Report on panel model results

Vicente López Díaz

July 22, 2021

Abstract

These note is to relate several models that can use panel data.

The objective of these note is to give a broad overview of the possible models that can use panel data. There are several usual features to consider in a model with panel data, for example, changes on parameters for time or individual. Also, specification on error term is relevant for interpretation.

The note is based on Hsiao (2014). It goes from the theory in the text, to the application.

1 Data plots

Here is a summary of the data available for the analysis.

Figure 1 presents the initial data points, these are used for the analysis.

I made a decision on which brands to include based on the number of observations on the period previous to the tax implementation, the tax started in january 2021, I made an exploratory analysis on the december 2020 data. This would ease the inclusion of brand explicitly in the analysis, brands with few observations, with difficulties to calculate most estimates could be analyzed after applying some criteria to make groups of brands out of the individual ones.

Figure 2 only considers the 7 most frequent brands. The graph provides some guidance on what to consider for the proposed descriptive model. In particular, there is clear trend over time and there are price adjustments in january.

Except for prices of other products, there are no other potential regressors to consider at the same level of the data.

2 Dummies for each level: city, brand, time

Estimations using areg, fixed effects are imposed. Using this method there is one category with parameters "absorbed", which are not estimated as a result of the procedure.

Specification with indicators for city, brand and trend for time. Same tax effect on all the brands.

