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## 1 Warmup: Probabilities

1. We are given that  $P(A_1) = P(A_2) = P(A_1/A_2) = \frac{1}{2}$ 

From bayes theorm:  $P(A_1 \cap A_2) = P(A_1 \mid A_2)P(A_2)$ 

From the above two expressions, we have  $P(A_1 \cap A_2) = P(A_1)P(A_2)$ 

We see that the probability of occurrence of both the events  $(A_1 \text{ and } A_2)$  is equal to the product of the probabilities of occurrence of each event. This means that occurrence of  $A_1$  does not affect the probability of  $A_2$  occurring. Hence the events  $A_1$  and  $A_2$  are independent.

2. 
$$P(A_1) = P(A_2) = P(A_3) = 1/3$$
  
 $P(A_4 \mid A_1) = 1/6$   $P(A_4 \mid A_2) = 2/6$   $P(A_4 \mid A_3) = 3/6$   
 $P(A_4 \cap A_1) = P(A_4 \mid A_1)P(A_1) = (1/6) * (1/3) = 1/18$   
 $P(A_4 \cap A_2) = P(A_4 \mid A_2)P(A_2) = (2/6) * (1/3) = 2/18$   
 $P(A_4 \cap A_3) = P(A_4 \mid A_3)P(A_3) = (3/6) * (1/3) = 3/18$   
 $P(A_4) = P(A_4 \cap A_1) + P(A_4 \cap A_2) + P(A_4 \cap A_3) = 6/18 = 1/3 = 0.333\overline{3}$ 

3. If a six sided die is tossed than the possible numbers that can occur on the top face of the die are either 1 or 2 or 3 or 4 or 5 or 6.  $n = \{1, 2, 3, 4, 5, 6\}$ 

Also, the probability of any of the 6 numbers appearing on the top face of the die is P(n) = 1/6 n is the number of times the coin is tossed

$$P(heads = 2 \mid n = 1) = 0$$

$$P(heads = 2 \mid n = 2) = 1/4$$

$$P(heads = 2 \mid n = 3) = 3/8$$

$$P(heads = 2 \mid n = 4) = 6/16$$

$$P(heads = 2 \mid n = 5) = 10/32$$

$$P(heads = 2 \mid n = 6) = 15/64$$

Probability of exactly 2 heads = 
$$P(heads = 2) = \sum_{n=1}^{6} P(heads = 2 \mid n)P(n)$$
  
=  $\frac{1}{6}\sum_{n=1}^{6} P(heads = 2 \mid n)$ 

$$P(heads = 2) = \frac{1}{6}(1/4 + 3/8 + 6/16 + 12/32 + 15/64) = 33/128 = 0.2578125$$

$$\begin{aligned} 4. \ \ & P(A_1) = a_1 \text{ and } P(A_2) = a_2 \\ & P(A_1 \cup A_2) \leq 1 \\ & P(A_1) + P(A_2) - P(A_1 \mid A_2) P(A_2) \, \leq \, 1 \\ & P(A_1 \mid A_2) \geq \frac{P(A_1) + P(A_2) - 1}{P(A_2)} \\ & P(A_1 \mid A_2) \geq \frac{a_1 + a_2 - 1}{a_2} \end{aligned}$$

5. (a) If  $A_1$  and  $A_2$  are independent events, show that  $E[A_1 + A_2] = E[A_1] + E[A_2]$ Consider that the random variable  $A_1$  takes the values  $a_{11}, a_{12}, a_{13}, ...., a_{1m}$ Consider that the random variable  $A_2$  takes the values  $a_{21}, a_{22}, a_{23}, ...., a_{2n}$   $E[X] = \sum_{i=1}^n x_i P(X = x_i)$ 

$$E[A_1 + A_2] = \sum_{i=1}^{m} \sum_{j=1}^{n} (a_{1i} + a_{2j}) P(A_1 = a_{1i}, A_2 = a_{2j})$$

$$= \sum_{i=1}^{m} \sum_{j=1}^{n} a_{1i} P(A_1 = a_{1i}, A_2 = a_{2j}) + \sum_{i=1}^{m} \sum_{j=1}^{n} a_{2j} P(A_1 = a_{1i}, A_2 = a_{2j})$$

$$= \sum_{i=1}^{m} \left( a_{1i} \sum_{j=1}^{n} P(A_1 = a_{1i}, A_2 = a_{2j}) \right) + \sum_{i=1}^{n} \left( a_{2j} \sum_{i=1}^{m} P(A_1 = a_{1i}, A_2 = a_{2j}) \right)$$

