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MATH 395 HW 4

Ch 6: 21, 23, 27, 39, 41(a) only, 48

- 21 Let $f_{X,Y}(x, y) = 24xy$, $0 \leq x \leq 1, 0 \leq y \leq 1, 0 \leq x + y \leq 1$

And let it equal 0 otherwise.

- a) Show that $f(x, y)$ is a joint probability density function

We know if it is a joint pdf if the double integral is equal to 1

If we integrate with respect to x first, x goes from 0 to $1 - y$

Then, if we integrate with respect to y , y goes from 0 to 1

This is

$$\int_0^1 \int_0^y 24xy dx dy$$

$$\int_0^1 [12(1 - y)^2] y dy$$

$$\int_0^1 12(y - 2y^2 + y^3) dy$$

$$12[\frac{1}{2}y^2 - \frac{2}{3}y^3 + \frac{1}{4}y^4]_0^1$$

$$12[\frac{1}{2} - \frac{2}{3} + \frac{1}{4}] = 1$$

So it is a joint pdf

- b) To find $E[X]$, we must find $\int_{-\infty}^{\infty} x f_X(x) dx$

This is $\int_0^1 \int_0^{1-x} 24x^2 y dy dx$

$$\int_0^1 12x^2(1 - x)^2 dx$$

$$\int_0^1 12x^2(1 - 2x + x^2) dx$$

$$\int_0^1 12x^2 - 24x^3 + 12x^4 dx$$

$$4x^3 - 6x^4 + \frac{12}{5}x^5 \Big|_0^1$$

$$4 - 6 + \frac{12}{5} = \frac{2}{5}$$

- c) Find $E[Y]$. We can tell that $E[Y]$ will be calculated the same way as $E[X]$, and so it will produce the same expected value

$$E[Y] = \frac{2}{5}$$

- 23 The random variables X, Y have joint density function

$$f_{X,Y}(x, y) = 12xy(1 - x), \quad 0 < x < 1, 0 < y < 1$$

and equal to 0 otherwise

- a) Are X, Y independent?

$$f_X(x) = \int_0^1 f_{X,Y}(x, y) dy$$

$$f_X(x) = 6x(1 - x)y^2 \Big|_0^1$$

$$f_X(x) = 6x(1 - x), \text{ for } 0 < x < 1$$

$$f_Y(y) = \int_0^1 f_{X,Y}(x, y) dx$$

$$f_Y(y) = 12y[\frac{1}{2}x^2 - \frac{1}{3}x^3]_0^1$$

$$f_Y(y) = 12y[\frac{1}{2} - \frac{1}{3}]$$

$$f_Y(y) = 2y \text{ for } 0 < y < 1$$

They are independent if $f_Y(y) * f_X(x) = f_{X,Y}(x, y)$

$$6x(1-x) * 2y = 12xy(1-x), \text{ for } 0 < x < 1, 0 < y < 1$$

So they are independent

b) Find $E[X]$

$$E[X] = \int_{-\infty}^{\infty} x f_X(x) dx$$

$$\int_0^1 x[6x(1-x)] dx$$

$$\int_0^1 6[x^2 - x^3] dx$$

$$6[\frac{1}{3}x^3 - \frac{1}{4}x^4]_0^1$$

$$6[\frac{1}{3} - \frac{1}{4}] = \frac{1}{2}$$

c) Find $E[Y]$

$$E[Y] = \int_{\mathbb{R}} y f_Y(y) dy$$

$$\int_0^1 2y^2 dy$$

$$\frac{2}{3}y^3|_0^1 = \frac{2}{3}$$

d) Find $\text{Var}(X)$

$$\text{Var}(X) = E[X^2] - (E[X])^2$$

Find $E[X^2]$

$$E[X^2] = \int_{-\infty}^{\infty} x^2 f_X(x) dx$$

$$\int_0^1 x^2[6x(1-x)] dx$$

$$\int_0^1 6[x^3 - x^4] dx$$

$$6[\frac{1}{4}x^4 - \frac{1}{5}x^5]_0^1$$

$$6[\frac{1}{4} - \frac{1}{5}] = \frac{6}{20}$$

$$\text{Then } \text{Var}(X) = \frac{6}{20} - [\frac{1}{2}]^2 = \frac{1}{20}$$

e) Find $\text{Var}(Y)$

$$\text{Var}(Y) = E[Y^2] - (E[Y])^2$$

Find $E[Y^2]$

$$E[Y^2] = \int_{\mathbb{R}} y^2 f_Y(y) dy$$

$$\int_0^1 2y^3 dy$$

$$= \frac{1}{2}$$

$$\text{Var}(Y) = \frac{1}{2} - [\frac{2}{3}]^2 = \frac{1}{18}$$

- 27 If X_1, X_2 are independent exponential random variables with respective parameters λ_1, λ_2 , find the distribution of $Z = \frac{X_1}{X_2}$. Also compute $P(X_1 < X_2)$

