

Predictive Assignment 2

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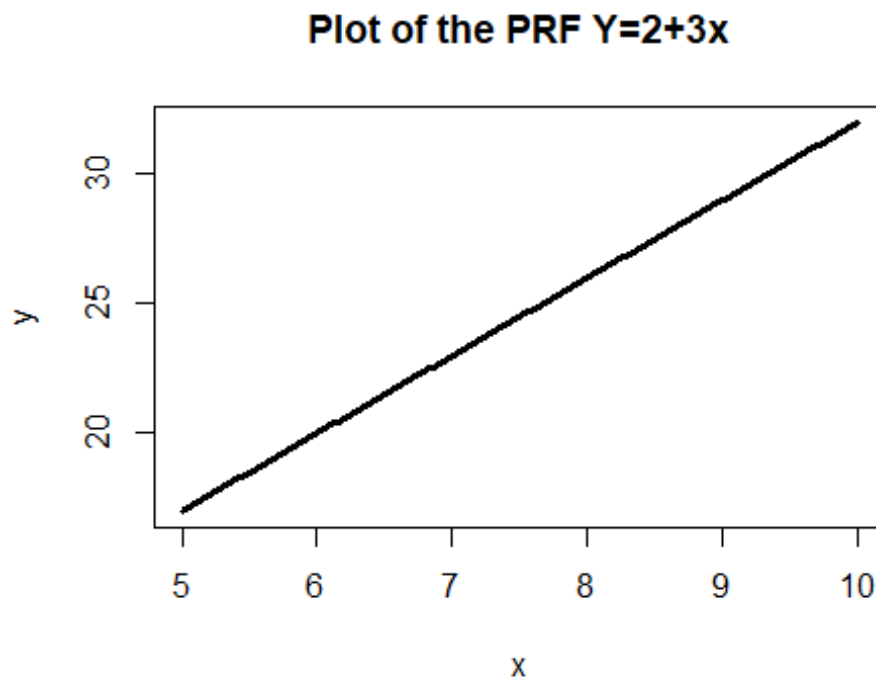
2026-02-05

1. Problem to demonstrate that the population regression line is fixed, but least square regression line varies

Model:

$$y = 2 + 3x + \varepsilon, \quad \varepsilon \sim \mathcal{N}(0, \sigma^2)$$
$$y_i \sim U(0,1), \quad \varepsilon_i \sim \mathcal{N}(0, 4^2), \quad \forall i = 1, 2, \dots, n$$

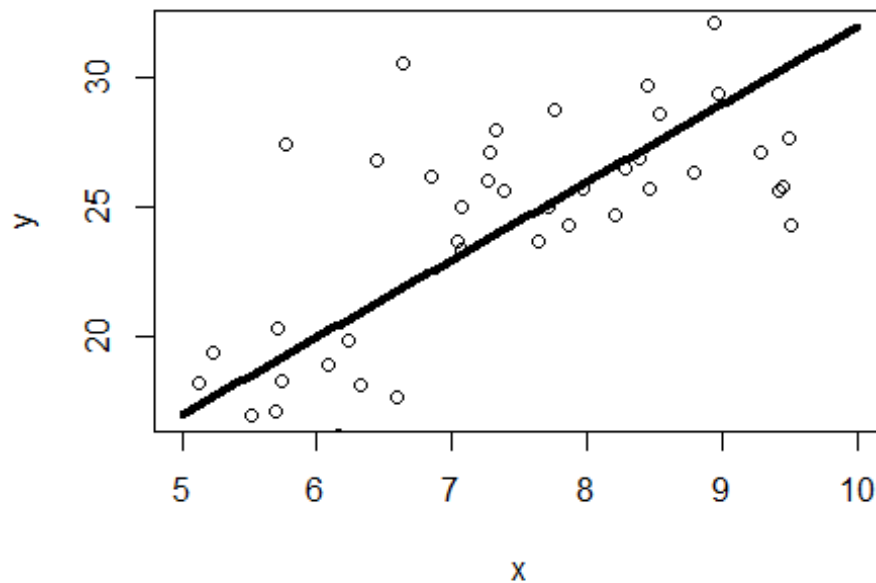
```
rm(list=ls())  
  
#Step 1: Take 200 equidistant points in the range [5:10]  
x=seq(5,10,length.out=200)  
y=2+3*x  
plot(x,y,type='l',lwd=3,main="Plot of the PRF Y=2+3x")
```



```
#Step 2: Generate xi from U(5,10) and ei from N(0,16)  
n=50  
set.seed(123)  
xi=runif(n,5,10)
```

```
ei=rnorm(n,0,4)
#SRF y=2+3x+e
yi=2+3*xi+ei
#Overlaying the the points on the PRF
plot(x,y,type='l',lwd=4,main="Plot of the PRF Y=2+3x and the Sample points")
points(xi,yi)
```

