Networks and Random Processes Assignment 1

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1 Question 1

This section considers a Simple Random Walk on 1, ..., L with probabilities $p \in [0, 1]$ and q = 1 - p to jump right and left respectively.

Different boundary conditions are considered.

1.1 Part A

1.1.1 Case 1 - Periodic

Periodic boundary conditions, so p(0, L) = q and p(L, 0) = p

The transition matrix is:

$$P = \begin{bmatrix} 0 & p & 0 & \dots & 0 & q \\ q & 0 & p & \dots & 0 & 0 \\ 0 & q & 0 & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & 0 & 0 & \dots & 0 & p \\ p & 0 & 0 & \dots & q & 0 \end{bmatrix}$$

The process is irreducible, \bar{i} .e. every state can, eventually, reach every other state. And there is a finite state space, so it has 1 unique stationary distribution.

The states can be laid out on a circle, and are symmetrical, so the stationary distribution is where all states have equal probabilities.

$$\pi = (1/L, 1/L, ..., 1/L)$$
 for all $p \in (0, 1)$

The stationary distribution is reversible only for the case $p=q=\frac{1}{2}$.

1.1.2 Case 2 - Closed

Transition matrix:

$$P = \begin{bmatrix} q & p & 0 & \dots & 0 & 0 \\ q & 0 & p & \dots & 0 & 0 \\ 0 & q & 0 & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & 0 & 0 & \dots & 0 & p \\ 0 & 0 & 0 & \dots & q & p \end{bmatrix}$$

If p=1, there is an absorbing state at L, and the stationary distribution is $\pi = (0, 0, ..., 0, 1)$. This stationary distribution is reversible, all terms in the detailed balance equations are zero. This is not irreducible, as the walk can only move from s_x to s_{x+1} .

If q=1, there is an absorbing state at 1, and the stationary distribution is $\pi = (1, 0, ..., 0, 0)$. This stationary distribution is reversible, all terms in the detailed balance equations are zero. This is not irreducible, as the walk can only move from s_x to s_{x-1} .

If $p=q=\frac{1}{2}$, the process is irreducible, as every state can, eventually, be reached from every other state. There is a finite state space and so there is onle 1 stationary distribution. The sum of all columns equal 1, so there is a constant left eigenvector, and so the stationary distribution is $\pi = (1/L, 1/L, ..., 1/L)$. This stationary distribution is reversible.

If $p, q \neq [0, 1/2, 1]$, then the process is irreducible, and it has a finite state space so there is only 1 stationary distribution. Looing at the detailed balance equations, we can find a recurrence relation of the form:

$$\pi_{x-1}p = \pi_x p$$
 over $x = \{2, 3, ..., L-1\}$
 $\pi_x = \pi_{x-1} \frac{p}{q}$
By induction, this suggests a solution like

 $\pi_x = \pi_1(\frac{p}{q})^{x-1}$

This is a reversible distribution, and so must also be stationary. As the process is ergodic, this is the unique stationary distribution.

We can check this distribution is staionary for a closed simple random walk with 4 states

$$\begin{bmatrix} \pi_1 \\ \pi_1 \frac{p}{q} \\ \pi_1 (\frac{p}{q})^2 \\ \pi_1 (\frac{p}{q})^3 \end{bmatrix} \begin{bmatrix} q & p & 0 & 0 \\ q & 0 & p & 0 \\ 0 & q & 0 & p \\ 0 & 0 & q & p \end{bmatrix} = \begin{bmatrix} \pi_1 \\ \pi_1 \frac{p}{q} \\ \pi_1 (\frac{p}{q})^2 \\ \pi_1 (\frac{p}{q})^3 \end{bmatrix}$$

We also need to normalise the distribution, so the stationary distribution

will be
$$\pi_x = \frac{\left(\frac{p}{q}\right)^{x-1}}{\sum_i^L \pi_i}$$

1.2 Part B - Absorbing

$$P = \begin{bmatrix} 1 & 0 & 0 & \dots & 0 & 0 \\ q & 0 & p & \dots & 0 & 0 \\ 0 & q & 0 & \dots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & 0 & 0 & \dots & 0 & p \\ 0 & 0 & 0 & \dots & 0 & 1 \end{bmatrix}$$

The process is not irreducible, as no other states can be reached from state 1 or state L.

