



CHA.4 AND 5 HYPOTHESES TESTING AND INFERENCE BASED ON TWO SAMPLE

□ The hypotheses to be test is called null hypotheses H_0 and negation is called alternative hypotheses H_a

□ A hypotheses is said to be a sample hypotheses when H completely specifies the density of the population (*i. e.* $\theta = 2, \dots$). Otherwise is called composite hypotheses(*i. e.* $\theta < 2, \dots$)

➤ Method of finding test

❖ The likelihood ratio test statistic denoted with (W)

and (C) refers to critical region or just rejection region is the set store the value for reject the null hypotheses H_0 . It is defined by

$$W(x_1, x_2, \dots, x_n) = \frac{\max_{\theta \in \Omega_0} L(\theta, x_1, \dots, x_n)}{\max_{\theta \in \Omega_a} L(\theta, x_1, \dots, x_n)} \text{ or } W = \frac{L(\theta_0)}{L(\theta_a)}$$

then $C = \{(x_1, x_2, \dots, x_n) | W(\theta, x_1, x_2, \dots, x_n) \leq k\}, k \in [0, 1]$

❖ We have two type of errors are

1. Reject H_0 when H_0 is true . It is called type I error
2. Accept H_0 when H_0 is false . It is called type 2 error

❖ The significance level α is the value of probability of type I error

❖ β is the value of probability of type II error

❖ The function or power function is denoted by π and defined by

$$\pi(\theta) = P(\text{type I error}) \text{ or } \pi(\theta) = 1 - P(\text{type 2 error})$$

□ The uniformly most powerful (UMP)

if T is UMP than w of the same test of level δ if : $\pi_T(\theta) \geq \pi_W(\theta)$

□ Neyman-Pearson Lemma is the set of critical region c defined by

$$C = \{(x_1, \dots, x_n) \mid \frac{L(\theta_0)}{L(\theta_a)} \leq k\}, \text{ where } L \text{ is the likelihood function}$$

□ If X_1, \dots, X_n is random sample from normal distribution that has parameter μ and σ^2 are known and $n < 30$

□ The z test for population mean is defined as the test statistic value is $z = \frac{\bar{x} - \mu_0}{\sigma_0 / \sqrt{n}}$

□ Three case of hypotheses test when $H_0: \mu = \mu_0$ and if

1. $H_a: \mu > \mu_a$ then critical region C or test is $C = \{z: z = \frac{\bar{x} - \mu_0}{\frac{\sigma_0}{\sqrt{n}}} \geq z_\alpha\}$

2. $H_a: \mu < \mu_a$ then $C = \{z: z \leq -z_\alpha\}$

3. $H_a: \mu \neq \mu_a$ then $C = \{z: |z| \geq z_\alpha\}$

- For large sample $n \geq 30$ and it is from a distribution with known μ , then

$$z = \frac{\bar{x} - \mu}{s/\sqrt{n}}, \text{ where } s \text{ is the standard deviation of the sample}$$

❑ We use T-test when it is a normal distribution with unknown parameter μ and σ^2 also $n < 30$

when $H_0: \mu = \mu_0$ then test statistic value for T-test : $t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$

❑ the critical region (C) or rejection region (RR)

1. when $H_a: \mu > \mu_0$, then $C = \{t: t \geq t_{\alpha, n-1}\}$

2. when $H_a: \mu < \mu_0$ then $C = \{t: t \leq -t_{\alpha, n-1}\}$

3. When $H_a: \mu \neq \mu_0$ then $C = \{t: |t| \geq t_{\frac{\alpha}{2}, n-1}\}$

❑ Without testing hypotheses by using critical region, we can also use P-value for testing the hypotheses for z test

P-value = $1 - \phi(z)$ for upper-tailed or right-tailed test ($H_a: \mu_a > \mu_0$)

= $\phi(z)$ for lower-tailed or left-tailed test ($H_a: \mu_a < \mu_0$)

= $2[1 - \phi(|z|)]$ for two-tailed test ($H_a: \mu_a \neq \mu_0$)

