

2) (a) Revision Solution

$$\mu = \begin{pmatrix} -3 \\ 1 \end{pmatrix} \quad \Sigma = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} = I_2$$

(b) Since Ωy is a linear combination of an MVN vector y , $\Omega y \sim \text{MVN}$ with mean

$$\begin{aligned} \Omega \mu &= \begin{pmatrix} -2 & 1 \\ 0.5 & 2 \end{pmatrix} \begin{pmatrix} -3 \\ 1 \end{pmatrix} \\ &= \begin{pmatrix} 7 \\ 0.5 \end{pmatrix} \end{aligned}$$

and covariance matrix

$$\Omega \Omega^T = \begin{pmatrix} 5 & 1 \\ 1 & 4.25 \end{pmatrix}$$

(c) ~~$\psi^T B$~~ Use theorem 3.14:

$$\begin{aligned} \psi^T I B &= \psi^T B \\ &= (-3.5 \ -7) \begin{pmatrix} 1 & -0.5 \\ -0.5 & 1 \end{pmatrix} \\ &= (0 \ -5.25) \neq (0 \ 0) \end{aligned}$$

$\therefore y^T B y$ and $\psi^T y$ are NOT indept.

Q3 (a) A full rank matrix X corresponds to Assumption II is Lect 4 Part 1.

By Gauss-Markov theorem (Thm 4.4), we ^{also} need

(I) The true relationship between X and y is

$y = X\beta + \epsilon$, where X is an $n \times (k+1)$ matrix and β is a $(k+1)$ -dimensional vector for any $k \geq 0$.

(II) The random errors are zero-centered, i.e.,
 $E(\epsilon) = 0$

(III) The random errors are uncorrelated and have homogeneous variance, i.e.,

$$\text{Var}(\epsilon) = \sigma^2 I_n$$

(b) $y_i \stackrel{\text{ind}}{\sim} N(x_i^T \beta, d_i)$, $i = 1, \dots, n$

$$\Rightarrow y \sim \text{MVN}(X\beta, D) \text{ where } D = \begin{bmatrix} d_1 & & & \\ & d_2 & & \\ & & \ddots & \\ 0 & & & d_n \end{bmatrix}$$

$$\text{Let } \tilde{y} = D^{-\frac{1}{2}} y \text{ and } \tilde{X} = D^{-\frac{1}{2}} X$$

$$\text{Then } \tilde{y} \sim \text{MVN}(\tilde{X}\beta, I)$$

By theorem 4.11,

$$\tilde{\beta} = (\tilde{X}^T \tilde{X})^{-1} \tilde{X}^T \tilde{y} \text{ is UMVUE for } \beta$$

But....

$$\tilde{\beta} = (\tilde{X}^T \tilde{X})^{-1} \tilde{X}^T \tilde{y}$$

$$= \cancel{X}^T (X^T D^{-1} X)^{-1} X^T D^{-1} y$$

$$\neq \hat{\beta}$$

$\therefore \hat{\beta}$ is not UMVUE for β .

Q4) (a) $p=6$ ~~so~~ df of $SS_{res} = 16$

$$\Rightarrow n-p=16$$

$$\Rightarrow n=16+p=22$$

(b)

$$F = \frac{y^T H y / p}{SS_{res} / (n-p)} \sim F_{p, n-p}$$

\therefore null dist is F -distribution with
6 and 16 degrees of freedom.

Assumptions required: By theorems 5.2 and 5.3 we

need: $y =$ vector of responses 22×1 $X =$ design matrix 22×6

(I) The true relationship between X and y is
 $y = X\beta + \epsilon$

(II) X is full rank

(V) $\epsilon \sim \text{MVN}(0, \sigma^2 I)$, for some $\sigma^2 > 0$.

(C) $H_0: \beta_{\text{wing}} = 0$ vs. $H_1: \beta_{\text{wing}} \neq 0$.

$$T = \frac{\hat{\beta}_{\text{wing}}}{\text{SE}(\hat{\beta}_{\text{wing}})} = -1.325 \sim t_{16}$$

$$p\text{-value} = 0.20378 > 0.1$$

Since $p\text{-value} > 0.1$, we don't reject H_0
and we conclude that β_{wing} is plausibly zero.

$$(d) \text{ Multiple } R^2 = 1 - \frac{SS_{\text{res}}}{\text{Corrected } SS_{\text{total}}}$$

$$\text{Since } \text{Adj-}R^2 = 1 - \frac{n-1}{n-p} (1 - R^2)$$

Since R^2 never decreases as we add new predictors in the model, it is unsuitable for model selection. On the other hand, $\text{adj-}R^2$ accounts for model complexity by penalising models based on number of coefficients. Hence, unlike R^2 , the $\text{adj-}R^2$ is suitable for model selection.

(f) Since $\beta_{\text{beaktonotch}}$ has the smallest absolute t -statistic value (equivalent to smallest F -statistic), we should drop it first.

$$\begin{aligned} \text{(e)} \quad \text{adj-}R^2 &= 1 - \frac{n-1}{n-1-k} (1 - R^2) \\ &= 1 - \frac{22-1}{22-1-5} (1 - 0.9633) \\ &= 0.9518313 \end{aligned}$$