Indicators: Now With Pair-wise Flavor!

• Recall I_i is indicator variable for event A_i when:

$$I_i = \begin{cases} 1 & \text{if } A_i \text{ occurs} \\ 0 & \text{otherwise} \end{cases}$$

• Let X = # of events that occur: $X = \sum_{i=1}^{n} I_i$

$$E[X] = E\left[\sum_{i=1}^{n} I_{i}\right] = \sum_{i=1}^{n} E[I_{i}] = \sum_{i=1}^{n} P(A_{i})$$

- Now consider pair of events A_i A_j occurring
 - $I_i I_j = 1$ if both events A_i and A_j occur, 0 otherwise
 - Number of pairs of events that occur is $\binom{X}{2} = \sum_{i \le j} I_i I_j$

From Event Pairs to Variance

Expected number of pairs of events:

$$E\begin{bmatrix} X \\ 2 \end{bmatrix} = E\begin{bmatrix} \sum_{i < j} I_i I_j \end{bmatrix} = \sum_{i < j} E[I_i I_j] = \sum_{i < j} P(A_i A_j)$$

$$E\begin{bmatrix} \frac{X(X-1)}{2} \end{bmatrix} = \frac{1}{2} (E[X^2] - E[X]) = \sum_{i < j} P(A_i A_j)$$

$$E[X^2] - E[X] = 2\sum_{i < j} P(A_i A_j) \implies E[X^2] = 2\sum_{i < j} P(A_i A_j) + E[X]$$

• Recall: $Var(X) = E[X^2] - (E[X])^2$

$$Var(X) = 2\sum_{i < j} P(A_i A_j) + E[X] - (E[X])^2$$

$$= 2\sum_{i < j} P(A_i A_j) + \sum_{i=1}^n P(A_i) - \left(\sum_{i=1}^n P(A_i)\right)^2$$

Let's Try It With the Binomial

$$E[X] = \sum_{i=1}^{n} P(A_i) = np$$

- Each trial: $X_i \sim Ber(p)$ $E[X_i] = p$
- Let event A_i = trial i is success (i.e., X_i = 1)

$$E\begin{bmatrix} X \\ 2 \end{bmatrix} = \sum_{i < j} E[X_i X_j] = \sum_{i < j} P(A_i A_j) = \sum_{i < j} p^2 = \binom{n}{2} p^2$$

$$E\begin{bmatrix} X(X-1) \\ 2 \end{bmatrix} = \frac{1}{2} (E[X^2] - E[X]) = \frac{n(n-1)}{2} p^2$$

$$Var(X) = E[X^2] - (E[X])^2 = (E[X^2] - E[X]) + E[X] - (E[X])^2$$

 $= n(n-1)p^{2} + np - (np)^{2} = n^{2}p^{2} - np^{2} + np - n^{2}p^{2}$

$$= np(1-p)$$

Computer Cluster Utilization

- Computer cluster with k servers
 - Requests independently go to server i with probability p_i
 - Let event A_i = server i receives no requests
 - X = # of events $A_1, A_2, ..., A_k$ that occur
 - Y = # servers that receive \geq 1 request = k X
 - E[Y] after first n requests?
 - Since requests independent: $P(A_i) = (1 p_i)^n$

$$E[X] = \sum_{i=1}^{k} P(A_i) = \sum_{i=1}^{k} (1 - p_i)^n$$

$$E[Y] = k - E[X] = k - \sum_{i=1}^{k} (1 - p_i)^n$$

when
$$p_i = \frac{1}{k}$$
 for $1 \le i \le k$, $E[Y] = k - \sum_{i=1}^k (1 - \frac{1}{k})^n = k \left(1 - (1 - \frac{1}{k})^n\right)^n$

Computer Cluster Utilization (cont.)