Decision for P-value

If P-value $\leq \alpha$, then reject H_0

If P-value $> \alpha$ not reject H_0

❑ β and sample size

Recall: β is probability of type II error

❑ To find the power π , we have $\pi(\mu') = 1 - \beta(\mu')$, where

-
1. If $H_a: \mu > \mu_0$ then $\beta(\mu') = \Phi\left(z_\alpha + \frac{\mu_0 - \mu'}{\frac{\sigma}{\sqrt{n}}}\right)$
 2. If $H_a: \mu < \mu_0$ then $\beta(\mu') = 1 - \Phi\left(-z_\alpha + \frac{\mu_0 - \mu'}{\frac{\sigma}{\sqrt{n}}}\right)$
 3. If $H_a: \mu \neq \mu_0$ then $\beta(\mu') = \Phi\left(\frac{z_\alpha}{2} + \frac{\mu_0 - \mu'}{\frac{\sigma}{\sqrt{n}}}\right) - \Phi\left(-\frac{z_\alpha}{2} + \frac{\mu_0 - \mu'}{\frac{\sigma}{\sqrt{n}}}\right)$

❑ Size determinant

1. For one tailed test $n = \left[\frac{\sigma(z_\alpha + z_\beta)}{\mu_0 - \mu'}\right]^2$
2. For two-tailed test $n = \left[\frac{\sigma\left(\frac{z_\alpha}{2} + z_\beta\right)}{\mu_0 - \mu'}\right]^2$, n is rounded up to the whole number

□ Test for a Variance and standard deviation

for random sample from normal distribution of size n with mean and variance are unknown

We use χ^2 test

Null hypotheses $H_0: \sigma^2 = \sigma_0^2$

And test statistic value $\chi^2 = \frac{(n-1)s^2}{\sigma_0^2}$

□ Determine the rejection region

1. If $H_a: \sigma^2 > \sigma_0^2$ then $C = \{\chi^2: \chi^2 \geq \chi_{\alpha, n-1}^2\}$
2. If $H_a: \sigma^2 < \sigma_0^2$ then $C = \{\chi^2: \chi^2 \leq \chi_{1-\alpha, n-1}^2\}$
3. If $H_a: \sigma^2 \neq \sigma_0^2$ then $C = \left\{ \chi^2: \chi^2 \leq \chi_{1-\frac{\alpha}{2}, n-1}^2 \text{ or } \chi^2 \geq \chi_{\frac{\alpha}{2}, n-1}^2 \right\}$

□ Test about a population proportion

❖ Assumption $x_1, \dots, x_n \sim \text{Bernulli}(p)$ sample, $x_i \in \{0,1\}$

then $X = \sum x_i \sim \text{Ber}(n, p)$, where $\hat{p} = \frac{x}{n}$ (probability of success)

We use z –test

➤ Null Hypotheses : $H_0: p = p_0$

➤ Test statistic value $z = \frac{\hat{p} - p_0}{\sqrt{\frac{p_0(1-p_0)}{n}}}$

➤ Critical region

1. For $H_a: p > p_0$ then $C = \{z: z \geq z_\alpha\}$

2. For $H_a: p < p_0$ then $C = \{z: z \leq -z_\alpha\}$

3. For $H_a: p \neq p_0$ then $C = \{z: z \leq -z_{\frac{\alpha}{2}} \text{ or } z \geq z_{\frac{\alpha}{2}}\} = \{z: |z| \geq z_{\frac{\alpha}{2}}\}$

Remark : test will provided that $np_0 \geq 5$ and $n(1 - p_0) \geq 5$

□ Beta and sample size

we have alternative hypotheses $\beta(p')$, where

$$1. \quad H_a: p > p_0 \text{ then } \beta(p') = \phi \left[\frac{p_0 - p' + z_{\alpha} \sqrt{\frac{p_0(1-p_0)}{n}}}{\sqrt{\frac{p'(1-p')}{n}}} \right]$$

$$2. \quad H_a: p < p_0 \text{ then } \beta(p') = 1 - \phi \left[\frac{p_0 - p' - z_{\alpha} \sqrt{\frac{p_0(1-p_0)}{n}}}{\sqrt{\frac{p'(1-p')}{n}}} \right]$$

$$3. \quad H_a: p \neq p_0 \text{ then } \beta(p') = \phi \left[\frac{p_0 - p' + \frac{z_{\alpha}}{2} \sqrt{\frac{p_0(1-p_0)}{n}}}{\sqrt{\frac{p'(1-p')}{n}}} \right] - \phi \left[\frac{p_0 - p' - \frac{z_{\alpha}}{2} \sqrt{\frac{p_0(1-p_0)}{n}}}{\sqrt{\frac{p'(1-p')}{n}}} \right]$$

