

## Solutions for problem set 07

1.(a)from

$$\begin{aligned}C_t &= \beta_1^0 + \beta_2^0(Y_t - T_t) + \varepsilon_t \\T_t &= \gamma_1^0 + \gamma_2^0 Y_t + v_t \\Y_t &= C_t + G_t\end{aligned}$$

we can get

$$\begin{aligned}Y_t - T_t &= \frac{\beta_1^0 - \gamma_1^0 - \gamma_2^0 \beta_1^0}{1 - \beta_2^0 + \beta_2^0 \gamma_2^0} + \frac{1 - \gamma_2^0}{1 - \beta_2^0 + \beta_2^0 \gamma_2^0}(G_t + \varepsilon_t) + \frac{1}{1 - \beta_2^0 + \beta_2^0 \gamma_2^0}v_t \\E[(Y_t - T_t)\varepsilon_t] &= \frac{1 - \gamma_2^0}{1 - \beta_2^0 + \beta_2^0 \gamma_2^0}\sigma_u^2\end{aligned}$$

if  $\gamma_2^0 = 1$ , OLS estimator of  $\beta$  is consistence. otherwise, the OLS estimator of  $\beta$  is NOT consistence.

(b)Yes, We can use  $G_t$  as a valid instrumental variable and get a consistent estimator  $\hat{\beta}$  by 2sls.since  $G_t$  is exogeneous variable.  $\text{cov}(G_t, \varepsilon_t) = 0$  and  $\text{cov}(G_t, Y_t - T_t) = \frac{1 - \gamma_2^0}{1 - \beta_2^0 + \beta_2^0 \gamma_2^0}$ .

Stage1:

$$\begin{aligned}Y_t - T_t &= \gamma_0 + \gamma_1 G_t + \mu_t \\Y - T &= G \cdot \gamma + \mu \quad G = (1, G_t) \\\hat{\gamma}_{ols} &= (G'G)^{-1}G'(Y - T) \\(Y - T)^\wedge &= G \cdot \hat{\gamma}_{ols} = G(G'G)^{-1}G'(Y - T)\end{aligned}$$

Stage2:

$$\begin{aligned}C_t &= \beta_1^0 + \beta_2^0 (Y_t - T_t)^\wedge + \tilde{\varepsilon}_t \\YI_t &= (Y_t - T_t)^\wedge \quad YI = (1, YI_t) \\C &= YI \cdot \beta + \tilde{\varepsilon} \\\hat{\beta}_{2sls} &= (YI'YI)^{-1}YI'C\end{aligned}$$

where

$$YI = (1, G(G'G)^{-1}G'(Y - T))$$

(c)

$$\begin{aligned}G_t &= \frac{\beta_1^0 \gamma_2^0 + \gamma_1^0 - \beta_2^0 \gamma_1^0}{1 - \beta_2^0 + \beta_2^0 \gamma_1^0 - \gamma_2^0} + \frac{\gamma_2^0}{1 - \beta_2^0 + \beta_2^0 \gamma_1^0 - \gamma_2^0}\varepsilon_t + \frac{1 - \beta_2^0}{1 - \beta_2^0 + \beta_2^0 \gamma_1^0 - \gamma_2^0}v_t + \frac{1 - \beta_2^0 + \beta_2^0 \gamma_2^0}{1 - \beta_2^0 + \beta_2^0 \gamma_1^0 - \gamma_2^0}w_t \\E(G_t \varepsilon_t) &= \frac{\gamma_2^0}{1 - \beta_2^0 + \beta_2^0 \gamma_1^0 - \gamma_2^0}\sigma_u^2\end{aligned}$$

for  $\gamma_2^0 \neq 0$ ,  $E(G_t \varepsilon_t) \neq 0$ . Therefore,  $G_t$  is not a valid instrumental variable.

2. (a)

$$Y_t = X'_t \beta^0 + \varepsilon_t = \beta_0^0 + X_{1t} \beta_1^0 + \varepsilon_t$$

$$E(X_{1t} \varepsilon_t) = E[(v_t + u_t)(w_t + u_t)] = E(u_t^2) = \sigma_u^2 = 1$$

therefore, OLS estimator  $\hat{\beta}$  is inconsistent for  $\beta^0$ . (b)

$$Z_{1t} = w_t - \varepsilon_t = -u_t$$

$$E(Z_{1t} \varepsilon_t) = E[(-u_t)(w_t + u_t)] = -E(u_t^2) = -\sigma_u^2 = -1$$

that means  $Z_t$  is NOT a valid instrumental vector.