Plot of the PRF $Y=2+3x$ and the Sample points



#Step 3: Least Square Regression Line

```
modell1=lm(yi~xi)
```

```
summary(modell1)
```

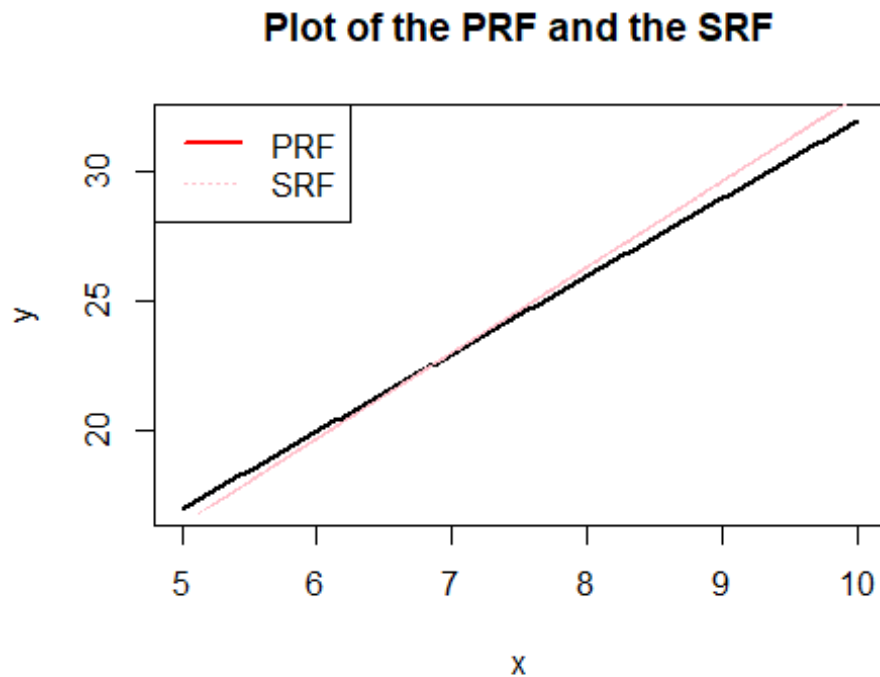
```
##
## Call:
## lm(formula = yi ~ xi)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -9.0231 -2.2314 -0.2627  2.1970  8.7445
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.09639    2.82610  -0.034   0.973
## xi           3.30540    0.36519   9.051 5.96e-12 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.761 on 48 degrees of freedom
```

```
## Multiple R-squared:  0.6306, Adjusted R-squared:  0.6229
## F-statistic: 81.93 on 1 and 48 DF,  p-value: 5.962e-12
```

$$\hat{y}_i = -0.09639 + 3.30540 x_i$$

#Overlaying the PRF and SRF

```
y.hat = -0.09639 + 3.30540*xi
plot(x,y,type='l',lwd=2,main="Plot of the PRF and the SRF")
lines(xi,y.hat,type='l',col="pink",lty="dotted")
legend("topleft",col=c("red","pink"),lwd=c(2,1),legend=c("PRF","SRF"),
      lty=c("solid","dotted"))
```



