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$$\ell = \sum_{i=1}^N \left[d_{iC} \log(P_{iC}) + \sum_{j \in J_B} d_{ij} \log(P_{iB} P_{ij|B}) \right]$$

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Expanding the likelihood function from the previous slide:

$$\ell = \sum_{i=1}^N \left[d_{iC} \log(P_{iC}) + (d_{iBB} + d_{iRB}) \log(P_{iB}) + \sum_{j \in J_B} d_{ij} \log(P_{ij|B}) \right]$$

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Key insight: The likelihood decomposes into separate components that can be estimated sequentially

For sequential estimation, decompose utility as:

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We estimate parameters $(\beta_B, \beta_{RB}, \gamma, \lambda)$ where γ corresponds to alternative-specific variables (Z 's)

The normalizations imply:

$$u_{iC} = 0$$

$$u_{iBB} = \beta_B X_i + \gamma(Z_{BB} - Z_C)$$

$$u_{iRB} = (\beta_B + \beta_{RB})X_i + \gamma(Z_{RB} - Z_C)$$

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Stage 1: Estimate β_{RB} and γ using only observations that chose bus (N_B):

$$\ell_1 = \sum_{i=1}^{N_B} \left\{ d_{iRB} \left[\frac{u_{iRB|B}}{\lambda} \right] - \log \left(1 + \exp \left(\frac{u_{iRB|B}}{\lambda} \right) \right) \right\}$$

Define the inclusive value l_{iB} :

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This allows us to rewrite the nest probability term as:

$$\begin{aligned} \left[\exp \left(\frac{u_{iBB}}{\lambda} \right) + \exp \left(\frac{u_{iRB}}{\lambda} \right) \right]^\lambda &= \exp \left(\frac{u_{iBB}}{\lambda} \right) \left[1 + \exp \left(\frac{u_{iRB|B}}{\lambda} \right) \right]^\lambda \\ &= \exp \left(\frac{u_{iBB}}{\lambda} + \lambda l_{iB} \right) \end{aligned}$$

Define the inclusive value I_{iB} :

$$I_{iB} = \log \left(1 + \exp \left(\frac{U_{iRB|B}}{\lambda} \right) \right)$$

This allows us to rewrite the nest probability term as:

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Note: λI_{iB} represents the expected utility associated with the bus nest

Stage 2: Estimate β_B and λ using the inclusive value from Stage 1:

$$\ell_2 = \sum_i \left\{ d_{iB} \left[\frac{u_{iBB}}{\lambda} + \lambda I_{iB} \right] - \log \left(1 + \exp \left(\frac{u_{iBB}}{\lambda} + \lambda I_{iB} \right) \right) \right\}$$

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This sequential estimation works because the log likelihood is additively separable:

$$\log(P_{iB}(\beta_B, \beta_{RB}, \gamma, \lambda)) + \log(P_{iRB|B}(\beta_{RB}, \gamma))$$

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Parameters β_{RB} and γ are identified using only within-nest variation

Then we treat them as given when estimating between-nest parameters β_B and λ

Instead of sequential estimation, can estimate all parameters simultaneously (FIML):

$$\ell(\beta_B, \beta_{RB}, \gamma, \lambda) = \sum_{i=1}^N \sum_{j=1}^J d_{ij} \log(P_{ij})$$

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Where the choice probabilities are:

$$P_{ij} = \Pr(\text{nest}) \times \Pr(j|\text{nest})$$

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e.g.

$$P_{i,\text{RB}} = \underbrace{\frac{[\exp(u_{i,\text{RB}}/\lambda) + \exp(u_{i,\text{BB}}/\lambda)]^\lambda}{[\exp(u_{i,\text{RB}}/\lambda) + \exp(u_{i,\text{BB}}/\lambda)]^\lambda + \exp(u_{i,\text{C}})}}_{\Pr(\text{Bus})} \times \underbrace{\frac{\exp(u_{i,\text{RB}}/\lambda)}{\exp(u_{i,\text{RB}}/\lambda) + \exp(u_{i,\text{BB}}/\lambda)}}_{\Pr(\text{RB}|\text{Bus})}$$

Advantages of FIML:

- Asymptotically efficient (achieves Cramér-Rao lower bound)
- Provides correct standard errors in one step
- Allows for joint hypothesis testing across stages

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Disadvantages of FIML:

- Computationally more demanding
- Less stable convergence (especially for λ near boundary)
- Requires good starting values

Further considerations:

- Log-likelihood function is not globally concave

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- Log-likelihood function is not globally concave
- Common parameters across submodels estimated separately in sequential approach
- FIML automatically enforces equality constraints
- Must satisfy $0 < \lambda_k \leq 1$ for model to be consistent with utility maximization