

# BABY STEPS IN TAX POLICY: IMPACTS OF THE FEDERAL ADOPTION TAX CREDIT\*

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## **Abstract**

The federal adoption tax credit is among the largest credits claimable by an individual filer in the US and is one of the primary tools used to facilitate child adoption. A temporary reform in 2010 and 2011 made the credit fully refundable, affording the opportunity to trace the heterogeneous adoption response of taxpayers across the distribution of tax liability. The estimates imply increasing access to the credit was effective at encouraging child adoption, inducing an additional 36,188 adoptions among those impacted by reform and that the average increase in fiscal transfers required to encourage an additional adoption ranges between \$15,500 and \$17,500.

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# 1 Introduction

While most children are cared for by their biological parents, roughly 2.2 percent of children in the United States are adopted according to the US Census. The costs of adoption can be significant, easily exceeding tens of thousands of dollars, and often involves becoming the parent of an especially vulnerable child who may need additional time and attention.<sup>1</sup> To offset these costs, subsidize the care of these children, and encourage more parents to adopt, the US federal government offers a tax credit worth up to \$16,810 per adoption as of 2025. Despite its high nominal amount, the credit has been completely nonrefundable for most of its history<sup>2</sup>, meaning the credit was only usable for households to the extent that they had positive tax liability to offset. The implications of this design are clear. First, households with low levels of tax liability are offered less incentive to adopt relative to households with higher tax liability. Second, those with low levels of tax liability who choose to adopt, along with their adoptive children, receive a lower benefit for doing so. Whether this design is optimal depends on a number of key considerations, but likely of first order importance, although certainly not exclusively, is the responsiveness of these different kinds of households to the incentive, the kinds of adoptions the credit ultimately encourages, and the average increase in expenditure associated with encouraging an additional adoption. Between 2010-2011, the credit was made fully refundable, differentially increasing the amount of credit available to households depending on their level of tax liability, affording the opportunity to quantitatively analyze each of these key considerations.

By utilizing the multi-valued difference in differences design suggested by Callaway et al. (2024), this paper compares the adopting behavior of those who gained access to different levels of additional credit to those who experienced only a small increase. The reform’s provisions enable the estimation of its level impact on adoptions, the average increase in expenditure required to encourage an additional adoption, and a partial characterization of the kinds of adoptions the credit encourages, across the distribution of tax liability. To make these comparisons, this paper introduces a new procedure that utilizes the demographic characteristics located in publicly available Census data to impute whether an adopted child is a new adoption, enabling the linkage of adopting behavior to contemporaneous parental and adopted child characteristics.<sup>3</sup> The estimates imply the reform increased adoptions

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<sup>1</sup>Despite these barriers, survey evidence suggests that roughly 1/3 of families have considered adoption at some point in their lives (“US Adoption Statistics — Adoption Network”, 2020).

<sup>2</sup>The credit was first instituted in its current form in 2002 and was worth \$10,000. It was set to automatically adjust for inflation annually.

<sup>3</sup>This data has been hitherto not used because it does not distinguish between step-parent adoptions and other adoption types. It also does not record when an adoption was completed. This measure imputes the latter and is likely robust to the former. While other national surveys have been done in the past, most

for those with zero tax liability by 58.6 percent and 46.3 percent for those with positive liability, relative to their pre-treatment averages, implying an average partial elasticity of 4 percent per \$1000 of additional credit for the 0 liability group and 5.6 percent per \$1000 of additional credit for the positive liability group. In total, the reform induced the treated groups to adopt 36,188 additional children than they would have otherwise in the absence of the refundability reform, and the estimates suggests these tended to be domestic adoptions of children aged less than 2.<sup>4</sup> The average additional expenditure required to encouraged an additional adoption varied across the distribution of tax liability<sup>5</sup>, costing on average \$15,801 for those with zero liability, \$15,533 for those with positive liability, and \$17,425 for those with \$1-\$5,000 of liability. If a policy maker were seeking to simply maximize adoptions among these groups at a minimum cost, targeting lower and higher liability groups may be more cost effective than targeting those in the center. Of course, the value of these expenditures ought not be conceptualized as merely a cost, given that they are transfers to households. For example, the fact that those with lower income and hence lower tax liability are typically more inclined to adopt from foster care implies an increase in expenditures to them may have a significantly high social marginal value, even if the increase resulted in no increase in adoptions.

While this paper concentrates on the effectiveness of adoption incentives on all adoption, these estimates are intuitively commensurate with other studies that have primarily focused on public adoptions (Buckles, 2013; Brehm, 2021; Doyle and Peters, 2007; Hansen, 2007).<sup>6</sup> Of particular relevance to this paper, Brehm (2021) estimates the impact of the 2010 adoption tax credit reform on adoptions from foster care, estimating that the credit’s imminent expiration induced an increase of roughly 2,400 additional public adoptions in the month of December 2011, constituting a 44% increase, roughly scaling with the estimates presented in this paper.<sup>7</sup>

Additionally, studies that have examined the relationship between household characteristics and other adoption types beyond public adoptions broadly corroborate the descriptive evidence found using this paper’s measure of new adoptions. This provides important validation that the measure presented in this paper tracks actual adopting behavior. Baccara

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notable the National Survey of Adoptive Parents, this imputation will allow future researchers to connect adopting behavior with the wealth of information available in the Census.

<sup>4</sup>These estimates do not suggest older children did not also benefit from reform, only that the measure used will most reliably detect the adoption of younger children.

<sup>5</sup>The average fiscal expenditure is how much tax expenditures have to increase to encourage an additional adoption, given there are household who would have adopted and claimed the credit regardless.

<sup>6</sup>While there is much more data on children adopted from foster care, very little information is recorded on the economic characteristics of the households conducting the adoption.

<sup>7</sup>The adoptions studied here all adoptions and the estimates are at the annual level, and so one should expect comparable magnitudes, but not the literal estimates to be the same.

et al. (2014) using data from a private adoption facilitator and matching theory estimate a higher willingness to pay for girls and white children relative to boys and black children, which tracks the aggregate finding of a higher proportion of new adoptions being composed of girls and a higher proportion being white at higher income levels. Similar findings are documented by Khun et al. (2020), which utilizes the National Survey of Adoptive Parents<sup>8</sup> and variation in adoption cost to estimate the rate of substitution of adoption types between 1990-2008 using a logit framework.

This paper is the first to examine how exogenous variation in financial incentives impacts adopting behavior across all households. Taken as a whole, this paper provides new evidence that financial incentives are an effective means of encouraging child adoption and that taxpayers across the distribution of tax liability are responsive to them.

## 2 Adoption in the US and the Federal Adoption Tax Credit

The process of adopting and the costs involved are highly heterogeneous and depend on the type of adoption being performed, the jurisdictions under which the adoptive parent resides, and the jurisdictions regulating the welfare of the adoptive child. Adoptions done through a private agency domestically or internationally can have expenses ranging from \$15,000 to \$50,000 (Malm and Vandivere, 2009). These avenues primarily result in the adoption of younger children, with more than 2/3 of these being less than a year old. These children are less likely to be diagnosed with a disability, are whiter, and more likely to be girls relative to the general population (Malm and Vandivere, 2009). In contrast, adoptions from foster care tend to have nearly zero upfront costs, with states and counties often subsidizing their care, even after the adoption is finalized (Bald et al., 2022). These children are much more likely to have a disability, are older, and come from more vulnerable backgrounds. In 2009, 45 percent of the children adopted from foster care lived in households with income less than twice the poverty threshold (Malm and Vandivere, 2009).

The credit is designed to encourage both private adoptions and adoptions from foster care. Passed in its current form in 2001 under the The Economic Growth and Tax Relief Reconciliation Act, the nonrefundable credit amount was set at \$10,000 per adoption and is recalculated every year to adjust for inflation. The credit was set to begin phasing out after \$150,000 in MAGI<sup>9</sup>, reducing proportionally to \$0 at \$190,000.<sup>10</sup> Interestingly, while

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<sup>8</sup>This is a nationally representative survey conducted in 2008, that contains a wealth of information on adoptive parents and children.

<sup>9</sup>A filer's MAGI is their AGI, plus additional sources of income that were excluded from their AGI, such as IRA contributions and student loan interest. For most taxpayers, AGI and MAGI are roughly equivalent. Instructions can be found on form 8839.

<sup>10</sup>In 2025, the credit amount is \$16,810 and the lower bound is set at a MAGI of \$252,150 with an upper

the lower bound is set to adjust for inflation yearly, the statutory upper bound—the MAGI level where the claimable credit becomes 0—is fixed at a nominal distance of \$40,000 above the lower bound<sup>11</sup>, implying that in real terms, the credit’s phaseout region becomes more steep over time. In the case of international and private domestic adoptions, the credit may be used for eligible adoption expenses, which include all costs associated with the process of adopting a child such as legal services, home study fees, and the cost of travel. In the case of “special needs” adoptions, which mostly comprise adoptions from foster care, the full credit may be claimed regardless of adoption expenses.<sup>12</sup> The credit is claimable the year the adoption is finalized or, if the adoption is private and domestic, the year after the expense is incurred. The credit possesses a carry forward of up to five years, implying it is usable for adoption expenses so long as a taxpayer has enough liability within that time horizon.<sup>13</sup>

## 2.1 The Reform

In 2010, under the Affordable Care Act, the tax credit was made fully refundable and increased by \$1,000. This means that taxpayers with zero tax liability who would have previously not been able to benefit from the credit, now qualified for up to \$13,360 to offset adoption expenses. Then, on January 1st 2012, the credit reverted back to its non-refundable state, and the \$1,000 increase was also allowed to expire.

The reform led to increases in both tax expenditures on the credit and the number of taxpayers who filed for it. From 2009 to 2010, tax expenditures on the credit rose from 278 million in 2009 to 1.2 billion in 2010, and the number of filers who claimed the credit increased from 80,676 in 2009 to 140,676 in 2010. While expenditures were also higher in 2011, 668.1 million relative to the pre-reform period, the number of filers declined to 51,539. Part of the reason for this may be due to the decreased value of the carry-forward provision, given the credit was made refundable. More concretely, under the regime in which the credit is nonrefundable, a fraction of the people filing for the credit in a given year are made up of individuals who incurred adoption expenses less than 5 years prior, but did not have enough tax liability to make full use of it. When the credit became refundable in 2010, tax liability no longer was a constraint, and so, unlike tax year 2009 and 2010, the filings in 2011 do

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bound at \$292,150.

<sup>11</sup>In other words, the point at which 0 credit is claimable is simply the point at which the phaseout region begins, which is adjusted for inflation annually, plus \$40,000.

<sup>12</sup>Each state has their own specific definition of special needs. Nonetheless, the vast majority of children are given this designation. In 2013, this included 91 percent of foster children Brehm, 2021

<sup>13</sup>In addition, adoption expenses paid for through qualified adoption assistance provided by an employer may not be covered by the credit and reduces the filer’s maximum credit amount. Special needs adoptions provide an exception to this.

not represent as many individuals carrying forward credit from the year previous. While the average filer in every income bracket received additional credit, the largest proportional increases occurred for those making under \$50,000. From 1997 to 2009, 10 percent of tax expenditures for the credit were attributable to those making below 50,000. In 2010-2011, this expenditure share increased to 46 percent. In the same years, those making less than 15,000 received 0 percent of tax expenditures, and in 2010 to 2011, the share increased to 13 percent (Brehm, 2021). These statistics suggest the reform substantially increased the usability of the credit for those with lower tax liability in addition to providing a sizable cash transfers for those already benefiting from it.

A full accounting of the potential social welfare consequences of this policy requires the evaluation of a number of additional considerations beyond its impact on the number of adoptions. Even if the credit induced no change in behavior<sup>14</sup>, whether or not the policy induced a net gain in social welfare hinges on the relative social benefit of the fiscal transfer to these families, the social value of public funds lost due to the increase in tax expenditure, and the resulting associated fiscal externalities, which includes the present value of future fiscal savings or losses it induced (Hendren and Sprung-Keyser, 2020).

Despite the large increase in tax expenditure associated with the policy, there is evidence that cash infusions to those with children and to those with foster children in particular can result in increased expenditures on childcare and improved educational outcomes (Buckles, 2013; Simon et al., 2024), which in turn could offset these costs in the form of higher receipts in the future (Hendren and Sprung-Keyser, 2020). Even more directly, the reform disproportionately increased payments to those with very low incomes, positively increasing the credit’s redistributive effect. If the social benefit of transferring these funds to these families is high enough, or the present social value of expenditure savings induced by them is high enough, even if the policy induced no change in adopting behavior, it could have resulted in a net gain in social welfare.<sup>15</sup> My current analysis does not speak to these considerations, but they are vital to evaluating the merits of the policy.<sup>16</sup>

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<sup>14</sup>Note: this is quite different than inducing no change the number of adoptions.

<sup>15</sup>There are great number of additional considerations that the present analysis also does not speak to. For example, if adopting behavior truly was inelastic in this way, it could serve as a tag for higher earning potential, given the home study that is often required (Akerlof, 1978).

<sup>16</sup>A note on administrative burden: During this reform period, the IRS, and hence filers as well, suffered a large increase in administrative burden. The complexity of the credit’s income limitations, the different state definition for “special needs”, the large amount of documentation required to validate expenses, and the large sums at stake induced an audit rate of 69 percent for adoption tax credit claims in the 2012 filing season. 55 percent of these resulted in no change in taxes owed or refund due, with the average correspondence audit taking 126 days (Bogadi, 2012). For 2011, only \$11 million of the 668.1 expenditure was disallowed, with the IRS having to pay 2.1 million in interest for withholding refunds due over 45 days. For any analysis of welfare, the administrative costs and increased burden of reform must be accurately factored in (Keen and

### 3 Data

The Census’s American Community Survey 1 % samples for the years 2008 to 2022 is utilized. This is a yearly, cross-sectional random sample of the US population. It contains detailed self-reported information on income, family structure, and demographic characteristics, all of which are vital for imputing federal tax liability and inferring adopting behavior. For all estimates and descriptive statistics, the ACS’s household survey weights are used so that the results are nationally representative, unless stated otherwise.

