

# Lab 2: Mixed-effects models

April 5, 2018

# Mixed-effects models

## Laws of probability

### 1. Axiom of conditional probability

$$\Pr(X, Y) = \Pr(Y|X) \Pr(X)$$

- Often easier to specify conditional probabilities than joint probabilities

### 2. Law of total probability

$$\Pr(X) = \int \Pr(X, Y) dY$$

- Used when justifying hierarchical models

## Laplace approximation

- Define joint log-likelihood:

$$f(\theta, \varepsilon; y) = \log(\Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2))$$

- Taylor series expansion of joint log-likelihood

$$f(\varepsilon|\theta, y) \approx f(\hat{\varepsilon}|\theta) + f'(\hat{\varepsilon}|\theta)(\hat{\varepsilon} - \varepsilon) + \frac{1}{2}f''(\hat{\varepsilon}|\theta)(\hat{\varepsilon} - \varepsilon)^2$$

- Evaluate Taylor series around “inner maximum”

$$\hat{\varepsilon} = \operatorname{argmax}_{\varepsilon} (f(\theta, \varepsilon))$$

- Implies that  $f'(\hat{\varepsilon}|\theta) = 0$

- Approximate joint likelihood via Taylor series expansion

$$\Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2) = e^{f(\varepsilon|\theta)} \approx e^{f(\hat{\varepsilon}|\theta) - \frac{1}{2}|f''(\hat{\varepsilon})|(\hat{\varepsilon} - \varepsilon)^2}$$

# Likelihood statistics

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## Laplace approximation

- Approximate joint likelihood via Taylor series expansion

$$\Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2) = e^{f(\varepsilon|\theta)} \approx e^{f(\hat{\varepsilon}|\theta) - \frac{1}{2}|f''(\hat{\varepsilon})|(\hat{\varepsilon}-\varepsilon)^2}$$

- Integrate both sides

$$\int \Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2) d\varepsilon = \int e^{f(\varepsilon|\theta)} d\varepsilon$$

$$\int \Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2) d\varepsilon \approx \int e^{f(\hat{\varepsilon}|\theta)} e^{-\frac{1}{2}|f''(\hat{\varepsilon})|(\hat{\varepsilon}-\varepsilon)^2} d\varepsilon$$

- And  $e^{f(\hat{\varepsilon}|\theta)}$  is a constant so:

$$\int e^{f(\hat{\varepsilon}|\theta)} e^{-\frac{1}{2}|f''(\hat{\varepsilon})|(\hat{\varepsilon}-\varepsilon)^2} d\varepsilon = e^{f(\hat{\varepsilon}|\theta)} \int e^{-\frac{1}{2}|f''(\hat{\varepsilon})|(\hat{\varepsilon}-\varepsilon)^2} d\varepsilon$$

- Looks like a normal distribution

- $\hat{\varepsilon}$  is the mean of the normal distribution
- $f''(\hat{\varepsilon})$  is the hessian of the normal distribution ( $f''(\hat{\varepsilon}) = \Sigma^{-1}$ )

$$\text{Normal PDF: } \Pr(\varepsilon|\mu, \Sigma) = \frac{1}{(2\pi)^{k/2} |\Sigma|^{1/2}} \exp\left(\frac{-(\varepsilon - \mu)^T \Sigma^{-1} (\varepsilon - \mu)}{2}\right)$$

## Chi-squared example

$$\Pr(x) = \frac{x^{\frac{k}{2}-1} e^{-\frac{x}{2}}}{c}$$

Defining the log-likelihood

$$\log(\Pr(x)) \equiv f(x)$$

Taking derivatives:

$$f(x) \propto \left(\frac{k}{2} - 1\right) \log(x) - \frac{x}{2}$$

$$f'(x) \propto \left(\frac{k}{2} - 1\right) x^{-1} - \frac{1}{2}$$

$$f''(x) \propto -\left(\frac{k}{2} - 1\right) x^{-2}$$

Solving for mode and Hessian:

