

CAUSAL INFERENCE WITH ORDINAL OUTCOMES

DENSITY ESTIMATION BASED APPROACH

CHANHYUK PARK
WASHINGTON UNIVERSITY IN ST. LOUIS

MOTIVATION

- Experiments to identify causal effects of treatments
- Outcomes are usually believed to be continuous in unidimensional space
 - Approval ratings (Canes-Wrone and De Marchi, 2002; Kriner and Schwartz, 2009)
 - Policy preferences (Scheve and Slaughter, 2001; Mayda and Rodrik, 2005; Wu, 2022)

MOTIVATION

- Experiments to identify causal effects of treatments
- Outcomes are usually believed to be continuous in unidimensional space
 - Approval ratings (Canes-Wrone and De Marchi, 2002; Kriner and Schwartz, 2009)
 - Policy preferences (Scheve and Slaughter, 2001; Mayda and Rodrik, 2005; Wu, 2022)

The True Outcome (Y^*)



MOTIVATION

- Experiments to identify causal effects of treatments
- Outcomes are usually believed to be continuous in unidimensional space
 - Approval ratings (Canes-Wrone and De Marchi, 2002; Kriner and Schwartz, 2009)
 - Policy preferences (Scheve and Slaughter, 2001; Mayda and Rodrik, 2005; Wu, 2022)

The True Outcome (Y^*)



Strongly Disagree

Disagree

Neither

Agree

Strongly Agree

MOTIVATION

- Experiments to identify causal effects of treatments
- Outcomes are usually believed to be continuous in unidimensional space
 - Approval ratings (Canes-Wrone and De Marchi, 2002; Kriner and Schwartz, 2009)
 - Policy preferences (Scheve and Slaughter, 2001; Mayda and Rodrik, 2005; Wu, 2022)



MOTIVATION

- Experiments to identify causal effects of treatments
- Outcomes are usually believed to be continuous in unidimensional space
 - Approval ratings (Canes-Wrone and De Marchi, 2002; Kriner and Schwartz, 2009)
 - Policy preferences (Scheve and Slaughter, 2001; Mayda and Rodrik, 2005; Wu, 2022)

The True Outcome (Y^*)



THE GOAL

- *Identification*
- *Estimation*

THE GOAL

- *Identification*

Naive causal identification with ordinal outcome "index" fails

====> Alternative estimand: **Normalized Latent Treatment Effect**

- *Estimation*

THE GOAL

- *Identification*

Naive causal identification with ordinal outcome "index" fails

====> Alternative estimand: **Normalized Latent Treatment Effect**

- *Estimation*

Parametric ordered probit and logit require strong distributional assumptions

====> **Flexible, density estimation** based estimators

WHY NAIVE IDENTIFICATION FAILS

The True Outcome ($Y_i^* = \tau D_i + \varepsilon_i$)



WHY NAIVE IDENTIFICATION FAILS

The True Outcome ($Y_i^* = \tau D_i + \varepsilon_i$)



	$Y^*(0)$	$Y^*(1)$
A	0.8	1.2
B	0.4	0.6
C	1.9	2.2
D	1.5	1.7
E	2.5	2.4

WHY NAIVE IDENTIFICATION FAILS

The True Outcome ($Y_i^* = \tau D_i + \varepsilon_i$)



	$Y^*(0)$	$Y^*(1)$
A	0.8	1.2
B	0.4	0.6
C	1.9	2.2
D	1.5	1.7
E	2.5	2.4

WHY NAIVE IDENTIFICATION FAILS

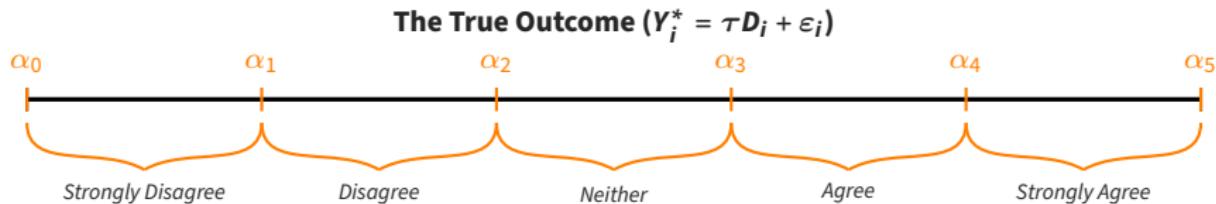
The True Outcome ($Y_i^* = \tau D_i + \varepsilon_i$)



