

MAST90104: A First Course in Statistical Learning

Week 9 Practical and Workshop Solution

1 Workshop questions

1. Verify that for the binomial regression model with logistic link

$$\begin{aligned}\mathbb{E} \frac{\partial l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_i} &= 0 \\ -\mathbb{E} \frac{\partial^2 l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_i \partial \theta_j} &= \mathbb{E} \left(\frac{\partial l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_i} \frac{\partial l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_j} \right)\end{aligned}$$

Solution:

Suppose that $Y_k, k = 1, \dots, n \sim \text{Binomial}(m_k, p_k = g^{-1}(\mathbf{x}_k^T \boldsymbol{\theta}))$ where $\mathbf{x}_k, i = k, \dots, n$ are explanatory predictors and $g(p) = \log(\frac{p}{1-p})$ is the logistic link function. Hence

$$\begin{aligned}\frac{\partial l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_i} &= \sum_{k=1}^n Y_k \frac{1}{p_k} \frac{\partial p_k}{\partial \theta_i} - (m_k - Y_k) \frac{1}{1-p_k} \frac{\partial p_k}{\partial \theta_i} \\ &= \sum_{k=1}^n \frac{\partial p_k}{\partial \theta_i} Z_k\end{aligned}$$

where $Z_k = Y_k/p_k - (m_k - Y_k)/(1 - p_k)$, so $\mathbb{E}(Z_k) = m_k p_k/p_k - m_k(1 - p_k)/(1 - p_k) = 0$ and $\text{var}(Z_k) = \text{var}(Y_k/(p_k(1 - p_k)) - m_k/(1 - p_k)) = m_k/(p_k(1 - p_k))$. Thus

$$\begin{aligned}\mathbb{E} \frac{\partial l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_i} &= \sum_{k=1}^n \mathbb{E} \left(\frac{\partial p_k}{\partial \theta_i} Z_k \right) \\ &= \sum_{k=1}^n \frac{\partial p_k}{\partial \theta_i} \mathbb{E}(Z_k) = 0.\end{aligned}$$

Now

$$\begin{aligned}\frac{\partial^2 l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_i \partial \theta_j} &= \sum_{k=1}^n \frac{\partial^2 p_k}{\partial \theta_i \partial \theta_j} Z_k + \frac{\partial p_k}{\partial \theta_i} \frac{\partial Z_k}{\partial \theta_j} \\ &= \sum_{k=1}^n \frac{\partial^2 p_k}{\partial \theta_i \partial \theta_j} Z_k + \frac{\partial p_k}{\partial \theta_i} \frac{\partial p_k}{\partial \theta_j} \left(\frac{-Y_k}{p_k^2} - \frac{m_k - Y_k}{(1 - p_k)^2} \right)\end{aligned}$$

so

$$\begin{aligned}-\mathbb{E} \frac{\partial^2 l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_i \partial \theta_j} &= -0 + \sum_{k=1}^n \frac{\partial p_k}{\partial \theta_i} \frac{\partial p_k}{\partial \theta_j} \left(\frac{m_k}{p_k} + \frac{m_k}{1 - p_k} \right) \\ &= \sum_{k=1}^n \frac{\partial p_k}{\partial \theta_i} \frac{\partial p_k}{\partial \theta_j} \frac{m_k}{p_k(1 - p_k)}.\end{aligned}$$

Whereas

$$\left(\frac{\partial l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_i} \frac{\partial l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_j} \right) = \sum_{k=1}^n \frac{\partial p_k}{\partial \theta_i} Z_k \sum_{l=1}^n \frac{\partial p_l}{\partial \theta_j} Z_l$$

and, given the $Z_k, k = 1 \cdots n$ are independent,

$$\begin{aligned}
\mathbb{E} \left(\frac{\partial l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_i} \frac{\partial l(\boldsymbol{\theta}; \mathbf{Y})}{\partial \theta_j} \right) &= \sum_{k=1}^n \frac{\partial p_k}{\partial \theta_i} \mathbb{E} \left(Z_k \sum_{l=1}^n \frac{\partial p_l}{\partial \theta_j} Z_l \right) \\
&= \sum_{k=1}^n \frac{\partial p_k}{\partial \theta_i} \sum_{l=1}^n \frac{\partial p_l}{\partial \theta_j} \mathbb{E}(Z_k Z_l) \\
&= \sum_{k=1}^n \frac{\partial p_k}{\partial \theta_i} \frac{\partial p_l}{\partial \theta_j} \left(\mathbb{E}(Z_k^2) + \sum_{l \neq k} \mathbb{E}(Z_k Z_l) \right) \\
&= \sum_{k=1}^n \frac{\partial p_k}{\partial \theta_i} \frac{\partial p_l}{\partial \theta_j} \left(\text{var}(Z_k) + \sum_{l \neq k} \mathbb{E}(Z_k) \mathbb{E}(Z_l) \right) \\
&= \sum_{k=1}^n \frac{\partial p_k}{\partial \theta_i} \frac{\partial p_l}{\partial \theta_j} \frac{m_k}{p_k(1-p_k)} + 0
\end{aligned}$$

as required. Notice that the proof does not rely on the form of the relationship between p_k and $\boldsymbol{\theta}$.