Find the distribution means find the cdf

a) Want $P(Z < a)$

$$P(\frac{X_1}{X_2} < a)$$

$$= P(X_1 < aX_2)$$

Exponential, so x_1 goes from 0 to aX_2

Exponential, so x_2 goes from 0 to ∞

$$F_Z(a) = \int_0^\infty \left[\int_0^{a(x_2)} \lambda_1 e^{-\lambda_1 x_1} dx_1 \right] \lambda_2 e^{-\lambda_2 x_2} dx_2$$

$$\int_0^\infty [-e^{-\lambda_1 x_1}]_0^{a x_2} \lambda_2 e^{-\lambda_2 x_2} dx_2$$

$$\int_0^\infty [1 - e^{-\lambda_1 a x_2}] \lambda_2 e^{-\lambda_2 x_2} dx_2$$

$$\lambda_2 \int_0^\infty e^{-\lambda_2 x_2} - e^{-[\lambda_1 a + \lambda_2] x_2} dx_2$$

$$-e^{-\lambda_2 x_2} \Big|_0^\infty + \frac{\lambda_2}{\lambda_1 a + \lambda_2} e^{-[\lambda_1 a + \lambda_2] x_2} \Big|_0^\infty$$

$$[0 + 1] + \frac{\lambda_2}{\lambda_1 a + \lambda_2} [0 - 1]$$

$$1 - \frac{\lambda_2}{\lambda_2 + \lambda_1 a}$$

b) Find $P(X_1 < X_2)$

This is the same as part a, but we set $a = 1$

This is just

$$1 - \frac{\lambda_2}{\lambda_2 + \lambda_1}$$

- 39 Two dice are rolled. Let X, Y denote respectively, the largest and smallest values obtained. Compute the conditional mass function of Y given $X = i$, for $i = 1, \dots, 6$. Are X, Y independent? Why?

Conditional mass function

$$P_{Y|X}(y|x) = P(Y = j|X = i)$$

There are two possibilities, either $i = j$, when the rolls are equal, or $j < i$, since X is the largest roll.

If $j = i$

$$P_{Y|X}(y|x) = \frac{P(Y=i|X=i)}{P(X=i)}$$

$$= \frac{P(Y=i, X=i)}{P(X=i)}$$

The larger roll is determined already, with probability $\frac{1}{6}$

The smaller roll must match the larger roll, with probability $\frac{1}{6}$

So probability of roll 1 matching roll 2 is $\frac{1}{36}$

$$= \frac{1}{36} \frac{1}{P(X=i)}$$

If $j < i$

Then either the first roll is the smaller roll, with value j , or the second roll is smaller, with value j

the probability of rolling specifically i or j is $\frac{1}{36}$, with equal probability of roll 1 or 2 having the smaller roll, j

$$\text{Then } P(Y = j, X = i) = \frac{2}{36} \frac{1}{P(X=i)}$$

Need to find $P(X = i)$

Each specific outcome for roll 1 and 2 have probability $\frac{1}{36}$

In the case where one value is strictly less than the other, we have $j < i$

If roll 1 has the larger value, i, then the possible values of the other roll may range from 1 to $i - 1$

And the larger value i may occur on roll 1 or 2 with equal probability, so probability of $j < i$ on any roll is $\frac{1}{36}$

Then the total probability of $X = i$ for $i < j$ on either roll is $\sum_1^{i-1} \frac{2}{36} = \frac{2i-2}{36}$

And in the case that the two rolls are equal, the probability is $\frac{1}{36}$

Then the probability that $X = i$ is equal to the sum of the probabilities that $i = j, j < i$

Which is $\frac{2i-2}{36} + \frac{1}{36}$

Not independent, because $Y \leq X$, so it cannot be independent.

41a The joint density function of X, Y is given by $f_{X,Y}(x,y)xe^{-x(y+1)} \quad x > 0, y > 0$

a) Find the conditional density of X , given $Y = y$, and of Y , given $X = x$

Want $f_{X|Y}(x|y)$

This is equal to $\frac{f_{X,Y}(x,y)}{f_Y(y)}$

Find $f_Y(y)$

$$f_Y(y) = \int_0^\infty xe^{-x(y+1)} dx$$

$$\int_0^\infty xe^{-x(y+1)} dx$$

$$[x * \frac{-1}{y+1} e^{-x(y+1)}]_0^\infty - \int_0^\infty \frac{-1}{y+1} e^{-x(y+1)}$$

$$[0 - 0] - [\frac{1}{(y+1)^2} e^{-x(y+1)}]_0^\infty$$

$$-[0 - \frac{1}{(y+1)^2}]$$

$$= \frac{1}{(y+1)^2}$$

$$f_{X|Y}(x|y) = \frac{xe^{-x(y+1)}}{\frac{1}{(y+1)^2}}$$

$$= (y+1)^2 xe^{-x(y+1)} \text{ for } x > 0$$

Want $f_{Y|X}(y|x)$

This is equal to $\frac{f_{X,Y}(x,y)}{f_X(x)}$

Find $f_X(x)$

$$f_X(x) = \int_0^\infty xe^{-x(y+1)} dy$$

$$xe^{-x} \int_0^\infty e^{-xy} dy$$

$$xe^{-x} \frac{-1}{x} e^{-xy} \Big|_0^\infty$$

$$xe^{-x} [0 + \frac{1}{x}]$$

$$f_X(x) = e^{-x}$$

$$f_{Y|X}(y|x) = \frac{xe^{-x(y+1)}}{e^{-x}}$$

$$= xe^{-xy} \text{ for } y > 0$$

48 If X_1, X_2, X_3, X_4, X_5 are independent and identically distributed exponential random variables with the parameter λ , compute

Hint for 48a, $P(\min < a) = 1 - P(\min > a)$ and note the minimum is $> a$ iff each of the five X is $> a$

a) $P(\min(X_1, \dots, X_5) \leq a)$

This is equal to $1 - P(\min(X_1, \dots, X_5) > a)$

There are 5 X_i 's that need to be all be greater than a

$$1 - \left[\int_a^\infty \lambda e^{-\lambda x} \right]^5$$

$$1 - [e^{-a\lambda}]^5$$

b) $P(\max(X_1, \dots, X_5) \leq a)$

Exponential has cdf $1 - e^{-\lambda a}$

Each X_i must be greater than a , there are 5 of them

$$P(\max(X_i) \leq a) = [1 - e^{-\lambda a}]^5$$