# Quantile Processes for Semi and Nonparametric Regression

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April 7, 2016

Abstract A collection of quantile curves provides a complete picture of conditional distributions. Properly centered and scaled versions of estimated curves at various quantile levels give rise to the so-called quantile regression process (QRP). In this paper, we establish weak convergence of QRP in a general series approximation framework, which includes linear models with increasing dimension, nonparametric models and partial linear models. An interesting consequence is obtained in the last class of models, where parametric and non-parametric estimators are shown to be asymptotically independent. Applications of our general process convergence results include the construction of non-crossing quantile curves and the estimation of conditional distribution functions. As a result of independent interest, we obtain a series of Bahadur representations with exponential bounds for tail probabilities of all remainder terms.

**Keywords:** Bahadur representation, quantile regression process, semi/nonparametric model, series estimation.

# 1 Introduction

Quantile regression is widely applied in various scientific fields such as Economics (Koenker and Hallock (2001)), Biology (Briollais and Durrieu (2014)) and Ecology (Cade and Noon (2003)). By focusing on a collection of conditional quantiles instead of a single conditional mean, quantile regression allows to describe the impact of predictors on the entire conditional distribution of the response. A properly scaled and centered version of these estimated curves form an underlying (conditional) quantile regression process (see Section 2 for a formal definition). Existing literature

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on QRP is either concerned with models of *fixed* dimension (Koenker and Xiao, 2002; Angrist et al., 2006), or with a linearly interpolated version based on kernel smoothing (Qu and Yoon, 2015).

In this paper, we study weak convergence of QRP in models of the following (approximate) form

$$Q(x;\tau) \approx \mathbf{Z}(x)^{\mathsf{T}} \gamma_n(\tau),$$
 (1.1)

where  $Q(x;\tau)$  denotes the  $\tau$ -th quantile of the distribution of Y conditional on  $X=x\in\mathbb{R}^d$  and  $\mathbf{Z}(x)\in\mathbb{R}^m$  is a transformation vector of x. As noted by Belloni et al. (2011), the above framework incorporates a variety of estimation procedures such as parametric (Koenker and Bassett, 1978), non-parametric (He and Shi, 1994) and semi-parametric (He and Shi, 1996) ones. For example,  $\mathbf{Z}(x)=x$  corresponds to a linear model (with potentially high dimension), while  $\mathbf{Z}(x)$  can be chosen as powers, trigonometrics or local polynomials in the non-parametric basis expansion (where m diverges at a proper rate). Partially linear and additive models are also covered by (1.1). Therefore, our weak convergence results are developed in a broader context than those available in the literature.

A noteworthy result in the present paper is obtained for partially linear models

$$Q(X;\tau) = V^{\top} \alpha(\tau) + h(W;\tau),$$

where  $X = (V^{\top}, W)^{\top} \in \mathbb{R}^{k+k'}$ ,  $\alpha(\tau)$  is an unknown Euclidean vector and  $h(W; \tau)$  is an unknown smooth function. Here, k and k' are both fixed. In the spirit of (1.1), we can estimate  $(\alpha(\tau), h(\cdot; \tau))$  based on the following series approximation

$$h(W;\tau) \approx \widetilde{\mathbf{Z}}(W)^{\top} \boldsymbol{\beta}_n^{\dagger}(\tau).$$

Our general theorem shows the weak convergence of the *joint* quantile process resulting from  $(\widehat{\alpha}(\tau), \widehat{h}(\cdot;\tau))$  in  $(\ell^{\infty}(\mathcal{T}))^{k+1}$ , where  $\ell^{\infty}(\mathcal{T})$  denotes the class of uniformly bounded real functions on  $\mathcal{T} \ni \tau$ . An interesting consequence is that  $\widehat{\alpha}(\tau)$  and  $\widehat{h}(w_0;\tau)$  (after proper centering and scaling) jointly converge to two *independent* Gaussian processes. This asymptotic independence result is useful for simultaneously testing parametric and nonparametric components in semi-nonparametric models<sup>i</sup>. To the best of our knowledge, this is the first time that such a joint asymptotic result for quantile regression is established – in fact, even the point-wise result is new. Therefore, we prove that the "joint asymptotics phenomenon" discovered by Cheng and Shang (2015) even holds for non-smooth loss functions with multivariate nonparametric covariates.