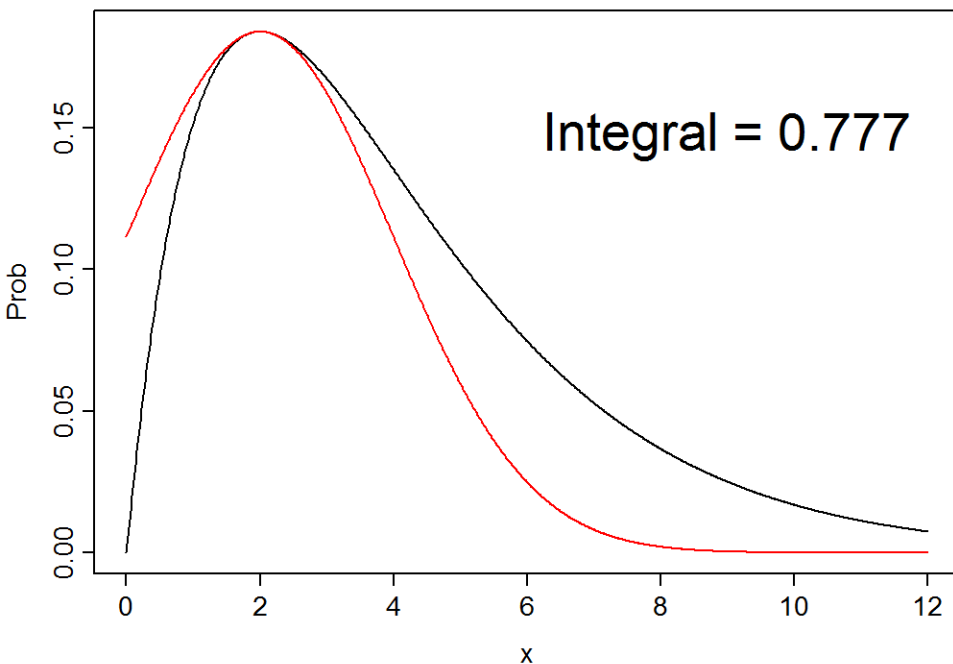
$$f'(x) = 0 \quad \rightarrow \quad \hat{x} = k - 2$$

$$f''(\hat{x}) = -\left(\frac{1}{2(k-2)}\right)$$

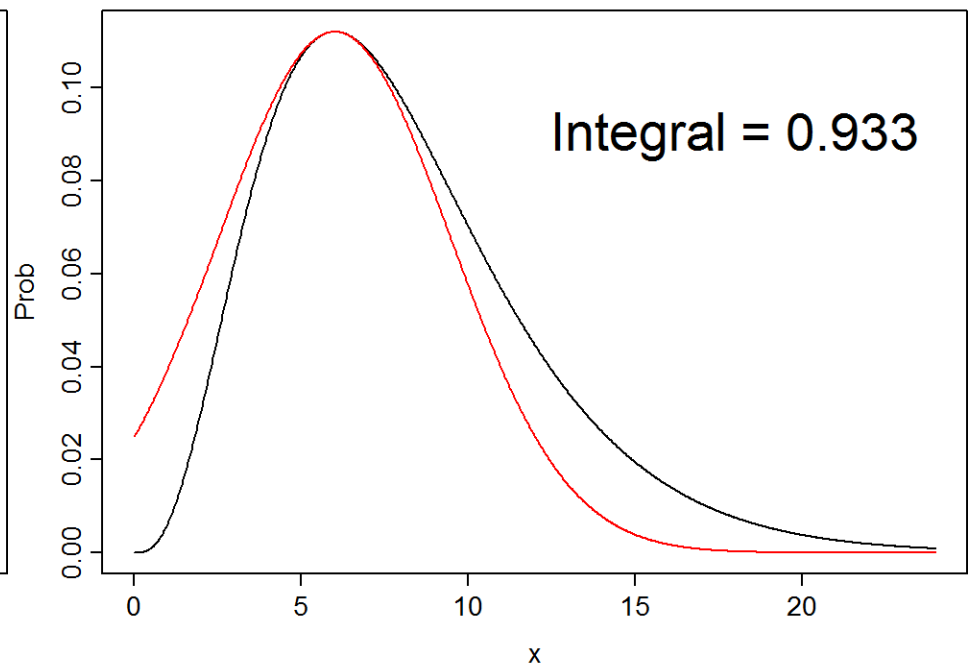
Hence:

