

1. X, Y independent random variables $X \sim N(0, 1)$ $Y \sim U[2, 3]$

a) Find joint density $p_{XY}(x, y)$

X, Y independent $p_{XY}(x, y) = p_X(x) p_Y(y)$

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) \quad p_Y(y) = \begin{cases} 1 & y \in [2, 3] \\ 0 & \text{else} \end{cases}$$

$$p_{XY}(x, y) = \begin{cases} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right) & y \in [2, 3], \forall x \\ 0 & \text{else} \end{cases}$$

b) Find the conditional density $p_{X|Y=y}(x|y)$

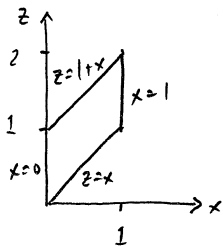
$$p_{X|Y}(x|y) = \frac{p_{XY}(x, y)}{p_Y(y)} = \frac{p_X(x) p_Y(y)}{p_Y(y)} = p_X(x)$$

$$p_{X|Y}(x|y) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$$

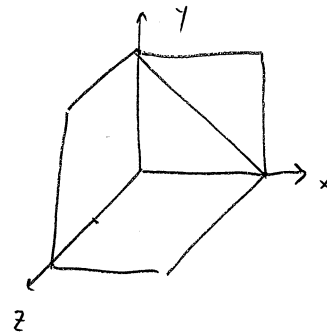
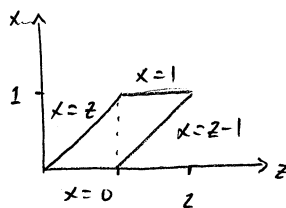
2 Let X, Y be independent $X \sim U[0, 1]$ $Y \sim U[0, 1]$

$$Z = X + Y$$

a) Find $p_{XZ}(x, z)$



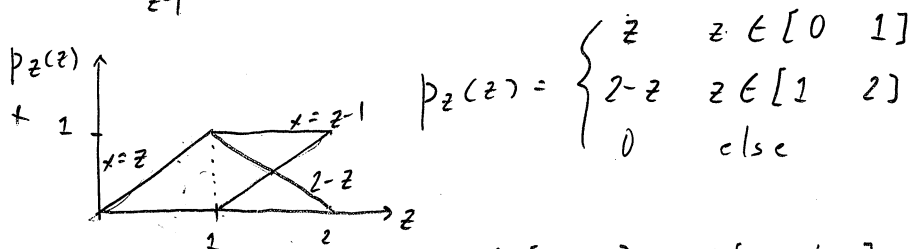
$$p_{XZ} = \begin{cases} 1 & x \in [0, 1], z \in [x, 1+x] \\ 0 & \text{else} \end{cases}$$



$$p_Z(z) = \int_x p_{XZ} dx$$

$$p_Z(z) \int_x 1 \quad x \in [0, 1], z \in [x, 1+x] dx \Rightarrow \int 1 \quad z \in [0, 2], x \in [z-1, z] dx$$

$$= \int_{z-1}^z 1 dx = x \Big|_{z-1}^z = 1 \quad z \in [0, 2], x \in [z-1, z]$$

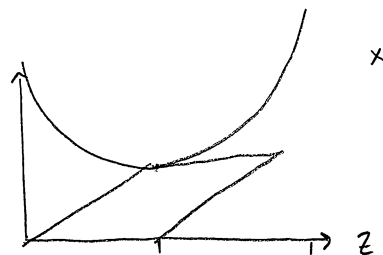


$$p_Z(z) = \begin{cases} z & z \in [0, 1] \\ 2-z & z \in [1, 2] \\ 0 & \text{else} \end{cases}$$

$$p_{X|Z} = \frac{p_{XZ}}{p_Z} = \begin{cases} 1 & x \in [0, 1], z \in [x, 1+x] \\ 0 & \text{else} \end{cases}$$