Weak convergence of QRP is very useful in developing statistical inference procedures such as hypothesis testing on conditional distributions (Bassett and Koenker, 1982), detection of treatment effect on the conditional distribution after an intervention (Koenker and Xiao, 2002; Qu and Yoon, 2015) and testing conditional stochastic dominance (Delgado and Escanciano, 2013). For additional examples, the interested reader is referred to Koenker (2005). Our paper focuses on the estimation of conditional distribution functions and construction of non-crossing quantile curves by means of monotone rearrangement (Dette and Volgushev, 2008; Chernozhukov et al., 2010).

<sup>&</sup>lt;sup>i</sup>In the literature, a statistical model is called semi-nonparametric if it contains both finite-dimensional and infinite-dimensional unknown parameters of interest; see Cheng and Shang (2015).

Our derivation of quantile process convergence relies upon a series of Bahadur representations which are provided in Section 5. Specifically, we obtain weak convergence results by examining the asymptotic tightness of the leading terms in these representations, as shown in Lemma A.3. This is achieved by combining with a new maximal inequality in Kley et al. (2015). These new Bahadur representations are more flexible than those available in the literature (see, for instance, Belloni et al. (2011)) in terms of choosing approximation model centers. This is crucial in deriving the joint process convergence in partial linear models. As a result of independent interest, we obtain bounds with exponential tail probability for the remainder terms in our Bahadur representations. Bounds of this kind are especially useful in analyzing statistical inference procedures under divide-and-conquer setup; see Zhao et al. (2016); Shang and Cheng (2015); Volgushev et al. (2016).

The rest of this paper is organized as follows. Section 2 presents the weak convergence of QRP under general series approximation framework. Section 3 discusses the QRP in quantile partial linear models. As an application of our weak convergence theory, Section 4 considers various functionals of the quantile regression process. A detailed discussion on our novel Bahadur representations is given in Section 5, and all proofs are deferred to the appendix.

**Notation.** Denote  $\{(X_i, Y_i)\}_{i=1}^n$  i.i.d. samples in  $\mathcal{X} \times \mathbb{R}$  where  $\mathcal{X} \subset \mathbb{R}^d$ . Here, the distribution of  $(X_i, Y_i)$  and the dimension d can depend on n, i.e. triangular arrays. For brevity, let  $\mathbf{Z} = \mathbf{Z}(X)$  and  $\mathbf{Z}_i = \mathbf{Z}(X_i)$ . Define the empirical measure of  $(Y_i, \mathbf{Z}_i)$  by  $\mathbb{P}_n$ , and the true underlying measure by P with the corresponding expectation as  $\mathbb{E}$ . Note that the measure P depends on n for triangular array cases, but this dependence is omitted in the notation. Denote by  $\|\mathbf{b}\|$  the  $L^2$ -norm of a vector  $\mathbf{b}$ .  $\lambda_{\min}(A)$  and  $\lambda_{\max}(A)$  are the smallest and largest eigenvalue of a matrix A.  $\mathbf{0}_k$  denotes a k-dimensional 0 vector, and  $I_k$  be the k-dimensional identity matrix for  $k \in \mathbb{N}$ . Define

$$\rho_{\tau}(u) = (\mathbf{1}(u \le 0) - \tau)u,$$

where  $\mathbf{1}(\cdot)$  is the indicator function.  $\mathcal{C}^{\eta}(\mathcal{X})$  denotes the class of  $\eta$ -continuously differentiable functions on a set  $\mathcal{X}$ .  $\mathcal{C}(0,1)$  denotes the class of continuous functions defined on (0,1). Define

