Portfolio Choice

Utility

Preferences and Utility

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- The relations $\succ$ and $\sim$ satisfy rational axioms as follows.

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7. (Continuity) For any $\mathcal{A}, \mathcal{B}, \mathcal{C} \in \mathbb{P}$, if $\mathcal{A} \succ \mathcal{C} \succ \mathcal{B}$ then there exists $p \in [0,1]$ with $p\,\mathcal{A} + (1-p)\,\mathcal{B} \sim \mathcal{C}$.

Uniqueness of Probability Values

Lemma

Suppose that $\mathcal{A}, \mathcal{B}, \mathcal{C} \in \mathbb{P}$ *with* $\mathcal{A} \succ \mathcal{C} \succ \mathcal{B}$ *and* $p\,\mathcal{A} + (1-p)\,\mathcal{B} \sim \mathcal{C}$*, then* $0 < p < 1$ *and* p *is unique.*

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- Trivially $p \neq 0$ or 1

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- Note that $\mathcal{B} = \left(\frac{p-q}{1-q}\right)\mathcal{B} + \left(\frac{1-p}{1-q}\right)\mathcal{B}$ and $\mathcal{A} \succ \mathcal{B}$

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Suppose that $\mathcal{A}, \mathcal{B}, \mathcal{C} \in \mathbb{P}$ *with* $\mathcal{A} \succ \mathcal{C} \succ \mathcal{B}$ *and* $p\,\mathcal{A} + (1-p)\,\mathcal{B} \sim \mathcal{C}$*, then* $0 < p < 1$ *and* p *is unique.*

Proof

- Trivially $p \neq 0$ or 1
- Suppose p is not unique: $\exists\, q$ with $q\,\mathcal{A} + (1-q)\,\mathcal{B} \sim \mathcal{C}$
- WLOG assume $q < p$, so $0 < p - q < 1 - q$
- Note that $\mathcal{B} = \left(\frac{p-q}{1-q}\right)\mathcal{B} + \left(\frac{1-p}{1-q}\right)\mathcal{B}$ and $\mathcal{A} \succ \mathcal{B}$
- By Independence – Preference Axiom $\left(\dfrac{p-q}{1-q}\right)\mathcal{A} + \left(\dfrac{1-p}{1-q}\right)\mathcal{B} \succ \mathcal{B}$

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- By Independence – Preference Axiom again $p\,\mathcal{A} + (1-p)\,\mathcal{B} \succ q\,\mathcal{A} + (1-q)\,\mathcal{B}$
- But this contradicts that both expressions are indifferent to $\mathcal{C}$

Existence of Utility Function

Theorem

There exists a real-valued function $f : \mathbb{P} \to \mathbb{R}$ *with*

$$f(\mathcal{A}) > f(\mathcal{B}) \quad \textit{if and only if} \quad \mathcal{A} \succ \mathcal{B},$$

and

$$f(p\,\mathcal{A} + (1-p)\,\mathcal{B}) = pf(\mathcal{A}) + (1-p)\,f(\mathcal{B})$$

for any $\mathcal{A}, \mathcal{B} \in \mathbb{P}$ *and* $0 \leqslant p \leqslant 1$. *Furthermore,* f *is unique up to affine transformations.*

Proof of Utility Function Existence (Part 1)

Proof

- If $\mathcal{A} \sim \mathcal{B}$ for all $\mathcal{A}, \mathcal{B} \in \mathbb{P}$, take $f(\mathcal{A}) \equiv 0$

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- For case (a): $\exists$ unique $p \in (0, 1)$ with $p\,\mathcal{A} + (1-p)\,\mathcal{D} \sim \mathcal{C}$; define $f(\mathcal{A}) = \dfrac{1}{p}$

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- For case (c): $\exists$ unique $q \in (0, 1)$ with $q\,\mathcal{C} + (1-q)\,\mathcal{D} \sim \mathcal{A}$; define $f(\mathcal{A}) = q$

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- For case (d): set $f(\mathcal{A}) = 0$
- For case (e): $\exists$ unique $r \in (0,1)$ with $r\mathcal{C} + (1-r)\mathcal{A} \sim \mathcal{D}$; define $f(\mathcal{A}) = \dfrac{-r}{1-r}$

Proof of Utility Function Existence (Part 2)

Proof (continued)

- To verify f satisfies the conditions requires checking 15 cases: 5 where both $\mathcal{A}$ and $\mathcal{B}$ are in the same case, and 10 where they're in different cases

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- We have $f(\mathcal{A}) = q_1$ and $f(\mathcal{B}) = q_2$ where $\mathcal{A} \sim q_1\,\mathcal{C} + (1 - q_1)\,\mathcal{D}$ and $\mathcal{B} \sim q_2\,\mathcal{C} + (1 - q_2)\,\mathcal{D}$

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- When $q_1 = q_2$: $\mathcal{A} \sim \mathcal{B}$ and condition is satisfied
- When $q_1 > q_2$: $q_1\,\mathcal{C} + (1 - q_1)\,\mathcal{D} \succ q_2\,\mathcal{C} + (1 - q_2)\,\mathcal{D}$, so $\mathcal{A} \succ \mathcal{B}$ and $f(\mathcal{A}) > f(\mathcal{B})$ as required

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- Similarly when $q_1 < q_2$: $\mathcal{B} \succ \mathcal{A}$
- For linearity, let $p \in (0, 1)$ and apply Independence – Indifference Axiom:

$$\begin{aligned} p\,\mathcal{A} + (1-p)\,\mathcal{B} &\sim p(q_1\,\mathcal{C} + (1-q_1)\,\mathcal{D}) + (1-p)(q_2\,\mathcal{C} + (1-q_2)\,\mathcal{D}) \\ &\sim (pq_1 + (1-p)q_2)\mathcal{C} + (p(1-q_1) + (1-p)(1-q_2))\mathcal{D} \end{aligned}$$

Proof of Utility Function Existence (Part 3)

Proof (continued)

- From definition of f:
  $f(p\,\mathcal{A} + (1-p)\,\mathcal{B}) = p\,q_1 + (1-p)\,q_2 = p\,f(\mathcal{A}) + (1-p)\,f(\mathcal{B})$

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- To verify f is unique up to affine transformations, suppose g also satisfies the conditions

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- To verify f is unique up to affine transformations, suppose g also satisfies the conditions
- Since $\mathcal{C} \succ \mathcal{D}$, we have $g(\mathcal{C}) > g(\mathcal{D})$

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- To verify f is unique up to affine transformations, suppose g also satisfies the conditions
- Since $\mathcal{C} \succ \mathcal{D}$, we have $g(\mathcal{C}) > g(\mathcal{D})$
- Define $\beta = g(\mathcal{D})$ and $\alpha = g(\mathcal{C}) - g(\mathcal{D}) > 0$

Proof of Utility Function Existence (Part 3)

Proof (continued)

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- For an $\mathcal{A}$ in case (c) with $f(\mathcal{A}) = q$, we have $\mathcal{A} \sim q\,\mathcal{C} + (1-q)\,\mathcal{D}$

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- For an $\mathcal{A}$ in case (c) with $f(\mathcal{A}) = q$, we have $\mathcal{A} \sim q\,\mathcal{C} + (1-q)\,\mathcal{D}$
- Therefore $g(\mathcal{A}) = g(q\,\mathcal{C} + (1-q)\,\mathcal{D}) = qg(\mathcal{C}) + (1-q)g(\mathcal{D}) =$
  $q(\alpha+\beta) + (1-q)\beta = q\alpha + \beta = \alpha f(\mathcal{A}) + \beta$

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Proof (continued)

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  d. *Logarithmic*: $v(x) = \ln x$. Follows from HARA as $\gamma \to 0$, using l'Hôpital's rule: $\dfrac{x^{\gamma} - 1}{\gamma} \to \ln x$

Mean-Variance Analysis

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- $\mathbf{R} \equiv (R_1, R_2, \dots, R_s)^\top$: the random vector representing the rate of return on the assets; $R_i = \dfrac{S_{i,1}}{S_{i,0}}$

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- $\mathbf{S}_0 \equiv (S_{1,0}, S_{2,0}, \dots, S_{s,0})^\top \neq \mathbf{0}$: the constant price vector at time 0
- $\mathbf{S}_1 \equiv (S_{1,1}, S_{2,1}, \dots, S_{s,1})^\top$: the random price vector at time 1
- $\mathbf{x} \equiv (x_1, x_2, \dots, x_s)^\top$: the proportion vector of the time-0 wealth invested in each asset; $\displaystyle\sum_{i=1}^{s} x_i = 1$.
- $\mathbf{R} \equiv (R_1, R_2, \dots, R_s)^\top$: the random vector representing the rate of return on the assets; $R_i = \dfrac{S_{i,1}}{S_{i,0}}$
- w: the (constant) wealth at time 0