$$\begin{cases} \frac{1}{z} & z \in [0, 1] \\ \frac{1}{2-z} & z \in [1, 2] \\ 0 & \text{else} \end{cases}$$

$$= \begin{cases} \frac{1}{z} & z \in [0, 1], x \in [0, z] \\ \frac{1}{2-z} & z \in [1, 2], x \in [z-1, 1] \\ 0 & \text{else} \end{cases}$$



$$c) E\{X|Z\} = \int_x x p_{X|Z}(X|Z) dx =$$

$$\int_0^z x \frac{1}{2} dx = \frac{1}{2} x^2 \frac{1}{2} \Big|_0^z = \frac{1}{2} z^2 z^{-1} = \frac{z}{2}$$

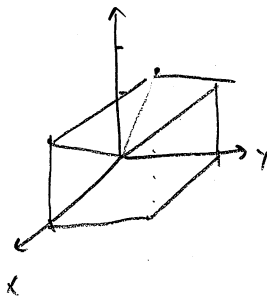
$$\int_{z-1}^1 x \frac{1}{2-z} dx = \frac{1}{2} x^2 \frac{1}{2-z} \Big|_{z-1}^1 = \frac{1}{2} \left(\frac{1}{2-z} \right) (1 - z^2 + 2z - 1)$$

$$= \frac{1}{2} \frac{2z - z^2}{2-z} = \frac{z}{2}$$

$$d) E\{Z|X\} = E\{X+Y|X\} = E\{X|X\} + E\{Y|X\}$$

↑ independent

$$= X + \frac{1}{2}$$



3 Let X, Y be RV

$$\begin{bmatrix} X \\ Y \end{bmatrix} \sim N\left(\begin{bmatrix} 1 \\ -1 \end{bmatrix}, \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}\right)$$

a) Find $p_X(x)$

$$X = [1 \ 0] \begin{bmatrix} X \\ Y \end{bmatrix} \quad m_X = [1 \ 0] \begin{bmatrix} 1 \\ -1 \end{bmatrix} = 1$$

$$\sigma_X^2 = [1 \ 0] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} 1 \\ 0 \end{bmatrix} = 1$$

$$X \sim N(1, 1)$$

b) Find $p_R(r)$ $R = X + Y$

$$R = [1 \ 1] \begin{bmatrix} X \\ Y \end{bmatrix} \quad m_R = [1 \ 1] \begin{bmatrix} 1 \\ -1 \end{bmatrix} = 0$$

$$\sigma_R^2 = [1 \ 1] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} 1 \\ 1 \end{bmatrix} = 2$$

$$R \sim N(0, 2)$$

c) Find $p_S(s)$ $S = X - Y$

$$S = [1 \ -1] \begin{bmatrix} X \\ Y \end{bmatrix} \quad m_S = [1 \ -1] \begin{bmatrix} 1 \\ -1 \end{bmatrix} = 2$$

$$\sigma_S^2 = [1 \ -1] \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} 1 \\ -1 \end{bmatrix} = 2$$

$$S \sim N(2, 2)$$

d) Find $E\{R|S=s\}$ $Z = [X \ Y \ R \ S]^T$

$$Z = A \begin{bmatrix} X \\ Y \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \\ 1 & 1 \\ 1 & -1 \end{bmatrix} \begin{bmatrix} X \\ Y \end{bmatrix} \quad m_Z = \begin{bmatrix} 1 \\ -1 \\ 0 \\ 2 \end{bmatrix}$$

$$\Lambda_{ZZ} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \\ 1 & 1 \\ 1 & -1 \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix} \begin{bmatrix} 1 & 0 & 1 & 1 \\ 0 & 1 & 1 & -1 \end{bmatrix} = \begin{bmatrix} 1 & 0 & 1 & 1 \\ 0 & 1 & 1 & -1 \\ 1 & 1 & 2 & 0 \\ 1 & -1 & 0 & 2 \end{bmatrix}$$