$$\Pr(x) \propto \text{Normal}(k - 2, 2(k - 2))$$

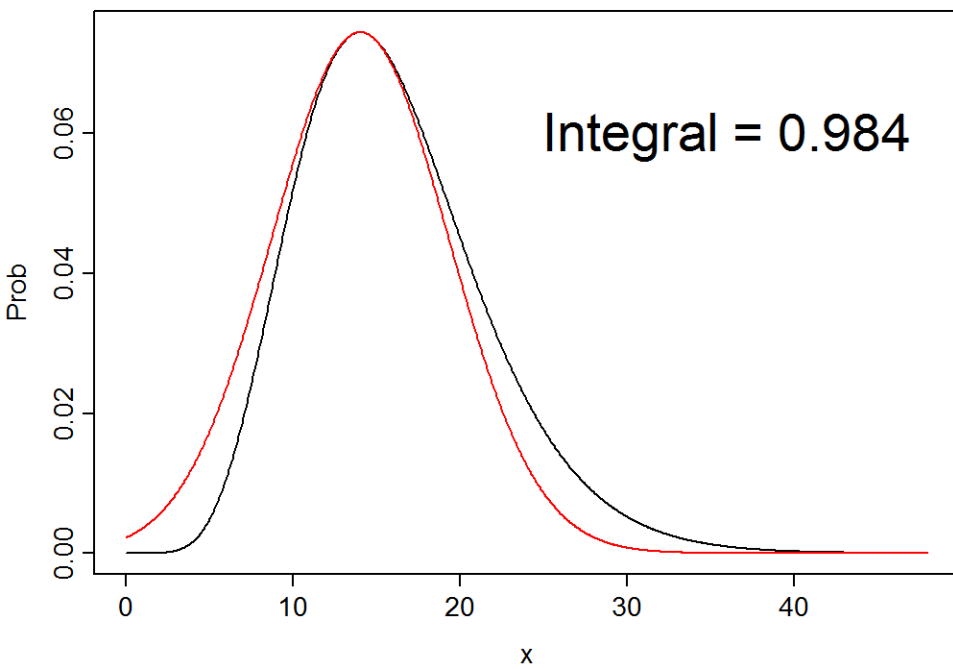
**DF = 4**



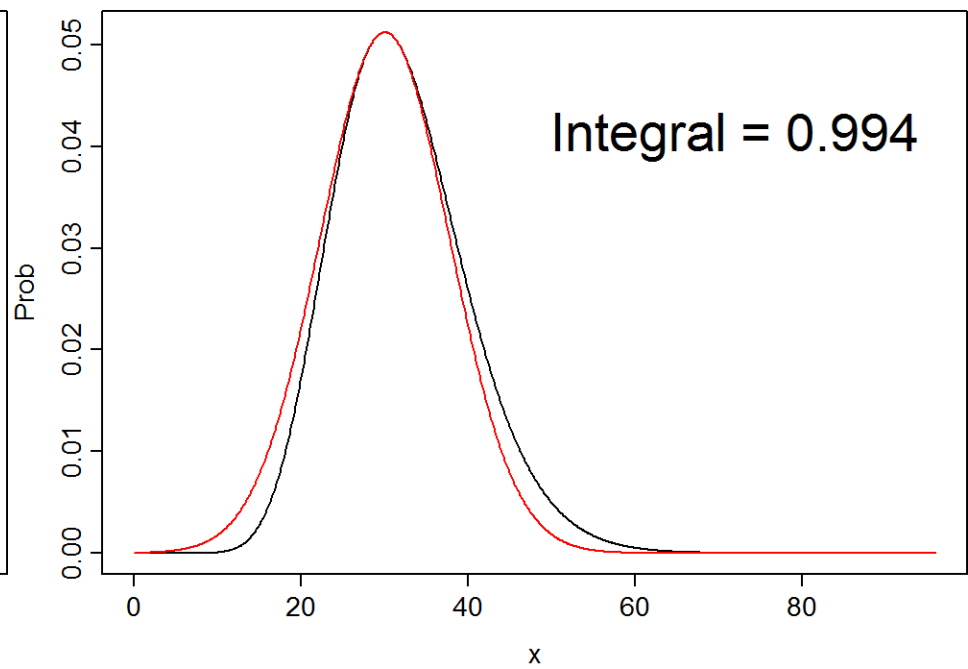
**DF = 8**



**DF = 16**



**DF = 32**



# Likelihood statistics

## Bottom line

$$\ln L(\theta; y) \equiv \int \Pr(y, \varepsilon | \theta) d\varepsilon \cong \log(\Pr(y, \varepsilon | \theta)) - \frac{1}{2} \log(|\mathbf{H}|)$$

