

Notes on Special Discrete Distributions

(1) Bernoulli(p) X is Bernoulli(p) if it has probability mass function given by

$$p(x) = P(X = x) = \begin{cases} p & \text{if } x = 1 \\ 1 - p & \text{if } x = 0 \\ 0 & \text{otherwise} \end{cases}$$

Thus, $X \in \{0, 1\}$. X can be thought of as an *indicator random variable*, which indicates if some event happens ($X = 1$) or not ($X = 0$).

It is easy to compute $E(X) = 0 \cdot (1 - p) + 1 \cdot p = p$. Also, $E(X^2) = 0^2 \cdot (1 - p) + 1^2 \cdot p = p$, so $\text{var}(X) = E(X^2) - [E(X)]^2 = p - p^2 = p(1 - p)$. We also get

$$F(x) = P(X \leq x) = \begin{cases} 0 & \text{if } x < 0 \\ 1 - p & \text{if } 0 \leq x < 1 \\ 1 & \text{if } x \geq 1 \end{cases}$$

(2) Binomial(n, p) X is Binomial(n, p) if its mass function is given by

$$p(x) = P(X = x) = \begin{cases} \binom{n}{x} p^x (1 - p)^{n-x} & \text{if } x \in \{0, 1, 2, \dots, n\} \\ 0 & \text{otherwise} \end{cases}$$

Thus, $X \in \{0, 1, 2, \dots, n\}$. X can be thought of as the number of *successes* in n independent trials, where the probability any one trial is a success is p .

We can compute the expectation

$$E(X) = \sum_{i=0}^n i \cdot \binom{n}{i} p^i (1 - p)^{n-i},$$

which evaluates to $E(X) = np$, after some manipulation. We can also compute

$$\text{var}(X) = E(X^2) - [E(X)]^2 = \sum_{i=0}^n i^2 \cdot \binom{n}{i} p^i (1 - p)^{n-i} - [np]^2 = np(1 - p)$$

(It is not obvious how to do the tricks to sum the series, but we will see another trick later for how to get $E(X)$ and $\text{var}(X)$ for a Binomial(n, p), based on observing that X can be written as a sum of n independent and identically distributed Bernoulli(p) random variables.)

One can use the mass function $p(x)$ to compute the cdf, $F(x)$, but there is no closed form expression, just an ugly summation.

(3) Poisson(λ) X is Poisson(λ) if its mass function is given by

$$p(x) = P(X = x) = \begin{cases} e^{-\lambda} \frac{\lambda^x}{x!} & \text{if } x \in \{0, 1, 2, 3, \dots\} \\ 0 & \text{otherwise} \end{cases}$$

Thus, $X \in \{0, 1, 2, \dots\}$. We can compute the expectation

$$E(X) = \sum_{i=0}^{\infty} i \cdot e^{-\lambda} \frac{\lambda^i}{i!} = \lambda$$

and $\text{var}(X) = E(X^2) - [E(X)]^2$ works out to give $\text{var}(X) = \lambda$. The cdf, $F(x)$, can be obtained from $p(x)$ in the usual way, as a summation, but has no nice closed formula.

A Poisson(λ) random variable X is often used to model the number of misprints on a page of text, or the number of customers that arrive at a store in a given interval of time, or the number of “events” that occur in a given interval of time, etc.

The Poisson(λ) distribution is an approximation to the Binomial(n, p) distribution for the case that n is large, p is small, and $\lambda = np$. In other words, if Y is Binomial(n, p), and n is large and p is small, and X is Poisson(λ) with $\lambda = np$, then

$$P(Y = i) = \binom{n}{i} p^i (1 - p)^{n-i} \approx P(X = i) = e^{-np} \frac{(np)^i}{i!},$$

Examples from the text: 7a (page 162), 7b (page 162), 7c (page 162), self-test problem 14, page 203.

Example: On average, there is 1 typo per 3 pages of a book. Find the probability that there is at least one typo on page 37.

Let X denote the number of typos on page 37. Since there is an average of 1 typo per 3 pages, we know that $E(X) = 1/3$. We can model X as a $\text{Poisson}(1/3)$ random variable. We want to compute

$$P(X \geq 1) = 1 - P(X = 0) = 1 - e^{-1/3}$$

(Alternatively, $P(X \geq 1) = \sum_{i=1}^{\infty} e^{-1/3} \frac{(1/3)^i}{i!}$.)

