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MATH 394 HW 4

3.79, 3.84 on p.97 of Chapter 3.

4, 11, 13, 19, 20, 21, 24, 29, 32, 33, 35 and 39 on page 163 of Chapter 4

$$3.79 \text{ Let A} = \{\text{sum} = 7\} = \{(1,6), (2,5), (3,4), (4,3), (5,2), (6,1)\}$$

Then
$$P(A) = \frac{6}{36} = \frac{1}{6}$$

Let
$$B = \{sum = even\}$$

Then $P(B) = \frac{1}{2}$, since half the time, both dice will be both even or both odd.

We want the probability of getting 2 sevens before 6 evens.

So in successive rolls, $P(A) = \frac{\frac{1}{6}}{\frac{1}{6} + \frac{1}{2}} = \frac{1}{4}$

And
$$P(B) = \frac{\frac{1}{2}}{\frac{1}{6} + \frac{1}{2}} = \frac{3}{4}$$

Call the roll a success if the sum is 7

We know that the probability of exactly k successes in m+n-1 trials is

$$\binom{m+n-1}{k} p^k (1-p)^{m+n-1-k}$$

So it follows that the desired probability of n successes before m failures is

$$P_{n,m} = \sum_{k=n}^{m+n-1} {m+n-1 \choose k} p^k (1-p)^{m+n-1-k}$$

So
$$P(2sevensbefore 6evens) = \sum_{k=2}^{2+7-1} {2+6-1 \choose k} p^k (1-p)^{2+6-1-k}$$

We know
$$\sum_{k=0}^{2+7-1} {2+6-1 \choose k} p^k (1-p)^{2+6-1-k} = 1$$

So
$$P(2sevensbefore6evens) = 1 - \sum_{k=0}^{1} {7 \choose k} p^k (1-p)^{7-k}$$

This is equal to
$$1 - \left[\binom{7}{0}p^0(1-p)^7 + \binom{7}{1}p^1(1-p)^6\right]$$

We call 7 a success, so
$$p = \frac{1}{4}$$
, $(1-p) = \frac{3}{4}$

So the desired probability is
$$1 - [\binom{7}{0} \frac{1}{4}^0 (\frac{3}{4})^7 + \binom{7}{1} \frac{1}{4}^1 (\frac{3}{4})^6]$$

Which is equal to
$$\frac{4547}{8192} \approx 0.55505$$

3.84 a) If the balls are replaced, the probability of grabbing a white ball is always $\frac{4}{12} = \frac{1}{3}$

$$P(A)$$
 winning = $\sum_{i=0}^{\infty} P(A \text{ wins on the } i^{th} \text{ draw})$

For A to win on any given i^{th} draw, A must draw a white with probability $\frac{1}{3}$, and everyone from previous hands must lose with probability $\frac{2}{3}$ each for each previous draw.

This is equal to
$$\sum_{i=0}^{\infty} \frac{1}{3} (\frac{2}{3})^{3i}$$

This is a geometric series, and converges to $\frac{a}{1-r}$

That is
$$\frac{1}{3} \frac{1}{1 - \frac{8}{27}} = \frac{9}{19}$$

P(B) winning =
$$\sum_{i=0}^{\infty} P(B \text{ wins on the } i^{th} \text{ draw})$$

For B to win on any given i^{th} draw, A must draw a non white ball with probability $\frac{2}{3}$, and B must draw a white ball with probability $\frac{2}{3}$, and all previous draws must be non white with probability $\frac{2}{3}$

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This is equal to
$$\frac{2}{3}\frac{1}{3}\frac{1}{1-\frac{8}{27}} = \frac{6}{19}$$

P(C) winning = 1 - P(A) - $P(B) = \frac{4}{19}$

b) If the balls are not replaced

 $P(A) = Sum P(A wins on i^{th} draw)$

This is equal to $\frac{4}{12}+[\frac{8}{12}\frac{7}{11}\frac{6}{10}]*\frac{4}{9}+[\frac{8}{12}\frac{7}{11}\frac{6}{10}\frac{5}{9}\frac{4}{8}\frac{3}{7}]*\frac{4}{6}=\frac{7}{15}\approx0.4666$

