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Time Series Analysis and Its Applications: With R Examples

Second Edition

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This is the site for the second edition of the text and is no longer maintained.

Follow this link if you're looking for the site of the third edition.

An R Time Series Tutorial

Here are some examples that may help you become familiar with analyzing time series using R. You can copy-and-paste the R commands (multiple lines are ok) from this page into R. Printed output is *blue*. I suggest that you have R up and running before you start this tutorial.

Please note that this is not a lesson in time series analysis. Also, the analyses performed on this page are simply demonstrations, they are not meant to be optimal or complete in any way. This is done intentionally so as not to spoil the fun you'll have working on the problems in the text.

If you're new to R/Splus, I suggest reading *R for Beginners* (a pdf file) first. Another good read for exploring time series is *Econometrics in R* (a pdf file). You may also want to poke around the *QuickR* website.

♦ Baby steps... your first R session. Get comfortable, then start her up and try some simple addition:

```
2+2
[1] 5
```

Ok, now you're an expert use \mathbf{R} . It's time to move on to time series. What you'll see in the following examples should be enough to get you through the first four chapters of the text.

Let's play with the Johnson & Johnson data set. Download <u>jj.dat</u> to a directory called mydata (or wherever you choose ... the examples below and in the text assume the data are in that directory).

```
jj = scan("/mydata/jj.dat")  # read the data
jj <- scan("/mydata/jj.dat")  # read the data another way
scan("/mydata/jj.dat") -> jj  # and another
```

The *R people* (yes, they exist) prefer that you use the second [<-] or third [->] assignment operator, but your wrists and health care professionals prefer that you use the simpler first [-] method if you can.

Next, print jj (to the screen)

```
jj

[1] 0.71 0.63 0.85 0.44

[5] 0.61 0.69 0.92 0.55

. . . . . .

[77] 14.04 12.96 14.85 9.99

[81] 16.20 14.67 16.02 11.61
```

and you see that jj is a collection of 84 numbers called an object. You can see all of your objects by typing

```
objects()
```

If you're a Matlab (or similar) user, you may think jj is an 84×1 vector, but it's not. It has order and length, but no dimensions (no rows, no columns). R call these objects vectors so you have to be careful. In R, matrices have dimensions but vectors do not. To wit:

```
jj[1]  # the first element
  [1] 0.71
jj[84]  # the last element
  [1] 11.61
jj[1:4]  # the first 4 elements
  [1] 0.71 0.63 0.85 0.44
jj[-(1:80)] # everything EXCEPT the first 80 elements
  [1] 16.20 14.67 16.02 11.61
length(jj)  # the number of elements
  [1] 84
```

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```
dim(jj)  # but no dimensions ...
  NULL
nrow(jj)  # ... no rows
  NULL
ncol(jj)  # ... and no columns
  NULL
#-- if you want it to be a column vector (in R, a matrix), an easy way to go is:
jj = as.matrix(jj)
dim(jj)
  [1] 84  1
```

Now, let's make jj a time series object.

```
jj = ts(jj, start=1960, frequency=4)
```

Note that the data are *quarterly* earnings, hence the frequency=4 statement. One nice thing about R is you can do a bunch of stuff (technical term) in one line. For example, you can read the data into jj and make it a time series object at the same time:

```
jj = ts(scan("/mydata/jj.dat"), start=1960, frequency=4)
```

In the lines above, you can replace scan by read.table. Inputting data using read.table is an easy way to read a data file that is laid out as a matrix and may have headers (column descriptions). At this point, you might want to find out about read.table, data frames, and time series objects:

```
jj = ts(read.table("/mydata/jj.dat"), start=1960, frequency=4)
help(read.table)
help(ts)
help(data.frame)
```

There is a difference between scan and read.table. The former produces a vector (no dimensions) while the latter produces a data frame (and has dimensions).

One final note on reading the data. If the data started on the third quarter of 1960, say, then you would have something like ts(x, start=c(1960,3), frequency=4) and so on. If you had monthly data that started from June, 1984, then you would have ts(x, start=c(1984,6), frequency=12).

Let's view the data again as a time series object:

```
jj

Qtr1 Qtr2 Qtr3 Qtr4

1960 0.71 0.63 0.85 0.44

1961 0.61 0.69 0.92 0.55

. . . . . .

