

### Ising model for multivariate binary response

Some results from Joe (1997), Multivariate models and dependence concepts.

$Y_1, \dots, Y_d$  dependent binary random variables in  $\{0, 1\}$ .

Exponential family model (analogous to multivariate normal) is

$$f_{\mathbf{Y}}(\mathbf{y}) = \zeta^{-1} \exp \left\{ \sum_j \alpha_j y_j + \sum_{1 \leq j < k \leq d} \gamma_{jk} y_j y_k \right\}, \quad y_j \in \{0, 1\} \forall j,$$

where  $\zeta$  is a normalizing constant that depends on the  $\alpha$  and  $\gamma$  parameters. This implies conditional distributions that are logistic in one variable given the remainder.

$$\begin{aligned} \frac{\mathbb{P}(Y_j = 1 | \mathbf{Y}_{-j} = \mathbf{y}_{-j})}{\mathbb{P}(Y_j = 0 | \mathbf{Y}_{-j} = \mathbf{y}_{-j})} &= \frac{\exp \{ \alpha_j + \sum_{k \neq j} [\alpha_k y_k + \gamma_{jk} y_k] + \sum_{k_1 < k_2, k_1 \neq j, k_2 \neq j} \gamma_{k_1 k_2} y_{k_1} y_{k_2} \}}{\exp \{ \sum_{k \neq j} \alpha_k y_k + \sum_{k_1 < k_2, k_1 \neq j, k_2 \neq j} \gamma_{k_1 k_2} y_{k_1} y_{k_2} \}} \\ &= \exp \left\{ \alpha_k + \sum_{k \neq j} \gamma_{jk} y_k \right\} \end{aligned}$$

or

$$\text{logit } \mathbb{P}(Y_j = 1 | \mathbf{Y}_{-j} = \mathbf{y}_{-j}) = \alpha_k + \sum_{k \neq j} \gamma_{jk} y_k.$$

The converse result [Joe and Liu, 1996, SAPL] is that if one has logistic regressions of any response variable on the remaining  $y$ 's, then these logistic regressions if the coefficient  $\gamma_{jk}$  of  $y_j$  regressing on  $y_k$  matches the the coefficient  $\gamma_{kj}$  of  $y_k$  regressing on  $y_j$ . That is,  $\gamma_{jk} = \gamma_{kj}$  for all  $j \neq k$ .

If  $\gamma_{jk} = 0$ , then it implies some conditional independence properties that are similar to multivariate Gaussian (it does not imply anything about conditional correlation).

Ising:  $\gamma_{jk} = 0$  implies

$$\text{logit } \mathbb{P}(Y_j = 1 | \mathbf{Y}_{-j} = \mathbf{y}_{-j}) = \text{logit } \mathbb{P}(Y_j = 1 | Y_\ell = y_\ell, \ell \neq j, k),$$

and

$$\text{logit } \mathbb{P}(Y_k = 1 | \mathbf{Y}_{-k} = \mathbf{y}_{-k}) = \text{logit } \mathbb{P}(Y_k = 1 | Y_\ell = y_\ell, \ell \neq j, k).$$

That is, if  $\gamma_{12} = 0$ , then conditional distribution of  $Y_1$  given  $Y_2, \dots, Y_d$  is the same as the conditional distribution of  $Y_1$  given  $Y_3, \dots, Y_d$ , etc.

This is analogous to a property of multivariate Gaussian: if  $\Sigma^{-1} = (\sigma^{jk})$ , then  $\sigma^{jk} = 0$  for  $j \neq k$  implies

$$[Y_j | \mathbf{Y}_{-j} = \mathbf{y}_{-j}] = [Y_j | Y_\ell = y_\ell, \ell \neq j, k], \quad (2)$$

$$[Y_k | \mathbf{Y}_{-k} = \mathbf{y}_{-k}] = [Y_k | Y_\ell = y_\ell, \ell \neq j, k].$$

That is, there is a simplification in some Bayesian network representation.

For multivariate Gaussian, this results follows from: if  $Y_j$  is regressed on  $Y_\ell$ ,  $\ell \neq j$ , then the sign of the regression coefficient  $\beta_k$  for  $Y_k$  is the same as the sign of  $\rho_{jk;rest}$ . Hence  $\beta_k = 0$  implies  $\rho_{jk;rest} = \sigma^{jk} = 0$  and result (1).

---

MTP<sub>2</sub> properties (multivariate totally positive of order 2)

Multivariate density  $f_{1:d}$  is MTP<sub>2</sub> if

$$f_{1:d}(\mathbf{y} \vee \mathbf{z}) f_{1:d}(\mathbf{y} \wedge \mathbf{z}) \geq f_{1:d}(\mathbf{y}) f_{1:d}(\mathbf{z}) \quad \forall \mathbf{y}, \mathbf{z}, \quad (1)$$

where  $\vee$  and  $\wedge$  denote coordinatewise max and min.

Multivariate normal density  $f_{1:d}(\mathbf{y}) \propto \exp\{-Q(\mathbf{y})\}$  is MTP<sub>2</sub> if all of the coefficients of cross terms  $\sigma^{jk}$  in the quadratic form  $Q$  are negative or if  $\rho_{jk;rest} \geq 0$  for all  $j, k$ .

Similarly, for the Ising exponential family multivariate binary density,  $f_{1:d}(\mathbf{y})$  is MTP<sub>2</sub> if all  $\gamma_{jk} \geq 0$ .