$$\sigma_{RS} = 0$$

$$\begin{bmatrix} R \\ S \end{bmatrix} \sim N\left(\begin{bmatrix} 0 \\ 2 \end{bmatrix}, \begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}\right)$$

$$p_{RS}(r, s) = \frac{1}{(2\pi) \left(\begin{vmatrix} 2 & 0 \\ 0 & 2 \end{vmatrix} \right)^{\frac{1}{2}}} \exp \left(-\frac{1}{2} [(r-0)(s-2)] \begin{bmatrix} 2 & 0 \\ 0 & 2 \end{bmatrix}^{-1} \begin{bmatrix} r \\ s-2 \end{bmatrix} \right)$$

$$= \frac{1}{2\pi(2)} \exp \left(-\frac{1}{2} [r (s-2)] \begin{bmatrix} \frac{1}{2} & 0 \\ 0 & \frac{1}{2} \end{bmatrix} \begin{bmatrix} r \\ s-2 \end{bmatrix} \right)$$

$$= \frac{1}{4\pi} \exp \left(-\frac{1}{4} (r^2 + (s-2)^2) \right)$$

$$p_s = \frac{1}{\sqrt{2\pi}(2)} \exp \left(-\frac{1}{2} \frac{(s-2)^2}{2} \right) = \frac{1}{2\sqrt{\pi}} \exp \left(-\frac{1}{4} (s-2)^2 \right)$$

$$p_{R|S=s}(r|s) = \frac{\frac{1}{4\pi} \exp \left(-\frac{1}{4} (r^2 + (s-2)^2) \right)}{\frac{1}{2\sqrt{\pi}} \exp \left(-\frac{1}{4} (s-2)^2 \right)}$$

$$= \frac{1}{2\sqrt{\pi}} \exp \left(-\frac{1}{4} r^2 \right) = \frac{1}{\sqrt{2\pi} \cdot 2} \exp \left(-\frac{1}{4} r^2 \right)$$