	$Y^*(0)$	$Y^*(1)$
A	0.8	1.2
B	0.4	0.6
C	1.9	2.2
D	1.5	1.7
E	2.5	2.4

- $\tau = \mathbb{E}[Y^*(1)] - \mathbb{E}[Y^*(0)] \approx +0.2$

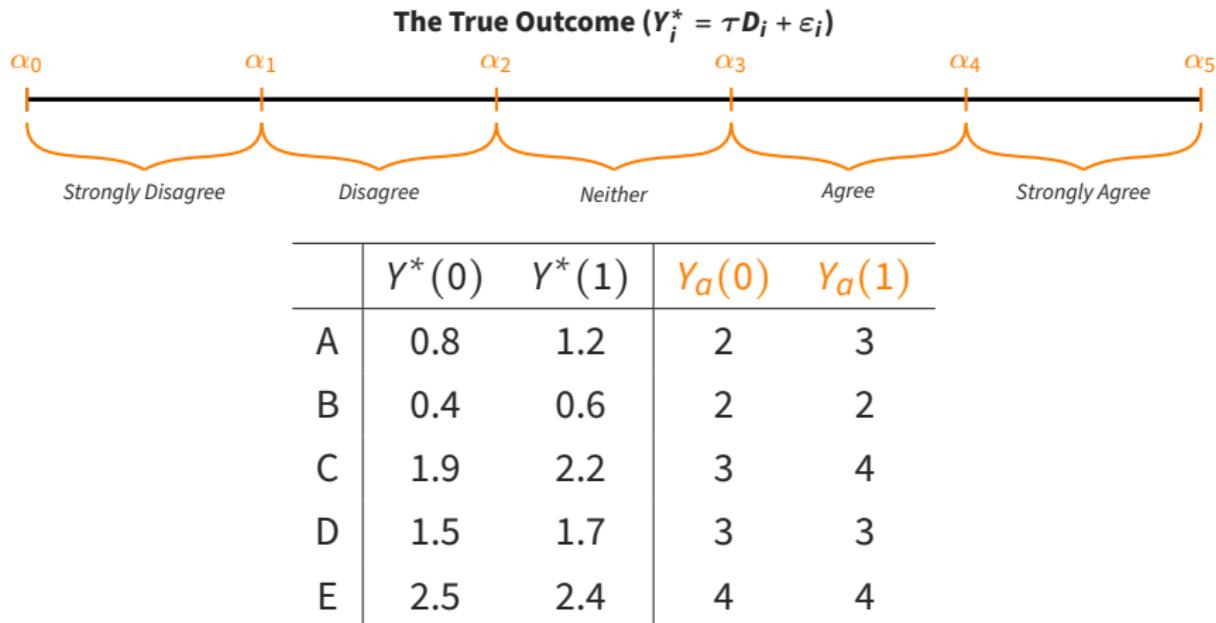
WHY NAIVE IDENTIFICATION FAILS



	$Y^*(0)$	$Y^*(1)$	$Y_a(0)$	$Y_a(1)$
A	0.8	1.2	Disagree	Neither
B	0.4	0.6	Disagree	Disagree
C	1.9	2.2	Neither	Agree
D	1.5	1.7	Neither	Neither
E	2.5	2.4	Agree	Agree

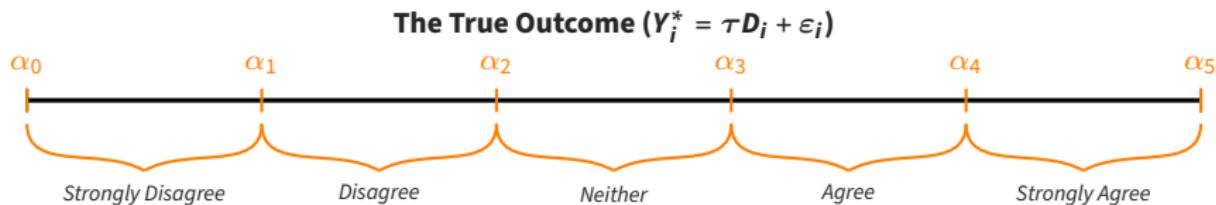
- $\tau = \mathbb{E}[Y^*(1)] - \mathbb{E}[Y^*(0)] \approx +0.2$

WHY NAIVE IDENTIFICATION FAILS



- $\tau = \mathbb{E}[Y^*(1)] - \mathbb{E}[Y^*(0)] \approx +0.2$

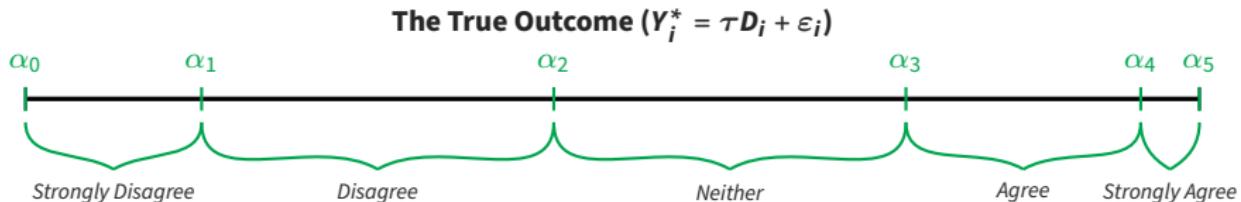
WHY NAIVE IDENTIFICATION FAILS



	$Y^*(0)$	$Y^*(1)$	$Y_a(0)$	$Y_a(1)$
A	0.8	1.2	2	3
B	0.4	0.6	2	2
C	1.9	2.2	3	4
D	1.5	1.7	3	3
E	2.5	2.4	4	4

- $\tau = \mathbb{E}[Y^*(1)] - \mathbb{E}[Y^*(0)] \approx +0.2$
- $\tau_a = \mathbb{E}[Y_a(1)] - \mathbb{E}[Y_a(0)] \approx +0.4$

WHY NAIVE IDENTIFICATION FAILS

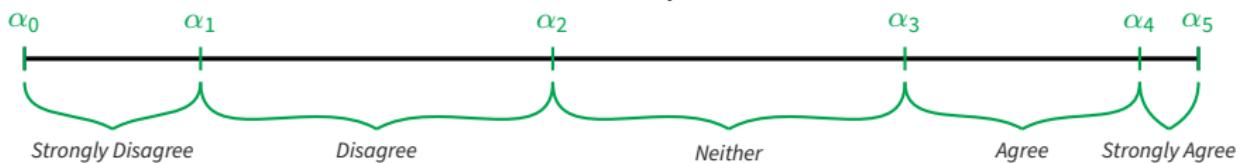


	$Y^*(0)$	$Y^*(1)$	$Y_b(0)$	$Y_b(1)$
A	0.8	1.2	Disagree	Disagree
B	0.4	0.6	Disagree	Disagree
C	1.9	2.2	Neither	Neither
D	1.5	1.7	Neither	Neither
E	2.5	2.4	Agree	Neither

- $\tau = \mathbb{E}[Y^*(1)] - \mathbb{E}[Y^*(0)] \approx +0.2$

WHY NAIVE IDENTIFICATION FAILS

The True Outcome ($Y_i^* = \tau D_i + \varepsilon_i$)



	$Y^*(0)$	$Y^*(1)$	$Y_b(0)$	$Y_b(1)$
A	0.8	1.2	2	2
B	0.4	0.6	2	2
C	1.9	2.2	3	3
D	1.5	1.7	3	3
E	2.5	2.4	4	3

- $\tau = \mathbb{E}[Y^*(1)] - \mathbb{E}[Y^*(0)] \approx +0.2$
- $\tau_b = \mathbb{E}[Y_b(1)] - \mathbb{E}[Y_b(0)] \approx -0.2$

IDENTIFICATION THROUGH LINK FUNCTION

- Cumulative probabilities such as $\mathbb{P}(Y \leq j)$
- If we can map this probability to the Y^* space
- Link Function = CDF of ε_i (F_{ε_i}) does this:

$$F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i \leq j)) = \underbrace{\alpha_j - \tau D_i}_{\text{Defined in } Y^* \text{ space}}$$

- Key Assumption: Distribution of ε_i is known

$$\underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(0) \leq j))}_{\text{Control}} - \underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(1) \leq j))}_{\text{Treated}} = \tau$$