If the inequality in (1) is reversed, this leads to MRR<sub>2</sub> (multivariate reverse rule of order 2), a weak concept of negative dependence.

Mathematical result: MTP<sub>2</sub> density implies all marginal densities of dimensions 2 or more are MTP<sub>2</sub>. There is no analogous result for MRR<sub>2</sub>.

Hence, multivariate normal density that is MTP<sub>2</sub> has correlations  $\rho_{jk} \geq 0$  for all  $j \neq k$ .

Ising density that is MTP<sub>2</sub> has bivariate margins that are positive dependent.

---

trivariate Ising model :

$\alpha_1$	$\alpha_2$	$\alpha_3$	$\gamma_{12}$	$\gamma_{13}$	$\gamma_{23}$	$\mu_1$	$\mu_2$	$\mu_3$	$\rho_{12}$	$\rho_{13}$	$\rho_{23}$	OR12	OR13	OR23
0.5	0.4	0.6	0.2	0.3	0	0.696	0.631	0.691	0.0448	0.0650	0.0029	1.221	1.350	1.013
0.5	0.4	0.6	0.2	0.3	0.2	0.699	0.664	0.718	0.0466	0.0653	0.0459	1.237	1.362	1.237
0.5	0.4	0.6	0.2	0.3	0.4	0.702	0.697	0.746	0.0479	0.0650	0.0855	1.251	1.374	1.511
0.5	0.4	0.6	0.2	0.3	0.8	0.708	0.760	0.799	0.0485	0.0629	0.1535	1.276	1.395	2.255
0.5	0.4	0.6	0.2	0.3	2.0	0.722	0.903	0.918	0.0387	0.0469	0.2750	1.321	1.434	7.486
0.5	0.4	0.6	0.2	0.3	5.0	0.730	0.994	0.995	0.0104	0.0123	0.3359	1.344	1.455	150.4
0.5	0.4	0.6	0.2	0.3	-0.2	0.693	0.600	0.664	0.0424	0.0642	-0.0429	1.205	1.337	0.829
0.5	0.4	0.6	0.2	0.3	-1.0	0.683	0.491	0.573	0.0297	0.0581	-0.2392	1.136	1.285	0.373
0.5	0.4	0.6	0.2	0.3	-5.0	0.669	0.349	0.454	0.0001	0.0415	-0.6455	1.001	1.195	0.007
0.5	0.4	0.6	0.2	-0.3	0	0.609	0.627	0.602	0.0474	-0.0710	-0.0034	1.221	0.741	0.986
0.5	0.4	0.6	0.2	-0.3	0.2	0.608	0.656	0.633	0.0434	-0.0677	0.0428	1.205	0.748	1.204
0.5	0.4	0.6	0.2	-0.3	0.4	0.607	0.686	0.664	0.0395	-0.0643	0.0862	1.189	0.754	1.471
0.5	0.4	0.6	0.2	-0.3	0.8	0.605	0.745	0.728	0.0322	-0.0573	0.1624	1.162	0.765	2.194
0.5	0.4	0.6	0.2	-0.3	2.0	0.602	0.891	0.883	0.0165	-0.0368	0.3044	1.113	0.787	7.285
0.5	0.4	0.6	0.2	-0.3	5.0	0.599	0.993	0.993	0.0034	-0.0090	0.3787	1.088	0.798	146.3
0.5	0.4	0.6	0.2	-0.3	-0.2	0.609	0.600	0.573	0.0515	-0.0742	-0.0517	1.239	0.734	0.807
0.5	0.4	0.6	0.2	-0.3	-1.0	0.612	0.511	0.478	0.0665	-0.0848	-0.2480	1.314	0.705	0.363
0.5	0.4	0.6	-0.2	-0.3	0	0.550	0.572	0.606	-0.0491	-0.0725	0.0036	0.819	0.741	1.015
0.5	0.4	0.6	-0.2	-0.3	0.2	0.547	0.602	0.635	-0.0519	-0.0737	0.0508	0.807	0.734	1.240
0.5	0.4	0.6	-0.2	-0.3	0.4	0.543	0.634	0.664	-0.0542	-0.0744	0.0956	0.797	0.727	1.514
0.5	0.4	0.6	-0.2	-0.3	0.8	0.535	0.700	0.724	-0.0568	-0.0740	0.1752	0.779	0.715	2.259
0.5	0.4	0.6	-0.2	-0.3	2.0	0.516	0.867	0.878	-0.0493	-0.0593	0.3269	0.746	0.693	7.499
0.5	0.4	0.6	-0.2	-0.3	5.0	0.501	0.992	0.992	-0.0141	-0.0164	0.4079	0.729	0.681	150.6
0.5	0.4	0.6	-0.2	-0.3	-0.2	0.554	0.543	0.580	-0.0459	-0.0709	-0.0455	0.831	0.748	0.831
0.5	0.4	0.6	-0.2	-0.3	-1	0.564	0.451	0.496	-0.0313	-0.0623	-0.2400	0.881	0.778	0.373
0.5	0.4	0.6	-1	-1	-1	0.422	0.394	0.451	-0.1803	-0.1875	-0.1837	0.464	0.459	0.462
0.5	0.4	0.6	-4	-4	-4	0.286	0.260	0.315	-0.3376	-0.3866	-0.3617	0.048	0.043	0.046