Classical PO: Mean-Variance (MV) Criterion

- W: the (random) wealth at time 1; $W = \left(\sum_{i=1}^{s} x_i R_i\right) w = \mathbf{x}^\top \mathbf{R}\, w$

  (For asset S_i, $\dfrac{x_i w}{S_{i,0}}$ denotes the "quantity" allocated at time 0; so at time 1 this part of wealth becomes $\dfrac{x_i w}{S_{i,0}}\, S_{i,1} = x_i R_i w$)

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- "For some fixed mean rate of return $\mu = \mathsf{E}\{\mathbf{x}^\top \mathbf{R}\}$, try to minimize the variance $\sigma^2 = \operatorname{var}\{\mathbf{x}^\top \mathbf{R}\}$ of the return over portfolios $\mathbf{x}$"

MV: All Risky Assets

$$\min_{\mathbf{x}} \ \frac{1}{2}\,\mathbf{x}^\top \mathbf{V}\mathbf{x} \quad \text{s.t.} \quad \begin{cases} \mathbf{x}^\top \mathbf{e} = 1 \\ \mathbf{x}^\top \mathbf{r} = \mu \end{cases} \qquad \mathbf{e} \equiv \underbrace{(1, 1, \dots, 1)^\top}_{s \text{ items}}$$

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- Set $\mathcal{L} \equiv \frac{1}{2}\mathbf{x}^\top \mathbf{V}\mathbf{x} + \lambda\left(1 - \mathbf{x}^\top \mathbf{e}\right) + \nu\left(\mu - \mathbf{x}^\top \mathbf{r}\right)$ with Lagrange multipliers λ, ν

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- By $\dfrac{\partial \mathcal{L}}{\partial \mathbf{x}} = \mathbf{V}\mathbf{x} - \lambda\,\mathbf{e} - \nu\,\mathbf{r} = 0 \implies \mathbf{x} = \lambda\,\mathbf{V}^{-1}\mathbf{e} + \nu\,\mathbf{V}^{-1}\mathbf{r}$
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- Substitute into $\begin{cases} \mathbf{x}^\top \mathbf{e} = 1 \\ \mathbf{x}^\top \mathbf{r} = \mu \end{cases} \implies \begin{cases} \lambda\,\mathbf{e}^\top \mathbf{V}^{-1}\mathbf{e} + \nu\,\mathbf{r}^\top \mathbf{V}^{-1}\mathbf{e} = 1 \\ \lambda\,\mathbf{e}^\top \mathbf{V}^{-1}\mathbf{r} + \nu\,\mathbf{r}^\top \mathbf{V}^{-1}\mathbf{r} = \mu \end{cases}$

- Set $\alpha = \mathbf{e}^\top \mathbf{V}^{-1}\mathbf{e}$, $\beta = \mathbf{r}^\top \mathbf{V}^{-1}\mathbf{e} = \mathbf{e}^\top \mathbf{V}^{-1}\mathbf{r}$, $\gamma = \mathbf{r}^\top \mathbf{V}^{-1}\mathbf{r}$, $\delta \equiv \alpha\gamma - \beta^2$, then

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$$\begin{aligned} &(\mathbf{r} - c\,\mathbf{e})^\top \mathbf{V}^{-1}(\mathbf{r} - c\,\mathbf{e}) > 0 \\ &\implies \mathbf{r}^\top \mathbf{V}^{-1}\mathbf{r} - c\,\mathbf{r}^\top \mathbf{V}^{-1}\mathbf{e} - c\,\mathbf{e}\mathbf{V}^{-1}\mathbf{r} + c^2\,\mathbf{e}^\top \mathbf{V}^{-1}\mathbf{e}^\top > 0 \\ &\implies \gamma - 2\,c\,\beta + c^2\,\alpha > 0 \\ &\implies -\delta = \beta^2 - \gamma\alpha < 0 \end{aligned}$$

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$$= \lambda + \nu\mu = \frac{\gamma - \beta\mu}{\delta} + \nu\frac{\alpha\mu - \beta}{\delta} = \frac{\alpha\mu^2 - 2\beta\mu + \gamma}{\delta}$$

$$\implies \frac{\sigma^2}{\left(\frac{1}{\sqrt{\alpha}}\right)^2} - \frac{\left(\mu - \frac{\beta}{\alpha}\right)^2}{\left(\frac{\sqrt{\delta}}{\alpha}\right)^2} = 1$$

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\text{equation:} \quad & \frac{(x-h)^2}{a^2} - \frac{(y-k)^2}{b^2} = 1 \\
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- Here we have (σ, μ) with $a = \dfrac{1}{\sqrt{\alpha}}$, $b = \dfrac{\sqrt{\delta}}{\alpha}$, $h = 0$, $k = \dfrac{\beta}{\alpha}$, the asymptotes are $\left(\mu - \dfrac{\beta}{\alpha}\right) = \pm\dfrac{\frac{\sqrt{\delta}}{\alpha}}{\frac{1}{\sqrt{\alpha}}}\sigma \implies \mu = \dfrac{\beta}{\alpha} \pm \sqrt{\dfrac{\delta}{\alpha}}\sigma$

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- $\mathbf{x} = \lambda \mathbf{V}^{-1}\mathbf{e} + \nu \mathbf{V}^{-1}\mathbf{r} = \lambda\,\alpha\,\mathbf{x}_g + \nu\,\beta\,\mathbf{x}_d$, so **every portfolio is the convex combination of $\mathbf{x}_g$ and $\mathbf{x}_d$**: note that $\lambda\alpha + \nu\beta = 1$ (constraint $\mathbf{x}^\top\mathbf{e} = 1$) !

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Theorem (Mutual Fund Theorem)

Any minimum-variance portfolio is equivalent to investing in the convex combination of $\mathbf{x}_g$ and $\mathbf{x}_d$.

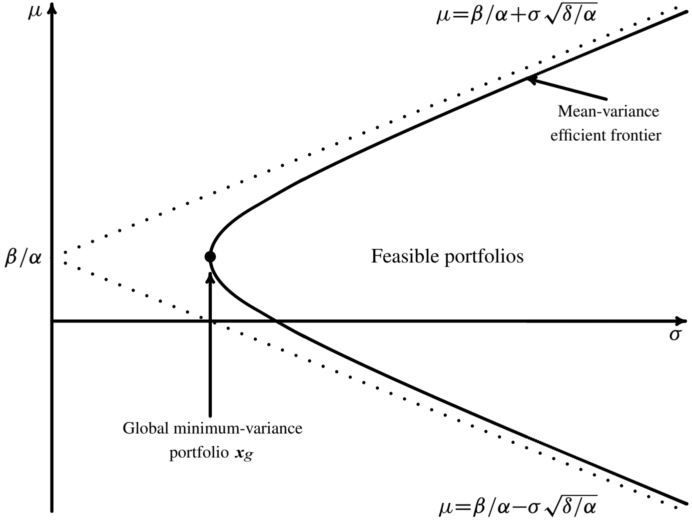

Figure: The Case of All Risky Assets

Theorem

Diversified portfolio $\mathbf{x}_d$ is the portfolio that maximize $s(\mathbf{x}) \equiv \dfrac{\mathbf{x}^\top \mathbf{r}}{\sqrt{\mathbf{x}^\top \mathbf{V} \mathbf{x}}}$.