(c)

$$E(v_t \varepsilon_t) = E[v_t(w_t + u_t)] = 0$$

$$E(V_{1t} v_t) = E[(v_t + u_t)v_t] = E(v_t^2) = 1$$

Thus,  $v_t$  is a valid instrumental variable,  $Z_t$  is a valid instrumental variable.  
stage1:

$$X = Z \cdot \gamma + \mu$$

$$\hat{\gamma} = (Z'Z)^{-1}Z'X$$

$$\hat{X} = X(Z'Z)^{-1}Z'X$$

stage2:

$$Y = \hat{X}\beta + \tilde{\varepsilon}$$

$$\begin{aligned} \hat{\beta}_{2sls} &= (\hat{X}'\hat{X})^{-1}\hat{X}'Y \\ &= \left[ X'Z(Z'Z)^{-1}Z'Z(Z'Z)^{-1}Z'X \right]^{-1} X'Z(Z'Z)^{-1}Z'Y \\ &= \left[ X'Z(Z'Z)^{-1}Z'X \right]^{-1} X'Z(Z'Z)^{-1}Z'Y \\ &= \left[ X'Z(Z'Z)^{-1}Z'X \right]^{-1} X'Z(Z'Z)^{-1}Z'(X\beta + \varepsilon) \\ &= \beta_0 + \left[ X'Z(Z'Z)^{-1}Z'X \right]^{-1} X'Z(Z'Z)^{-1}Z'\varepsilon \end{aligned}$$

$$\begin{aligned} \sqrt{n}(\hat{\beta}_{2sls} - \beta_0) &= \left[ \frac{X'Z}{n} \left( \frac{Z'Z}{n} \right)^{-1} \frac{Z'X}{n} \right]^{-1} \frac{X'Z}{n} \left( \frac{Z'Z}{n} \right)^{-1} \frac{Z'\varepsilon}{\sqrt{n}} \\ &= \left[ \hat{Q}_{XZ}(\hat{Q}_{ZZ})^{-1}\hat{Q}_{ZX} \right]^{-1} \hat{Q}_{XZ}(\hat{Q}_{ZZ})^{-1} \frac{Z'\varepsilon}{\sqrt{n}} \end{aligned}$$

By WLLN,  $\hat{Q}_{XZ} \xrightarrow{p} Q_{XZ}$ ,  $\hat{Q}_{ZZ} \xrightarrow{p} Q_{ZZ}$ ,  $\hat{Q}_{ZX} \xrightarrow{p} Q_{ZX}$ ,

$$\left[ \hat{Q}_{XZ} \hat{Q}_{ZZ}^{-1} \hat{Q}_{ZX} \right]^{-1} \hat{Q}_{XZ} \hat{Q}_{ZZ}^{-1} \xrightarrow{p} \left[ Q_{XZ} Q_{ZZ}^{-1} Q_{ZX} \right]^{-1} Q_{XZ} Q_{ZZ}^{-1} \triangleq A$$

By CLT,

$$\frac{Z' \varepsilon}{\sqrt{n}} = \frac{1}{\sqrt{n}} \sum_{t=1}^n Z_t \varepsilon_t \xrightarrow{p} N(0, V_1)$$

$$\begin{aligned} V_1 &= E(Z_t Z_t' \varepsilon_t^2) \\ &= E(Z_t Z_t' E(\varepsilon_t^2 | Z_t)) \\ &= 2E(Z_t Z_t') = 2Q_{ZZ} \end{aligned}$$

$$\sqrt{n}(\hat{\beta}_{2sls} - \beta_0) \rightarrow N(0, V)$$

$$\begin{aligned} V &= AV_1 A = 2(Q_{XZ} Q_{ZZ}^{-1} Q_{ZX})^{-1} \\ &= 2 \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} \\ &= 2I_2 \end{aligned}$$

$$\sqrt{n}(\hat{\beta}_{2sls} - \beta_0) \rightarrow N(0, 2I_2)$$

(d)

$$JF = (R\hat{\beta} - r) \left( S^2 R(X'X)^{-1} R' \right)^{-1} (R\hat{\beta} - r) \sim \chi_J^2$$

where  $S^2 = \frac{\hat{e}'\hat{e}}{n-K}$ ,  $\hat{e} = Y - X\hat{\beta}_{2sls}$ . for the second stage,  $\tilde{e} = Y - \hat{X}\hat{\beta}_{2sls}$ ,  $\tilde{S}^2 = \frac{\tilde{e}'\tilde{e}}{n-K}$ .

