

Let's start with our usual skill choice problem. One dimension we want to leave open is individual heterogeneity in the skill constraint. That is, we think of relative weights θ^e as given by the worker's education group but the total constraint w as heterogeneous on the individual level and potentially correlated with education and job choice. Also, we're letting A vary by job.

For reference, J is a job, i and j are skill dimensions, n and m are workers, and various e s are education levels.

1 Le Within

Given a job J , a worker with education e solves

$$\max y(A \circ S, J) \tag{1}$$

$$\text{s.t. } \sum_i S_i \theta_i^e \leq w. \tag{2}$$

The first order condition in dimension i is

$$A_i y'_i(A \circ S^*, J) = \lambda \theta_i^e. \tag{3}$$

Now, assume a CES production function

$$y(A \circ S, J) = \left(\sum_i J_i (A_i S_i)^\sigma \right)^{\frac{1}{\sigma}}. \tag{4}$$

Our first-order condition is now

$$J_i S_i^{\sigma-1} A_i^\sigma \left(\sum_k J_k (A_k S_k)^\sigma \right)^{\frac{1-\sigma}{\sigma}} = \lambda \theta_i^e. \tag{5}$$

Dividing across different skills and raising to the power $\frac{1}{\sigma-1}$, we have

$$\frac{S_k}{S_i} = \left(\frac{J_k}{J_i} \right)^{-\frac{1}{\sigma-1}} \left(\frac{A_k}{A_i} \right)^{-\frac{\sigma}{\sigma-1}} \left(\frac{\theta_k^e}{\theta_i^e} \right)^{\frac{1}{\sigma-1}} \tag{6}$$

Let's multiply both sides with θ_i^e .

$$\frac{\theta_k^e S_k}{S_i} = \left(\frac{J_k}{J_i}\right)^{-\frac{1}{\sigma-1}} \left(\frac{A_k}{A_i}\right)^{-\frac{\sigma}{\sigma-1}} \left(\frac{\theta_k^e}{\theta_i^e}\right)^{\frac{1}{\sigma-1}} \quad (7)$$

Now, let's sum across i s and use the fact that $\sum_i \theta_i^e S_i = w$. Rearranging,

$$S_i = \frac{J_i^{-\frac{1}{\sigma-1}} A_i^{-\frac{\sigma}{\sigma-1}} (\theta_i^e)^{\frac{1}{\sigma-1}}}{\sum_k J_k^{-\frac{1}{\sigma-1}} A_k^{-\frac{\sigma}{\sigma-1}} (\theta_k^e)^{\frac{\sigma}{\sigma-1}}} w \quad (8)$$

... which looks terrible. Instead of losing hope, take a derivative with respect to A_k where $k \neq i$.

$$\frac{dS_i}{dA_k} = \frac{\sigma}{\sigma-1} \frac{J_i^{-\frac{1}{\sigma-1}} A_i^{-\frac{\sigma}{\sigma-1}} (\theta_i^e)^{\frac{1}{\sigma-1}}}{\sum_l J_l^{-\frac{1}{\sigma-1}} A_l^{-\frac{\sigma}{\sigma-1}} (\theta_l^e)^{\frac{\sigma}{\sigma-1}}} w \frac{J_k^{-\frac{1}{\sigma-1}} A_k^{-\frac{\sigma}{\sigma-1}-1} (\theta_k^e)^{\frac{\sigma}{\sigma-1}}}{\sum_l J_l^{-\frac{1}{\sigma-1}} A_l^{-\frac{\sigma}{\sigma-1}} (\theta_l^e)^{\frac{\sigma}{\sigma-1}}} \quad (9)$$

which simplifies, rather miraculously, to

$$\frac{dS_i}{d \ln A_k} = \frac{\sigma}{\sigma-1} \frac{S_i S_k \theta_k^e}{w}. \quad (10)$$