#### 3.1 Imputing Tax Liability

While the ACS does not directly provide tax data, the tax liability of each household may be imputed by utilizing the wealth of economic variables contained in the ACS paired with the National Bureau of Economic Research’s (NBER) TaxSim.<sup>17</sup> This program takes as inputs household-level income variables and, utilizing the state and federal income tax schedules, computes a variety of tax-relevant household-level characteristics. In order to calculate each household’s level of exposure to reform, proxied by the increase in credit access each household experiences due to the credit becoming refundable, reasonable estimates of each household’s modified adjusted gross income and their federal tax liability is necessary. One interesting silver lining of the American Community survey’s sample design is that income variables are reported for the year prior, while demographic variables are contemporaneous. What this means is if a household were induced to adopt within a given year due to reform and subsequently endogenously change their tax liability that year, so long as in the absence of treatment their liability the year prior would have been roughly the same, their true level of exposure to the refundability reform will be computed accurately.<sup>18</sup>

In order to reduce the number of assumptions that must be made in imputation, each household’s federal tax liability before credits assuming the standard deduction is used. The reasons for this are twofold. First, and most importantly, the ACS does not record take-up of credits and does not record deductible expenditures for each household. Such an imputation would require overly strong assumptions on household behavior. Secondly, while in theory the credits and deductions each household chooses to utilize could be imputed, using a household’s federal tax liability before credits measures their maximum capacity to make use of the non-refundable version of the credit, which is not contaminated by the

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Slemrod, 2017). Although presently outside the scope of this analysis, these considerations will be vital for deciding whether it is optimal to re-implement this reform in the future.

<sup>17</sup>While the ACS provides their own estimates of tax liability using TaxSim, they do not provide values before 2009. The estimates are roughly equivalent and if used, do not change the results

<sup>18</sup>Of course, panel data would be used in the ideal scenario, as this would allow the use of lagged income outside the treatment time. This isn’t possible given the ACS’s structure

endogenous take up response of other credits. This best captures the underlying motivation for the definition of treatment: being offered additional credit, abstracting from the take up responses of other credits that may be available to the household.

There are some additional important simplifications noted here. First, the Census does not report whether households file jointly or separately. Therefore, it is assumed households file separately if they are separated but not divorced, but jointly if they are married even if the spouse is not currently living in the household. This is likely a reasonable assumption, since in 2010 only 2.31% of married people filed separately (“SOI tax stats - Individual income tax returns filed and sources of income — Internal Revenue Service”, n.d.). In addition, while the ACS has data on interest and dividend income, it does not distinguish between them despite the fact they are treated differently by the tax code. It is assumed that dividend income is zero and the remainder is interest income. In addition, the ACS has no information on realized long or short terms capital gains or losses. These are set to zero as well.

To provide an assessment of the accuracy of the imputation, Figure 1 compares the imputed mean average level of federal tax liability before credits, relative to the aggregate average reported by the IRS. The top blue line is the average level of imputed federal tax liability before credits per household among all households imputed as having positive liability. The bottom blue line is the average level of imputed federal tax liability before credits for all households. Since households with very low levels of income are not required to file, these measures ought to serve as rough upper and lower bounds for the true mean level of federal tax liability before credits respectively.<sup>19</sup> Despite the simplifications required for the imputation, the true aggregate hovers between these two measures during the treatment period, lending credence to the validity of the imputation exercise.

### 3.2 Imputing Adopting Behavior

Since 2008, the American Community Survey has collected data on whether or not a child is adopted.<sup>20</sup> This measure has two important limitations. First, while the ACS does allow for a child to be categorized as a step-child, it does not distinguish between step-parent adoptions and other adoptions. As these types of adoptions would not qualify for the credit, ideally, these types of adoptions should be excluded when estimating the credit’s impact. Second, while the age of adopted children is recorded, the date of adoption is not. Given

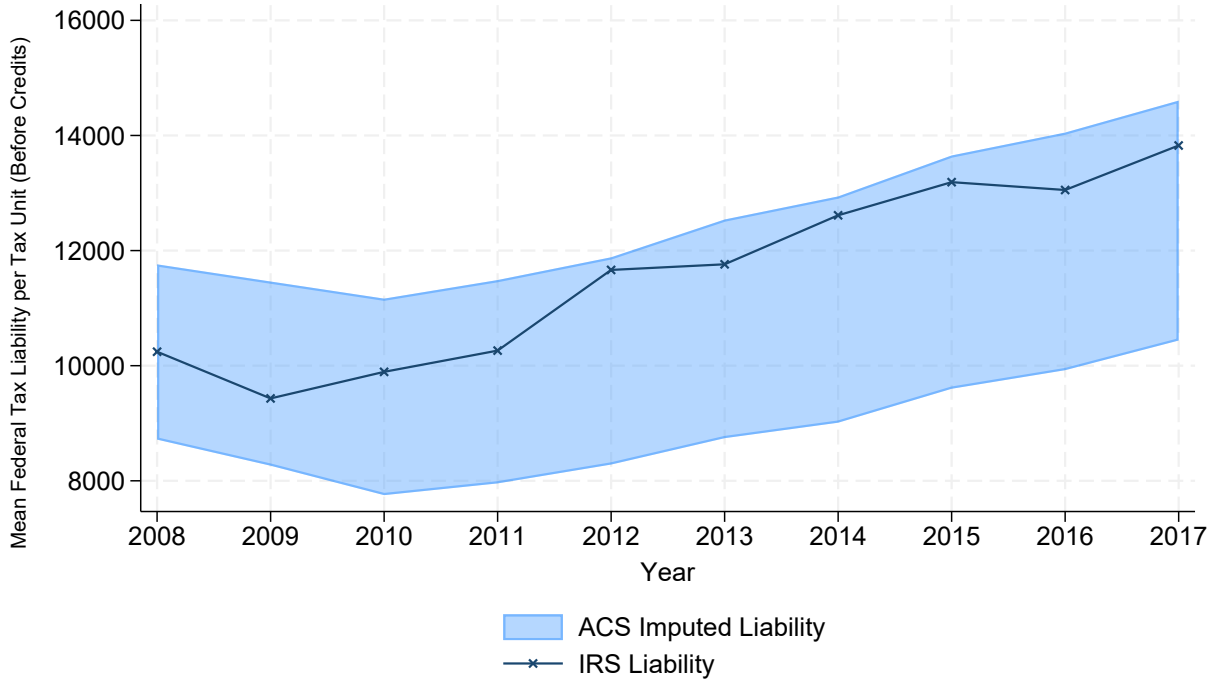
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<sup>19</sup>Technically, even individuals are imputed as possessing a very low level of tax liability, they still may not be required to file if their income is low enough. Hence, the “upper bound” is only approximate

<sup>20</sup>The Census does have information on whether or not a child was adopted in 2000, but the categorization was slightly different. For example, the step-child category was entirely omitted, meaning household heads with step-children may have called them their “adopted child”.



Figure 1: Imputed Liability Against Aggregate IRS Liability



Note: The black line plots the average level of income tax liability before credits and comes from the Internal Revenue Service's Statistics of Income, Individual Income Tax Returns Table 1 for tax years 2008-2017. The top blue line is the average level of imputed federal tax liability before credits per household among all households I impute as having positive liability. The bottom blue line is the average level of imputed federal tax liability before credits for all households. Since households with very low levels of income are not required to file, these measures ought to serve as upper and lower bounds for the true mean level of federal tax liability before credits respectively. Replacing the number of returns with the imputed number of households using the aggregate IRS data does not qualitatively change the graph.

that children can be adopted at any age, using merely the total number of adopted children only provides information on the stock rather than the flow of adoptions. It is not obvious which adoptions were performed prior to treatment by merely using the stock.

The primary goal of this imputation procedure is to connect yearly household-level adopting behavior to the household-level characteristics that determine exposure to reform. A graphical representation of the imputation procedure is given in Figure 2 and a detailed description is located in the Appendix. This procedure uses the age of the adopted child, their immigration status, and their migration status relative to other household members to ascertain whether or not they were adopted within the census year they are observed. With respect to step-parent adoptions, to the extent that step-children likely lived with their adoptive parent the year prior to adoption, these children will not be categorized as newly adopted. Nonetheless, to ensure the trends observed in the data are not being driven by trends in step-parent adoption, children are also identified using an additional measure that is more likely to exclude step parent adoptions, using the migration status of the spouse of

the household head.

A potential drawback to this procedure is its potential to misclassify adoptions as not “new” when the child lives with their adoptive family for longer than a year prior to adoption. For private adoptions and for adoptions of very young children, this is unlikely to be a problem.<sup>21</sup> However, for adoptions of certain types, such as grandparent adoptions or adoptions of older children from foster care, this procedure is more likely to misclassify these children.<sup>22</sup> Given that the majority of studies analyzing the impact of financial incentives on adoption tend to focus on adoptions from foster care, the results presented here should be seen as complementary with existing results, best able to track the adopting of younger children, but understating the adoption of older, publicly adopted children.<sup>23</sup>

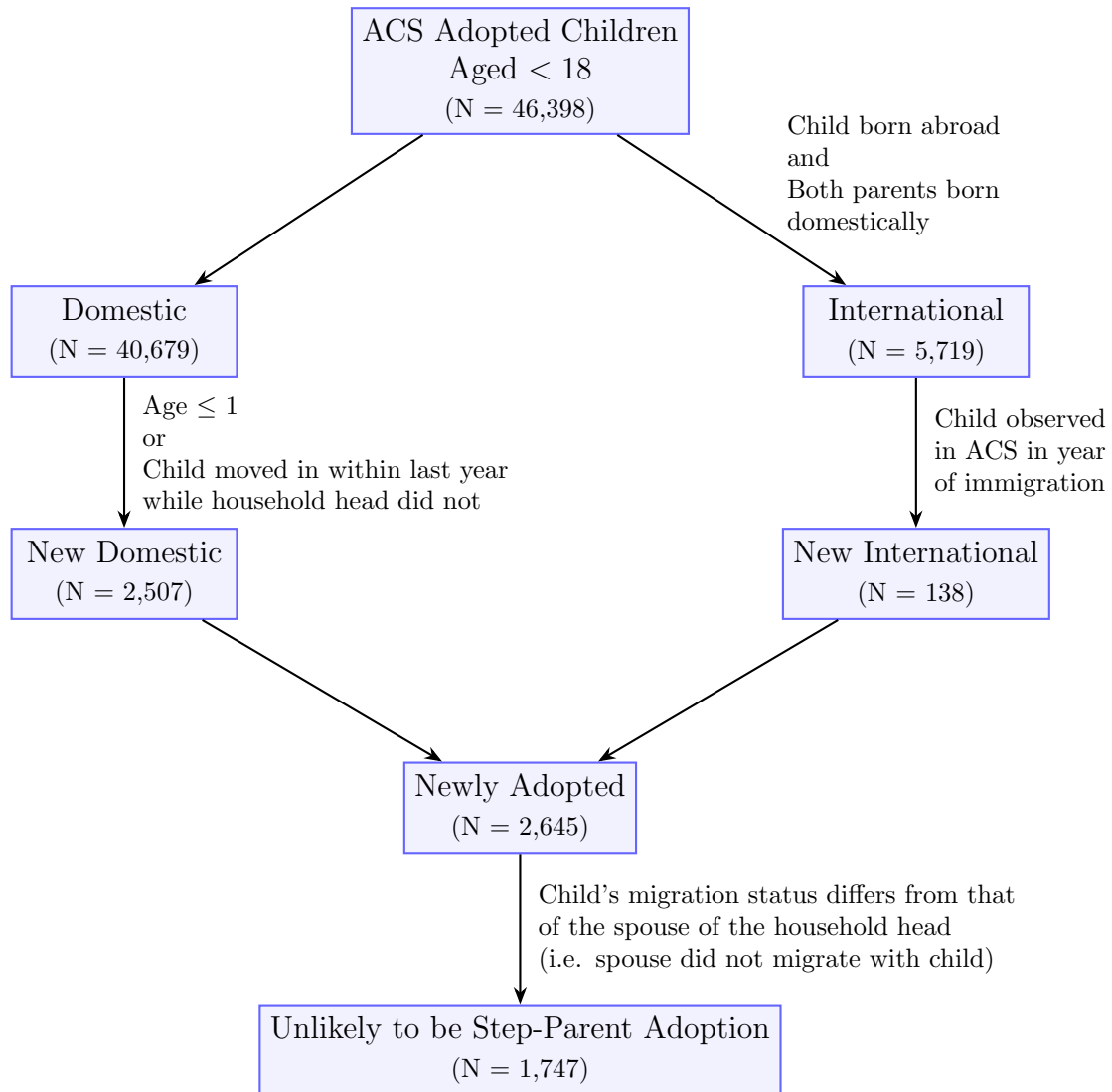
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<sup>21</sup>The National Survey of Adoptive Parents records that most private adoptions are young children and that most internationally adopted children were not living with their adoptive family the year prior.

<sup>22</sup>The average length of time to adoption for an adoptable child in foster care is highly heterogeneous. The average time is around two years (Bald et al., 2022), but notably, this is not the length of time the child is with their adoptive family before adoption. This information is not recorded in the national census of children in foster care. However, according to the National Survey of Adoptive Parents, 44 percent of publicly adopted children lived with a foster family other than their adoptive family immediately prior to their adoptive placement, and 11 percent lived in congregate care.

