (0.001) | 0.046*** (0.001) |
| $\Delta\varepsilon_{jt} \times$ more constrained | 0.068*** (0.001) | 0.070*** (0.001) | 0.066*** (0.001) | 0.118*** (0.001) |
| Within firm variation | No | No | No | Yes |
| N | 11,052,040 | 11,052,040 | 11,052,040 | 11,052,040 |

Table 11: We split the sample and report estimates of pass-through coefficients for firms that we label as constrained and unconstrained. We use alternative methods to classify firms as more or less financially constrained. Robust standard errors in parentheses. The superscripts ***, **, and * denote statistical significance at the 1, 5, and 10 percent level, respectively.

lower drop in investment.

The INVIND survey provides information on whether firms in our sample signed any “second level contract”. In the Italian labor market, “second-level contracts” (also known as decentralized bargaining agreements) refer to collective agreements negotiated at a level below the national sectoral agreements –typically at the company or regional level. These contracts complement the national collective agreements (known as first-level contracts), which set the baseline conditions for wages, hours, and benefits across an entire sector. These contracts enhance wage flexibility by allowing firms to adjust wages based on specific performance or productivity metrics, as well as regional economic conditions. First, we estimate equation (24) separately for firms that did and did not have any “second level contract” in place during 2008. We plot estimates of γ_s in Figure 5 for these two groups of firms. The sample differs from the one used in Section 4.3 as the information on second level contracts is not available for all firms in our sample. Estimates from (24) show that firms with high leverage and a second level contract during 2008 back-loaded wages of newly hired workers substantially more than highly levered firms without a second level contract during 2008. In other words, financially constrained firms with greater flexibility in wage setting offered steeper wage-tenure profiles to workers hired during the Great Recession.

Then, we turn to study the investment dynamics of these different groups of firms during the Great Recession. We estimate a triple-difference specification, as described in equation (55), to measure the differential response of investment during the Great

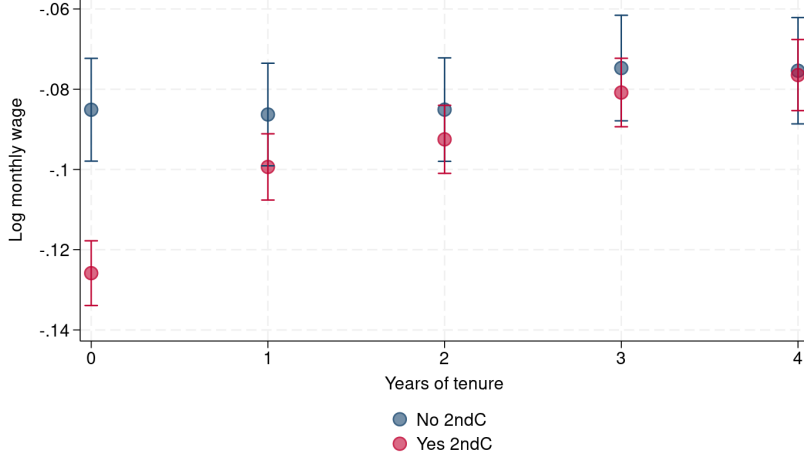


Figure 5: We plot estimates for the coefficients $\{\gamma_s\}_{s=0}^4$ in equation (24), setting $t_0 = 2009$ for firms that had “second level contract” (red) and firms who did not (in blue).

Recession for financially constrained firms with and without second level contracts.

$$\begin{aligned} \log(i/k)_{jt} = & \beta_0 + \sum_s \alpha_s year_s 1(lev_{jt_0-1} > \text{median}) + \sum_s \beta_s year_s 1(2ndC_{jt_0}) \\ & \sum_s \pi_s year_s 1(2ndC_{jt_0}) 1(lev_{jt_0-1} > \text{median}) + \sum_s \delta_s X'_j + u_{jt} \end{aligned} \quad (55)$$

On the left-hand side of (55) there is the logarithm of investment rate of firm j at time t . On the left-hand side, $2ndC_{jt_0}$ is a dummy equal to one if firm j has a second level contract at time t_0 . The coefficient π_s measures the difference in $\log(i/k)$ between firms with leverage above median with a second level contract and firms with leverage above median without a second level contract. We control for sector-year fixed effects by including firms’ sectors in X'_j , as the dynamics of investment rates over time can vary substantially across sectors.