 $P(B) = Sum P(B wins on i^{th} draw)$

$$\frac{8}{12} \frac{4}{11} + \left[\frac{8}{12} \frac{7}{11} \frac{6}{10} \right] \frac{5}{9} \frac{4}{8} + \left[\frac{8}{12} \frac{7}{11} \frac{6}{10} \frac{5}{9} \frac{4}{8} \frac{3}{7} \right] \frac{2}{6} \frac{4}{5} = \frac{53}{165} \approx 0.3212$$

$$P(C) = 1 - P(A) - P(B) \approx 0.212$$

4.4 There are 5 men, 5 women, they are ranked, all unique, all rankings equally likely

Let X=i denote women's highest rank achieved at rank at i

P(X=1) is when the highest ranked woman is rank 1

Out of 5 women, pick 1 to be rank 1. There are 9! remaining permutations for the remaining people

So
$$P(X=1) = \frac{\binom{5}{1}9!}{10!} = \frac{1}{2}$$

P(X=2) when the highest ranked woman is rank 2

The rank 1 person is male. 5 males to choose one to be rank 1

The rank 2 is female, 5 females to choose one to be rank 2

8! permutations for the remaining rankings

$$P(X=2) = \frac{5\binom{5}{1}8!}{10!} = \frac{5}{18}$$

Rank 1 and 2 are male, 5 choices male rank 1, 4 choices male rank 2

Rank 3 female, 5 females choose 1 for rank 3, 7! permutations for others

$$P(X=3) = \frac{5*4\binom{5}{1}7!}{10!} = \frac{5}{36}$$

Rank 1,2, and 3 are male. 5 choices for rank 1, 4 for 2, 3 for 3. 5 choices for female rank 4. 6! permutations for remaining people.

$$P(X=4) = \frac{5*4*3\binom{5}{1}6!}{10!} = \frac{5}{84}$$

Rank 1,2,3, and 4 male. 5 choices female for rank 5. 5! permutations for remaining people.

$$P(X=5) = \frac{5*4*3*2\binom{5}{1}5!}{10!} = \frac{5}{252}$$

Rank 1,2,3,4, and 5 are male. 5 choices for female rank 6. 4! permutations for remaining.

$$P(X=6) = \frac{5!\binom{5}{1}4!}{10!} = \frac{1}{252}$$

Rank 1,2,3,4,5, and 6 are male. impossible.

$$P(X=7) = P(X=8) = P(X=9) = P(X=10) = 0$$

4.11 a) For 3, $10^3 = 3 * 333 + 1$, so 333 numbers between 1 and 10^3 are divisible by 3

So
$$P(3) = \frac{333}{1000}$$

For 5, $10^3 = 5 * 200$, so 200 numbers between 1 and 1000 are divisible

$$P(5) = \frac{200}{1000}$$

For 7, $10^3 = 7 * 142 + 6$, so 142 numbers between 1 and 1000 are divisible

$$P(7) = \frac{142}{1000}$$

For 15, $10^3 = 15 * 66 + 10$, so 66 numbers

$$P(15) = \frac{66}{1000}$$

For 105, $10^3 = 105 * 9 + 55$, so 9 numbers

$$P(105) = \frac{9}{1000}$$

For k large enough, P(a divides N) converges to $\frac{1}{a}$

b) Find
$$P(\mu(N) = 0)$$
, as $k \to \infty$

$$P(\mu(N) = 0) = P(N \text{ is not divisible by } P_i^2, i \ge 1)$$

=
$$\prod_{i=1}^{\infty} P(N \text{ is not divisible by } p_i^2)$$

$$=\prod_{i=1}^{\infty} P(1 - N \text{ is divisible by } p_i^2)$$

Know that the probability of a number N being divisible by a number a from 10^k for k large converges to $\frac{1}{a}$

In this case, $a = p_i^2, i \ge 1$

$$= \prod_{i=1}^{\infty} \left(1 - \frac{1}{p_i^2}\right)$$

Know that this is equal to $\frac{6}{\pi^2}$

4.13 Let s be a successful sale

Let 1st be the first sale, and 2nd be the second sale

Let dlx be a deluxe sale, and std be a standard sale

$$P(s|1st) = 0.3$$

$$P(s|2nd) = 0.6$$

$$P(dlx|s) = 0.5$$

$$P(std|s) = 0.5$$

Let X be the value of all sales.

