Arrouncewits

- (vow). fixed).... · - HW4: Typo in Vist
 - Quiz 6 (11/22) will be optioned.
- Cheaty: Hw must be writer <u>alove</u>.

 Quittes vo tallary, looking at ôter people, etc.

Parametre Inferme

 $F = \{ f_{\theta}(x) : \theta \in \Theta \}$ paramean noch

Ex. F = \(\frac{1}{2\tau_0^2} \exp(-\left(\frac{X-\mu)^2}{2\sigma^2}\right), \left(\mu_0^2) \exp(-\left(\frac{X-\mu)^2}{2\sigma^2}\right), \left(\mu_0^2) \exp(-\left(\frac{X-\mu)^2}{2\sigma^2}\right), \left(\mu_0^2) \exp(-\left(\frac{X-\mu)^2}{2\sigma^2}\right), \left(\mu_0^2) \exp(-\left(\frac{X-\mu)^2}{2\sigma^2}\right), \left(\mu_0^2) \exp(-\left(\frac{X-\mu}{2\sigma^2}\right), \left(\mu

. Θ= (μ, δ). Θ= R×R,

God: learn & from dorten.

pet. A parameter et intest is some quantity (depend on O) that we cave about.

 $F_{\mathbf{x}} \cdot \mu = T(\theta)$

 $E_{X} = 1 - \underbrace{\overline{D}\left(\frac{1-\mu}{5}\right)}_{5}, \quad \underline{\overline{D}(x)} = \underbrace{\int_{-\infty}^{X} e^{x} \rho\left(-\frac{3}{2}^{2}\right) dz}_{\infty}$

Deil.	The:	j-th.	moineil.	of a	Rv. X.	is Œ[xi]
• •		• • •				

Det. The j-th moment of a distribut $\overline{F_0}$ is $\alpha_j = \alpha_j(\theta) = \int x^i d\overline{F_0}(x) = \mathbb{E}[z^i]$

Det. The jth sample moment of iid data

X1, ..., Xn is $\alpha_j = \frac{1}{n} \sum_{i=1}^{n} X_i^{i}$

I deas find the value of the thrut gives the same moments as wheat are observed in the data.

Det. Let $F = \{f_{\theta}(x) : \theta \in \Theta\}$, where $\Theta \in \mathbb{R}^k$.

The method of mounds estimate is

the parameter $\widehat{\theta}_n \in \mathbb{R}^k$ such that $\widehat{\theta}_n = X_1(\widehat{\theta}_n) = X_1(\widehat{\theta}_{n_1}, \dots, \widehat{\theta}_{n_r k})$ $\widehat{\chi}_1 = \chi_1(\widehat{\theta}_n) = \chi_1(\widehat{\theta}_{n_1}, \dots, \widehat{\theta}_{n_r k})$

 $\widehat{A}_{k} = \alpha_{k}(\widehat{\theta}_{n}) = \alpha_{k}(\widehat{\theta}_{n,1}, \dots, \widehat{\theta}_{n,k}).$

Mrs is k equatous and & unknowns.

$$\hat{x}_{i} = P, \quad \hat{H} = (0,1)$$

$$\hat{x}_{i} = \frac{1}{n} \hat{z}_{i} \quad \hat{x}_{i} \quad \hat{x}_{i} = 1 - p + 0 \cdot (1-p) = p$$

Solve:
$$\hat{p}_n = \frac{1}{n} \sum_{i=1}^{n} X_i$$
 for \hat{p}

Ex.
$$F = \{F_{\mu,\sigma}: \mu \in \int \chi dF(x), \sigma^{2} = \int (x-\mu)^{2} dF(x), \mu \in \mathbb{R}^{n}, \sigma^{2} > 0\}$$

$$\hat{\alpha}_1 = \frac{1}{n} \sum_{i=1}^{n} X_i$$

$$\hat{\alpha}_2 = \frac{1}{n} \sum_{i=1}^{n} X_i$$

$$\int_{\Lambda}^{\Lambda} x^{2} = \frac{1}{\Lambda} \sum_{i=1}^{N} x_{i}^{2} - \left(\frac{1}{\Lambda} \sum_{i=1}^{N} x_{i}\right)^{2} = \frac{1}{\Lambda} \sum_{i=1}^{N} \left(x_{i} - x_{\Lambda}\right)^{2}$$

Theorem

Let En he the method et voumes estimation. Under appropriée conditions:

 $2: \stackrel{P}{\theta_n} \xrightarrow{P} \emptyset \qquad ((onlish f))$

 $3. \sqrt{n} (\hat{\theta}_{n} - \theta) \rightarrow N(0, \Sigma).$

where $\Sigma = g \mathbb{E} [YY^{\dagger}] g^{\dagger}$

$$y = (x, x^2, ..., x^k)^T, g = (g_1, ..., g_k)$$

$$g_{i} = \frac{\partial \alpha_{i}(\theta)}{\partial \theta}$$