$$\psi(Y_i, \mathbf{Z}_i; \mathbf{b}, \tau) := \mathbf{Z}_i(\mathbf{1}\{Y_i \leq \mathbf{Z}_i^{\top} \mathbf{b}\} - \tau), \quad \mu(\mathbf{b}, \tau) := \mathbb{E}\big[\psi(Y_i, \mathbf{Z}_i; \mathbf{b}, \tau)\big] = \mathbb{E}\big[\mathbf{Z}_i\big\{F_{Y|X}(\mathbf{Z}_i^{\top} \mathbf{b}|X) - \tau\big\}\big],$$

and for a vector  $\gamma_n(\tau) \in \mathbb{R}^m$ , we define the following quantities

$$g_n := g_n(\boldsymbol{\gamma}_n) := \sup_{\tau \in \mathcal{T}} \|\mu(\boldsymbol{\gamma}_n(\tau), \tau)\| = \sup_{\tau \in \mathcal{T}} \|\mathbb{E}\left[\mathbf{Z}_i \left\{ F_{Y|X}(\mathbf{Z}_i^\top \boldsymbol{\gamma}_n(\tau)|X) - \tau \right\} \right] \|$$
(1.2)

Let  $\mathcal{S}^{m-1} := \{\mathbf{u} \in \mathbb{R}^m : ||\mathbf{u}|| = 1\}$  denote the unit sphere in  $\mathbb{R}^m$ . For a set  $\mathcal{I} \subset \{1, ..., m\}$ , define

$$\mathbb{R}^m_{\mathcal{I}} := \{ \mathbf{u} = (u_1, ..., u_m)^\top \in \mathbb{R}^m : u_j \neq 0 \text{ if and only if } j \in \mathcal{I} \}$$
$$\mathcal{S}^{m-1}_{\mathcal{I}} := \{ \mathbf{u} = (u_1, ..., u_m)^\top \in \mathcal{S}^{m-1} : u_j \neq 0 \text{ if and only if } j \in \mathcal{I} \}$$

Finally, consider the class of functions

$$\Lambda_{c}^{\eta}(\mathcal{X}, \mathcal{T}) := \left\{ f_{\tau} \in \mathcal{C}^{\lfloor \eta \rfloor}(\mathcal{X}) : \tau \in \mathcal{T}, \sup_{|\boldsymbol{j}| \leq \lfloor \eta \rfloor} \sup_{x, \tau \in \mathcal{T}} |D^{\boldsymbol{j}} f_{\tau}(x)| \leq c, \sup_{|\boldsymbol{j}| = \lfloor \eta \rfloor} \sup_{x \neq y, \tau \in \mathcal{T}} \frac{|D^{\boldsymbol{j}} f_{\tau}(x) - D^{\boldsymbol{j}} f_{\tau}(y)|}{\|x - y\|^{\eta - \lfloor \eta \rfloor}} \leq c \right\},$$
(1.3)

where  $\lfloor \eta \rfloor$  denotes the integer part of a real number  $\eta$ , and  $|\boldsymbol{j}| = j_1 + ... + j_d$  for d-tuple  $\boldsymbol{j} = (j_1, ..., j_d)$ . For simplicity, we sometimes write  $\sup_{\tau} (\inf_{\tau})$  and  $\sup_{x \in \mathcal{X}} (\inf_x)$  instead of  $\sup_{\tau \in \mathcal{T}} (\inf_{\tau \in \mathcal{T}})$  and  $\sup_{x \in \mathcal{X}} (\inf_{x \in \mathcal{X}})$  throughout the paper.

# 2 Weak Convergence Results

In this section, we first present our weak convergence results of QRP in a general series approximation framework that covers linear models with increasing dimension, nonparametric models and partial linear models. Furthermore, we demonstrate that the use of polynomial splines with local support, such as B-splines, significantly weakens the sufficient conditions required in the above general framework.