Since the random variables are independent

$$= \sum_{i=1}^{m} \left( a_{1i} \sum_{j=1}^{n} P(A_1 = a_{1i}) P(A_2 = a_{2j}) \right) + \sum_{j=1}^{n} \left( a_{2j} (\sum_{i=1}^{m} P(A_1 = a_{1i}) P(A_2 = a_{2j})) \right)$$

$$= \sum_{i=1}^{m} \left( a_{1i} P(A_1 = a_{1i}) \sum_{j=1}^{n} P(A_2 = a_{2j}) \right) + \sum_{j=1}^{n} \left( a_{2j} P(A_2 = a_{2j}) \sum_{i=1}^{m} P(A_1 = a_{1i}) \right)$$

$$\text{we know that} \quad \sum_{j=1}^{n} P(A_2 = a_{2j}) = 1 \quad \text{and} \quad \sum_{i=1}^{m} P(A_1 = a_{1i}) = 1$$

$$= \sum_{i=1}^{m} (a_{1i}) P(A_1 = a_{1i}) + \sum_{j=1}^{n} (a_{2j}) P(A_2 = a_{2j})$$

$$= E[A_1] + E[A_2] \tag{1}$$

(b) If  $A_1$  and  $A_2$  are independent events, show that  $var[A_1 + A_2] = var[A_1] + var[A_2]$ 

$$var[X] = \sum_{i=1}^{n} (x_i - E[X])^2 P(X = x_i)$$

$$var[A_{1} + A_{2}] = \sum_{i=1}^{m} \sum_{j=1}^{n} (a_{1i} - E[A_{1}] + a_{2j} - E[A_{2}])^{2} P(A_{1} = a_{1i}, A_{2} = a_{2j})$$

$$= \sum_{i=1}^{m} \sum_{j=1}^{n} (a_{1i} - E[A_{1}])^{2} P(A_{1} = a_{1i}, A_{2} = a_{2j})$$

$$+ \sum_{i=1}^{m} \sum_{j=1}^{n} (a_{2j} - E[A_{2}])^{2} P(A_{1} = a_{1i}, A_{2} = a_{2j})$$

$$- 2 \sum_{i=1}^{m} \sum_{j=1}^{n} (a_{1i} - E[A_{1}])(a_{2j} - E[A_{2}]) P(A_{1} = a_{1i}, A_{2} = a_{2j})$$

$$(2)$$

Consider the first two terms of the above equation (2)

$$\sum_{i=1}^{m} \sum_{j=1}^{n} (a_{1i} - E[A_1])^2 P(A_1 = a_{1i}, A_2 = a_{2j}) + \sum_{i=1}^{m} \sum_{j=1}^{n} (a_{2j} - E[A_2])^2 P(A_1 = a_{1i}, A_2 = a_{2j})$$

$$= \sum_{i=1}^{m} \sum_{j=1}^{n} (a_{1i} - E[A_1])^2 P(A_1 = a_{1i}, A_2 = a_{2j})$$

$$+ \sum_{i=1}^{m} \sum_{j=1}^{n} (a_{2j} - E[A_2])^2 P(A_1 = a_{1i}, A_2 = a_{2j})$$

$$= \sum_{i=1}^{m} (a_{1i} - E[A_1])^2 \sum_{j=1}^{n} P(A_1 = a_{1i}, A_2 = a_{2j})$$

$$+ \sum_{i=1}^{n} (a_{2j} - E[A_2])^2 \sum_{i=1}^{m} P(A_1 = a_{1i}, A_2 = a_{2j})$$

(Since the random variables are independent)

$$= \sum_{i=1}^{m} (a_{1i} - E[A_1])^2 \sum_{j=1}^{n} P(A_1 = a_{1i}) P(A_2 = a_{2j})$$

$$+ \sum_{j=1}^{n} (a_{2j} - E[A_2])^2 \sum_{i=1}^{m} P(A_1 = a_{1i}) P(A_2 = a_{2j})$$

$$= \sum_{i=1}^{m} (a_{1i} - E[A_1])^2 P(A_1 = a_{1i}) \sum_{j=1}^{n} P(A_2 = a_{2j})$$

$$+ \sum_{j=1}^{n} (a_{2j} - E[A_2])^2 P(A_2 = a_{2j}) \sum_{i=1}^{m} P(A_1 = a_{1i})$$

$$\text{(we know that } \sum_{j=1}^{n} P(A_2 = a_{2j}) = 1 \quad \text{and} \quad \sum_{i=1}^{m} P(A_1 = a_{1i}) = 1)$$

$$= \sum_{i=1}^{m} (a_{1i} - E[A_1])^2 P(A_1 = a_{1i}) + \sum_{j=1}^{n} (a_{2j} - E[A_2])^2 P(A_2 = a_{2j})$$

$$= var[A_1] + var[A_2]$$

(3)