1979 14.04 12.96 14.85 9.99

1980 16.20 14.67 16.02 11.61
```

Notice the difference? You also get some nice things with the ts object, for example, the corresponding time values:

By the way, you could have put the data into jj and printed it at the same time by enclosing the command:

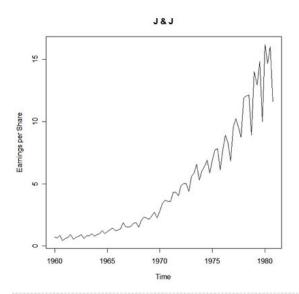
```
(jj = ts(scan("/mydata/jj.dat"), start=1960, frequency=4))
```

Now try a plot of the data:

```
plot(jj, ylab="Earnings per Share", main="J & J")
```

with the result being:

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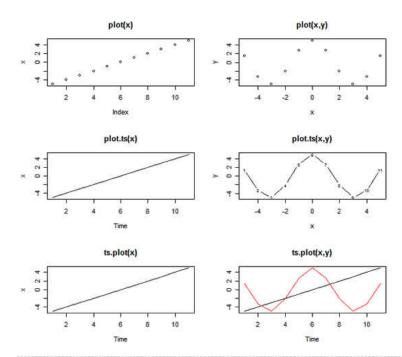
Try these and see what happens:

```
plot(jj, type="o", col="blue", lty="dashed")
plot(diff(log(jj)), main="logged and diffed")
```

and while you're here, check out plot.ts and ts.plot:

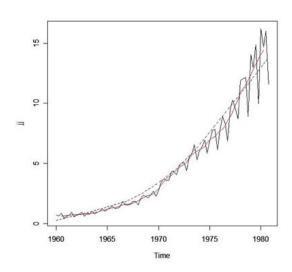
```
x = -5:5
                                # sequence of integers from -5 to 5
y = 5*cos(x)
                                # guess
                                # multifigure setup: 3 rows, 2 cols
par(mfrow=c(3,2))
#--- plot:
plot(x, main="plot(x)")
plot(x, y, main="plot(x,y)")
#--- plot.ts:
plot.ts(x, main="plot.ts(x)")
plot.ts(x, y, main="plot.ts(x,y)")
#--- ts.plot:
ts.plot(x, main="ts.plot(x)")
\texttt{ts.plot}(\texttt{ts}(\texttt{x})\texttt{, ts}(\texttt{y})\texttt{, col=1:2, main="ts.plot}(\texttt{x},\texttt{y})") \quad \texttt{\# note-} \texttt{ x and } \texttt{y} \texttt{ are ts objects}
\#--- the help files [? and help() are the same]:
?plot.ts
help(ts.plot)
              # might as well skim the graphical parameters help file while you're here
?par
```

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Note that if your data are a time series object, plot() will do the trick (for a simple time plot, that is). Otherwise, plot.ts() will coerce the graphic into a time plot.

... and the result:



Let's difference the logged data and call it dljj. Then we'll play with dljj:

```
dljj = diff(log(jj))  # difference the logged data
plot(dljj)  # plot it if you haven't already
shapiro.test(dljj)  # test for normality
    Shapiro-Wilk normality test
    data: dljj
```

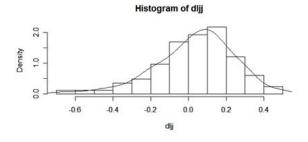
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```
W = 0.9725, p-value = 0.07211
```

