# Lecture 8 Multivariate Distributions

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#### **Conditional distributions**

Let X and Y have a joint discrete distribution with PMF f(x, y) on space S.

Say the marginal PMF are  $f_X(x)$  and  $f_Y(y)$  respectively. Let event

$$A = \{X = x\}$$
 and event  $B = \{Y = y\}$ . Thus  $A \cap B = \{X = x, Y = y\}$ .

Because  $P(A \cap B) = P(X = x, Y = y) = f(x, y)$  and

 $P(B) = P(Y = y) = f_Y(y)$ , the conditional probability of A given B is

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{f(x,y)}{f_Y(y)}$$

**Definition**: The conditional probability mass function of X, given that Y = y, is defined by

$$g(x|y) = \frac{f(x,y)}{f_Y(y)}$$

provided that  $f_Y(y) > 0$ 

# **Conditional distributions**

**Example**: Let *X* and *Y* have the joint PMF

$$f(x,y) = \frac{x+y}{21}, \quad x = 1,2,3, \quad y = 1,2.$$

Find the conditional distribution g(x|y).

We first calculate marginal PMF of *y*:

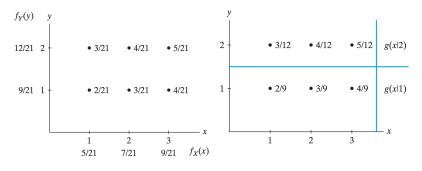
$$f_Y(y) = \sum_{x=1}^3 \frac{x+y}{21} = \frac{y+2}{7}, \quad y = 1, 2$$

Thus, the conditional PMF of X given Y is

$$g(x|y) = \frac{f(x,y)}{f_Y(y)} = \frac{(x+y)/21}{(y+2)/7} = \frac{x+y}{3y+6}$$

#### **Conditional distribution**

Similar to conditional probability, we can visualize the joint, marginal, and conditional PMF.



(Graphic illustration of joint, marginal and conditional PMF.)

# **Conditional expectation**

Because conditional PMF is a PMF, we thus can define conditional expectation the same way we define mathematical expectation:

$$E[u(Y)|X=x] = \sum_{y} u(y)g(y|x)$$

Conditional mean and conditional variance are defined by

$$\mu_{Y|X} = E(Y|X) = \sum_{y} yg(y|X)$$
 $\sigma_{Y|X}^{2} = E[(Y - \mu_{Y|X})^{2}|X] = \sum_{y} (y - \mu_{Y|X})^{2}g(y|X)$ 

# **Conditional expectation**

**Example**: Let X and Y have a multinomial PMF with parameters n,  $p_X$ , and  $p_Y$ . That is,

$$f(x,y) = \frac{n!}{x!y!(n-x-y)!} p_X^x p_Y^y (1 - p_X - p_Y)^{n-x-y}$$

What is the conditional mean of X given Y?

We know that the marginal distribution of Y is binomial,. i.e.,

$$f_Y(y) = \frac{n!}{y!(n-y)!}p_Y^y(1-p_Y)^{n-y}$$

Thus, the conditional PMF of X given Y is

$$g(x|y) = \frac{f(x,y)}{f_Y(y)} = \frac{(n-y)!}{x!(n-y-x)!} \left(\frac{p_X}{1-p_Y}\right)^x \left(1 - \frac{p_X}{1-p_Y}\right)^{n-y}$$

This is a binomial distribution with parameters n-y and  $\frac{\rho_X}{1-\rho_Y}$ . Thus, the conditional mean is  $(n-y)\frac{\rho_X}{1-\rho_Y}$ .