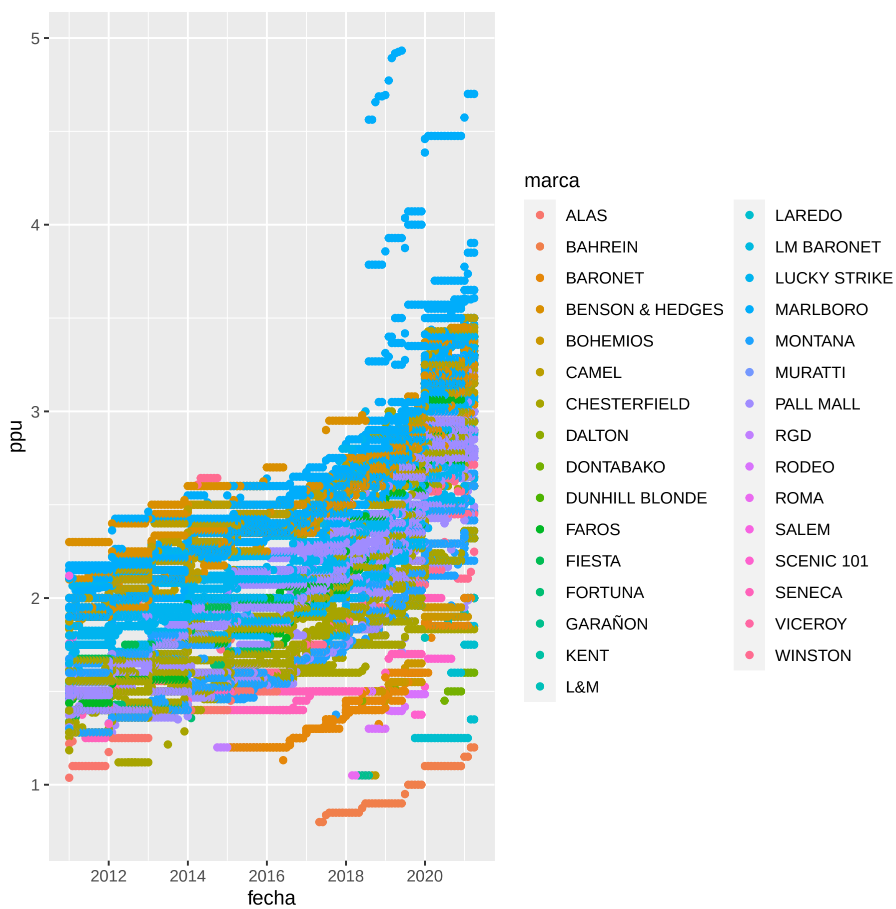

Figure 1: All brands average price per unit

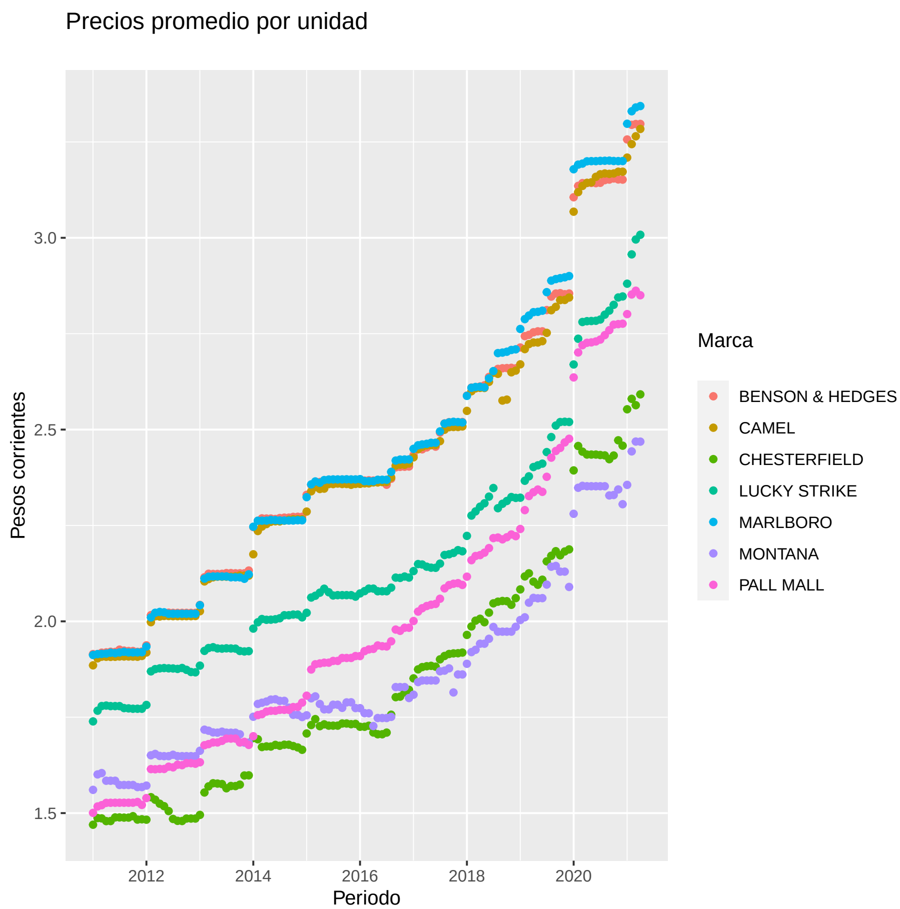

Figure 2: Seven brands average price per unit

$$y_{ctm} = \alpha_i^* + \gamma_m^* + \lambda * t + \beta_0^{'} jan + \beta_1^{'} tax2020 + \beta_1^{'} tax2021 + u_{ctm}; \tag{2.1}$$

$c = 1, \ldots, N; t = 1, \ldots, T; m = 1, \ldots, M.$

Specification with indicators for city, brand and trend for time. Effect interacted for each brand.

$$y_{ctm} = \alpha_i^* + \gamma_m^* + \lambda * t + \beta_0^{'} jan + \beta_{1m}^{'} tax2020 + \beta_{2m}^{'} tax2021 + u_{itm}; \tag{2.2}$$

$c = 1, \ldots, N; t = 1, \ldots, T; m = 1, \ldots, M.$

Results with data for the 7 brands:

Table 1: Fixed Effects for city and brand

VARIABLES	(1) ppu	(2) ppu	(3) ppu	(4) ppu
2.marca	-0.012***	-0.011***	-0.006**	-0.006**
	(0.003)	(0.003)	(0.003)	(0.003)
3.marca	-0.595***	-0.593***	-0.592***	-0.591***
	(0.004)	(0.004)	(0.003)	(0.003)
4.marca	-0.270***	-0.268***	-0.268***	-0.266***
	(0.003)	(0.003)	(0.003)	(0.003)
5.marca	0.011***	0.010***	0.011***	0.010***
	(0.003)	(0.003)	(0.002)	(0.002)
6.marca	-0.524***	-0.521***	-0.527***	-0.524***
	(0.005)	(0.005)	(0.004)	(0.004)
7.marca	-0.435***	-0.435***	-0.441***	-0.440***
	(0.003)	(0.003)	(0.002)	(0.002)
1.m1_20		0.238***		
		(0.024)		
1.m1_20#2.marca		-0.007		-0.012
		(0.042)		(0.032)
1.m1_20#3.marca		-0.111**		-0.112***
		(0.043)		(0.033)
1.m1_20#4.marca		-0.147***		-0.150***
		(0.039)		(0.030)
1.m1_20#5.marca		0.052		0.051**
		(0.032)		(0.024)
1.m1_20#6.marca		-0.239***		-0.235***
		(0.062)		(0.048)
1.m1_20#7.marca		-0.028		-0.024
		(0.034)		(0.026)
1.m1_21		0.285***		
		(0.024)		

