

# Package ‘spbridge’

December 2, 2025

**Title** Marginally interpretable spatial logistic regression with bridge processes

**Version** 1.0.0.2

**Description** Provides functions for marginally interpretable spatial logistic regression with bridge processes. For details, see Lee and Dunson (2025+) <[doi:10.48550/arXiv.2412.04744](https://doi.org/10.48550/arXiv.2412.04744)>.

**License** GPL (>= 3)

**URL** <https://github.com/changwoo-lee/spbridge>

**BugReports** <https://github.com/changwoo-lee/spbridge/issues>

**Encoding** UTF-8

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**Imports** BayesLogit,  
coda,  
fields,  
lme4,  
Matrix,  
matrixStats,  
methods,  
spam,  
stats,  
utils

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gambia

*Gambia Malaria Data*

## Description

Malaria prevalence in children recorded at villages in The Gambia, Africa. This dataset and documentation is imported from R package GeoR version 1.9-6, inheriting the GPL-3 license.

## Usage

```
data(gambia)
```

## Format

Two objects are made available:

### 1. gambia

A data frame with 2035 observations on the following 8 variables.

**x** x-coordinate of the village (UTM).

**y** y-coordinate of the village (UTM).

**pos** presence (1) or absence (0) of malaria in a blood sample taken from the child.

**age** age of the child, in days

**netuse** indicator variable denoting whether (1) or not (0) the child regularly sleeps under a bed-net.

**treated** indicator variable denoting whether (1) or not (0) the bed-net is treated (coded 0 if netuse=0).

**green** satellite-derived measure of the green-ness of vegetation in the immediate vicinity of the village (arbitrary units).

**phc** indicator variable denoting the presence (1) or absence (0) of a health center in the village.

### 2. gambia.borders

A data frame with 2 variables:

**x** x-coordinate of the country borders.

**y** y-coordinate of the country borders.

## References

Thomson, M., Connor, S., D Alessandro, U., Rowlingson, B., Diggle, P., Cresswell, M. & Greenwood, B. (1999). Predicting malaria infection in Gambian children from satellite data and bednet use surveys: the importance of spatial correlation in the interpretation of results. *American Journal of Tropical Medicine and Hygiene* 61: 2–8.

Diggle, P., Moyeed, R., Rowlingson, B. & Thomson, M. (2002). Childhood malaria in The Gambia: a case-study in model-based geostatistics, *Applied Statistics*.

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splogi_bridge	<i>Spatial logistic model with bridge process random effect (empirical Bayes)</i>
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**Description**

$$\text{logit} [\Pr(y_{ij} = 1 \mid X_{ij}, u(s_i))] = X_{ij}^\top \beta + u(s_i)$$

where  $u(s) \sim \text{Bridge process}$  with parameter  $\phi$  and Matern correlation kernel,  $i=1,\dots,n$  corresponds to  $n$  spatial locations and  $j=1,\dots, N_i$  correspond to  $n_i$  responses at location  $i$ , resulting data of size  $N = N_1 + \dots + N_n$ . The parameter  $\phi$  is estimated using empirical Bayes approach.

**Usage**

```
splogi_bridge(
  y,
  X,
  id,
  coords,
  priors = list(beta_intercept_scale = 10, beta_scale = 2.5, beta_df = Inf, logpriorphi =
    NULL, rho_lb = NULL, rho_ub = NULL),
  smoothness = 1.5,
  nburn = 100,
  nsave = 1000,
  nthin = 1,
  verbose = TRUE
)
```

**Arguments**

y	$N \times 1$ binary vector
X	$N \times p$ fixed-effect design matrix, including intercept
id	$N \times 1$ vector of spatial location id. When $N=n$ , it is point-referenced data
coords	$n \times 2$ matrix of spatial coordinates
priors	list of prior hyperparameters, see details
smoothness	positive numeric, Matern smoothness parameter
nburn	number of burn-in iteration
nsave	number of posterior samples
nthin	thin-in rate
verbose	logical, whether to print progress

**Details**

Priors are specified by "priors" argument, which is a list of hyperparameters. Specifically, we set zero centered normal or t prior for beta, uniform prior for Matern range parameter rho (see `fields::Matern`), and for phi, we use prior that induces half-Cauchy prior on the standard deviation of  $u$ .