Sample size

$$\text{For one tailed test: } n = \left[\frac{z_{\alpha} \sqrt{p_0(1-p_0)} + z_{\beta} \sqrt{p'(1-p')}}{p' - p_0} \right]^2$$

$$\text{For two tailed test : } n = \left[\frac{\frac{z_{\alpha}}{2} \sqrt{p_0(1-p_0)} + z_{\beta} \sqrt{p'(1-p')}}{p' - p_0} \right]^2$$

CH.5

Test for normal distribution with known the variance

□ The null hypotheses $H_0: \Delta_0 = \mu_1 - \mu_2$

□ We use z-test and test statistic value is $z = \frac{\bar{x} - \bar{y} - \Delta_0}{\sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}}$

□ We have the critical region , when

1. $H_a: \mu_1 - \mu_2 > \Delta_0$ then $C = \{z: z \geq z_\alpha\}$
2. $H_a: \mu_1 - \mu_2 < \Delta_0$ then $C = \{z: z \leq -z_\alpha\}$
3. $H_a: \mu_1 - \mu_2 \neq \Delta_0$ then $C = \{z: |z| \geq z_{\frac{\alpha}{2}}\}$

□ Alternative hypotheses $\beta(\Delta')$

Recall $\beta(\Delta') = P(\text{type II error})$, when $\Delta' = \mu_1 - \mu_2$ is defined by

1. If $H_a: \mu_1 - \mu_2 > \Delta_0$ then $\beta(\Delta') = \phi\left(z_\alpha - \frac{\Delta' - \Delta_0}{\sigma}\right)$
2. If $H_a: \mu_1 - \mu_2 < \Delta_0$ then $\beta(\Delta') = 1 - \phi\left(-z_\alpha - \frac{\Delta' - \Delta_0}{\sigma}\right)$
3. If $H_a: \mu_1 - \mu_2 \neq \Delta_0$ then $\beta(\Delta') = \phi\left(\frac{z_\alpha}{2} - \frac{\Delta' - \Delta_0}{\sigma}\right) - \phi\left(-\frac{z_\alpha}{2} - \frac{\Delta' - \Delta_0}{\sigma}\right)$

Where $\sigma = \sigma_{\bar{X} - \bar{Y}} = \sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}}$

Sample size

We use $\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n} = \frac{(\Delta' - \Delta_0)^2}{(z_\alpha + z_\beta)^2}$ for one -tailed test , where m and n is the sample size

We use $\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n} = \frac{(\Delta' - \Delta_0)^2}{\left(\frac{z_\alpha}{2} + z_\beta\right)^2}$ for two tailed test

Large Sample Tests

□ A assumption normal distribution when $n \geq 30$ and $m \geq 30$,

from Central limit theorem

□ Test statistic value $z = \frac{\bar{x} - \bar{y} - \Delta_0}{\sqrt{\frac{S_1^2}{m} + \frac{S_2^2}{n}}}$, where $\Delta_0 = \mu_1 - \mu_2$ (null hypotheses)

□ Confidence Interval for $\mu_1 - \mu_2$

$$\text{we get } CI = \left[\bar{x} - \bar{y} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\sigma_1^2}{m} + \frac{\sigma_2^2}{n}} \right]$$

$$\text{For Large sample } CI = \left[\bar{x} - \bar{y} \pm z_{\frac{\alpha}{2}} \sqrt{\frac{S_1^2}{m} + \frac{S_2^2}{n}} \right]$$

T-test for two sample

When distribution are both normal and $\sigma_X^2 \neq \sigma_Y^2$, n and m < 30

□ Null hypotheses $H_0: \Delta_0 = \mu_1 - \mu_2$

□ Test statistic value $T = \frac{\bar{x} - \bar{y} - \Delta_0}{\sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}}}$

□ The critical region is defined by

1. $H_a: \mu_1 - \mu_2 > \Delta_0$ then $C = \{t: t \geq t_{\alpha, \nu}\}$
2. $H_a: \mu_1 - \mu_2 < \Delta_0$ then $C = \{t \leq -t_{\alpha, \nu}\}$
3. $H_a: \mu_1 - \mu_2 \neq \Delta_0$ then $C = \{t: |t| \geq t_{\frac{\alpha}{2}, \nu}\}$

