

Portfolio Optimization

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(For asset S_i , $\frac{x_i w}{S_{i,0}}$ denotes the “quantity” allocated at time 0; so at time 1

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- “For some fixed mean rate of return $\mu = \mathbf{E}\{\mathbf{x}^\top \mathbf{R}\}$, try to minimize the variance $\sigma^2 = \text{var}\{\mathbf{x}^\top \mathbf{R}\}$ of the return over portfolios \mathbf{x} ”

MV: All Risky Assets

$$\min_{\mathbf{x}} \frac{1}{2} \mathbf{x}^\top \mathbf{V} \mathbf{x} \quad \text{s.t.} \quad \begin{cases} \mathbf{x}^\top \mathbf{e} = 1 \\ \mathbf{x}^\top \mathbf{r} = \mu \end{cases} \quad \mathbf{e} \equiv \underbrace{(1, 1, \dots, 1)^\top}_{s \text{ items}}$$

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- Set $\mathcal{L} \equiv \frac{1}{2} \mathbf{x}^\top \mathbf{V} \mathbf{x} + \lambda (1 - \mathbf{x}^\top \mathbf{e}) + \nu (\mu - \mathbf{x}^\top \mathbf{r})$ with Lagrange multipliers λ, ν

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- Substitute into $\begin{cases} \mathbf{x}^\top \mathbf{e} = 1 \\ \mathbf{x}^\top \mathbf{r} = \mu \end{cases} \implies \begin{cases} \lambda \mathbf{e}^\top \mathbf{V}^{-1} \mathbf{e} + \nu \mathbf{r}^\top \mathbf{V}^{-1} \mathbf{e} = 1 \\ \lambda \mathbf{e}^\top \mathbf{V}^{-1} \mathbf{r} + \nu \mathbf{r}^\top \mathbf{V}^{-1} \mathbf{r} = \mu \end{cases}$

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$$\begin{aligned} &(\mathbf{r} - c\mathbf{e})^\top \mathbf{V}^{-1} (\mathbf{r} - c\mathbf{e}) > 0 \\ \implies &\mathbf{r}^\top \mathbf{V}^{-1} \mathbf{r} - c \mathbf{r}^\top \mathbf{V}^{-1} \mathbf{e} - c \mathbf{e}^\top \mathbf{V}^{-1} \mathbf{r} + c^2 \mathbf{e}^\top \mathbf{V}^{-1} \mathbf{e} > 0 \\ \implies &\gamma - 2c\beta + c^2\alpha > 0 \\ \implies &-\delta = \beta^2 - \gamma\alpha < 0 \end{aligned}$$

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- Here we have (σ, μ) with $a = \frac{1}{\sqrt{\alpha}}$, $b = \frac{\sqrt{\delta}}{\alpha}$, $h = 0$, $k = \frac{\beta}{\alpha}$, the asymptotes

$$\text{are } \left(\mu - \frac{\beta}{\alpha}\right) = \pm \frac{\frac{\sqrt{\delta}}{\alpha}}{\frac{1}{\sqrt{\alpha}}} \sigma \Rightarrow \mu = \frac{\beta}{\alpha} \pm \sqrt{\frac{\delta}{\alpha}} \sigma$$

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- Global minimum-variance portfolio \mathbf{x}_g

- First find μ_g that minimizes $\sigma^2 = \frac{\alpha\mu^2 - 2\beta\mu + \gamma}{\delta}$:

By differentiation $2\alpha\mu_g - 2\beta = 0 \implies \mu_g = \frac{\beta}{\alpha}$

- $\lambda_g = \frac{\gamma - \beta\mu_g}{\delta} = \frac{\gamma - \beta\frac{\beta}{\alpha}}{\delta} = \frac{\gamma\alpha - \beta^2}{\alpha\delta} = \frac{1}{\alpha}$

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Theorem (Mutual Fund Theorem)

Any minimum-variance portfolio is equivalent to investing in the convex combination of \mathbf{x}_g and \mathbf{x}_d .

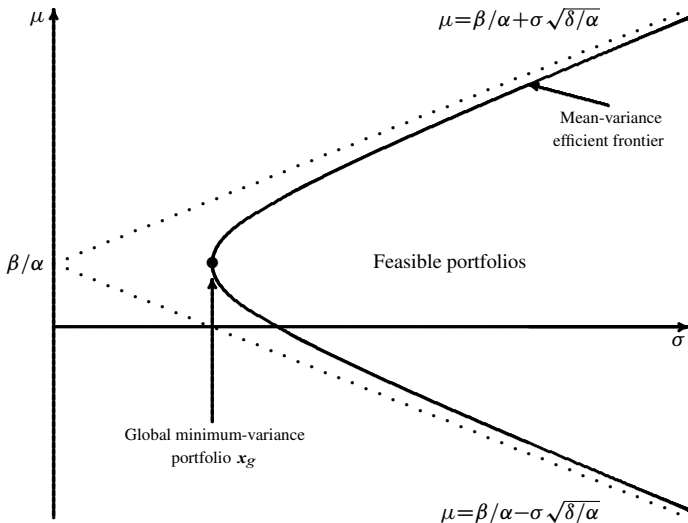


Figure: The Case of All Risky Assets

Theorem

Diversified portfolio \mathbf{x}_d is the portfolio that maximize $s(\mathbf{x}) \equiv \frac{\mathbf{x}^\top \mathbf{r}}{\sqrt{\mathbf{x}^\top \mathbf{V} \mathbf{x}}}$.