Proof

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- The covariance between the return of the global mininum-variance portfolio and other minimum-variance portfolio is constant:
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  $= \dfrac{\lambda}{\alpha}\mathbf{e}^\top \mathbf{V}^{-1}\mathbf{e} + \dfrac{\nu}{\alpha}\mathbf{e}^\top \mathbf{V}^{-1}\mathbf{r} = \dfrac{\lambda\,\alpha + \nu\,\beta}{\alpha} = \dfrac{1}{\alpha}$

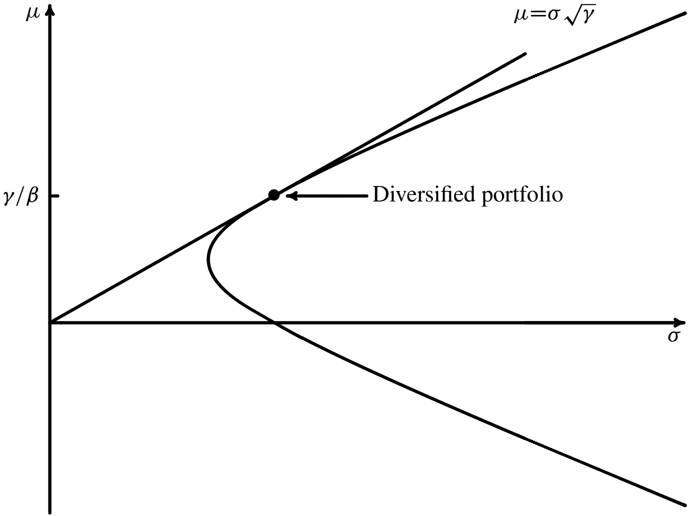

Figure: The Diversified Portfolio

MV: All But One Risky Assets

WLOG add riskless asset 0 with constant return r_0; the portfolio becomes $(x_0, x_1, x_2, \ldots, x_s)^\top$

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- Set $\overline{\mathcal{L}} \equiv \frac{1}{2}\mathbf{x}^\top \mathbf{V}\mathbf{x} + \overline{\lambda}\left(1 - x_0 - \mathbf{x}^\top \mathbf{e}\right) + \overline{\nu}\left(\mu - x_0 r_0 \mathbf{x}^\top \mathbf{r}\right)$ with Lagrange multipliers $\overline{\lambda}$, $\overline{\nu}$

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- By $\frac{\partial \overline{\mathcal{L}}}{\partial x_0} = -\overline{\lambda} - \overline{\nu} r_0 = 0 \implies \overline{\nu} = -\frac{\overline{\lambda}}{r_0}$

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- Set $\alpha = \mathbf{e}^\top \mathbf{V}^{-1} \mathbf{e},\ \ \beta = \mathbf{r}^\top \mathbf{V}^{-1} \mathbf{e} = \mathbf{e}^\top \mathbf{V}^{-1} \mathbf{r},\ \ \gamma = \mathbf{r}^\top \mathbf{V}^{-1} \mathbf{r},\ \ \delta \equiv \alpha\gamma - \beta^2$, the above becomes

$$\begin{cases} x_0 + \overline{\lambda}\alpha + \overline{\nu}\beta = x_0 + \overline{\lambda}\alpha - \frac{\overline{\lambda}}{r_0}\beta = 1 \\ x_0 r_0 + \overline{\lambda}\beta + \overline{\nu}\gamma = x_0 r_0 + \overline{\lambda}\beta - \frac{\overline{\lambda}}{r_0}\gamma = \mu \end{cases}$$

with solutions $x_0 = \dfrac{\alpha\mu r_0 - \beta r_0 + \gamma - \beta\mu}{\epsilon^2}$, $\overline{\lambda} = \dfrac{(r_0 - \mu) r_0}{\epsilon^2}$, $\overline{\nu} = -\dfrac{r_0 - \mu}{\epsilon^2}$, where $\epsilon^2 = \alpha r_0^2 - 2\beta r_0 + \gamma = \alpha\Big(r_0 - \dfrac{\beta}{\alpha}\Big)^2 + \dfrac{\delta}{\alpha}$

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- The relation of σ with μ

$$\begin{aligned} \sigma^2 = \mathbf{x}^\top \mathbf{V}\mathbf{x} &= \mathbf{x}^\top \mathbf{V}(\overline{\lambda}\mathbf{V}^{-1}\mathbf{e} + \overline{\nu}\mathbf{V}^{-1}\mathbf{r}) = \overline{\lambda}(\mathbf{x}^\top \mathbf{e}) + \overline{\nu}(\mathbf{x}^\top \mathbf{r}) \\ &= \overline{\lambda}(1 - x_0) + \overline{\nu}(\mu - x_0 r_0) = \overline{\lambda} + \overline{\nu}\mu = \frac{(\mu - r_0)^2}{\epsilon^2} \end{aligned}$$

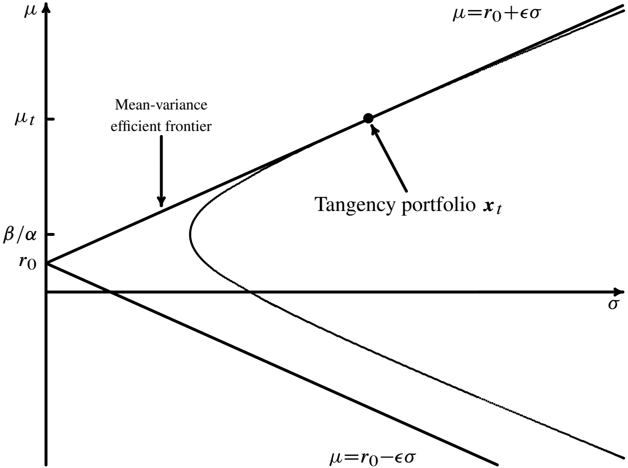

Figure: The Case of All But One Risky Assets

Property

If $r_0 < \frac{\beta}{\alpha}$, then $\mu = r_0 + \epsilon\sigma$ touches the hyperbola $\sigma^2 = \frac{\alpha\mu^2 - 2\beta\mu + \gamma}{\delta}$ at $\left(\frac{\epsilon}{\beta - \alpha r_0}, \frac{\gamma - \beta r_0}{\beta - \alpha r_0}\right)$

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Proof

On $\sigma - \mu$ plane the slope of the tangent is obtained by implicit differentiation of $\sigma^2 = \frac{\alpha\mu^2 - 2\beta\mu + \gamma}{\delta}$ w.r.t σ (let $\mu \equiv \mu(\sigma)$): $2\sigma = \frac{2\alpha\mu\mu' - 2\beta\mu'}{\delta} \implies \mu' = \frac{\delta\sigma}{\alpha\mu - \beta}$. The tangent line is $\mu = r_0 + \epsilon\sigma$ with slope ϵ, so $\epsilon = \frac{\delta\sigma}{\alpha\mu - \beta} \implies \delta\sigma = \alpha\mu\epsilon - \beta\epsilon \implies \delta\sigma = \alpha\epsilon(r_0 + \epsilon\sigma) - \beta\epsilon \implies (\delta - \alpha\epsilon^2)\sigma = \epsilon(\alpha r_0 - \beta)$. Note that $\epsilon^2 = \alpha r_0^2 - 2\beta r_0 + \gamma = \alpha\left(r_0 - \frac{\beta}{\alpha}\right)^2 + \frac{\delta}{\alpha}$, so $\sigma = \frac{\epsilon(\alpha r_0 - \beta)}{\delta - \alpha\epsilon^2} = \frac{\epsilon(\alpha r_0 - \beta)}{-\alpha^2\left(r_0 - \frac{\beta}{\alpha}\right)^2} = \frac{\epsilon}{\beta - \alpha r_0}$, $\mu = r_0 + \epsilon\frac{\epsilon}{\beta - \alpha r_0} = \frac{\beta r_0 - \alpha r_0^2 + \epsilon^2}{\beta - \alpha r_0} = \frac{\gamma - \beta r_0}{\beta - \alpha r_0}$.

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$$\mathbf{x}_t = \frac{1}{\beta - \alpha r_0}\mathbf{V}^{-1}(\mathbf{r} - r_0\mathbf{e}) = \frac{\beta}{\beta - \alpha r_0}\mathbf{x}_d - \frac{\alpha r_0}{\beta - \alpha r_0}\mathbf{x}_g$$

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- $\mathbf{x} = \overline{\lambda}\mathbf{V}^{-1}\mathbf{e} + \overline{\nu}\mathbf{V}^{-1}\mathbf{r} = \overline{\nu}\mathbf{V}^{-1}(\mathbf{r} - r_0\mathbf{e}) \equiv (1 - x_0)\mathbf{x}_t$

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- $\mu_t = \mathbf{x}_t^\top\mathbf{r} = \mathbf{r}^\top\mathbf{x}_t = \frac{\beta}{\beta - \alpha r_0}\mathbf{r}^\top\mathbf{x}_d - \frac{\alpha r_0}{\beta - \alpha r_0}\mathbf{r}^\top\mathbf{x}_g$
  $= \frac{\beta}{\beta - \alpha r_0}\mu_d - \frac{\alpha r_0}{\beta - \alpha r_0}\mu_g = \frac{\gamma - \beta r_0}{\beta - \alpha r_0}$ for $\mu_d = \frac{\gamma}{\beta}$, $\mu_g = \frac{\beta}{\alpha}$

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Mean-Variance Pricing Equation

- $\mathbf{V} = \mathsf{E}\left\{(\mathbf{R}-\mathbf{r})(\mathbf{R}-\mathbf{r})^\top\right\} = \mathsf{E}\left\{\mathbf{R}\,\mathbf{R}^\top - \mathbf{R}\,\mathbf{r}^\top - \mathbf{r}\,\mathbf{R}^\top + \mathbf{r}\,\mathbf{r}^\top\right\} = \mathsf{E}\left\{\mathbf{R}\,\mathbf{R}^\top - \mathbf{R}\,\mathbf{r}^\top\right\}$