$$\begin{aligned} \tilde{e} &= Y - \hat{X}\hat{\beta}_{2sls} \\ &= (Y - X\hat{\beta}_{2sls}) + (X - \hat{X})\hat{\beta}_{2sls} \\ &= \hat{e} + \hat{\mu}\hat{\beta}_{2sls} \end{aligned}$$

where  $\hat{\mu}$  is the estimate residual from the first stage.

$$\begin{aligned} \tilde{e}'\tilde{e} &= \left( \hat{e} + \hat{\mu}\hat{\beta}_{2sls} \right)' \left( \hat{e} + \hat{\mu}\hat{\beta}_{2sls} \right) \\ &= \hat{e}'\hat{e} + 2\hat{e}'\hat{\mu}\hat{\beta}_{2sls} + \hat{\beta}_{2sls}'\hat{\mu}'\hat{\mu}\hat{\beta}_{2sls} \\ &= \sum \hat{e}_t^2 + 2\hat{\beta}_{2sls}^1 \sum \hat{e}_t \hat{\mu}_t + \left( \hat{\beta}_{2sls}^1 \right)^2 \sum \hat{\mu}_t^2 \end{aligned}$$

$$\frac{\tilde{e}'\tilde{e}}{n-K} = \frac{\sum \hat{e}_t^2}{n-K} + \frac{2}{n-K} \hat{\beta}_{2sls}^1 \sum \hat{e}_t \hat{\mu}_t + \frac{1}{n-K} \left( \hat{\beta}_{2sls}^1 \right)^2 \sum \hat{\mu}_t^2 \xrightarrow{P} \sigma^2 + 2\beta_0^1 \text{cov}(\mu, e) + (\beta_0^1)^2 \sigma_\mu^2$$

thus,  $\tilde{S}^2$  is not a consistent estimation for  $S^2$ . therefore,  $J\tilde{F}$  from the second stage does not converge to  $\chi_J^2$ .

3.(a) because  $E(X_t \varepsilon_t) = 0$  and  $E(X_t X_t') = Q_{XX}$  is nonsingular. we can regard  $X_t$  as an instrumental variable for itself. That is,  $Z = X$ .

$$\begin{aligned} \hat{\beta}_{2sls} &= \left[ X'X(X'X)^{-1}X'X \right]^{-1} X'X(X'X)^{-1}X'Y \\ &= (X'X)^{-1}X'Y \\ &= \hat{\beta} \end{aligned}$$

$\hat{\beta}$  is a special 2SLS estimator  $\hat{\beta}_{2sls}$  with instrumental variable  $X$ .

(b)

$$\begin{aligned} \hat{\beta} &= (X'X)^{-1}X'Y \\ &= (X'X)^{-1}X'(X\beta^0 + \varepsilon) \\ &= \beta^0 + (X'X)^{-1}X'\varepsilon \end{aligned}$$

$$\begin{aligned} \hat{\beta} - \beta^0 &\xrightarrow{P} Q_{XX} E(X_t \varepsilon_t) = 0 \\ \sqrt{n}(\hat{\beta} - \beta^0) &= \left( \frac{X'X}{n} \right)^{-1} \frac{X'\varepsilon}{\sqrt{n}} \xrightarrow{P} Q_{XX}^{-1} \frac{\sum X_t \varepsilon_t}{\sqrt{n}} \\ \text{var} \left( \sqrt{n}(\hat{\beta} - \beta^0) \right) &\xrightarrow{P} Q_{XX}^{-1} E(X_t X_t' \varepsilon_t^2) Q_{XX}^{-1} = \sigma^2 Q_{XX}^{-1} \end{aligned}$$

By CLT,

$$\sqrt{n}(\hat{\beta} - \beta^0) \xrightarrow{d} N(0, \sigma^2 Q_{XX}^{-1})$$

$$\begin{aligned} \hat{\beta}_{2sls} &= \left[ X'Z(Z'Z)^{-1}Z'X \right]^{-1} X'Z(Z'Z)^{-1}Z'Y \\ &= \beta_0 + \left[ X'Z(Z'Z)^{-1}Z'X \right]^{-1} X'Z(Z'Z)^{-1}Z'\varepsilon \\ &= \beta_0 + \left[ \frac{X'Z}{n} \left( \frac{Z'Z}{n} \right)^{-1} \frac{Z'X}{n} \right]^{-1} \frac{X'Z}{n} \left( \frac{Z'Z}{n} \right)^{-1} \frac{Z'\varepsilon}{n} \\ &\xrightarrow{P} 0 \end{aligned}$$