– Where

$$\Pr(y, \varepsilon | \theta) = \Pr(y | \theta_1, \varepsilon) \Pr(\varepsilon | \theta_2)$$

– And

$$\mathbf{H} = \frac{\partial^2}{\partial \varepsilon^2} (\log(\Pr(y, \varepsilon | \theta)))$$

- Definitions

- $\log(L(\theta; y))$  is the marginal log-likelihood
- $\Pr(y, \varepsilon | \theta)$  is the joint likelihood
- $|\mathbf{H}|$  is the determinant of the Hessian matrix

# Mixed-effects models

## Bayes rule

- By the Axiom of conditional probability

$$\Pr(\theta|y) \Pr(y) = \Pr(y, \theta) = \Pr(y|\theta) \Pr(\theta)$$

- Therefore

$$\Pr(\theta|y) = \frac{\Pr(y|\theta) \Pr(\theta)}{\Pr(y)}$$

- By the Law of total probability

$$\Pr(y) = \int \Pr(y, \theta) dy = \int \Pr(y|\theta) \Pr(\theta) dy$$

- Therefore

$$\Pr(\theta|y) = \frac{\Pr(y|\theta) \Pr(\theta)}{\int \Pr(y|\theta) \Pr(\theta) dy}$$

- MCMC gives you

$$\Pr(\theta|y) \propto \Pr(y|\theta) \Pr(\theta)$$



# Mixed-effects models

## Empirical Bayes

- By the definition of a likelihood

$$L(\theta; y) = \Pr(y|\theta)$$

- By the Law of total probability

$$\Pr(y|\theta) = \int \Pr(y, \varepsilon|\theta) d\varepsilon$$

- By the Axiom of conditional probability

$$\Pr(y, \varepsilon|\theta) = \Pr(y | \varepsilon, \theta) \Pr(\varepsilon|\theta)$$

- Therefore

$$\Pr(y|\theta) = \int \Pr(y | \varepsilon, \theta) \Pr(\varepsilon|\theta) d\varepsilon$$

# Mixed-effects models

## Generalized linear mixed model

1. Specify distribution for response variable

$$c_i \sim \text{Poisson}(\lambda_i)$$

2. Specify function for expected value

$$g^{-1}(\lambda_i) = x_0 + \mathbf{x}_i^T \boldsymbol{\beta} + \mathbf{z}_i^T \boldsymbol{\varepsilon}$$

3. Specify a link function

$$g^{-1}(a) = \log(a) \rightarrow g(a) = \exp(a)$$

4. Specify distribution for random effects

$$\boldsymbol{\varepsilon} \sim \text{Normal}(0, \sigma_{\boldsymbol{\varepsilon}}^2)$$

= General linear model + mixed effect(s)

# Mixed-effects models

How to estimate standard errors?

- Estimate the “Hessian” at the log-marginal likelihood

$$H(\boldsymbol{\theta}; \mathbf{y}) = \begin{bmatrix} \frac{\partial^2 \ln L(\boldsymbol{\theta}; \mathbf{y})}{\partial \theta_1^2} & \frac{\partial^2 \ln L(\boldsymbol{\theta}; \mathbf{y})}{\partial \theta_1 \partial \theta_2} \\ \frac{\partial^2 \ln L(\boldsymbol{\theta}; \mathbf{y})}{\partial \theta_1 \partial \theta_2} & \frac{\partial^2 \ln L(\boldsymbol{\theta}; \mathbf{y})}{\partial \theta_2^2} \end{bmatrix}$$

- Calculate its inverse

$$\widehat{\mathbf{V}}(\boldsymbol{\theta}; \mathbf{y}) = \mathbf{H}^{-1}$$

- Extract element and take square root

$$\widehat{\text{SE}}(\theta_i; \mathbf{y}) = \sqrt{\widehat{\mathbf{V}}(\boldsymbol{\theta}; \mathbf{y})_{i,i}}$$

