New stochastic sketching methods for Big Data Ridge Regression

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July, 2017

Abstract

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1. Randomized Newton Method

1.1 Algorithm

1.2 Convergence rate (draft)

1.2.1 General case

$$\lambda_{min}(A^{-\frac{1}{2}}E[Z]A^{-\frac{1}{2}})\geqslant \lambda_{min}(\mathbf{C}^TA\mathbf{C})\lambda_{min}(D^2)$$

$$\lambda_{min}(D^2)=\min_i p_i \frac{1}{\lambda_{max}(I_{C_i}AI_{C_i}^T)}=\min_i \frac{p_i}{\lambda_{max}(I_{C_i}^TI_{C_i})\lambda_{max}(A)}\geqslant \frac{p}{\lambda_{max}(A)}, \text{ since for any } i\in\{1,\ldots,n\}, \text{ for any } x \text{ in } \mathbb{R}^n\left\langle I_{C_i}^TI_{C_i}x \,|\, x\right\rangle=\|I_{C_i}x\|^2\leqslant \|x\|^2 \text{ and then } \lambda_{max}(I_{C_i}^TI_{C_i})\leqslant 1.$$

Therefore,
$$\lambda_{min}(A^{-\frac{1}{2}}E[Z]A^{-\frac{1}{2}})\geqslant p\frac{\lambda_{min}(\mathbf{C}^TA\mathbf{C})}{\lambda_{max}(A)}=p\frac{\lambda_{min}(A)\lambda_{min}(\mathbf{C}\mathbf{C}^T)}{\lambda_{max}(A)}\geqslant p\frac{\lambda_{min}(A)}{\lambda_{max}(A)}$$

1.2.2 Uniform case

$$Z = AI_C^T (I_C AI_C^T)^{-1} I_C A$$

$$\rho = 1 - \lambda_{min} (A^{-\frac{1}{2}} E[Z] A^{-\frac{1}{2}}).$$

$$A^{-\frac{1}{2}} E[Z] A^{-\frac{1}{2}} = \sum_{i} p_{i} A^{\frac{1}{2}} I_{C_{i}}^{T} (I_{C_{i}} A I_{C_{i}}^{T})^{-1} I_{C_{i}} A^{\frac{1}{2}}$$

for any $i \in \{1, ..., n\}$, $A^{\frac{1}{2}}I_{C_i}^T(I_{C_i}AI_{C_i}^T)^{-1}I_{C_i}A^{\frac{1}{2}}$ is a projection matrix and then its eigenvalues are a nonempty subset of $\{0, 1\}$.

Since λ_{max} is convex, we obtain that :

$$0 \leqslant \lambda_{min}(A^{-\frac{1}{2}}E[Z]A^{-\frac{1}{2}}) \leqslant \lambda_{max}(A^{-\frac{1}{2}}E[Z]A^{-\frac{1}{2}}) \leqslant \sum_{i} p_{i}\lambda_{max}(A^{\frac{1}{2}}I_{C_{i}}^{T}(I_{C_{i}}AI_{C_{i}}^{T})^{-1}I_{C_{i}}A^{\frac{1}{2}}) \leqslant 1.$$

$$\mathbf{C} = (I_{C_{1}}^{T}, \dots, I_{C_{-}}^{T}).$$

 $A^{-\frac{1}{2}}E[Z]A^{-\frac{1}{2}}=(A^{\frac{1}{2}}\mathbf{C}D)(D\mathbf{C}^TA^{\frac{1}{2}}) \text{ where } D=\sqrt{p}\operatorname{diag}((S_1AS_1^T)^{-\frac{1}{2}},\ldots,(S_rAS_r^T)^{-\frac{1}{2}}) \text{ where } p=\frac{1}{\binom{n}{s}} \text{ is the uniform probability of choosing } s \text{ rows uniformly on } \{1,\ldots,n\}, \text{ knowing that } s \text{ is the sketchsize.}$