"priors" is a named list with the following possible elements:

**beta\_intercept\_scale** scale of intercept parameter (default 10)  
**beta\_scale** scale of other beta parameters (default 2.5)  
**beta\_df** degrees of freedom for t prior on beta (default Inf, normal prior)  
**rho\_lb** lower bound for range parameter rho (default min distance between coords)  
**rho\_ub** upper bound for range parameter rho (default max distance between coords)

### Value

Returns list of

post_save	a matrix of posterior samples (coda::mcmc) with nsave rows
u_save	a matrix of posterior samples (coda::mcmc) of random effects, with nsave rows
betam_save	a matrix of posterior samples (coda::mcmc) of population-averaged log odds, with nsave rows
loglik_save	a nsave x n matrix of pointwise log-likelihood values, can be used for WAIC calculation.
priors	list of hyperprior information
nsave	number of MCMC samples
t_mcmc	wall-clock time for running MCMC
t_premcmc	wall-clock time for preprocessing before MCMC
y	response vector
X	fixed effect design matrix
coords	a n x 2 matrix of Euclidean coordinates

### Examples

```
library(spbridge)
data(gambia)
N = length(gambia$pos) # 2035
y = gambia$pos # binary response, N = 2035 by 1 vector
# define id based on spatial coords unique values
id = as.numeric(factor(paste(gambia$x, gambia$y)))
n = length(unique(id)) # 65
coords = unique(cbind(gambia$x, gambia$y)/1000) # n by 2 matrix, in km
# standardized covariates following Gelman et al (2008)
intercept = rep(1,N) # intercept
age = scale(gambia$age/365, scale = 2*sd(gambia$age/365)) # in years
netuse = gambia$netuse - mean(gambia$netuse)
treated = gambia$treated - mean(gambia$treated)
green = scale(gambia$green, scale = 2*sd(gambia$green))
green2 = scale(gambia$green^2, scale = 2*sd(gambia$green^2))
healthctr = gambia$phc - mean(gambia$phc)

X = cbind(intercept, age, netuse, treated, green, green2, healthctr)
colnames(X) = c("(Intercept)", "age", "netuse", "treated", "green", "green2", "healthctr")

centers = c(attr(age, "scaled:center"), mean(gambia$netuse), mean(gambia$treated),
           attr(green, "scaled:center"), attr(green2, "scaled:center"), mean(gambia$phc))
scales = c(attr(age, "scaled:scale"), 1, 1, attr(green, "scaled:scale"), attr(green2, "scaled:scale"), 1)

#
```

```
# commented out due to a long run time
#
# fit_bridge = spbridge::splogi_bridge(y = y,
#                                     X = X,
#                                     id = id,
#                                     priors = list(beta_intercept_scale = 10,
#                                     beta_scale = 2.5, beta_df = Inf,
#                                     rho_lb = 0.01, rho_ub = 100),
#                                     coords = coords,
#                                     smoothness = 0.5, nburn = 1000, nsave = 10000, nthin = 1)
```

**splogi\_bridge\_fb***Spatial logistic model with bridge process random effect (fully Bayes)***Description**

$$\text{logit} [\Pr(y_{ij} = 1 \mid X_{ij}, u(s_i))] = X_{ij}^\top \beta + u(s_i)$$

where  $u(s) \sim \text{Bridge process}$  with parameter  $\phi$  and Matern correlation kernel,  $i=1,\dots,n$  corresponds to  $n$  spatial locations and  $j=1,\dots, N_i$  correspond to  $n_i$  responses at location  $i$ , resulting data of size  $N = N_1 + \dots + N_n$ . The parameter  $\phi$  is estimated using fully Bayesian approach by placing prior on it.