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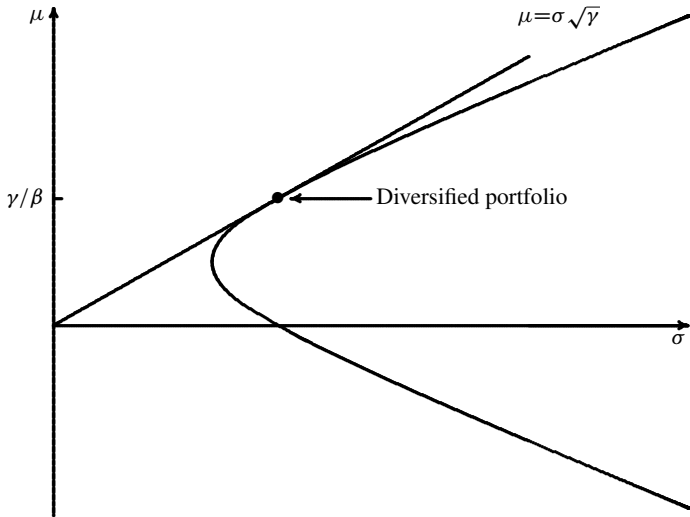


Figure: The Diversified Portfolio

MV: All But One Risky Assets

WLOG add riskless asset 0 with constant return r_0 ; the portfolio becomes $(x_0, x_1, x_2, \dots, x_s)^\top$

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- Set $\bar{\mathcal{L}} \equiv \frac{1}{2} \mathbf{x}^\top \mathbf{V} \mathbf{x} + \bar{\lambda} (1 - x_0 - \mathbf{x}^\top \mathbf{e}) + \bar{\nu} (\mu - x_0 r_0 - \mathbf{x}^\top \mathbf{r})$ with Lagrange multipliers $\bar{\lambda}, \bar{\nu}$

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- By $\frac{\partial \bar{\mathcal{L}}}{\partial x_0} = -\bar{\lambda} - \bar{\nu} r_0 = 0 \implies \bar{\nu} = -\frac{\bar{\lambda}}{r_0}$

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- Set $\alpha = \mathbf{e}^\top \mathbf{V}^{-1} \mathbf{e}$, $\beta = \mathbf{r}^\top \mathbf{V}^{-1} \mathbf{e} = \mathbf{e}^\top \mathbf{V}^{-1} \mathbf{r}$, $\gamma = \mathbf{r}^\top \mathbf{V}^{-1} \mathbf{r}$, $\delta \equiv \alpha\gamma - \beta^2$, the above becomes

$$\begin{cases} x_0 + \bar{\lambda}\alpha + \bar{\nu}\beta = x_0 + \bar{\lambda}\alpha - \frac{\bar{\lambda}}{r_0}\beta = 1 \\ x_0 r_0 + \bar{\lambda}\beta + \bar{\nu}\gamma = x_0 r_0 + \bar{\lambda}\beta - \frac{\bar{\lambda}}{r_0}\gamma = \mu \end{cases}$$

with solutions $x_0 = \frac{\alpha\mu r_0 - \beta r_0 + \gamma - \beta\mu}{\epsilon^2}$, $\bar{\lambda} = \frac{(r_0 - \mu)r_0}{\epsilon^2}$,

$\bar{\nu} = -\frac{r_0 - \mu}{\epsilon^2}$, where $\epsilon^2 = \alpha r_0^2 - 2\beta r_0 + \gamma = \alpha \left(r_0 - \frac{\beta}{\alpha}\right)^2 + \frac{\delta}{\alpha}$

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- The relation of σ with μ

$$\begin{aligned} \sigma^2 &= \mathbf{x}^\top \mathbf{V} \mathbf{x} = \mathbf{x}^\top \mathbf{V} (\bar{\lambda} \mathbf{V}^{-1} \mathbf{e} + \bar{\nu} \mathbf{V}^{-1} \mathbf{r}) = \bar{\lambda} (\mathbf{x}^\top \mathbf{e}) + \bar{\nu} (\mathbf{x}^\top \mathbf{r}) \\ &= \bar{\lambda} (1 - x_0) + \bar{\nu} (\mu - x_0 r_0) = \bar{\lambda} + \bar{\nu} \mu = \frac{(\mu - r_0)^2}{\epsilon^2} \end{aligned}$$

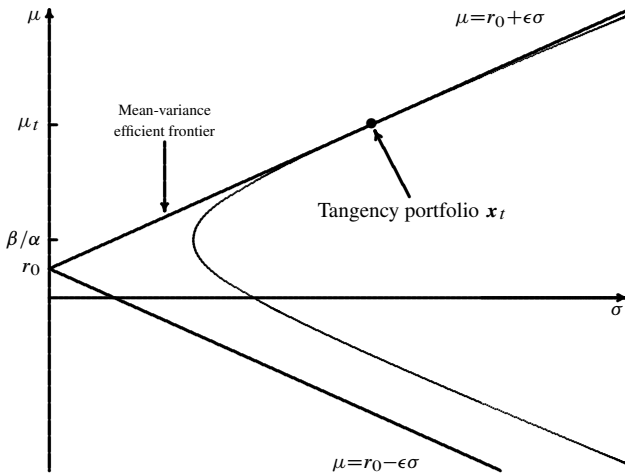


Figure: The Case of All But One Risky Assets

Property

If $r_0 < \frac{\beta}{\alpha}$, then $\mu = r_0 + \epsilon\sigma$ touches the hyperbola $\sigma^2 = \frac{\alpha\mu^2 - 2\beta\mu + \gamma}{\delta}$ at $\left(\frac{\epsilon}{\beta - \alpha r_0}, \frac{\gamma - \beta r_0}{\beta - \alpha r_0}\right)$

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Proof

On $\sigma - \mu$ plane the slope of the tangent $\mu'(\sigma)$ is obtained by implicit differentiation of $\sigma^2 = \frac{\alpha\mu^2 - 2\beta\mu + \gamma}{\delta}$ w.r.t σ (let $\mu \equiv \mu(\sigma)$):

$$2\sigma = \frac{2\alpha\mu\mu' - 2\beta\mu'}{\delta} \implies \mu' = \frac{\delta\sigma}{\alpha\mu - \beta}. \text{ Solve } \sigma, \mu \text{ from } \mu = r_0 + \epsilon\sigma \text{ and } \epsilon = \frac{\delta\sigma}{\alpha\mu - \beta}, \text{ we obtain } (\sigma, \mu) = \left(\frac{\epsilon}{\beta - \alpha r_0}, \frac{\gamma - \beta r_0}{\beta - \alpha r_0}\right).$$