$$\Rightarrow p_{R|S}(r|s) \sim \mathcal{N}(0, 2)$$

Also uncorrelated \Rightarrow independence if Gaussian

$$\Rightarrow p_{R|S}(r|s) = p_R(r)$$

~~~~~

$$E(R|S=s) = m_R + \Lambda_{RS} \Lambda_{SS}^{-1} (s - m_S) = m_R = 0$$

4 Can  $\Lambda = \begin{bmatrix} 2 & 3 \\ 3 & 4 \end{bmatrix}$  be a covariance matrix?

$$\Lambda = \Lambda^* \quad |\Lambda_{11}| = 2 > 0 \quad |\Lambda_{22}| = 8 - 9 < 0 \Rightarrow \Lambda \neq 0$$

5 Let  $X_1, \dots, X_n$  be IID RV with mean  $m$  covariance  $\sigma^2$

$$Z = \frac{1}{n} \sum_{i=1}^n X_i \quad \begin{bmatrix} X_1 \\ \vdots \\ X_i \\ \vdots \\ X_n \end{bmatrix} \sim \begin{bmatrix} m \\ \vdots \\ m \\ \vdots \\ m \end{bmatrix}, \begin{bmatrix} \sigma^2 & & \\ & \sigma^2 & \\ & & \ddots \\ & & & \sigma^2 \end{bmatrix}$$

a) Compute  $E\{Z\}$

$$E\{Z\} = E\left\{\frac{1}{n} \sum_{i=1}^n X_i\right\} = \frac{1}{n} E\left\{\sum_{i=1}^n X_i\right\} = \frac{1}{n} \sum_{i=1}^n E\{X_i\}$$

$$= \frac{1}{n} \sum_{i=1}^n m = \frac{1}{n} nm = m$$

b) Compute  $E\{Z^2\}$

$$E\{Z^2\} = E\left\{\left(\frac{1}{n} \sum X_i\right) \left(\frac{1}{n} \sum X_i\right)^*\right\} = \frac{1}{n^2} E\left\{\left(\sum X_i\right) \left(\sum X_i\right)\right\}$$

$$= \frac{1}{n^2} E\left\{(X_1 + \dots + X_n)(X_1 + \dots + X_n)\right\} = \frac{1}{n^2} E\{X_1^2 + \dots + X_n^2 + 2X_1X_2 + \dots\}$$

$$= \frac{1}{n^2} E\left\{\sum_{i=1}^n \sum_{j=1}^n X_i X_j\right\} = \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n E\{X_i X_j\}$$

$$= \frac{1}{n^2} \left[ n(\sigma^2 + m^2) \right] + \frac{1}{n^2} \sum_{i=1}^n \sum_{j=1}^n E\{(1 - \delta_{ij}) X_i X_j\}$$

$$= \frac{1}{n} (\sigma^2 + m^2) + \frac{1}{n^2} 2nm^2 = \frac{1}{n} (\sigma^2 + 3m^2) ?$$

5  $X_1, \dots, X_n$  IID variables mean  $m$ , variance  $\sigma^2$

$$Z = \frac{1}{n} \sum_{i=1}^n X_i \quad \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{bmatrix} \sim \begin{bmatrix} m \\ m \\ \vdots \\ m \end{bmatrix}, \sigma^2 I_n$$

$$X = \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{bmatrix}$$

a) Compute  $E\{Z\}$

$$Z = \frac{1}{n} \sum_{i=1}^n X_i = \frac{1}{n} [1 \dots 1] \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{bmatrix} = \frac{1}{n} a^* X$$

$$E\{Z\} = E\left\{\frac{1}{n} a^* X\right\} = \frac{a^*}{n} E\{X\} = \frac{a^*}{n} \begin{bmatrix} m \\ m \\ \vdots \\ m \end{bmatrix} = \frac{nm}{n} = m$$

b) Compute  $E\{Z^2\}$

$$\begin{aligned} E\{Z^2\} &= E\left\{\frac{1}{n} a^* X X^* a \frac{1}{n}\right\} = \frac{1}{n^2} a^* E\{X X^*\} a \\ &= \frac{1}{n^2} a^* \sigma^2 I a = \frac{\sigma^2 n}{n^2} = \frac{\sigma^2}{n} \end{aligned}$$