<sup>23</sup>This also may not be too much of a concern if policy makers are more interested in determining the credit’s impact on the capacity of adoptive children to find their permanent home. If a child were already living with what will be their permanent family, and the credit merely induced them financially to complete the formal process, this “new adoption” is clearly distinct to the case of a child with no permanent home gaining access to one.

Figure 2: Adoption Imputation Tree



Note: N denotes the unweighted number of observations from the 2009-2011 American Community Survey 1% samples.

### 3.3 Characteristics of Newly Adopted Children and What it Implies About Imputation

Table 1 show the demographic characteristics of children flagged as newly adopted across the distribution of tax liability. As anticipated, these children are more similar to the average privately adopted child than the average publicly adopted child as reported in the National Survey of Adoptive Parents (Malm and Vandivere, 2009). They are quite young, averaging less than two years of age, more likely to be classified as white, and more likely to be female (Malm and Vandivere, 2009). In addition, and consistent with aggregate data (Khun and Lahiri, 2017), the share of internationally adopted children is declining. This may be why the average age of adopted children and the fraction who are white is rising over time, as internationally adopted children tend to be younger and less white. The average age of adopted children tends to hover between 1.5 and 2.5. While it is also generally true that the age of the adoptive child is decreasing in tax liability, the group with the youngest newly adopted children is the zero liability group. Given that publicly adoptable children tend to be adopted by lower income households and that this measure may be missing some older, publicly adopted kids, this is not isn't necessarily surprising.<sup>24</sup> What this implies is that the adoption response estimated is best seen as capturing the response of the adoption of younger children, rather than older publicly adoptable children.

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<sup>24</sup>Another possibility is that lower-income households may be more likely to adopt the younger children of other family or community members. Since these households have less economic resources, the pooling together of informal resources, such as providing child care, through relationships of mutual reciprocity may be an important means through which these households weather their own idiosyncratic shocks. In other words, a household may adopt a neighbor's or family member's unexpected child to help them weather a hard time, which furnishes the expectation that when they face an idiosyncratic shock that neighbor or family member will help them. The tax credit would apply to expenses incurred in performing such adoptions and so the reform may have helped subsidize this type of informal community support.

Table 1: Characteristics of Newly Adopted Children Before and During Reform

	2008-2009	2010-2011	2012-2013	2014-2015
<b>Average Age</b>				
Total	2.199 (0.109)	2.381 (0.134)	2.365 (0.148)	2.411 (0.15)
Control	2.060 (0.165)	2.588 (0.281)	2.372 (0.292)	2.276 (0.278)
0 Liab	1.685 (0.21)	1.996 (0.251)	1.927 (0.228)	2.723 (0.36)
1-5000	2.484 (0.242)	2.818 (0.288)	2.394 (0.29)	2.333 (0.294)
5001-10,000	2.526 (0.262)	1.789 (0.213)	3.101 (0.422)	2.348 (0.315)
Positive Liab	2.518 (0.173)	2.495 (0.191)	2.575 (0.224)	2.346 (0.201)
<b>Fraction International</b>				
Total	0.113 (0.00794)	0.0721 (0.00818)	0.0575 (0.00725)	0.0441 (0.00677)
Control	0.217 (0.0183)	0.153 (0.023)	0.0964 (0.0191)	0.0897 (0.0172)
0 Liab	0.016 (0.007)	0.0113 (0.0066)	0.0103 (0.00537)	0.0126 (0.00962)
1-5000	0.0555 (0.0115)	0.0516 (0.0142)	0.0451 (0.012)	0.0311 (0.0134)
5001-10,000	0.138 (0.0209)	0.089 (0.0224)	0.0805 (0.0192)	0.0303 (0.0118)
Positive Liab	0.0928 (0.0109)	0.0673 (0.0113)	0.0625 (0.0105)	0.0339 (0.00869)
<b>Fraction White</b>				
Total	0.646 (0.0133)	0.656 (0.0153)	0.707 (0.0159)	0.705 (0.0167)
Control	0.674 (0.0216)	0.671 (0.0284)	0.73 (0.0314)	0.725 (0.0309)
0 Liab	0.541 (0.0346)	0.557 (0.0337)	0.617 (0.0354)	0.667 (0.0381)
1-5000	0.675 (0.0243)	0.669 (0.0299)	0.744 (0.0284)	0.699 (0.0337)
5001-10,000	0.663 (0.0298)	0.753 (0.0294)	0.726 (0.0359)	0.739 (0.035)
Positive Liab	0.676 (0.018)	0.697 (0.0207)	0.743 (0.0211)	0.712 (0.023)
<b>Fraction Black</b>				
Total	0.209 (0.0119)	0.224 (0.0134)	0.227 (0.0153)	0.235 (0.0157)
Control	0.138 (0.0161)	0.153 (0.0192)	0.178 (0.0285)	0.213 (0.0275)
0 Liab	0.328 (0.035)	0.32 (0.0312)	0.285 (0.034)	0.244 (0.0352)
1-5000	0.194 (0.0193)	0.251 (0.0288)	0.192 (0.0253)	0.244 (0.0328)
5001-10,000	0.207 (0.0254)	0.167 (0.0247)	0.244 (0.0354)	0.251 (0.0351)
Positive Liab	0.199 (0.0148)	0.207 (0.0186)	0.216 (0.0208)	0.244 (0.0225)
<b>Fraction Female</b>				
Total	0.537 (0.0131)	0.524 (0.0154)	0.522 (0.0162)	0.504 (0.0178)
Control	0.51 (0.0217)	0.545 (0.0284)	0.543 (0.0304)	0.496 (0.0337)
0 Liab	0.548 (0.0342)	0.527 (0.0332)	0.545 (0.0356)	0.514 (0.0385)
1-5000	0.56 (0.0233)	0.528 (0.0299)	0.502 (0.0307)	0.522 (0.0369)
5001-10,000	0.516 (0.0302)	0.478 (0.0359)	0.536 (0.035)	0.471 (0.0383)
Positive Liab	0.548 (0.0178)	0.515 (0.0217)	0.504 (0.0223)	0.504 (0.0251)

Note: Taken from the American Community Survey 1% samples. All estimates are weighted. Standard errors of the means are reported in parentheses.

### 3.4 Characteristics of New Adopters vs Non Adopters

The descriptive statistics implied by these measures tend to corroborates other descriptive findings on who adopts (Khun et al., 2020; Khun and Lahiri, 2017; Baccara et al., 2014). Table 2 reports three of the most salient—income, age, and same-sex status—across what will constitute the control and treatment groups, over time. Table A1 reports additional demographic and economic statistics in the Appendix.

The columns of Table 2 Panel A reveal new adopters tend to be higher income within tax liability groups and that this difference has persisted over time. This makes sense given the cost of adoption and that many forms of adoption require the household to complete a home-study, the passage of which is likely associated household income.

The columns of Table 2 Panel B report the mean age of household heads over time. New Adopters tend to be younger, and this difference in age is most striking among those households with zero tax liability. Zero tax liability households are more likely to be retirees relative to positive liability households, so for non-adopters, which are a good enough stand-in for the general population, this makes sense. The fact that this trend reverses among new adopters may mean that retirees, who likely already have children of their own, are simply less likely to adopt *given* they have zero liability and hence less resources.<sup>25</sup>

The columns of Table 2 Panel C report the number of same-sex couples per 1000 households.<sup>26</sup> New adopters are around 5-6 times more likely to live in same-sex households relative to non-adopters, a finding also reported in Baccara et al., 2014 when examining a private adoption agency, and this difference has only grown over time, likely due to the growing acceptance of same-sex marriage and same-sex families during this period.<sup>27</sup>

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<sup>25</sup>This fact likely also implies some “new” grandparent adoptions are not being counted. If a child were to live with a grandparent for more than 2 years before being adopted, the child would not be categorized as newly adopted. This may not be a substantial concern, if the primary adoptions policy makers care about are non-relative adoptions. However, as will be discussed below, this may mean certain types of public adopts are being miss using this measure.

<sup>26</sup>While the ACS provides its own measure for whether a couple is a same-sex couple, it only started doing so in 2013. Hence, I impute whether a couple is a same-sex couple by simply looking at whether the household head and their partner are the same sex. This is the SSC Imputed variable.

<sup>27</sup>During the reform period, it is interesting to note the share of same-sex couples who were new adopters fell for the control but tended to weakly rise among most of the treated groups. This may imply treatment tended to be especially beneficial to same-sex couples of particular socioeconomic classes, but not monotonically. Given the fall in the share of same-sex couples among the 1-5000 liability group category, this benefit was likely not monotonically increasing in income. This may be due to differences in the types of kids being adopted by same sex couples across the distribution of tax-liability.

Table 2: Main Summary Statistics

<b>PANEL A: Mean Annual Household Income (\$1,000s)</b>					
		<i>2008-2009</i>	<i>2010-2011</i>	<i>2012-2013</i>	<i>2014-2015</i>
<i>New Adopters</i>	Control	188.23 (4.73)	189.04 (5.05)	204.25 (7.40)	221.29 (8.48)
	0 Liab	24.31 (1.52)	24.21 (1.23)	23.53 (1.30)	26.24 (4.01)
	1-5000 Liab	52.51 (1.33)	54.68 (1.43)	56.15 (1.49)	54.83 (1.53)
	5001-10,000	81.55 (1.03)	87.13 (1.53)	90.42 (3.25)	86.35 (1.22)
	Positive Liab	65.60 (1.09)	71.48 (1.28)	71.62 (1.67)	74.25 (2.29)
<i>Non-Adopters</i>	Control	171.92 (0.18)	182.26 (0.19)	186.94 (0.20)	198.21 (0.21)
	0 Liab	22.36 (0.03)	23.27 (0.03)	23.89 (0.03)	25.25 (0.04)
	1-5000 Liab	44.44 (0.03)	45.58 (0.03)	46.71 (0.04)	48.60 (0.04)
	5001-10,000	75.61 (0.04)	77.57 (0.05)	79.41 (0.05)	81.84 (0.05)
	Positive Liab	56.81 (0.03)	59.79 (0.03)	60.82 (0.04)	63.89 (0.04)
<b>PANEL B: Mean Age of Household Head</b>					
		<i>2008-2009</i>	<i>2010-2011</i>	<i>2012-2013</i>	<i>2014-2015</i>
<i>New Adopters</i>	Control	40.633 (0.326)	41.215 (0.456)	40.874 (0.471)	41.114 (0.541)
	0 Liab	33.255 (0.644)	33.879 (0.609)	34.228 (0.743)	35.739 (0.922)
	1-5000 Liab	36.417 (0.424)	36.268 (0.697)	35.719 (0.51)	35.632 (0.567)
	5001-10,000	38.607 (0.468)	37.564 (0.499)	38.216 (0.559)	37.121 (0.658)
	Positive Liab	37.409 (0.305)	37.073 (0.429)	36.697 (0.37)	36.647 (0.407)
<i>Non-Adopters</i>	Control	49.400 (0.0205)	50.088 (0.0239)	50.433 (0.0233)	50.801 (0.023)
	0 Liab	56.891 (0.0322)	56.467 (0.0296)	57.059 (0.03)	57.608 (0.0301)
	1-5000 Liab	47.152 (0.0223)	47.868 (0.0232)	48.148 (0.0239)	48.454 (0.0244)
	5001-10,000	47.430 (0.026)	48.193 (0.028)	48.776 (0.0287)	49.052 (0.0289)
	Positive Liab	47.276 (0.0166)	48.008 (0.017)	48.406 (0.0176)	48.732 (0.0177)
<b>PANEL C: Number of Same-Sex Couples per 1,000 Households</b>					
		<i>2008-2009</i>	<i>2010-2011</i>	<i>2012-2013</i>	<i>2014-2015</i>
<i>New Adopters</i>	Control	36.7 (8.59)	26.3 (8.9)	82.4 (15.2)	118 (23.7)
	0 Liab	5.15 (2.98)	12.6 (4.94)	19.6 (6.61)	24.6 (12.2)
	1-5000 Liab	32.5 (8.03)	22.8 (7.01)	46 (12.3)	53.2 (16.1)
	5001-10,000	21.3 (7.48)	55.3 (16.5)	46.8 (13)	33.3 (10.9)
	Positive Liab	29.6 (6.03)	33.2 (7.05)	48.8 (9.43)	40.5 (9.38)
<i>Non-Adopters</i>	Control	6.59 (0.131)	6.57 (0.145)	8.25 (0.158)	11 (0.178)
	0 Liab	3.87 (0.0908)	3.91 (0.089)	4.23 (0.0911)	4.37 (0.0919)
	1-5000 Liab	5.09 (0.0916)	5.36 (0.098)	5.91 (0.105)	6.59 (0.113)
	5001-10,000	4.83 (0.123)	4.96 (0.128)	5.62 (0.138)	7.32 (0.159)
	Positive Liab	5.12 (0.0722)	5.38 (0.0755)	5.97 (0.0811)	6.99 (0.0879)

Note: Taken from the American Community Survey 1% samples. All estimates are weighted. Standard errors of the means are reported in parentheses.

## 4 Estimating the Adoption Response to Reform

The reform granted households access to substantially different levels of additional credit, depending on their levels of tax liability and income. The core of the design makes use of this variation by comparing the adopting behavior of those who gained access to different levels of credit. The control group includes those households who experienced no or only a small increase in credit access. This intuition is implemented concretely through the discrete-multi-valued treatment difference-in-differences specification proposed by Callaway et. al. (2024). This design enables the estimation of group-level<sup>28</sup> treatment effects, under the identifying assumption that the average change in adopting behavior of the control group is a good proxy for the average change in adopting behavior of each treated group, had the adoption tax credit reform never taken place.