**Usage**

```
splogi_bridge_fb(
  y,
  X,
  id,
  coords,
  priors = list(beta_intercept_scale = 10, beta_scale = 2.5, beta_df = Inf, logpriorphi =
    NULL, rho_lb = NULL, rho_ub = NULL),
  smoothness = 1.5,
  nburn = 100,
  nsave = 1000,
  nthin = 1,
  nparticle = 20,
  verbose = TRUE
)
```

**Arguments**

<b>y</b>	$N \times 1$ binary vector
<b>X</b>	$N \times p$ fixed-effect design matrix, including intercept
<b>id</b>	$N \times 1$ vector of spatial location id. When $N=n$ , it is point-referenced data
<b>coords</b>	$n \times 2$ matrix of spatial coordinates
<b>priors</b>	list of prior hyperparameters, see details

<b>smoothness</b>	positive numeric, Matern smoothness parameter
<b>nburn</b>	number of burn-in iteration
<b>nsave</b>	number of posterior samples
<b>nthin</b>	thin-in rate
<b>nparticle</b>	number of particles in particle marginal Metropolis-Hastings
<b>verbose</b>	logical, whether to print progress

## Details

Priors are specified by "priors" argument, which is a list of hyperparameters. Specifically, we set zero centered normal or t prior for beta, uniform prior for Matern range parameter rho (see fields::Matern), and for phi, we use prior that induces half-Cauchy prior on the standard deviation of u.

"priors" is a named list with the following possible elements:

<b>beta_intercept_scale</b>	scale of intercept parameter (default 10)
<b>beta_scale</b>	scale of other beta parameters (default 2.5)
<b>beta_df</b>	degrees of freedom for t prior on beta (default Inf, normal prior)
<b>rho_lb</b>	lower bound for range parameter rho (default min distance between coords)
<b>rho_ub</b>	upper bound for range parameter rho (default max distance between coords)
<b>logpriorphi</b>	function for log-prior on phi (default prior that induces half-Cauchy prior on the standard deviation of u.)

## Value

Returns list of

<b>post_save</b>	a matrix of posterior samples (coda::mcmc) with nsave rows
<b>u_save</b>	a matrix of posterior samples (coda::mcmc) of random effects, with nsave rows
<b>betam_save</b>	a matrix of posterior samples (coda::mcmc) of population-averaged log odds, with nsave rows
<b>loglik_save</b>	a nsave x n matrix of pointwise log-likelihood values, can be used for WAIC calculation.
<b>priors</b>	list of hyperprior information
<b>nsave</b>	number of MCMC samples
<b>t_mcmc</b>	wall-clock time for running MCMC
<b>t_premcmc</b>	wall-clock time for preprocessing before MCMC
<b>y</b>	response vector
<b>X</b>	fixed effect design matrix
<b>coords</b>	a n x 2 matrix of Euclidean coordinates