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$$\mathbf{x}_t = \frac{1}{\beta - \alpha r_0} \mathbf{V}^{-1}(\mathbf{r} - r_0 \mathbf{e}) = \frac{\beta}{\beta - \alpha r_0} \mathbf{x}_d - \frac{\alpha r_0}{\beta - \alpha r_0} \mathbf{x}_g$$

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 $= \frac{\beta}{\beta - \alpha r_0} \mu_d - \frac{\alpha r_0}{\beta - \alpha r_0} \mu_g = \frac{\gamma - \beta r_0}{\beta - \alpha r_0}$ for $\mu_d = \frac{\gamma}{\beta}$, $\mu_g = \frac{\beta}{\alpha}$

Theorem

Tangency portfolio \mathbf{x}_t is the portfolio that maximize $s(\mathbf{x}) \equiv \frac{\mathbf{x}^\top \mathbf{r} - r_0}{\sqrt{\mathbf{x}^\top \mathbf{V} \mathbf{x}}}$.

Proof

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Set $\beta_t \equiv (\beta_{1,t}, \beta_{2,t}, \dots, \beta_{s,t})^\top$ where $\beta_{i,t} = \frac{\text{cov}(R_i, \mathbf{x}_t^\top \mathbf{R})}{\text{var}(\mathbf{x}_t^\top \mathbf{R})}$, or

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Mean-Variance Analysis and Expected Utility

$f(\sigma, \mu) = E v(W)$ where $W = (x_0 r_0 + \mathbf{x}^\top \mathbf{R})w$ with $\mu = x_0 r_0 + \mathbf{x}^\top \mathbf{r}$,
 $\sigma^2 = \mathbf{x}^\top \mathbf{V} \mathbf{x}$. Assume that $\frac{\partial f}{\partial \sigma} < 0$, $\frac{\partial f}{\partial \mu} > 0$ with $x_0 + \mathbf{x}^\top \mathbf{e} = 1$,

$$\max_{\mathbf{x}} f(\sqrt{\mathbf{x}^\top \mathbf{V} \mathbf{x}}, r_0 + \mathbf{x}^\top (\mathbf{r} - r_0 \mathbf{e})).$$

$$\frac{\partial f}{\partial \mathbf{x}} = \frac{1}{\sigma} \frac{\partial f}{\partial \sigma} \mathbf{V} \mathbf{x} + \frac{\partial f}{\partial \mu} (\mathbf{r} - r_0 \mathbf{e}) = 0 \implies \mathbf{x} = -\frac{\sigma \frac{\partial f}{\partial \mu}}{\frac{\partial f}{\partial \sigma}} \mathbf{V}^{-1} (\mathbf{r} - r_0 \mathbf{e}) \propto \mathbf{x}_t$$

quadratic utility: $v(x) = ax + bx^2$ where $a, b \in \mathbb{R}$, $b \leq 0$. \implies

$$E v(W) = E v((x_0 r_0 + \mathbf{x}^\top \mathbf{R})w) = aw\mu + bw^2(\mu^2 + \sigma^2) = f(\sigma, \mu).$$

normally distributed returns: $\mathbf{R} \sim N(\mathbf{r}, \mathbf{V}) \implies \mathbf{x}^\top \mathbf{R} \sim N(\mathbf{x}^\top \mathbf{r}, \mathbf{x}^\top \mathbf{V} \mathbf{x}) \implies$

$$E v(W) = E v((\mu + \sigma Y)w), \text{ where } Y \sim N(0, 1).$$

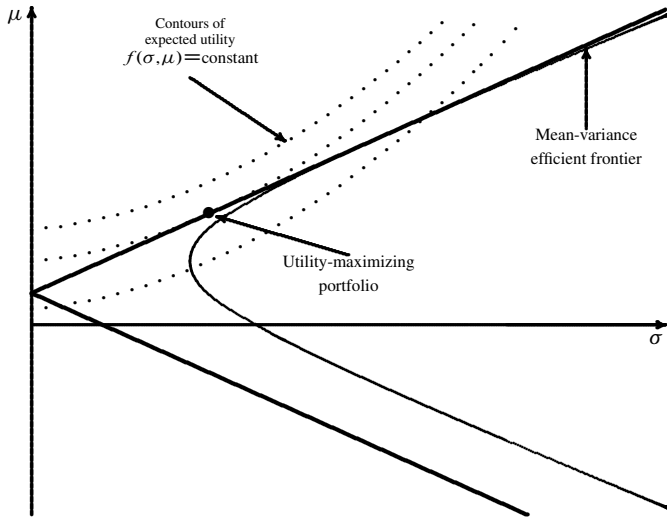


Figure: Determining the Utility-Maximizing Portfolio

Equilibrium: The Capital-Asset Pricing Model

Market with investors indexed by $j \in \mathcal{J}$ with proportions of wealth $x_{0,j}$,

$$\mathbf{x}_j = (x_{1,j}, x_{2,j}, \dots, x_{s,j})^\top$$

When each investor j has the utility function as above, the optimal $\mathbf{x}_j \propto \mathbf{x}_t \implies$

$$\mathbf{x}_j = (1 - x_{0,j})\mathbf{x}_t \quad \forall j \in \mathcal{J}.$$

The total value of the demand for risky asset i :

$$\sum_{j \in \mathcal{J}} w_j x_{i,j} = \left(\sum_{j \in \mathcal{J}} (1 - x_{0,j}) w_j \right) (\mathbf{x}_t)_i$$

Market portfolio of risky assets \mathbf{x}_m :

$$(\mathbf{x}_m)_i \equiv \frac{\text{The total value of the supply of risky asset } i}{\text{The total value of the supply of all risky assets}}, \quad \mathbf{x}_m^\top \mathbf{e} = 1.$$