Where degree of freedom $\nu = \frac{\left(\frac{s_1^2}{m} + \frac{s_2^2}{n}\right)^2}{\frac{\left(\frac{s_1^2}{m}\right)^2}{m-1} + \frac{\left(\frac{s_2^2}{n}\right)^2}{n-1}}$ (rounded it to nearest integer)

□ For $100(1 - \alpha)\%$ confidence interval is defined by

$$CI = \left[\bar{x} - \bar{y} \pm t_{\frac{\alpha}{2}, \nu} \sqrt{\frac{s_1^2}{m} + \frac{s_2^2}{n}} \right]$$

□ Pooled t procedures

❖ Alternative to the two sample t procedures assuming not only normal distributed and we don't know the variance for the population but we know that $\sigma_1^2 = \sigma_2^2$

❖ The standardized variable $T = \frac{\bar{X} - \bar{Y} - (\mu_1 - \mu_2)}{S_p \sqrt{\frac{1}{m} + \frac{1}{n}}} \sim t$ distributed with df $\nu = m + n - 2$

Where, S_p is called the pooled estimator or standard deviation of σ and defined by

$$S_p = \sqrt{\frac{(m-1)s_1^2 + (n-1)s_2^2}{m+n-2}}$$

□ Analysis for Paired Data

❖ Using one sample t-test

➤ Null Hypotheses $H_0: \mu_D = \Delta_0$

➤ D_i is the different $d = \text{before} - \text{after} = X_i - Y_i$

Test statistic value $t = \frac{\bar{d} - \Delta_0}{\frac{s_D}{\sqrt{n}}}$

➤ Alternative hypotheses and the critical region

1. When $H_a: \mu_D > \Delta_0$ then $C = \{t: t \geq t_{\alpha, n-1}\}$

2. $H_a: \mu_D < \Delta_0$ then $C = \{t: t \leq -t_{\alpha, n-1}\}$

3. $H_a: \mu_D \neq \Delta_0$ then $C = \{t: |t| \geq t_{\frac{\alpha}{2}, n-1}\}$

□ Confidence Interval for the Mean Difference

The paired t CI for μ_D is : $CI = \left[\bar{d} \pm t_{\frac{\alpha}{2}, n-1} \times \frac{S_D}{\sqrt{n}} \right]$

□ Inferences about two population variance

➤ X_1, \dots, X_n be a random from normal distribution with variance σ_1^2 , and let Y_1, \dots, Y_n be a random sample from normal distribution with variance σ_2^2 and let S_1^2 and S_2^2 denote the two sample variance .

❖ we use Fisher distribution F that is $F = \frac{\frac{S_1^2}{\sigma_1^2}}{\frac{S_2^2}{\sigma_2^2}}$ with df $\nu_1 = m - 1$ and $\nu_2 = n - 1$

❑ The test for equality of Variances

➤ Use Fisher distribution F

➤ Null hypotheses $H_0: \sigma_1^2 = \sigma_2^2$

➤ Test statistic value : $f = \frac{S_1^2}{S_2^2}$

➤ Alternative Hypotheses and rejection region

1. When $H_a: \sigma_1^2 > \sigma_2^2$ then C or $RR = \{f: f \geq F_{\alpha, m-1, n-1}\}$

2. When $H_a: \sigma_1^2 < \sigma_2^2$ then $C = \{f: f \leq F_{1-\alpha, m-1, n-1}\}$

3. When $H_a: \sigma_1^2 \neq \sigma_2^2$ then $C = \{f: f \geq F_{\frac{\alpha}{2}, m-1, n-1} \text{ or } f \leq F_{1-\frac{\alpha}{2}, m-1, n-1}\}$

□ A confidence Interval for σ_1^2/σ_2^2 and σ_1/σ_2

➤ A $100(1 - \alpha)\%$ CI for σ_1^2/σ_2^2 is $CI = \left[\frac{s_1^2}{s_2^2} \cdot \frac{1}{F_{\frac{\alpha}{2}, m-1, n-1}}, \frac{s_1^2}{s_2^2} \cdot F_{\frac{\alpha}{2}, n-1, m-1} \right]$