The (normalised) stationary distributions are

$$\pi_1 = [1, 0, 0, ..., 0]$$

$$\pi_2 = [0, 0, 0, ..., 1]$$

 $\pi_3 = [a, 0, 0, ..., 0, 1 - a]$ where $a \in [0, 1]$

These distributions are reversible, looking at the detailed balance conditions: $\pi(x)p(x,y) = \pi(y)p(y,x)$

All terms for all equations are zero, therefore it is reversible.

The absorption probability in site L is

$$h_k^L = P(X_n = L \text{ for some } n \ge 0 | X_0 = k)$$

$$h_k^L = P(X_n = L \text{ for some } n \ge 0 | X_0 = k)$$

 $h_k^L = P(X_n = L \text{ for some } n \ge 0 | X_0 = k)$
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$$h_k^L = P(X_n = L | X_0 = k)$$

Considering k+1, k-1, and by law of total probability:

$$h_k^L = P(X_n = L | X_1 = k+1, X_0 = k) * p + P(X_n = L | X_1 = k-1, X_0 = k) * q$$

Using the Markov property:

$$h_k^L = P(X_n = L | X_1 = k+1) \times p + P(X_n = L | X_1 = k-1) \times q$$

$$h_k^L = h_{k+1}^L \times p + h_{k-1}^L \times q$$

 $\begin{aligned} h_k^L &= h_{k+1}^L \times p + h_{k-1}^L \times q \\ \text{And the boundary conditions are:} \\ h_1^L &= 0 \text{ and } h_L^L = 1 \end{aligned}$

$$h_1^L = 0 \text{ and } h_L^L = 1$$

If p = q then this recursion relation becomes

$$h_k^L = \frac{h_{k+1}^L + h_{k-1}^L}{2}$$

This is linear interpolation between the two surrounding states, so the absorption probability is linear in k. Consdiering the boundary conditions, the solution is:

$$h_k^L = \frac{k-1}{L-1}$$

If $p \neq q$, we can solve the recursion realtion by considering the ansatz:

$$h_k^L = \lambda^k$$

$$\lambda = p\lambda^2 + q$$

This has roots:

$$\lambda_1 = 1$$
 and $\lambda_2 = q/p$

The general solution is of the form:

$$h_1^L = a\lambda_1 + b\lambda_2$$

$$h_{L}^{L} = a + b(\frac{q}{a})^{k}$$

 $h_k^L=a\lambda_1+b\lambda_2$ $h_k^L=a+b(\frac{q}{p})^k$ Looking at the boundary conditions:

$$h_1^L = 0 = a + b(\frac{q}{p})$$

$$h_L^L = 1 = a + b(\frac{q}{r})^L$$

 $h_L^L=1=a+b(\frac{q}{p})^L$ Subtracting the first equation from the second equation, and solving for b:

$$b = \frac{1}{\left(\frac{q}{p}\right)^{L} - \frac{q}{p}}$$
$$a = \frac{-1}{\left(\frac{q}{p}\right)^{L-1} - 1}$$

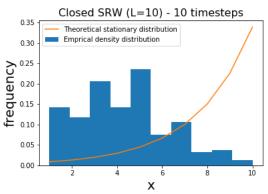


Figure 1. The frequency distri-

bution over 500 different realisations after 10 timesteps.

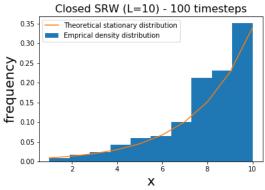


Figure 2. The frequency distri-

bution over 500 different realisations after 100 timesteps.

1.3 Part C - simulations

A simple random walk with L=10 and closed boundary conditions was simulated 500 times, with a p = 0.6 and starting at x=1 at t=0.