#Step 4: Repeat step 2 and 3 for 5 times.

```
set.seed(123)
x1=runif(n,5,10)
e1=rnorm(n,0,4)
y1=2+3*x1+e1
fit1=lm(y1~x1)
summary(fit1)

##
## Call:
## lm(formula = y1 ~ x1)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -9.0231 -2.2314 -0.2627  2.1970  8.7445
```

```
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) -0.09639    2.82610  -0.034    0.973
## x1          3.30540    0.36519   9.051 5.96e-12 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.761 on 48 degrees of freedom
## Multiple R-squared:  0.6306, Adjusted R-squared:  0.6229
## F-statistic: 81.93 on 1 and 48 DF,  p-value: 5.962e-12

x2=runif(n,5,10)
e2=rnorm(n,0,4)
y2=2+3*x2+e2
fit2=lm(y2~x2)

x3=runif(n,5,10)
e3=rnorm(n,0,4)
y3=2+3*x3+e3
fit3=lm(y3~x3)

x4=runif(n,5,10)
e4=rnorm(n,0,4)
y4=2+3*x4+e4
fit4=lm(y4~x4)

x5=runif(n,5,10)
e5=rnorm(n,0,4)
y5=2+3*x5+e5
fit5=lm(y5~x5)

#Data frame containing the five models' coefficients
coeff=data.frame(coef(fit1),
                 coef(fit2),
                 coef(fit3),
                 coef(fit4),
                 coef(fit5))

coeff

##             coef.fit1. coef.fit2. coef.fit3. coef.fit4. coef.fit5.
## (Intercept) -0.09638929  2.792188  1.392997  2.823089  2.032506
## x1          3.30539569  2.761042  3.073267  3.023608  3.028097

#Plot of the PRF and the Five SRFs

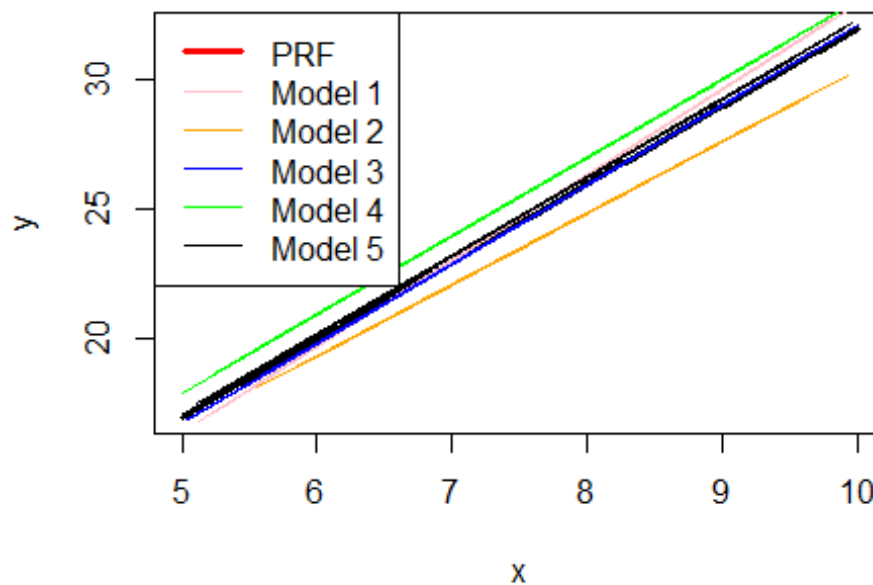
plot(x,y,type='l',lwd=3,main="Plot of the PRF and the Five SRFs")
lines(x1,predict(fit1),type='l',col="pink")
lines(x2,predict(fit2),type='l',col="orange")
lines(x3,predict(fit3),type='l',col="blue")
```

```

lines(x4,predict(fit4),type='l',col="green")
lines(x5,predict(fit5),type='l',col="black")
legend("topleft",legend=c("PRF","Model 1","Model 2","Model 3",
                          "Model 4","Model 5"),
      col=c("red","pink","orange","blue","green","black"),
      lwd=c(3,1,1,1,1,1))

```

Plot of the PRF and the Five SRFs



#Interpretation: PRF is fixed but SRF varies

2. Problem to demonstrate that $\hat{\beta}_0$ and $\hat{\beta}_1$ minimises RSS

Model

$$\hat{y}_i = -0.09639 + 3.30540 x_i$$

Here

$$\varepsilon_i \sim \mathcal{N}(0,1), \quad \forall i = 1, 2, \dots, n$$

```

rm(list=ls())
#step 1: Generate  $x_i \sim U(5,10)$  of size 50, and  $\varepsilon_i \sim N(0,1)$ 
# $y = 2 + 3x + \varepsilon$ 
n=50
set.seed(123)
x=runif(n,5,10)
xm=x-mean(x)

```

```

e=rnorm(n)
y=2+3*xm+e
fit=lm(y~xm)
summary(fit)

##
## Call:
## lm(formula = y ~ xm)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.25578 -0.55786 -0.06567  0.54926  2.18613
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   2.0562     0.1330   15.46  <2e-16 ***
## xm           3.0764     0.0913   33.70  <2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.9404 on 48 degrees of freedom
## Multiple R-squared:  0.9594, Adjusted R-squared:  0.9586
## F-statistic: 1135 on 1 and 48 DF, p-value: < 2.2e-16

```

$$\hat{\beta}_0 = 2.0562, \quad \hat{\beta} = 3.0764$$

$$RSS = (y_1 - \beta_0 - \beta x_1)^2 + (y_2 - \beta_0 - \beta x_2)^2 + \cdots + (y_n - \beta_0 - \beta x_n)^2$$