Proposition 1.2.1

$$\lambda_{min}(A^{-\frac{1}{2}}E[Z]A^{-\frac{1}{2}}) \geqslant \frac{1}{\binom{n}{s}} \frac{\lambda_{min}(A)}{\lambda_{max}(A)}$$

$$\lambda_{min}(A^{-\frac{1}{2}}E[Z]A^{-\frac{1}{2}})\geqslant \lambda_{min}(\mathbf{C}^TA\mathbf{C})\lambda_{min}(D^2)\\ \lambda_{min}(D^2) = p\min_{i}\frac{1}{\lambda_{max}(I_{C_i}AI_{C_i}^T)} = p\min_{i}\frac{1}{\lambda_{max}(I_{C_i}^TI_{C_i})\lambda_{max}(A)}\geqslant \frac{p}{\lambda_{max}(A)}, \text{ since for any } i\in\{1,\dots,n\}, \text{ for any } x \text{ in } \mathbb{R}^n\left\langle I_{C_i}^TI_{C_i}x \,|\, x\right\rangle = \|I_{C_i}x\|^2\leqslant \|x\|^2 \text{ and then } \lambda_{max}(I_{C_i}^TI_{C_i})\leqslant 1.$$

Therefore,
$$\lambda_{min}(A^{-\frac{1}{2}}E[Z]A^{-\frac{1}{2}})\geqslant p\frac{\lambda_{min}(\mathbf{C}^TA\mathbf{C})}{\lambda_{max}(A)}=p\frac{\lambda_{min}(A)\lambda_{min}(\mathbf{C}\mathbf{C}^T)}{\lambda_{max}(A)}\geqslant p\frac{\lambda_{min}(A)}{\lambda_{max}(A)}$$

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Hadamard Sketches

Algorithm 2.1

2.2 **Convergence rate (draft)**

$$\begin{split} Z &= AS^T(SAS^T)^{-1}SA \\ &\rho = 1 - \lambda_{min}(A^{-\frac{1}{2}}E[Z]A^{-\frac{1}{2}}) \\ &S_i = I_{C_i}H. \\ &\tilde{A} = \frac{1}{n}HAH^T = \frac{H}{\sqrt{n}}A\frac{H^T}{\sqrt{n}}. \\ &A^{-\frac{1}{2}}E[Z]A^{-\frac{1}{2}} = E[A^{\frac{1}{2}}S^T(SAS^T)^{-1}SA^{\frac{1}{2}}] = \sum_i p_iA^{\frac{1}{2}}H^TI_{C_i}^T(I_{C_i}HAH^TI_{C_i}^T)^{-1}I_{C_i}HA^{\frac{1}{2}} \\ &= A^{\frac{1}{2}}H^T\frac{1}{n}E[I_C^T(I_C\tilde{A}I_C^T)^{-1}I_C]HA^{\frac{1}{2}} = H^{-1}\tilde{A}^{\frac{1}{2}}E[I_C^T(I_C\tilde{A}I_C^T)^{-1}I_C]\tilde{A}^{\frac{1}{2}}(\frac{H^T}{n})^{-1} = H^{-1}\tilde{A}^{\frac{1}{2}}E[I_C^T(I_C\tilde{A}I_C^T)^{-1}I_C]\tilde{A}^{\frac{1}{2}}H. \\ &\text{Hence}: \end{split}$$

$$\rho = 1 - \lambda_{min} (A^{-\frac{1}{2}} E[Z] A^{-\frac{1}{2}}) = 1 - \lambda_{min} (\tilde{A}^{\frac{1}{2}} E[I_C^T (I_C \tilde{A} I_C^T)^{-1} I_C] \tilde{A}^{\frac{1}{2}})$$

 $\rho=1-\lambda_{min}(A^{-\frac{1}{2}}E[Z]A^{-\frac{1}{2}})=1-\lambda_{min}(\tilde{A}^{\frac{1}{2}}E[I_C^T(I_C\tilde{A}I_C^T)^{-1}I_C]\tilde{A}^{\frac{1}{2}})$ We recognize the convergence rate in the Randomized Newton Method and then, denoting by $ho_{Newton}(M)$ the convergence rate of the Newton method associated with the definite positive matrix M, we obtain that :

$$\rho = 1 - \lambda_{min}(\tilde{A}^{\frac{1}{2}}E[I_C^T(I_C\tilde{A}I_C^T)^{-1}I_C]\tilde{A}^{\frac{1}{2}}) = 1 - (1 - \rho_{Newton}(\tilde{A})) = \rho_{Newton}(\tilde{A})$$

3. Count-min Sketches

- 3.1 Algorithm
- 3.2 Convergence rate

$$Z = AS^T (SAS^T)^{-1} SA$$

4. Conclusion

References

[1] ROBERT GOWER AND PETER RICHTARIK, <u>Randomized iterative methods for linear systems</u>, SIAM, (2015).