Copula: An Introduction

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An *n*-dimensional copula $C(u_1, \ldots, u_n)$ is a joint cumulative distribution function in the unit hypercube $[0,1]^n$ with uniform margins.

Example

Let n = 2, a copula function $C(u_1, u_2)$ can be written as

$$C(u_1,u_2) \equiv P(U_1 \leqslant u_1, U_2 \leqslant u_2),$$

where the random variables U_1 , U_2 are uniformly distributed on [0,1].

Given a random variable X, the cumulative distribution function (abbreviated as cdf) F_X of X is defined as

$$F_X(x) \equiv P(X \leqslant x).$$

The quantile function (generalized inverse) function F_X^{-1} of F_X is defined as

$$F_X^{-1}(u) = \inf\{x | F_X(x) \ge u\}, \quad u \in (0, 1).$$

Theorem

- Given a [0,1]-uniformly distributed random variable U and a cdf F, the random variable $X \equiv F^{-1}(U)$ has the cdf F.
- **2** The random variable $U \equiv F_X(X)$ is uniformly distributed on [0,1].

Proof.

1

$$F_X(x) = P(X \leqslant x) = P(F^{-1}(U) \leqslant x) = P(U \leqslant F(x)) = F(x).$$

The last equality holds for U is [0,1]-uniformly distributed.

2

$$F_U(u) = P(U \le u) = P(F_X(X) \le u)$$

= $P(X \le F_X^{-1}(u)) = F_X(F_X^{-1}(u)) = u, \quad u \in [0, 1]$

which is the cdf of a $\left[0,1\right]$ -uniformly distributed random variable.

Let X_1, X_2, \ldots, X_n be random variables with cdfs $F_{X_1}, F_{X_2}, \ldots, F_{X_n}$; the random variables $U_i \equiv F_{X_i}(X_i), i=1,2,\ldots,n$ are uniformly distributed on [0,1]. Furthermore, the function

$$C_{(X_{1},X_{2},...,X_{n})}(u_{1}, u_{2},..., u_{n})$$

$$\equiv P(U_{1} \leqslant u_{1}, U_{2} \leqslant u_{2},..., U_{n} \leqslant u_{n})$$

$$= P(F_{X_{1}}(X_{1}) \leqslant u_{1}, F_{X_{2}}(X_{2}) \leqslant u_{2},..., F_{X_{n}}(X_{n}) \leqslant u_{n})$$

$$= P\left(X_{1} \leqslant F_{X_{1}}^{-1}(u_{1}), X_{2} \leqslant F_{X_{2}}^{-1}(u_{2}),..., X_{n} \leqslant F_{X_{n}}^{-1}(u_{n})\right)$$

is a copula. The joint cdf of (X_1, X_2, \dots, X_n) can be recover as

$$\begin{split} &\mathsf{P}\left(X_{1} \leqslant x_{1}, X_{2} \leqslant x_{2}, \dots, X_{n} \leqslant x_{n}\right) \\ &= \mathsf{P}\left(F_{X_{1}}(X_{1}) \leqslant F_{X_{1}}(x_{1}), F_{X_{2}}(X_{2}) \leqslant F_{X_{2}}(x_{2}), \dots, F_{X_{n}}(X_{n}) \leqslant F_{X_{n}}(x_{n})\right) \\ &= \mathsf{P}\left(U_{1} \leqslant F_{X_{1}}(x_{1}), U_{2} \leqslant F_{X_{2}}(x_{2}), \dots, U_{n} \leqslant F_{X_{n}}(x_{n})\right) \\ &= C_{(X_{1}, X_{2}, \dots, X_{n})}(F_{X_{1}}(x_{1}), F_{X_{2}}(x_{2}), \dots, F_{X_{n}}(x_{n})) \end{split}$$

The above discussion culminates in the

Theorem (Sklar [31]; c.f. Durante and Sampi [9], Joe [17])

For a n-variate distribution F with j-th univariate margin F_j , the copula associated with F is a distribution function $C: [0,1]^n \to [0,1]$ with U(0,1) margins that satisfies

$$F(y) = C(F_1(y_1), F_2(y_2), \dots, F_n(y_n)), y \in \mathbb{R}^n.$$

If F is a continuous n-variate distribution function with univariate margins F_1, F_2, \ldots, F_n and quantile functions $F_1^{-1}, F_2^{-1}, \ldots, F_n^{-1}$, then

$$C(u) = F(F_1^{-1}(u_1), F_2^{-1}(u_2), \dots, F_n^{-1}(u_n))$$

is the unique choice.

The copula uniquely determined in $[0,1]^n$ for distributions F under absolutely continuous margins F_i has the density function

$$f(x_1,...,x_n) = c(F_1(x_1),...,F_n(x_n)) \prod_{i=1}^n f_i(x_i)$$

where f_i are the marginal densities and c is the density function of the copula given by

$$c(u_1,\ldots,u_n) = \frac{f(F_1^{-1}(u_1),\ldots,F_n^{-1}(u_n))}{\prod_{i=1}^n f_i(F_i^{-1}(u_i))}.$$

The copula function $C: [0,1]^n \to [0,1]$ has the following properties:

- $C(1,\ldots,1,u_i,1,\ldots,1)=u_i.$
- **③** For each *n*-dimensional rectangle $\mathbb{T} \equiv [a_1, b_1] \times [a_2, b_2] \times \cdots \times [a_n, b_n], a_i, b_i \in [0, 1], a_i < b_i,$

$$0 \leqslant \sum_{(d_1,d_2,\ldots,d_n)\in\mathbb{T}} (-1)^{|\{j:d_j=a_j\}|} C(d_1,d_2,\ldots,d_n) \leqslant 1.$$

For example, if n = 2, the condition reads

$$0 \leqslant (-1)^{0} C(b_{1}, b_{2}) + (-1)^{2} C(a_{1}, a_{2}) + (-1)^{1} C(b_{1}, a_{2}) + (-1)^{1} C(a_{1}, b_{2})$$

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