Now a histogram and a Q-Q plot, one on top of the other:

```
par(mfrow=c(2,1))  # set up the graphics
hist(dljj, prob=TRUE, 12)  # histogram
lines(density(dljj))  # smooth it - ?density for details
qqnorm(dljj)  # normal Q-Q plot
qqline(dljj)  # add a line
```

and the results:

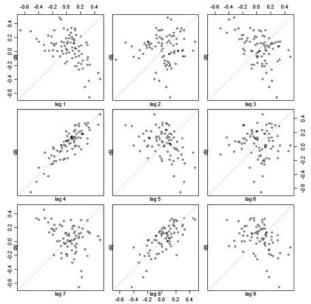


Normal Q-Q Plot Selection of the control of the co

Let's check out the correlation structure of dljj using various techniques. First, we'll look at a grid of scatterplots of dljj(t-lag) vs dljj(t) for $lag=1,2,\ldots,9$.

```
lag.plot(dljj, 9, do.lines=FALSE) # why the do.lines=FALSE? ... try leaving it out
```

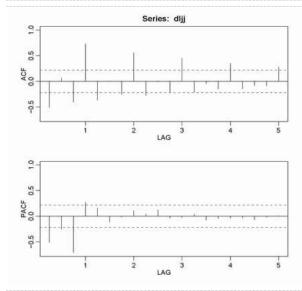
Notice the large positive correlation at lags 4 and 8 and the negative correlations at a few other lags:



Now let's take a look at the ACF and PACF of dljj:

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!!NOTE!! acf2 on the line below is NOT available in R... details follow the graph below
acf2(dljj) # this is what you'll see below



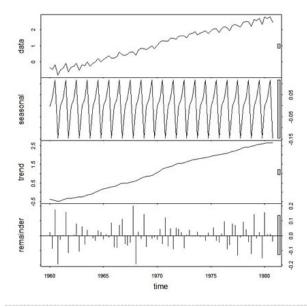
Note that the LAG axis is in terms of frequency, so 1,2,3,4,5 correspond to lags 4,8,12,16,20 because frequency=4 here. If you don't like this type of labeling, you can replace dljj in any of the above by ts(dljj, freq=1); e.g., acf(ts(dljj, freq=1), 20)

♦ Ok- here's the story on acf2. I like my ACF and PACF a certain way, which is not the default R way. So, I wrote a little script called acf2.R that you can read about and obtain here: Examples (there are other goodies there).

Moving on, let's try a structural decomposition of log(jj) = trend + season + error using lowess. Note, this example works only if jj is dimensionless (i.e., you didn't read it in using read.table ... thanks to Jon Moore of the University of Reading, U.K., for pointing this out.)

```
plot(dog <- stl(log(jj), "per"))</pre>
```

Here's what you get:



If you want to inspect the residuals, for example, they're in dog\$time.series[,3], the third column of the resulting series (the seasonal and trend components are in columns 1 and 2). Check out the ACF of the residuals, acf(dog\$time.series[,3]); the residuals aren't white- not even close. You can do a little (very little) better using a local seasonal window, plot(dog <- stl(log(jj), s.win=4)), as opposed to the global one used by specifying "per". Type ?stl for details. There's also something called StructTS that will fit parametric structural models. We don't use these functions in the text when we present structural modeling in Chapter 6 because we prefer to use our own programs.

♦ This is a good time to explain \$. In the above, dog is an object containing a bunch of things (technical term). If you type dog, you'll see the components, and if you type summary(dog) you'll get a little summary of the results. One of the components of dog is time.series, which contains the resulting series (seasonal, trend, remainder). To see this component of the object dog, you type dog\$time.series (and you'll see 3 series, the last of which contains the residuals). And that's the story of \$... you'll see more examples as we move along.