## Examples

```

library(spbridge)
data(gambia)
N = length(gambia$pos) # 2035
y = gambia$pos # binary response, N = 2035 by 1 vector
# define id based on spatial coords unique values
id = as.numeric(factor(paste(gambia$x, gambia$y)))
n = length(unique(id)) # 65
coords = unique(cbind(gambia$x, gambia$y)/1000) # n by 2 matrix, in km
# standardized covariates following Gelman et al (2008)
intercept = rep(1,N) # intercept
age = scale(gambia$age/365, scale = 2*sd(gambia$age/365)) # in years
netuse = gambia$netuse - mean(gambia$netuse)
treated = gambia$treated - mean(gambia$treated)
green = scale(gambia$green, scale = 2*sd(gambia$green))
green2 = scale(gambia$green^2, scale = 2*sd(gambia$green^2))
healthctr = gambia$phc - mean(gambia$phc)

X = cbind(intercept, age, netuse, treated, green, green2, healthctr)
colnames(X) = c("(Intercept)", "age", "netuse", "treated", "green", "green2", "healthctr")

centers = c(attr(age, "scaled:center"), mean(gambia$netuse), mean(gambia$treated),
           attr(green, "scaled:center"), attr(green2, "scaled:center"), mean(gambia$phc))
scales = c(attr(age, "scaled:scale"), 1, 1, attr(green, "scaled:scale"), attr(green2, "scaled:scale"), 1)

#
# commented out due to a long run time
#
# fit_bridge_fb = spbridge::splogi_bridge_fb(y = y,
#                                              X = X,
#                                              id = id,
#                                              priors = list(beta_intercept_scale = 10,
#                                                            beta_scale = 2.5, beta_df = Inf,
#                                                            rho_lb = 0.01, rho_ub = 100),
#                                              coords = coords,
#                                              smoothness = 0.5, nburn = 1000, nsave = 10000, nthin = 1)

```

## splogi\_bridge\_fblowrank

*Spatial logistic model with bridge process random effect with low-rank structure (fully Bayes)*

## Description

$$\text{logit} [\Pr(y_{ij} = 1 \mid X_{ij}, u(s_i))] = X_{ij}^\top \beta + u(s_i)$$

where  $u(s) \sim \text{Bridge process}$  with parameter  $\phi$  and Matern correlation kernel with low-rank structure,  $i=1,\dots,n$  corresponds to  $n$  spatial locations and  $j=1,\dots, N_i$  correspond to  $n_i$  responses at location  $i$ , resulting data of size  $N = N_1 + \dots + N_n$ . The parameter  $\phi$  is estimated using fully Bayesian approach by placing prior on it.

## Usage

```
splogi_bridge_fblowrank(
  y,
  X,
  id,
  coords,
  coords_knot,
  priors = list(beta_intercept_scale = 10, beta_scale = 2.5, beta_df = Inf, logpriorphi =
    NULL, rho_lb = NULL, rho_ub = NULL),
  smoothness = 1.5,
  nburn = 100,
  nsave = 1000,
  nthin = 1,
  nparticle = 20,
  verbose = TRUE
)
```

## Arguments

y	N x 1 binary vector
X	N x p fixed-effect design matrix, including intercept
id	N x 1 vector of spatial location id. When N=n, it is point-referenced data
coords	n x 2 matrix of spatial coordinates
coords_knot	q x 2 matrix of knot coordinates
priors	list of prior hyperparameters, see details
smoothness	positive numeric, Matern smoothness parameter
nburn	number of burn-in iteration
nsave	number of posterior samples
nthin	thin-in rate
nparticle	number of particles in particle marginal Metropolis-Hastings
verbose	logical, whether to print progress

## Details

Priors are specified by "priors" argument, which is a list of hyperparameters. Specifically, we set zero centered normal or t prior for beta, uniform prior for Matern range parameter rho (see fields::Matern), and for phi, we use prior that induces half-Cauchy prior on the standard deviation of u.

priors is a named list with the following possible elements:

- beta\_intercept\_scale** scale of intercept parameter (default 10)
- beta\_scale** scale of other beta parameters (default 2.5)
- beta\_df** degrees of freedom for t prior on beta (default Inf, normal prior)
- logpriorphi** function for log-prior on phi (default prior that induces half-Cauchy prior on the standard deviation of u.)
- rho\_lb** lower bound for range parameter rho (default min distance between coords)
- rho\_ub** upper bound for range parameter rho (default max distance between coords)



```

# id = id,
# priors = list(beta_intercept_scale = 10,
#                 beta_scale = 2.5, beta_df = Inf,
#                 rho_lb = 0.01, rho_ub = 100),
# coords = coords,
# coords_knot = coords_knot,
# smoothness = 0.5, nburn = 1000, nsave = 10000, nthin = 1)