#### 2.1 General Series Estimator

Consider a general series estimator  $\widehat{Q}(x;\tau) := \widehat{\gamma}(\tau)^{\top} \mathbf{Z}(x)$ , where for each fixed  $\tau$ 

$$\widehat{\gamma}(\tau) := \underset{\boldsymbol{\gamma} \in \mathbb{R}^m}{\operatorname{argmin}} \sum_{i=1}^n \rho_{\tau}(Y_i - \boldsymbol{\gamma}^{\top} \mathbf{Z}_i), \tag{2.1}$$

and m is allowed to grow as  $n \to \infty$ , and assume the following conditions:

- (A1) Assume that  $\|\mathbf{Z}_i\| \leq \xi_m = O(n^b)$  almost surely with b > 0, and that  $1/M \leq \lambda_{\min}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^\top]) \leq \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^\top]) \leq M$  holds uniformly in n for some fixed constant M > 0.
- (A2) The conditional distribution  $F_{Y|X}(y|x)$  is twice differentiable w.r.t. y. Denote the corresponding derivatives by  $f_{Y|X}(y|x)$  and  $f'_{Y|X}(y|x)$ . Assume that  $\bar{f} := \sup_{y,x} |f_{Y|X}(y|x)| < \infty$  and  $\bar{f'} := \sup_{y,x} |f'_{Y|X}(y|x)| < \infty$  uniformly in n.
- (A3) Assume that uniformly in n, there exists a constant  $f_{\min} > 0$  such that

$$\inf_{\tau \in \mathcal{T}} \inf_{x} f_{Y|X}(Q(x;\tau)|x) \ge f_{\min}.$$

In the above assumptions, uniformity in n is necessary as we consider triangular arrays. Assumptions (A2) and (A3) are fairly standard in the quantile regression literature. Hence, we only make a few comments on Assumption (A1). In linear models where  $\mathbf{Z}(X) = X$  and m = d, it holds that  $\xi_m \lesssim \sqrt{m}$  if each component of X is bounded almost surely. If B-splines  $\widetilde{\mathbf{B}}(x)$  defined in Section 4.3 of Schumaker (1981) are adopted, then one needs to use its re-scaled version  $\mathbf{B}(x) = m^{1/2}\widetilde{\mathbf{B}}(x)$  as  $\mathbf{Z}(x)$  such that (A1) holds (cf. Lemma 6.2 of Zhou et al. (1998)). In this case, we have  $\xi_m \asymp \sqrt{m}$ . In addition, Assumptions (A1) and (A3) imply that for any sequence of  $\mathbb{R}^m$ -valued (non-random) functions  $\gamma_n(\tau)$  satisfying  $\sup_{\tau \in \mathcal{T}} \sup_x |\gamma_n(\tau)^{\top} \mathbf{Z}(x) - Q(x;\tau)| = o(1)$ , the smallest eigenvalues of the matrices

$$\widetilde{J}_m(\tau) := \mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top} f_{Y|X}(\boldsymbol{\gamma}_n(\tau)^{\top}\mathbf{Z}|X)], \quad J_m(\tau) := \mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top} f_{Y|X}(Q(X;\tau)|X)]$$

are bounded away from zero uniformly in  $\tau$  for all n. Define for any  $\mathbf{u} \in \mathbb{R}^m$ ,

$$\chi_{\gamma_n}(\mathbf{u}, \mathbf{Z}) := \sup_{\tau \in \mathcal{T}} \left| \mathbf{u}^\top J_m(\tau)^{-1} \mathbb{E} \left[ \mathbf{Z}_i \left( \mathbf{1} \{ Y_i \le Q(X_i; \tau) \} - \mathbf{1} \{ Y_i \le \mathbf{Z}_i^\top \gamma_n(\tau) \} \right) \right] \right|.$$

We are now ready to state our weak convergence result for QRP based on the general series estimators.

**Theorem 2.1.** Suppose (A1)-(A3) hold and  $m^3\xi_m^2(\log n)^3 = o(n)$ . Let  $\gamma_n(\cdot): \mathcal{T} \to \mathbb{R}^m$  be a sequence of functions. Let  $g_n := g_n(\gamma_n(\tau)) = o(n^{-1/2})$  (see (1.2)) and  $c_n = c_n(\gamma_n) := \sup_{x,\tau \in \mathcal{T}} |Q(x;\tau) - \mathbf{Z}(x)^\top \gamma_n(\tau)|$  and assume that  $mc_n \log n = o(1)$ . Then for any  $\mathbf{u}_n \in \mathbb{R}^m$  satisfying  $\chi_{\gamma_n}(\mathbf{u}_n, \mathbf{Z}) = o(\|\mathbf{u}_n\|n^{-1/2})$  and  $\widehat{\gamma}(\tau)$  defined in (2.1),