Estimates of the coefficients π_s are plotted in Figure 6. While there is no significant difference between investment rates of firms with and without second level contracts before the Great Recession, estimates of the coefficients π_s during the Great Recession are positive and statistically significant. These estimates imply that highly levered firms with second level contracts experienced a less pronounced drop in investment during the Great Recession compared to highly levered firms without second level contracts.

This descriptive evidence provides additional empirical support for the model mechanism, as firms with greater flexibility in wage setting back-loaded wages more during the Great Recession and experienced a less pronounced drop in investment. These results are consistent with the model mechanism illustrated in Section 3, namely that wage back-loading frees resources for investment. In the next section we use our quantitative model to study the macro-economic implications of dynamic wage contracts during financial

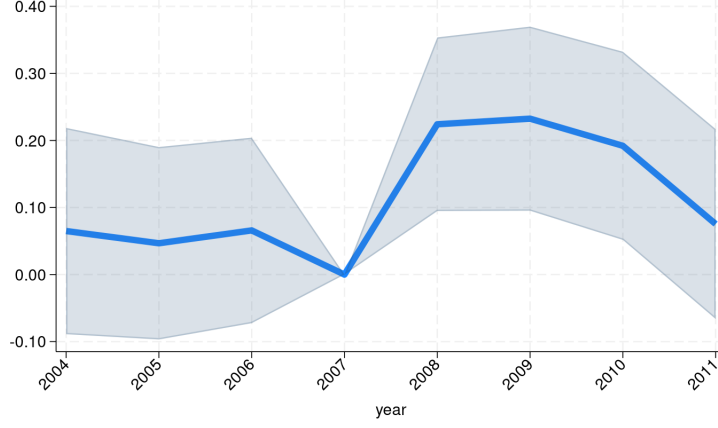


Figure 6: The blue line interpolates estimates for the coefficients $\{\pi_s\}_{s=0}^4$ in equation (55), setting $t_0 = 2009$. The gray shaded area corresponds to the 95% confidence intervals.

crises.

B.9 Estimation of productivity process

Let consider the following Error Correction Model (ECM)

$$\begin{aligned}
 y_{jt} &= B_j + z_{jt} + \nu_{jt} \\
 z_{jt} &= \rho z_{jt-1} + \varepsilon_{jt} \\
 v &\sim (0, \sigma_v), \quad \varepsilon \sim (0, \sigma_\varepsilon), \quad B \sim (0, \sigma_B)
 \end{aligned} \tag{56}$$

where log-productivity y_{jt} is the sum of a firm-specific component B_j , a persistent component z_{jt} and a purely idiosyncratic component ν_{jt} . Using panel data and assumptions to recover firms' productivity from balance sheet data, one can estimate the parameters of the stochastic process described in (56) using a generalized method of moments estimator. Indeed, the parameters are over-identified from auto-covariances of firms' productivity y_{jt} at different horizons, as in standard income processes (Heathcote, Perri, and Violante, 2010).

The goal of this exercise is to recover estimates of ρ and σ_ε , that are the parameters disciplining the stochastic process for idiosyncratic productivity in our model. At the same time, this exercise allows us to isolate variation in productivity driven by a persistent components from cross-sectional variation in firms' productivity driven by fixed heterogeneity and purely idiosyncratic shocks, as these features are not part of our model. Once we have a proxy for firms' log-productivity y_{jt} , the stochastic process specified in equation (56) is identified from panel data and can be estimated using a generalized method of moments estimator.

