Discovering Sparse Covariance Structures with the Isomap

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Graphical Models Reading Group

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Overview

- We talked about covariance estimation with Cholesky decomposition and GLM.
- A large class of methods with assumption: variables have a natural ordering, such as longitudinal data, time series, spatial data.
- We will first talk about some other methods in that class, then how to discover a structured ordering with Isomap, as proposed by Wagaman, A. S., & Levina, E. (2009).

Banding

• Given a $p \times p$ sample covariance matrix $\hat{\mathbf{\Sigma}} = [\hat{\sigma}_{ij}]$ and an integer k, $0 \le k \le p$, the k-banded estimator is defined by

$$B_k(\hat{\mathbf{\Sigma}}) = [\hat{\sigma}_{ij}\mathbf{1}\{|i-j| \leq k\}].$$

 Asymptotic results: the banded estimator and its inverse are consistent as long as

$$\frac{\log p}{n}\to 0.$$

Not necessarily positive definite.

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Tapering

• Given a tapering matrix $\mathbf{W} = [w_{ij}]$, a tapered estimator is defined by

$$B_{\mathbf{W}}(\hat{\mathbf{\Sigma}}) = \hat{\mathbf{\Sigma}} \circ \mathbf{W} = [\hat{\sigma}_{ij} w_{ij}].$$

- Schur product theorem: if **W** is positive definite, $B_{\mathbf{W}}(\hat{\Sigma})$ is positive definite.
- Banding corresponds to

$$\mathbf{W} = [\mathbf{1}\{|i-j| \le k\}]$$

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Tapering

• For example, the trapezoidal tapering matrix

$$w_{ij} = \begin{cases} 1 & \text{if } |i-j| \leq I_h, \\ 2 - |i-j|/I_h & \text{if } I_h < |i-j| < I, \\ 0 & \text{otherwise,} \end{cases}$$

for a given tapering parameter I and $I_h = I/2$.

 Asymptotic results: the tapered estimator and its inverse are consistent as long as

$$\frac{\log p}{n} \to 0.$$

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Thresholding

• The thresholded estimator for a $\lambda \geq 0$ is defined by

$$T_{\lambda}(\hat{\mathbf{\Sigma}}) = [\hat{\sigma}_{ij}\mathbf{1}\{\hat{\sigma}_{ij} \geq \lambda\}].$$

- Different from banding and tapering, thresholding is permutation-invariant.
- Asymptotic results: the thresholded estimator and its inverse are consistent as long as

$$\frac{\log p}{n}\to 0.$$

 Wagaman, A. S., & Levina, E. (2009) claim that banding has a better convergence rate than thresholding. While Pourahmadi, M. (2013) does not completely agree.

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Multidimensional Scaling (MDS)

- Mapping data in \mathbb{R}^p to a lower-dimensional \mathbb{R}^k , preserving the pairwise dissimilarity as well as possible.
- Observations $x_1, x_2, \ldots, x_N \in \mathbb{R}^p$, and let d_{ij} be the dissimilarities between observations i and j. Often we choose Euclidean distance $d_{ij} = ||x_i x_j||$.
- Metric multidimensional scaling: find values $z_1, z_2, \dots, z_N \in \mathbb{R}^k$ to minimize

$$S_M(z_1, z_2, \dots, z_N) = \sum_{i \neq j} (d_{ij} - ||z_i - z_j||)^2.$$

• A gradient descent algorithm is used to minimized S_M .

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Multidimensional Scaling (MDS)

• Classical multidimensional scaling: Start with similarities $s_{ij} = \langle x_i - \bar{x}, x_j - \bar{x} \rangle$. Minimize

$$S_C(z_1, z_2, \ldots, z_N) = \sum_{i,j} (s_{ij} - \langle z_i - \overline{z}, z_j - \overline{z} \rangle)^2.$$

 Classical multidimensional scaling is equivalent to principal components analysis.

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Multidimensional Scaling (MDS)

- Metric MDS and Classical MDS approximate the actual dissimilarity or similarities. Nonmetric multidimensional scaling effectively uses only ranks.
- Minimize

$$S_{NM}(z_1, z_2, \ldots, z_N, \theta) = \frac{\sum_{i \neq j} (\theta(d_{ij}) - ||z_i - z_j||)^2}{\sum_{i \neq j} ||z_i - z_j||^2},$$

where θ is an increasing function.