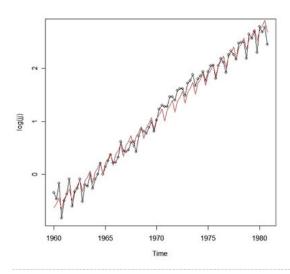
```
And now, we'll do some of Problem 2.1. We're going to fit the regression \log(jj) = \beta * time + \alpha_1 * Q1 + \alpha_2 * Q2 + \alpha_3 * Q3 + \alpha_4 * Q4 + \epsilon
```

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```
where Qi is an indicator of the quarter i = 1, 2, 3, 4. Then we'll inspect the residuals.
Q = factor(rep(1:4,21))
                          # make (Q)uarter factors [that's repeat 1,2,3,4, 21 times]
trend = time(jj)-1970  # not necessary to "center" time, but the results look nicer
reg = lm(log(jj) \sim 0 + trend + Q, na.action=NULL) # run the regression without an intercept
 #-- the na.action statement is to retain time series attributes
 summary(reg)
 lm(formula = log(jj) \sim 0 + trend + Q, na.action = NULL)
 Residuals:
              1Q Median
                                30
 -0.29318 -0.09062 -0.01180 0.08460 0.27644
 Coefficients:
      Estimate Std. Error t value Pr(>|t|)
 trend 0.167172 0.002259 74.00 <2e-16 ***
 Q1 1.052793 0.027359 38.48 <2e-16 ***
 Q2 1.080916 0.027365 39.50 <2e-16 ***
 Q3 1.151024 0.027383 42.03 <2e-16 ***
 Q4 0.882266 0.027412 32.19 <2e-16 ***
 Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.1254 on 79 degrees of freedom
Multiple R-squared: 0.9935, Adjusted R-squared: 0.9931
F-statistic: 2407 on 5 and 79 DF, p-value: < 2.2e-16
You can view the model matrix (with the dummy variables) this way:
model.matrix(reg)
    trend Q1 Q2 Q3 Q4
                        (remember trend is time centered at 1970)
 1 -10.00 1 0 0
    -9.75 0 1
                 0 0
    -9.50 0 0
                 1
    -9.25 0 0 0
    -9.00 1 0 0
    -8.75 0 1
                 0 0
    -8.50 0 0 1 0
    -8.25 0 0 0 1
 81 10.00 1 0 0 0
 82 10.25 0 1 0 0
 83 10.50 0 0 1 0
 84 10.75 0 0 0 1
Now check out what happened. Look at a plot of the observations and their fitted values:
plot(log(jj), type="o")  # the data in black with little dots
lines(fitted(reg), col=2) # the fitted values in bloody red - or use lines(reg$fitted, col=2)
```

you get:

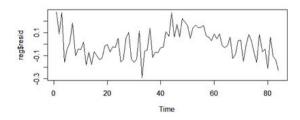
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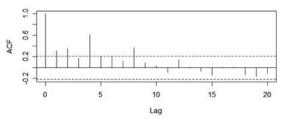
... and a plot of the residuals and the ACF of the residuals:

```
par(mfrow=c(2,1))
plot(resid(reg))  # residuals - reg$resid is same as resid(reg)
acf(resid(reg),20)  # acf of the resids
```

and you get:



Series reg\$resid



Do those residuals look white? [Ignore the 0-lag correlation, it's always 1.]

You have to be careful when you **regress one time series on lagged components of another** using lm(). There is a package called dynlm that makes it easy to fit lagged regressions, and I'll discuss that right after this example. If you use lm(), then what you have to do is "tie" the series together using ts.intersect. If you don't tie the series together, they won't be aligned properly. Here's an example regressing weekly cardiovascular mortality (cmort.dat) on particulate pollution (cmort.dat) at the present value and lagged four weeks (about a month). For details about the data set, see Chapter 2.