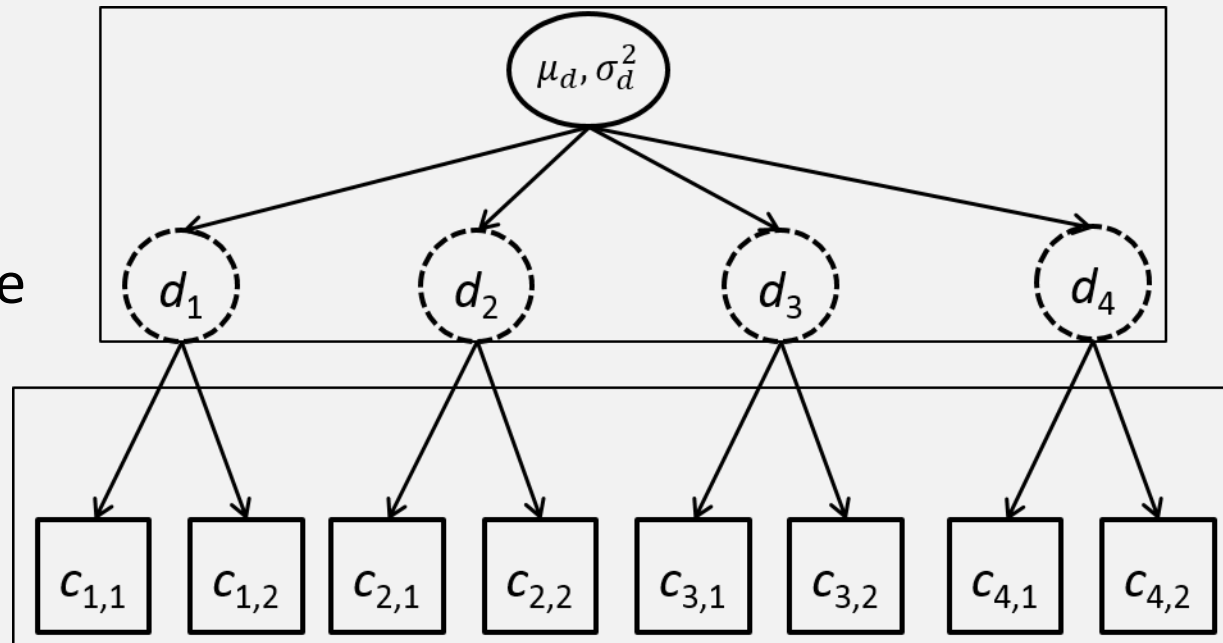
# Mixed-effects models

Example – Hierarchical count samples

$$\log(d_j) \sim \text{Normal}(\mu_d, \sigma_d^2)$$

$$c_{i,j} \sim \text{Poisson}(d_j)$$

- Counts
  - 4 sites
  - 2 observations/site
  - 3 fixed effects
  - 4 random effects



# *Mixed-effects models*

- Simulating data
  - [See R code]

# Mixed-effects models

## Fit using R

- Using *lme4* package
- *formula*: way to specify model

### 1. Linear model – *lm(formula= ... )*

- $\text{Count} \sim 0 + \text{factor}(\text{Site})$
- “Count” – response variable
- “0” – Don’t include intercept
- “factor(Site)” – Include a fixed effect for each site

### 2. Linear mixed model – *lm(formula = ... | ... )*

- $\text{Count} \sim ( 1 \mid \text{factor}(\text{Site}))$
- “( 1 | factor(Site) )” – Include a random effect for each site

# *Mixed-effects models*

**Fit using R**

– [See R code]

# Mixed-effects models

## Fit using TMB

### Steps during optimization

1. Write joint log-likelihood  $\Pr(y, \varepsilon | \theta)$  in CPP file

$$f(\theta, \varepsilon) = \log(\Pr(y | \theta_1, \varepsilon) \Pr(\varepsilon | \theta_2))$$