1.m1_21#2.marca		-0.035		-0.040
		(0.040)		(0.031)
1.m1_21#3.marca		-0.116**		-0.118***
		(0.046)		(0.035)
1.m1_21#4.marca		-0.132***		-0.134***
		(0.038)		(0.029)
1.m1_21#5.marca		0.022		0.021
		(0.032)		(0.024)
1.m1_21#6.marca		-0.361***		-0.357***
		(0.063)		(0.048)
1.m1_21#7.marca		-0.036		-0.031
		(0.034)		(0.026)
m1	-0.024***	-0.024***	-0.069***	-0.069***
	(0.004)	(0.004)	(0.004)	(0.004)
ym	0.009***	0.009***		
	(0.000)	(0.000)		
m1_20	0.209***		-0.017*	0.012
	(0.011)		(0.009)	(0.018)
m1_21	0.242***		-0.011	0.032*
	(0.011)		(0.010)	(0.019)
Constant	-3.925***	-3.926***	1.991***	1.990***
	(0.018)	(0.018)	(0.004)	(0.004)
Observations	24,010	24,010	24,010	24,010
R-squared	0.897	0.897	0.940	0.940

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

The columns 1 and 3 consider the same effect for each brand, the columns 2 and 4 estimate a different effect for each brand. The columns 1 and 2 consider a trend, columns 3 and 4 use a combination of dummy variables for year and month.

MANUAL: REMOVE THE CATEGORIES ZERO IN m1-20.marca

2.1 Comparisons by segment

Results by brand type: premium, columns 1 and 2; medium, columns 3 and 4; low, columns 5 and 6. The omitted identifier corresponds to the reference brand.

Table 2: Fixed Effects by brand type

	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	ppu	ppu	ppu	ppu	ppu	ppu
2.marca	-0.008***	-0.007***				

	(0.003)	(0.003)				
5.marca	0.008***	0.007***				
	(0.002)	(0.002)				
1.m1_20		0.198***		0.119***		0.225***
		(0.019)		(0.033)		(0.040)
1.m1_20#2.marca		-0.009				
		(0.032)				
1.m1_20#5.marca		0.051**				
		(0.025)				
1.m1_21		0.235***		0.188***		0.310***
		(0.019)		(0.033)		(0.042)
1.m1_21#2.marca		-0.037				
		(0.031)				
1.m1_21#5.marca		0.022				
		(0.025)				
m1	-0.018***	-0.018***	-0.040***	-0.040***	-0.013	-0.013
	(0.004)	(0.004)	(0.007)	(0.007)	(0.009)	(0.009)
ym	0.010***	0.010***	0.009***	0.009***	0.007***	0.007***
	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)
m1_20	0.219***		0.182***		0.191***	
	(0.012)		(0.021)		(0.034)	
m1_21	0.238***		0.235***		0.231***	
	(0.012)		(0.021)		(0.036)	
7.marca			-0.165***	-0.166***		
			(0.004)	(0.004)		
0b.m1_20#7o.marca				0.000		
				(0.000)		
1.m1_20#7.marca				0.102**		
				(0.041)		
1.m1_21#7.marca				0.076*		
				(0.040)		
6.marca					0.101***	0.102***
					(0.007)	(0.007)
1.m1_20#6.marca						-0.119*
						(0.072)
1.m1_21#6.marca						-0.252***
						(0.073)
Constant	-4.429***	-4.429***	-4.040***	-4.041***	-2.892***	-2.895***
	(0.019)	(0.019)	(0.037)	(0.037)	(0.051)	(0.051)
Observations	13,396	13,396	6,700	6,700	3,914	3,914
R-squared	0.917	0.918	0.846	0.846	0.760	0.761

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

Table with results to test for difference of coeficients in brands.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	Equality of Intercept			Equality of Tax 2020			Equality of Tax 2021		
Equation	Numerator	Denominator	F	Numerator	Denominator	F	Numerator	Denominator	F
(2.1)	6	23954	10163.44						
(2.2)	6	23942	10006.08	6	23942	8.15	6	23942	9.14
(2.1)	2	13344	19.05						
(2.2)	2	13340	16.81	2	13340	2.89	2	13340	1.88
(2.1)	1	6649	1604.66						
(2.2)	1	6647	1613.81	1	6647	6.09	1	6647	3.52
(2.1)	1	3867	187						
(2.2)	1	3865	193.13	1	3865	2.74	1	3865	11.8