IDENTIFICATION THROUGH LINK FUNCTION

- Cumulative probabilities such as $\mathbb{P}(Y \leq j)$
- If we can map this probability to the Y^* space
- Link Function = CDF of ε_i (F_{ε_i}) does this:

$$F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i \leq j)) = \underbrace{\alpha_j - \tau D_i}_{\text{Defined in } Y^* \text{ space}}$$

- Key Assumption: Distribution of ε_i is known

$$\underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(0) \leq j))}_{\text{Control}} - \underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(1) \leq j))}_{\text{Treated}} = \tau$$

IDENTIFICATION THROUGH LINK FUNCTION

- Cumulative probabilities such as $\mathbb{P}(Y \leq j)$
- If we can map this probability to the Y^* space
- Link Function = CDF of ε_i (F_{ε_i}) does this:

$$F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i \leq j)) = \underbrace{\alpha_j - \tau D_i}_{\text{Defined in } Y^* \text{ space}}$$

- Key Assumption: Distribution of ε_i is known

$$\underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(0) \leq j))}_{\text{Control}} - \underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(1) \leq j))}_{\text{Treated}} = \tau$$

IDENTIFICATION THROUGH LINK FUNCTION

- Cumulative probabilities such as $\mathbb{P}(Y \leq j)$
- If we can map this probability to the Y^* space
- Link Function = CDF of ε_i (F_{ε_i}) does this:

$$F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i \leq j)) = \underbrace{\alpha_j - \tau D_i}_{\text{Defined in } Y^* \text{ space}}$$

- Key Assumption: Distribution of ε_i is known

$$\underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(0) \leq j))}_{\text{Control}} - \underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(1) \leq j))}_{\text{Treated}} = \tau$$

IDENTIFICATION THROUGH LINK FUNCTION

- Cumulative probabilities such as $\mathbb{P}(Y \leq j)$
- If we can map this probability to the Y^* space
- Link Function = CDF of ε_i (F_{ε_i}) does this:

$$F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i \leq j)) = \underbrace{\alpha_j - \tau D_i}_{\text{Defined in } Y^* \text{ space}}$$

- Key Assumption: Distribution of ε_i is known

$$\underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(0) \leq j))}_{\text{Control}} - \underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(1) \leq j))}_{\text{Treated}} = \tau$$

NEW ESTIMANDS: NORMALIZED TREATMENT EFFECT

- One caveat: for a positive constant C ,

$$\begin{aligned}\mathbb{P}(Y_i^* \leq \alpha_4) &= \mathbb{P}(C \cdot Y_i^* \leq C \cdot \alpha_4) \\ &= \mathbb{P}(C \cdot \tau D_i + C \cdot \varepsilon_i \leq C \cdot \alpha_4)\end{aligned}$$

- Thus τ becomes indistinguishable from $C \cdot \tau$
- Normalization to cancel out C

② Scale to fix $\sigma_{\varepsilon_i} = 1$

③ If there are multiple treatments, fixing $\tau_a = 1$

NEW ESTIMANDS: NORMALIZED TREATMENT EFFECT

- One caveat: for a positive constant C ,

$$\begin{aligned}\mathbb{P}(Y_i^* \leq \alpha_4) &= \mathbb{P}(C \cdot Y_i^* \leq C \cdot \alpha_4) \\ &= \mathbb{P}(C \cdot \tau D_i + C \cdot \varepsilon_i \leq C \cdot \alpha_4)\end{aligned}$$

- Thus τ becomes indistinguishable from $C \cdot \tau$
- Normalization to cancel out C
 - ② Scale to fix $\sigma_{\varepsilon_i} = 1$
 - ③ If there are multiple treatments, fixing $\tau_d = 1$

NEW ESTIMANDS: NORMALIZED TREATMENT EFFECT

- One caveat: for a positive constant C ,

$$\begin{aligned}\mathbb{P}(Y_i^* \leq \alpha_4) &= \mathbb{P}(C \cdot Y_i^* \leq C \cdot \alpha_4) \\ &= \mathbb{P}(C \cdot \tau D_i + C \cdot \varepsilon_i \leq C \cdot \alpha_4)\end{aligned}$$

- Thus τ becomes indistinguishable from $C \cdot \tau$
- Normalization to cancel out C

- ① Scale to fix $\sigma_{\varepsilon_j} = 1$
- ② If there are multiple treatments, fixing $\tau_a = 1$

NEW ESTIMANDS: NORMALIZED TREATMENT EFFECT

- One caveat: for a positive constant C ,

$$\begin{aligned}\mathbb{P}(Y_i^* \leq \alpha_4) &= \mathbb{P}(C \cdot Y_i^* \leq C \cdot \alpha_4) \\ &= \mathbb{P}(C \cdot \tau D_i + C \cdot \varepsilon_i \leq C \cdot \alpha_4)\end{aligned}$$

- Thus τ becomes indistinguishable from $C \cdot \tau$
- Normalization to cancel out C
 - ① Scale to fix $\sigma_{\varepsilon_i} = 1$
 - ② If there are multiple treatments, fixing $\tau_a = 1$

NEW ESTIMANDS: NORMALIZED TREATMENT EFFECT

- One caveat: for a positive constant C ,

$$\begin{aligned}\mathbb{P}(Y_i^* \leq \alpha_4) &= \mathbb{P}(C \cdot Y_i^* \leq C \cdot \alpha_4) \\ &= \mathbb{P}(C \cdot \tau D_i + C \cdot \varepsilon_i \leq C \cdot \alpha_4)\end{aligned}$$

- Thus τ becomes indistinguishable from $C \cdot \tau$
- Normalization to cancel out C
 - ① Scale to fix $\sigma_{\varepsilon_i} = 1$
 - ② If there are multiple treatments, fixing $\tau_a = 1$

NEW ESTIMANDS: NORMALIZED TREATMENT EFFECT

- One caveat: for a positive constant C ,

$$\begin{aligned}\mathbb{P}(Y_i^* \leq \alpha_4) &= \mathbb{P}(C \cdot Y_i^* \leq C \cdot \alpha_4) \\ &= \mathbb{P}(C \cdot \tau D_i + C \cdot \varepsilon_i \leq C \cdot \alpha_4)\end{aligned}$$

- Thus τ becomes indistinguishable from $C \cdot \tau$
- Normalization to cancel out C

① Scale to fix $\sigma_{\varepsilon_i} = 1$

$$\implies \frac{C \cdot \tau}{C \cdot \sigma_{\varepsilon_i}} = \frac{\tau}{\sigma_{\varepsilon_i}} \text{ is identified}$$

② If there are multiple treatments, fixing $\tau_a = 1$

HOW CAN WE ESTIMATE?