In equilibrium

$$(\mathbf{x}_m)_i = \frac{\left(\sum_{j \in \mathcal{J}} (1 - x_{0,j}) w_j \right) (\mathbf{x}_t)_i}{\sum_{j \in \mathcal{J}} \sum_{k=1}^s w_j x_{k,j}} = \frac{\left(\sum_{j \in \mathcal{J}} (1 - x_{0,j}) w_j \right) (\mathbf{x}_t)_i}{\left(\sum_{j \in \mathcal{J}} (1 - x_{0,j}) w_j \right) \sum_{k=1}^s (\mathbf{x}_t)_k} = (\mathbf{x}_t)_i$$

$$\text{since } \mathbf{x}^\top \mathbf{e} = 1. \implies \mathbf{r} = r_0 \mathbf{e} + (\mu_m - r_0) \boldsymbol{\beta}_m, \quad \boldsymbol{\beta}_m \equiv (\beta_{1,m}, \beta_{2,m}, \dots, \beta_{s,m})^\top,$$

$$\beta_{i,m} = \frac{\text{cov}(R_i, \mathbf{x}_m^\top \mathbf{R})}{\text{var}(\mathbf{x}_m^\top \mathbf{R})} \quad \text{— capital-asset-pricing equation}$$

Problem

Suppose that an investment X has either (i) the uniform distribution $U[0, 2\mu]$ or (ii) the exponential distribution with $E X = \mu$, and the investor has a utility function which is either (a) logarithmic, $v(x) = \log x$ (b) power form, $v(x) = x^\theta$. Show that both the compensatory risk premium and the investment risk premium are proportional to μ in all 4 possible cases.

Solution I

For distributions (i)(ii) of X , the r.v. $Y \equiv \frac{X}{\mu}$ does not depend on μ , so the equation $E v(X + \alpha) = v(\mu)$ for the compensatory risk premium α reduces to $E v(Y + c) = v(1)$ in cases (a)(b) when $\alpha = c\mu$. For the insurance risk premium when $\beta = d\mu$, d is the solution of $E v(Y) = v(1 - d)$.

For case (i)(a),

$$E v(Y + c) = \int_0^2 \frac{\log(y + c)}{2} dy = \frac{1}{2}((2 + c) \log(2 + c) - c \log c - 2), \text{ and}$$

$v(1) = \log 1 = 0$, so $\alpha = c\mu$ where c is the unique positive root of $(2 + c) \log(2 + c) - c \log c - 2 = 0$. Using `rmaxima` we have $c = 0.176965531$.

For the insurance premium $\beta = d\mu$, $E \log Y = \log 2 - 1 = \log(1 - d)$, so

$$d = 1 - \frac{2}{e} = 0.264.$$

Problem

An investor has a utility function $v(x) = \sqrt{x}$ and is considering three investments with random outcomes X, Y, Z . Here X has the uniform distribution $U[0, a]$, Y has the gamma distribution $\Gamma(\gamma, \lambda)$ with probability density function $\frac{e^{-\lambda y} \lambda^\gamma y^{\gamma-1}}{\Gamma(\gamma)}$ for $y > 0$, where $\gamma > 0$, $\lambda > 0$, and Z is log-normal, i.e. $Z \sim N(\nu, \sigma^2)$. The parameter of the distributions are such that $E X = E Y = E Z = \mu$ and $\text{var } X = \text{var } Y = \text{var } Z$. Recall that the gamma function $\Gamma(\gamma) = \int_0^\infty u^{\gamma-1} e^{-u} du$ that satisfies $\Gamma(\gamma + 1) = \gamma \Gamma(\gamma)$ and $\Gamma(1/2) = \sqrt{\pi}$. Determine the investor's preference ordering of X, Y, Z for all values of μ .

Solution I

- $X \sim U[0, a] \implies \mathbb{E} X = \frac{a}{2}, \text{ var } X = \frac{a^2}{12}$
- $Y \sim \Gamma(\gamma, \lambda) \implies \mathbb{E} Y = \frac{\gamma}{\lambda}, \text{ var } Y = \frac{\gamma}{\lambda^2}$
- $Z \sim \text{lognormal}(\nu, \sigma^2) \implies \mathbb{E} Z = e^{\nu + \frac{\sigma^2}{2}}, \text{ var } Z = e^{2\nu + \sigma^2}(e^{\sigma^2} - 1)$ by the formula $\mathbb{E} e^{\theta W} = e^{\mu\theta + \frac{\theta^2 \sigma^2}{2}}$ for $W \sim N(\mu, \sigma^2)$

$$\begin{aligned}\mathbb{E} e^{\theta W} &= e^{\mu\theta + \frac{\theta^2 \sigma^2}{2}} \text{ as } W \sim N(\mu, \sigma^2): \sqrt{2\pi}\sigma \mathbb{E} e^{\theta W} = \int_{-\infty}^{\infty} e^{\theta x} \cdot e^{-\frac{1}{2} \frac{x^2 - 2\mu x + \mu^2}{\sigma^2}} dx \\&= \int_{-\infty}^{\infty} e^{-\frac{1}{2} \frac{x^2 - 2(\mu + \theta\sigma^2)x + \mu^2}{\sigma^2}} dx = \int_{-\infty}^{\infty} e^{-\frac{1}{2} \frac{(x - (\mu + \theta\sigma^2))^2 + \mu^2 - (\mu + \theta\sigma^2)^2}{\sigma^2}} dx \\&= \int_{-\infty}^{\infty} e^{-\frac{1}{2} \frac{(x - (\mu + \theta\sigma^2))^2 - 2\mu\theta\sigma^2 - (\theta\sigma^2)^2}{\sigma^2}} dx = e^{\mu\theta + \frac{\theta^2 \sigma^2}{2}} \cdot \int_{-\infty}^{\infty} e^{-\frac{1}{2} \frac{(x - (\mu + \theta\sigma^2))^2}{\sigma^2}} dx \\&= e^{\mu\theta + \frac{\theta^2 \sigma^2}{2}} \cdot \int_{-\infty}^{\infty} e^{-\frac{x^2}{2\sigma^2}} dx = \sqrt{2\pi}\sigma \cdot e^{\mu\theta + \frac{\theta^2 \sigma^2}{2}} \text{ by } \int_{-\infty}^{\infty} e^{-x^2} dx = \sqrt{\pi}.\end{aligned}$$

The conditions $\mathbb{E} X = \mathbb{E} Y = \mathbb{E} Z = \mu$ and $\text{var } X = \text{var } Y = \text{var } Z$ imply

- $a = 2\mu$, so that $\text{var } X = \frac{\mu^2}{3}$.

