

1.)

x_1, \dots, x_n ARE IID GAUSSIAN.

unknown mean θ

known variance σ^2

(A) $\theta \sim N(\mu, \sigma^2)$,

So for a single observation, probability is,

$$P(x_i | \theta) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x_i - \theta)^2}{2\sigma^2}}$$

as x_1, x_2, \dots, x_n are independent,

For n observations, in a dataset D : -

$$P(D | \theta) = \prod_{i=1}^n P(x_i | \theta)$$

$$= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x_i - \theta)^2}{2\sigma^2}}$$

$$P(\theta) = \frac{1}{\sqrt{2\pi\sigma^2}} \times e^{-\frac{(\theta - \mu)^2}{2\sigma^2}}$$

Maximum posterior,

$$\theta_{\text{MAP}} = \operatorname{argmax} \{ P(D|\theta) \cdot P(\theta) \}$$

$$= \operatorname{argmax} \left\{ \prod_{i=1}^n P(x_i|\theta) \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(\theta-\mu)^2}{2\sigma^2}} \right\}$$

$$= \operatorname{argmax} \left\{ \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma_0^2}} e^{-\frac{(x_i-\theta)^2}{2\sigma_0^2}} \times \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(\theta-\mu)^2}{2\sigma^2}} \right\}$$

$$= \operatorname{argmax} \left\{ \left(\frac{1}{\sqrt{2\pi\sigma_0^2}} \right)^n e^{-\sum_{i=1}^n \frac{(x_i-\theta)^2}{2\sigma_0^2}} \times \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(\theta-\mu)^2}{2\sigma^2}} \right\}$$

$$= \operatorname{argmax} \left\{ \frac{1}{(2\pi)^{\frac{n+1}{2}} \sigma_0^n \sigma} e^{-\sum_{i=1}^n \frac{(x_i-\theta)^2}{2\sigma_0^2}} e^{-\frac{(\theta-\mu)^2}{2\sigma^2}} \right\}$$

Taking log (to smoothe) & log is monotonically increasing.

$$\begin{aligned} \log(P(D|\theta) \cdot P(\theta)) &= \log \left(\frac{1}{(2\pi)^{\frac{n+1}{2}} \sigma_0^n \sigma} \right) \\ &\quad - \sum_{i=1}^n \frac{(x_i-\theta)^2}{2\sigma_0^2} \\ &\quad - \frac{(\theta-\mu)^2}{2\sigma^2} \end{aligned}$$

Maximizing by derivation

$$\frac{\partial}{\partial \theta} [\log (P(D|\theta) \cdot P(\theta))] = \frac{\partial}{\partial \theta} \left[\log \left(\frac{1}{(2\pi)^{n/2} \sigma_0^2} e^{-\sum_{i=1}^n \frac{(x_i - \theta)^2}{2\sigma_0^2}} \right) \right]$$

$$= \sum_{i=1}^n \frac{(x_i - \theta)}{\sigma_0^2} - \frac{(\theta - \mu)}{\sigma_0^2}$$

Now equating to zero: -

$$\sum_{i=1}^n \frac{(x_i - \theta)}{\sigma_0^2} - \frac{(\theta - \mu)}{\sigma_0^2} = 0$$

$$\sum_{i=1}^n \frac{(x_i - \theta)}{\sigma_0^2} = \frac{\theta - \mu}{\sigma_0^2}$$

$$\sigma_0^2 \sum_{i=1}^n (x_i - \theta) = \sigma_0^2 (\theta - \mu)$$

$$\sigma_0^2 \left(\sum_{i=1}^n x_i - n\theta \right) = \sigma_0^2 \theta - \sigma_0^2 \mu$$

$$\sigma_0^2 \sum_{i=1}^n x_i - \sigma_0^2 n\theta = \sigma_0^2 \theta - \sigma_0^2 \mu$$

$$\sum_{i=1}^n x_i \sigma_0^2 + \sigma_0^2 \mu = \sigma_0^2 \theta + \sigma_0^2 n\theta$$

$$\sum_{i=1}^n x_i \sigma_0^2 + \sigma_0^2 \mu = \theta (\sigma_0^2 + \sigma_0^2 n)$$

$$\therefore \theta = \frac{\sigma_0^2 \mu + \sigma^2 \sum_{i=1}^n (x_i)}{\sigma_0^2 + \sigma^2 n}$$

$$\therefore \theta_{\text{MAP}} = \frac{\sigma_0^2 \mu + \sigma^2 \sum_{i=1}^n x_i}{\sigma_0^2 + \sigma^2 n}$$