We use data for a balanced panel of firms between 2007 and 2018. Consistently with

the production function in our model we measure firms' productivity as

$$y_{jt} = \log(\text{VA/worker})_{jt} - \alpha \log(k)_{jt}$$

We use fixed assets at book value to measure k_{jt} . We report also estimates obtained using total assets at book value to measure k_{jt} , as well estimates in which we simply measure y_{jt} as the log of value-added per worker. Estimates are reported in Table 12. In all these cases we obtain similar estimates. We obtain estimates of the parameters disciplining the stochastic process at quarterly frequencies as:

$$\sigma_{\varepsilon, \text{quarter}}^2 = \frac{\sigma_{\varepsilon, \text{annual}}^2}{4}$$

$$\rho_{\text{quarter}} = \rho_{\text{annual}}^{1/4}$$

| | (1) | (2) | (3) |
|------------------------|--------------|--------------|-----------|
| | Fixed assets | Total assets | VA/worker |
| ρ | 0.86 | 0.86 | 0.88 |
| σ_{ε} | 0.16 | 0.13 | 0.13 |
| Quarterly frequencies | | | |
| | Fixed assets | Total assets | VA/worker |
| ρ | 0.96 | 0.96 | 0.96 |
| σ_{ε} | 0.08 | 0.07 | 0.07 |

Table 12: Estimates of $\rho, \sigma_{\varepsilon}$ obtained using different specifications. The first panel reports estimates at annual frequency, as balance sheet data. The second panel reports the implied estimates at quarterly frequency.

C Quantitative details and additional results

C.1 Numerical Algorithm

We solve the model using standard projection methods and we approximate the law of motion of the aggregate state S following [Krusell and Smith \(1998\)](#). Because of [Proposition 1](#), the problem of entrepreneurs depends on the aggregate state S only through the realization of the aggregate shock ξ and the price of bonds q . While the law of motion for ξ is exogenous, we rely on the following approximation to characterize the law of motion of q as:

$$q_{t+1} = \beta_0(\xi_t, \xi_{t+1}) + \beta_1(\xi_t)q_t \quad (57)$$

that is defined by the six coefficients: $\beta_0(\xi_L, \xi_L), \beta_0(\xi_L, \xi_H), \beta_0(\xi_H, \xi_L), \beta_0(\xi_H, \xi_H), \beta_1(\xi_H), \beta_1(\xi_L)$. In other words, agents forecast future prices conditional on future realizations of the aggregate shock ξ_{t+1} using information on current prices and the current value of the aggregate shock ξ_t .

In this sense, we summarize the aggregate state in period $t + 1$ as (ξ_{t+1}, ξ_t, q_t) , where q_t depends on the history of previous shocks ξ according to [equation \(57\)](#). We start the algorithm with an initial guess for the coefficients in [\(57\)](#). As there is not an explicit characterization for q as a function of a one-dimensional aggregate state variable, we follow [Krusell and Smith \(1997\)](#) and we include the price of bonds q as a state variable in the entrepreneurs' decision problem. This means that in a first step we solve for entrepreneurs' decision problem as a function of q , and then we solve for the market clearing price period by period using entrepreneurs' decision rules during simulation.

We simplify the computation by solving the problem of matched entrepreneurs defined in [\(2\)](#) using the multiplier γ rather than the promised utility W as a state variable. The advantage from doing so is that, according to the optimality conditions for $W'(z', \xi')$, the multiplier γ has to be constant throughout the length of a match. Moreover, it's also easier to solve for the optimal policy for state contingent wages $w'(z', \xi')$ using the risk-sharing condition given the multiplier γ .