More formally, following Callaway et al. (2024), let  $Y_{itg}$  be the adoption outcome of interest for a household  $i$  at time  $t$  in treatment group  $g$ . Let  $D_{it}$  be a multi-valued treatment indicator, taking on a discrete set of values that are group specific given by  $d_g$ . All estimates are relative to the control group who received only a marginal increase<sup>29</sup> in credit access, given by  $d_0$ . The main specification, given by equation (1) below, is estimated separately through OLS for each treatment year, 2010 and 2011, relative to 2009.

$$Y_{itg} = \alpha_g + \delta_t + \sum_{g=0}^G 1(D_{it} = d_g) \beta_g + \epsilon_{itg} \quad (1)$$

The parameter  $\beta_g$  is interpreted as the average treatment effect on group  $g$  among those receiving  $d_g$ , defined in the subsequent section, under the assumption that the mean change in adopting behavior of the control group between 2009 and the treatment year of interest, is equal to the mean change in adopting behavior of treated group  $g$ , had the reform never taken place.<sup>30</sup>

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<sup>28</sup>The groups will be defined by their level of federal tax liability before credits. For example, one set of groups are those with 0 tax liability and those with positive liability who are not in the control.

<sup>29</sup>These include households who benefited from the 1000 dollars increase, but did not benefit from the credit being made refundable due to their high level of tax liability, as well as households with high enough incomes to not qualify for any of the credit.

<sup>30</sup>In addition, given the data is cross sectional, it must be assumed that changes that the unobservable and observable characteristics that are relevant to adopting behavior would have been stationary in the absence of treatment. If, for example, the group of households with zero tax liability was markedly different between 2009 and 2011 due to some factor unrelated to treatment, changes in outcomes due to this compositional change would influence the estimated treatment effects.



## 4.1 Calculating the Average Relative Level of Treatment Received by Groups

The estimates of  $\beta_g$  capture a non-parametric average adoption response of each group due to reform. To measure the average level of treatment actually received by each group,  $d_g$ , the relative level of additional credit households gained access to must be estimated. To do this, the level of credit each household would qualify for under the new regime in a given treatment year is calculated, based on their imputed federal tax liability before credits and their imputed MAGI. Then, the level of credit the household would qualify for under the counterfactual regime where the credit is non-refundable is calculated. The difference between these two quantities is the imputed household-level increase in available credit.<sup>31</sup> Within each group, this quantity is averaged over households. To make the average group-level treatment relative to the control group, the control group's average increase in available credit is subtracted.<sup>32</sup>

This average relative increase in credit access can be combined with the estimated average treatment effects to estimate average treatment effects per dollar of increased available credit. It is important to note that this is distinct from estimating the marginal treatment effect of increasing credit access by an additional dollar. In order to estimate this quantity, a much stronger version of parallel trends would have to be imposed (Callaway et al., 2024). Namely, that the change in adopting behavior of the group receiving  $d_g$  is a valid proxy for all other groups,  $g'$ , had they received  $d_g$  instead of  $d_{g'}$ .

## 4.2 Choosing Treatment and Control Groups

Treatment groups are defined by their level of federal tax liability before credits assuming the standard deduction, as this tracks the level of maximum credit they could qualify for within a given year and allows for a clean identification of household groups over time. The control group is defined as those households with tax liability above \$13,360, the maximum credit available in 2011. These households experienced a modest increase in credit access of \$700 on average, owing to the \$1000 increase in the credit due to reform.<sup>33</sup>

Figure 3 shows the average imputed increase in credit across \$1000 bins. The reason

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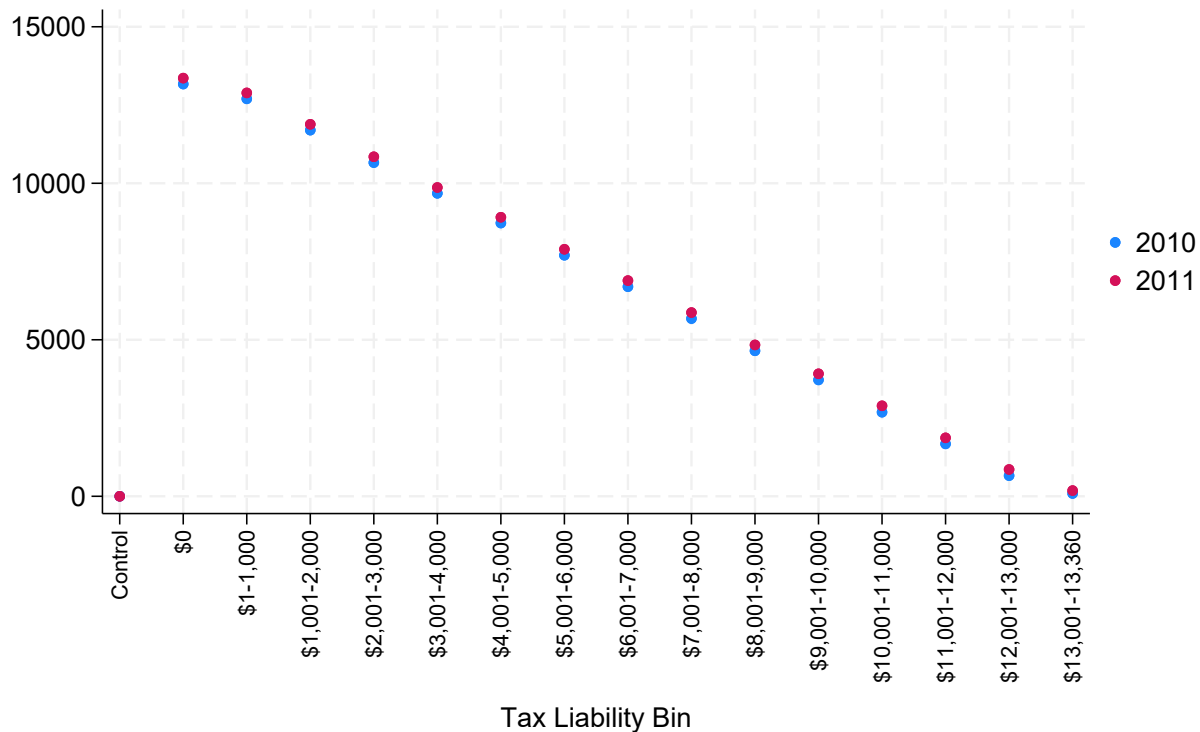
<sup>31</sup>Again, with the caveat that their liability is computed using income variable reported for the prior year

<sup>32</sup>As an example, consider a group of households with \$5000 of tax liability with sufficiently low MAGI to not enter the phaseout range. In 2009, these households that year would have access to \$5000 of credit. However, after the reform, households in 2011 with \$5000 of tax liability had access to \$13,360 of credit. This constitutes an increase of \$7360 dollars of additional credit after reform. Since the control group gained, on average, roughly \$700 dollars of credit, this translates into a group-level average relative credit increase of \$6660.

<sup>33</sup>A natural alternative to this group would be the set of households with MAGI above the maximum threshold, as these households never qualified for the credit. However, this alternative group exhibited trends that were non-parallel after repeal, implying the parallel trends assumption necessary to use them would not be credible.

why tax liability does not perfectly track exposure to additional credit is because the MAGI income limitations reduce the maximum credit available to households.

Figure 3: Imputed Mean Increase in Credit Access by Liability Bin Due to the Credit Becoming Refundable



Grouping taxpayers by their level of tax liability requires balancing several tradeoffs. By choosing finer bins, households are more similar within a group and receive a more similar level of treatment, allowing for a better characterization of the treatment effect heterogeneity induced by the policy. However, since finer bins also contain fewer households, their mean adopting behavior over-time is less stable, making the identifying assumption of parallel trends less credible. Finer bins may also make estimates more susceptible to bias induced by compositional change unrelated to treatment over time.<sup>34</sup>

Given these considerations, two sets of treatment groups are defined. The first is composed of the group of households with zero liability and the group of households with positive

<sup>34</sup>Another alternative to using bins would be to attempt to estimate treatment effects using a flexible function of the imputed increased level of credit access, as proposed by Callaway et al. (2024). This approach was not taken for two reasons. First, it requires strong parallel trends, which is unlikely to be valid in this setting. Second, evaluating the plausibility of parallel trends as well as whether or not enough statistical power is available to form valid counterfactuals is simply more transparent using the binned approach.

liability who are not in the control. These are termed the coarse liability bins. The second set is composed of four treatment groups: those with zero liability, those with \$1-\$5000 of liability, those with \$5001-\$10000 of liability, and those with \$10001-\$13360 of liability. These are termed the \$5000 liability bins.

### 4.3 Visually Inspecting Trends

Figure 4 shows the mean evolution of adopting behavior for two treated groups as well as the control. These treated groups are those with 0 tax liability, and those with positive tax liability below the maximum credit amount in 2011. The dashed lines represent when the treatment occurred and when the credit reverted back to its previous nonrefundable state.<sup>35</sup> While the most common practice to assess the plausibility of parallel trends is the examination of pre-trends, as mentioned previously, the ACS does not record whether or not a child is adopted before 2008. Nonetheless, it is apparent that between 2008 and 2009 both the treated and control groups experienced parallel sharp declines in adopting behavior. Given there are not additional pre-treatment periods over which to observe pre-treatment trends, to garner further support that the treated and control groups would have evolved in parallel in the absence of treatment, their behavior after the reform expired is analyzed.

The danger in using this evidence, relative to using pre-trends, is that treatment may have induced the treated group to evolve in parallel to the control group when, in the absence of treatment, the treated groups would have evolved in a non-parallel fashion. In this sense, pre-trends provide less “contaminated” evidence that the control and treated group would have actually evolved in parallel in the absence of treatment, compared to using post-trends after a repeal. Nonetheless, examining the trend in behavior after repeal still lends credibility to the parallel trends assumption for two reasons. First, as can be seen in Figure 4, the treated and control group exhibit qualitatively quite similar behavior many years after treatment. None of the deviations in trend year-over-year, which will be measured in a variety of different ways, are statistically significant. Second, while treatment may have impacted long-run adopting behavior both directly through its 5 year carry forward provision<sup>36</sup> or possibly indirectly, through other mechanisms (i.e. long run-network effects<sup>37</sup>, a now larger adoption industry due to the subsidy, etc.), this does not imply that it induced what would have been a non-parallel relationship to be parallel. For example, if the credit

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<sup>35</sup>The expiry of the refundability regime occurred in January 2012, when the credit again became non-refundable and the \$1000 increase was reversed.

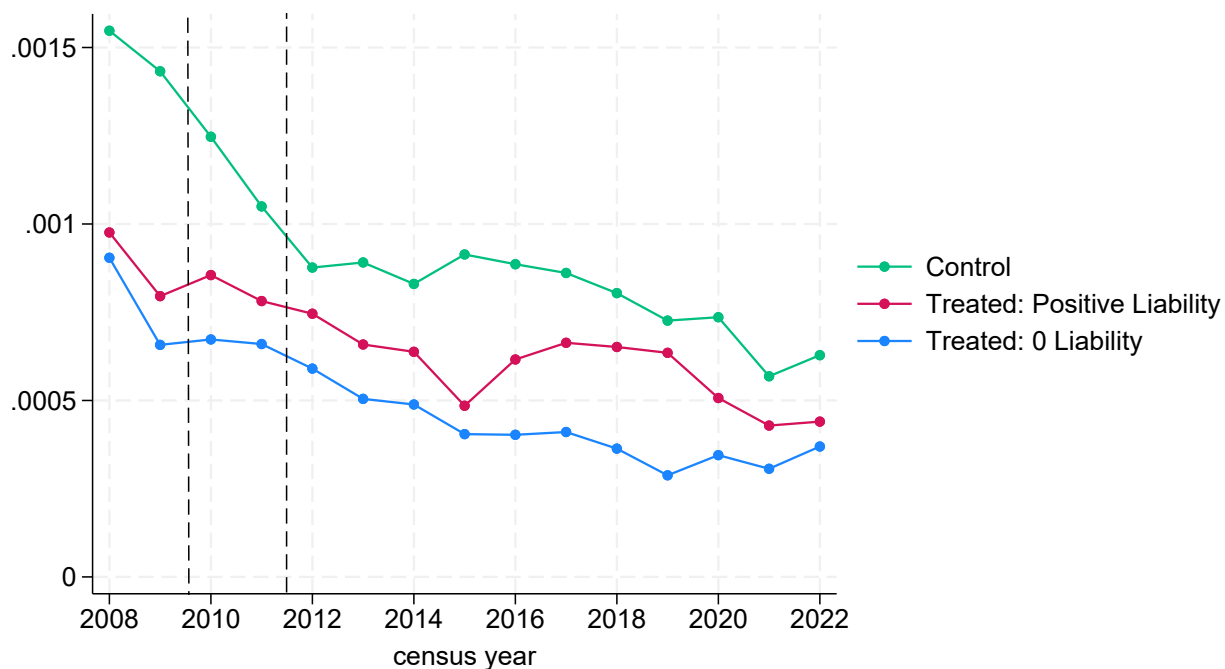
<sup>36</sup>Expenditures on the credit declined to their pre-repeal state, suggesting payments of carried forward credit were not substantial during this time (Brehm, 2021)

<sup>37</sup>For example, if the credit encouraged people who never would have adopted to adopt, they could influence friends and family members to adopt through networking. The National Survey of Adoptive Parents reveals that many who choose to adopt previously knew and were influenced by other adopters

induced a level increase in adopting behavior among the treated groups, post-repeal trends would still provide a similar type of evidence as pre-trends.

All this is to say, it is impossible to evaluate directly parallel trends. Using pre-trends provides intuitive support because, the argument goes, why would, in the absence of treatment, the treated and control have evolved in a non-parallel fashion when they were evolving in parallel before?<sup>38</sup> To the extent that this argument is convincing, a parallel argument exists in this context. Why, after the major provisions of the reform are repealed, would the treated and control group exhibit statistically similar trends, but not have done so in the past in the absence of those major provisions?