- The key is **Link Function** (F_{ϵ_i})
- Ordered Probit, Ordered Logit, and IRT
- They assume F_{ϵ_i} is a specific distribution → MLE
- Distributional Assumption can be violated
 - Pure misspecification
 - Unaccounted covariates

\implies Become *inconsistent*
- Two Flexible Estimators without rigid Distributional Assumptions

HOW CAN WE ESTIMATE?

- The key is **Link Function** (F_{ϵ_i})
- Ordered Probit, Ordered Logit, and IRT
- They assume F_{ϵ_i} is a specific distribution → MLE
- Distributional Assumption can be violated
 - Pure misspecification
 - Unaccounted covariates

\implies Become *inconsistent*
- Two Flexible Estimators without rigid Distributional Assumptions

HOW CAN WE ESTIMATE?

- The key is **Link Function** (F_{ϵ_i})
- Ordered Probit, Ordered Logit, and IRT
- They assume F_{ϵ_i} is a specific distribution → MLE
- Distributional Assumption can be violated
 - Pure misspecification
 - Unaccounted covariates

 ➡ Become *inconsistent*
- Two Flexible Estimators without rigid Distributional Assumptions

HOW CAN WE ESTIMATE?

- The key is **Link Function** (F_{ϵ_i})
- Ordered Probit, Ordered Logit, and IRT
- They assume F_{ϵ_i} is a specific distribution → MLE
- Distributional Assumption can be violated
 - Pure misspecification
 - Unaccounted covariates

\implies Become *inconsistent*
- Two Flexible Estimators without rigid Distributional Assumptions

HOW CAN WE ESTIMATE?

- The key is **Link Function** (F_{ϵ_i})
- Ordered Probit, Ordered Logit, and IRT
- They assume F_{ϵ_i} is a specific distribution → MLE
- Distributional Assumption can be violated
 - Pure misspecification
 - Unaccounted covariates

\implies Become *inconsistent*
- Two Flexible Estimators without rigid Distributional Assumptions

HOW CAN WE ESTIMATE?

- The key is **Link Function** (F_{ϵ_i})
- Ordered Probit, Ordered Logit, and IRT
- They assume F_{ϵ_i} is a specific distribution → MLE
- Distributional Assumption can be violated
 - Pure misspecification
 - Unaccounted covariates

\implies Become *inconsistent*
- Two Flexible Estimators without rigid Distributional Assumptions

HOW CAN WE ESTIMATE?

- The key is **Link Function** (F_{ϵ_i})
- Ordered Probit, Ordered Logit, and IRT
- They assume F_{ϵ_i} is a specific distribution → MLE
- Distributional Assumption can be violated
 - Pure misspecification
 - Unaccounted covariates

\implies Become *inconsistent*
- Two Flexible Estimators without rigid Distributional Assumptions

HOW CAN WE ESTIMATE?

- The key is **Link Function** (F_{ϵ_i})
- Ordered Probit, Ordered Logit, and IRT
- They assume F_{ϵ_i} is a specific distribution → MLE
- Distributional Assumption can be violated
 - Pure misspecification
 - Unaccounted covariates

\implies Become *inconsistent*
- **Two Flexible Estimators** without rigid Distributional Assumptions

ALTERNATIVE: ESTIMATE THE DISTRIBUTION

- Estimated \hat{F}_{ε_i} from the Data
- Nonparametric method: Kernel Density Estimation
 - Use Kernels to smooth each observations
- Parametric Generative Model: Normalizing Flows
 - Transform complex distributions to simple one
- Same MLE

ALTERNATIVE: ESTIMATE THE DISTRIBUTION

- Estimated \hat{F}_{ε_i} from the Data
- Nonparametric method: **Kernel Density Estimation**
 - Use Kernels to smooth each observations
- Parametric Generative Model: **Normalizing Flows**
 - Transform complex distributions to simple one
- Same MLE

ALTERNATIVE: ESTIMATE THE DISTRIBUTION

- Estimated \hat{F}_{ε_i} from the Data
- Nonparametric method: **Kernel Density Estimation**
 - Use Kernels to smooth each observations
- Parametric Generative Model: **Normalizing Flows**
 - Transform complex distributions to simple one
- Same MLE