6. Define  $X \sim U[0, 1]$ ,  $Y = X^2$ ,  $Z = [X \ Y]^T$

Calculate  $\Lambda_{ZZ}$   $p_X = 1$   $0 \leq x \leq 1$

$$m_X = \int_{-\infty}^{+\infty} x \cdot 1 \, dx = \frac{1}{2}$$

$$m_Y = E\{Y\} = \int_{-\infty}^{+\infty} x^2 \cdot 1 \, dx = \frac{1}{3}$$

$$\Lambda_{ZZ} = E\left\{ \begin{bmatrix} X - \frac{1}{2} \\ Y - \frac{1}{3} \end{bmatrix} \begin{bmatrix} X - \frac{1}{2} & Y - \frac{1}{3} \end{bmatrix} \right\} = E\left\{ \begin{bmatrix} (X - \frac{1}{2})^2 & (X - \frac{1}{2})(Y - \frac{1}{3}) \\ (Y - \frac{1}{3})(X - \frac{1}{2}) & (Y - \frac{1}{3})^2 \end{bmatrix} \right\}$$

$$E\left\{ (X - \frac{1}{2})^2 \right\} = \int_0^1 (x - \frac{1}{2})^2 \, dx = \frac{1}{3} (x - \frac{1}{2})^3 \Big|_0^1 = \frac{1}{3} \left( \frac{1}{4} + \frac{1}{4} \right) = \frac{1}{12}$$

$$E\left\{ (X - \frac{1}{2})(Y - \frac{1}{3}) \right\} = \int_0^1 (x - \frac{1}{2})(x^2 - \frac{1}{3}) \, dx = \int_0^1 x^3 - \frac{1}{2}x^2 - \frac{1}{3}x + \frac{1}{6} \, dx$$

$$= \frac{1}{4}x^4 - \frac{1}{6}x^3 - \frac{1}{6}x^2 + \frac{1}{6}x \Big|_0^1 = \frac{1}{4} - \frac{1}{6} - \frac{1}{6} + \frac{1}{6} = \frac{1}{12}$$

$$E\left\{ (Y - \frac{1}{3})^2 \right\} = \int_0^1 (x^2 - \frac{1}{3})^2 \, dx = \int_0^1 x^4 - \frac{2}{3}x^2 + \frac{1}{9} \, dx$$

$$= \frac{1}{5}x^5 - \frac{2}{9}x^3 + \frac{1}{9}x \Big|_0^1 = \frac{1}{5} - \frac{2}{9} + \frac{1}{9} = \frac{4}{45}$$

$$\Lambda_{ZZ} = \begin{bmatrix} \frac{1}{12} & \frac{1}{12} \\ \frac{1}{12} & \frac{4}{45} \end{bmatrix}$$



7. Let  $X \sim U[0, 1]$ . Find  $f: \mathbb{R} \rightarrow \mathbb{R}$  so that  
 $Y = f(X) \sim N(0, \sigma^2)$

$$p_X(x) = \begin{cases} 1 & 0 \leq x \leq 1 \\ 0 & \text{else} \end{cases}$$

$$p_Y(y) = \frac{1}{(2\pi\sigma^2)^{\frac{1}{2}}} \exp\left(-\frac{y^2}{2\sigma^2}\right)$$

$$p_Y(y) = p_X(f^{-1}(y)) \left| \frac{df^{-1}(y)}{dy} \right|$$

$$p_Y(y) = p_X(g(y)) \left| \frac{dg(y)}{dy} \right|$$

$$\frac{1}{(2\pi\sigma^2)^{\frac{1}{2}}} \exp\left(-\frac{y^2}{2\sigma^2}\right) = 1 \left| \frac{d}{dy} g(y) \right|$$

$$g(y) = \int \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{y^2}{2\sigma^2}\right) dy = \frac{1}{\sqrt{2\pi\sigma^2}} \frac{\sqrt{\pi}}{\sqrt{2}} \operatorname{erf}\left(\frac{y}{\sqrt{2}\sigma}\right)$$

$$g(y) = \frac{1}{2\sigma^2} \operatorname{erf}\left(\frac{y}{\sqrt{2}\sigma}\right) = x$$

$$\operatorname{erf}\left(\frac{y}{\sqrt{2}\sigma}\right) = 2\sigma^2 x$$

$$\Rightarrow y = \sqrt{2}\sigma \operatorname{erf}^{-1}(2\sigma^2 x)$$