```
mort = ts(scan("/mydata/cmort.dat"),start=1970, frequency=52) # make these time series objects
  Read 508 items
part = ts(scan("/mydata/part.dat"), start=1970, frequency=52)
  Read 508 items
ded = ts.intersect(mort,part,part4=lag(part,-4), dframe=TRUE) # tie them together in a data frame
fit = lm(mort~part+part4, data=ded, na.action=NULL)
                                                            # now the regression will work
summary(fit)
  Call:
  lm(formula = mort ~ part + part4, data = ded, na.action = NULL)
  Residuals:
      Min
                10
                    Median
                                   30
                                           Max
  -22.7429 -5.3677 -0.4136
                               5.2694 37.8539
```

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```
Coefficients:

Estimate Std. Error t value Pr(>|t|)

(Intercept) 69.01020   1.37498  50.190  < 2e-16 ***

part   0.15140  0.02898  5.225  2.56e-07 ***

part4   0.26297  0.02899  9.071  < 2e-16 ***

---

Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1

Residual standard error: 8.323 on 501 degrees of freedom

Multiple R-Squared: 0.3091, Adjusted R-squared: 0.3063

F-statistic: 112.1 on 2 and 501 DF, p-value: < 2.2e-16
```

Note: There was no need to rename lag(part, -4) to part4, it's just an example of what you can do.

An alternative to the above is the package dynlm, which has to be installed [for details, in R type help(INSTALL) or help("install.packages")]. After the package is installed, you can do the previous example as follows:

```
library(dynlm)
                                     # load the package
fit = dynlm(mort~part + lag(part,-4)) # assumes mort and part are ts objects
# fit = dynlm(mort~part + L(part, 4)) is the same thing.
summary(fit)
 dynlm(formula = mort ~ part + lag(part, -4))
  Residuals:
      Min
               1Q Median
                                 3Q
  -22.7429 -5.3677 -0.4136 5.2694 37.8539
  Coefficients:
               Estimate Std. Error t value Pr(>|t|)
  (Intercept) 69.01020 1.37498 50.190 < 2e-16 ***
               0.15140
                          0.02898
                                   5.225 2.56e-07 ***
  lag(part, -4) 0.26297
                          0.02899
                                   9.071 < 2e-16 ***
  Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
 Residual standard error: 8.323 on 501 degrees of freedom
 Multiple R-Squared: 0.3091, Adjusted R-squared: 0.3063
 F-statistic: 112.1 on 2 and 501 DF, p-value: < 2.2e-16
```

Well, it's time to simulate. The workhorse for ARIMA simulations is arima.sim(). Here are some examples; no output is shown here so you're on your own.

```
# some AR1s
x1 = arima.sim(list(order=c(1,0,0), ar=.9), n=100)
x2 = arima.sim(list(order=c(1,0,0), ar=-.9), n=100)
par(mfrow=c(2,1))
plot(x1, main=(expression(AR(1)~~~phi==+.9)))
                                                    # ~ is a space and == is equal
plot(x2, main=(expression(AR(1) \sim \sim phi==-.9)))
x11()
                                                # open another graphics device if you wish
par(mfcol=c(2,2))
acf(x1, 20)
acf(x2, 20)
pacf(x1, 20)
pacf(x2, 20)
# you could have, for example, used acf2(x1)
\# to get the ACF and PACF of x1 (or x2)... if you had acf2.R, of course.
# an MA1
x = arima.sim(list(order=c(0,0,1), ma=.8), n=100)
par(mfcol=c(3,1))
plot(x, main=(expression(MA(1) \sim \sim theta== .8)))
acf(x,20)
pacf(x, 20)
```

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```
# an AR2
x = arima.sim(list(order=c(2,0,0), ar=c(1,-.9)), n=100)
par(mfcol=c(3,1))
plot(x, main=(expression(AR(2)~~~phi[1]==1~~~phi[2]==-.9)))
acf(x, 20)
pacf(x, 20)