All tests are significant at 1 percent level

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	Equality of Intercept			Equality of Tax 2020			Equality of Tax 2021		
Equation	Numerator	Denominator	F	Numerator	Denominator	F	Numerator	Denominator	F
(2.1)	6	23954	10163.44						
(2.2)	6	23942	10006.08	6	23942	8.15	6	23942	9.14
(2.1)	2	13344	19.05						
(2.2)	2	13340	16.81	2	13340	2.89	2	13340	1.88
(2.1)	1	6649	1604.66						
(2.2)	1	6647	1613.81	1	6647	6.09	1	6647	3.52
(2.1)	1	3867	187						
(2.2)	1	3865	193.13	1	3865	2.74	1	3865	11.8

All tests are significant at 1 percent level

3 Different parameters for each brand

Uses xtsur, user-defined, command. One estimate of each parameter for each brand. The intention was to make a unique model of Seemingly Unrelated Regressions to test the coefficients of the tax change for equality. Unfortunately, it is impossible (using the xtsur routine, in a 4th gen i7 with 16ram) to make the estimation based on the complete sample, with 7 brands. I present the test based on three groups of brands.

$$y_{itm} = \alpha_i^* + \lambda_1 * t + \beta_0^{'} janDummy_m + \beta_1^{'} taxDummy_m + u_{itm};$$

$i = 1, \dots, N; t = 1, \dots, T.m = 1, 2, \dots, 7$

3.1 Comparisons by segment

Results for premium brands

VARIABLES	(1) ppu1	(2) ppu2	(3) ppu5
m1_20	0.156***	0.135***	0.151***
	(0.018)	(0.018)	(0.012)
m1_21	0.077	0.027	0.070
	(0.000)	(0.000)	(0.000)
m1	0.024***	0.000	0.003
	(0.004)	(0.004)	(0.002)
ym	0.009***	0.009***	0.009***
	(0.000)	(0.000)	(0.000)
Observations	2,543	2,543	2,543
Number of cve_ciudad	45	45	45

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

Results for lower segment brands

VARIABLES	(1) ppu3	(2) ppu6
m1_20	0.322***	0.204***
	(0.047)	(0.030)
m1_21	0.364***	0.446***
	(0.047)	(0.030)
m1	0.009	-0.035***
	(0.015)	(0.010)
ym	0.007***	0.005***
	(0.000)	(0.000)
Observations	614	614
Number of cve_ciudad	43	43

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

Results for mid-range segment brands

VARIABLES	(1) ppu4	(2) ppu7
m1_20	0.150***	0.071***
	(0.018)	(0.018)
m1_21	0.050***	0.191***
	(0.018)	(0.018)
m1	-0.013***	-0.018***
	(0.005)	(0.005)
ym	0.009***	0.010***
	(0.000)	(0.000)
Observations	2,112	2,112
Number of cve_ciudad	43	43

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

4 Constant Parameters over time

Estimations using xtreg, first static estimations, second dynamic estimates. Separate regression for each brand.

The estimation routine has the possibility to distinguish between fixed or random individual coefficients.

Separate regression for each individual defined by city and brand.

$$y_{it} = \alpha_i^* + \beta_i^{'} x_{it} + u_{it}; i = 1, \dots, N; t = 1, \dots, T.$$

Where i represents a combination of city and brand.

4.1 Static models

The proposed model only uses fixed regressors, the effect of the price change in every january, january 2020 and january 2021,

$$y_{it} = \alpha_i^* + \gamma_m^* + \lambda * t + \beta_0^{'} jan + \beta_1^{'} tax2020 + \beta_1^{'} tax2021 + u_{it}; \tag{4.1}$$

$c = 1, \ldots, N; t = 1, \ldots, T; m = 1, \ldots, M.$

It includes interactions, for the effect of the price change in january 2020 and january 2021, for different brand-types.

$$y_{it} = \alpha_i^* + \gamma_m^* + \lambda * t + \beta_0^{'} jan + \beta_{1m}^{'} tax2020 + \beta_{2m}^{'} tax2021 + u_{it}; \tag{4.2}$$

$i = 1, \ldots, N; t = 1, \ldots, T; m = 1, \ldots, M.$

Because there are many omitted variables captured in the individual effects, there is the question of the relevance of them as fixed or random.