➤ A random $100(1 - \alpha)\%$ CI for σ_1/σ_2 is

$$CI = \left[\frac{s_1}{s_2} \cdot \frac{1}{\sqrt{F_{\frac{\alpha}{2}, m-1, n-1}}}, \frac{s_1}{s_2} \sqrt{F_{\frac{\alpha}{2}, n-1, m-1}} \right]$$

□ Inferences about two population proportion

- Test statistic value $Z = \frac{\hat{p}_1 - \hat{p}_2 - (p_1 - p_2)}{\sqrt{\frac{p_1 q_1}{m} + \frac{p_2 q_2}{n}}}$ approximate to standard normal
-

- For null hypotheses $H_0: p_1 = p_2$, then

- the test Statistic value $z = \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}\hat{q}\left(\frac{1}{m} + \frac{1}{n}\right)}}$, where

$$\hat{p}_1 = \frac{x}{m} \text{ and } \hat{p}_2 = \frac{y}{n}, \hat{p} = \frac{x+y}{m+n}, \hat{q} = 1 - \hat{p}$$

For test safely be used as long as $m\hat{p}_1, m\hat{q}_1, n\hat{p}_2, n\hat{q}_2$ are all at least 5 (≥ 5)

□ Alternative hypotheses and Rejection region

1. $H_a: p_1 - p_2 > 0$ then $R = \{z: z \geq z_\alpha\}$
2. $H_a: p_1 - p_2 < 0$ then $R = \{z: z \leq -z_\alpha\}$
3. $H_a: p_1 - p_2 \neq 0$ then $R = \{z: |z| \geq z_{\frac{\alpha}{2}}\}$

□ Alternative Hypotheses $\beta(p_1, p_2)$, if

1. $H_a: p_1 - p_2 > 0$ then $\beta(p_1, p_2) = \phi \left[\frac{z_\alpha \sqrt{\bar{p}\bar{q}\left(\frac{1}{m} + \frac{1}{n}\right)} - (p_1 - p_2)}{\sigma} \right]$

2. $H_a: p_1 - p_2 < 0$ then $\beta(p_1, p_2) = 1 - \phi \left[\frac{-z_\alpha \sqrt{\bar{p}\bar{q}\left(\frac{1}{m} + \frac{1}{n}\right)} - (p_1 - p_2)}{\sigma} \right]$

3. $H_a: p_1 - p_2 \neq 0$ then $\beta(p_1, p_2) = \phi \left[\frac{z_\alpha \sqrt{\bar{p}\bar{q}\left(\frac{1}{m} + \frac{1}{n}\right)} - (p_1 - p_2)}{\sigma} \right] - \phi \left[\frac{-\frac{z_\alpha}{2} \sqrt{\bar{p}\bar{q}\left(\frac{1}{m} + \frac{1}{n}\right)} - (p_1 - p_2)}{\sigma} \right]$

Where $\bar{p} = \frac{(mp_1 + np_2)}{(m+n)}$, $\bar{q} = \frac{mq_1 + nq_2}{m+n}$ and $\sigma = \sigma_{\hat{p}_1 - \hat{p}_2} = \sqrt{\frac{p_1 q_1}{m} + \frac{p_2 q_2}{n}}$

□ For sample size

➤ In case $m = n$ and $d = p_1 - p_2$ then

■ Sample size $n = \frac{\left[z_{\alpha} \sqrt{\frac{(p_1+p_2)(q_1+q_2)}{2}} + z_{\beta} \sqrt{p_1 q_1 + p_2 q_2} \right]^2}{d^2}$ for one tailed test

■ For two-tailed test $n = \frac{\left[z_{\frac{\alpha}{2}} \sqrt{\frac{(p_1+p_2)(q_1+q_2)}{2}} + z_{\beta} \sqrt{p_1 q_1 + p_2 q_2} \right]^2}{d^2}$

□ A large sample confidence interval

■ A $100(1 - \alpha)\%$ CI for $p_1 - p_2$ is

$$CI = \left[\hat{p}_1 - \hat{p}_2 \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\hat{p}_1 \hat{q}_1}{m} + \frac{\hat{p}_2 \hat{q}_2}{n}} \right]$$