# an ARIMA(1,1,1)
x = arima.sim(list(order=c(1,1,1), ar=.9, ma=-.5), n=200)
par(mfcol=c(3,1))
plot(x, main=(expression(ARIMA(1,1,1)~~~phi==.9~~~theta==-.5)))
acf(x, 30) # the process is not stationary, so there is no population [P]ACF ...
pacf(x, 30) # but look at the sample values to see how they differ from the examples above
```

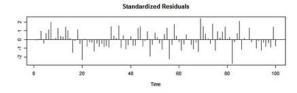
 \Diamond Next, we're going to do some ARIMA estimation. This gets a bit tricky because R is not useR friendly when it comes to fitting ARIMA models. Much of the story is spelled out in our R Issues page. I'll be as gentle as I can at first.

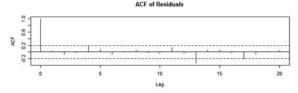
First, we'll fit an ARMA model to some simulated data (with diagnostics and forecasting):

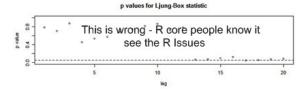
... diagnostics:

tsdiag(x.fit, gof.lag=20) # you know the routine- ?tsdiag for details

... and the output







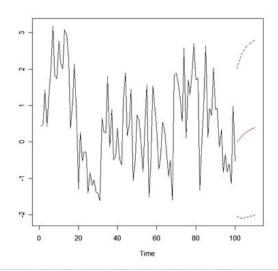
... forecast 10 ahead:

```
x.fore = predict(x.fit, n.ahead=10)
# plot the forecasts
U = x.fore$pred + 2*x.fore$se
L = x.fore$pred - 2*x.fore$se
minx=min(x,L)
maxx=max(x,U)
```

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```
ts.plot(x,x.fore$pred,col=1:2, ylim=c(minx,maxx))
lines(U, col="blue", lty="dashed")
lines(L, col="blue", lty="dashed")
```

... and here's the plot of the data and the forecasts (with error bounds):



That wasn't too bad... but hold on. We're going to work with the global temperature data from Chapter 3. The data are in the file globtemp2.dat. There are three columns in the file, the second column has the yearly global temperature deviations from 1880 to 2004. If you download the data to the mydata directory, you can pull out the global temperatures like this:

```
u = read.table("/mydata/globtemp2.dat")  # read the data
gtemp = ts(u[,2], start=1880, freq=1)  # yearly temp in col 2
plot(gtemp)  # graph the series (not shown here)
```

... long story short, the data appear to be an ARIMA(1,1,1) with a drift of about +.6 °C per century (and hence the global warming hypothesis). Let's fit the model:

```
arima(gtemp, order=c(1,1,1))

Coefficients:

ar1 ma1

0.2545 -0.7742

s.e. 0.1141 0.0651
```

So what's wrong? well, there's no estimate of the drift!! With no drift, the global warming hypothesis is kaput (technical term)... that is, the temps are just basically taking a random walk. How do you get the estimate of drift?... do this:

What happened? The two runs should have given the same results, but the default models for the two cases are different. I won't go into detail here because the details can be found on the <u>R Issues</u> page. And, of course, this problem continues if you try to do forecasting. There are remedies. One remedy is to do the following:

and then make sure you continue the along these lines when you forecast. Another remedy is to use the scripts called sarima.R for model fitting, and sarima.for.R for forecasting. You can get those scripts with some details on this page: Examples.

 \Diamond You may not have understood all the details of this example, but at least you should realize that you may get into trouble when fitting ARIMA models with R. In particular, you should come away from this realizing that, in R, arima(x, order=c(1,1,1)) is different than arima(diff(x), order=c(1,0,1)) and arima calls the estimate of the *mean* the *intercept*. Again, much of the story is spelled out in our <u>R Issues</u> page.