Now, consider the third term of the equation (2)

$$-2\sum_{i=1}^{m}\sum_{j=1}^{n}(a_{1i}-E[A_{1}])(a_{2j}-E[A_{2}])P(A_{1}=a_{1i},A_{2}=a_{2j})$$

$$=-2\sum_{i=1}^{m}\sum_{j=1}^{n}(a_{1i}-E[A_{1}])(a_{2j}-E[A_{2}])P(A_{1}=a_{1i})P(A_{2}=a_{2j})$$

$$=-2\sum_{i=1}^{m}\left((a_{1i}-E[A_{1}])P(A_{1}=a_{1i})\sum_{j=1}^{n}\left((a_{2j}-E[A_{2}])P(A_{2}=a_{2j})\right)\right)$$

$$=-2\sum_{i=1}^{m}\left((a_{1i}-E[A_{1}])P(A_{1}=a_{1i})\sum_{j=1}^{n}\left((a_{2j}P(A_{2}=a_{2j})-E[A_{2}]P(A_{2}=a_{2j})\right)\right)$$

$$=-2\sum_{i=1}^{m}\left((a_{1i}-E[A_{1}])P(A_{1}=a_{1i})\left(\sum_{j=1}^{n}(a_{2j}P(A_{2}=a_{2j})-\sum_{j=1}^{n}E[A_{2}]P(A_{2}=a_{2j})\right)\right)$$

$$=-2\sum_{i=1}^{m}\left((a_{1i}-E[A_{1}])P(A_{1}=a_{1i})\left(E[A_{2}]-E[A_{2}]\sum_{j=1}^{n}P(A_{2}=a_{2j})\right)\right)$$

$$=-2\sum_{i=1}^{m}\left((a_{1i}-E[A_{1}])P(A_{1}=a_{1i})\left(E[A_{2}]-E[A_{2}]*1\right)\right)$$

$$=-2\sum_{i=1}^{m}\left((a_{1i}-E[A_{1}])P(A_{1}=a_{1i})\left(E[A_{2}]-E[A_{2}]*1\right)\right)$$

Now, from equation 2, equation 3 and equation 4,  $var[A_1 + A_2] = var[A_1] + var[A_2]$ 

## 2 Naive Bayes

1. (a) If we have infinite data drawn from the distribution, then the distribution of the sampled data will be the same as the true distribution and hence the predicted values will be equal to true values.  $\hat{P}(x_1 \mid y) = P(x_1 \mid y)$  and  $\hat{P}(y) = P(y)$ .

	Input $x_1$	$\hat{P}(x_1, y = -1)$	$\hat{P}(x_1, y = 1)$	<b>Prediction:</b> $y' = arg \max_{y} \hat{P}(x_1, y)$
(b)	-1	0.08	0.09	1
	1	0.02	0.81	1

(c) Error of classifier is 10%:

$$P(y' \neq y) = P(y' \neq y, x_1 = -1) + P(y' \neq y, x_1 = 1) = 0.08 + 0.02 = 0.1 = 10\%.$$

2. (a) It is given that  $x_2$  is identical to  $x_1$ , hence,  $P(x_1, x_2 \mid y) = P(x_1 \mid y)$ If  $x_1$  and  $x_2$  were conditionally independent, then  $P(x_1, x_2 \mid y) = P(x_1 \mid y)P(x_2 \mid y)$ For the above equations to agree, we need to have  $P(x_2 \mid y) = 1$ , But this may not be possible. Hence  $x_1$  and  $x_2$  are not conditionally independent.

(4)

	$x_1$	$x_2$	$\hat{P}(x_1, x_2, y = -1)$	$\hat{P}(x_1, x_2, y = 1)$	<b>Prediction:</b> $y' = arg \max_{y} \hat{P}(x_1, x_2, y)$
	-1	-1	0.064	0.009	-1
(b)	-1	1	0.016	0.081	1
	1	-1	0.016	0.081	1
	1	1	0.004	0.729	1

(c) Error of the classifier is 11%:

$$P(y' \neq y) = P(y' \neq y, x_1 = -1, x_2 = -1) + P(y' \neq y, x_1 = -1, x_2 = 1) + P(y' \neq y, x_1 = 1, x_2 = 1) + P(y' \neq y, x_1 = 1, x_2 = -1)$$

It is given that  $x_1$  and  $x_2$  are identical, hence,  $(x_1 = -1 \text{ and } x_2 = 1)$  cannot occur simultaneously. And  $(x_1 = 1 \text{ and } x_2 = -1)$  cannot occur simultaneously.