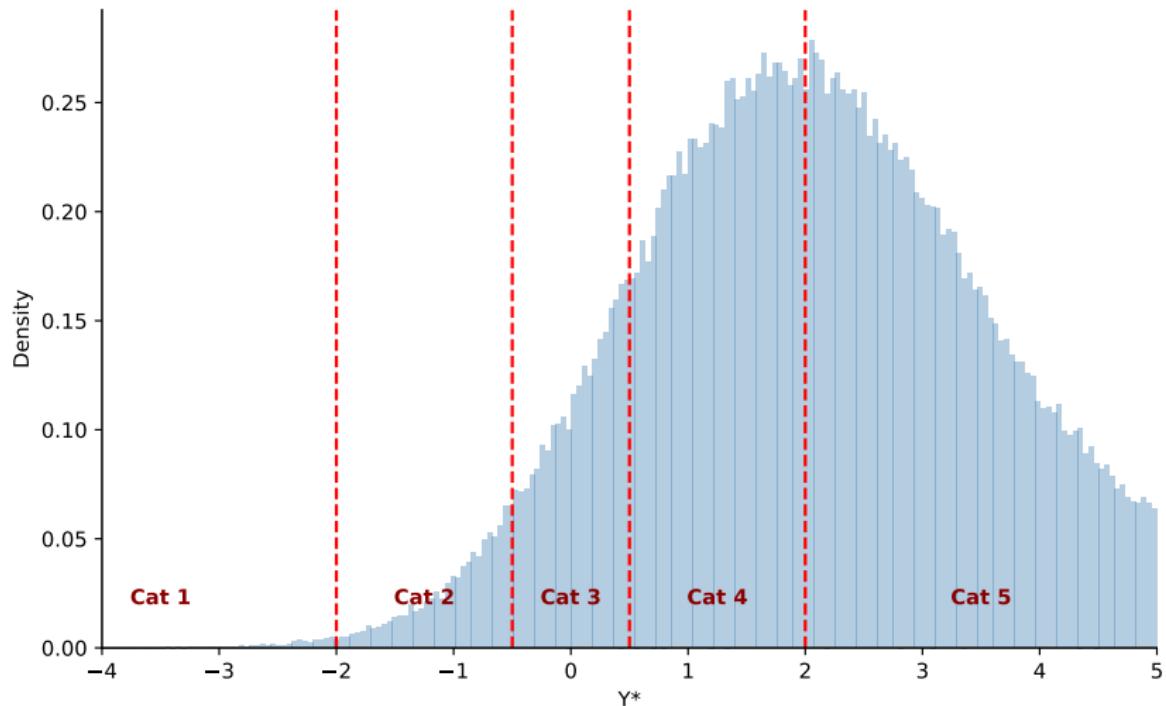
ALTERNATIVE: ESTIMATE THE DISTRIBUTION

- Estimated \hat{F}_{ε_i} from the Data
- Nonparametric method: **Kernel Density Estimation**
 - Use Kernels to smooth each observations
- Parametric Generative Model: **Normalizing Flows**
 - Transform complex distributions to simple one
- Same MLE