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- $(\mathbf{V}\mathbf{x}_t)_i = \dfrac{1}{\beta - \alpha r_0}(r_i - r_0);$

Mean-Variance Pricing Equation

- $\mathbf{V} = \mathsf{E}\left\{(\mathbf{R}-\mathbf{r})(\mathbf{R}-\mathbf{r})^\top\right\} = \mathsf{E}\left\{\mathbf{R}\,\mathbf{R}^\top - \mathbf{R}\,\mathbf{r}^\top - \mathbf{r}\,\mathbf{R}^\top + \mathbf{r}\,\mathbf{r}^\top\right\} = \mathsf{E}\left\{\mathbf{R}\,\mathbf{R}^\top - \mathbf{R}\,\mathbf{r}^\top\right\}$
- $\operatorname{cov}(R_i, \mathbf{x}_t^\top \mathbf{R}) = \mathsf{E}\left\{(R_i - r_i)(\mathbf{x}_t^\top \mathbf{R} - \mathbf{x}_t^\top \mathbf{r})\right\} = \mathsf{E}\left\{R_i\,\mathbf{x}_t^\top \mathbf{R} - R_i\,\mathbf{x}_t^\top \mathbf{r} - r_i\,\mathbf{x}_t^\top \mathbf{R} + r_i\,\mathbf{x}_t^\top \mathbf{r}\right\} = \mathsf{E}\left\{R_i\,\mathbf{x}_t^\top \mathbf{R} - R_i\,\mathbf{x}_t^\top \mathbf{r}\right\} = \mathsf{E}\left\{R_i\,\mathbf{R}^\top \mathbf{x}_t - R_i\,\mathbf{r}^\top \mathbf{x}_t\right\}$
- $\mathbf{V}\mathbf{x}_t = \mathsf{E}\left\{\mathbf{R}\,\mathbf{R}^\top \mathbf{x}_t - \mathbf{R}\,\mathbf{r}^\top \mathbf{x}_t\right\}$
- $(\mathbf{V}\mathbf{x}_t)_i = \dfrac{1}{\beta - \alpha r_0}(r_i - r_0);$
- $\operatorname{var}(\mathbf{x}_t^\top \mathbf{R}) = \mathsf{E}\{\mathbf{x}_t^\top \mathbf{R}\cdot(\mathbf{x}_t^\top \mathbf{R})^\top\} - (\mathsf{E}\{\mathbf{x}_t^\top \mathbf{R}\})^2 = \mathsf{E}\left\{\mathbf{x}_t^\top \mathbf{R}\,\mathbf{R}^\top \mathbf{x}_t\right\} - \mathsf{E}\left\{\mathbf{x}_t^\top \mathbf{R}\right\}\mathsf{E}\left\{\mathbf{R}^\top \mathbf{x}_t\right\} = \mathsf{E}\left\{\mathbf{x}_t^\top \mathbf{R}\,\mathbf{R}^\top \mathbf{x}_t\right\} - \mathbf{x}_t^\top \mathbf{r}\,\mathbf{r}^\top \mathbf{x}_t = \mathbf{x}_t^\top\,\mathsf{E}\left\{\mathbf{R}\,\mathbf{R}^\top - \mathbf{r}\,\mathbf{r}^\top\right\}\mathbf{x}_t = \mathbf{x}_t^\top V \mathbf{x}_t = \dfrac{\mu_t - r_0}{\beta - \alpha r_0}.$

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- $\beta_{i,t} = \dfrac{\operatorname{cov}(R_i, \mathbf{x}^\top \mathbf{R})}{\operatorname{var}(\mathbf{x}^\top \mathbf{R})} = \operatorname{cor}(R_i, \mathbf{x}^\top \mathbf{R})\sqrt{\dfrac{\operatorname{var} R_i}{\operatorname{var}(\mathbf{x}^\top \mathbf{R})}}$; define $\boldsymbol{\beta}_t \equiv (\beta_{1,t}, \beta_{2,t}, \ldots, \beta_{s,t})^\top$

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- $\boldsymbol{\beta}_t = \dfrac{1}{\mu_t - r_0}(\mathbf{r} - r_0\mathbf{e}) \implies \mathbf{r} = r_0\mathbf{e} + (\mu_t - r_0)\boldsymbol{\beta}_t$

Mean-Variance Analysis and Expected Utility

- Define $f(\sigma, \mu) = \mathsf{E}\, v(W)$ where $W = (x_0 r_0 + \mathbf{x}^\top \mathbf{R})w$, $\sigma^2 = \mathbf{x}^\top \mathbf{V}\mathbf{x}$, $\mu = x_0 r_0 + \mathbf{x}^\top \mathbf{r} = \mathbf{x}^\top(\mathbf{r} - r_0\mathbf{e})$

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- Assume that $\dfrac{\partial f}{\partial \sigma} < 0$, $\dfrac{\partial f}{\partial \mu} > 0$ with $x_0 + \mathbf{x}^\top \mathbf{e} = 1$, perform $\max\limits_{\mathbf{x}} f\left(\sqrt{\mathbf{x}^\top \mathbf{V}\mathbf{x}}, r_0 + \mathbf{x}^\top(\mathbf{r} - r_0\mathbf{e})\right)$

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- Example:
  - For quadratic utility $v(x) = ax + bx^2$ where $a, b \in \mathbb{R}$, $b \leqslant 0$: $\mathsf{E}\, v(W) = \mathsf{E}\, v((x_0 r_0 + \mathbf{x}^\top \mathbf{R})w) = aw\mu + bw^2(\mu^2 + \sigma^2) = f(\sigma, \mu)$

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  - For normally distributed returns $\mathbf{R} \sim N(\mathbf{r}, \mathbf{V})$, $\mathbf{x}^\top \mathbf{R} \sim N(\mathbf{x}^\top \mathbf{r}, \mathbf{x}^\top V \mathbf{x})$: $\mathsf{E}\, v(W) = \mathsf{E}\, v((\mu + \sigma Y)w)$, where $Y \sim N(0, 1)$

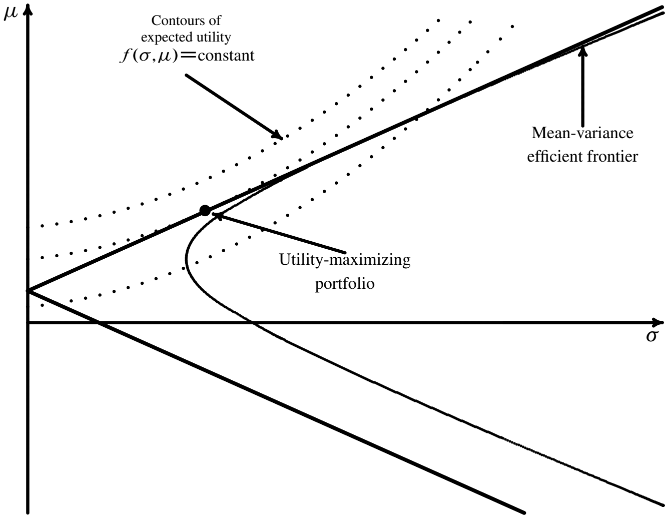

Figure: Determining the Utility-Maximizing Portfolio

Equilibrium: The Capital-Asset Pricing Model

- Investors indexed by $j \in \mathcal{J}$, each with proportions of wealth $x_{0,j}$ and $\mathbf{x}_j = (x_{1,j}, x_{2,j}, \ldots, x_{s,j})^\top$

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- The total value of the demand for risky asset i: $$\sum_{j \in \mathcal{J}} w_j x_{i,j} = \Big(\sum_{j \in \mathcal{J}} (1 - x_{0,j}) w_j \Big) (\mathbf{x}_t)_i$$

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$$\frac{\Big(\sum_{j\in\mathcal{J}}(1-x_{0,j})w_j\Big)(\mathbf{x}_t)_i}{\Big(\sum_{j\in\mathcal{J}}(1-x_{0,j})w_j\Big)\sum_{k=1}^{s}(\mathbf{x}_t)_k} = (\mathbf{x}_t)_i, \text{ since } \mathbf{x}_t^\top \mathbf{e} = 1$$

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- $\mathbf{r} = r_0\mathbf{e} + (\mu_m - r_0)\boldsymbol{\beta}_m$, $\boldsymbol{\beta}_m \equiv (\beta_{1,m}, \beta_{2,m}, \dots, \beta_{s,m})^\top$, $$\beta_{i,m} = \frac{\operatorname{cov}(R_i, \mathbf{x}_m^\top \mathbf{R})}{\operatorname{var}(\mathbf{x}_m^\top \mathbf{R})}$$ — capital-asset-pricing equation

Problems and Solutions

Problem 1

Suppose that an investment X has either (i) the uniform distribution $U[0, 2\mu]$ or (ii) the exponential distribution with $\mathsf{E}\, X = \mu$, and the investor has a utility function which is either (a) logarithmic, $v(x) = \log x$ (b) power form, $v(x) = x^\theta$. Show that both the compensatory risk premium and the investment risk premium are proportional to μ in all 4 possible cases.