This way, one can solve for the value functions without explicitly characterizing the state-contingent future promised values $W'(z', \xi')$. More formally, one can define the Pareto problem $P(m, \gamma, z, S)$ as

$$P(m, \gamma, z) = \max_W \left[J(m, W, z) + \gamma W \right]$$

and one can easily show that the policy functions that solve the program defined in [\(58\)](#)

are a solution to (2).

$$P(m, \gamma, z, S) = \max_{\substack{c^m, b', k', \\ m'(z', \xi'), w(z', \xi')}} \left\{ v(c^m) + \beta \phi \mathbb{E}[V(m'(z', \xi'), z', S') | z, S] + \gamma \mathbb{E}[u(w'(z', \xi'))] \right. \\ \left. + \beta(1 - \phi) \left\{ \mathbb{E}[P(m', \gamma, z', S') | z, S] \right\} \right\} \quad (58)$$

$$(\text{Budget constraint} : \lambda^e) \quad c^m + k' \leq m + qb'$$

$$(\text{Net worth} : \eta(z', \xi')) \quad m'(z', \xi') \leq z'f(k') + (1 - \delta)k' - w'(z', \xi') - b'$$

$$(\text{Collateral constraint} : \mu) \quad b' \leq \xi k'$$

We solve the problem of matched entrepreneurs using grids for the state variables (m, γ, z, ξ, q) , and the problem of vacant entrepreneurs with using grids for the states (m, z, ξ, q) . We use a GPU to iterate over policy functions and value functions, given the relatively large number of states. Similarly to [Menzio and Shi \(2011\)](#), we impose $\lambda_f(\theta) = \min(1, B\theta^{-\eta})$. Note that the participation constraint of workers searching in sub-market (θ, W) always implies $\lambda_w(\theta) \in (0, 1)$ ⁵¹. We initialize the algorithm with an initial guess $W(m, \gamma, z, \xi, q)$ for the promised utility as a function of the states, and we update this guess at each iteration using the wage policy and the previous guess of W .

Then, at each step we combine the value function of matched entrepreneurs obtained as a function of (m, γ, z, ξ, q) with the corresponding promised utility $W(m, \gamma, z, \xi, q)$ to obtain policy functions and value functions that depends on the states (m, W, z, ξ, q) . We then use these as inputs to update the value functions and the policy functions of vacant entrepreneurs.

Once we have solved for the policy functions and the value functions, we simulate the model by approximating the distribution of idiosyncratic states $\Lambda^m(m, W, z)$, $\Lambda^v(m, z)$ on a grid. We solve for the market clearing price of risk-free bonds q period-by-period during simulation, that is we solve for q such that

$$\sum_{m, W, z} b'(m, W, z, \xi, q) \times \Lambda^m(m, W, z) = \sum_{m, z} a'(m, z, \xi, q) \times \Lambda^v(m, z)$$

We simulate the economy for T periods, we drop the first T_0 observations, and we use the simulated series of prices q to update the coefficients of the forecasting rule (57). We keep iterating until the root mean squared error obtained from using the initial guess for

⁵¹We calibrated and solved an alternative version of our model using the matching function proposed by [den Haan, Ramey, and Watson \(2000\)](#), that implies job finding rates and vacancy filling rates always below one. We find very similar results.

(57) to predict q is small enough, that is when the agents forecasting rule is accurate and consistent with the time series of prices. We stop the algorithm when the R^2 from the forecasting regression (57) on newly simulated data is greater than 0.999.

When we solve the model with payroll subsidies we also include government debt as an aggregate state variable in agents' decision problem. Entrepreneurs also need to forecast future values of public debt. Assuming that government debt follows an AR(1) with persistency ρ_B greatly simplifies the forecasting problem.

C.2 Model with subsidies: details

The problem of matched entrepreneurs in the economy with payroll and investment subsidies have been described in (32). Compared to the problem of matched entrepreneur in the baseline model, there are two additional idiosyncratic state variables (k, e) . The state variable e is equal to one if the entrepreneur is eligible for the subsidies, and zero otherwise. Its law of motion is given by

$$\begin{aligned} e' &= e & \text{if } \xi' = \xi_L \\ e' &= 0 & \text{if } \xi' = \xi_H \end{aligned}$$

meaning that eligible entrepreneurs remain eligible during recessions, and all matched entrepreneurs become ineligible at the end of a recession.