Figure 4: Mean Number of New Adoptions per Household



#### 4.4 Estimating Year-Over-Year Placebo ATTs

In order to quantify whether the treated and control groups evolve in parallel both in the pre-period and during the post-repeal periods, a variety of placebo tests are performed to verify whether or not these trends are statistically distinguishable.

The first set of tests involve estimating equation (1), but acting as if treatment occurred in different years. During the pre-period, a pseudo-ATT is estimated by assuming treatment occurred in 2009, relative to 2008. During the post-repeal period, the same is done for each

<sup>38</sup>Post hoc ergo propter hoc?

year, estimating a treatment effect for 2012 relative to 2011, a treatment effect in 2013 relative to 2012, and so on up to 2022. A statistically significant pseudo-ATT would be equivalent to rejecting the hypothesis that the trend in the outcome between the treated groups and the control group are the same (Callaway and Sant’Anna, 2021).

The next set of tests involves estimating pseudo-ATTs over 2-year time horizons. In this case, equation (1) is estimated to create pseudo-ATTs but is estimated over two year intervals. Again, a statistically significant treatment effect in these contexts is equivalent to rejecting the hypothesis that the treated and control group are evolving in parallel over that two year time horizon.

Finally, to obtain more conservative estimated placebo-ATTs using larger samples, as well as to better evaluate the parallel trends assumption underlining the treatment effect estimate that pools 2010 and 2011 together, equation (1) is estimated but with observations over two year sets being pooled together. Given the larger number of observations, these bi-annual pseudo-ATTs ought to have more statistical power, and so are more able to reject the hypothesis that the treated and control groups are exhibiting non-parallel trends in the post-reform period.

## 5 Results

### 5.1 Impact on the Mean Number of Adoptions Across Coarse Liability Bins

Table 3 column 1 shows the point estimates from specification (1) for  $\beta_g$  for years 2010 and 2011. All regressions utilize the ACS’s survey weights, making the sample and treatment effects nationally representative.<sup>39</sup>

The other columns of Table 3 provide regression coefficients estimated including state fixed effects, state by year fixed effects, as well as a set of carefully chosen demographic characteristics that would not be impacted by treatment.<sup>40</sup> Estimates without controls are typically preferable, as including these would require both conditional parallel trends, as well as an implicit homogeneity and linearity assumption in how the covariates impact adopting behavior (Karim and Webb, 2024; Deb et al., 2024).

The point estimates in Table 3 may be interpreted as the average treatment effect of  $d_g$  among those in group  $g$ , under parallel trends. Figure A5 plots these ATTs, along with the

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<sup>39</sup>Point estimates are of comparable size when making no use of the survey weights. Survey weights allow for the proper correction of heteroskedasticity in standard error calculations. See Solon et. al. (2013) for more details and “IPUMS USA”, n.d. for how the weights are constructed to incorporate sampling methodology.

<sup>40</sup>These include whether or not the household head has been married for 2 years, whether or not there are any non-adopted children 2 years old or greater in the household, and race

year over year pseudo-ATTs. Consistent with the findings in Brehm (2021), only significant effects in 2011 are detected. This is consistent with the view that it may have taken time for individuals to respond to the increased incentive to adopt. Table A3 presents specification (1) when pooling 2010 and 2011 together. These results partially validate this explanation, as point estimates for each liability group are reduced, but remain significant and are of the same order of magnitude. The treatment effects do not change much with the iterative addition of various controls, which speaks to their robustness under alternative conditional parallel trends assumptions. Neither the pre-period nor any of the post repeal pseudo ATTs are statistically significant, indicating that during the period when the reform expired, the treated and control groups do not exhibit statistically different year-over-year trends.

The point estimate of 3.854 implies an average treatment effect of 3.854 additional adoptions per 10,000 households in the 0 liability group in 2011. Similarly, the point estimate of 3.69 in the positive liability group implies a treatment effect of 3.69 additional adoptions per 10,000 households in 2011. These increases constitute a 58.6 percent increase and a 46.3 percent increase in the average number of new adoptions per household, relative to their pre-treatment 2009 average, respectively.

The approximate relative increase in credit access, relative to the control, among these two groups was 12660 dollars and 8164 dollars in 2011, respectively. This implies a partial elasticity of 4 percent per \$1000 of additional credit for the 0 liability group and 5.6 percent per \$1000 of additional credit for the positive liability group. Scaling these estimates to their relative population sizes implies treatment increased the total number of adoptions in 2011 by 13,643 for the 0 liability group and 22,545 for the positive liability group. These effect sizes are large<sup>41</sup> and heterogenous. The lower per dollar responsiveness of the 0 liability group relative to the positive liability group makes sense in light of the often high upfront costs associated with adoption, above and beyond the credit amount. For example, if it costs, say \$20,000 to complete a particular adoption, additional resources beyond the credit would be needed to carry it out. Given their level of resources this adoption would likely not be feasible among those in the 0 liability group but would likely be possible for those in the positive liability group.

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<sup>41</sup>The measure for new adoptions implies there were 100,000 total adoption in 2009. As mentioned previously, this is likely an under count due to the nature of public adoptions

Table 3: Treatment Effects for Coarse Liability Bins: Number of Newly Adopted Children

	(1)	(2)	(3)	(4)	(5)
ATT 0 Liab 2010	2.009 (1.731)	1.992 (1.732)	1.644 (1.730)	1.631 (1.730)	1.444 (1.746)
ATT Pos Liab 2010	2.451 (1.588)	2.450 (1.588)	2.401 (1.587)	2.403 (1.587)	2.205 (1.600)
ATT 0 Liab 2011	3.854** (1.640)	3.846** (1.640)	3.540** (1.639)	3.530** (1.639)	3.590** (1.662)
ATT Pos Liab 2011	3.690** (1.527)	3.674** (1.527)	3.687** (1.527)	3.671** (1.527)	3.723** (1.545)
Demographic Controls	No	No	Yes	Yes	Yes
State FE	No	Yes	No	Yes	Yes
State x Year FE	No	No	No	No	Yes
Observations	2,395,597	2,396,627	2,395,597	2,396,627	2,396,627

Note: The table displays the dynamic results for the regression design with coarse bins, where the outcome of interest is the reform's impact on the mean number of newly adopted children per household. The coefficients are transformed to represent the treatment effect per 10,000 households. The demographic controls include the household head's age and age squared, race, whether they were married for 2 or more years, and whether or not the household head has any non-adopted children over 2. Under unconditional parallel trends, the coefficients in model (1) may be interpreted as the average treatment effect on the treated, where treatment is conceptualized as the quantity of additional credit the group gained access to, though not necessarily actually used. Models (2)-(5) may be interpreted similarly, though under separate conditional parallel trends assumptions.

## 5.2 Impact on the Mean Number of Adoptions Across \$5000 Liability Bins

In order to better parse the potentially heterogeneous response to treatment across tax liability, the treatment effects across \$5000 liability bins are estimated. These results, along with the relevant year over year placebo ATTs are displayed in Table A4. The treatment effect in 2010 continues to remain insignificant across liability bins. However, the treatment effect for 2011 is significant for those with \$1-\$5000 of tax liability and those with \$10,001-\$13,360 of tax liability, although the latter group fails its pre-period placebo test. Interestingly, the 5,001-10,000 is insignificant in 2011. While this may simply be driven by lower power, this may be indicative of a smaller adoption response given the lower level of treatment.

A treatment effect of 3.67 for the \$1-\$5000 liability group constitutes a 51 percent increase relative to their pre-treatment average. Since this group experiences, on average, \$10567 additional credit relative to the control, this suggests an average responsiveness of 4.8 percent per \$1000 of additional credit. A lower average per dollar responsiveness of

the 1-5000 liability group relative to the positive liability group, which had an average per \$1000 responsiveness of 5.6 percent, is consistent with the hypothesis that those with more means may have a greater adoption response due to the high fixed costs associated with adoption.<sup>42</sup> Scaling the \$1-\$5000 liability treatment effects by their population suggests this group adopted 13,909 additional children than they would have, in the absence of treatment.

### 5.3 Estimating Average Expenditure Increase to Encourage an Additional Adoption

When the government modified the credit, additional tax expenditures flowed both to individuals who would have claimed the credit in the absence of reform and to individuals induced to perform a new adoption by the reform.<sup>43</sup> Of particular interest to policymakers is the average fiscal cost of encouraging an additional adoption, given they are not able to distinguish between these two groups. Aggregate tax expenditures on the adoption tax credit across income groups are reported by the IRS yearly. However, these are not reported across tax liability. In order to translate this expenditure data into this context, the expenditures on households across liability bins needs to be imputed. To do so, first, consider all households that performed a new adoption within a given year. Each household is then assigned to an income group, the groups being the level at which the aggregate expenditures are reported. Naturally, within an income group, there are households that belong to different tax liability bins. The fractions of households that are within each liability bin are computed for each income group. That fraction of total expenditure is then assigned to each liability group.<sup>44</sup> This gives an imputed measure for the level of tax expenditures flowing to each liability bin. These results are presented in Table 4.

Only the treatment effect in 2011 is incorporated. However, the average change in expenditures, both in 2010 and 2011 are used, as the change in expenditures in 2010 may have been expenses for an adoption that was finalized in 2011. Assuming the total change in expenditure on each liability group are attributable to reform, which seems to be a valid assumption given expenditures on the credit were relatively level before reform (Brehm, 2021), the average fiscal cost per new adoption for the zero liability group was approximately

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<sup>42</sup>While I am reticent to interpret the 10001-13360 group, due to it failing the pre-period placebo test, its somewhat unstable mean, and high treatment effect, the large response I observe is also consistent with this hypothesis.

<sup>43</sup>To use the standard terminology, these are the always takers and the compliers (Imbens and Angrist, 1994).

<sup>44</sup>As an example, suppose the IRS reports \$10,000 of tax expenditure was spent on those households making \$20,000-30,000. If 50 percent of these households have 0 liability and the rest have positive liability, 50 percent of the expenditures is allocated to the 0 liability group and 50 percent on the positive liability group. It's important to remember this is among newly adoptive households, not all households.



$\$215,581,000/13,643 = \$15,801$ . The average fiscal cost per new adoption for the positive liability group was approximately  $\$350,202,000/22,545 = \$15,533$ . The average fiscal cost per new adoption for the \$1-\$5000 liability group was  $242377000/13,909 = \$17,425$  implying it took more funds flowing to this group to induce more adoptions, relative to the zero liability group and the higher liability groups. This suggests a non-monotonic per dollar cost per adoption across the distribution of tax liability.

Table 4: Imputed Total Expenditures Flowing to Liability Groups Over Time

	2008-2009	2010-2011	Difference
<b>Actual Yearly Average Total Tax Expenditure (\$1000s)</b>	315972	908905	592933
<b>Imputed Yearly Average Tax Expenditure (\$1000s) on:</b>			
Control Group	134565	162184	27619
0 Liability Group	1308	216889	215581
1-5000 Liability Group	69596	311973	242377
5001-10,000 Liability Group	88764	164311	75547
10,001-13,360 Liability Group	21194	53472	32279
Treated Positive Liability Group	179554	529756	350202

Note: For the actual totals I use the Internal Revenue Service, Statistics of Income, Individual Income Tax Returns, Publication 1304, Table 3.3, Tax Years 2008-2011. To impute the share of tax expenditures flowing to each liability group, I use the disaggregated expenditures that are broken down by AGI bins provided by the above publication. I then look at the fraction of new adopters within each liability bin that compose each AGI bin. I then assign each liability bin that share of the aggregate AGI level expenditure.

## 5.4 Impacts on Other Outcomes

In addition to investigating the reform's impact on the number of new adoptions, other outcomes can also be investigated.

The first is treatment's impact on the probability of conducting a new adoption. Figure A1 plots this average for the control group, the zero liability group, and the positive liability group. Qualitatively, the trends appear roughly parallel in the pre-treatment period and post-repeal periods, with the possible exception of the period between 2011-2012. Table A6 plots both the estimated treatment effects and the various pseudo-ATTs, all of the latter not being statistically significant. The first model of Table A2 reports the treatment effect point estimates for the coarse liability bins. The treatment effect of .000297 for the zero liability group is roughly 50 percent of their 2009 pre treatment probability of adoption, implying the reform had a large extensive margin effects.<sup>45</sup> The point estimate

<sup>45</sup>These incredibly small numbers are one of the reasons why large, nationally representative data sets are so important when investigating aggregate policy geared to influencing child adoption.

for the positive liability group of .000248, implies a 33 percent increase relative to their pre-treatment average probability of adoption. Table A5 further decomposes these treatment effects by 5000 liability bins. None of the positive liability bins remain significant except for the \$10000-\$13,360 liability group, but this fails the pre-period placebo test.

Given the average age of newly adopted children using this measure is quite young<sup>46</sup>, the impact of treatment on the number of adopted children aged less than 2 is calculated. Figure A2 plots this outcome over time for the 0 liability and positive liability treated groups as well as the control group. Here, the trends between 2013 and 2015 appear somewhat divergent. Nonetheless, the year over year trends are not statistically distinguishable, and the break between treated and control groups is quite pronounced during the treatment period. The treatment effects and pseudo ATTs are plotted in Figure A2, while the second model of Table A2 displays the point estimates. Again, the treatment effects are positive and significant for both liability groups in 2011. The point estimate of 3.40 for the 0 liability group constitutes a 59 percent increase relative to their 2009 level. For the positive liability group, a point estimate of 3.36 constitutes this increase a 53 percent increase relative to their 2009 level. Scaling both of these estimates suggests the zero liability group adopted 12036 more babies and the positive liability group adopted 20496 more babies than they would have under the zero treatment counterfactual. This suggests the majority of the measured treatment effect, about 90 percent, is being driven by the adoption of children less than 2. This treatment effect, again, may be indicative of my measure of new adoptions primarily picking up non-foster care adoptions. As mentioned previously, 2/3 of private adoptions are of children less than 1 year old (Malm and Vandivere, 2009). Table A6 reports treatment effect estimates for the \$5000 bin estimates as well as their placebos. Here, effects remain significant for all bins in 2011. Absolute treatment effects are declining, but the average responsiveness per dollar of credit exposure rises in tax liability.