Also, 
$$P(y' \neq y, x_1 = -1, x_2 = -1) = P(y' \neq y, x_1 = -1) = P(y' \neq y, x_2 = -1)$$
  
Similarly,  $P(y' \neq y, x_1 = 1, x_2 = 1) = P(y' \neq y, x_1 = 1) = P(y' \neq y, x_2 = 1)$   
Then,  $P(y' \neq y) = P(y' \neq y, x_1 = -1) + 0 + 0 + P(y' \neq y, x_1 = 1)$   
 $P(y' \neq y) = 0.09 + 0.02 = 0.11 = 11\%$ 

(d) From the above problem, the performance of the naive bayse classifier decreases when the variables were duplicated. Naive bayse classifier assumes that the features are independent, but logistic regression does not make any such assumption. Logistic regression will perform better than naive bayse since it is not affected by duplication of features.

## 3 Naïve Bayes and Linear Classifiers

The classifier will predict the label 1 if  $\Pr(y=1|\mathbf{x}) \ge \Pr(y=0|\mathbf{x})$ . Or equivalently,

$$\frac{\Pr(\mathbf{x}|y=1)\Pr(y=1)}{\Pr(\mathbf{x}|y=0)\Pr(y=0)} \ge 1$$

The Gaussian distribution is represented by the following probability density function:

$$f(x_j | \mu_{j,y}, \sigma) = \frac{1}{\sqrt{2\sigma^2 \pi}} e^{-\frac{(x_j - \mu_{j,y})^2}{2\sigma^2}}$$

From the naïve Bayes assumption:  $\Pr(\mathbf{x}|y) = \prod_{j=0}^{d} \Pr(x_j|y)$ .

$$\frac{\Pr(y=1)}{\Pr(y=0)} \prod_{j=0}^{d} \frac{\Pr(x_j|y=1)}{\Pr(x_j|y=0)} \ge 1$$

To simplify the notations, let P(y=1)=p, P(y=0)=1-p, and also substitute the probability density function in the above equation.

Let  $\mu_{0j}$  be the mean of the *jth* variable when y=0

Let  $\mu_{1j}$  be the mean of the *jth* variable when y=1

$$\frac{p}{1-p} \prod_{j=0}^{d} \frac{\frac{1}{\sqrt{2\sigma^{2}\pi}} e^{-\frac{(x_{j}-\mu_{1j})^{2}}{2\sigma^{2}}}}{\frac{1}{\sqrt{2\sigma^{2}\pi}} e^{-\frac{(x_{j}-\mu_{1j})^{2}}{2\sigma^{2}}}} \ge 1$$

$$\frac{p}{1-p} \prod_{j=0}^{d} e^{\frac{-(x_{j}-\mu_{1j})^{2}}{2\sigma^{2}} + \frac{(x_{j}-\mu_{0j})^{2}}{2\sigma^{2}}} \ge 1$$

$$\frac{p}{1-p} \prod_{j=0}^{d} e^{\frac{-(x_{j}-\mu_{1j})^{2} + (x_{j}-\mu_{0j})^{2}}{2\sigma^{2}}} \ge 1$$

$$\frac{p}{1-p} \prod_{j=0}^{d} e^{\frac{-x_{j}^{2}-\mu_{1j}^{2} + 2x_{j}\mu_{1j} + x_{j}^{2} + \mu_{0j}^{2} - 2x_{j}\mu_{0j}}{2\sigma^{2}}} \ge 1$$

$$\frac{p}{1-p} \prod_{j=0}^{d} e^{\frac{(\mu_{0j}^{2}-\mu_{1j}^{2}) + x_{j}(2\mu_{1j}-2\mu_{0j})}{2\sigma^{2}}} \ge 1$$

$$\frac{p}{1-p} \prod_{j=0}^{d} e^{\frac{(\mu_{0j}^{2}-\mu_{1j}^{2})}{2\sigma^{2}}} e^{\frac{x_{j}(2\mu_{1j}-2\mu_{0j})}{2\sigma^{2}}} \ge 1$$

$$\frac{p}{1-p} \prod_{j=0}^{d} e^{\frac{(\mu_{0j}^{2}-\mu_{1j}^{2})}{2\sigma^{2}}} \prod_{j=0}^{d} e^{\frac{x_{j}(2\mu_{1j}-2\mu_{0j})}{2\sigma^{2}}} \ge 1$$

$$log(\frac{p}{1-p} \prod_{i=0}^{d} e^{\frac{(\mu_{0j}^{2}-\mu_{1j}^{2})}{2\sigma^{2}}}) + \sum_{i=0}^{d} x_{j} \frac{(2\mu_{1j}-2\mu_{0j})}{2\sigma^{2}} \ge 0$$