$$\sqrt{n}(\hat{\beta}_{2sls} - \beta_0) = \left[ \frac{X'Z}{n} \left( \frac{Z'Z}{n} \right)^{-1} \frac{Z'X}{n} \right]^{-1} \frac{X'Z}{n} \left( \frac{Z'Z}{n} \right)^{-1} \frac{Z'\varepsilon}{\sqrt{n}} \xrightarrow{P} (Q_{XZ} Q_{ZZ}^{-1} Q_{ZX}) Q_{XZ} Q_{ZZ}^{-1} \frac{Z'\varepsilon}{\sqrt{n}}$$

$$\text{var} \left( \frac{Z' \varepsilon}{\sqrt{n}} \right) = E(Z_t Z'_t \varepsilon_t^2) = \sigma^2 Q_{ZZ}$$

$$\text{var} \left( \hat{\beta}_{2sls} \right) = (Q_{XZ} Q_{ZZ}^{-1} Q_{ZX})^{-1} Q_{XZ} Q_{ZZ}^{-1} Q_{ZZ} Q_{ZZ}^{-1} Q_{XZ} (Q_{XZ} Q_{ZZ}^{-1} Q_{ZX})^{-1} \sigma^2 = \sigma^2 (Q_{XZ} Q_{ZZ}^{-1} Q_{ZX})^{-1}$$

By CLT,

$$\sqrt{n}(\hat{\beta}_{2sls} - \beta_0) \xrightarrow{d} N \left( 0, \sigma^2 (Q_{XZ} Q_{ZZ}^{-1} Q_{ZX})^{-1} \right)$$

$$Q_{XX} - Q_{XZ} Q_{ZZ}^{-1} Q_{ZX} = E(X_t X'_t) - E(X_t Z'_t) [E(Z_t Z'_t)]^{-1} E(Z_t X'_t)$$

$$\begin{aligned} \frac{X'X}{n} - \frac{X'Z}{n} \left( \frac{Z'Z}{n} \right)^{-1} \frac{Z'X}{n} &= \frac{1}{n} X' (I - Z(Z'Z)^{-1} Z') X \\ &= \frac{1}{n} X' M_Z X \\ &= \frac{1}{n} (M_Z X)' M_Z X \\ &= \frac{1}{n} D' D \end{aligned}$$

So  $Q_{XX} - Q_{XZ} Q_{ZZ}^{-1} Q_{ZX}$  is positive semi-definite.  $\hat{\beta}$  is more asymptotically efficient than  $\hat{\beta}_{2sls}$

(c)

$$\begin{aligned} \hat{\beta} - \beta^0 &= (X'X)^{-1} X' \varepsilon = \left( \frac{X'X}{n} \right)^{-1} \frac{X' \varepsilon}{n} \\ \sqrt{n}(\hat{\beta} - \beta^0) &= \left( \frac{X'X}{n} \right)^{-1} \frac{X' \varepsilon}{\sqrt{n}} = \left( \frac{\sum X'_t X_t}{n} \right)^{-1} \frac{\sum X'_t \varepsilon_t}{\sqrt{n}} \\ &= \left( \frac{\sum X'_t X_t}{n} \right)^{-1} \xrightarrow{P} Q_{XX}^{-1} \\ \frac{\sum X_t \varepsilon_t}{\sqrt{n}} &= E \left( \frac{\sum X_t \varepsilon_t}{\sqrt{n}} \right) + o_p \left( \sqrt{\text{var} \left( \frac{\sum X_t \varepsilon_t}{\sqrt{n}} \right)} \right) \end{aligned}$$

since  $E \left( \frac{\sum X_t \varepsilon_t}{\sqrt{n}} \right) = 0$  and  $\text{var} \left( \frac{\sum X_t \varepsilon_t}{\sqrt{n}} \right) = O_p(1)$ ,

$$\frac{\sum X_t \varepsilon_t}{\sqrt{n}} = O_p(1)$$

$$\left( \frac{X'X}{n} \right)^{-1} \frac{X' \varepsilon}{\sqrt{n}} - Q_{XX}^{-1} \frac{\sum X_t \varepsilon_t}{\sqrt{n}} \xrightarrow{P} 0$$

$$\sqrt{n}(\hat{\beta} - \beta^0) - Q_{XX}^{-1} \frac{\sum X_t \varepsilon_t}{\sqrt{n}} \xrightarrow{P} 0$$