The result of the Hausman test for fixed effects does not rule out the non systematic difference in coefficients, this is in favour of the random effects model: Chi2(4) = 0.60, $Prob \geq chi2 = 0.9628$

Similarly the Breusch-Pagan test for random effects does rule out the alternative of OLS : chibar2(01) = 7.5e+05, $Prob \geq chibar2 = 0.0000$

The test of unit root, using the Fischer type estimation from Choi (2001): Inverse chi-squared(500) = 1112.8056 , $Prob \geq chi2 = 0.0000$, does not rule out the presence of unit root for any panel (defined as a combination of city and brand), except for the model that includes a drift. The result suggests to consider different trends for each brand or city, there is an estimation by brand to test for unit roots by specifications of the panel.

Next are the results from the complete sample. First two columns and last two columns are estimated using random effects. The third column using fixed effects, since the next estimates by brand suggest that most brands have fixed effects for city.

Table 3: Fixed Effects for brand and city by brand type

VARIABLES	(1) ppu	(2) ppu	(3) ppu	(4) ppu	(5) ppu
2.marca	-0.013	-0.013		-0.009	-0.008
	(0.024)	(0.022)		(0.018)	(0.017)
3.marca	-0.605***	-0.604***		-0.603***	-0.602***
	(0.024)	(0.022)		(0.018)	(0.017)
4.marca	-0.274***	-0.272***		-0.269***	-0.268***
	(0.024)	(0.021)		(0.017)	(0.017)
5.marca	0.004	0.004		0.007	0.007
	(0.022)	(0.020)		(0.016)	(0.016)

6.marca	-0.502***	-0.500***		-0.505***	-0.503***
	(0.028)	(0.026)		(0.021)	(0.020)
7.marca	-0.426***	-0.425***		-0.443***	-0.442***
	(0.023)	(0.021)		(0.017)	(0.017)
m1_20	0.202***	0.230***	0.230***	-0.015**	0.012
	(0.010)	(0.021)	(0.021)	(0.007)	(0.015)
m1_21	0.232***	0.276***	0.275***	-0.012	0.027*
	(0.010)	(0.021)	(0.021)	(0.008)	(0.016)
1.m1_20#2.marca					-0.011
					(0.026)
1.m1_20#3.marca					-0.103***
					(0.027)
1.m1_20#4.marca					-0.129***
					(0.024)
1.m1_20#5.marca					0.052***
					(0.020)
1.m1_20#6.marca					-0.195***
					(0.039)
1.m1_20#7.marca					-0.040*
					(0.021)
1.m1_21#2.marca					-0.039
					(0.025)
1.m1_21#3.marca					-0.100***
					(0.029)
1.m1_21#4.marca					-0.113***
					(0.024)
1.m1_21#5.marca					0.022
					(0.020)
1.m1_21#6.marca					-0.360***
					(0.039)
1.m1_21#7.marca					-0.033
					(0.021)
m1	-0.023***	-0.023***	-0.023***	-0.070***	-0.070***
	(0.003)	(0.003)	(0.003)	(0.003)	(0.003)
ym	0.009***	0.009***	0.009***		
	(0.000)	(0.000)	(0.000)		
2.marca#1.m1_20		-0.012	-0.011		
		(0.037)	(0.037)		
3.marca#1.m1_20		-0.109***	-0.108***		
		(0.039)	(0.039)		
4.marca#1.m1_20		-0.125***	-0.124***		
		(0.034)	(0.034)		
5.marca#1.m1_20		0.056**	0.056**		
		(0.028)	(0.028)		
6.marca#1.m1_20		-0.194***	-0.191***		
		(0.056)	(0.056)		

7.marca#1.m1_20		-0.052*	-0.053*		
		(0.030)	(0.030)		
2.marca#1.m1_21		-0.047	-0.051		
		(0.036)	(0.036)		
3.marca#1.m1_21		-0.101**	-0.101**		
		(0.041)	(0.041)		
4.marca#1.m1_21		-0.117***	-0.118***		
		(0.034)	(0.034)		
5.marca#1.m1_21		0.027	0.027		
		(0.028)	(0.028)		
6.marca#1.m1_21		-0.364***	-0.364***		
		(0.056)	(0.056)		
7.marca#1.m1_21		-0.053*	-0.051*		
		(0.030)	(0.030)		
Constant	-3.940***	-3.940***	-4.137***	1.989***	1.989***
	(0.023)	(0.022)	(0.017)	(0.012)	(0.012)
Observations	23,926	23,926	23,926	23,926	23,926
Number of gr_marca_ciudad	263	263	263	263	263
R-squared			0.866		

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

F tests show in the specification in this stage does reject equality of coeficients.