Problem 1 Solution

- For distributions (i)(ii) of X, the r.v. $Y \equiv \dfrac{X}{\mu}$ does not depend on μ, so $\mathsf{E}\, v(X+\alpha) = v(\mu)$ for the compensatory risk premium α reduces to $\mathsf{E}\, v(Y+c) = v(1)$ in cases (a)(b) when $\alpha = c\mu$. For the insurance risk premium when $\beta = d\mu$, d is the solution of $\mathsf{E}\, v(Y) = v(1-d)$.

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- For case (i)(a), $\mathsf{E}\, v(Y+c) = \displaystyle\int_0^2 \frac{\log(y+c)}{2}\, \mathrm{d}y = \frac{1}{2}\big((2+c)\log(2+c) - c\log c - 2\big)$, and $v(1) = \log 1 = 0$, so $\alpha = c\mu$ where c is the unique positive root of $(2+c)\log(2+c) - c\log c - 2 = 0$. Using `rmaxima`

```
find_root((2 + x) * log(2 + x) - x * log(x) - 2, x, 0.01, 20);
```

we have $c = 0.176965531$. For the insurance premium $\beta = d\mu$, $\mathsf{E}\log Y = \log 2 - 1 = \log(1-d)$, so $d = 1 - \dfrac{2}{e} = 0.264$.

Problem 2

An investor has a utility function $v(x) = \sqrt{x}$ and is considering three investments with random outcomes X, Y, Z. Here X has the uniform distribution $U[0, a]$, Y has the gamma distribution $\Gamma(\gamma, \lambda)$ with probability density function $\dfrac{e^{-\lambda y}\lambda^{\gamma} y^{\gamma-1}}{\Gamma(\gamma)}$ for $y > 0$, where $\gamma > 0$, $\lambda > 0$, and Z is log-normal, i.e $Z \sim N(\nu, \sigma^2)$. The parameter of the distributions are such that $\mathsf{E}\, X = \mathsf{E}\, Y = \mathsf{E}\, Z = \mu$ and $\operatorname{var} X = \operatorname{var} Y = \operatorname{var} Z$. Recall that the gamma function $\Gamma(\gamma) = \displaystyle\int_0^{\infty} u^{\gamma-1} e^{-u}\, \mathsf{d}u$ that satisfies $\Gamma(\gamma + 1) = \gamma\, \Gamma(\gamma)$ and $\Gamma(1/2) = \sqrt{\pi}$. Determine the investor's preference ordering of X, Y, Z for all values of μ.

Problem 2 Solution I

- $X \sim U[0,a] \implies \mathsf{E}\,X = \dfrac{a}{2}$, $\operatorname{var} X = \dfrac{a^2}{12}$
- $Y \sim \Gamma(\gamma,\lambda) \implies \mathsf{E}\,Y = \dfrac{\gamma}{\lambda}$, $\operatorname{var} Y = \dfrac{\gamma}{\lambda^2}$
- $Z \sim \text{lognormal}(\nu,\sigma^2) \implies \mathsf{E}\,Z = e^{\nu+\frac{\sigma^2}{2}}$, $\operatorname{var} Z = e^{2\nu+\sigma^2}(e^{\sigma^2}-1)$ by the formula $\mathsf{E}\,e^{\theta W} = e^{\mu\theta+\frac{\theta^2\sigma^2}{2}}$ for $W \sim N(\mu,\sigma^2)$

$\mathsf{E}\,e^{\theta W} = e^{\mu\theta+\frac{\theta^2\sigma^2}{2}}$ as $W \sim N(\mu,\sigma^2)$: $\sqrt{2\pi}\sigma\;\mathsf{E}\,e^{\theta W} = \int_{-\infty}^{\infty} e^{\theta x}\cdot e^{-\frac{1}{2}\frac{x^2-2\mu+\mu^2}{\sigma^2}}\,\mathrm{d}x$

$$= \int_{-\infty}^{\infty} e^{-\frac{1}{2}\frac{x^2-2(\mu+\theta\sigma^2)+\mu^2}{\sigma^2}}\,\mathrm{d}x = \int_{-\infty}^{\infty} e^{-\frac{1}{2}\frac{(x-(\mu+\theta\sigma^2))^2+\mu^2-(\mu+\theta\sigma^2)^2}{\sigma^2}}\,\mathrm{d}x$$

$$= \int_{-\infty}^{\infty} e^{-\frac{1}{2}\frac{(x-(\mu+\theta\sigma^2))^2-2\mu\theta\sigma^2-(\theta\sigma^2)^2}{\sigma^2}}\,\mathrm{d}x = e^{\mu\theta+\frac{\theta^2\sigma^2}{2}}\cdot\int_{-\infty}^{\infty} e^{-\frac{1}{2}\frac{(x-(\mu+\theta\sigma^2))^2}{\sigma^2}}\,\mathrm{d}x$$

$= e^{\mu\theta+\frac{\theta^2\sigma^2}{2}}\cdot\int_{-\infty}^{\infty} e^{-\frac{x^2}{2\sigma^2}}\,\mathrm{d}x = \sqrt{2\pi}\sigma\cdot e^{\mu\theta+\frac{\theta^2\sigma^2}{2}}$ by $\int_{-\infty}^{\infty} e^{-x^2}\,\mathrm{d}x = \sqrt{\pi}$.

The conditions $\mathsf{E}\,X = \mathsf{E}\,Y = \mathsf{E}\,Z = \mu$ and $\operatorname{var} X = \operatorname{var} Y = \operatorname{var} Z$ imply

- $a = 2\mu$, so that $\operatorname{var} X = \dfrac{\mu^2}{3}$.

Problem 2 Solution II

- $\mathsf{E}\, Y = \frac{\gamma}{\lambda} = \mu$, so that $\operatorname{var} Y = \frac{\gamma}{\lambda^2} = \frac{\mu^2}{\gamma} = \operatorname{var} X = \frac{\mu^2}{3} \implies \gamma = 3$
- $\mathsf{E}\, Z = e^{\nu + \frac{\sigma^2}{2}} = \mu$, $\operatorname{var} Z = e^{2\nu+\sigma^2}(e^{\sigma^2} - 1) = \mu^2(e^{\sigma^2} - 1) = \operatorname{var} X = \frac{\mu^2}{3}$
  $\implies \sigma^2 = \log \frac{4}{3}$.
- $\mathsf{E}\, \sqrt{X} = \int_0^{2\mu} \frac{\sqrt{x}}{2\mu}\, \mathsf{d}x = \frac{2^{\frac{3}{2}}}{3}\sqrt{\mu} \approx 0.943\sqrt{\mu}$
- $\mathsf{E}\, \sqrt{Y} = \int_0^{\infty} \sqrt{y}\, \frac{1}{2} e^{-\lambda y} \lambda^3 y^2\, \mathsf{d}y = \frac{\Gamma(\frac{7}{2})}{2\sqrt{\lambda}} = \frac{\Gamma(\frac{7}{2})}{2\sqrt{3}}\sqrt{\mu} = \frac{5\sqrt{3\pi}}{16}\sqrt{\mu} \approx 0.959\sqrt{\mu}$
- $\mathsf{E}\, \sqrt{Z} = e^{\frac{\nu}{2} + \frac{\sigma^2}{8}} = e^{-\frac{\sigma^2}{8}}\sqrt{\mu} = \left(\frac{3}{4}\right)^{\frac{1}{8}}\sqrt{\mu} \approx 0.965\sqrt{\mu}$

So $Z \succ Y \succ X$.

Problem 3

Suppose that an investor has the utility function $v(x) = 1 - e^{-ax}$ with $a > 0$, and the outcome of an investment is a r.v. X with mean μ, finite variance and finite moment-generating function $\psi(a) = \mathsf{E}\left\{e^{-aX}\right\}$ for $a > 0$. Show that

1. The compensatory risk premium and the insurance risk premium have the same value α, and express α in terms of μ and the moment generating function ψ.
2. Both the Arrow-Pratt and global risk aversions are a.
3. As $a \downarrow 0$, $\alpha = \dfrac{a}{2} \operatorname{var} X + \mathcal{O}(a)$. Under what circumstances is it true that $\alpha = \dfrac{a}{2} \operatorname{var} X$ for $a > 0$?
4. $\psi'' \psi - (\psi')^2 \geqslant 0$ and hence α is an increasing function of a. This shows that the more risk-averse the investor is, the higher the value of the premium that is required.