So,

$$\sqrt{n}(\hat{\beta} - \beta^0) = Q_{XX}^{-1} \frac{\sum X_t \varepsilon_t}{\sqrt{n}} + O_p(1)$$

(d)

$$\sqrt{n}(\hat{\beta}_{2sls} - \beta_0) = \left[ \frac{X'Z}{n} \left( \frac{Z'Z}{n} \right)^{-1} \frac{Z'X}{n} \right]^{-1} \frac{X'Z}{n} \left( \frac{Z'Z}{n} \right)^{-1} \frac{Z'\varepsilon}{\sqrt{n}}$$

by WLLN,

$$\left[ \frac{X'Z}{n} \left( \frac{Z'Z}{n} \right)^{-1} \frac{Z'X}{n} \right]^{-1} \frac{X'Z}{n} \left( \frac{Z'Z}{n} \right)^{-1} \xrightarrow{P} (Q_{XZ}Q_{ZZ}^{-1}Q_{ZX}) Q_{XZ}Q_{ZZ}^{-1}$$

$$\frac{\sum Z_t \varepsilon_t}{\sqrt{n}} = E \left( \frac{\sum Z_t \varepsilon_t}{\sqrt{n}} \right) + O_p \left( \sqrt{\text{var} \left( \frac{\sum Z_t \varepsilon_t}{\sqrt{n}} \right)} \right)$$

$$\text{var} \left( \frac{\sum Z_t \varepsilon_t}{\sqrt{n}} \right) = O_p(1)$$

$$\frac{\sum Z_t \varepsilon_t}{\sqrt{n}} = O_p(1)$$

$$\sqrt{n}(\hat{\beta}_{2sls} - \beta^0) = (Q_{XZ}Q_{ZZ}^{-1}Q_{ZX})^{-1} Q_{XZ}Q_{ZZ}^{-1} \frac{\sum X_t \varepsilon_t}{\sqrt{n}} + o_p(1)$$

(e) By (c) and (d)

$$\begin{aligned} \sqrt{n}(\hat{\beta} - \hat{\beta}_{2sls}) &= \sqrt{n}(\hat{\beta} - \beta^0) - \sqrt{n}(\hat{\beta}_{2sls} - \beta^0) \\ &= Q_{XX}^{-1} \frac{\sum X_t \varepsilon_t}{\sqrt{n}} + o_p(1) - (Q_{XZ}Q_{ZZ}^{-1}Q_{ZX})^{-1} Q_{XZ}Q_{ZZ}^{-1} \frac{\sum X_t \varepsilon_t}{\sqrt{n}} - o_p(1) \\ &= \frac{1}{\sqrt{n}} \sum \left( Q_{XX}^{-1} X_t - (Q_{XZ}Q_{ZZ}^{-1}Q_{ZX})^{-1} Q_{XZ}Q_{ZZ}^{-1} X_t \right) \varepsilon_t + o_p(1) \end{aligned}$$

(f) Denote  $\Lambda_{XZ} = (Q_{XZ}Q_{ZZ}^{-1}Q_{ZX})^{-1} Q_{XZ}Q_{ZZ}^{-1}$  and  $A_t = Q_{XX}^{-1} X_t - \Lambda_{XZ} X_t$  since the asymptotic distribution of  $\sqrt{n}(\hat{\beta} - \hat{\beta}_{2sls})$  is determined by the leading term.

$$\frac{1}{n} \sum (Q_{XX}^{-1} X_t - \Lambda_{XZ} X_t) \varepsilon_t = Q_{XX}^{-1} \sum \frac{X_t \varepsilon_t}{n} - \Lambda_{XZ} \sum \frac{Z_t \varepsilon_t}{n} \xrightarrow{P} 0$$