2. A orthopaedics surgeon is investigating the functioning score of post-knee replacement surgery patients that underwent three different post-surgical intervention under three participating physiotherapists. The data collected includes the response `FunctionScore` and two factor variables `Intervention` and `Therapist`. Unfortunately, a computer virus corrupted the output file. The corrupted output is as follows:

```
str(DataF)
'data.frame': 20 obs. of 3 variables:
 $ FunctionScore: num 12.35 11.82 7.22 13.69 13.05 ...
 $ Intervention : Factor w/ 3 levels "1","2","3": 1 1 3 2 2 1 1 1 2 2 ...
 $ Therapist : Factor w/ 3 levels "A","B","C": 3 1 1 3 3 2 2 2 1 1 ...
> summary(imodel)
```

Call:

```
lm(formula = FunctionScore ~ Intervention * Therapist, data = DataF)
```

Residuals:

Min	1Q	Median	3Q	Max
-0.8285	-0.1937	0.0000	0.3317	0.6139

Coefficients:

Estimate	Std. Error	t value	Pr(> t)
(Intercept)	11.7250	0.3467	33.823 1.81e-12 ***
Intervention2	-3.2753	0.4903	-6.681 3.46e-05 ***
Intervention3	-4.1635	0.4475	-9.303 1.51e-06 ***
TherapistB	2.5520	0.4246	6.011 8.79e-05 ***
TherapistC	0.2161	0.4903	0.441 0.668
Intervention2:TherapistB	-0.8392	0.7354	-1.141 0.278
Intervention3:TherapistB	-0.2600	0.6169	-0.421 0.682
Intervention2:TherapistC	4.9021	0.6638	7.385 1.39e-05 ***
Intervention3:TherapistC	6.1056	0.7489	8.153 5.45e-06 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: (d) on (a) degrees of freedom

Multiple R-squared: 0.9788, Adjusted R-squared: 0.9634

F-statistic: 63.49 on 8 and (a) DF, p-value: 4.075e-08

```
> sum(imodel$residuals^2)
```

```

[1] 2.643836

summary(amodel)

Call:
lm(formula = FunctionScore ~ Intervention + Therapist, data = DataF)

Residuals:
Min      1Q  Median      3Q      Max
-2.91203 -0.68993  0.07554  1.02621  2.34695

Coefficients:
Estimate Std. Error t value Pr(>|t|)
(Intercept)  10.7563      0.7389   14.558 2.96e-10 ***
Intervention2 -1.7656      0.8253   -2.139 0.049253 *
Intervention3 -2.9097      0.8053   -3.613 0.002555 **
TherapistB      2.7526      0.8020    3.432 0.003705 **
TherapistC      3.6897      0.8361    4.413 0.000504 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 1.455 on (c) degrees of freedom
Multiple R-squared:  0.7455, Adjusted R-squared:  0.6777
F-statistic: 10.99 on 4 and (c) DF, p-value: 0.0002298

> anova(amodel, imodel)
Analysis of Variance Table

Model 1: FunctionScore ~ Intervention + Therapist
Model 2: FunctionScore ~ Intervention * Therapist
Res.Df  RSS  Df  Sum of Sq    F    Pr(>F)
1      (c) (e)
2      (a) (b) (f)  (g)      (h)  6.991e-06 ***
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

Deduce the missing output (a) to (h).

Solution:

- (a) Since there are 9 parameters in the reparameterised interaction model, $(a) = n - 9 = 20 - 9 = 11$.
- (b) Let e_i denote the i -th residual for the reparameterised interaction model. Then $(b) = \sum e_i^2 = 2.643836$.
- (c) Since there are 5 parameters in the reparameterised additive model, $(c) = n - 5 = 20 - 5 = 15$.
- (d) $s = \sqrt{\sum e_i^2 / 11} = 0.4902538$.
- (e) For additive model, $SS_{res}^{H_0} = (s^{H_0})^2 \times (n - 5) = 1.455^2 \times 15 = 31.75538$.
- (f) DF of $SS_{res}^{H_0} - SS_{res}^{H_1} = 15 - 11 = 4$.
- (g) $SS_{res}^{H_0} - SS_{res}^{H_1} = 31.75538 - 2.643836 = 29.11154$.
- (h) F-statistic $= \{(SS_{res}^{H_0} - SS_{res}^{H_1}) / 4\} / \{0.4902538^2\} = 30.28051$.