Solution II

- $EY = \frac{\gamma}{\lambda} = \mu$, so that $\text{var } Y = \frac{\gamma}{\lambda^2} = \frac{\mu^2}{\gamma} = \text{var } X = \frac{\mu^2}{3} \implies \gamma = 3$
- $EZ = e^{\nu + \frac{\sigma^2}{2}} = \mu$, $\text{var } Z = e^{2\nu + \sigma^2}(e^{\sigma^2} - 1) = \mu^2(e^{\sigma^2} - 1) = \text{var } X = \frac{\mu^2}{3}$
 $\implies \sigma^2 = \log \frac{4}{3}$.
- $E\sqrt{X} = \int_0^{2\mu} \frac{\sqrt{x}}{2\mu} dx = \frac{2^{\frac{3}{2}}}{3} \sqrt{\mu} \approx 0.943\sqrt{\mu}$
- $E\sqrt{Y} = \int_0^\infty \sqrt{y} \frac{1}{2} e^{-\lambda y} \lambda^3 y^2 dy = \frac{\Gamma(\frac{7}{2})}{2\sqrt{\lambda}} = \frac{\Gamma(\frac{7}{2})}{2\sqrt{3}} \sqrt{\mu} = \frac{5\sqrt{3}\pi}{16} \sqrt{\mu} \approx 0.959\sqrt{\mu}$
- $E\sqrt{Z} = e^{\frac{\nu}{2} + \frac{\sigma^2}{8}} = e^{-\frac{\sigma^2}{8}} \sqrt{\mu} = \left(\frac{3}{4}\right)^{\frac{1}{8}} \sqrt{\mu} \approx 0.965\sqrt{\mu}$

So $Z \succ Y \succ X$.

Problem

Suppose that an investor has the utility function $v(x) = 1 - e^{-ax}$ with $a > 0$, and the outcome of an investment is a r.v. X with mean μ , finite variance and finite moment-generating function $\psi(a) = E\{e^{-ax}\}$ for $a > 0$.

- 1 Show that the compensatory risk premium and the insurance risk premium have the same value α , and express α in terms of μ and the moment generating function ψ .
- 2 Both the Arrow-Pratt and global risk aversions are a . Confirm directly that as $a \downarrow 0$, $\alpha = a \frac{\text{var } X}{2} + o(a)$. Under what circumstances is it true that $\alpha = a \frac{\text{var } X}{2}$ for all $a > 0$?
- 3 Prove that $\psi''\psi - (\psi')^2 \geq 0$ and hence α is an increasing function of a . This shows that the more risk-averse the investor is, the higher the value of the premium that is required.

Solution I

The compensatory risk premium α solves $\mathbb{E} v(\alpha + X) = v(\mu)$ while the insurance risk premium β solves $\mathbb{E} v(X) = v(\mu - \beta)$ giving the common value:

$$\alpha = \beta = \mu + \frac{1}{a} \ln(\psi(a)).$$

The expansion for small a is straightforward; when $\alpha = \frac{a \operatorname{var} X}{2}$ for all $a > 0$ we have:

$$\psi(a) = \mathbb{E} e^{-aX} = e^{-a\mu + \frac{a^2 \operatorname{var} X}{2}}$$

which is true only when X has a normal distribution, using (A.16). For the final part:

$$\psi''\psi - (\psi')^2 = \mathbb{E} X^2 e^{-aX} \mathbb{E} e^{-aX} - (\mathbb{E} X e^{-aX})^2 \geq 0$$

by the Cauchy-Schwarz inequality applied to the random variables $A = X e^{-aX/2}$ and $B = e^{-aX/2}$. To see that α is increasing:

$$\frac{d\alpha}{da} = \frac{1}{a^2} \left[\frac{a\psi'}{\psi} - \ln(\psi) \right] = \frac{1}{a^2} f(a), \text{ say.}$$

Solution II

But $f(0) = 0$ and $f' = a[\psi''\psi - (\psi')^2]/\psi^2 \geq 0$ and the conclusion follows.

Problem

Consider a one-period investment model in which there are only two risky assets. The returns on these assets have means 3, 4 respectively and variances 2, 3 respectively with the covariance between the returns being 2.

- 1 Calculate the mean-variance efficient frontier and the minimum-variance portfolio in terms of the mean return.
- 2 Calculate the mean return of the global minimum-variance portfolio and of diversified portfolio.
- 3 Suppose that in addition to the two risky assets there is a riskless asset with return $3/2$. Find the minimum-variance portfolio in terms of the mean return and hence calculate the mean return of the tangency portfolio.

Solution I

The inverse matrix of $\begin{pmatrix} a & b \\ c & d \end{pmatrix}$ is $\frac{1}{ad-bc} \begin{pmatrix} d & -b \\ -c & a \end{pmatrix}$, so if $V = \begin{pmatrix} 2 & 2 \\ 2 & 3 \end{pmatrix}$,
 $V^{-1} = \frac{1}{2} \begin{pmatrix} 3 & -2 \\ -2 & 2 \end{pmatrix}$. $\alpha = \mathbf{e}^\top \mathbf{V}^{-1} \mathbf{e} = \frac{1}{2}$, $\beta = \mathbf{e}^\top \mathbf{V}^{-1} \mathbf{r} = \frac{3}{2}$, $\gamma = \mathbf{r}^\top \mathbf{V}^{-1} \mathbf{r} = \frac{11}{2}$,
 $\delta = \alpha\gamma - \beta^2 = \frac{1}{2}$.

① $\min_{x_1, x_2} (x_1 \ x_2) \begin{pmatrix} 2 & 2 \\ 2 & 3 \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \end{pmatrix} = \min_{x_1, x_2} 2x_1^2 + 4x_1x_2 + 3x_2^2 \text{ s.t.}$
 $\begin{cases} x_1 + x_2 = 1 \\ 3x_1 + 4x_2 = \mu \end{cases}$. From constraints $x_1 = 4 - \mu$, $x_2 = \mu - 3$, so the
mean-variance efficient frontier is $\sigma^2 = \mu^2 - 6\mu + 11$.