• With θ fixed, minimize over z_i by gradient descent. With z_i fixed, use isotonic regression to find best monotonic approximation θ .

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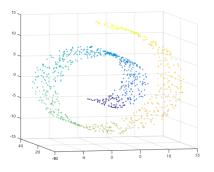
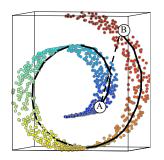


Figure: Swiss roll dataset.

- *k*-nearest neighbors (*k*-NN) adjacency graph: a sparse graph representing local structure.
- Distance or similarity can be defined based on the graph.

- Euclidean distance between two neighboring nodes as weight of edge.
- Shortest paths: Djikstra's algorithm.
- Geodesic distance as dissimilarity, apply MDS.



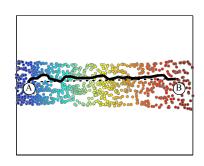


Figure: Euclidean distance vs geodesic distance.

The Isomap algorithm

- 1. For each point, find its r nearest neighbors using the dissimilarities d(i,j). Construct a neighborhood graph by connecting each point to its r neighbors, with dissimilarities as the edge weights.
- 2. Estimate the geodesic distance $\tilde{d}_r(i,j)$ between each pair of points i,j by computing the shortest-path distance from i to j through the neighborhood graph.
- 3. Apply metric MDS to the matrix of pairwise shortest-path distances to obtain an embedding in \mathbb{R}^d . In our case, this means find $z_1, \ldots, z_p \in \mathbb{R}^1$ that minimize the stress function (known as stress 1 in the literature)

$$S(z_1, \ldots, z_p) = rac{\sum_i \sum_j (|z_i - z_j| - \tilde{d}_r(i, j))^2}{\sum_i \sum_j |z_i - z_j|^2} \; .$$

This minimization reduces to computing the first eigenvector of a matrix derived from pairwise distances.

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There are many ways to define a dissimilarity measure.

- $d_{ij} = 1 |\hat{\rho}_{ij}|$, where $\hat{\rho}_{ij}$ is the sample correlation coefficient between variable i and j.
- $\bullet \ d_{ij}=1-\hat{\rho}_{ij}.$
- $d_{ij} = C |\hat{\sigma}_{ij}|$, where $\hat{\sigma}_{ij}$ is the sample covariance between variable i and j.



Isoband (Isomap+banding)

• From the ordering, we construct a $p \times p$ permutation matrix **P**. The covariance matrix is reordered

$$\hat{\pmb{\Sigma}}_0 = P\hat{\pmb{\Sigma}}P^T.$$

• The banding operator B_k is applied to $\hat{\Sigma}_0$, and then reorder back.

$$\hat{\mathbf{\Sigma}}_1 = \mathbf{P}^T B_k(\hat{\mathbf{\Sigma}}_0) \mathbf{P}.$$

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Bootstrap

Bootstrapped isoband algorithm: modify Step 1 as follows:

- 1a Resample the observations with replacement T times and construct T bootstrap sample covariance matrices $\hat{\Sigma}_1^*, \dots, \hat{\Sigma}_T^*$.
- 1b For each matrix $\hat{\Sigma}_t^*$, t = 1, ..., T, construct a neighborhood graph by connecting each variable to its r nearest neighbors using dissimilarities $d_t^*(i,j)$ based on $\hat{\Sigma}_t^*$.
- 1c Construct the final neighborhood graph by putting an edge between variables i and j if an edge is present between i and j in at least cT of the bootstrap graphs, where $c \in (0,1)$ is a tuning parameter, and assign weight d(i,j) (original dissimilarity) to the edge.

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References



Wagaman, A. S., & Levina, E. (2009). Discovering sparse covariance structures with the isomap. Journal of Computational and Graphical Statistics, 18(3), 551-572.



Pourahmadi, M. (2013). High-Dimensional Covariance Estimation: With High-Dimensional Data. John Wiley & Sons.



Saul, L. K., Weinberger, K. Q., Ham, J. H., Sha, F., & Lee, D. D. (2006). Spectral methods for dimensionality reduction. Semisupervised learning, 293-308.



Trevor J.. Hastie, Tibshirani, R. J., & Friedman, J. H. (2009). The elements of statistical learning: data mining, inference, and prediction. Springer.

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