$$\mathbf{u}_n^{\top}(\widehat{\boldsymbol{\gamma}}(\tau) - \boldsymbol{\gamma}_n(\tau)) = -\frac{1}{n}\mathbf{u}_n^{\top}J_m(\tau)^{-1}\sum_{i=1}^n \mathbf{Z}_i(\mathbf{1}\{Y_i \le Q(X_i; \tau)\} - \tau) + o_P\left(\frac{\|\mathbf{u}_n\|}{\sqrt{n}}\right)$$
(2.2)

where the remainder term is uniform in  $\tau \in \mathcal{T}$ . In addition, if the following limit

$$H(\tau_1, \tau_2; \mathbf{u}_n) := \lim_{n \to \infty} \|\mathbf{u}_n\|^{-2} \mathbf{u}_n^{\mathsf{T}} J_m^{-1}(\tau_1) \mathbb{E}[\mathbf{Z}\mathbf{Z}^{\mathsf{T}}] J_m^{-1}(\tau_2) \mathbf{u}_n(\tau_1 \wedge \tau_2 - \tau_1 \tau_2)$$

$$(2.3)$$

exists for any  $\tau_1, \tau_2 \in \mathcal{T}$ , then

$$\frac{\sqrt{n}}{\|\mathbf{u}_n\|} \left( \mathbf{u}_n^{\top} \widehat{\boldsymbol{\gamma}}(\cdot) - \mathbf{u}_n^{\top} \boldsymbol{\gamma}_n(\tau) \right) \rightsquigarrow \mathbb{G}(\cdot) \text{ in } \ell^{\infty}(\mathcal{T}), \tag{2.4}$$

where  $\mathbb{G}(\cdot)$  is a centered Gaussian process with the covariance function H defined as (2.3). In particular, there exists a version of  $\mathbb{G}$  with almost surely continuous sample paths.

The proof of Theorem 2.1 is given in Section A.1. Theorem 2.1 holds under very general conditions. For transformations **Z** that have a specific local structure, the assumptions on  $m, \xi_m$  can be relaxed considerably. Details are provided in Section 2.2.

In the end, we illustrate Theorem 2.1 in linear quantile regression models with *increasing* dimension, in which  $g_n$ ,  $c_n$  and  $\chi_{\gamma_n}(\mathbf{u}, \mathbf{Z})$  are trivially zero. As far as we are aware, this is the first quantile process result for linear models with increasing dimension.

Corollary 2.2. (Linear models with increasing dimension) Suppose (A1)-(A3) hold with  $\mathbf{Z}(X) = X$  and  $Q(x;\tau) = x^{\top} \gamma_n(\tau)$  for any x and  $\tau \in \mathcal{T}$ . Assume that  $m^3 \xi_m^2(\log n)^3 = o(n)$ . In addition, if  $\mathbf{u}_n \in \mathbb{R}^m$  is such that the following limit

$$H_1(\tau_1, \tau_2; \mathbf{u}_n) := \lim_{n \to \infty} \|\mathbf{u}_n\|^{-2} \mathbf{u}_n^{\top} J_m^{-1}(\tau_1) \mathbb{E}[XX^{\top}] J_m^{-1}(\tau_2) \mathbf{u}_n(\tau_1 \wedge \tau_2 - \tau_1 \tau_2)$$
(2.5)

exists for any  $\tau_1, \tau_2 \in \mathcal{T}$ , then (2.4) holds with the covariance function  $H_1$  defined in (2.5). Moreover, by setting  $\mathbf{u}_n = x_0$ , we have for any fixed  $x_0$ 

$$\frac{\sqrt{n}}{\|x_0\|} \left\{ \widehat{Q}(x_0; \cdot) - Q(x_0; \cdot) \right\} \leadsto \mathbb{G}(x_0; \cdot) \text{ in } \ell^{\infty}(\mathcal{T}),$$

where  $\mathbb{G}(x_0;\cdot)$  is a centered Gaussian process with covariance function  $H_1(\tau_1,\tau_2;x_0)$ . In particular, there exists a version of  $\mathbb{G}$  with almost surely continuous sample paths.