② μ_g is the root of $\frac{d\sigma^2}{d\mu} = 0$, so $2\mu_g - 6 = 0 \implies \mu_g = 3$. $\mu_d = \frac{\gamma}{\beta} = \frac{11}{3}$.

Solution II

③ Now the problem is $\min_{x_0, x_1, x_2} 2x_1^2 + 4x_1x_2 + 3x_2^2$ s.t.

$$\begin{cases} x_0 + x_1 + x_2 = 1 \\ \frac{3}{2}x_0 + 3x_1 + 4x_2 = \mu \end{cases}. \text{ Form the Lagrangian}$$

$$\mathcal{L} = 2x_1^2 + 4x_1x_2 + 3x_2^2 + \lambda(1 - x_0 - x_1 - x_2) + \nu(\mu - \frac{3}{2}x_0 - 3x_1 - 4x_2).$$

By solving $\frac{\partial \mathcal{L}}{\partial x_0} = 0$, $\nu = -\frac{2\lambda}{3}$. From $\frac{\partial \mathcal{L}}{\partial x_1} = 0$ and $\frac{\partial \mathcal{L}}{\partial x_2} = 0$ we have

$$4x_1 + 4x_2 - \lambda - 3\nu = 0 \text{ and } 4x_1 + 6x_2 - \lambda - 4\nu = 0; \text{ so } x_1 = \frac{\lambda}{12}, x_2 = -\frac{\lambda}{3}.$$

Substitute into the constraints yields $\lambda = \frac{12(3 - 2\mu)}{17}$, and so $x_0 = \frac{26 - 6\mu}{17}$,

$$x_1 = \frac{3 - 2\mu}{17}, x_2 = -\frac{4(3 - 2\mu)}{17}. \text{ The tangency portfolio corresponds to}$$

$$x_0 = 0 \text{ or } \mu_t = \frac{13}{3}.$$

Problem

Suppose that v is concave, $X \sim N(\mu, \sigma^2)$ and $f(\sigma, \mu) = E v(X)$.

- 1 Show that $\frac{\partial f}{\partial \mu} > 0$ when v is strictly increasing, and $\frac{\partial f}{\partial \sigma} \leq 0$. Hence show in the context of mean-variance analysis that, when all returns are jointly normally distributed, an investor maximizing the expected utility of his final wealth will choose a mean-variance-efficient optimal portfolio.
- 2 Show that f is concave in μ and σ . Deduce that this optimal portfolio corresponds to a point in the (σ, μ) plane where an indifference contour is tangent to the efficient frontier.

Solution I

Write $X = \mu + \sigma Y$ where $Y \sim N(0, 1)$. Then it follows that:

$$\frac{\partial f}{\partial \mu} = E\{v'(\mu + \sigma Y)\} > 0 \text{ when } v' > 0,$$

and using the relation (A.14):

$$\frac{\partial f}{\partial \sigma} = E\{Y v'(\mu + \sigma Y)\} = \sigma E\{v''(\mu + \sigma Y)\} \leq 0,$$

by the concavity of v . Now when returns are normally distributed then the wealth created by each portfolio has a normal distribution; this argument shows that maximizing in σ for fixed μ gives a value of (σ, μ) on the efficient frontier. To see the concavity of f , note that:

$$\frac{\partial^2 f}{\partial \mu^2} = E\{v''(\mu + \sigma Y)\} \leq 0 \text{ and } \frac{\partial^2 f}{\partial \sigma^2} = E\{Y^2 v''(\mu + \sigma Y)\} \leq 0,$$

$$\frac{\partial^2 f}{\partial \mu \partial \sigma} = E\{Y v''(\mu + \sigma Y)\},$$

and then:

$$\frac{\partial^2 f}{\partial \mu^2} \frac{\partial^2 f}{\partial \sigma^2} \geq \left(\frac{\partial^2 f}{\partial \mu \partial \sigma} \right)^2$$

follows by applying the Cauchy-Schwarz inequality to the random variables $A = Y \sqrt{-v''(\mu + \sigma Y)}$ and $B = \sqrt{-v''(\mu + \sigma Y)}$; this shows that the 2×2 matrix of second derivatives has non-positive diagonal entries and a non-negative determinant which is sufficient for the matrix to be negative semi-definite. The fact that f is concave means that sets of the form $\{(\sigma, \mu) : f(\sigma, \mu) > c\}$ are convex which gives the last statement.

Problem

Suppose that an investor has a concave utility function v . The investor seeks to maximize $E v(W)$ where $W = (x_0 r_0 + \mathbf{x}^\top \mathbf{R})w$ is his final wealth.

- ① Show that, when \bar{W} is his optimal final wealth, then

$$E\{v'(\bar{W})(R_j - r_0)\} = 0, \quad \forall j = 1, 2, \dots, s.$$

- ② Show that, when \mathbf{R} has a multivariate normal distribution, then

$$r_j - r_0 = \alpha \operatorname{cov}(\bar{W}, R_j), \quad \forall j = 1, 2, \dots, s, \quad \text{where } \alpha = -\frac{E\{v''(\bar{W})\}}{E\{v'(\bar{W})\}} \text{ is his global risk aversion.}$$

- ③ Now suppose that the market is determined by investors $i = 1, 2, \dots, n$, where investor i has concave utility v_i , initial wealth w_i , optimal final wealth \bar{W}_i and global risk aversion α_i . With the normality assumption, show that

$$E M - r_0 = \bar{w} \bar{\alpha} \operatorname{var} M, \quad \text{where } M = \frac{\sum_{i=1}^n \bar{W}_i}{\sum_{i=1}^n w_i} \text{ is the market rate of return,}$$

$$\bar{w} = \frac{\sum_{i=1}^n w_i}{n} \text{ is the average initial wealth of investors, and } \bar{\alpha} \text{ is the harmonic mean of } \alpha_i.$$

Solution 1

The objective function to maximize is:

$$f(\mathbf{x}) = \mathbb{E} v \left(w \left(r_0 + \sum_{j=1}^s x_j (R_j - r_0) \right) \right)$$

where $\mathbf{x} = (x_1, \dots, x_s)^\top$ and we have used the condition that $x_0 + \sum_{j=1}^s x_j = 1$.