$$Y = f(X) \quad X = f^{-1}(Y)$$

$$X = g(Y) \quad Y = g^{-1}(X)$$

$$X = f^{-1}(g^{-1}(X))$$

$$\Rightarrow g^{-1} = f(X)$$

8 Let  $X_1, \dots, X_n$  be IID RV

$$\begin{bmatrix} X_1 \\ \vdots \\ X_n \end{bmatrix} \sim \begin{bmatrix} m \\ \vdots \\ m \end{bmatrix}, \begin{bmatrix} \sigma^2 & & 0 \\ & \ddots & \\ 0 & & \sigma^2 \end{bmatrix}$$

a) Compute  $E\{X_1 | X_1 + X_2 + \dots + X_n = s\}$

$$s = [1 \ 1 \ \dots \ 1] \begin{bmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{bmatrix} \Rightarrow m_s = nm$$

$$\sigma_s^2 = [1 \ 1 \ \dots \ 1] \sigma^2 I \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix} = n\sigma^2$$

$$m_s = nm \quad \sigma_s^2 = n\sigma^2$$

$$Z = \begin{bmatrix} X_1 \\ \vdots \\ X_n \\ s \end{bmatrix} = \begin{bmatrix} I_n \\ 1 \ 1 \ \dots \ 1 \end{bmatrix} \begin{bmatrix} X_1 \\ \vdots \\ X_n \end{bmatrix} \quad m_Z = \begin{bmatrix} m \\ m \\ \vdots \\ m \\ nm \end{bmatrix}$$

$$\Lambda_{ZZ} = \begin{bmatrix} \sigma^2 & & & & \\ & \sigma^2 & & & \\ & & \ddots & & \\ & & & \sigma^2 & \\ \sigma^2 & \dots & \sigma^2 & & n\sigma^2 \end{bmatrix} \quad S = \sum_{i=1}^n X_i$$

$$\text{IID} \Rightarrow E\{X_1 | S=s\} = E\{X_2 | S=s\} = \dots = E\{X_n | S=s\} = \gamma$$

$$n\gamma = \sum_{i=1}^n E\{X_i | S=s\} = E\left\{\sum_{i=1}^n X_i | S=s\right\} = E\{X_1 + X_2 + \dots + X_n | S=s\} = s$$

$$n\gamma = s \Rightarrow \gamma = s/n$$

b) Compute for  $h \leq n$   $E\left\{\frac{X_1 + \dots + X_h}{X_1 + \dots + X_n}\right\}$

$$E\left\{\frac{X_1 + \dots + X_h}{s}\right\} = \alpha$$

1 Let  $X, Y$  be RV with joint density

$$p_{XY}(x, y) = \frac{1}{2\pi} \exp\left(-\frac{1}{2}(x^2 + y^2)\right)$$

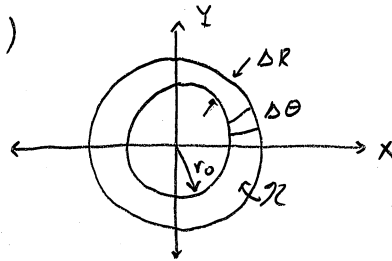
$$R^2 = X^2 + Y^2$$

$$\begin{bmatrix} X \\ Y \end{bmatrix} \sim \mathcal{N}\left(\begin{bmatrix} 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}\right)$$

Define  $R > 0$   $\theta \in [0, 2\pi)$

$$X = R \cos \theta, \quad Y = R \sin \theta$$

Find  $p_{R\theta}(r, \theta)$



$$\text{Prob}((X, Y) \in \mathcal{R}) = \int_{\mathcal{R}} p_{XY} dx dy = \int_{\mathcal{R}} \frac{1}{2\pi} \exp\left(-\frac{x^2 + y^2}{2}\right) dx dy$$

$$= \int_{\mathcal{R}} \frac{1}{2\pi} \exp\left(-\frac{r^2}{2}\right) dx dy \approx \frac{1}{2\pi} \exp\left(-\frac{r_0^2}{2}\right) \int_{\mathcal{R}} dx dy$$

$$= \frac{1}{2\pi} \exp\left(-\frac{r^2}{2}\right) A(\mathcal{R}) = \text{Prob}((R, \theta) \in \mathcal{R}) = \int_{\mathcal{R}} p_{R\theta}(r, \theta) 1 r dr d\theta$$

10 Let  $X \sim N(m, \Lambda)$