After 10 time steps (Figure 1), the state of the process is still heavily influenced by the starting condition of x(0) = 1.

After 100 time steps (Figure 2), the empirical distribution is similar to the theoretical stationary distribution. Ergodic processes tend towards the stationary distribution after a large number of time steps.

After 500 time steps of 1 realisation (Figure 3), the empirical distribution is similar to the theoretical stationary distribution. This is just 1 realisation, so there is a lot of stochasticity in the specific distribution generated. This is a reasonably good representative example.

Closed SRW (L=10)-1 realisation with 500 timesteps

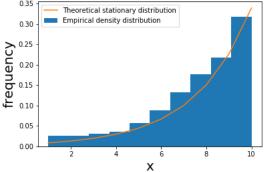


Figure 3. The frequency distri-

bution of states over 500 timesteps of 1 realisation of the simulated simple random walk.

2 Question 2

 X_1, X_2, \dots is a sequence of independent, identically distributed random variables (iirdvs) with

$$X_i \sim N(\mu, \sigma^2)$$

$$\mu \in R \text{ and } \sigma^2 > 0$$

Discrete time random walk on state space R

$$(Y_n : n \ge 0)$$
 with $Y_{n+1} = Y_n + X_{n+1}$ and $Y_0 = 0$

2.1 Part A

The weak law of large numbers for Y_n :

$$\frac{1}{n}Y_n = \frac{1}{n}\sum_{k=1}^n X_k - > \mu \text{ as } n - > \infty$$

The expected value of $\frac{Y_n}{n}$ will converge to μ with large n.

The central limit theorem for Y_n :

$$\frac{Y_n - n\mu}{\sigma n^{\frac{1}{2}}} = \frac{1}{\sigma n^{\frac{1}{2}}} \sum_{k=1}^n (X_k - \mu) - > \xi \text{ as } n - > \infty$$
where $\xi \sim N(0, 1)$

2.2 Part B - distribution of Y_n

$$Y_n \sim N(n\mu, n\sigma^2)$$

 Y_n is approximately normally distributed with mean equal to $n\mu$, where μ is the mean of the random variable X, and with variance $n\sigma^2$, where σ^2 is the variance of the random variable X.

2.3 Part C

 \mathbb{Z}_n has a recurive relationship given by

$$\begin{split} Z_{n+1} &= Z_n exp(X_{n+1}) \\ \text{and} \\ Z_n &= \sum_{i=1}^n exp(X_i) \\ \text{TO DO - Derive the pdf, look at Matt's link on whatsapp} \\ E(Z_n) &= exp(n\mu + \frac{n\sigma^2}{2}) \\ Var(Z_n) &= exp(2n\mu + n\sigma^2)(exp(n\sigma^2) - 1) \\ \text{Where } \mu \text{ and } \sigma^2 \text{ are the mean and variance of X.} \\ \text{TO DO - MEDIAN TO DO - check all this} \end{split}$$

2.4Part E

We want

$$E(Z_n) = 1 for all n$$

$$E(Z_n) = exp(n\mu + \frac{n\sigma^2}{2}) = 1$$

$$\mu + \frac{\sigma^2}{2} = 0$$

$$\mu = -\frac{\sigma^2}{50}$$

$$\mu = -\frac{1}{50}$$

3 Question 3

3.1 Part a

The state space is $X(i) \in X_0(j)$ where $i, j \in [1, L]$

It is not irreducible, as once a type goes extinct it stays extinct, and so not every state can be reached from every other state.

The stationary distributions are where all individuals are of one type Q- is that right?

3.2Part b

The future state of the process depends only on the current state at each time step, so yes it is a Markov process.

The state space is

$$N_i \in [0, L]$$
 and $\sum_{i=1}^{L} N_i = L$
The transition probabilities are ???

The process is not irreducible, as once a state has $N_i = 0$ for any i, i.e. a type goes extinct, the state space where $N_i > 0$ is no longer accessible.

The stationary distributions are where

$$N_i = \delta_{i,j} L$$
 where $j \in [1, L]$

Q. Limiting distribution?