- Computer cluster with k servers
 - Requests independently go to server i with probability p_i
 - Let event A_i = server i receives no requests
 - X = # of events $A_1, A_2, ..., A_k$ that occur
 - Y = # servers that receive \geq 1 request = k X
 - Var(Y) after first n requests? (= (-1)² Var(X) = Var(X))
 - Independent requests: $P(A_i A_i) = (1 p_i p_i)^n$, $i \neq j$

$$E[X(X-1)] = E[X^{2}] - E[X] = 2\sum_{i < j} P(A_{i}A_{j}) = 2\sum_{i < j} (1 - p_{i} - p_{j})^{n}$$

$$Var(X) = 2\sum_{i < j} (1 - p_i - p_j)^n + E[X] - (E[X])^2 \qquad E[X] = \sum_{i=1}^k (1 - p_i)^n$$

$$=2\sum_{i< j}(1-p_i-p_j)^n+\sum_{i=1}^k(1-p_i)^n-\left(\sum_{i=1}^k(1-p_i)^n\right)^2=\mathrm{Var}(Y)$$

Computer Cluster = Coupon Collecting

- Computer cluster with k servers
 - Requests independently go to server i with probability p_i
 - Let event A_i = server i receives no requests
 - X = # of events $A_1, A_2, ..., A_k$ that occur
 - Y = # servers that receive \geq 1 request = k X
- This is really another "Coupon Collector" problem
 - Each server is a "coupon type"
 - Request to server = collecting a coupon of that type
- Hash table version
 - Each server is a bucket in table
 - Request to server = string gets hashed to that bucket

Product of Expectations

 Say X and Y are <u>independent</u> random variables, and g(●) and h(●) are real-valued functions

$$E[g(X)h(Y)] = E[g(X)]E[h(Y)]$$

Proof:

$$E[g(X)h(Y)] = \int_{y=-\infty}^{\infty} \int_{x=-\infty}^{\infty} g(x)h(y)f_{X,Y}(x,y) dx dy$$

$$= \int_{y=-\infty}^{\infty} \int_{x=-\infty}^{\infty} g(x)h(y)f_X(x)f_Y(y) dx dy$$

$$= \int_{x=-\infty}^{\infty} g(x)f_X(x) dx \cdot \int_{y=-\infty}^{\infty} h(y)f_Y(y) dy$$

$$= E[g(X)]E[h(Y)]$$

The Dance of the Covariance

- Say X and Y are arbitrary random variables
- Covariance of X and Y:

$$Cov(X,Y) = E[(X - E[X])(Y - E[Y])]$$

Equivalently:

$$Cov(X,Y) = E[XY - E[X]Y - XE[Y] + E[Y]E[X]]$$

$$= E[XY] - E[X]E[Y] - E[X]E[Y] + E[X]E[Y]$$

$$= E[XY] - E[X]E[Y]$$

- X and Y independent, E[XY] = E[X]E[Y] → Cov(X,Y) = 0
- But Cov(X,Y) = 0 does <u>not</u> imply X and Y independent!

Dependence and Covariance

X and Y are random variables with PMF:

Y	-1	0	1	p _Y (y)	
0	1/3	0	1/3	2/3	$Y = \begin{cases} 0 \end{cases}$
1	0	1/3	0	1/3	1 1
$p_{X}(x)$	1/3	1/3	1/3	1	