Next are the results for brand type.

Table 4: Fixed Effects/Random Effects for brand and city by brand type

	(1)	(2)	(3)	(4)	(5)	(6)
VARIABLES	ppu	ppu	ppu	ppu	ppu	ppu
4.marca			-4.013***	0.150***		
			(0.041)	(0.028)		
1.m1_20		0.196***		0.132***		0.212***
		(0.018)		(0.030)		(0.037)
1.m1_20#7.marca				0.071*		
				(0.037)		
1.m1_21		0.233***		0.251***		0.295***
		(0.018)		(0.024)		(0.040)
1.m1_21#4.marca				-0.062*		
				(0.037)		
m1	-0.018***	-0.018***	-0.039***	-0.039***	-0.012	-0.012
	(0.004)	(0.004)	(0.006)	(0.006)	(0.008)	(0.008)
ym	0.010***	0.010***	0.009***	0.009***	0.007***	0.007***
	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)

1.m1_20#2.marca		-0.012				
		(0.031)				
1.m1_20#5.marca		0.053**				
		(0.024)				
1.m1_21#2.marca		-0.049				
		(0.030)				
1.m1_21#5.marca		0.024				
		(0.024)				
m1_20	0.218***		0.176***		0.190***	
	(0.011)		(0.019)		(0.032)	
m1_21	0.235***		0.228***		0.217***	
	(0.011)		(0.019)		(0.034)	
7.marca			-4.162***			
			(0.041)			
1.m1_20#6.marca						-0.078
						(0.068)
1.m1_21#6.marca						-0.245***
						(0.069)
Constant	-4.417***	-4.416***		-4.164***	-2.961***	-2.963***
	(0.019)	(0.019)		(0.041)	(0.050)	(0.049)
Observations	13,396	13,396	6,620	6,620	3,910	3,910
R-squared	0.916	0.916			0.720	0.721
Number of gr_marca_ciudad	126	126	80	80	57	57

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

By brand type, for premium(1) and low(3) brand types the Hausman test rejects the alternative of random effects in favour of fixed individual effects. The test for type 1 and 3 are Chi2(4) = 0.60, $Prob \geq chi2 = 0.9628$ and Chi2(4) = 0.60, $Prob \geq chi2 = 0.9628$, respectively.

Similarly the Breusch-Pagan test for random effects does rule out the alternative of OLS : chibar2(01) = 7.5e+05, $Prob \geq chibar2 = 0.0000$

The test of unit root, using the Fischer type estimation from Choi (2001): Inverse chi-squared(500) = 1112.8056 , $Prob \geq chi2 = 0.0000$, does not rule out the presence of unit root for any panel (defined as a combination of city and brand), except for the model that includes a drift. The result suggests to consider different trends for each brand or city, there is an estimation by brand to test for unit roots by specifications of the panel.

F tests or Chi2 estimated for equality of coeficients.

Testing difference in brand effects, F corresponds to models by fixed effects, chi2 corresponds to models estimated by random effects.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	Equality of Intercept			Equality of Tax 2020			Equality of Tax 2021		
Equation	Num/gl-chi2	Den	F/Chi2	Num/gl-chi2	Den	F/Chi2	Num/gl-chi2	Den	F/Chi2
(4.1)	6		1287.39***						
(4.2)	6		1547.03***	6		51.86***	6		67.31***
(4.1)	125	13266	22.32***						
(4.2)	125	13262	22.32***	2	13262	3.51**	2	13262	3.08***
(4.1)	2		11224.48***						
(4.2)	1		28.63***	1		3.62*	1		2.85*
(4.1)	56	3849	41.93***						
(4.2)	56	3847	42.17***	1	3847	1.32	1	3847	12.43*