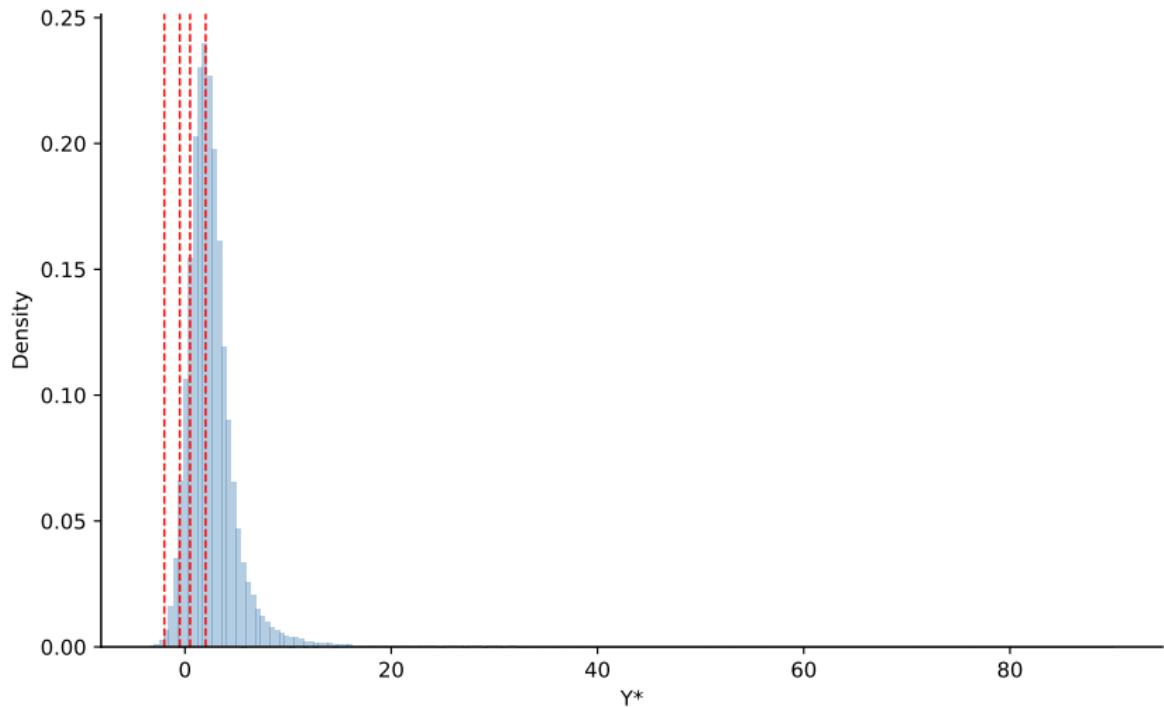
▶ Little More on Density Estimation

SIMULATION SETTING

SIMULATION SETTING

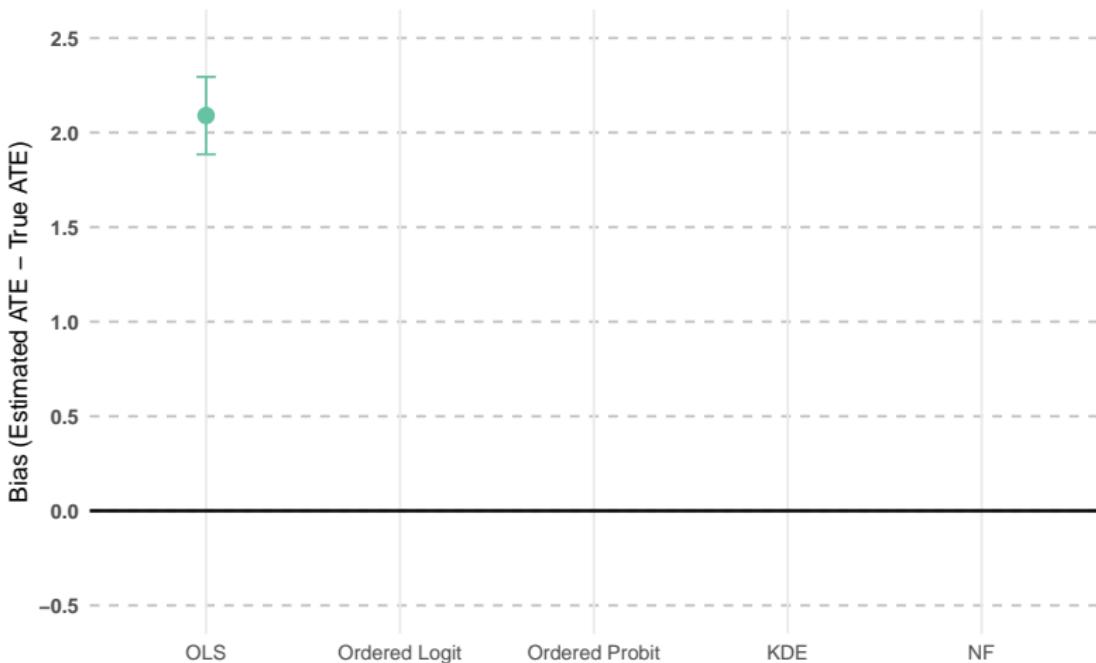


SIMULATION SETTING

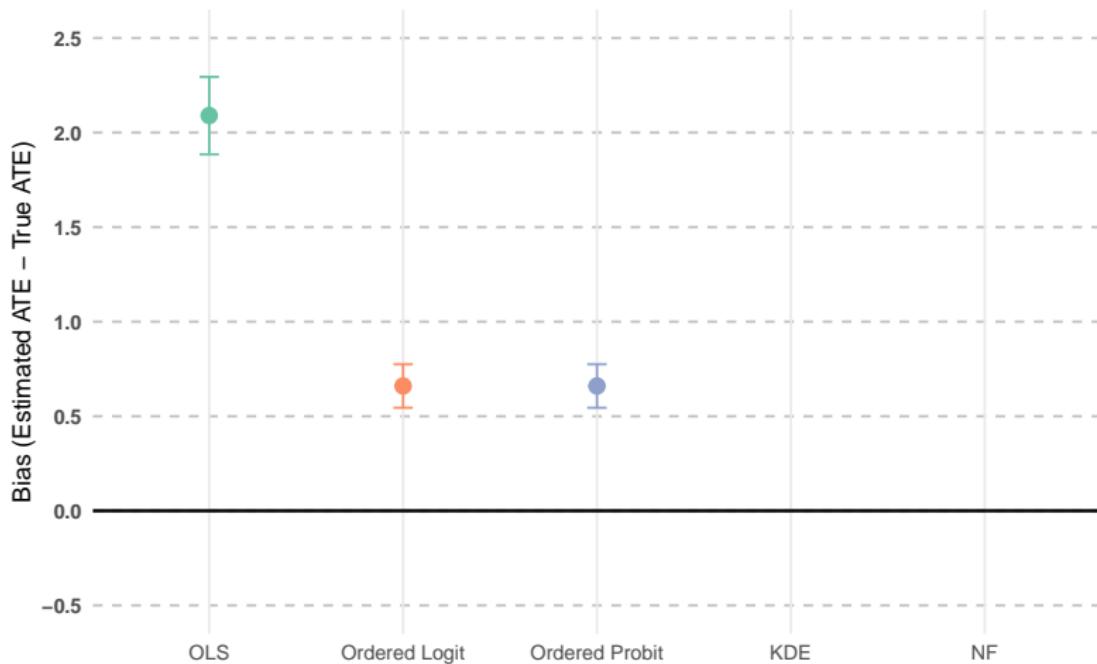


SIMULATION RESULTS: BIAS AND SIMULATION SD

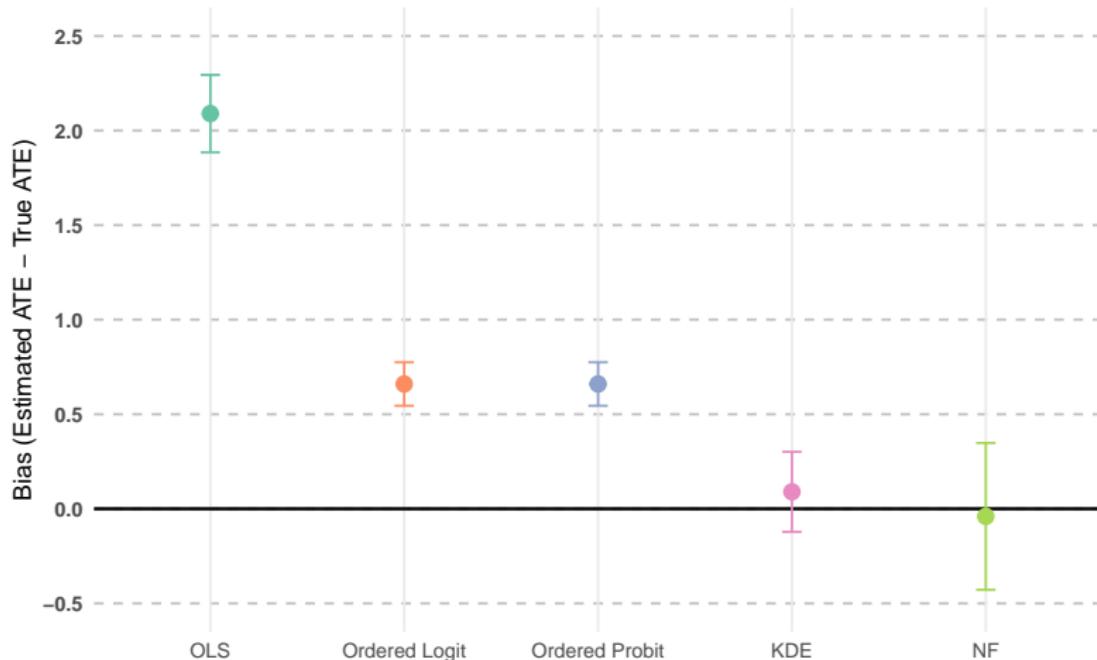
SIMULATION RESULTS: BIAS AND SIMULATION SD



SIMULATION RESULTS: BIAS AND SIMULATION SD



SIMULATION RESULTS: BIAS AND SIMULATION SD



REPLICATION – MATTINGLY ET AL. (2025)

- State-produced promotional media → preference on political system
- 19 different countries with $n = 6,000$
- Three treatments:
 - ① China: Two CCP produced videos
 - ② USA: Two US produced videos
 - ③ Competition: One from CCP and the other from US
- Outcome: Preference on Political System
 - Strongly Prefer the US to Strongly Prefer China*

REPLICATION – MATTINGLY ET AL. (2025)

- State-produced promotional media → preference on political system
- 19 different countries with $n = 6,000$
- Three treatments:
 - ① China: Two CCP produced videos
 - ② USA: Two US produced videos
 - ③ Competition: One from CCP and the other from US
- Outcome: Preference on Political System
 - Strongly Prefer the US to Strongly Prefer China*

REPLICATION – MATTINGLY ET AL. (2025)

- State-produced promotional media → preference on political system
- 19 different countries with $n = 6,000$
- Three treatments:
 - ① China: Two CCP produced videos
 - ② USA: Two US produced videos
 - ③ Competition: One from CCP and the other from US
- Outcome: Preference on Political System
Strongly Prefer the US to Strongly Prefer China

REPLICATION – MATTINGLY ET AL. (2025)

- State-produced promotional media → preference on political system
- 19 different countries with $n = 6,000$
- Three treatments:
 - ① China: Two CCP produced videos
 - ② USA: Two US produced videos
 - ③ Competition: One from CCP and the other from US
- Outcome: Preference on Political System