if $X \neq 0$

otherwise

•
$$E[X] = -1(1/3) + 0(1/3) + 1(1/3) = 0$$

•
$$E[Y] = 0(2/3) + 1(1/3) = 1/3$$

•
$$Cov(X, Y) = E[XY] - E[X]E[Y] = 0 - 0 = 0$$

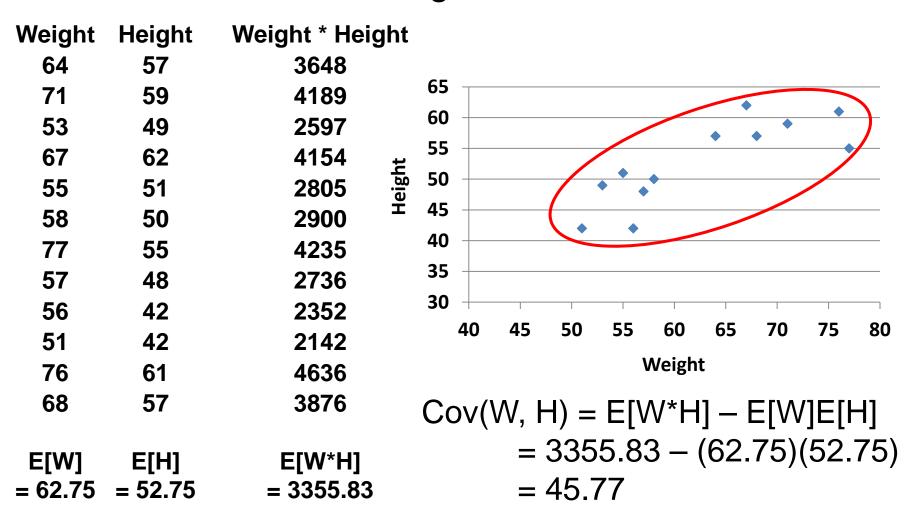
But, X and Y are clearly dependent!

Example of Covariance

- Consider rolling a 6-sided die
 - Let indicator variable X = 1 if roll is 1, 2, 3, or 4
 - Let indicator variable Y = 1 if roll is 3, 4, 5, or 6
- What is Cov(X, Y)?
 - E[X] = 2/3 and E[Y] = 2/3
 - E[XY] = $\sum_{x} \sum_{y} xy \ p(x, y)$ = (0 * 0) + (0 * 1/3) + (0 * 1/3) + (1 * 1/3) = 1/3
 - Cov(X, Y) = E[XY] E[X]E[Y] = 1/3 4/9 = -1/9
 - Consider: P(X = 1) = 2/3 and P(X = 1 | Y = 1) = 1/2
 - Observing Y = 1 makes X = 1 less likely

Another Example of Covariance

Consider the following data:



Properties of Covariance

- Say X and Y are arbitrary random variables
 - Cov(X,Y) = Cov(Y,X)
 - $Cov(X, X) = E[X^2] E[X]E[X] = Var(X)$
 - Cov(aX + b, Y) = aCov(X, Y)
- Covariance of sums of random variables
 - $X_1, X_2, ..., X_n$ and $Y_1, Y_2, ..., Y_m$ are random variables

•
$$Cov \left(\sum_{i=1}^{n} X_i, \sum_{j=1}^{m} Y_j \right) = \sum_{i=1}^{n} \sum_{j=1}^{m} Cov(X_i, Y_j)$$

Variance of Sum of Variables

•
$$\operatorname{Var}\left(\sum_{i=1}^{n} X_{i}\right) = \sum_{i=1}^{n} \operatorname{Var}(X_{i}) + 2\sum_{i=1}^{n} \sum_{j=i+1}^{n} \operatorname{Cov}(X_{i}, X_{j})$$

Proof:

$$\operatorname{Var}\left(\sum_{i=1}^{n} X_{i}\right) = \operatorname{Cov}\left(\sum_{i=1}^{n} X_{i}, \sum_{j=1}^{n} X_{j}\right)$$

$$= \sum_{i=1}^{n} \sum_{j=1}^{n} \operatorname{Cov}(X_{i}, X_{j})$$

$$= \sum_{i=1}^{n} \operatorname{Var}(X_i) + \sum_{i=1}^{n} \sum_{i=1}^{n} \operatorname{Cov}(X_i, X_j)$$

$$= \sum_{i=1}^{n} \text{Var}(X_i) + 2\sum_{i=1}^{n} \sum_{j=i+1}^{n} \text{Cov}(X_i, X_j)$$

• If all
$$X_i$$
 and X_j independent $(i \neq j)$: $Var\left(\sum_{i=1}^n X_i\right) = \sum_{i=1}^n Var(X_i)$

Note: Cov(X, X) = Var(X)

Note. $Cov(\Lambda, \Lambda) = var(\Lambda)$

By symmetry:

 $Cov(X_i, X_j) = Cov(X_j, X_i)$

Hola Compadre: La Distribución Binomial

- Let Y ~ Bin(n, p)
 - n independent trials
 - Let $X_i = 1$ if *i*-th trial is "success", 0 otherwise
 - $X_i \sim Ber(p)$ $E[X_i] = p$
 - $Var(Y) = Var(X_1) + Var(X_2) + ... + Var(X_n)$
 - $Var(X_i) = E[X_i^2] (E[X_i])^2$ = $E[X_i] - (E[X_i])^2$ since $X_i^2 = X_i$ = $p - p^2 = p(1 - p)$
 - $Var(Y) = nVar(X_i) = np(1 p)$