*** p<0.01, ** p<0.05, * p<0.1

The next table presents the static results by brand. Considering the Hausman test, the model was estimated using random or fixed effects, for each brand. The column number corresponds to the brand label.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
VARIABLES	ppu	ppu	ppu	ppu	ppu	ppu	ppu
m1_20	0.193***	0.202***	0.202***	0.202***	0.202***	0.202***	0.202***
	(0.018)	(0.010)	(0.010)	(0.010)	(0.010)	(0.010)	(0.010)
m1_21	0.231***	0.232***	0.232***	0.232***	0.232***	0.232***	0.232***
	(0.018)	(0.010)	(0.010)	(0.010)	(0.010)	(0.010)	(0.010)
m1	-0.013**	-0.023***	-0.023***	-0.023***	-0.023***	-0.023***	-0.023***
	(0.006)	(0.003)	(0.003)	(0.003)	(0.003)	(0.003)	(0.003)
ym	0.010***	0.009***	0.009***	0.009***	0.009***	0.009***	0.009***
	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)
Observations	4,650	23,926	23,926	23,926	23,926	23,926	23,926
R-squared	0.921	0.865	0.865	0.865	0.865	0.865	0.865
Number of gr_marca_ciudad	44	263	263	263	263	263	263

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

4.2 Dynamic models

Since we cannot rule the presence of unit roots for each panel, by city, except for a model with drift, an alternative is to consider dynamics in the equation, in particular the lagged dependent variable.

$$y_{it} = \delta * y_{i,t-1} + \alpha_i^* + \gamma_m^* + \lambda * t + \beta_0^{'} jan + \beta_1^{'} tax2020 + \beta_1^{'} tax2021 + u_{it}; \quad (4.3)$$

$c = 1, \ldots, N; t = 1, \ldots, T; m = 1, \ldots, M.$

Similar to the previous section, estimations include interactions, to consider the effect of the price change in every january, january 2020 and january 2021.

$$y_{it} = \delta * y_{i,t-1} + \alpha_i^* + \gamma_m^* + \lambda * t + \beta_0^{'} jan + \beta_{1m}^{'} tax2020 + \beta_{2m}^{'} tax2021 + u_{it}; \quad (4.4)$$

$i = 1, \ldots, N; t = 1, \ldots, T; m = 1, \ldots, M.$

The table shows the results fixed effects, prefered by the Hausmann test. It shows also the estimates by type. F tests for the equality of the effects by brand and by segment.

VARIABLES	(1) ppu	(2) ppu	(3) ppu	(4) ppu	(5) ppu	(6) ppu
m1_20	0.182***	0.202***	0.207***	0.230***	0.202***	
	(0.003)	(0.010)	(0.006)	(0.021)	(0.010)	
m1_21	0.044***	0.232***	0.083***	0.275***	0.232***	
	(0.003)	(0.010)	(0.006)	(0.021)	(0.010)	
m1	0.022***	-0.023***	0.022***	-0.023***	-0.023***	-0.023***
	(0.001)	(0.003)	(0.001)	(0.003)	(0.003)	(0.003)
ym			0.000***	0.009***	0.009***	0.009***
			(0.000)	(0.000)	(0.000)	(0.000)
1b.marca#0b.m1_20			0.000	0.000		
			(0.000)	(0.000)		
2.marca#1.m1_20			-0.018*	-0.011		
			(0.010)	(0.037)		
3.marca#1.m1_20			-0.043***	-0.108***		
			(0.010)	(0.039)		
4.marca#1.m1_20			-0.088***	-0.124***		
			(0.009)	(0.034)		
5.marca#1.m1_20			0.031***	0.056**		
			(0.008)	(0.028)		
6.marca#1.m1_20			-0.049***	-0.191***		
			(0.015)	(0.056)		
7.marca#1.m1_20			-0.073***	-0.053*		
			(0.008)	(0.030)		
1b.marca#0b.m1_21			0.000	0.000		

			(0.000)	(0.000)		
2.marca#1.m1_21			-0.069***	-0.051		
			(0.010)	(0.036)		
3.marca#1.m1_21			-0.023**	-0.101**		
			(0.011)	(0.041)		
4.marca#1.m1_21			-0.083***	-0.118***		
			(0.009)	(0.034)		
5.marca#1.m1_21			-0.013*	0.027		
			(0.008)	(0.028)		
6.marca#1.m1_21			-0.059***	-0.364***		
			(0.015)	(0.056)		
7.marca#1.m1_21			-0.069***	-0.051*		
			(0.008)	(0.030)		
2.tipo#1.m1_20						-0.103***
						(0.021)
3.tipo#1.m1_20						-0.155***
						(0.030)
2.tipo#1.m1_21						-0.079***
						(0.021)
3.tipo#1.m1_21						-0.187***
						(0.032)
Constant			-0.171***	-4.137***		
			(0.009)	(0.017)		
Observations	23,616	23,926	23,616	23,926	23,926	23,926
R-squared	0.990	0.865	0.990	0.866	0.865	0.866
Number of gr_marca_ciudad	262	263	262	263	263	263

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

Estimates by type.