```

**splogi\_bridge\_lowrank** *Spatial logistic model with bridge process random effect with low-rank structure (empirical Bayes)*

## Description

$$\text{logit} [\Pr(y_{ij} = 1 \mid X_{ij}, u(s_i))] = X_{ij}^\top \beta + u(s_i)$$

where  $u(s) \sim \text{Bridge process}$  with parameter  $\phi$  and Matern correlation kernel with low-rank structure,  $i=1,\dots,n$  corresponds to  $n$  spatial locations and  $j=1,\dots, N_i$  correspond to  $n_i$  responses at location  $i$ , resulting data of size  $N = N_1 + \dots + N_n$ . The parameter  $\phi$  is estimated using empirical Bayes approach.

## Usage

```

splogi_bridge_lowrank(
  y,
  X,
  id,
  coords,
  coords_knot,
  priors = list(beta_intercept_scale = 10, beta_scale = 2.5, beta_df = Inf, logpriorphi =
    NULL, rho_lb = NULL, rho_ub = NULL),
  smoothness = 1.5,
  nburn = 100,
  nsave = 1000,
  nthin = 1,
  verbose = TRUE
)

```

## Arguments

y	$N \times 1$ binary vector
X	$N \times p$ fixed-effect design matrix, including intercept
id	$N \times 1$ vector of spatial location id. When $N=n$ , it is point-referenced data
coords	$n \times 2$ matrix of spatial coordinates
coords_knot	$q \times 2$ matrix of knot coordinates
priors	list of prior hyperparameters, see details
smoothness	positive numeric, Matern smoothness parameter

nburn	number of burn-in iteration
nsave	number of posterior samples
nthin	thin-in rate
verbose	logical, whether to print progress

## Details

Priors are specified by "priors" argument, which is a list of hyperparameters. Specifically, we set zero centered normal or t prior for beta, uniform prior for Matern range parameter rho (see fields::Matern), and for phi, we use prior that induces half-Cauchy prior on the standard deviation of u.

priors is a named list with the following possible elements:

- beta\_intercept\_scale** scale of intercept parameter (default 10)
- beta\_scale** scale of other beta parameters (default 2.5)
- beta\_df** degrees of freedom for t prior on beta (default Inf, normal prior)
- rho\_lb** lower bound for range parameter rho (default min distance between coords)
- rho\_ub** upper bound for range parameter rho (default max distance between coords)

## Value

Returns list of

post_save	a matrix of posterior samples (coda::mcmc) with nsave rows
u_save	a matrix of posterior samples (coda::mcmc) of random effects, with nsave rows
betam_save	a matrix of posterior samples (coda::mcmc) of population-averaged log odds, with nsave rows
loglik_save	a nsave x n matrix of pointwise log-likelihood values, can be used for WAIC calculation.
priors	list of hyperprior information
nsave	number of MCMC samples
t_mcmc	wall-clock time for running MCMC
t_premcmc	wall-clock time for preprocessing before MCMC
y	response vector
X	fixed effect design matrix
coords	a n x 2 matrix of Euclidean coordinates
coords_knot	a q x 2 matrix of knot coordinates