Now suppose we are also asked to find the probability that there are at least 3 typos on page 37, given that there is at least one typo on page 37 that you have already seen. We want to compute

$$P(X \geq 3 \mid X \geq 1) = \frac{P(X \geq 3, X \geq 1)}{P(X \geq 1)} = \frac{P(X \geq 3)}{P(X \geq 1)} = \frac{1 - e^{-1/3}(1 + (1/3) + (1/3)^2/2)}{1 - e^{-1/3}}$$

Example: Suppose that the probability that a radio lasts more than 15 years is $p = e^{-3}$. (a). What is the (exact) probability that, of 1000 such radios that you purchase, at least four of them last more than 15 years?

Let X be the number (among the 1000) that last more than 15 years. Then, X is $\text{Binomial}(1000, p)$, where $p = e^{-3}$. We want

$$P(X \geq 4) = \sum_{i=4}^{1000} \binom{1000}{i} p^i (1-p)^{1000-i}$$

(b). Give an alternative expression that approximates your answer to part (b).

Since 1000 is “large” and $p = e^{-3}$ is “small”, we know that X has a distribution that is approximated by that of Y , a $\text{Poisson}(1000p)$ random variable. Thus,

$$P(X \geq 4) \approx P(Y \geq 4) = \sum_{i=4}^{\infty} e^{-1000p} \frac{(1000p)^i}{i!}$$

(Note: you could also stop the summation at $i = 1000$; it does not really matter for the approximation.)

Example: People enter a store, on average one every two minutes. (a). What is the probability that no people enter between 12:00 and 12:05?

Let X be the number of people that enter between 12:00 and 12:05. We model X as a $\text{Poisson}(\lambda)$ random variable. What should λ be? It should be such that $\lambda = E(X)$, the average number of people that arrive in the 5-minute interval. Well, if 1 person arrives every 2 minutes, on average (so $1/2$ a person per minute), in 5 minutes and average of 2.5 people will arrive. Thus, $\lambda = 2.5$.

We want to compute $P(X = 0) = e^{-2.5} \frac{(2.5)^0}{0!} = e^{-2.5}$.

(b). Find the probability that at least 4 people enter during [12:00, 12:05].

We want

$$P(X \geq 4) = 1 - e^{-2.5} - \frac{(2.5)^1}{1!} e^{-2.5} - \frac{(2.5)^2}{2!} e^{-2.5} - \frac{(2.5)^3}{3!} e^{-2.5}$$

(4) Geometric(p) X is $\text{Geometric}(p)$ if its mass function is given by

$$p(x) = P(X = x) = \begin{cases} (1-p)^{x-1} \cdot p & \text{if } x \in \{1, 2, 3, \dots\} \\ 0 & \text{otherwise} \end{cases}$$

Thus, $X \in \{1, 2, \dots\}$. We can compute the expectation

$$E(X) = \sum_{i=1}^{\infty} i \cdot (1-p)^{i-1} \cdot p = \frac{1}{p}$$

and $\text{var}(X) = E(X^2) - [E(X)]^2$ works out to give $\text{var}(X) = (1-p)/p^2$.

The cdf, $F(x)$, can be obtained from $p(x)$ in the usual way, as a summation, and actually does have a nice closed formula. For $i = 1, 2, 3, \dots$, we get

$$P(X \leq i) = 1 - P(X \geq i+1) = 1 - \sum_{j=i+1}^{\infty} (1-p)^{j-1} \cdot p = 1 - p(1-p)^i \sum_{j=0}^{\infty} (1-p)^j = 1 - (1-p)^i.$$

Thus, the cdf is

$$F(x) = P(X \leq x) = \begin{cases} 0 & \text{if } x < 1 \\ 1 - (1-p)^i & \text{if } i \leq x < i+1, \text{ for } i = 1, 2, 3, \dots \end{cases}$$

A Geometric(p) random variable X can be thought of as the number of trials until first success when independent trials are repeated over and over until the first success (up to and including the success; thus, $X \geq 1$).

Example: Suppose parts are of two varieties: good (with probability 90/92) and slightly defective (with probability 2/92). Parts are produced one after the other. What is the probability that at least 5 parts must be produced until there is a slightly defective part produced?

Let X be the number of parts produced up to (and including) the first slightly defective part. Then, X is Geometric(2/92).