	(1)	(2)	(3)
VARIABLES	ppu	ppu	ppu
1.m1_20	0.199***	0.199***	0.199***
	(0.005)	(0.005)	(0.005)
1.m1_20#2.marca	-0.018**	-0.018**	-0.018**
	(0.009)	(0.009)	(0.009)
1.m1_20#5.marca	0.031***	0.031***	0.031***
	(0.007)	(0.007)	(0.007)
1.m1_21	0.076***	0.076***	0.076***
	(0.005)	(0.005)	(0.005)
1.m1_21#2.marca	-0.069***	-0.069***	-0.069***
	(0.008)	(0.008)	(0.008)
1.m1_21#5.marca	-0.013**	-0.013**	-0.013**
	(0.007)	(0.007)	(0.007)
m1	0.030***	0.030***	0.030***
	(0.001)	(0.001)	(0.001)
ym	0.000***	0.000***	0.000***
	(0.000)	(0.000)	(0.000)
L.ppu	0.961***	0.961***	0.961***
	(0.002)	(0.002)	(0.002)
Constant	-0.195***	-0.195***	-0.195***
	(0.012)	(0.012)	(0.012)
Observations	13,255	13,255	13,255
R-squared	0.994	0.994	0.994
Number of gr_marca_ciudad	125	125	125

Standard errors in parentheses
*** $p<0.01$, ** $p<0.05$, * $p<0.1$

MANUAL: REMOVE THE ZEROS

Interpretation: With premium for the first label, it shows that the medium brands have the smallest tax impact , although, counterintuitively the lowest impact is estimated for the medium brands with a decrease of 8.9 cents while the lower brand only decreased 5 cents, both with respect to the premium brands average.

F tests for equality of coeficients, the significant indicate the

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
	Equality of Intercept			Equality of Tax 2020			Equality of Tax 2021		
Equation	Num	Den	F	Num	Den	F	Num	Den	F
(4.3)	261	23349	1.6***						
(4.4)	261	23337	1.62***	6	23337	47.36***	6	23337	25.88***
(4.3)	124	13125	0.59						
(4.4)	124	13121	0.6	2	13121	20.34***	2	13121	35.05***
(4.3)	79	6439	0.92						
(4.4)	79	6437	0.93	1	6437	2.17	1	6437	1.9***
(4.3)	56	3775	0.94						
(4.4)	56	3773	0.95	1	3773	0.1	1	3773	2.93

*** $p<0.01$, ** $p<0.05$, * $p<0.1$

5 Consistent estimation for Variable Intercept

This models are based on Andrews, et al. (2006). The initial model comes from the transformation of:

$$y_{it} = x_{it}\beta_i + w_{j(i,t)t}\gamma + u_{it}\eta + q_{j(i,t)}\rho + \alpha_i + \phi_{j(i,t)} + \mu_t + \epsilon_{i,t};$$

$$i = 1, \ldots, N; t = 1, \ldots, T$$

Given the interest only on the fixed independent variables, we can define an heterogeneity measure on brand and city (s), take the averages at that level, and make the transformation of variables, following:

$$y_{it} - \bar{y_s} = (x_{it} - \bar{x_s})\beta_i + (w_{j(i,t)t} - \bar{w_s})\gamma + (\epsilon_{i,t} - \bar{\epsilon_s});$$

$$i = 1, \ldots, N; t = 1, \ldots, T$$

Results in the left have the same estimate for the effect. Estimates in second column correspond to the first labeled brand, Benson.

	(1)	(2)
VARIABLES	dm_ppu_cm	dm_ppu_cm
dm_m1_20_cm	0.202***	0.230***
	(0.010)	(0.021)
dm_m1_21_cm	0.232***	0.275***
	(0.010)	(0.021)
dm_m1_cm	-0.023***	-0.023***
	(0.003)	(0.003)
ym	0.009***	0.009***
	(0.000)	(0.000)
Observations	23,926	23,926
R-squared	0.865	0.866
Number of gr_marca_ciudad	263	263

Standard errors in parentheses
*** p<0.01, ** p<0.05, * p<0.1

RESULTADOS DE PRUEBAS DE DIFERENCIA DEL EFECTO POR MARCA FUERON NO SIGNIFICATIVOS.