$$X = \{0, 500, 1000, 1500, 2000\}$$

$$P(X = 0) = P(s^c|1st) * P(s^c|2nd) = (1 - 0.3)(1 - 0.6) = 0.28$$

$$P(\mathbf{X} = 500) = P(s^c|1st) * P(s|2nd) * P(std|s) + P(s|1st)(std|s) * P(s^c|2nd) = (1-0.3)(0.6)(0.5) + (0.3)(0.5) * (1-0.6) = 0.27$$

$$P(X = 1000) =$$

$$P(s|1st)*P(std|s)*P(s|2nd)P(std|s)+P(s|1st)P(dlx|s)P(s^c|2nd)+P(s^c|1st)P(s|2nd)P(dlx|s)=\\ [0.3*0.5*0.6*0.5]+[0.3*0.5*0.4]+[0.7*0.6*0.5]=0.315$$

$$P(X = 1500) = P(s|1st) * P(dlx|s) * P(s|2nd) * P(std|s) + P(s|1st) * P(std|s) + P(s|2nd) * P(dlx|s) = [0.3*0.5*0.6*0.5] + [0.3*0.5*0.6*0.5] = 0.09$$

$$P(X = 2000) = P(s|1st) * P(dlx|s) * P(s|2nd) * P(dlx|s) = [0.3*0.5*0.6*0.5] = 0.045$$

4.19
$$P(b < 0) = 0$$

$$P(0 \le b < 1) = \frac{1}{2}$$

$$P(1 \le b < 2) = \frac{3}{5} - \frac{1}{2} = \frac{1}{10}$$

$$P(2 \le b < 3) = \frac{4}{5} - \frac{3}{5} = \frac{1}{5}$$

$$P(3 \le b < 3.5) = \frac{9}{10} - \frac{4}{5} = \frac{1}{10}$$

$$p(b \ge 3.5) = 1 - \frac{9}{10} = \frac{1}{10}$$

4.20 a)
$$P(X > 0) = P(Win first) + P(lose, win, win)$$

= $\frac{18}{38} + \frac{20}{38} \frac{18}{38}^2 \approx 0.5918$

- b) No, this is not a foolproof strategy because it is possible to win 1, or lose 1 or 3.
- c) To win 1, player can win first, or lose and win 2nd and 3rd.

To end with 0, impossible.

To lose 1, player loses first, and then plays a second and third round. There are 2 choices to pick 1 to win, the other is a loss.

To lose 2, it is impossible.

To lose 3, the player loses the first, and loses the second and third.

$$E[X] = \left[1 * \frac{18}{38} + \frac{20}{38} \frac{18}{38}^{2}\right] + \left[(-1) * \frac{20}{38} {2 \choose 1} \frac{20}{38} \frac{18}{38}\right] + \left[(-3) * \frac{20}{38}^{3}\right] \approx -0.10803$$

- 4.21 148 students
 - a) It is more likely to select a student in a bigger bus than it is to select a student from any bus with equal probability. So E[X] > E[Y]

b)
$$E[X] = 40 * \frac{40}{148} + 33 * \frac{33}{148} + 25 * \frac{25}{148} + 50 * \frac{50}{148} \approx 39.28$$

 $E[Y] = \frac{(40+33+25+50)}{4} = 37$

4.24 P(b guess 1) = p P(b guess 2) = 1 - p

Loss = 3/4 unit paid

Win = i units received

a) Determine expected if A has written down number 1

$$E[X] = p * 1 - (1 - p)\frac{3}{4}$$

b) Determine expected if A has written down number 2

$$E[X] = p * -\frac{3}{4} + (1-p) * 2$$

Set the two equations equal to each other to find best p

$$p * 1 - (1 - p)\frac{3}{4} = p * -\frac{3}{4} + (1 - p) * 2$$

Solving for p, we get $p = \frac{11}{18}$

So maximum at $p = \frac{11}{18}$, so max = $\frac{23}{72}$

c) P(a writes 1) = q

P(a writes 2) = 1-q

$$E[X] = q * 1 - (1 - q) * \frac{3}{4}$$

d)
$$E[X] = -\frac{3}{4} * q + (1 - q) * 2$$

Setting the equations equal and solving for q,

We get $q = \frac{11}{18}$. At this value, A's maximum loss is equal to B's maximum gain.