Problem 3 Solution I

1. - The compensatory risk premium α solves $\mathsf{E}\{v(\alpha+X)\} = v(\mu) \implies \mathsf{E}\{1-e^{-a(\alpha+X)}\} = 1-e^{-a\mu} \implies 1-e^{-a\alpha}\,\mathsf{E}\{e^{-aX}\} = 1-e^{-a\mu} \implies -a\alpha + \ln\psi(a) = -a\mu \implies \alpha = \mu + \frac{1}{a}\ln(\psi(a))$
   - The insurance risk premium β solves $\mathsf{E}\,v(X) = v(\mu-\beta) \implies \mathsf{E}\{1-e^{-aX}\} = 1-e^{-a(\mu-\beta)} \implies 1-\psi(a) = 1-e^{-a(\mu-\beta)} \implies \ln\psi(a) = -a(\mu-\beta) \implies \beta = \mu + \frac{1}{a}\ln(\psi(a))$

   So $\alpha = \beta = \mu + \frac{1}{a}\ln(\psi(a))$.
2. Note that $v'(x) = ae^{-ax}$, $v''(x) = -a^2e^{-ax}$, the Arrow-Pratt absolute risk aversion is $-\frac{v''(\mu)}{v'(\mu)} = \frac{a^2e^{-a\mu}}{ae^{-a\mu}} = a$, the global absolute risk aversion is $-\frac{\mathsf{E}\{v''(X)\}}{\mathsf{E}\{v'(X)\}} = \frac{a^2\,\mathsf{E}\{e^{-aX}\}}{a\,\mathsf{E}\{e^{-aX}\}} = a$.

Problem 3 Solution II

3. From $\alpha(a) = \mu + \dfrac{1}{a}\ln(\psi(a)) \implies \psi(a) = e^{a(\alpha(a)-\mu)}$; Differentiation yields $\psi'(a) = e^{a(\alpha(a)-\mu)}(\alpha(a) - \mu + a\,\alpha'(a))$, $\psi''(a) = e^{a(\alpha(a)-\mu)}(2\alpha'(a) + a\,\alpha''(a)) + e^{a(\alpha(a)-\mu)}(\alpha(a) - \mu + a\,\alpha'(a))^2$. Note that $\psi(0) = 1$, $\psi'(0) = -\mu$, $\psi''(0) = \mathsf{E}\,X^2$; setting $a = 0$ yields $\psi'(0) = e^{0(\alpha(0)-\mu)}(\alpha(0) - \mu + 0\,\alpha'(0)) \implies -\mu = \alpha(0) - \mu \implies \alpha(0) = 0$, $\psi''(0) = e^{0(\alpha(0)-\mu)}(2\alpha'(0) + 0\,\alpha''(0)) + e^{0(\alpha(0)-\mu)}(\alpha(0) - \mu + 0\,\alpha'(0))^2 \implies \mathsf{E}\,X^2 = 2\alpha'(0) + \mu^2 \implies \alpha'(0) = \dfrac{1}{2}(\mathsf{E}\,X^2 - \mu^2) = \dfrac{1}{2}\operatorname{var} X$. For small $a > 0$, the Taylor expansion of $\alpha(a) = \alpha(0) + a\,\alpha'(0) + \mathcal{O}(a^2) = \dfrac{a}{2}\operatorname{var} X + \mathcal{O}(a^2)$. When $\alpha(a) = \dfrac{a}{2}\operatorname{var} X$ exactly for $a > 0$, then $\psi(a) = \mathsf{E}\left\{e^{-aX}\right\} = e^{-a\mu + \frac{a^2}{2}\operatorname{var} X}$, which is true only when X is normally distributed.
4. $\psi''\psi - (\psi')^2 = \mathsf{E}\left\{X^2 e^{-aX}\right\}\mathsf{E}\left\{e^{-aX}\right\} - (\mathsf{E}\left\{Xe^{-aX}\right\})^2 \geqslant 0$ by the Cauchy-Schwarz inequality applied to r.v.s $A = Xe^{-\frac{a}{2}X}$ and $B = e^{-\frac{a}{2}X}$. To see that α is increasing $\dfrac{\mathrm{d}\alpha}{\mathrm{d}a} = \dfrac{1}{a^2}\left(\dfrac{a\psi'}{\psi} - \ln(\psi)\right) \equiv \dfrac{1}{a^2}f(a)$, but $f(0) = 0$ and $f' = \dfrac{a(\psi''\psi - (\psi')^2)}{\psi^2} \geqslant 0$ and the conclusion follows.

Problem 4

Consider a one-period investment model in which there are only two risky assets. The returns on these assets have means 3, 4 respectively and variances 2, 3 respectively with the covariance between the returns being 2.

1. Calculate the mean-variance efficient frontier and the minimum-variance portfolio in terms of the mean return.
2. Calculate the mean return of the global minimum-variance portfolio and of diversified portfolio.
3. Suppose that in addition to the two risky assets there is a riskless asset with return $3/2$. Find the minimum-variance portfolio in terms of the mean return and hence calculate the mean return of the tangency portfolio.

Problem 4 Solution I

The inverse matrix of $\begin{pmatrix} a & b \\ c & d \end{pmatrix}$ is $\dfrac{1}{ad-bc}\begin{pmatrix} d & -b \\ -c & a \end{pmatrix}$, so if $V = \begin{pmatrix} 2 & 2 \\ 2 & 3 \end{pmatrix}$, $V^{-1} = \dfrac{1}{2}\begin{pmatrix} 3 & -2 \\ -2 & 2 \end{pmatrix}$. $\alpha = \mathbf{e}^\top \mathbf{V}^{-1}\mathbf{e} = \dfrac{1}{2}$, $\beta = \mathbf{e}^\top \mathbf{V}^{-1}\mathbf{r} = \dfrac{3}{2}$, $\gamma = \mathbf{r}^\top \mathbf{V}^{-1}\mathbf{r} = \dfrac{11}{2}$, $\delta = \alpha\gamma - \beta^2 = \dfrac{1}{2}$.

1. $\min\limits_{x_1,x_2} \begin{pmatrix} x_1 & x_2 \end{pmatrix}\begin{pmatrix} 2 & 2 \\ 2 & 3 \end{pmatrix}\begin{pmatrix} x_1 \\ x_2 \end{pmatrix} = \min\limits_{x_1,x_2} 2x_1^2 + 4x_1x_2 + 3x_2^2$ s.t. $\begin{cases} x_1 + x_2 = 1 \\ 3x_1 + 4x_2 = \mu \end{cases}$. From constraints $x_1 = 4 - \mu$, $x_2 = \mu - 3$, so the mean-variance efficient frontier is $\sigma^2 = \mu^2 - 6\mu + 11$.
2. μ_g is the root of $\dfrac{\mathrm{d}\sigma^2}{\mathrm{d}\mu} = 0$, so $2\mu_g - 6 = 0 \implies \mu_g = 3$. $\mu_d = \dfrac{\gamma}{\beta} = \dfrac{11}{3}$.

Problem 4 Solution II

3. Now the problem is $\min_{x_0,x_1,x_2} 2x_1^2 + 4x_1x_2 + 3x_2^2$ s.t. $\begin{cases} x_0 + x_1 + x_2 = 1 \\ \frac{3}{2}x_0 + 3x_1 + 4x_2 = \mu \end{cases}$. Form the Lagrangian $\mathcal{L} = 2x_1^2 + 4x_1x_2 + 3x_2^2 + \lambda(1 - x_0 - x_1 - x_2) + \nu(\mu - \frac{3}{2}x_0 - 3x_1 - 4x_2)$. By solving $\frac{\partial \mathcal{L}}{\partial x_0} = 0$, $\nu = -\frac{2\lambda}{3}$. From $\frac{\partial \mathcal{L}}{\partial x_1} = 0$ and $\frac{\partial \mathcal{L}}{\partial x_2} = 0$ we have $4x_1 + 4x_2 - \lambda - 3\nu = 0$ and $4x_1 + 6x_2 - \lambda - 4\nu = 0$; so $x_1 = \frac{\lambda}{12}$, $x_2 = -\frac{\lambda}{3}$. Substitute into the constraints yields $\lambda = \frac{12(3-2\mu)}{17}$, and so $x_0 = \frac{26-6\mu}{17}$, $x_1 = \frac{3-2\mu}{17}$, $x_2 = -\frac{4(3-2\mu)}{17}$. The tangency portfolio corresponds to $x_0 = 0$ or $\mu_t = \frac{13}{3}$.