## 6 Robustness Checks

### 6.1 Trends in International Adoptions

International adoptions have been declining precipitously since 2008 (Khun and Lahiri, 2017)<sup>47</sup>, and given that richer households are more likely to perform an international adop-

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<sup>46</sup>As noted previously, this may be indicative of my measure missing some types of public adoptions.

<sup>47</sup>This decline is due to a variety of reasons, including greater restrictions placed by two of the most common countries parents adopted from at the time in 2007. These included China, who put in place additional regulations on who may adopt from them internationally and Guatemala, the adoptions from which peaked in 2007 and then declined after an anti-corruption campaign to reduce the potential for human trafficking.

tion, this presents the risk of my control group being differentially exposed to an adoption shock during my treatment period. To account for this, treatment effects on the number of new adoptions flagged as domestic are calculated. Figure A3 shows this plot for the coarser liability bins, as well as the control group. The decline in the treated and control groups pre-treatment is still readily apparent, and each of the groups still qualitatively evolved in parallel after repeal, perhaps even a bit more so. The 3rd model in Table A2 shows the point estimates for these treatment effects, while Figure A8 plots these treatment effects and the relevant placebo tests. The estimated treatment effects remain positive and statistically significant for 2011, and are of similar magnitude to the primary treatment effect estimates, indicating that the results are being driven primarily by domestic adoptions, and not the decline in international adoptions. Table A7 reports these estimates across the \$5000 bins.

## 6.2 A Measure Even More Robust to Step-Parent Adoptions

Finally, to ensure that the trends in the data are being driven by non-step parent adoptions, estimates are calculated again using only new adoptions where the partner of the household head is also not living in another household 1 year prior. This definition is used because when undertaking a step parent adoption, if the child moved in within the last year, it is likely that their biological parent did as well. The ACS records the adoption status of children relative to the household head, and not to both members of the household. If the partner of the household head and the adopted child are both living outside the household 1 year prior, this may suggest the partner is the biological parent of the child and the household head performed a step-parent adoption after their partner and the child moved in. Figure A4 plots the mean evolution of outcomes, no longer counting these children as newly adopted. Figure A9 plot the treatment effects excluding these children as well. While some power is lost by dropping these observations, making the estimates less precise, it is clear the trends that compose the identifying assumption are not being driven by these observations and hence, it is unlikely trends in step-parent adoptions are driving the results.

## 6.3 A Discussion of Substitution Effects

One concern is that some of the treatment effects detected are partially the product of those in the control group losing access to children they themselves would have adopted anyway. Indeed, these substitution patterns could occur between the control group and treated groups and between the treated groups themselves. If the entire decline in the control group were being driven by this substitution, then the control group would not be a valid counterfactual for how the treated groups would have evolved in the absence of treatment. While this possibility can't be ruled out, as it would require knowledge of the

counterfactual, there is suggestive evidence that the decline is not being primarily driven by substitution. First, adoptions are declining in the control group at a similar rate before treatment is enacted, and this is not explained by the aggregate decline in international adoptions, as shown in Figure A3. Second, as shown in Table 1, since each of these liability groups adopt different kinds of children, if the trend in the control group were being driven entirely by substitution, this ought to show up in the characteristics of newly adopted children. However, the characteristics of adopted children in the control do not change in a statistically significant way, save for the secular decline in international adoptions.

Other forms of substitution are also possible, but their presence is less of a concern. For example, substitution effects occurring within a treated group will be correctly incorporated into the treatment effects estimated. Substitution effects between liability bins could lead to over or under estimation of their group specific treatment effects, depending on the substitution patterns involved, but these effects would tend to counteract one another.<sup>48</sup> Of course, when evaluating the welfare and efficiency consequences of this policy, determining whether there is shifting going on is quite important.

## 7 Conclusion

This paper has shed light on the behavioral response of child adoption to financial incentives across the distribution of tax liability, using a reform that took place in 2010-2011 that differentially impacted households across the tax liability distribution. These estimates are particularly relevant to studying the adoption of younger children from private settings. The credit was recently reformed to allow the first \$5000 to be fully refundable. These estimates imply the reform will result in roughly 4,112 additional adoptions among those with 0 tax liability.<sup>49</sup> The magnitude of these estimates suggests tax incentives matter for determining household adoption behavior and that households across the distribution of tax liability respond in an empirically significant way.

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<sup>48</sup>This provides another justification for adopting a coarser bin strategy

<sup>49</sup>My ATT estimate for these tax payers in response to \$12660 of relative credit exposure is .00038. \$12660 in 2011 is worth roughly \$18,480.12 today. Scaling the real value of treatment exposure so that it matches the value of the current reform and scaling the treatment effect downward by the same quantity implies an average treatment effect of .0001028. Multiplying this quantity across the size of the 0 liability population gives 4112.

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## A Adoption Imputation Procedure

Consider all children flagged as adopted by the ACS who are less than age 18. A child is flagged as an international adoption if both parents are born domestically, and their child is born internationally. The rest are flagged as domestic adoptions.

Next, to classify whether an adoption is “new”, two procedures are used depending on if the child is flagged as international or domestic. If the child is international, a child is flagged as newly adopted if their year of immigration is the same as the year they are observed.<sup>50</sup> If the child is domestic, a combination of their age and migration status relative to the head of household is used. If the child is young enough less than or equal to age 1 in the ACS, they are flagged as newly adopted.<sup>51</sup> For all other children less than this age, the MIGRATE variable is used. If an adopted child’s MIGRATE variable indicates that they lived outside the household the year prior and the household head did not move, they are flagged as newly adopted.

Finally, in order to create a measure more robust to the inclusion of step-parent adoptions, children are grouped into the category “unlikely to be a step-parent adoption”. If both the child and the household head’s spouse lived in a different household the year prior, they are classified as a possible step-parent adoption and not included in this category. Although many step-parent adoptions probably occur while the child is living with the household head for a longer period of time, this provides a means of assessing whether changes in step-parent adoptions are driving the adoption trends in the data.

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<sup>50</sup>An alternative measure using the migration status and relative immigration status of the child alone produces similar results.

<sup>51</sup>As explained in detail by Baccara et. al. (2014), it is possible for a parent to be the expected adopted parent before the child is born. Of course, the adoption itself can only be finalized after birth, and depending on the state, the birth mother may be able to change her mind.

## B Descriptive Tables

Table A1: Summary Statistics

<b>PANEL A: Share of Household Heads who are White</b>					
		<i>2008-2009</i>	<i>2010-2011</i>	<i>2012-2013</i>	<i>2014-2015</i>
<i>New Adopters</i>	Control	0.873 (0.0171)	0.87 (0.0183)	0.859 (0.0262)	0.876 (0.0243)
	0 Liab	0.562 (0.0328)	0.596 (0.0344)	0.626 (0.0346)	0.691 (0.0369)
	1-5000 Liab	0.772 (0.0224)	0.743 (0.0306)	0.814 (0.0244)	0.756 (0.0332)
	5001-10,000	0.827 (0.0255)	0.854 (0.026)	0.871 (0.0273)	0.911 (0.0236)
	Positive Liab	0.802 (0.0161)	0.796 (0.0199)	0.839 (0.0171)	0.823 (0.0204)
<i>Non-Adopters</i>	Control	0.871 (0.000596)	0.865 (0.000715)	0.86 (0.00069)	0.855 (0.00068)
	0 Liab	0.745 (0.000692)	0.744 (0.000653)	0.74 (0.000668)	0.735 (0.000679)
	1-5000 Liab	0.776 (0.000605)	0.778 (0.00063)	0.774 (0.000647)	0.765 (0.000668)
	5001-10,000	0.84 (0.000745)	0.84 (0.0008)	0.836 (0.000812)	0.83 (0.000815)
	Positive Liab	0.8 (0.000456)	0.803 (0.000469)	0.798 (0.000484)	0.792 (0.00049)
<b>PANEL B: Share of Household Heads who are Married</b>					
		<i>2008-2009</i>	<i>2010-2011</i>	<i>2012-2013</i>	<i>2014-2015</i>
<i>New Adopters</i>	Control	0.901 (0.0134)	0.925 (0.0141)	0.901 (0.0239)	0.914 (0.0247)
	0 Liab	0.482 (0.0339)	0.462 (0.0343)	0.429 (0.0345)	0.438 (0.0401)
	1-5000 Liab	0.696 (0.0244)	0.662 (0.0315)	0.734 (0.0308)	0.684 (0.0341)
	5001-10,000	0.893 (0.0215)	0.879 (0.0234)	0.871 (0.0242)	0.914 (0.0214)
	Positive Liab	0.774 (0.0171)	0.757 (0.0209)	0.794 (0.0209)	0.787 (0.0208)
<i>Non-Adopters</i>	Control	0.76 (0.000748)	0.769 (0.000853)	0.758 (0.000831)	0.76 (0.000803)
	0 Liab	0.338 (0.000708)	0.345 (0.000682)	0.34 (0.000688)	0.336 (0.000688)
	1-5000 Liab	0.43 (0.000679)	0.433 (0.000716)	0.421 (0.000724)	0.411 (0.00073)
	5001-10,000	0.658 (0.000927)	0.661 (0.00099)	0.655 (0.000999)	0.654 (0.000984)
	Positive Liab	0.511 (0.000543)	0.518 (0.000565)	0.509 (0.000575)	0.505 (0.000571)
<b>PANEL C: Mean Number of Children per Household</b>					
		<i>2008-2009</i>	<i>2010-2011</i>	<i>2012-2013</i>	<i>2014-2015</i>
<i>New Adopters</i>	Control	1.970 (0.0601)	2.034 (0.0753)	1.939 (0.0777)	1.925 (0.069)
	0 Liab	2.760 (0.09)	2.819 (0.105)	2.731 (0.122)	2.783 (0.121)
	1-5000 Liab	2.308 (0.0674)	2.287 (0.0758)	2.197 (0.0808)	2.361 (0.0974)
	5001-10,000	1.987 (0.0711)	2.224 (0.101)	2.034 (0.0926)	1.970 (0.099)
	Positive Liab	2.185 (0.0489)	2.233 (0.0568)	2.156 (0.0591)	2.175 (0.065)
<i>Non-Adopters</i>	Control	0.816 (0.00179)	0.825 (0.00212)	0.812 (0.00201)	0.806 (0.00195)
	0 Liab	0.725 (0.00193)	0.738 (0.00183)	0.726 (0.00184)	0.704 (0.00184)
	1-5000 Liab	0.724 (0.00152)	0.714 (0.00159)	0.701 (0.00162)	0.695 (0.00163)
	5001-10,000	0.771 (0.00206)	0.766 (0.00223)	0.751 (0.00223)	0.748 (0.0022)
	Positive Liab	0.74 (0.00118)	0.728 (0.00122)	0.718 (0.00124)	0.709 (0.00123)

Note: Taken from the American Community Survey 1% samples. All estimates are weighted. Standard errors of the means are reported in parentheses.



## C Results

Table A2: Treatment Effects for Coarse Liability Bins: Probability of Adoption, Infant Adoption, Domestic Adoption

	(1) Prob New Adoption	(2) Num New Infants	(3) Num New Domestic
ATT 0 Liab 2010	0.000231 (0.000148)	1.55 (1.53)	1.43 (1.69)
ATT Pos Liab 2010	0.000247* (0.000135)	2.13 (1.36)	1.94 (1.54)
ATT 0 Liab 2011	0.000297** (0.000151)	3.40** (1.44)	3.25** (1.60)
ATT Pos Liab 2011	0.000248* (0.000141)	3.36*** (1.28)	3.09** (1.47)
Observations	11,166,155	11,166,155	11,166,155

Note: The table displays the dynamic results for the regression design with coarse bins for several additional outcomes: the probability of performing a new adoption, the number of newly adopted children less than 2, and the number of newly adopted children not flagged as being an international adoptions. The coefficients in columns (2) and (3) are transformed to represent the treatment effect per 10,000 households. Under unconditional parallel trends, these may be interpreted as the average treatment effect on the treated, where treatment is conceptualized as the amount of additional credit the group gained access, though not actually used.

Table A3: Treatment Effects for Coarse Liability Bins Pooled Across Treatment Years 2010-2011

	(1)	(2)	(3)	(4)
	Num New Adoption	Prob New Adoption	Num New Infants	Num New Domestic
ATT 0 Liab Pooled	2.942** (1.47)	0.000234** (0.000110)	1.81* (1.06)	1.84 (1.18)
ATT Pos Liab Pooled	3.083** (1.346)	0.000253** (0.0000993)	2.47*** (0.932)	2.27** (1.07)
Observations	11,166,155	11,166,155	11,166,155	11,166,155

Note: The table displays the regression results when pooling samples together across treatment years for each of my outcomes of interest. The coefficients in columns (1), (3), and (4) are transformed to represent the treatment effect per 10,000 households. By pooling across time, these estimates provide a summary for the average treatment effect on the treated over the whole reform period. The previous analysis suggest, however, that these effect are likely being driven by the adoption response taking place in 2011.

