

Stat314/461Term 4: Logistic Regression from a Bayesian viewpoint

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Logistic regression

Regression modelling for a binary response variable

$$Y_i \sim \text{Bernoulli}(\theta_i); \quad i = 1, \dots, N$$

$$\log \left(\frac{\theta_i}{(1 - \theta_i)} \right) = \beta_0 + X_{1i}\beta_1 + X_{2i}\beta_2 + X_{3i}\beta_3 + \dots$$

or equivalently

$$\theta_i = \frac{\exp(\beta_0 + X_{1i}\beta_1 + X_{2i}\beta_2 + X_{3i}\beta_3 + \dots)}{1 + \exp(\beta_0 + X_{1i}\beta_1 + X_{2i}\beta_2 + X_{3i}\beta_3 + \dots)}$$

This model ensures $0 < \theta_i < 1$, for all i , or, equivalently, for all combinations of the X 's.

We are modelling the relationship of the success probabilities to the covariates X_1, X_2, \dots i.e “regression.”

Interpretation of logistic regression coefficients (i)

Let

$$\theta = \frac{\exp(\beta_0 + X_1\beta_1 + X_2\beta_2 + X_3\beta_3 + \dots)}{1 + \exp(\beta_0 + X_1\beta_1 + X_2\beta_2 + X_3\beta_3 + \dots)}$$
$$\log\left(\frac{\theta}{1 - \theta}\right) = \beta_0 + X_1\beta_1 + X_2\beta_2 + X_3\beta_3 + \dots \quad (1)$$

Effect on $\text{logit}(\theta)$ on changing from, say, X_1 to $X_1 + 1$ is

$$\beta_0 + (X_1 + 1)\beta_1 + X_2\beta_2 + X_3\beta_3 + \dots \quad (2)$$

$$- \beta_0 + (X_1)\beta_1 + X_2\beta_2 + X_3\beta_3 + \dots \quad (3)$$

$$= \beta_1 \quad (4)$$

Interpretation of logistic regression coefficients (ii)

$$\log \left(\frac{\theta}{1-\theta} | (X_1 + 1) \right) - \log \left(\frac{\theta}{1-\theta} | X_1 \right) = \beta_1 \quad (5)$$

Hence

$$\exp \left[\log \left(\frac{\theta}{1-\theta} | (X_1 + 1) \right) - \log \left(\frac{\theta}{1-\theta} | X_1 \right) \right] = \exp(\beta_1) \quad (6)$$

$$\frac{[\theta/(1-\theta)] | (X_1 + 1)}{[\theta/(1-\theta)] | X_1} = \exp(\beta_1) \quad (7)$$

exponentiated logistic regression coefficients are interpretable as odds ratios (except for intercept)

$\text{invlogit}(\beta_0)$ is the probability of the event when $X_1 = X_2 = \dots = 0$

Example

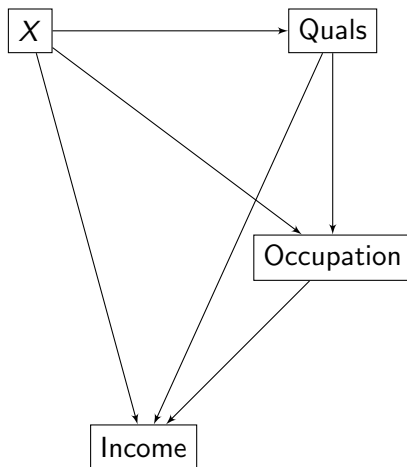
Example: $\Pr(\text{income} > \$1250/\text{week} | \text{age}, \text{sex}, \text{quals})$

- from the income survey data

$$p(\beta | \mathbf{Y}) \propto p(\beta) \prod_i \theta_i^{Y_i} (1 - \theta_i)^{(1 - Y_i)}$$
$$\theta_i = \frac{\exp(\beta_0 + X_{1i}\beta_1 + X_{2i}\beta_2 + X_{3i}\beta_3 + \dots)}{1 + \exp(\beta_0 + X_{1i}\beta_1 + X_{2i}\beta_2 + X_{3i}\beta_3 + \dots)}$$

and $X_1 = \text{age}$, $X_2 = \text{sex}$, $X_3, X_4, X_5, X_6 = \text{indicators for qualification categories}$.

Directed Acyclic Graph for Income model



Likelihood for the logistic regression example (i)

X fixed; Y random (or analysing conditional on X)
really just a sequence of Bernoulli trials but with different success probabilities on each trial.

Don't panic!

Let $\mathbf{X}_i = (1, X_{1i}, X_{2i}, \dots)'$, $\boldsymbol{\beta} = (\beta_0, \beta_1, \beta_2, \dots)'$

$\mathbf{X}'_i \boldsymbol{\beta} = \sum_{j=0}^{i=K} X_{ji} \beta_j$ where $X_{0i} = 1$

$$\theta_i = \text{expit}(\mathbf{X}'_i \boldsymbol{\beta}) = \text{invlogit}(\mathbf{X}'_i \boldsymbol{\beta}) = \frac{\exp(\mathbf{X}'_i \boldsymbol{\beta})}{1 + \exp(\mathbf{X}'_i \boldsymbol{\beta})}.$$

Likelihood for the logistic regression example (ii)