The idea of joint distributions of discrete random variables can be extended to that of continuous random variables. The **joint probability density function** of two continuous random variables is an integrable function f(x, y) such that

- $f(x,y) \le 0$ , where f(x,y) = 0 when (x,y) is not in the space of X and Y;
- $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = 1;$
- $P(X, Y) \in A = \int \int_A f(x, y) dx dy$

The marginal probability density function of X and Y are given by

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy, \quad x \in S_X;$$
  
 $f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx, \quad y \in S_Y;$ 

X and Y are **independent** if and only if  $f(x, y) = f_X(x)f_Y(y)$ 

The correlation coefficient of two continuous random variables X and Y is defined in the same way as the discrete random variables as

$$\rho = \frac{Cov(X, Y)}{\sigma_X \sigma_Y} = \frac{\sigma_{XY}}{\sigma_X \sigma_Y}$$

The Conditional probability density function of X, given that Y = y, is

$$f(x|y) = \frac{f(x,y)}{f_Y(y)},$$

provided that  $f_Y(y) > 0$ .

**Example**: Let *X* and *Y* have the joint PDF

$$f(x,y) = 1, \quad x \leqslant y \leqslant x + 1, \quad 0 \leqslant x \leqslant 1.$$

Find the marginal PDF and the correlation coefficient of X and Y.

The marginal PDFs of X and Y are

$$f_X(x) = \int_x^{x+1} 1 \, dy = 1, \quad 0 \leqslant x \leqslant 1$$

$$f_Y(y) = \begin{cases} \int_0^y 1 \, dx = y, \quad 0 \leqslant y \leqslant 1, \\ \int_{y-1}^1 1 \, dx = 2 - y, \quad 1 \leqslant y \leqslant 2. \end{cases}$$

The mean and variance of X and Y are

$$\mu_X = \int_0^1 x \cdot 1 dx = \frac{1}{2}$$

$$\mu_Y = \int_0^1 y \cdot y dy + \int_1^2 y \cdot (2 - y) dy = \frac{1}{3} + \frac{2}{3} = 1$$

$$E(X^2) = \int_0^1 x^2 \cdot 1 dx = \frac{1}{3}$$

$$E(Y^2) = \int_0^1 y^2 \cdot y dy + \int_1^2 y^2 \cdot (2 - y) dy = \frac{7}{6}$$

$$E(XY) = \int_0^1 \int_x^{x+1} xy \cdot 1 dy dx = \int_0^1 \frac{1}{2} x(2x+1) dx = \frac{7}{12}$$

$$\sigma_X^2 = \frac{1}{3} - \left(\frac{1}{2}\right)^2 = \frac{1}{12}$$

$$\sigma_Y^2 = \frac{7}{6} - 1^2 = \frac{1}{6}$$

$$\sigma_{XY} = \frac{7}{12} - \left(\frac{1}{2}\right)(1) = \frac{1}{12}$$

Therefore, the correlation coefficient is

$$\rho = \frac{\sigma_{XY}}{\sigma_X \sigma_Y} = \frac{1/12}{\sqrt{(1/12)(1/6)}} = \frac{\sqrt{2}}{2}$$

A very commonly used multivariate distribution is the multivariate normal distribution. Random variables *X* and *Y* have a bivariate normal distribution if its joint PDF is

$$f(x,y) = \frac{1}{2\pi\sigma_X\sigma_Y\sqrt{1-\rho^2}}\exp\Big[-\frac{q(x,y)}{2}\Big],$$

where

$$q(x,y) = \frac{1}{1-\rho^2} \left[ \left( \frac{x-\mu_X}{\sigma_X} \right)^2 - 2\rho \left( \frac{x-\mu_X}{\sigma_X} \right) \left( \frac{y-\mu_Y}{\sigma_Y} \right) + \left( \frac{y-\mu_Y}{\sigma_Y} \right)^2 \right]$$

Here,  $\mu_X$  and  $\mu_Y$  are the mean of X and Y,  $\sigma_X$  and  $\sigma_Y$  are the standard deviation of X and Y, and  $\rho$  is the correlation coefficient.