Variance of Sample Mean

- Consider n I.I.D. random variables X₁, X₂, ... X_n
 - X_i have distribution F with $E[X_i] = \mu$ and $Var(X_i) = \sigma^2$
 - We call sequence of X_i a <u>sample</u> from distribution F
 - Recall sample mean: $\overline{X} = \sum_{i=1}^{n} \frac{X_i}{n}$ where $E[\overline{X}] = \mu$
 - What is $Var(\overline{X})$?

$$\operatorname{Var}(\overline{X}) = \operatorname{Var}\left(\sum_{i=1}^{n} \frac{X_{i}}{n}\right) = \left(\frac{1}{n}\right)^{2} \operatorname{Var}\left(\sum_{i=1}^{n} X_{i}\right)$$

$$= \left(\frac{1}{n}\right)^{2} \sum_{i=1}^{n} \operatorname{Var}(X_{i}) = \left(\frac{1}{n}\right)^{2} \sum_{i=1}^{n} \sigma^{2} = \left(\frac{1}{n}\right)^{2} n \sigma^{2}$$

$$= \frac{\sigma^{2}}{n}$$

Sample Variance

- Consider n I.I.D. random variables X₁, X₂, ... X_n
 - X_i have distribution F with $E[X_i] = \mu$ and $Var(X_i) = \sigma^2$
 - We call sequence of X_i a <u>sample</u> from distribution F
 - Recall sample mean: $\overline{X} = \sum_{i=1}^{n} \frac{X_i}{n}$ where $E[\overline{X}] = \mu$
 - Sample deviation: $\overline{X} X_i$ for i = 1, 2, ..., n
 - Sample variance: $S^2 = \sum_{i=1}^n \frac{(X_i X)^2}{n-1}$
 - What is $E[S^2]$?
 - $E[S^2] = \sigma^2$
 - We say S^2 is "unbiased estimate" of σ^2

Proof that $E[S^2] = \sigma^2$ (just for reference)

$$E[S^{2}] = E\left[\sum_{i=1}^{n} \frac{(X_{i} - \overline{X})^{2}}{n - 1}\right] \Rightarrow (n - 1)E[S^{2}] = E\left[\sum_{i=1}^{n} (X_{i} - \overline{X})^{2}\right]$$

$$(n - 1)E[S^{2}] = E\left[\sum_{i=1}^{n} (X_{i} - \overline{X})^{2}\right] = E\left[\sum_{i=1}^{n} ((X_{i} - \mu) + (\mu - \overline{X}))^{2}\right]$$

$$= E\left[\sum_{i=1}^{n} (X_{i} - \mu)^{2} + \sum_{i=1}^{n} (\mu - \overline{X})^{2} + 2\sum_{i=1}^{n} (X_{i} - \mu)(\mu - \overline{X})\right]$$

$$= E\left[\sum_{i=1}^{n} (X_{i} - \mu)^{2} + n(\mu - \overline{X})^{2} + 2(\mu - \overline{X})\sum_{i=1}^{n} (X_{i} - \mu)\right]$$

$$= E\left[\sum_{i=1}^{n} (X_{i} - \mu)^{2} + n(\mu - \overline{X})^{2} + 2(\mu - \overline{X})n(\overline{X} - \mu)\right]$$

$$= E\left[\sum_{i=1}^{n} (X_{i} - \mu)^{2} - n(\mu - \overline{X})^{2}\right] = \sum_{i=1}^{n} E[(X_{i} - \mu)^{2}] - nE[(\mu - \overline{X})^{2}]$$

$$= n\sigma^{2} - nVar(\overline{X}) = n\sigma^{2} - n\frac{\sigma^{2}}{n} = n\sigma^{2} - \sigma^{2} = (n - 1)\sigma^{2}$$

• So, $E[S^2] = \sigma^2$