If

$$\log \left(\frac{\theta_i}{(1 - \theta_i)} \right) = \mathbf{X}_i' \beta$$

$$\theta_i = \text{invlogit}(\mathbf{X}_i' \beta) \Rightarrow (1 - \theta_i) = (1 + \exp(\mathbf{X}_i' \beta))^{-1}$$

and

$$\begin{aligned} p(\mathbf{Y}|\mathbf{X}, \beta) &= \prod_i \theta_i^{Y_i} (1 - \theta_i)^{(1 - Y_i)} \\ &= \prod_i \left(\frac{\theta_i}{1 - \theta_i} \right)^{Y_i} (1 - \theta_i) \end{aligned}$$

Likelihood for the logistic regression example (ii)

If

$$\log \left(\frac{\theta_i}{(1 - \theta_i)} \right) = \mathbf{X}_i' \beta$$

$$\theta_i = \text{invlogit}(\mathbf{X}_i' \beta) \Rightarrow (1 - \theta_i) = (1 + \exp(\mathbf{X}_i' \beta))^{-1}$$

and

$$\begin{aligned} p(\mathbf{Y}|\mathbf{X}, \beta) &= \prod_i \theta_i^{Y_i} (1 - \theta_i)^{(1 - Y_i)} \\ &= \prod_i \left(\frac{\theta_i}{1 - \theta_i} \right)^{Y_i} (1 - \theta_i) \end{aligned}$$

$$\begin{aligned} \log(p(\mathbf{Y}|\mathbf{X}, \beta)) &= \sum_i Y_i \log \left(\frac{\theta_i}{1 - \theta_i} \right) + \log((1 + \exp(\mathbf{X}_i' \beta))^{-1}) \\ &= \sum_i Y_i \times \mathbf{X}_i' \beta - \log(1 + \exp(\mathbf{X}_i' \beta)) \end{aligned}$$

Prior for logistic regression coefficients (ii)

- We have a vector of parameters $\beta = (\beta_0, \beta_1, \beta_2, \dots)'$. Even if we model these parameters as *a priori* independent they will be correlated in the posterior: (To get the idea try

```
logitmodel <- glm(highincome ~ hoursfactor + sexfactor +  
qualfactor,family=binomial(link="logit") )  
vcov(logitmodel)
```

which returns the variance *matrix* for the mle.

Prior for logistic regression coefficients (ii)

- It is convenient to treat $\beta = (\beta_0, \beta_1, \beta_2, \dots)'$ as a single entity in computations.
- If we do regard the components of β as *a priori* independent and adopt a normal prior we are saying

$$p(\beta) = \prod_{j=0}^{j=K} \text{normal}(\beta_j | \mu_j, \sigma_j^2) \quad (8)$$

This is equivalent to saying $\beta \sim \text{MVN}(\mu, \Omega)$ where

$$\mu = (\mu_0, \mu_1, \dots, \mu_K)' \text{ and } \Omega = \begin{bmatrix} \sigma_0^2 & 0 & 0 & \dots \\ 0 & \sigma_1^2 & 0 & \dots \\ \vdots & \vdots & \vdots & \vdots \\ 0 & 0 & 0 \dots & \sigma_K^2 \end{bmatrix}$$

Prior for logistic coefficients (iii)

- In general the multivariate normal (MVN) distribution is characterised by a mean *vector* and variance *matrix*.
- Off-diagonal elements in the variance matrix are *covariances*.
Non-zero covariances indicate correlation between the corresponding elements of the vector being modelled

$$\begin{bmatrix} \sigma_0^2 & \sigma_{01} & \sigma_{02} & \dots \sigma_{0K} \\ \sigma_{01} & \sigma_1^2 & \sigma_{12} & \dots \sigma_{1K} \\ \vdots & \vdots & \vdots & \vdots \\ \sigma_{0K} & \sigma_{1K} & \sigma_{2K} \dots & \sigma_K^2 \end{bmatrix}$$

$$\rho_{12} = \frac{\sigma_{12}}{\sigma_1 \sigma_2} \text{ etc}$$

Prior for logistic coefficients (iv)

If we assume *a priori* independence we can construct the joint prior for β by considering each component separately.

- $\pi_0 = \text{invlogit}(\beta_0)$ is the probability of the event if the X 's are all zero. Could
 - 1 set, say, 95% prior limits on π_0
 - 2 transform to logit scale
 - 3 construct a normal that matches the transformed limits:
 $\mu_0 = 0.5(\text{logit}(\pi_{0,low}) + \text{logit}(\pi_{0,high}))$; variance s.t transformed limits are 2.5% and 97.5% quantiles of a normal with mean μ_0
- $\gamma_j = \exp(\beta_j)$ is the odds ratio for the effect of $X_j, j = 1, \dots, K$.
 - 1 set 95% prior limits on γ_j (remember 1 is the “null” value)
 - 2 transform to log-scale
 - 3 construct a normal that matches the transformed limits (2.5% and 97.5% quantiles of a normal.)

Importance sampling for logistic regression

- If we can:

- ① write a function to compute the log-likelihood for any value of β
- ② write a function to compute the log prior density for any value of β

we can compute the log unnormalised posterior density at any setting of the parameters.

- We can obtain a multivariate normal approximation to the posterior by maximizing the log unnormalised posterior density wrt β to find the posterior mode and the curvature of the log-posterior at the model. The approximating MVN is centred at the posterior mode and has variance based on the inverse of the curvature.
- Standard optimisation functions can find these quantities. The `laplace` function from the `LearnBayes` package is useful; `bayesglm` from the `ARM` package can also be used.

Example: logistic regression for analysis of effect of sex and tertiary education on probability of high income

- see logistic regression importance sampling R code for an application based on the income survey data
`logisticregression_importance2021.r`