If random variables X and Y have a bivariate normal distribution, then the marginal distribution of X and Y are both normal.

$$\begin{aligned} q(x,y) &= \frac{1}{1-\rho^2} \left[ \left( \frac{x-\mu_X}{\sigma_X} \right)^2 - 2\rho \left( \frac{x-\mu_X}{\sigma_X} \right) \left( \frac{y-\mu_Y}{\sigma_Y} \right) + \left( \frac{y-\mu_Y}{\sigma_Y} \right)^2 \right] \\ &= \frac{1}{1-\rho^2} \left[ \left( \frac{x-\mu_X}{\sigma_X} - \rho \frac{y-\mu_Y}{\sigma_Y} \right)^2 + (1-\rho^2) \left( \frac{y-\mu_Y}{\sigma_Y} \right)^2 \right] \\ &= \frac{1}{\sigma_X^2 (1-\rho^2)} \left( x - \mu_X - \rho \frac{\sigma_X}{\sigma_Y} (y-\mu_Y) \right)^2 + \left( \frac{y-\mu_Y}{\sigma_Y} \right)^2 \end{aligned}$$

Thus, the marginal distribution of Y is

$$f_{Y}(y) = \int_{-\infty}^{\infty} f(x, y) dx = \int_{-\infty}^{\infty} \frac{1}{2\pi\sigma_{X}\sigma_{Y}\sqrt{1 - \rho^{2}}} \exp\left[-\frac{q(x, y)}{2}\right] dx$$

$$= \frac{1}{2\pi\sigma_{X}\sigma_{Y}\sqrt{1 - \rho^{2}}} \exp\left[-\frac{(y - \mu_{Y})^{2}}{2\sigma_{Y}^{2}}\right]$$

$$\int_{-\infty}^{\infty} \exp\left[-\frac{1}{2\sigma_{X}^{2}(1 - \rho^{2})} \left(x - \mu_{X} - \rho\frac{\sigma_{X}}{\sigma_{Y}}(y - \mu_{Y})\right)^{2}\right] dx$$

$$= \frac{1}{2\pi\sigma_{X}\sigma_{Y}\sqrt{1 - \rho^{2}}} \exp\left[-\frac{(y - \mu_{Y})^{2}}{2\sigma_{Y}^{2}}\right] (\sigma_{X}\sqrt{2\pi}\sqrt{1 - \rho^{2}})$$

$$= \frac{1}{\sigma_{Y}\sqrt{2\pi}} \exp\left[-\frac{(y - \mu_{Y})^{2}}{2\sigma_{Y}^{2}}\right]$$

Thus, the marginal distribution of Y is  $N(\mu_Y, \sigma_Y^2)$ . Using the procedure, it is obvious that  $X \sim N(\mu_X, \sigma_X^2)$ .

If If random variables X and Y have a bivariate normal distribution, then the conditional distribution of X given Y is normal.

The joint PDF is

$$f(x,y) = \frac{1}{2\pi\sigma_X\sigma_Y\sqrt{1-\rho^2}}\exp\Big[-\frac{q(x,y)}{2}\Big],$$

where

$$q(x,y) = \frac{1}{\sigma_X^2(1-\rho^2)} \left(x - \mu_X - \rho \frac{\sigma_X}{\sigma_Y}(y - \mu_Y)\right)^2 + \left(\frac{y - \mu_Y}{\sigma_Y}\right)^2$$

The marginal PDF of Y is

$$f_Y(y) = \frac{1}{\sigma_Y \sqrt{2\pi}} \exp\left[-\frac{(y - \mu_Y)^2}{2\sigma_Y^2}\right]$$

The conditional distribution of *X* given *Y* is thus

$$g(x|y) = \frac{f(x,y)}{f_Y(y)} = \frac{1}{\sigma_X \sqrt{2\pi} \sqrt{1 - \rho^2}} \exp \left[ -\frac{\left[ x - \mu_X - \rho (\sigma_X/\sigma_Y)(y - \mu_Y) \right]^2}{2\sigma_X^2 (1 - \rho^2)} \right]$$

Thus, g(x|y) is  $N(\mu_X + \rho \frac{\sigma_X}{\sigma_Y}(y - \mu_Y), (1 - \rho^2)\sigma_X^2)$ .