For any input  $\mathbf{x}$ , the first term in this summation is a constant because it does not have any  $x_j$  terms. Let us denote it by  $b = log \left( \frac{p}{1-p} \prod_{j=0}^d e^{\frac{(\mu_{0j}^2 - \mu_{1j}^2)}{2\sigma^2}} \right)$ . Further, let us denote  $\frac{(2\mu_{1j} - 2\mu_{0j})}{2\sigma^2}$  by  $w_j$ . Substituting these, we get the following expression:

$$b + \sum_{j=0}^{d} x_j w_j \ge 0$$

We obtained this condition for predicting that the label is 1. This means that our classifier is a linear classifier.

# 4 Experiment

1. 
$$g(\mathbf{w}) = \log(1 + \exp(-y_i \mathbf{w}^T \mathbf{x}_i))$$
  

$$\frac{dg(\mathbf{w})}{d\mathbf{w}} = \frac{1}{1 + \exp(-y_i \mathbf{w}^T \mathbf{x}_i)} \exp(-y_i \mathbf{w}^T \mathbf{x}_i)(-y_i \mathbf{x}_i)$$

$$\frac{dg(\mathbf{w})}{d\mathbf{w}} = \frac{-y_i \mathbf{x}_i}{\exp(y_i \mathbf{w}^T \mathbf{x}_i) + 1}$$

2. If the entire dataset is composed of a single example, then the onjective function will be:

$$J = log(1 + \exp(-y_i \mathbf{w}^T \mathbf{x}_i)) + \frac{1}{\sigma^2} \mathbf{w}^T \mathbf{w}$$

The gradient of this objective with respect to the weight vector is:

$$\frac{dJ(\mathbf{w})}{d\mathbf{w}} = \frac{-y_i \mathbf{x}_i}{\exp(y_i \mathbf{w}^T \mathbf{x}_i) + 1} + \frac{2\mathbf{w}}{\sigma^2}$$

3. Pseudo code for the stochastic gradient algorithm using the gradient from previous part:

### initialize w

for each epoch, 
$$t=0,\ 1,\ 2,\dots T)$$
 do Shuffle the data for each training example  $(x_i,y_i)$  do 
$$| \mathbf{w}^{t+1} = \mathbf{w}^t - r(\frac{-y_i x_i}{\exp(y_i \mathbf{w}^T \mathbf{x}_i) + 1} + \frac{2\mathbf{w}^t}{\sigma^2})$$
 end end return  $\mathbf{w}$ 

4. Results of 10-fold Cross validation with 5 epochs and learning rate of 0.01 are:

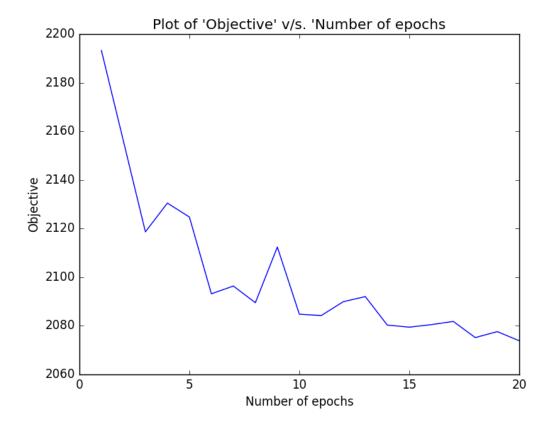
$\sigma$	Cross validation Accuracy(%)	
1	75.586	
50	84.019	
100	84.316	
125	84.051	
150	84.393	
175	84.126	
200	84.390	
225	84.191	
250	84.003	

Best Sigma through Cross Validation is 150, with Average accuracy of 84.393%

Number of Epochs used for final run: 20

Accuracy on training data: 85.27%

Accuracy on test data: 84.64%



From the above figure, it is seen that the Magnitude of the Objective decreases with increase in the number of epochs.