## Examples

```
library(spbbridge)
data(gambia)
N = length(gambia$pos) # 2035
y = gambia$pos # binary response, N = 2035 by 1 vector
# define id based on spatial coords unique values
id = as.numeric(factor(paste(gambia$x, gambia$y)))
n = length(unique(id)) # 65
coords = unique(cbind(gambia$x, gambia$y)/1000) # n by 2 matrix, in km
# standardized covariates following Gelman et al (2008)
```

```

intercept = rep(1,N) # intercept
age = scale(gambia$age/365, scale = 2*sd(gambia$age/365)) # in years
netuse = gambia$netuse - mean(gambia$netuse)
treated = gambia$treated - mean(gambia$treated)
green = scale(gambia$green, scale = 2*sd(gambia$green))
green2 = scale(gambia$green^2, scale = 2*sd(gambia$green^2))
healthctr = gambia$phc - mean(gambia$phc)

X = cbind(intercept, age, netuse, treated, green, green2, healthctr)
colnames(X) = c("(Intercept)", "age", "netuse", "treated", "green", "green2", "healthctr")

centers = c(attr(age, "scaled:center"), mean(gambia$netuse), mean(gambia$treated),
           attr(green, "scaled:center"), attr(green2, "scaled:center"), mean(gambia$phc))
scales = c(attr(age, "scaled:scale"), 1, 1, attr(green, "scaled:scale"), attr(green2, "scaled:scale"), 1)

coords_knot = expand.grid(seq(quantile(coords[,1], 0.1), quantile(coords[,1], 0.9), length.out = 5),
                           seq(quantile(coords[,2], 0.1), quantile(coords[,2], 0.9), length.out = 5))

#
# uncomment below
#
# fit_bridge_lowrank = spbridge::splogi_bridge_lowrank(y = y,
#                                                       X = X,
#                                                       id = id,
#                                                       priors = list(beta_intercept_scale = 10,
#                                                       beta_scale = 2.5, beta_df = Inf,
#                                                       rho_lb = 0.01, rho_ub = 100),
#                                                       coords = coords,
#                                                       coords_knot = coords_knot,
#                                                       smoothness = 0.5, nburn = 1000, nsave = 10000, nthin = 1)
#

```

## splogi\_gaussian

*Spatial logistic model with Gaussian process random effect***Description**

$$\text{logit} [\Pr(y_{ij} = 1 \mid X_{ij}, u(s_i))] = X_{ij}^\top \beta + u(s_i)$$

where  $u(s) \sim$  Gaussian process with marginal variance  $\text{sigu2}$  and Matern correlation kernel,  $i=1,\dots,n$  corresponds to  $n$  spatial locations and  $j=1,\dots, N_i$  correspond to  $n_i$  responses at location  $i$ , resulting data of size  $N = N_1 + \dots + N_n$ . The parameter  $\text{sigu2}$  is estimated using fully Bayesian approach by placing prior on it.

**Usage**

```

splogi_gaussian(
  y,
  X,
  id,
  coords,
  priors = list(beta_intercept_scale = 10, beta_scale = 2.5, beta_df = Inf, logpriorsigu2
  = NULL, rho_lb = NULL, rho_ub = NULL),

```

```

smoothness = 1.5,
nburn = 100,
nsave = 1000,
nthin = 1,
verbose = TRUE
)

```

## Arguments

y	N x 1 binary vector
x	N x p fixed-effect design matrix, including intercept
id	N x 1 vector of spatial location id. When N=n, it is point-referenced data
coords	n x 2 matrix of spatial coordinates
priors	list of prior hyperparameters, see details
smoothness	positive numeric, Matern smoothness parameter
nburn	number of burn-in iteration
nsave	number of posterior samples
nthin	thin-in rate
verbose	logical, whether to print progress

## Details

Priors are specified by "priors" argument, which is a list of hyperparameters. Specifically, we set zero centered normal or t prior for beta, uniform prior for Matern range parameter rho (see fields::Matern), and for sigu2, we use prior that induces half-Cauchy prior on the standard deviation of u.

"priors" is a named list with the following possible elements:

- beta\_intercept\_scale** scale of intercept parameter (default 10)
- beta\_scale** scale of other beta parameters (default 2.5)
- beta\_df** degrees of freedom for t prior on beta (default Inf, normal prior)
- logpriorsigu2** function for log-prior on sigu2 (default prior that induces half-Cauchy prior on the standard deviation of u.)
- rho\_lb** lower bound for range parameter rho (default min distance between coords)
- rho\_ub** upper bound for range parameter rho (default max distance between coords)