We want to compute

$$P(X \geq 5) = \sum_{i=5}^{\infty} (90/92)^{i-1} (2/92) = \frac{2}{92} \left(\frac{90}{92}\right)^4 \cdot \frac{1}{1 - \frac{90}{92}} = \left(\frac{90}{92}\right)^4$$

(Alternatively, we could have written $P(X \geq 5) = 1 - P(1 \leq X \leq 4)$, and written out a short summation of terms; it gives the same answer, of course.)

(5) Negative Binomial (r, p) X is NegBin(r, p) if its mass function is given by

$$p(x) = P(X = x) = \begin{cases} \binom{x-1}{r-1} p^r (1-p)^{x-r} & \text{if } x \in \{r, r+1, r+2, \dots\} \\ 0 & \text{otherwise} \end{cases}$$

Here, r is a positive integer and $0 < p < 1$. Thus, $X \in \{r, r+1, r+2, \dots\}$. (A NegBin random variable is also known as a “Pascal” random variable.)

We can compute the expectation

$$E(X) = \sum_{i=1}^{\infty} i \cdot \binom{i-1}{r-1} p^r (1-p)^{i-r} = \frac{r}{p}$$

and $\text{var}(X) = E(X^2) - [E(X)]^2$ works out to give $\text{var}(X) = r(1-p)/p^2$. The cdf, $F(x)$, can be obtained from $p(x)$ in the usual way, as a summation, but has no nice closed formula.

A NegBin(r, p) random variable X can be thought of as the number of trials until the r th success when independent trials are repeated over and over until the r th success (up to and including the r th success; thus, $X \geq r$).

Example: [problem 64, page 195, Ross] The suicide rate in a certain state is 1 suicide per 100,000 inhabitants per month. (a). Find the probability that in a city of 400,000 inhabitants within this state, there will be 8 or more suicides in the month of June this year.

Let X be the number of suicides in this city in the month of June this year. Then, X is counting the number of “successes” in 400,000 independent trials, where “success” here means that a person commits suicide during the month of June, which happens with probability 1/100,000. We are making the assumption of independence: one person’s act of suicide does not influence another person to commit (or not) suicide. Thus, X is Binomial(n, p), where $n = 400,000$, $p = 1/100,000$.

We want to compute

$$\alpha = P(X \geq 8) = 1 - P(X \leq 7) = 1 - \left(\binom{n}{0} p^0 (1-p)^{n-0} + \binom{n}{1} p^1 (1-p)^{n-1} + \dots + \binom{n}{7} p^7 (1-p)^{n-7} \right)$$

Now, we can also make a very good estimate of this probability, since $n = 400,000$ is large and $p = 1/100,000$ is small: X is approximately Poisson(λ), where $\lambda = 400,000(1/100,000) = 4$. (Note that X is **not** a Poisson random variable; we are just using it as an approximation.) Thus,

$$\alpha = P(X \geq 8) = 1 - P(X \leq 7) = 1 - \left(1 + \frac{4}{1!} + \frac{4^2}{2!} + \frac{4^3}{3!} + \dots + \frac{4^7}{7!} \right) e^{-4}$$

(b). What is the probability that there will be at least two months during the year in which the city has 8 or more suicides?

Let Y be the number of months this year in which the city has 8 or more suicides (a “bad month”). Then $Y \in \{0, 1, 2, \dots, 12\}$, and, assuming independence among the months, Y is $\text{Binomial}(12, \alpha)$, where α is the number from part (a). We want to compute

$$P(Y \geq 2) = 1 - P(Y = 0) - P(Y = 1) = 1 - \binom{12}{0} \alpha^0 (1 - \alpha)^{12} - \binom{12}{1} \alpha^1 (1 - \alpha)^{11}.$$

(c). Counting the present month as month number 1, what is the probability that the first month to have 8 or more suicides (in the city) will be month number i , $i \geq 1$?

Let Z be the number of the first month that has 8 or more suicides. Then Z is $\text{Geometric}(\alpha)$. We want to compute $P(Z = i) = (1 - \alpha)^{i-1} \alpha$, for $i = 1, 2, 3, \dots$

(d). Today is January 1. What is the probability that the fourth month to have 8 or more suicides in the city is December?

Let U be the month number of the fourth month to have 8 or more suicides in the city. Then U is $\text{NegBin}(4, \alpha)$, since it counts how many “trials” (months) until the 4th “success” (a bad month, with 8 or more suicides). We want to compute

$$P(U = 12) = \binom{12-1}{4-1} \alpha^4 (1 - \alpha)^{12-4}$$