Strongly Prefer the US to Strongly Prefer China

REPLICATION – MATTINGLY ET AL. (2025)

- State-produced promotional media → preference on political system
- 19 different countries with $n = 6,000$
- Three treatments:
 - ① China: Two CCP produced videos
 - ② USA: Two US produced videos
 - ③ Competition: One from CCP and the other from US
- Outcome: Preference on Political System

Strongly Prefer the US to Strongly Prefer China

REPLICATION – MATTINGLY ET AL. (2025)

- State-produced promotional media → preference on political system
- 19 different countries with $n = 6,000$
- Three treatments:
 - ① China: Two CCP produced videos
 - ② USA: Two US produced videos
 - ③ Competition: One from CCP and the other from US
- Outcome: Preference on Political System
Strongly Prefer the US to Strongly Prefer China

REPLICATION – MATTINGLY ET AL. (2025)

- State-produced promotional media → preference on political system
- 19 different countries with $n = 6,000$
- Three treatments:
 - ① China: Two CCP produced videos
 - ② USA: Two US produced videos
 - ③ Competition: One from CCP and the other from US
- Outcome: Preference on Political System

Strongly Prefer the US to Strongly Prefer China

REPLICATION RESULTS

	ATE	NLTE			
	Original – OLS	Ordered Logit	Ordered Probit	KDE-based	NF-based
China	1.04*** (0.05)	0.73 *** (0.04)	0.76 *** (0.04)	0.36 *** (0.07)	0.37 *** (0.04)
USA	-0.43*** (0.04)	-0.35 *** (0.04)	-0.38 *** (0.04)	-0.18 *** (0.05)	-0.17 *** (0.03)
Competition	0.36*** (0.05)	0.21 *** (0.04)	0.25 *** (0.04)	0.13* (0.06)	0.11** (0.04)

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

REPLICATION RESULTS

	ATE	NLTE			
	Original – OLS	Ordered Logit	Ordered Probit	KDE-based	NF-based
China	1.04*** (0.05)	0.73*** (0.04)	0.76*** (0.04)	0.36*** (0.07)	0.37*** (0.04)
USA	-0.43*** (0.04)	-0.35*** (0.04)	-0.38*** (0.04)	-0.18*** (0.05)	-0.17*** (0.03)
Competition	0.36*** (0.05)	0.21*** (0.04)	0.25*** (0.04)	0.13* (0.06)	0.11** (0.04)

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

REPLICATION RESULTS

	ATE	NLTE			
	Original – OLS	Ordered Logit	Ordered Probit	KDE-based	NF-based
China	1.04*** (0.05)	0.73*** (0.04)	0.76*** (0.04)	0.36*** (0.07)	0.37*** (0.04)
USA	-0.43*** (0.04)	-0.35*** (0.04)	-0.38*** (0.04)	-0.18*** (0.05)	-0.17*** (0.03)
Competition	0.36*** (0.05)	0.21*** (0.04)	0.25*** (0.04)	0.13* (0.06)	0.11** (0.04)

*** $p < 0.001$; ** $p < 0.01$; * $p < 0.05$

CONCLUSION

- *Identification*
- *Estimation*

CONCLUSION

- *Identification*

Naive use of ordinal "index" can not give us what we want

====> Alternative Estimand: **Normalized Latent Treatment Effect**

- *Estimation*

CONCLUSION

- *Identification*

Naive use of ordinal "index" can not give us what we want

==> Alternative Estimand: **Normalized Latent Treatment Effect**

- *Estimation*

Parametric ordinal regressions require strong assumptions

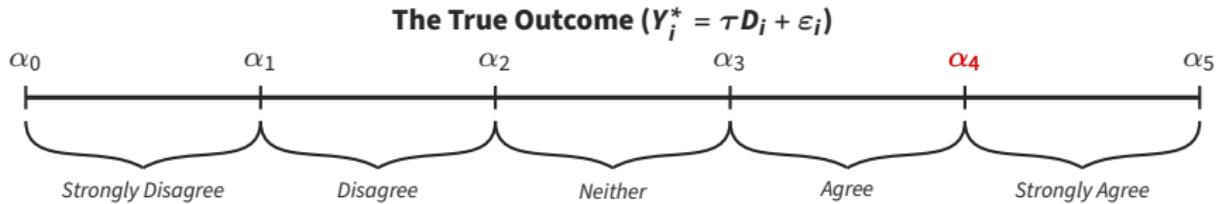
==> **KDE** or **Normalizing Flows** based estimators

MORE ON LINK FUNCTION

- Cumulative probabilities such as $\mathbb{P}(Y \leq \text{Agree})$
- Link function = CDF of the error term (F_{ε_i})
- Suppose we know F_{ε_i} :

$$\underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(0) \leq \text{Agree}))}_{\text{Control}} - \underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(1) \leq \text{Agree}))}_{\text{Treated}} = \tau$$

MORE ON LINK FUNCTION

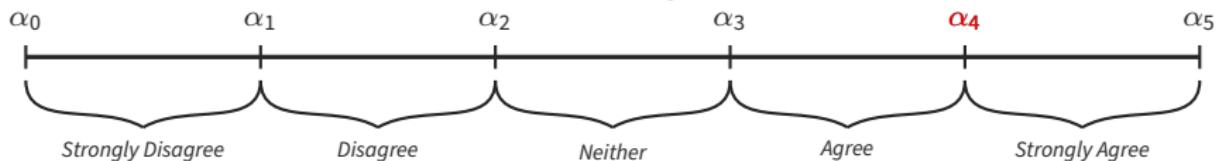


- Link function = CDF of the error term (F_{ε_i})

$$F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i \leq \text{Agree})) = F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i^* \leq \color{red}{\alpha_4}))$$

MORE ON LINK FUNCTION

The True Outcome ($Y_i^* = \tau D_i + \varepsilon_i$)

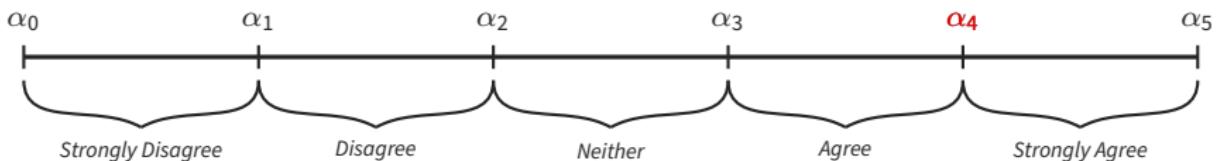


- Link function = CDF of the error term (F_{ε_i})

$$\begin{aligned} F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i \leq \text{Agree})) &= F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i^* \leq \alpha_4)) \\ &= F_{\varepsilon_i}^{-1}(\mathbb{P}(\tau D_i + \varepsilon_i \leq \alpha_4)) \end{aligned}$$