OR

$$RSS = \sum_{i=1}^n (y_i - \beta_0 - \beta x_i)^2$$

```

beta0=2.0562
beta=3.0764
#Creating the beta and beta0 grid values
beta0.grid=seq(-1,1,length.out=101)+beta0
beta.grid=seq(-1,1,length.out=101)+beta
#Computing RSS
RSS=c()
for(i in 1:length(beta.grid))
{
  RSS[i]=sum((y-beta0.grid[i]-beta.grid[i]*xm)^2)
}
RSS

## [1] 198.52541 192.34441 186.28828 180.35703 174.55065 168.86915
## [8] 157.88076 152.57388 147.39187 142.33473 137.40247 132.59508
## [15] 127.91256

```

```
## [15] 123.35492 118.92215 114.61425 110.43123 106.37308 102.43981
98.63140
## [22] 94.94788 91.38922 87.95544 84.64653 81.46249 78.40333
75.46905
## [29] 72.65963 69.97509 67.41542 64.98063 62.67071 60.48566
58.42549
## [36] 56.49019 54.67976 52.99421 51.43353 49.99772 48.68679
47.50073
## [43] 46.43954 45.50323 44.69179 44.00522 43.44353 43.00671
42.69477
## [50] 42.50769 42.44550 42.50817 42.69572 43.00814 43.44544
44.00761
## [57] 44.69465 45.50656 46.44335 47.50502 48.69155 50.00296
51.43925
## [64] 53.00040 54.68643 56.49734 58.43311 60.49376 62.67929
64.98968
## [71] 67.42495 69.98510 72.67012 75.48001 78.41477 81.47441
84.65892
## [78] 87.96831 91.40257 94.96170 98.64570 102.45458 106.38833
110.44696
## [85] 114.63046 118.93883 123.37208 127.93020 132.61319 137.42106
142.35380
## [92] 147.41141 152.59390 157.90126 163.33349 168.89060 174.57258
180.37943
## [99] 186.31116 192.36776 198.54924
```

```
df=data.frame(beta0.grid,beta.grid,RSS)
head(df) #For viewing purpose
```

```
## beta0.grid beta.grid RSS
## 1 1.0562 2.0764 198.5254
## 2 1.0762 2.0964 192.3444
## 3 1.0962 2.1164 186.2883
## 4 1.1162 2.1364 180.3570
## 5 1.1362 2.1564 174.5507
## 6 1.1562 2.1764 168.8692
```

```
df[which.min(RSS),]
```

```
## beta0.grid beta.grid RSS
## 51 2.0562 3.0764 42.4455
```

#We can verify the LSE minimises the RSS

RSS is minimum at $\beta_0 = 2.0562$, $\beta = 3.0764$, with RSS = 42.4455, which are the least squares estimates.

3. Problem to demonstrate that least square estimators are unbiased

```
rm(list=ls())
#Step 1: Generate  $x_i \sim U(5,10)$  of size 50, and  $e_i \sim N(0,1)$ 
n=50
set.seed(123)
x=runif(n,0,1)
e=rnorm(n)
y=2+3*x+e
#Step 2: obtain LSE
fit=lm(y~x)
summary(fit)

##
## Call:
## lm(formula = y ~ x)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -2.25578 -0.55786 -0.06567  0.54926  2.18613
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)   1.8576      0.2721   6.827 1.36e-08 ***
## x             3.3817      0.4565   7.408 1.75e-09 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.9404 on 48 degrees of freedom
## Multiple R-squared:  0.5334, Adjusted R-squared:  0.5237
## F-statistic: 54.88 on 1 and 48 DF,  p-value: 1.745e-09
```

$$\hat{\beta}_0 = 1.8576, \quad \hat{\beta} = 3.3817$$

```
beta0.hat=1.8576
beta.hat=3.3817
#Repeat R=1000 times
set.seed(123)
R=10000
beta0=numeric(R)
beta=numeric(R)
n=50
for(i in 1:R) {
  x=runif(n,0,1)
  e=rnorm(n,0,1)
  y=2+3*x+e
  fit=lm(y~x)
  beta0[i]=coef(fit)[1]
  beta[i]=coef(fit)[2]
}
```



```
list(Expected_beta0=mean(beta0),Expected_beta=mean(beta))
```

```
## $Expected_beta0  
## [1] 2.005203  
##  
## $Expected_beta  
## [1] 2.991338
```

#comment: the Long run means of beta0.hat and beta hat converges to true beta and beta0

$$\text{As } n \rightarrow \infty, \quad \mathbb{E}(\hat{\beta}_0) \rightarrow \beta_0 \quad \text{and} \quad \mathbb{E}(\hat{\beta}) \rightarrow \beta$$

```
list(var(beta0),var(beta))
```

```
## [[1]]  
## [1] 0.08292927  
##  
## [[2]]  
## [1] 0.2495031
```

#Comment: var approaches 0 as R goes up. Hence LSE is consistent.

$$\text{As } n \rightarrow \infty, \quad \text{Var}(\hat{\beta}_0) \rightarrow 0 \quad \text{and} \quad \text{Var}(\hat{\beta}) \rightarrow 0.$$

Hence, by the sufficient condition for consistency, the least squares estimators $\hat{\beta}_0$ and $\hat{\beta}$ are consistent.