Vacant entrepreneurs face a discrete choice problem between posting a vacancy or not. Entrepreneurs that decide to open a vacancy have to choose a sub-market (θ, W) where to open it. Their problem is described in (59). If a vacant entrepreneur is matched to a worker, he obtains the continuation value $J(m, W, z, 0, 1, S)$, that is the value of being a matched entrepreneur with 0 capital and eligible for subsidies ($e = 1$).

$$\widehat{V}(m, z, S) = \max \left(\max_{(\theta, W)} \left\{ [\lambda_f(\theta)J(m, W, z, S) + (1 - \lambda_f(\theta))V(m, z, S)] \right\}, V(m, z, S) \right) \quad (59)$$

After matching and separation, vacant entrepreneurs decide how much to consume and how much to save, according to (60). The main difference compared to the model

with no subsidies is that lump sum transfers enter the law of motion of net worth.

$$V(m, z, S) = \max_{a', c^e, m'} \left\{ v(c^e) + \beta \mathbb{E} \left[\widehat{V}(m', z', S') \mid z, S \right] \right\} \quad (60)$$

$$(\text{Budget constraint}) : \quad c^e + qa' \leq m$$

$$(\text{Net worth}) : \quad m' \leq a' + \bar{b} + T'$$

The decision problem of workers is unchanged from the baseline model described in Section 2.

The market clearing condition for risk-free bonds is now:

$$\int b'(m, W, z, k, e, S) d\Lambda^m(m, W, z, k, e) + B' = \int a'(m, z, S) d\Lambda^v(m, z)$$

Government expenditure G on payroll subsidies and investment subsidies is equal to

$$G_t = \int \mathbb{E} [w(m_{t-1}, W_{t-1}, z_{t-1}, k_{t-1}, e_{t-1}, S_{t-1}, z_t, S_t) \mid z_{t-1}] \times \tau_N(\xi_t) d\Lambda_{t-1}^m(m_{t-1}, W_{t-1}, z_{t-1}, k_{t-1}, 1) \\ + \int i(m_t, W_t, z_t, k_t, e_t, S_t) \times \tau_I(\xi_t) d\Lambda_t^m(m_t, W_t, z_t, k_t, 1)$$

C.3 Additional counterfactual: changing workers' risk aversion

Wage contracts in this model solve a risk-sharing problem between entrepreneurs and workers. Section 3 highlighted that the magnitude of wage adjustments depend on the ratio between the risk-aversion coefficient of workers and entrepreneurs σ_W/σ_E . When this ratio is high, wage adjustments must be low and entrepreneurs have to bear more risk. The structure of wage contract affects entrepreneurs' investment and hiring decisions, as illustrated in Section 3, and the magnitudes are quantitatively relevant, as shown in Section 5. Here we use the quantitative model to further illustrate the role played by wage adjustments in shaping the dynamics of investment, employment and output after an aggregate financial shock. We compare the calibrated model to an economy where workers have a higher coefficient of relative risk aversion. Figure 7 plots the impulse response functions to a drop in ξ in the baseline model ($\sigma_W = 11$, solid line) and in a model with a higher value for the workers' relative risk aversion coefficient ($\sigma_W = 22$, dashed line). We re-calibrated the value of ξ_L in the counterfactual economy with higher σ_W as to obtain the same drop in aggregate debt. While the two economies experience the same drop in aggregate debt, output falls more in the economy with higher σ_W , that is when firms adjust wages less. The results are qualitatively similar to those presented in Figure 2 from Section 5: output falls more in response to an aggregate financial shock when we constrain firms' ability to back-load wage payments and to adjust wages in

response to shocks.