4.29 There are two possibilities: Machine 1 is broken, with probability p, or Machine 2 is broken with probability 1-p

If we check machine 1 first, cost is C_1 , and there is a (1-p) probability that machine 2 is broken instead, and we need to check that for C_2 and pay R_2 for repairs. Otherwise there is a p probability that 1 is the problem so we pay R_1 to fix it

Then the expected value for cost is $C_1 + (1-p)(C_2 + R_2) + (p)(R_1)$

If we check machine 2 first, cost is C_2 , and there is a p probability that machine 2 is broken, and we need to check that for C_1 and pay R_1 for repair. Otherwise there is a (1-p) probability that 2 is the problem, so we pay R_2 to fix it

Then the expected value for cost is $C_2 + (p)(C_1 + R_1) + (1-p)(R_2)$

We would want to it in ascending order if the expected value for cost of checking 1 first is less than the expected value of cost for checking 2 first

So we check when
$$C_1 + (1-p)(C_2 + R_2) + (p)(R_1) \le C_2 + (p)(C_1 + R_1) + (1-p)(R_2)$$

This is equivalent to $C_1 \leq p(C_1 + C_2)$

So when the inequality $C_1 \leq p(C_1 + C_2)$ is not satisfied, it is better to reverse the checking order.

4.32 100 people, 10 groups of 10 people

0.1 probability for disease, each person, independent

Let p = probability disease = 0.1

If someone in the group of 10 is positive, with a probability $(p)^{10}$, then 11 tests must be conducted

If no one in the group of 10 is positive, with a probability $(1-p)^{10}$, then only 1 test is performed

$$E[X] = 1[(1 - 0.1)^{10}] + 11[(0.1)^{10}] \approx 0.3486$$

4.33 Purchase for 10

Sell for 15

Demand is a binomial variable, n = 10

Probability of purchase $p = \frac{1}{3}$

Binomial, so we know $E[X] = np = \frac{10}{3} \approx 3.333$

So he can expect to sell 3 papers (can't buy and sell fractions of paper), so he should purchase 3 papers to maximize his profit.

4.35 5 red marbles, 5 blue marbles

Draw 2 randomly

If they are the same color, win 1.10

If different colors, lose 1

a) Find the expected value of win

Two events can happen, either they are both the same color, or not the same

Let p = both same color

Then
$$E[X] = p(1.1) + (1-p)(-1) = 2.1p - 1$$

There are 5 red marbles, and 5 blue marbles. Then the probability of picking two different colors = $1 - p = \frac{\frac{5}{1} \frac{5}{1}}{\binom{10}{2}} = \frac{5}{9} \approx 0.555$

So
$$p = \frac{4}{9} \approx 0.4444$$

So
$$E[X] = 2.1(\frac{4}{9}) - 1 = \frac{-1}{15} \approx -0.0666$$

b) Find the variance of the amount you win

$$Var(X) = E[X^2] - (E[X])^2$$

Find
$$E[X^2]$$

$$E[X^2] = (1.1)^2 \frac{4}{9} + (-1)^2 \frac{5}{9} = \frac{82}{75} \approx 1.09333$$

So
$$Var(X) = \frac{82}{75} - \left[\frac{-1}{15}\right]^2 = \frac{49}{45} \approx 1.0888$$

4.39 Urn contains 3 white balls, 3 black balls

Draw a ball, with replacement.

What is the probability that of the first 4 balls drawn, exactly 2 are white?

We know
$$P(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$$

Probability of drawing a black or white ball is 0.5 each

So
$$n = 4, k = 2, p = 0.5$$

So
$$P(X=2) = {4 \choose 2} 0.5^2 (1-0.5)^{4-2} = \frac{3}{8} \approx 0.375$$