Similarly, we have that

$$\frac{dS_i}{d \ln A_i} = \frac{\sigma}{\sigma-1} \frac{S_i^2 \theta_i^e}{w} - \frac{\sigma}{\sigma-1} S_i. \quad (11)$$

Thus, we have for each worker and each skill dimension i a first-order expansion of the form

$$\Delta \ln S_i = \frac{\sigma}{\sigma-1} \left[\sum_k \frac{S_k \theta_k^e}{w} \Delta \ln A_k - \Delta \ln A_i \right] \quad (12)$$

Notice this bakes in a few things we like - the Slutsky symmetry property, for instance. Now, there's just one problem - w is heterogeneous at the individual level, whereas we don't have panel data.

Subtracting across different dimensions, we have

$$\Delta \ln S_i - \Delta \ln S_k = \frac{\sigma}{\sigma-1} [\Delta \ln A_k - \Delta \ln A_i] \quad (13)$$

This time around, we're allowing A to vary at the job level; so we'll be writing A_{iJ} for the technological coefficient of skill i in job J . Of course,

these equations have quite a bit of redundancy, as everything is HOD0. So, let's pick a skill - manual - and set $\Delta \ln A_{\text{manual}J} = 0$ in every job.

Unfortunately, we don't observe individual-level skill changes. So we can't actually compute $\Delta \ln S_i$. But, under the assumption that the distribution of w is the same within the occupation-education-job cell before and after the change, we can instead sum across workers within the occ-ed-job cell.

Thus in period t for workers in job J with education e in skill $i \neq \text{manual}$ we have

$$\Delta \ln \overline{S_{iJ\text{etn}}^e} - \Delta \ln \overline{S_{\text{manuale}J\text{tn}}^e} = -\frac{\sigma}{\sigma - 1}(\Delta \ln A_{iJ})_t = -\pi_{iJt}. \quad (14)$$

As we now are in a position to produce one estimate of π_{iJt} for each education group, though, we take a simple average.

$$\hat{\pi}_{iJt} = \frac{1}{3} \sum_e [\Delta \ln \overline{S_{\text{manuale}J\text{tn}}^e} - \Delta \ln \overline{S_{iJ\text{etn}}^e}] \quad (15)$$

Now, we plug (15) into (12) and have

$$\Delta \ln \overline{S_{iJ\text{etn}}^e} + \hat{\pi}_{iJt} = \sum_k \frac{S_{kJ}\theta_k^e}{w} \hat{\pi}_{kJt}. \quad (16)$$

This must hold at the *individual* level, as the model predicts that $\frac{S_{kn}\theta_k^e}{w_n}$ is constant over individuals in the occupation-education cell. Therefore,

$$\frac{1}{N_{J\text{et}}} \sum_n^{N_{J\text{et}}} S_{kJ\text{etn}}\theta_k^e = \frac{1}{N_{J\text{et}}} \sum_n^{N_{J\text{et}}} \frac{S_{kJ\text{etn}}\theta_k^e}{w_n} w_n = \frac{S_k\theta_k^e}{w} \frac{1}{N_{J\text{et}}} \sum_n^{N_{J\text{et}}} w_n \quad (17)$$

so that using the fact that we've assumed $E[w_n] = 1$ in every job,

$$\theta_k^e \overline{S_{kJ\text{etn}}} = \frac{1}{N_{J\text{et}}} \sum_n^{N_{J\text{et}}} S_{kn}\theta_k^e = \frac{S_k\theta_k^e}{w} \quad (18)$$

finally arriving at our regression equation

$$\Delta \ln \overline{S_{iJ\text{etn}}^e} + \hat{\pi}_{iJt} = \sum_k \theta_k^e \overline{S_{kJ\text{etn}}} \hat{\pi}_{kJt} \quad (19)$$

Concern: we're assuming not just that the average ability within each group is $\bar{w} = 1$ but that this is true of the average ability within each occ-educ pair. This is actually falsifiable with our data alone, as we can compute this within-occ average using observed Ss and our estimated θ s. There are alternatives. We could iterate, solving for θ s and computing individual w s, plugging them back in to (16) and so on. I know iterating thetas left a bad taste in our mouth last time around, but *this time is different...*