4. Comparing several simple linear regressions

```
rm(list=ls())  
library(MASS)  
attach(Boston)
```

Model 1: $\text{medv}_i = \beta_0 + \beta_1 \text{lstat}_i + \varepsilon_i,$
Model 2: $\text{medv}_i = \beta_0 + \beta_1 \text{crim}_i + \varepsilon_i,$
Model 3: $\text{medv}_i = \beta_0 + \beta_1 \text{nox}_i + \varepsilon_i,$
Model 4: $\text{medv}_i = \beta_0 + \beta_1 \text{black}_i + \varepsilon_i,$

$$\varepsilon_i \sim \mathcal{N}(0, \sigma^2), \quad i = 1, 2, \dots, n$$

β_0 and

β_1 in each case are unknown parameters and are estimated using the method of least squares.

```
fit1=lm(medv~lstat)  
fit2=lm(medv~crim)  
fit3=lm(medv~nox)  
fit4=lm(medv~black)
```

(a) Selecting the predictors one by one, run four separate linear regressions to the data. Present the output in a single table.

```
library(stargazer)

## Warning: package 'stargazer' was built under R version 4.5.2

##
## Please cite as:
## Hlavac, Marek (2022). stargazer: Well-Formatted Regression and Summary
## Statistics Tables.
## R package version 5.2.3. https://CRAN.R-project.org/package=stargazer

stargazer(fit1,fit2,fit3,fit4,type="text",out="f.txt")

##
## =====
##                               Dependent variable:
##                               -----
##                               medv
##                               (1)      (2)      (3)      (4)
## -----
## lstat      -0.950***
##              (0.039)
##
## crim              -0.415***
##                  (0.044)
##
## nox              -33.916***
##                  (3.196)
##
## black              0.034***
##                  (0.004)
##
## Constant      34.554***  24.033***  41.346***  10.551***
##              (0.563)   (0.409)   (1.811)   (1.557)
## -----
## Observations      506      506      506      506
## R2                0.544      0.151      0.183      0.111
## Adjusted R2       0.543      0.149      0.181      0.109
## Residual Std. Error (df = 504)  6.216      8.484      8.323      8.679
## F Statistic (df = 1; 504)    601.618***  89.486***  112.591***  63.054***
## =====
## Note:                *p<0.1; **p<0.05; ***p<0.01
```

(b) Which model gives the best fit?

Model 1 gives the best fit among the four simple linear regression models since it has the highest coefficient of determination.

The R^2 values for the models are:

$$R_1^2 = 0.544, \quad R_2^2 = 0.151, \quad R_3^2 = 0.183, \quad R_4^2 = 0.111.$$

Since R^2 measures the variability in the dependent variable explained by the predictor, the higher R^2 value for Model 1 indicates that *lstat* is the best explanatory variable among the four.

Hence, Model 1 provides the best fit to the data.

(c) Comparison and usefulness of predictors:

From the four simple linear regression models with *medv* as the dependent variable, all predictors are statistically significant at the 1% level, indicating that each variable has a meaningful linear association with median house value.

The coefficient of *lstat* is -0.950 , implying that an increase in the percentage of lower-status population is associated with a substantial decrease in *medv*. This model also has the highest R^2 value (0.544), suggesting that *lstat* is the most useful predictor among the four in explaining variability in *medv*.

The coefficient of *crim* is -0.415 , indicating that higher crime rates are associated with lower median house values. However, its explanatory power is relatively weak, with an R^2 of 0.151, making it less useful compared to *lstat*.

The coefficient of *nox* is -33.916 , showing a strong negative effect of air pollution on *medv*. Although the magnitude of the coefficient is large due to the scale of *nox*, the R^2 value (0.183) suggests moderate usefulness in explaining variation in house values.

The coefficient of *black* is 0.034, indicating a positive association with *medv*. Despite being statistically significant, this predictor has the lowest explanatory power with an R^2 of 0.111, making it the least useful among the four.