#### 2.2 Local Basis Series Estimator

In this section, we assume that  $\mathbf{Z}(\cdot)$  corresponds to a basis expansion with "local" support. Our main motivation for considering this setting is that it allows to considerably weaken assumptions on  $m, \xi_m$  made in the previous section. To distinguish such basis functions from the general setting in the previous section, we shall use the notation  $\mathbf{B}$  instead of  $\mathbf{Z}$ . Let  $\widehat{\boldsymbol{\beta}}(\tau)$  be defined as

$$\widehat{\boldsymbol{\beta}}(\tau) := \underset{\mathbf{b} \in \mathbb{R}^m}{\operatorname{argmin}} \sum_{i} \rho_{\tau} \{ Y_i - \mathbf{b}^{\top} \mathbf{B}_i \}, \tag{2.6}$$

where  $\mathbf{B}_i = \mathbf{B}(X_i)$ . The notion of "local support" is made precise in the following sense.

(L) For each x, the basis vector  $\mathbf{B}(x)$  has zeroes in all but at most r consecutive entries, where r is fixed. Moreover,  $\sup_{x,\tau} \mathbb{E}[\mathbf{B}(x)^{\top} \widetilde{J}_m(\tau)^{-1} \mathbf{B}(X)] = O(1)$ .

The above assumption holds for certain choices of basis functions, e.g., univariate B-splines.

**Example 2.3.** Let  $\mathcal{X} = [0, 1]$ , assume that (A2)-(A3) hold and that the density of X over  $\mathcal{X}$  is uniformly bounded away from zero and infinity. Consider the space of polynomial splines of order q with k uniformly spaced knots  $0 = t_1 < ... < t_k = 1$  in the interval [0, 1]. The space of such splines can be represented through linear combinations of the basis functions  $B_1, ..., B_{k-q-1}$  with each basis function  $B_j$  having support contained in the interval  $[t_j, t_{j+q+1})$ . Let  $\mathbf{B}(x) := (B_1(x), ..., B_{k-q-1}(x))^{\top}$ . Then the first part of assumption (L) holds with r = q. The condition  $\sup_{x,\tau} \mathbb{E}[|\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau)\mathbf{B}(X)|] = O(1)$  is verified in the Appendix, see Section A.2.

Condition (L) ensures that the matrix  $\widetilde{J}_m(\tau)$  has a band structure, which is useful for bounding the off-diagonal entries of  $\widetilde{J}_m^{-1}(\tau)$ . See Lemma 6.3 in Zhou et al. (1998) for additional details.

Throughout this section, consider the specific centering

$$\boldsymbol{\beta}_n(\tau) := \underset{\mathbf{b} \in \mathbb{R}^m}{\operatorname{argmin}} \mathbb{E} \left[ (\mathbf{B}^\top \mathbf{b} - Q(X; \tau))^2 f_{Y|X}(Q(X; \tau)|X) \right], \tag{2.7}$$

where  $\mathbf{B} = \mathbf{B}(X)$ . For basis functions satisfying condition (L), assumptions in Theorem 2.1 in the previous section can be replaced by the following weaker version.

(B1) Assume that  $\xi_m^4(\log n)^6 = o(n)$  and letting  $\widetilde{c}_n := \sup_{x,\tau} |\beta_n(\tau)^\top \mathbf{B}(x) - Q(x;\tau)|$  with  $\widetilde{c}_n^2 = o(n^{-1/2})$ , where  $||\mathbf{B}(X_i)|| \leq \xi_m$  almost surely.

Note that the condition  $\xi_m^4(\log n)^6 = o(n)$  in (B1) is less restrictive than  $m^3\xi_m^2(\log n)^3 = o(n)$  required in Theorem 2.1. For instance, in the setting of Example 2.3 where  $\xi_m \approx \sqrt{m}$ , we only require  $m^2(\log n)^6 = o(n)$ , which is weaker than  $m^4(\log n)^3 = o(n)$  in Theorem 2.1. This improvement is made possible based on the local structure of the spline basis.