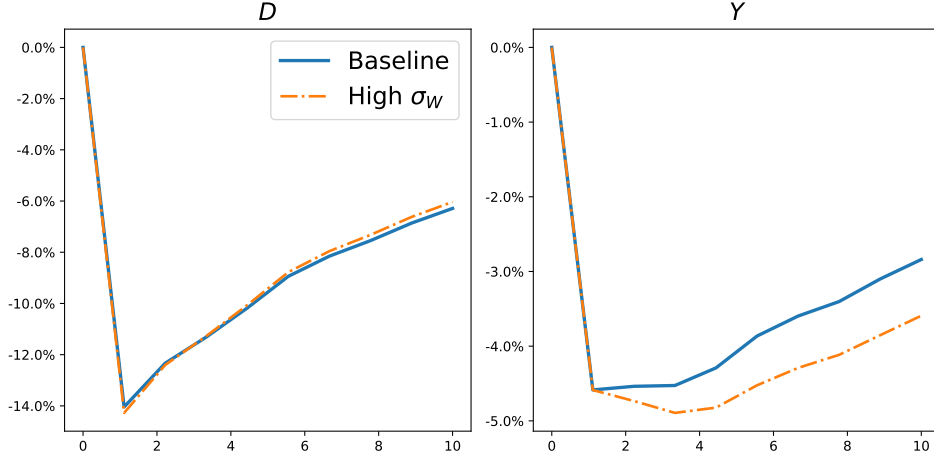


Figure 7: We compute $2 \times M$ simulations of length T . We draw M sequences of uniform random numbers that we use to simulate realizations of ξ . In the first M simulations we set $\xi = \xi_L$ at $T - 10$. The IRFs are computed taking the difference in logs between the first and second set of simulations from $T - 10$ to T , averaging across M .

C.4 Wage cyclicality

The measurement of wage cyclicality remains a significant area of research, with findings being highly sensitive to the specific definition of "wage" employed (Basu and House, 2016) and the availability of granular data on worker compensation (Grigsby, Hurst, and Yildirmaz, 2021). Nevertheless, there is broad consensus regarding well established stylized facts.

In this section we emphasize that the quantitative model is consistent with these stylized facts on wage cyclicality. First, in our model the average wage moves little during recession (Grigsby, 2022). Second, the user cost of labor is more cyclical than the average wage (Kudlyak, 2014; Basu and House, 2016). Third, the wage of new hires is a cyclical as the wage of incumbent worker (Gertler, Huckfeldt, and Trigari, 2020; Grigsby, Hurst, and Yildirmaz, 2021).

We present results from a simulated panel of entrepreneurs and workers in our model. We estimate equation (61) using different measures of wages on the left-hand side.

$$\log w_{ijt} = \beta_0 + \beta_1(\text{Recession}_t = 1) + u_{ijt} \quad (61)$$

We construct a measure of the risk-neutral user cost of labor similar to the one proposed by Kudlyak (2014); Basu and House (2016). We call this the *risk-neutral* user cost of labor as it does not reflect the true user cost of labor in our model, where the present discounted value of wages should be discounted with the entrepreneurs' specific stochastic discount factor. We define the risk-neutral user cost of labor UC_t^{RN} in equation (62) as

| | Wage | New hire wage | User cost |
|-------------------|-------|---------------|-----------|
| | (1) | (2) | (3) |
| $Recession_t = 1$ | 0.003 | 0.003 | 0.008 |

Table 13: Estimates of equation (62) using different measures of log-wage on the left-hand side. Estimates are obtained from a simulated panel of 20,000 entrepreneurs for 2000 periods (500 years).

the difference between the average present discounted value of wages for a match created at t minus the average present discounted value of wages for a match created at $t + 1$.

$$UC_t^{RN} = E_t[PDV_{jt}^{RN}] - \beta(1 - \phi)E_{t+1}[PDV_{jt+1}^{RN}] \quad (62)$$

To construct a measure of this present discounted value PDV_{jt}^{RN} for each firm j we use the present discounted value of all wages paid over the employment relationship discounted using $\beta(1 - \phi)$ as a discount factor and truncating the summation after seven years as in Kudlyak (2014).