The first-order conditions give:

$$\frac{\partial f}{\partial x_j} = w \mathbb{E} v'(\overline{W})(R_j - r_0) = 0, \text{ for } 1 \leq j \leq s.$$

Since $r_j = \mathbb{E} R_j$ and the fact that \overline{W} and R_j have a joint normal distribution we have that:

$$\begin{aligned} 0 &= \mathbb{E} v'(\overline{W})(R_j - r_0) = \mathbb{E}\{v'(\overline{W})(R_j - r_j)\} + \mathbb{E}\{v'(\overline{W})\}(r_j - r_0) \\ &= \text{cov}(v'(\overline{W}), R_j) + \mathbb{E}\{v'(\overline{W})\}(r_j - r_0) \\ &= \mathbb{E}\{v''(\overline{W})\} \text{cov}(\overline{W}, R_j) + \mathbb{E}\{v'(\overline{W})\}(r_j - r_0), \end{aligned}$$

Solution II

where the last equality uses (A.21), and this now gives the relation:

$$r_j - r_0 = \alpha \operatorname{cov}(\overline{W}, R_j),$$

as required.

For the final part, recall that for random variables X and Y and a a constant $\operatorname{cov}(X, Y + a) = \operatorname{cov}(X, Y)$ and $\operatorname{cov}(aX, Y) = a \operatorname{cov}(X, Y)$. Now for each i :

$$\alpha_i^{-1}(r_j - r_0) = \operatorname{cov}(\overline{W}_i, R_j)$$

and summing this on i yields:

$$\left(\sum_{i=1}^n \alpha_i^{-1} \right) (r_j - r_0) = \left(\sum_{i=1}^n w_i \right) \operatorname{cov}(M, R_j).$$

Divide through this relation by n and multiply by $\bar{\alpha}$, where $(\bar{\alpha})^{-1} = \sum_{i=1}^n \alpha_i^{-1} / n$, to obtain:

$$\mathbb{E} R_j - r_0 = w \bar{\alpha} \operatorname{cov}(M, R_j).$$

Solution III

When \bar{x}_{ij} is the optimal proportion invested by investor i in asset j then:

$$\bar{W}_i = w_i \left[r_0 + \sum_{j=1}^s \bar{x}_{ij}(R_j - r_0) \right]$$

which when summed on i gives:

$$(M - r_0) \left(\sum_{i=1}^n w_i \right) = \sum_{i=1}^n \sum_{j=1}^s w_i \bar{x}_{ij}(R_j - r_0).$$

Take the expectation in (B.3), multiply (B.2) by $w_i \bar{x}_{ij}$, sum on i and j , rearrange the expression using the two properties of covariance mentioned above and the result (1.21) follows. This shows that the risk premium for the market is proportional to $\bar{\alpha}$ which is a measure of the risk aversion in the economy.

Problem

Consider an investor with the utility function $v(x) = 1 - e^{-ax}$, $a > 0$, who is faced with a riskless asset with return r_0 and s risky assets with returns $\mathbf{R} \sim N(\mathbf{r}, \mathbf{V})$.

- 1 Show that when he seeks to maximize the expected utility of his final wealth, he will hold the risky assets in the same proportion as the tangency portfolio.
- 2 Show that when $\beta > \alpha r_0$, the more risk averse that he is, the smaller amount of his wealth that he invests in the risky assets.

Solution I

Suppose that the investor's initial wealth is $w > 0$ and that he wishes to minimize $\mathbb{E} e^{-aW}$ where:

$$W = w \left(r_0 + \sum_{j=1}^s x_j (R_j - r_0) \right) = w[r_0(1 - \mathbf{x}^\top \mathbf{e}) + \mathbf{x}^\top \mathbf{R}],$$

where $\mathbf{x} = (x_1, \dots, x_s)^\top$ and $\mathbf{e} = (1, \dots, 1)^\top$ as usual; the proportion of his wealth in the riskless asset is $x_0 = 1 - \mathbf{x}^\top \mathbf{e}$. Note that the linear combination $\mathbf{x}^\top \mathbf{R}$ has the $N(\mathbf{r}^\top \mathbf{x}, \mathbf{x}^\top V \mathbf{x})$ -distribution, then use the expression (A.16) for the moment-generating function of a normally distributed random variable to see that:

$$\mathbb{E} e^{-aW} = \exp \left(-awr_0(1 - \mathbf{x}^\top \mathbf{e}) - aw\mathbf{r}^\top \mathbf{x} + \frac{1}{2}a^2w^2\mathbf{x}^\top V \mathbf{x} \right).$$

It is necessary to minimize the expression:

$$\frac{1}{2}aw\mathbf{x}^\top V \mathbf{x} - \mathbf{x}^\top (\mathbf{r} - r_0\mathbf{e}),$$

for which the minimum occurs when $\mathbf{x} = (1/aw)V^{-1}(\mathbf{r} - r_0\mathbf{e})$, and the conclusion follows from (1.17). The amount of his wealth invested in the risky assets is $(\mathbf{x}^\top \mathbf{e})w = (\beta - \alpha r_0)/a$, which decreases in $a > 0$ when $\beta > \alpha r_0$.

Problem

Consider an investor with $\mathbf{R} = (R_1, R_2, \dots, R_s)^\top$ where R_i s are independent r.v. with R_i having gamma distribution, $E R_i = r_i$ and $\text{var } R_i = \sigma_i^2$. Suppose that he has the utility function $v(x) = 1 - e^{-ax}$, $a > 0$, and he seeks to maximize the expected utility of his final wealth.