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And now for some **regression with autocorrelated errors**. This can be accomplished two different ways. First, we'll use gls() from the package nlme, which you have to load. We're going to fit the model $M_t = \alpha + \beta t + \gamma P_t + e_t$ where M_t and P_t are the mortality and particulates series from a previous example, and e_t is autocorrelated error.

```
library(nlme)
                       # load the package
trend = time(mort)
                       \# assumes mort and part are there from previous examples
fit.lm = lm(mort~trend + part) # ols
acf(resid(fit.lm))
                       # check acf and pacf of the resids
                    # or use acf2(resid(fit.lm)) if you have acf2
pacf(resid(fit.lm))
# resids appear to be AR(2) ... now use gls() from nlme:
fit.gls = gls(mort~trend + part, correlation=corARMA(p=2), method="ML")
# take 5 ......
#.....
# done .
summary(fit.gls)
     Parameter estimate(s):
          Phil Phi2
     0.3980566 0.4134305
  Coefficients: Value Std.Error t-value p-value
  (Intercept) 3131.5452 857.2141 3.653166
 trend
                -1.5444 0.4340 -3.558021
                                              4e-04
                 0.1503
                          0.0210 7.162408
                                              0e+00
 part
\# resid analysis- we assumed e_t = \phi_1 e_{t-1} + \phi_2 e_{t-2} + w_t where w_t is white.
w = filter(resid(fit.gls), filter=c(1,-.3980566, -.4134305), sides=1) # get resids
w = w[-2:-1]
                             # first two are NA
Box.test(w, 12, type="Ljung") # check whiteness via Ljung-Box-Pierce statistic
  X-squared = 8.6074, df = 12, p-value = 0.736
pchisq(8.6074, 10, lower=FALSE)  # the p-value (they are resids from an ar2 fit)
  [1] 0.569723
```

Now, we'll doing the same thing using arima(), which is easier and a little quicker.

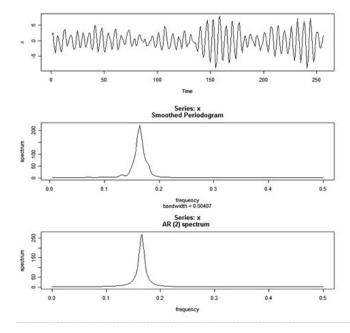
♦ ARMAX: If you want to fit an ARMAX model you have to do it via a state space model... more details will follow on the Chapter 6 page when I have the time. As seen above, using xreg in arima() does NOT fit an ARMAX model, which is too bad, but the help file (?arima) didn't say it did. For more info, head on over to the R Issues page and check out Issue 2.

Finally, a spectral analysis quicky:

```
x = arima.sim(list(order=c(2,0,0), ar=c(1,-.9)), n=2^8) # some data
 (u = polyroot(c(1, -1, .9)))
                                       # x is AR(2) w/complex roots
  [1] 0.5555556+0.8958064i 0.5555556-0.8958064i
 Arg(u[1])/(2*pi)
                                       # dominant frequency around .16:
  [1] 0.1616497
par(mfcol=c(3,1))
plot.ts(x)
spec.pgram(x, spans=c(3,3), log="no") # nonparametric spectral estimate; also see spectrum()
                                       # some help
  'spec.pgram' calculates the periodogram using a fast Fourier
  transform, and optionally smooths the result with a series of
  modified Daniell smoothers (moving averages giving half weight to
  the end values).
spec.ar(x, log="no")
                                       # parametric spectral estimate
```

and the graph:

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Finally, note that R tapers and logs by default, so if you simply want the periodogram of a series, the command is spec.pgram(x), taper=0, fast=FALSE, detrend=FALSE, log="no"). If you just asked for spec.pgram(x), you wouldn't get the RAW periodogram because the data are detrended, possibly padded, and tapered in spec.pgram, even though the title of the resulting graphic would say Raw Periodogram. An easier way to get a raw periodogram is this: $per=abs(fft(x))^2$ duh. The moral of this story ... and the bottom line: pay special attention to the defaults of the functions you're using.

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