We report estimates of equation (61) in Table 13. The average wage falls by only 0.3% in an average recession, meaning that the average wage is only mildly procyclical in our model. Moreover, by comparing Columns 1 and 2 the average wage of new hires is a cyclical as the average wage of incumbents workers (Gertler and Karadi, 2011; Grigsby, Hurst, and Yildirmaz, 2021). Finally, we find that drop in the average risk-neutral user cost of labor is almost three times larger than the drop in the average wage, consistent with evidence that the user cost of labor is substantially more cyclical than the average wage (Kudlyak, 2014; Basu and House, 2016).

D Model extensions

D.1 Model with Epstein-Zin Preferences

The baseline model assumes that workers have standard CRRA preferences, and in Section 3 we show that the key properties of the optimal wage contracts between entrepreneurs and workers depend on the ratio between their relative risk aversion coefficients. As it is well known, with CRRA preferences the relative risk aversion coefficient is also the inverse of the elasticity of intertemporal substitution (EIS). In principle, both the risk-aversion coefficient and the EIS of workers should matter. If workers have a large coefficient of RRA, we should expect them to be less inclined to accept wage contracts that are sensitive to idiosyncratic shocks. On the other hand, if workers have small EIS, we should expect them to be less inclined to accept wage payments that vary over time. In this section we derive a version of the baseline model where workers have Epstein-Zin preferences, that we use to better highlight the role played by risk and intertemporal substitution. In order to simplify the exposition we illustrate the model when Proposition 1 holds.

The value of a matched worker with promised utility W at the end of period (i.e. in the afternoon) is defined recursively as

$$W \leq E \left[\left\{ (1 - \beta)w'(z', \xi')^{1-\rho} + \beta \left[(1 - \phi)W'(z', \xi') + \phi \mathcal{U}^{1-RRA} \right]^{\frac{1-\rho}{1-RRA}} \right\}^{\frac{1-RRA}{1-\rho}} \right]$$

Note that this definition is equivalent to define the value of a matched worker with promised utility W in the morning, *before* wages are paid, as

$$W \leq \left\{ (1 - \beta)w^{1-\rho} + \beta E_t \left[(1 - \phi)W'(z', S')^{1-RRA} + \phi \mathcal{U}^{1-RRA} \right]^{\frac{1-\rho}{1-RRA}} \right\}^{\frac{1}{1-\rho}}$$

that is more similar to the standard timing used with these preferences.

The problem of matched entrepreneurs in the afternoon is identical to (2), but with a modified version of the promise keeping constraint.

$$J(m, W, z, S) = \max_{\substack{c^m, b', k', m'(z', \xi'), \\ w'(z', \xi'), W'(z', \xi')}} \left\{ v(c^m) + \beta(1 - \phi) \underbrace{\mathbb{E} [J(m'(z', \xi'), W'(z', \xi'), z', S') | z, S]}_{\text{not separate}} \right. \\ \left. + \beta \phi \underbrace{\mathbb{E} [V(m'(z', \xi'), z', S') | z, S]}_{\text{separate}} \right\}$$

$$(\text{Budget constraint : } \lambda^e) \quad c^m + k' \leq m + qb'$$

$$(\text{Net worth : } \eta(z', \xi')) \quad m'(z', \xi') \leq y(k', z') + (1 - \delta)k' - w'(z', \xi') - b'$$

$$(\text{Collateral constraint : } \mu) \quad b' \leq \xi(S)k'$$

$$(\text{Promise keeping : } \gamma) \quad W \leq E \left[\left\{ (1 - \beta)w'(z', \xi')^{1-\rho} + \beta \left[(1 - \phi)W'(z', \xi') + \phi \mathcal{U}^{1-RRA} \right]^{\frac{1-\rho}{1-RRA}} \right\}^{\frac{1-RRA}{1-\rho}} \right]$$