- 1 Show that he divides his wealth between the diversified portfolio and a second portfolio which should be identified; determine the amounts that he invests in each.
- 2 If he may invest in a risky asset with return r_0 , show that he will again divide his wealth between these two portfolios and the riskless asset; determine the amounts that he invests in each.
- 3 Give a necessary and sufficient condition, expressed in terms of the parameters r_i , $i = 0, 1, 2, \dots, s$ and σ_i^2 , $i = 1, 2, \dots, s$, that he is long in the risky assets.

Solution I

When R_i has the gamma distribution $\Gamma(\gamma_i, \lambda_i)$ we have that $E R_i = r_i = \gamma_i / \lambda_i$ and $\text{var } R_i = \gamma_i / \lambda_i^2$, from which it follows that $\gamma_i = r_i^2 / \sigma_i^2$ and $\lambda_i = r_i / \sigma_i^2$. For $\phi + \lambda_i > 0$, note that:

$$E(e^{-\phi R_i}) = \int_0^\infty \frac{e^{-\phi x} e^{-\lambda_i x} \lambda_i^{\gamma_i} x^{\gamma_i-1}}{\Gamma(\gamma_i)} dx = \left(\frac{\lambda_i}{\phi + \lambda_i} \right)^{\gamma_i} \int_0^\infty \frac{e^{-\phi x} e^{-\lambda_i x} (\phi + \lambda_i)^{\gamma_i} x^{\gamma_i-1}}{\Gamma(\gamma_i)} dx$$

because the integrand in the latter integral is a probability density function, and so the value of the integral is 1. The investor wishes to solve the constrained optimization problem:

$$\begin{aligned} & \text{maximize} && E(1 - e^{-aw(\mathbf{x}^\top \mathbf{R})}) \\ & \text{subject to} && \mathbf{x}^\top \mathbf{e} = 1, \end{aligned}$$

but this is equivalent to minimizing:

$$E(e^{-aw(\mathbf{x}^\top \mathbf{R})}) = \prod_{i=1}^s E(e^{-aw x_i R_i}) = \prod_{i=1}^s \left(\frac{\lambda_i}{aw x_i + \lambda_i} \right)^{\gamma_i},$$

Solution II

subject to the constraint. Taking logarithms, we need to:

$$\begin{aligned} & \text{maximize} && \sum_{i=1}^s \gamma_i \ln(awx_i + \lambda_i) \\ & \text{subject to} && \sum_{i=1}^s x_i = 1. \end{aligned}$$

Maximizing the Lagrangian:

$$\mathcal{L} = \sum_{i=1}^s \gamma_i \ln(awx_i + \lambda_i) + \theta \left(1 - \sum_{i=1}^s x_i \right)$$

in x_i gives $x_i = (\gamma_i/\theta) - \lambda_i/(aw)$. Substituting back into the constraint shows that the Lagrange multiplier is given as:

$$\theta = \frac{\sum_{j=1}^s \gamma_j}{1 + (aw)^{-1} \sum_{j=1}^s \lambda_j},$$

from which it follows that the optimal portfolio may be expressed as:

Solution III

$$\mathbf{x} = \left(1 + (aw)^{-1} \sum_{j=1}^s \lambda_j\right) \bar{\mathbf{x}} - \left((aw)^{-1} \sum_{j=1}^s \lambda_j\right) \mathbf{x}_d,$$

where the two portfolios $\bar{\mathbf{x}}$ and \mathbf{x}_d are:

$$(\bar{\mathbf{x}})_i = \frac{\gamma_i}{\sum_j \gamma_j} = \frac{r_i^2/\sigma_i^2}{\sum_j r_j^2/\sigma_j^2} \text{ and } (\mathbf{x}_d)_i = \frac{\lambda_i}{\sum_j \lambda_j} = \frac{r_i/\sigma_i^2}{\sum_j r_j/\sigma_j^2},$$

with the latter portfolio being the diversified portfolio (see Example 1.1). As his initial wealth is w , the investor invests the amount $w + \sum_j \lambda_j/a$ in $\bar{\mathbf{x}}$ and the amount $-\sum_j \lambda_j/a$ in the diversified portfolio.

Note that in the case when the random variables R_i have exponential distributions, then $\gamma_i = 1$, or $r_i^2 = \sigma_i^2$, for each $1 \leq i \leq s$, so that the portfolio $\bar{\mathbf{x}}$ is just the uniform portfolio $\bar{\mathbf{x}} = (1/s, \dots, 1/s)^\top$ which apportions wealth equally between the s risky assets.

For the final part, when there is a riskless asset and we set $x_0 = 1 - \mathbf{x}^\top \mathbf{e}$, we see that we wish to minimize the expression:

Solution IV

$$\mathbb{E} \left(e^{-aw(r_0(1-\mathbf{x}^\top \mathbf{e}) + \mathbf{x}^\top \mathbf{R})} \right) = e^{awr_0(\sum_j x_j - 1)} \prod_{i=1}^s \mathbb{E} \left(e^{-awx_i R_i} \right) = e^{awr_0(\sum_j x_j - 1)} \prod_{i=1}^s \left(\frac{1}{awx_i} \right)$$

in $\mathbf{x} = (x_1, \dots, x_s)^\top$, which is equivalent to maximizing:

$$\sum_{i=1}^s \gamma_i \ln(awx_i + \lambda_i) - awr_0 \sum_{i=1}^s x_i.$$

Deduce that for $1 \leq i \leq s$, the optimal $x_i = (aw)^{-1}((\gamma_i/r_0) - \lambda_i)$, and the optimal investment in the risky assets is determined by:

$$\mathbf{x} = \left((awr_0)^{-1} \sum_{j=1}^s \gamma_j \right) \bar{\mathbf{x}} - \left((aw)^{-1} \sum_{j=1}^s \lambda_j \right) \mathbf{x}_d.$$

The investor is long in the particular risky asset i when $x_i > 0$, which is true if and only if $r_i > r_0$; he is long overall in risky assets if and only if $\sum_{j=1}^s x_j > 0$ which is equivalent to the condition that $r_0^{-1} > \sum_{j=1}^s (r_j/\sigma_j^2) / \sum_{j=1}^s (r_j^2/\sigma_j^2)$.