Problem 5

Suppose that v is concave, $X \sim N(\mu, \sigma^2)$ and $f(\sigma, \mu) = \mathsf{E}\, v(X)$.

1. Show that $\dfrac{\partial f}{\partial \mu} > 0$ when v is strictly increasing, and $\dfrac{\partial f}{\partial \sigma} \leqslant 0$. Hence show in the context of mean-variance analysis that, when all returns are jointly normally distributed, an investor maximizing the expected utility of his final wealth will choose a mean-variance-efficient optimal portfolio.
2. Show that f is concave in μ and σ. Deduce that this optimal portfolio corresponds to a point in the (σ, μ) plane where an indifference contour is tangent to the efficient frontier.

Problem 5 Solution I

Write $X = \mu + \sigma Y$ where $Y \sim N(0,1)$.

1. $$\frac{\partial f}{\partial \mu} = \mathsf{E}\{v'(\mu + \sigma Y)\} > 0 \text{ when } v' > 0$$
   $$\frac{\partial f}{\partial \sigma} = \mathsf{E}\{Y v'(\mu + \sigma Y)\} = \sigma\, \mathsf{E}\{v''(\mu + \sigma Y)\} \leqslant 0$$
   by the concavity of v ($v'' < 0$). When returns are normally distributed, the wealth created by each portfolio is normally distributed; this shows that maximizing in σ for fixed μ gives a value of $(\sigma,\, \mu)$ on the efficient frontier.
2. To see the concavity of f, note that
   $$\frac{\partial^2 f}{\partial \mu^2} = \mathsf{E}\{v''(\mu + \sigma Y)\} \leqslant 0$$
   $$\frac{\partial^2 f}{\partial \sigma^2} = \mathsf{E}\left\{Y^2 v''(\mu + \sigma Y)\right\} \leqslant 0$$
   $$\frac{\partial^2 f}{\partial \mu \partial \sigma} = \mathsf{E}\{Y v''(\mu + \sigma Y)\}$$

Problem 5 Solution II

and then

$$\frac{\partial^2 f}{\partial \mu^2}\frac{\partial^2 f}{\partial \sigma^2} \geqslant \left(\frac{\partial^2 f}{\partial \mu \partial \sigma}\right)^2$$

follows by applying the Cauchy-Schwarz inequality to the r.v.s $A = Y\sqrt{-v''(\mu + \sigma Y)}$ and $B = \sqrt{-v''(\mu + \sigma Y)}$; this shows that the 2×2 matrix of second derivatives has non-positive diagonal entries and a non-negative determinant which is sufficient for the matrix to be negative semi-definite. The fact that f is concave means that sets of the form $\{(\sigma, \mu) : f(\sigma, \mu) \geqslant c\}$ are convex which gives the last statement.

Problem 6

Suppose that an investor has a concave utility function v. The investor seeks to maximize $\mathsf{E}\, v(W)$ where $W = (x_0 r_0 + \mathbf{x}^\top \mathbf{R})\, w$ is his final wealth.

1. Show that, when $\overline{W}$ is his optimal final wealth, then $\mathsf{E}\left\{v'(\overline{W})(R_j - r_0)\right\} = 0, \ \forall\, j = 1, 2, \dots, s.$
2. Show that, when $\mathbf{R}$ has a multivariate normal distribution, then $r_j - r_0 = \alpha \operatorname{cov}(\overline{W}, R_j), \ \forall\, j = 1, 2, \dots, s$, where $\alpha = -\dfrac{\mathsf{E}\left\{v''(\overline{W})\right\}}{\mathsf{E}\left\{v'(\overline{W})\right\}}$ is his global risk aversion.
3. Now suppose that the market is determined by investors $i = 1, 2, \dots, n$, where investor i has concave utility v_i, initial wealth w_i, optimal final wealth $\overline{W}_i$ and global risk aversion α_i. With the normality assumption, show that $\mathsf{E}\, M - r_0 = \overline{w}\,\overline{\alpha} \operatorname{var} M$, where $M = \dfrac{\sum_{i=1}^n \overline{W}_i}{\sum_{i=1}^n w_i}$ is the market rate of return, $\overline{w} = \dfrac{\sum_{i=1}^n w_i}{n}$ is the average initial wealth of investors, and $\overline{\alpha}$ is the harmonic mean of α_i.

Problem 6 Solution I

1. The objective function to maximize is

$$f(\mathbf{x}) = \mathsf{E}\, v\left(w\left(r_0 + \sum_{j=1}^{s} x_j(R_j - r_0)\right)\right)$$

where $\mathbf{x} = (x_1, \ldots, x_s)^\top$ and we have used the condition that $x_0 + \sum_{j=1}^{s} x_j = 1$. The first-order conditions give

$$\frac{\partial f}{\partial x_j} = w\,\mathsf{E}\, v'(\overline{W})(R_j - r_0) = 0, \text{ for } 1 \leqslant j \leqslant s$$

Since $r_j = \mathsf{E}\, R_j$ and the fact that $\overline{W}$ and R_j have a joint normal distribution we have that

$$\begin{aligned}
0 &= \mathsf{E}\left\{v'(\overline{W})(R_j - r_0)\right\} = \mathsf{E}\left\{v'(\overline{W})(R_j - r_j)\right\} + \mathsf{E}\left\{v'(\overline{W})\right\}(r_j - r_0) \\
&= \operatorname{cov}(v'(\overline{W}), R_j) + \mathsf{E}\left\{v'(\overline{W})\right\}(r_j - r_0) \\
&= \mathsf{E}\left\{v''(\overline{W})\right\}\operatorname{cov}(\overline{W}, R_j) + \mathsf{E}\left\{v'(\overline{W})\right\}(r_j - r_0)
\end{aligned}$$

Problem 6 Solution II

where the last equality uses, and this gives

$$r_j - r_0 = \alpha \operatorname{cov}(\overline{W}, R_j)$$

Note that for r.v.s X and Y and constant a, $\operatorname{cov}(X, Y + a) = \operatorname{cov}(X, Y)$ and $\operatorname{cov}(aX, Y) = a \operatorname{cov}(X, Y)$. Now for each i

$$\frac{1}{\alpha_i}(r_j - r_0) = \operatorname{cov}(\overline{W}_i, R_j)$$

and summing this on i yields

$$\left(\sum_{i=1}^{n} \frac{1}{\alpha_i}\right)(r_j - r_0) = \left(\sum_{i=1}^{n} w_i\right) \operatorname{cov}(M, R_j)$$

Divide through by n and multiply by $\overline{\alpha}$, where $\dfrac{1}{\overline{\alpha}} = \dfrac{\sum_{i=1}^{n} \frac{1}{\alpha_i}}{n}$, to obtain

$$\mathsf{E}\, R_j - r_0 = w\, \overline{\alpha} \operatorname{cov}(M, R_j) \tag{1}$$

Problem 6 Solution III

When $\overline{x}_{ij}$ is the optimal proportion invested by investor i in asset j then

$$\overline{W_i} = w_i \left(r_0 + \sum_{j=1}^{s} \overline{x}_{ij}(R_j - r_0) \right)$$

which when summed on i gives

$$(M - r_0) \left(\sum_{i=1}^{n} w_i \right) = \sum_{i=1}^{n} \sum_{j=1}^{s} w_i \overline{x}_{ij}(R_j - r_0) \tag{2}$$

2. Take the expectation in (2), multiply (1) by $w_i \overline{x}_{ij}$, sum on i and j, rearrange the expression using the two properties of covariance mentioned above and $\mathsf{E}\, M - r_0 = \overline{w}\, \overline{\alpha} \operatorname{var} M$ follows. This shows that the risk premium for the market is proportional to $\overline{\alpha}$ which is a measure of the risk aversion in the economy.

Problem 7

Consider an investor with the utility function $v(x) = 1 - e^{-ax}$, $a > 0$, who is faced with a riskless asset with return r_0 and s risky assets with returns $\mathbf{R} \sim N(\mathbf{r}, \mathbf{V})$.

1. Show that when he seeks to maximize the expected utility of his final wealth, he will hold the risky assets in the same proportion as the tangency portfolio.
2. Show that when $\beta > \alpha\, r_0$, the more risk averse that he is, the smaller amount of his wealth that he invests in the risky assets.