2. Choose initial values for fixed  $\theta_0$  and random  $\varepsilon_0$
3. “Inner optimization” – Optimize random effects with  $\theta_0$  held constant

$$\hat{\varepsilon} = \operatorname{argmax}_{\varepsilon} (f(\theta_0, \varepsilon))$$

4. Calculate Laplace approx. for marginal likelihood of fixed effects

$$\ln L(\theta_0; y) \cong f(\theta_0, \hat{\varepsilon}) - \frac{1}{2} \log(|\mathbf{H}|)$$

- TMB also provides the gradient of the penalized likelihood with respect to fixed effects
5. “Outer optimization” – Repeat steps 2-3
- Outer optimization is done in R using the function value and gradient provided by TMB



# *Mixed-effects models*

## **Fit using TMB**

[See R code]

# *Mixed-effects models*

- Benefits of using linear mixed models
  - Separate estimate of measurement and between-site variability
  - Include covariates for either one
  - Improved precision
  - “Shrinkage”
- Draw-backs
  - Biased if random effects aren’t “exchangeable”

# Mixed-effects models

## Restricted maximum likelihood models (REML)

- Maximum likelihood (ML) estimates of variance parameters are biased

- ML estimate  $\hat{\sigma}_{ML}^2$

$$\hat{\sigma}_{ML}^2 = \frac{1}{n_i} \sum_{i=1}^{n_i} (y - \hat{\mu}_i)^2$$

- Expectation

$$\hat{\sigma}_{unbiased}^2 = \frac{1}{n_i - 1} \sum_{i=1}^{n_i} (y - \hat{\mu}_i)^2$$

- Same problem arises for variance estimates of random effects
- REML gives unbiased estimates of random-effect variances
  - Also sometimes helps convergence
  - Important when log-likelihood function is correlated with respect to random and fixed effects

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## Confidence interval:

- Parameter estimates are normally distributed
- Computation

$$CI_{x\%}(\hat{\theta}) = \hat{\theta} \pm \widehat{SE}(\hat{\theta}) \times \Phi^{-1}\left(\frac{x}{2}\right)$$

- Where  $CI_{x\%}$  contains the true value  $x\%$  of the time if the model is correct
- $\Phi^{-1}$  is the inverse cumulative distribution for a normal distribution
- $\hat{\theta}$  is the estimate for parameter  $\theta$
- $\widehat{SE}(\hat{\theta})$  is the estimated standard error for parameter  $\theta$

## Confidence interval coverage

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- *Coverage* – the expected proportion of times that an estimated  $x\%$  confidence interval contains the true value given an estimation model and true “data-generating process”

### *Estimation:*

1. Simulate data with a known value for parameter  $\theta$
2. Record true parameter values
3. Apply estimator
4. Record confidence interval  $CI_{x\%}(\hat{\theta})$  for parameter  $\theta$
5. Repeat steps 1-4 hundreds of times
6. Compute the proportion of times where  $CI_{x\%}(\hat{\theta})$  contains the true value for parameter  $\theta$

## Separability

- What if different components of the model are statistically independent?

$$\Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2) = \prod_{i=1}^N \Pr(y|\theta_1, \varepsilon_i) \Pr(\varepsilon_i|\theta_2)$$

- Examples:

- Overdispersed samples

$$C_i \sim \text{Poisson}(\lambda_i)$$

$$\log(\lambda_i) \sim \text{Normal}(\mu, \sigma^2)$$

- Each  $\lambda_i$  is independent conditional on  $\mu, \sigma^2$

$$\Pr(C|\lambda) \Pr(\lambda|\mu, \sigma^2) = \prod_{i=1}^N \Pr(C_i|\lambda_i) \Pr(\lambda_i|\mu, \sigma^2)$$

## **Separability**

- Then we can factor the integral

$$\int \Pr(y|\theta_1, \varepsilon) \Pr(\varepsilon|\theta_2) d\varepsilon = \prod_{i=1}^N \int \Pr(y|\theta_1, \varepsilon_i) \Pr(\varepsilon_i|\theta_2) d\varepsilon_i$$

- Where we replace a  $N$ -dimensional integral with  $N$  1-dimensional integrals

## **Uses**

1. Meta-analysis: species are often independent
2. Time series: years are often “conditionally” independent

# *Mixed-effects models*

[Explore “map” argument to TMB]