## Value

Returns list of

post_save	a matrix of posterior samples (coda::mcmc) with nsave rows
u_save	a matrix of posterior samples (coda::mcmc) of random effects, with nsave rows
loglik_save	a nsave x n matrix of pointwise log-likelihood values, can be used for WAIC calculation.
priors	list of hyperprior information
nsave	number of MCMC samples
t_mcmc	wall-clock time for running MCMC

t_premcmc	wall-clock time for preprocessing before MCMC
y	response vector
X	fixed effect design matrix
coords	a n x 2 matrix of Euclidean coordinates

## Examples

```

library(spbridge)
data(gambia)
N = length(gambia$pos) # 2035
y = gambia$pos # binary response, N = 2035 by 1 vector
# define id based on spatial coords unique values
id = as.numeric(factor(paste(gambia$x, gambia$y)))
n = length(unique(id)) # 65
coords = unique(cbind(gambia$x, gambia$y)/1000) # n by 2 matrix, in km
# standardized covariates following Gelman et al (2008)
intercept = rep(1,N) # intercept
age = scale(gambia$age/365, scale = 2*sd(gambia$age/365)) # in years
netuse = gambia$netuse - mean(gambia$netuse)
treated = gambia$treated - mean(gambia$treated)
green = scale(gambia$green, scale = 2*sd(gambia$green))
green2 = scale(gambia$green^2, scale = 2*sd(gambia$green^2))
healthctr = gambia$phc - mean(gambia$phc)

X = cbind(intercept, age, netuse, treated, green, green2, healthctr)
colnames(X) = c("(Intercept)", "age", "netuse", "treated", "green", "green2", "healthctr")

centers = c(attr(age, "scaled:center"), mean(gambia$netuse), mean(gambia$treated),
           attr(green, "scaled:center"), attr(green2, "scaled:center"), mean(gambia$phc))
scales = c(attr(age, "scaled:scale"), 1, 1, attr(green, "scaled:scale"), attr(green2, "scaled:scale"), 1)

#
# please uncomment
#
# fit_gaussian = spbridge:::splogi_gaussian(y = y,
#                                             X = X,
#                                             id = id,
#                                             priors = list(beta_intercept_scale = 10,
#                                                           beta_scale = 2.5, beta_df = Inf,
#                                                           rho_lb = 0.01, rho_ub = 100),
#                                             coords = coords,
#                                             smoothness = 0.5, nburn = 1000, nsave = 10000, nthin = 1)

```

## Description

$$\text{logit} [\Pr(y_{ij} = 1 \mid X_{ij}, u(s_i))] = X_{ij}^\top \beta + u(s_i)$$

where  $u(s) \sim \text{Gaussian process}$  with marginal variance  $\text{sigu2}$  and Matern correlation kernel with low-rank structure,  $i=1,\dots,n$  corresponds to  $n$  spatial locations and  $j=1,\dots, N_i$  correspond to  $n_i$  responses at location  $i$ , resulting data of size  $N = N_1 + \dots + N_n$ . The parameter  $\text{sigu2}$  is estimated using fully Bayesian approach by placing prior on it.

## Usage

```
splogi_gaussian_lowrank(
  y,
  X,
  id,
  coords,
  coords_knot = NULL,
  priors = list(beta_intercept_scale = 10, beta_scale = 2.5, beta_df = Inf, logpriorsigu2
    = NULL, rho_lb = NULL, rho_ub = NULL),
  smoothness = 1.5,
  nburn = 100,
  nsave = 1000,
  nthin = 1,
  verbose = TRUE
)
```

## Arguments

<code>y</code>	$N \times 1$ binary vector
<code>X</code>	$N \times p$ fixed-effect design matrix, including intercept
<code>id</code>	$N \times 1$ vector of spatial location id. When $N=n$ , it is point-referenced data
<code>coords</code>	$n \times 2$ matrix of spatial coordinates
<code>coords_knot</code>	$q \times 2$ matrix of knot coordinates
<code>priors</code>	list of prior hyperparameters, see details
<code>smoothness</code>	positive numeric, Matern smoothness parameter
<code>nburn</code>	number of burn-in iteration
<code>nsave</code>	number of posterior samples
<code>nthin</code>	thin-in rate
<code>verbose</code>	logical, whether to print progress

## Details

Priors are specified by "priors" argument, which is a list of hyperparameters. Specifically, we set zero centered normal or t prior for beta, uniform prior for Matern range parameter rho (see `fields::Matern`), and for `sigu2`, we use prior that induces half-Cauchy prior on the standard deviation of  $u$ .