In the setting of Example 2.3, bounds on  $\tilde{c}_n$  can be obtained provided that the function  $x \mapsto Q(x;\tau)$  is smooth for all  $\tau \in \mathcal{T}$ . For instance, assuming that  $Q(\cdot;\cdot) \in \Lambda_c^{\eta}(\mathcal{X},\mathcal{T})$  with  $\mathcal{X} = [0,1]$  and integer  $\eta$ , Remark B.1 shows that  $\tilde{c}_n = O(m^{-\lfloor \eta \rfloor})$ . Thus the condition  $\tilde{c}_n^2 = o(n^{-1/2})$  holds provided that  $m^{-2\lfloor \eta \rfloor} = o(n^{1/2})$ . Since for splines we have  $\xi_m \sim m^{1/2}$ , this is compatible with the restrictions imposed in assumption (B1) provided that  $\eta \geq 1$ .

**Theorem 2.4.** (Nonparametric models with local basis functions) Assume that conditions (A1)-(A3) hold with  $\mathbf{Z} = \mathbf{B}$ , (L) holds for  $\mathbf{B}$  and (B1) for  $\beta_n(\tau)$ . Assume that the set  $\mathcal{I}$  consists of at most L consecutive integers, where  $L \geq 1$  is fixed. Then for any  $\mathbf{u}_n \in \mathbb{R}^m_{\mathcal{I}}$ , (2.2) holds with  $\widehat{\gamma}(\tau)$ ,  $\gamma_n(\tau)$  and  $\mathbf{Z}$  being replaced by  $\widehat{\beta}(\tau)$ ,  $\beta_n(\tau)$  and  $\mathbf{B}$ . In addition, if the following limit

$$\widetilde{H}(\tau_1, \tau_2; \mathbf{u}_n) := \lim_{n \to \infty} \|\mathbf{u}_n\|^{-2} \mathbf{u}_n^{\mathsf{T}} J_m^{-1}(\tau_1) \mathbb{E}[\mathbf{B}\mathbf{B}^{\mathsf{T}}] J_m^{-1}(\tau_2) \mathbf{u}_n(\tau_1 \wedge \tau_2 - \tau_1 \tau_2)$$
(2.8)

exists for any  $\tau_1, \tau_2 \in \mathcal{T}$ , then (2.4) holds with the same replacement as above, and the limit  $\mathbb{G}$  is a centered Gaussian process with covariance function  $\widetilde{H}$  defined as (2.8). Moreover, for any  $x_0$ , let  $\widehat{Q}(x_0; \tau) := \mathbf{B}(x_0)^{\top} \widehat{\beta}(\tau)$  and assume that  $\widetilde{c}_n = o(\|\mathbf{B}(x_0)\|n^{-1/2})$ . Then

$$\frac{\sqrt{n}}{\|\mathbf{B}(x_0)\|} \left\{ \widehat{Q}(x_0; \cdot) - Q(x_0; \cdot) \right\} \leadsto \mathbb{G}(x_0; \cdot) \text{ in } \ell^{\infty}(\mathcal{T}), \tag{2.9}$$

where  $\mathbb{G}(x_0;\cdot)$  is a centered Gaussian process with covariance function  $\widetilde{H}(\tau_1,\tau_2;\mathbf{B}(x_0))$ . In particular, there exists a version of  $\mathbb{G}$  with almost surely continuous sample paths.

The proof of Theorem 2.4 is given in Section A.2.