$$\begin{aligned} \text{var} \left( \sqrt{n}(\hat{\beta} - \hat{\beta}_{2sls}) | XZ \right) &= \sum_{t=1}^n A_t \frac{1}{n} \text{var}(\varepsilon_t | X_t Z_t) A_t' \\ &= \frac{1}{n} \sigma^2 \sum_{t=1}^n A_t A_t' = \frac{1}{n} \sigma^2 \sum_{t=1}^n (Q_{XX}^{-1} X_t - \Lambda_{XZ} Z_t) (Q_{XX}^{-1} X_t - \Lambda_{XZ} Z_t)' \\ &= \frac{1}{n} \sigma^2 \sum_{t=1}^n (Q_{XX}^{-1} X_t X_t' Q_{XX}^{-1} - \Lambda_{XZ} Z_t X_t' Q_{XX}^{-1} - Q_{XX}^{-1} X_t Z_t' \Lambda_{XZ}' + \Lambda_{XZ} Z_t Z_t' \Lambda_{XZ}') \\ &\xrightarrow{P} \sigma^2 [Q_{XX}^{-1} Q_{XX} Q_{XX}^{-1} - \Lambda_{XZ} Q_{ZX} Q_{XX}^{-1} - Q_{XX}^{-1} Q_{XZ} \Lambda_{XZ}' + \Lambda_{XZ} Q_{ZZ} \Lambda_{XZ}'] \\ &= \sigma^2 \left[ (Q_{XZ}Q_{ZZ}^{-1}Q_{ZX})^{-1} - Q_{XX}^{-1} \right] \end{aligned}$$

$$\sqrt{n}(\hat{\beta} - \hat{\beta}_{2sls}) \xrightarrow{d} N \left( 0, \sigma^2 \left[ (Q_{XZ}Q_{ZZ}^{-1}Q_{ZX})^{-1} - Q_{XX}^{-1} \right] \right)$$

(g)

$$\frac{\sqrt{n}(\hat{\beta} - \hat{\beta}_{2sls})}{\sqrt{\sigma^2 \left( (Q_{XZ}Q_{ZZ}^{-1}Q_{ZX})^{-1} - Q_{XX}^{-1} \right)}} \xrightarrow{d} N(0, 1)$$

$$n(\hat{\beta} - \hat{\beta}_{2sls}) \left[ (Q_{XZ}Q_{ZZ}^{-1}Q_{ZX})^{-1} - Q_{XX}^{-1} \right]^{-1} (\hat{\beta} - \hat{\beta}_{2sls})' \sim \chi_J^2$$

where  $J$  is the number of endogeneous variable.

4. (a)

$$e_t = Y_t - \hat{X}_t' \hat{\beta}_{2sls} = Y_t - X_t' \hat{\beta}_{2sls} + (X_t' - \hat{X}_t') \hat{\beta}_{2sls}$$

$$e = Y - X \hat{\beta}_{2sls} + (X - \hat{X}) \hat{\beta}_{2sls} = \hat{e} + (X - \hat{X}) \hat{\beta}_{2sls}$$

$$S^2 = \frac{e'e}{n-K} = \frac{(\hat{e} + \hat{\mu} \hat{\beta}_{2sls})' (\hat{e} + \hat{\mu} \hat{\beta}_{2sls})}{n-K}$$

$$= \frac{1}{n-K} \sum_{t=1}^n \left[ \hat{e}_t^2 + 2 \hat{\beta}_{2sls}' \hat{\mu}_t \hat{e}_t + \left( \hat{\beta}_{2sls}' \hat{\mu}_t \right)^2 \right]$$

$$\hat{e}_t = X_t' (\hat{\beta} - \hat{\beta}_{2sls}) + \varepsilon_t$$

$$\hat{\mu}_t = Z_t' (\gamma - \hat{\gamma}) + \mu_t$$

$$\frac{1}{n-K} \sum_{t=1}^n \hat{e}_t^2 \xrightarrow{P} \sigma^2$$

$$\frac{1}{n-K} \sum_{t=1}^n \hat{\beta}_{2sls}' \hat{\mu}_t \hat{e}_t \xrightarrow{P} \beta_0 E(X_t \varepsilon_t)$$

$$\frac{1}{n-K} \sum_{t=1}^n \left( \hat{\beta}_{2sls}' \hat{\mu}_t \right)^2 \xrightarrow{P} \beta_0 \Sigma_{\mu} \beta_0$$

$$S^2 = \frac{e'e}{n-K} \xrightarrow{P} \sigma^2 + \beta_0 E(X_t \varepsilon_t) + \beta_0 \Sigma_{\mu} \beta_0 \neq \sigma^2$$

Thus,  $S^2$  is not consistent for  $\sigma^2$ .

(b)

since  $S^2$  is not consistent for  $\sigma^2$ ,  $S^2 R \hat{Q}_{\hat{X}\hat{X}}^{-1} R'$  is not a consistent estimator for the variance of  $\sqrt{n}(R\hat{\beta}_{2sls} - r)$ , so  $\chi_J^2$  is not valid under  $H_0$  let  $\hat{\mu} = X - \hat{X}$ , then  $e = \hat{e} + \hat{\mu} \hat{\beta}_{2sls}$