`priors` is a named list with the following possible elements:

**`beta_intercept_scale`** scale of intercept parameter (default 10)

**`beta_scale`** scale of other beta parameters (default 2.5)

**beta\_df** degrees of freedom for t prior on beta (default Inf, normal prior)  
**logpriorsigu2** function for log-prior on sigu2 (default prior that induces half-Cauchy prior on the standard deviation of u.)  
**rho\_lb** lower bound for range parameter rho (default min distance between coords)  
**rho\_ub** upper bound for range parameter rho (default max distance between coords)

### Value

Returns list of

post_save	a matrix of posterior samples (coda::mcmc) with nsave rows
u_save	a matrix of posterior samples (coda::mcmc) of random effects, with nsave rows
loglik_save	a nsave x n matrix of pointwise log-likelihood values, can be used for WAIC calculation.
priors	list of hyperprior information
nsave	number of MCMC samples
t_mcmc	wall-clock time for running MCMC
t_premcmc	wall-clock time for preprocessing before MCMC
y	response vector
X	fixed effect design matrix
coords	a n x 2 matrix of Euclidean coordinates
coords_knot	a q x 2 matrix of knot coordinates

### Examples

```
library(spbridge)
data(gambia)
N = length(gambia$pos) # 2035
y = gambia$pos # binary response, N = 2035 by 1 vector
# define id based on spatial coords unique values
id = as.numeric(factor(paste(gambia$x, gambia$y)))
n = length(unique(id)) # 65
coords = unique(cbind(gambia$x, gambia$y)/1000) # n by 2 matrix, in km
# standardized covariates following Gelman et al (2008)
intercept = rep(1,N) # intercept
age = scale(gambia$age/365, scale = 2*sd(gambia$age/365)) # in years
netuse = gambia$netuse - mean(gambia$netuse)
treated = gambia$treated - mean(gambia$treated)
green = scale(gambia$green, scale = 2*sd(gambia$green))
green2 = scale(gambia$green^2, scale = 2*sd(gambia$green^2))
healthctr = gambia$phc - mean(gambia$phc)

X = cbind(intercept, age, netuse, treated, green, green2, healthctr)
colnames(X) = c("(Intercept)", "age", "netuse", "treated", "green", "green2", "healthctr")

centers = c(attr(age, "scaled:center"), mean(gambia$netuse), mean(gambia$treated),
           attr(green, "scaled:center"), attr(green2, "scaled:center"), mean(gambia$phc))
scales = c(attr(age, "scaled:scale"), 1, 1, attr(green, "scaled:scale"), attr(green2, "scaled:scale"), 1)

coords_knot = expand.grid(seq(quantile(coords[,1], 0.1), quantile(coords[,1], 0.9), length.out = 5),
                           seq(quantile(coords[,2], 0.1), quantile(coords[,2], 0.9), length.out = 5))
```

```
#  
# uncomment below  
#  
# fit_gaussian_lowrank = spbridge::splogi_gaussian_lowrank(y = y,  
#                                         X = X,  
#                                         id = id,  
#                                         priors = list(beta_intercept_scale = 10,  
#                                         beta_scale = 2.5, beta_df = Inf,  
#                                         rho_lb = 0.01, rho_ub = 100),  
#                                         coords = coords,  
#                                         coords_knot = coords_knot,  
#                                         smoothness = 0.5, nburn = 1000, nsave = 10000, nthin = 1)
```

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