$$m = \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}, \quad \Lambda = \begin{bmatrix} 2 & 1 & 1 \\ 1 & 2 & 0 \\ 1 & 0 & 2 \end{bmatrix}$$

a) Find  $Q$  :  $X = QZ$   $Z$  being independent,  $m_z = 0$   
 $\sigma_z^2 = 1$

$$Z \sim N\left(\begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \end{bmatrix}\right)$$

$$m_x = Q m_z \quad \underline{0} = Q \underline{0} \Rightarrow Q = \text{Anything } 3 \times 3$$

$$\Lambda_{zz} = I_{3 \times 3} \quad \Lambda_{xx} = Q \Lambda_{zz} Q^* = Q Q^*$$

$$\lambda(\Lambda_{xx}) = [0.5858, 2, 3.4142] \Rightarrow \Lambda_{xx} > 0$$

$$\Lambda_{xx} = U H U^* \quad \Lambda_{xx}^{\frac{1}{2}} = U H^{\frac{1}{2}} U^*$$

$$Q = U H^{\frac{1}{2}} U^* \Rightarrow \Lambda_{xx} = U H^{\frac{1}{2}} U^* (U H^{\frac{1}{2}} U^*)^*$$

$$U = \begin{bmatrix} -\frac{\sqrt{2}}{2} & 0 & \frac{1}{\sqrt{2}} \\ \frac{1}{2} & -\frac{1}{\sqrt{2}} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{\sqrt{2}} & \frac{1}{2} \end{bmatrix} \quad \Lambda_{xx} = U H^{\frac{1}{2}} H^{\frac{1}{2}} U^* = U H U^*$$

$$H^{\frac{1}{2}} = \begin{bmatrix} 0.7654 & & \\ & \sqrt{2} & \\ & & 1.4478 \end{bmatrix}$$

$$\Rightarrow Q = \begin{bmatrix} 1.3066 & 0.3827 & 0.3827 \\ 0.3827 & 1.3604 & -0.0538 \\ 0.3827 & -0.0538 & 1.3604 \end{bmatrix}$$

11 Let  $c_k \sim N(0, \Lambda_{EE})$ , and suppose  $c_k, c_j$  uncorrelated for  $k \neq j$

$$x_{k+1} = Ax_k + Bc_k \quad \text{with } x_0 = 0$$

a) Find an expression of  $x_k$  in terms of  $c_j$

$$x_1 = Ax_0 + Bc_0 = Bc_0$$

$$x_2 = Ax_1 + Bc_1 = ABc_0 + Bc_1$$

$$x_3 = Ax_2 + Bc_2 = AABc_0 + ABc_1 + Bc_2$$

$$x_k = A^{k-1}Bc_0 + A^{k-2}Bc_1 + \dots + ABc_{k-2} + Bc_{k-1}$$

$$x_k = \sum_{j=0}^{k-1} A^j B c_{k-1-j} = \sum_{i=0}^{k-1} A^{k-1-i} B c_i$$

b) Find the density function for  $x_k$

$x_k$  Gaussian - Gaussian

$$E\{x_k\} = E\left\{\sum_{i=0}^{k-1} A^{k-1-i} B c_i\right\} = \sum_{i=0}^{k-1} A^{k-1-i} B E\{c_i\} = 0$$

$$E\{x_k x_k^*\} = E\left\{\sum_{i=0}^{k-1} A^{k-1-i} B c_i \sum_{j=0}^{k-1} c_j^* B^* A^{k-1-j}\right\}$$

$$= E\left\{\sum_{i=0}^{k-1} A^{k-1-i} B c_i c_i^* B^* A^{k-1-i}\right\}$$

$$\Lambda_{xx}^k = \sum_{i=0}^{k-1} A^{k-1-i} B E\{c_i c_i^*\} B^* A^{k-1-i} = \sum_{i=0}^{k-1} A^{k-1-i} B \Lambda_{EE} B^* A^{k-1-i}$$

$$P_k = \sum_{i=0}^{k-1} A^{k-1-i} B \Lambda_{EE} B^* A^{k-1-i}$$

c) What happens as  $k \rightarrow \infty$ , Assumptions

$$P_{k+1} = \sum_{i=0}^k A^{k-i} B \Lambda_{EE} B^* A^{k-i} = \sum_{i=0}^{k-1} A A^{k-1-i} B \Lambda_{EE} B^* A^{k-1-i} A^* + B \Lambda_{EE} B^*$$

$$P_{k+1} = A P_k A^* + B \Lambda_{EE} B^*$$