2 Le Between: Discrete Choice of the Proletariat

Unlike in the DOT paper, we want a theory of sorting across jobs other than indifference. To that end, we'll employ a discrete choice framework in which a worker n in job J with wage w_J gets utility

$$\ln w_J + \xi_J + \eta_{Jn} \quad (20)$$

where ξ_J is a parameter expressing a common preference for job J , and η_{Jn} expresses a combination of idiosyncratic preference for and productivity in J and is i.i.d Gumbel-distributed.

Thus, the probability job J is chosen by a worker of type e is equal to

$$\frac{w_e(J)e^{\xi_J}}{\sum_{J'} w_e(J')e^{\xi_{J'}}} \quad (21)$$

Where do wages come from? As workers solve the problem in [crossref], and each intermediate output J is priced by the market somehow, $w_e(J) = y_e^*(J)P(J)$. Therefore the ratio of e workers in job J to those in J' is

$$\frac{q_e(J)}{q_e(J')} = \frac{P(J)y_e^*(J)e^{\xi_J}}{P(J')y_e^*(J')e^{\xi_{J'}}} \quad (22)$$

Summing across jobs, if the total amount of workers of type e is q_e^{Total} , we get

$$q_e(J) = \frac{P(J)y_e^*(J)e^{\xi_J}}{\sum_{J'} P(J')y_e^*(J')e^{\xi_{J'}}} q_e^{Tot} \quad (23)$$

Instead of imposing assumptions and estimating the discrete choice model at hand, we take a different tack: comparing across types.

$$\frac{q_e(J)}{q_{e'}(J)} = \frac{y_e^*(J)}{y_{e'}^*(J)} \frac{q_e^{Tot}}{q_{e'}^{Tot}} \frac{\sum_{J'} P(J') y_{e'}^*(J') e^{\xi_{J'}}}{\sum_{J'} P(J') y_e^*(J') e^{\xi_{J'}}} \quad (24)$$

From here, we can exploit changes over time.

$$\Delta \ln \frac{q_e(J)}{q_{e'}(J)} = \sum_i \left(\frac{\partial \ln y_e^*(J)}{\partial \ln A_{iJ}} - \frac{\partial \ln y_{e'}^*(J)}{\partial \ln A_{iJ}} \right) \Delta \ln A_{iJ} \quad (25)$$

$$+ \Delta \ln \left[\frac{q_e^{Tot} \sum_{J'} P(J') y_{e'}^*(J') e^{\xi_{J'}}}{q_{e'}^{Tot} \sum_{J'} P(J') y_e^*(J') e^{\xi_{J'}}} \right] \quad (26)$$

Using the envelope theorem, and noticing that the last term is a constant across J , we arrive at

$$\Delta \ln \frac{q_e(J)}{q_{e'}(J)} = \sum_i \left(\frac{S_{iJ\epsilon n} \theta_i^e}{w_{en}} - \frac{S_{iJ\epsilon' n} \theta_i^{e'}}{w_{e'n}} \right) \Delta \ln A_{iJ} + const \quad (27)$$

For our purposes, we'll use the mean $S\theta/w$ within the educ-occ cell. Substituting in π ,

$$\Delta \left[\ln \frac{q_e(J)}{q_{e'}(J)} \right]_t = \beta_J \sum_k \left(\theta_k^e \overline{S_{kJ\epsilon n}} - \theta_k^{e'} \overline{S_{kJ\epsilon' n}} \right) \hat{\pi}_{kJt} + const_{e,e',t} \quad (28)$$

The interpretation here is that $\beta_J = \frac{\sigma-1}{\sigma}$, which may or may not be interesting to know. A test of our model is that each $\sigma_J < 1$. More importantly, this will allow us to test our model's goodness of fit. Or something.