(b)

$$P(\theta) = \frac{1}{2b} e^{\frac{-(\theta - \mu)}{b}}$$

By MAP as we did previously,

$$\sum_{i=1}^n \frac{(x_i - \theta)}{\sigma_0^2} + \frac{(\theta - \mu)}{(\theta - \mu)b} = 0$$

$$\therefore \sum_{i=1}^n \frac{(x_i - \theta)}{\sigma_0^2} = - \frac{(\theta - \mu)}{(\theta - \mu)b} \quad \text{--- (1)}$$

Since it is assumed that $\mu = 0$
we can rewrite the above equation (1)
as:

$$\sum_{i=1}^n \frac{(x_i - \theta)}{\sigma_0^2} = - \frac{\theta}{(\theta)b}$$

$$\sum_{i=1}^n (x_i) - n\theta = \frac{(\text{sign } \theta) \sigma_0^2}{b}$$

$$\sum_{i=1}^n (x_i) + \frac{(\text{sign } \theta) \sigma_0^2}{b} = n\theta$$

$$\theta = \frac{\sum_{i=1}^n (x_i) + \frac{\text{sign } (\theta) \sigma_0^2}{b}}{n}$$

(c) $\theta \sim N(0, \Sigma = \sigma^2 I)$,

$$\theta_{MAP} = \text{argmax}_{\theta} \left\{ \frac{1}{\sqrt{2\pi} |\Sigma_0|^{n/2}} e^{(-1/2 (-\theta + x)_0^T \Sigma_0^{-1} (-\theta + x)_0)} \right. \\ \left. \times \frac{1}{(2\pi)^{n/2} \sigma^2} e^{(-1/2 \sigma^{-2} \theta^T \theta)} \right\}$$

Applying log:—

$$\log \left(\frac{1}{(2\pi)^{n/2} (2\pi)^{n/2} |\Sigma_0|^{n/2} \sigma^2} \right) - \\ - 1/2 (x - \theta)_0^T \Sigma_0^{-1} (x - \theta)_0 \\ - 1/2 \sigma^{-2} \theta^T \theta$$

So, now apply derivative and equate it to zero: -

$$\frac{\partial}{\partial \theta} \left(\frac{1}{2} (x - \theta)^T \Sigma_0^{-1} (x - \theta) \right) + \frac{\partial}{\partial \theta} \left(\frac{1}{2\sigma^2} \theta^T \theta \right) = 0$$

$$\frac{\partial}{\partial \theta} \left(x^T \Sigma_0^{-1} x - x \theta^T \Sigma_0^{-1} - x^T \theta \Sigma_0^{-1} + \theta \theta^T \Sigma_0^{-1} \right) + \frac{\partial}{\partial \theta} \left(\frac{1 \cdot \theta^T \theta}{\sigma^2} \right) = 0$$

$$- \left(x^T \Sigma_0^{-1} \right)^T - x \Sigma_0^{-1} + \Sigma_0^{-1} \theta + \theta \Sigma_0^{-1} + \frac{1}{\sigma^2} \theta = 0$$

$$-\frac{1}{\sigma^2} x \cdot \theta = - \left(x^T \Sigma_0^{-1} \right)^T - \Sigma_0^{-1} x + \Sigma_0^{-1} \theta + \theta \Sigma_0^{-1}$$

$$\frac{1}{\sigma^2} \theta = - \Sigma_0^{-1} x - x \Sigma_0^{-1} + \theta \Sigma_0^{-1} + \Sigma_0^{-1} \theta$$

$$- 2 \Sigma_0^{-1} x + 2 \Sigma_0^{-1} \theta = \frac{-2\theta}{\sigma^2}$$

$$\text{Since } \Sigma_0^{-1} = \Sigma_0^{-1T}$$

$$-2 \Sigma_0^{-1} X = + 2 \Sigma_0^{-1} \theta - \frac{2\theta}{\sigma^2}$$

$$-2 \Sigma_0^{-1} X = \left(2 \Sigma_0^{-1} \cdot \frac{2}{\sigma^2} \right) \theta$$

$$\therefore \theta = \frac{\Sigma_0^{-1} X}{\Sigma_0^{-1} \left(\frac{1}{\sigma^2} \right)}$$