Table A4: Treatment Effects and Placebos 5000 Bins: Number of Newly Adopted Kids Aged &lt; 18

	(1) Placebo ATT 2009	(2) ATT 2010	(3) ATT 2011	(4) Placebo ATT 2012	(5) Placebo ATT 2013	(6) Placebo ATT 2014
0 Liab	-1.22 (1.77)	2.08 (1.74)	3.76** (1.64)	1.05 (1.49)	-0.972 (1.46)	0.398 (1.45)
1-5000 Liab	-0.962 (1.64)	2.14 (1.65)	3.67** (1.61)	2.18 (1.58)	-1.56 (1.54)	0.420 (1.48)
5001-10000 Liab	0.926 (1.87)	2.65 (1.99)	2.73 (1.92)	-0.143 (1.80)	0.530 (1.70)	0.686 (1.68)
10001-13360 Liab	-5.76* (3.13)	4.13 (2.52)	6.88** (2.89)	1.27 (3.52)	-2.63 (3.19)	-0.764 (2.30)
Observations	11,166,155	11,166,155	11,166,155	17,802,417	15,528,557	15,528,557

Note: These results display the average treatment effect among the treated, given they received their average level of treatment relative to the control. The coefficients are transformed to represent the treatment effect per 10,000 households. The placebo ATTs in 2009 and years after 2011 are year over year estimates, assuming treatment took place in the given year, relative to the year prior. Insignificant results imply I fail to reject the hypothesis that the treated group is experiencing a differential trend relative to the control group in that given year.

Table A5: Treatment Effects and Placebos 5000 Bins: Probability of Performing New Adoption

	(1) Placebo ATT 2009	(2) ATT 2010	(3) ATT 2011	(4) Placebo ATT 2012	(5) Placebo ATT 2013	(6) Placebo ATT 2014
0 Liab	-0.0000775 (0.000159)	0.000231 (0.000148)	0.000297** (0.000151)	0.000160 (0.000140)	-0.000122 (0.000132)	0.0000413 (0.000134)
1-5000 Liab	-0.00000944 (0.000147)	0.000210 (0.000141)	0.000235 (0.000149)	0.000283* (0.000144)	-0.000196 (0.000135)	0.0000857 (0.000135)
5001-10000 Liab	0.000128 (0.000168)	0.000273 (0.000173)	0.000202 (0.000178)	0.0000282 (0.000166)	0.0000167 (0.000149)	0.0000825 (0.000154)
10001-13360 Liab	-0.000533* (0.000285)	0.000432** (0.000219)	0.000498** (0.000241)	0.000221 (0.000287)	-0.000246 (0.000266)	-0.0000329 (0.000216)
Observations	11,166,155	11,166,155	11,166,155	17,802,417	15,528,557	15,528,557

Note: These results display the average treatment effect among the treated, given they received their average level of treatment relative to the control. The placebo ATTs in 2009 and years after 2011 are year over year estimates, assuming treatment took place in the given year, relative to the year prior. Insignificant results imply I fail to reject the hypothesis that the treated group is experiencing a differential trend relative to the control group in that given year.

Table A6: Treatment Effects and Placebos 5000 Bins: Number of Newly Adopted Kids Aged  $< 2$ 

	(1) Placebo ATT 2009	(2) ATT 2010	(3) ATT 2011	(4) Placebo ATT 2012	(5) Placebo ATT 2013	(6) Placebo ATT 2014
0 Liab	-1.47 (1.58)	1.55 (1.53)	3.40** (1.44)	-0.0663 (1.29)	-0.171 (1.27)	-0.975 (1.28)
1-5000 Liab	-0.674 (1.39)	1.62 (1.41)	3.18** (1.37)	1.35 (1.35)	-0.921 (1.32)	-0.114 (1.32)
5001-10000 Liab	0.422 (1.61)	3.15* (1.75)	2.96* (1.62)	-1.81 (1.45)	1.57 (1.35)	0.0450 (1.47)
10001-13360 Liab	-5.02* (2.60)	2.59 (1.92)	6.04** (2.38)	1.09 (3.17)	-2.34 (2.86)	-1.16 (1.98)
Observations	11,166,155	11,166,155	11,166,155	17,802,417	15,528,557	15,528,557

Note: These results display the average treatment effect among the treated, given they received their average level of treatment relative to the control. The coefficients are transformed to represent the treatment effect per 10,000 households. The placebo ATTs in 2009 and years after 2011 are year over year estimates, assuming treatment took place in the given year, relative to the year prior. Insignificant results imply I fail to reject the hypothesis that the treated group is experiencing a differential trend relative to the control group in that given year.

Table A7: Treatment Effects and Placebos 5000 Bins: Number of Domestic Adoptions

	(1)	(2)	(3)	(4)	(5)	(6)
	Placebo ATT 2009	ATT 2010	ATT 2011	Placebo ATT 2012	Placebo ATT 2013	Placebo ATT 2014
0 Liab	-1.16 (1.72)	1.43 (1.69)	3.25** (1.60)	0.565 (1.47)	-1.22 (1.46)	0.445 (1.44)
1-5000 Liab	-0.891 (1.58)	1.52 (1.61)	3.05** (1.55)	1.73 (1.56)	-1.73 (1.54)	0.556 (1.47)
5001-10000 Liab	1.19 (1.81)	2.31 (1.94)	2.04 (1.83)	-0.486 (1.74)	0.635 (1.68)	0.806 (1.67)
10001-13360 Liab	-6.13** (2.94)	4.09* (2.33)	6.94** (2.72)	0.849 (3.55)	-3.08 (3.19)	-0.357 (2.24)
Observations	11,166,155	11,166,155	11,166,155	17,802,417	15,528,557	15,528,557

Note: These results display the average treatment effect among the treated, given they received their average level of treatment relative to the control. The coefficients are transformed to represent the treatment effect per 10,000 households. The placebo ATTs in 2009 and years after 2011 are year over year estimates, assuming treatment took place in the given year, relative to the year prior. Insignificant results imply I fail to reject the hypothesis that the treated group is experiencing a differential trend relative to the control group in that given year.