MORE ON LINK FUNCTION

The True Outcome ($Y_i^* = \tau D_i + \varepsilon_i$)

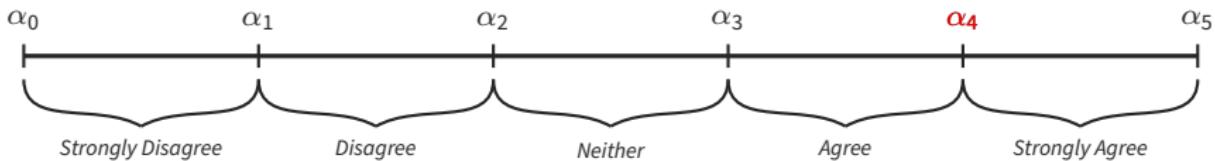


- Link function = CDF of the error term (F_{ε_i})

$$\begin{aligned} F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i \leq \text{Agree})) &= F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i^* \leq \alpha_4)) \\ &= F_{\varepsilon_i}^{-1}(\mathbb{P}(\tau D_i + \varepsilon_i \leq \alpha_4)) \\ &= F_{\varepsilon_i}^{-1}(\mathbb{P}(\varepsilon_i \leq \alpha_4 - \tau D_i)) \end{aligned}$$

MORE ON LINK FUNCTION

The True Outcome ($Y_i^* = \tau D_i + \varepsilon_i$)

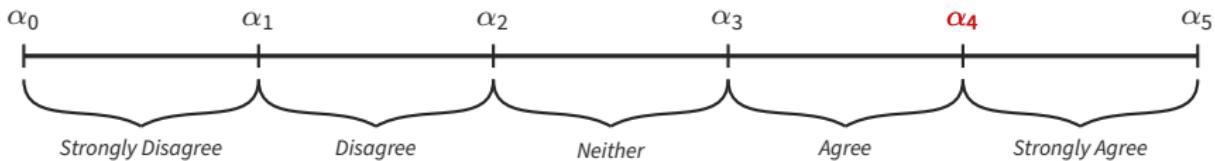


- Link function = CDF of the error term (F_{ε_i})

$$\begin{aligned} F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i \leq \text{Agree})) &= F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i^* \leq \alpha_4)) \\ &= F_{\varepsilon_i}^{-1}(\mathbb{P}(\tau D_i + \varepsilon_i \leq \alpha_4)) \\ &= F_{\varepsilon_i}^{-1}(\mathbb{P}(\varepsilon_i \leq \alpha_4 - \tau D_i)) \\ &= F_{\varepsilon_i}^{-1}(F_{\varepsilon_i}(\alpha_4 - \tau D_i)) \end{aligned}$$

MORE ON LINK FUNCTION

The True Outcome ($Y_i^* = \tau D_i + \varepsilon_i$)



- Link function = CDF of the error term (F_{ε_i})

$$\begin{aligned} F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i \leq \text{Agree})) &= F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i^* \leq \alpha_4)) \\ &= F_{\varepsilon_i}^{-1}(\mathbb{P}(\tau D_i + \varepsilon_i \leq \alpha_4)) \\ &= F_{\varepsilon_i}^{-1}(\mathbb{P}(\varepsilon_i \leq \alpha_4 - \tau D_i)) \\ &= F_{\varepsilon_i}^{-1}(F_{\varepsilon_i}(\alpha_4 - \tau D_i)) \\ &= \alpha_4 - \tau D_i \end{aligned}$$

MORE ON LINK FUNCTION

- Cumulative probabilities such as $\mathbb{P}(Y \leq \text{Agree})$
- Link function = CDF of the error term (F_{ε_i})

$$F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i \leq \text{Agree})) = \alpha_4 - \tau D_i$$

- Suppose we know F_{ε_i} :

$$\underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(0) \leq \text{Agree}))}_{\text{Control}} - \underbrace{F_{\varepsilon_i}^{-1}(\mathbb{P}(Y_i(1) \leq \text{Agree}))}_{\text{Treated}} = \tau$$

LITTLE MORE ON DENSITY ESTIMATION

- Kernel Density Estimation

$$\text{Let } V_i(\beta, \tau) = f(X_i, \beta) + D_i^\top \tau$$

By the Bayes' rule,

$$\mathbb{P}(Y_i \leq j \mid V_i = v) = \frac{\mathbb{P}(Y_i \leq j)g_1(v \mid Y \leq j)}{\mathbb{P}(Y_i \leq j)g_1(v \mid Y \leq j) + \mathbb{P}(Y_i > j)g_0(v \mid Y > j)}$$

Estimate g_i using KDE

$$\hat{g}_1(v \mid Y \leq j) = \frac{1}{n_1(j)h_{1j}} \sum_{i:Y_i \leq j} K\left(\frac{v - V_i}{h_{1j}}\right)$$

$$\hat{g}_0(v \mid Y > j) = \frac{1}{n_0(j)h_{0j}} \sum_{i:Y_i > j} K\left(\frac{v - V_i}{h_{0j}}\right)$$

LITTLE MORE ON DENSITY ESTIMATION

- Normalizing Flows
 - Use the change-of-variable formula

$$f_{\theta}(\varepsilon) = f_Z(T_{\theta}^{-1}(\varepsilon)) \left| \det \left(\frac{\partial T_{\theta}^{-1}(\varepsilon)}{\partial \varepsilon} \right) \right|,$$

- $\varepsilon_i = T_{\theta}(Z_i)$
- T_{θ} is a set of *invertible* transformation
- Maps ε_i to simple Z (e.g. Standard Normal)
- Estimate based on Z and then translate it back to ε_i

WHY NOT JUST USE BINARY OUTCOMES?

- Loss of Power
 - Collapsing categories wastes information
 - Require larger samples ($\approx 5\times$)
- Aggregation Bias
 - Arbitrary grouping may fail to identify effects or *flip the sign* of the effect
 - Depends on shifts in middle categories
- Example:
 - Treatment moves "Strongly Disagree" → "Disagree"
 - Binary ("Positive" vs "Negative") sees zero effect
 - Ordinal model captures the improvement