The optimality conditions for state-contingent wages and promised values are

$$\begin{aligned} \eta(z', \xi') &= \gamma \left\{ (1 - \beta)w'(z', \xi')^{1-\rho} + \beta \left[(1 - \phi)W'(z', \xi') + \phi \mathcal{U}^{1-RRA} \right]^{\frac{1-\rho}{1-RRA}} \right\}^{\frac{\rho-RRA}{1-\rho}} \\ &\quad (1 - RRA)(1 - \beta)w'(z', \xi')^{-\rho} \\ \gamma' &= \gamma \left\{ (1 - \beta)w'(z', \xi')^{1-\rho} + \beta \left[(1 - \phi)W'(z', \xi') + \phi \mathcal{U}^{1-RRA} \right]^{\frac{1-\rho}{1-RRA}} \right\}^{\frac{\rho-RRA}{1-\rho}} \\ &\quad \times \left[(1 - \phi)W'(z', \xi') + \phi \mathcal{U}^{1-RRA} \right]^{\frac{RRA-\rho}{1-RRA}} \end{aligned}$$

Conceptually these optimality conditions are similar to the simple risk-sharing condition from Section 3, where now the marginal value of a dollar for a worker earning $w'(z', \xi')$ does not depend only on the current wage payment, but also on future promised utilities. In order to highlight the role played here by the EIS and the RRA coefficient, it is useful to consider three illustrative examples.

Example 1: static problem. In order to highlight the role of the RRA coefficient, consider a two period economy where agents contract in the first period the wages will be paid in the second period. We have that for any pair $(z'_1, \xi'_1), (z'_2, \xi'_2)$:

$$\frac{\eta(z'_1, \xi'_1)}{\eta(z'_2, \xi'_2)} = \frac{w'(z'_1, \xi'_1)^{-RRA}}{w'(z'_2, \xi'_2)^{-RRA}}$$

In this setting wages solve a pure infra-temporal risk-sharing problem, as there is no dynamic, that depends solely on the risk aversion coefficient of workers RRA , and on the risk aversion of entrepreneurs through the usual multiplier η .

Example 2: deterministic problem with no separation. In order to highlight the role of the EIS, consider an economy with no idiosyncratic and aggregate risk, so that z and S are constant and $\phi = 1$. We can re-arrange the optimality conditions to obtain

$$\frac{\eta_t}{\eta_{t+1}} = \frac{w_t^{-\rho}}{w_{t+1}^{-\rho}}$$

In this setting wages solve a pure inter-temporal problem, as there are no shocks, that depends solely on the EIS coefficient of workers ρ , and on the EIS of entrepreneurs through the usual multiplier η .

Example 3: deterministic problem with separation. In order to highlight the role of separation, consider an economy where z and S are constant, but matches are subject to the idiosyncratic risk of separation. The optimality conditions for wages read:

$$\frac{\eta_t}{\eta_{t+1}} = \left[\frac{(1 - \phi)W' + \phi\mathcal{U}^{1-RAA}}{W'} \right]^{\frac{\rho-RAA}{1-RAA}} \frac{w_t^{-\rho}}{w_{t+1}^{-\rho}}$$

Note that whenever workers would prefer to be employed, we have that $W' < \mathcal{U}^{1-RAA}$. Then we have that

$$\frac{\eta_t}{\eta_{t+1}} > \frac{w_t^{-\rho}}{w_{t+1}^{-\rho}} \quad \Leftrightarrow \quad RRA > \rho$$

which means workers wages are less back-loaded than they would in the CRRA case if $RAA > \rho$. Intuitively, when $RAA > \rho$ workers have a preference for early resolution of uncertainty, where here uncertainty comes from separation occurring with probability ϕ . Thus, workers are less willing to accept back-loaded wages as there is a positive probability ϕ that they won't get the future promised utility W' .