## D Figures

Figure A1: Probability of Undertaking a New Adoption

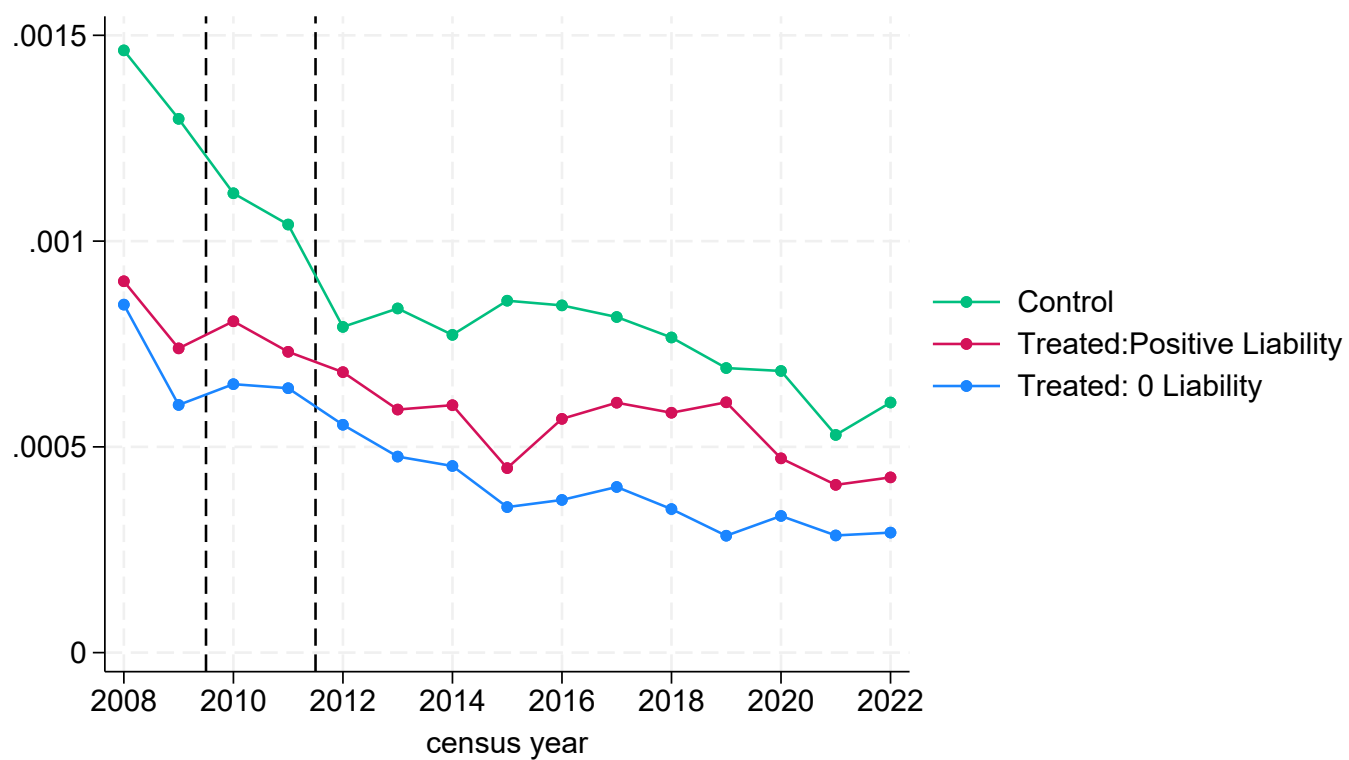


Figure A2: Mean Number of New Adoptions Aged <2 per Household

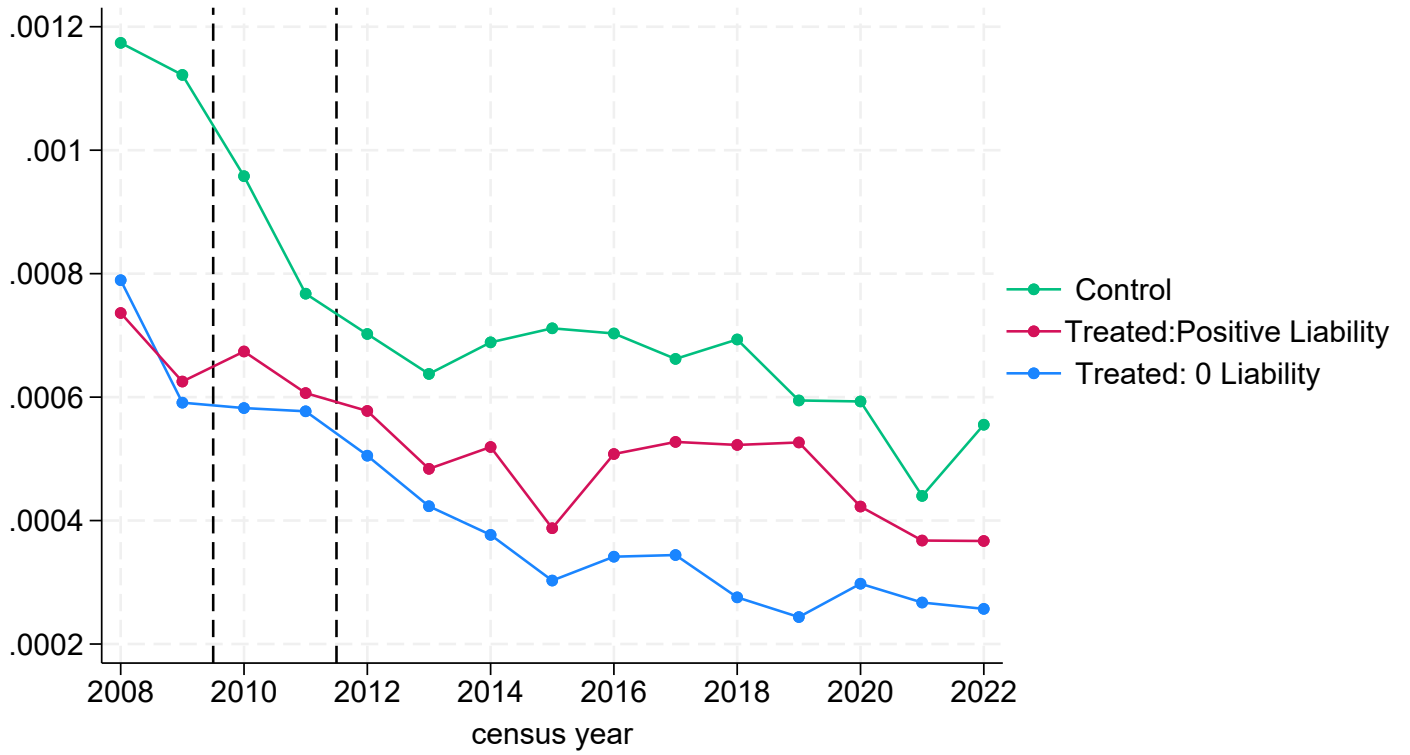




Figure A3: Mean Number of New Domestic Adoptions per Household

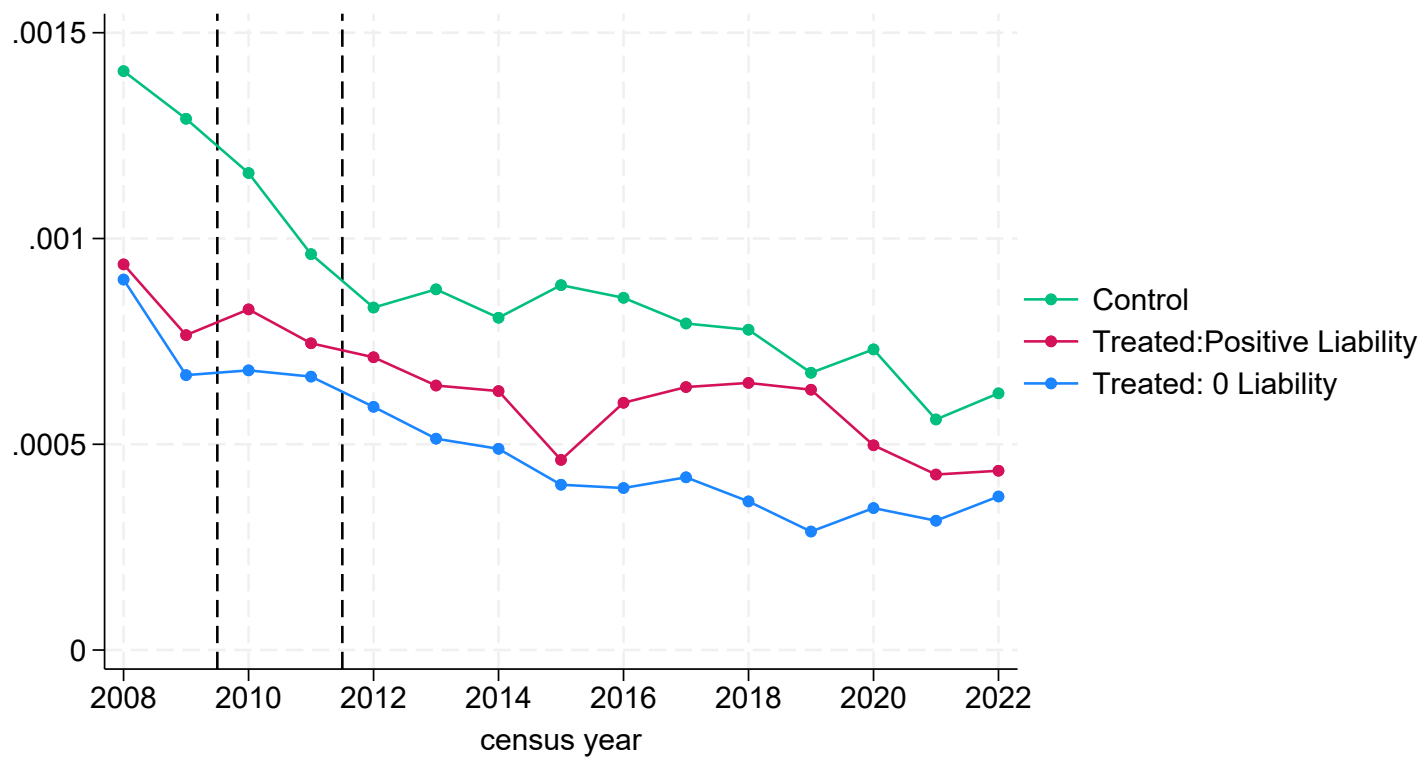


Figure A4: Mean Number of New Non-Step Adoptions per Household

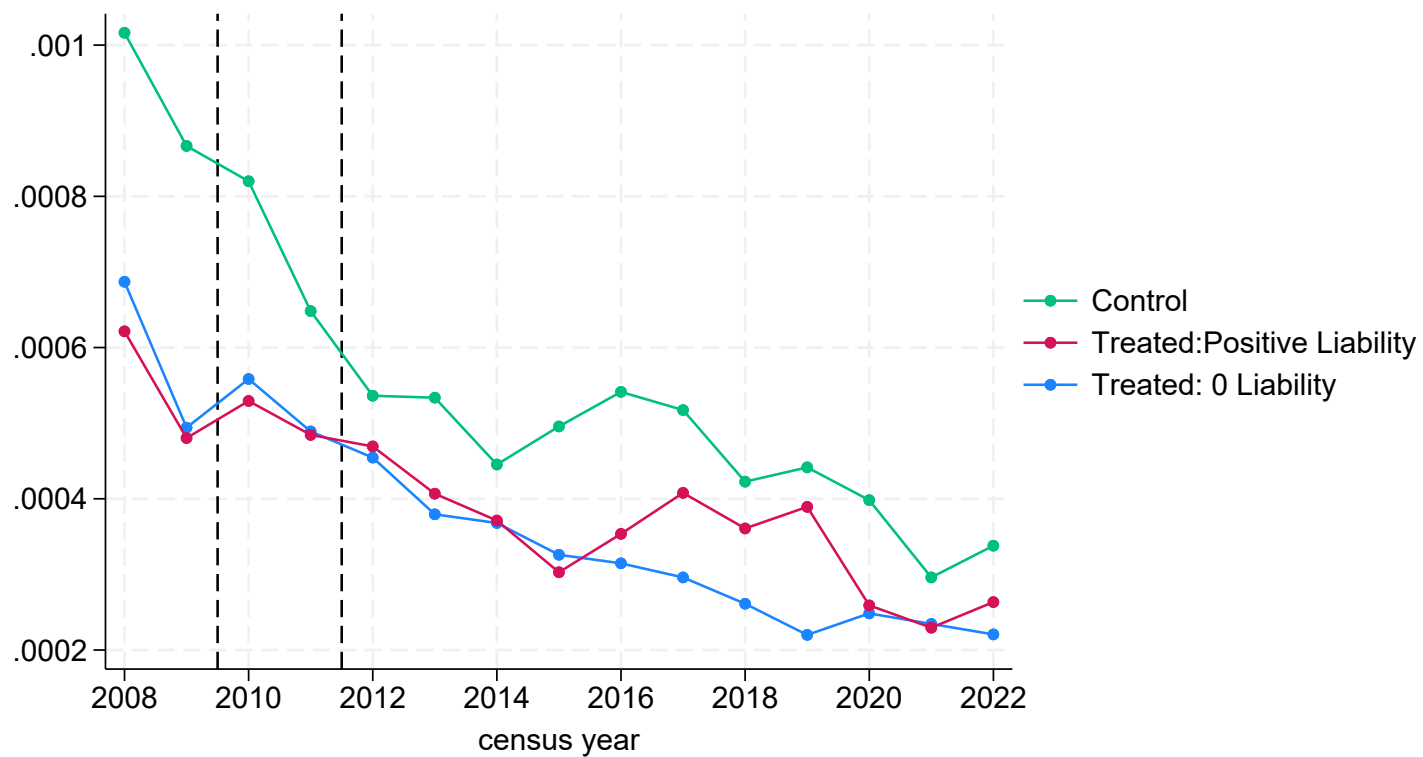


Figure A5: ATT Estimates With Placebo Tests for Number of New Adoptions

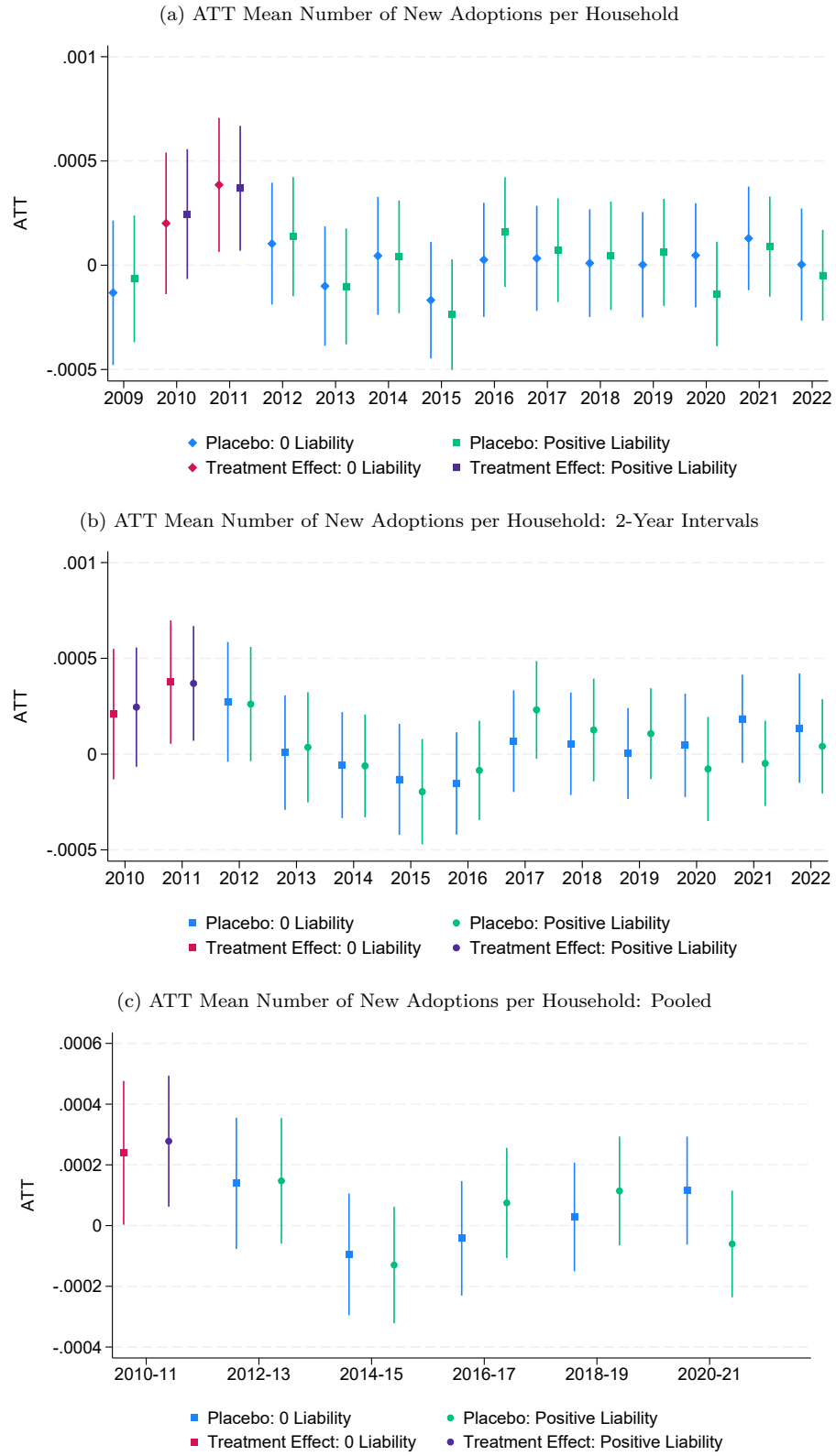


Figure A6: ATT Estimates With Placebo Tests for Probability of Performing New Adoption



Figure A7: ATT Estimates With Placebo Tests for Number of New Adoptions of Children <2



Figure A8: ATT Estimates With Placebo Tests for Number of New Domestic Adoptions

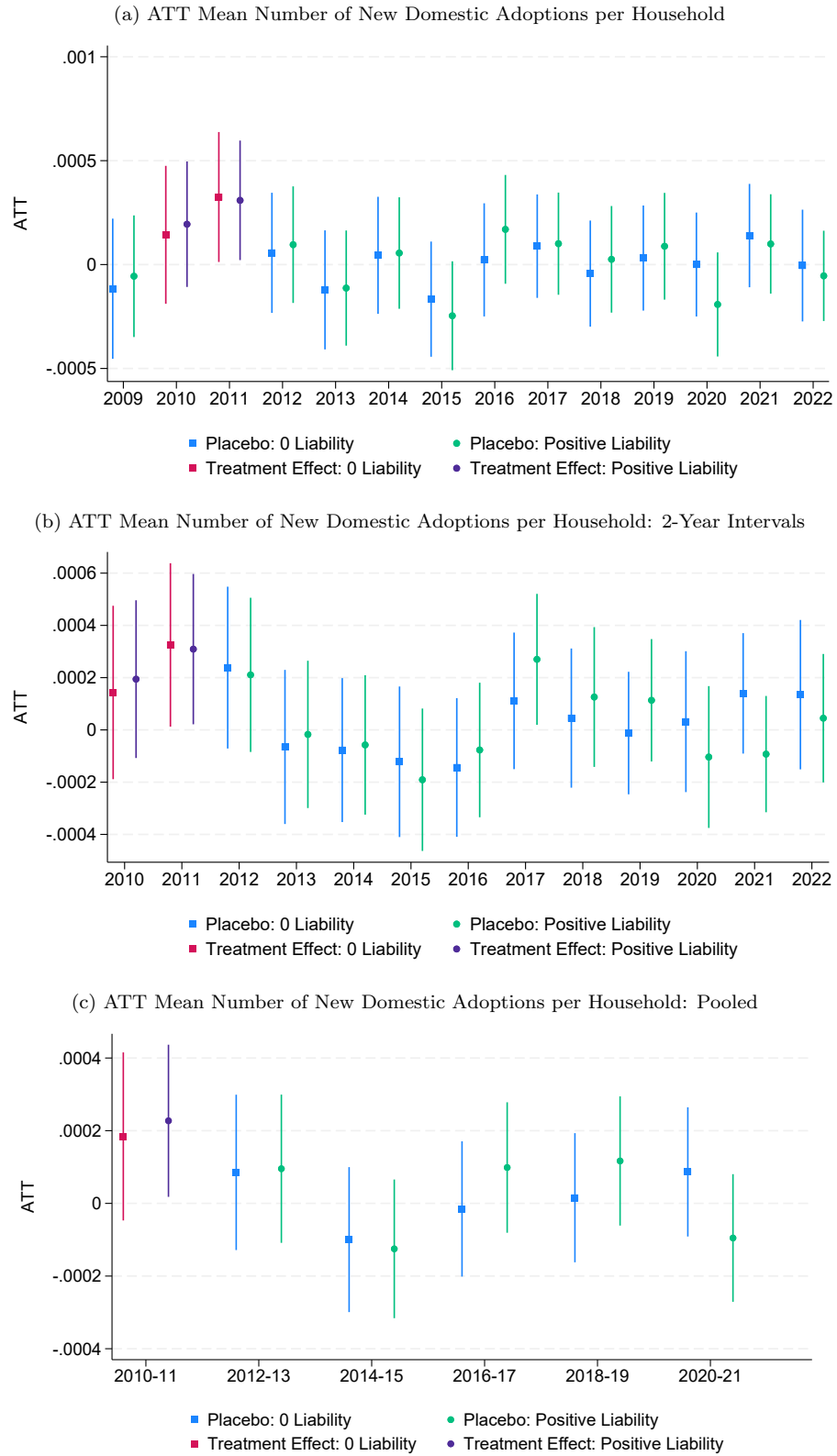


Figure A9: ATT Estimates With Placebo Tests for Number of New Non-Step Adoptions