Problem 7 Solution

1. Suppose that the investor's initial wealth is $w > 0$ and that he wishes to minimize $\mathsf{E}\, e^{-aW}$ where

$$W = w\left(r_0 + \sum_{j=1}^{s} x_j(R_j - r_0)\right) = w\left(r_0(1 - \mathbf{x}^\top \mathbf{e}) + \mathbf{x}^\top \mathbf{R}\right)$$

and $\mathbf{x} = (x_1, \dots, x_s)^\top$, $\mathbf{e} = (1, \dots, 1)^\top$, $x_0 = 1 - \mathbf{x}^\top \mathbf{e}$. Note that $\mathbf{x}^\top \mathbf{R} \sim N(\mathbf{r}^\top \mathbf{x}, \mathbf{x}^\top \mathbf{V} \mathbf{x})$, so

$$\mathsf{E}\, e^{-aW} = \exp\left\{-aw\, r_0(1 - \mathbf{x}^\top \mathbf{e}) - aw\, \mathbf{r}^\top \mathbf{x} + \frac{1}{2} a^2 w^2 \mathbf{x}^\top \mathbf{V} \mathbf{x}\right\}$$

It amounts to minimize $\frac{1}{2} aw\, \mathbf{x}^\top \mathbf{V} \mathbf{x} - \mathbf{x}^\top (\mathbf{r} - r_0 \mathbf{e})$ for which the minimum occurs when $\mathbf{x} = \frac{1}{aw} \mathbf{V}^{-1}(\mathbf{r} - r_0 \mathbf{e})$, and the conclusion follows.

2. The amount of his wealth invested in the risky assets is $(\mathbf{x}^\top \mathbf{e})w = \dfrac{\beta - \alpha r_0}{a}$, which decreases in $a > 0$ when $\beta > \alpha r_0$.

Problem 8

Consider an investor with $\mathbf{R} = (R_1, R_2, \dots, R_s)^\top$ where R_is are independent r.v. with R_i having gamma distribution, $\mathsf{E}\, R_i = r_i$ and $\operatorname{var} R_i = \sigma_i^2$. Suppose that he has the utility function $v(x) = 1 - e^{-ax}$, $a > 0$, and he seeks to maximize the expected utility of his final wealth.

1. Show that he divides his wealth between the diversified portfolio and a second portfolio which should be identified; determine the amounts that he invests in each.
2. If he may invest in a risky asset with return r_0, show that he will again divide his wealth between these two portfolios and the riskless asset; determine the amounts that he invests in each.
3. Give a necessary and sufficient condition, expressed in terms of the parameters r_i, $i = 0, 1, 2, \dots, s$ and σ_i^2, $i = 1, 2, \dots, s$, that he is long in the risky assets.

Problem 8 Solution I

1. When $R_i \sim \Gamma(\gamma_i, \lambda_i)$, $\operatorname{E} R_i = r_i = \dfrac{\gamma_i}{\lambda_i}$ and $\operatorname{var} R_i = \sigma_i^2 = \dfrac{\gamma_i}{\lambda_i^2} \implies \gamma_i = \dfrac{r_i^2}{\sigma_i^2}$, $\lambda_i = \dfrac{r_i}{\sigma_i^2}$. For $\phi + \lambda_i > 0$, note that

$$
\begin{aligned}
\operatorname{E} e^{-\phi R_i} &= \int_0^\infty \frac{e^{-\phi x} e^{-\lambda_i x} \lambda_i^{\gamma_i} x^{\gamma_i - 1}}{\Gamma(\gamma_i)} \, \mathrm{d}x \\
&= \left(\frac{\lambda_i}{\phi + \lambda_i} \right)^{\gamma_i} \int_0^\infty \frac{e^{-\phi x} e^{-\lambda_i x} (\phi + \lambda_i)^{\gamma_i} x^{\gamma_i - 1}}{\Gamma(\gamma_i)} \, \mathrm{d}x = \left(\frac{\lambda_i}{\phi + \lambda_i} \right)^{\gamma_i}
\end{aligned}
$$

because the integrand in the latter integral is a probability density function, and so the value of the integral is 1. The investor wishes to solve the constrained optimization problem

$$
\text{maximize} \quad \operatorname{E} \left\{ 1 - e^{-aw(\mathbf{x}^\top \mathbf{R})} \right\} \quad \text{subject to} \quad \mathbf{x}^\top \mathbf{e} = 1
$$

Problem 8 Solution II

which is equivalent to minimizing

$$\mathsf{E}\left\{e^{-aw(\mathbf{x}^\top \mathbf{R})}\right\} = \prod_{i=1}^{s} \mathsf{E}\left\{e^{-awx_i R_i}\right\} = \prod_{i=1}^{s} \left(\frac{\lambda_i}{awx_i + \lambda_i}\right)^{\gamma_i}$$

subject to the constraint. Taking logarithms, we need to

$$\text{maximize} \quad \sum_{i=1}^{s} \gamma_i \ln(awx_i + \lambda_i) \quad \text{subject to} \quad \sum_{i=1}^{s} x_i = 1$$

Maximizing the Lagrangian

$$\mathcal{L} = \sum_{i=1}^{s} \gamma_i \ln(awx_i + \lambda_i) + \theta \left(1 - \sum_{i=1}^{s} x_i\right)$$

Problem 8 Solution III

in x_i gives $x_i = \frac{\gamma_i}{\theta} - \frac{\lambda_i}{aw}$. Substituting back into the constraint shows that the Lagrange multiplier is given as

$$\theta = \frac{\sum_{j=1}^{s} \gamma_j}{1 + \frac{1}{aw}\sum_{j=1}^{s} \lambda_j}$$

from which it follows that the optimal portfolio may be expressed as

$$\mathbf{x} = \left(1 + \frac{1}{aw}\sum_{j=1}^{s} \lambda_j\right)\overline{\mathbf{x}} - \left(\frac{1}{aw}\sum_{j=1}^{s} \lambda_j\right)\mathbf{x}_d$$

where the two portfolios $\overline{\mathbf{x}}$ and $\mathbf{x}_d$ are

$$(\overline{\mathbf{x}})_i = \frac{\gamma_i}{\sum_j \gamma_j} = \frac{\frac{r_i^2}{\sigma_i^2}}{\sum_j \frac{r_j^2}{\sigma_j^2}} \text{ and } (\mathbf{x}_d)_i = \frac{\lambda_i}{\sum_j \lambda_j} = \frac{\frac{r_i}{\sigma_i^2}}{\sum_j \frac{r_j}{\sigma_j^2}}$$

Problem 8 Solution IV

with the latter portfolio being the diversified portfolio. As his initial wealth is w, the investor invests the amount $w + \dfrac{1}{a}\sum_j \lambda_j$ in $\overline{\mathbf{x}}$ and the amount $-\dfrac{1}{a}\sum_j \lambda_j$ in the diversified portfolio. Note that in the case when the r.v.s R_i have exponential distributions, then $\gamma_i = 1$, or $r_i^2 = \sigma_i^2$, for each $1 \leqslant i \leqslant s$, so that the portfolio $\overline{\mathbf{x}}$ is just the uniform portfolio $\overline{\mathbf{x}} = \left(\dfrac{1}{s}, \ldots, \dfrac{1}{s}\right)^\top$ which apportions wealth equally between the s risky assets.

2. When there is a riskless asset, set $x_0 = 1 - \mathbf{x}^\top \mathbf{e}$ and we wish to minimize

$$\begin{aligned}
\mathsf{E}\left\{e^{-aw(r_0(1-\mathbf{x}^\top\mathbf{e})+\mathbf{x}^\top\mathbf{R})}\right\} &= e^{awr_0(\sum_j x_j - 1)} \prod_{i=1}^{s} \mathsf{E}\left\{e^{-awx_iR_i}\right\} \\
&= e^{awr_0(\sum_j x_j - 1)} \prod_{i=1}^{s} \left(\frac{\lambda_i}{awx_i + \lambda_i}\right)^{\gamma_i}
\end{aligned}$$

Problem 8 Solution V

in $\mathbf{x} = (x_1, \ldots, x_s)^\top$, which is equivalent to maximizing

$$\sum_{i=1}^{s} \gamma_i \ln(awx_i + \lambda_i) - awr_0 \sum_{i=1}^{s} x_i$$

Deduce that for $1 \leqslant i \leqslant s$, the optimal $x_i = \frac{1}{aw}\left(\frac{\gamma_i}{r_0} - \lambda_i\right)$, and the optimal investment in the risky assets is determined by

$$\mathbf{x} = \left(\frac{1}{awr_0} \sum_{j=1}^{s} \gamma_j\right) \overline{\mathbf{x}} - \left(\frac{1}{aw} \sum_{j=1}^{s} \lambda_j\right) \mathbf{x}_d$$

3. The investor is long in the particular risky asset i when $x_i > 0$, which is true if and only if $r_i > r_0$; he is long overall in risky assets if and only if $\sum_{j=1}^{s} x_j > 0$ which is equivalent to the condition that $\frac{1}{r_0} > \frac{\sum_{j=1}^{s} \frac{r_j}{\sigma_j^2}}{\sum_{j=1}^{s} \frac{r_j^2}{\sigma_j^2}}$.