Remark 2.5. The proof of Theorem 2.4 and the related Bahadur representation result in Section 5.2 crucially rely on the fact that the elements of  $\widetilde{J}_m(\tau)^{-1}$  decay exponentially fast in their distance from the diagonal, i.e. a bound of the form  $|(\widetilde{J}_m(\tau)^{-1})_{i,j}| \leq C\gamma^{|i-j|}$  for some  $\gamma < 1$ . Assumption (L) provides one way to guarantee such a result. We conjecture that similar results can be obtained for more classes of basis functions as long as the entries of  $\widetilde{J}_m(\tau)^{-1}$  decay exponentially fast in their distance from suitable subsets of indices in  $(j,j') \in \{1,...,m\}^2$ . This kind of result can be obtained for matrices  $\widetilde{J}_m(\tau)$  with specific sparsity patterns, see for instance Demko et al. (1984). In particular, we conjecture that such arguments can be applied for tensor product B-splines, see Example 1 in Section 5 of Demko et al. (1984). A detailed investigation of this interesting topic is left to future research.

We conclude this section by discussing a special case where the limit in (2.9) can be characterized more explicitly.

**Remark 2.6.** The covariance function  $\widetilde{H}$  can be explicitly characterized under  $\mathbf{u}_n = \mathbf{B}(x)$  and univariate B-splines  $\mathbf{B}(x)$  on  $x \in [0,1]$ , with an order r and equidistant knots  $0 = t_1 < ... < t_k = 1$ . Assume additional to (A3) that

$$\sup_{t \in \mathcal{X}, \tau \in \mathcal{T}} \left| \partial_x |_{x=t} f_{Y|X} (Q(x;\tau)|x) \right| < C, \text{ where } C > 0 \text{ is a constant},$$
 (2.10)

and the density  $f_X(x)$  for X is bounded above, then under  $\tilde{c}_n = o(\|\mathbf{B}(x_0)\|n^{-1/2})$ , (2.9) in Theorem 2.4 can be rewritten as

$$\sqrt{\frac{n}{\mathbf{B}(x_0)^{\top} \mathbb{E}[\mathbf{B}\mathbf{B}^{\top}]^{-1} \mathbf{B}(x_0)}} \Big( \mathbf{B}(x_0)^{\top} \widehat{\boldsymbol{\beta}}(\cdot) - Q(x_0; \cdot) \Big) \leadsto \mathbb{G}(\cdot; x_0) \text{ in } \ell^{\infty}(\mathcal{T}), \tag{2.11}$$

where the Gaussian process  $\mathbb{G}(\cdot; x_0)$  is defined by the following covariance function

$$\widetilde{H}(\tau_1, \tau_2; x_0) = \frac{\tau_1 \wedge \tau_2 - \tau_1 \tau_2}{f_{Y|X}(Q(x_0; \tau_1)|x_0) f_{Y|X}(Q(x_0; \tau_2)|x_0)}.$$

Although we only show the univariate case here, the same arguments are expected to hold for tensor-product B-spline based on the same reasoning. See Section A.2 for a proof of this remark.

# 3 Joint Weak Convergence for Partial Linear Models

In this section, we consider partial linear models of the form

$$Q(X;\tau) = V^{\top} \alpha(\tau) + h(W;\tau), \tag{3.1}$$

where  $X = (V^{\top}, W^{\top})^{\top} \in \mathbb{R}^{k+k'}$  and  $k, k' \in \mathbb{N}$  are fixed. An interesting *joint* weak convergence result is obtained for  $(\widehat{\alpha}(\tau), \widehat{h}(w_0; \tau))$  at any fixed  $w_0$ . More precisely,  $\widehat{\alpha}(\tau)$  and  $\widehat{h}(w; \tau)$  (after proper scaling and centering) are proved to be asymptotically independent at any fixed  $\tau \in \mathcal{T}$ . Therefore, the "joint asymptotics phenomenon" first discovered in Cheng and Shang (2015) persists even for *non-smooth* quantile loss functions. Such a theoretical result is practically useful for joint inference on  $\alpha(\tau)$  and  $h(W; \tau)$ ; see Cheng and Shang (2015).

Expanding  $w \mapsto h(w; \tau)$  in terms of basis vectors  $w \mapsto \widetilde{\mathbf{Z}}(w)$ , we can approximate (3.1) through the series expansion  $\mathbf{Z}(x)^{\top} \gamma_n^{\dagger}(\tau)$  by setting  $\mathbf{Z}(x) = (v^{\top}, \widetilde{\mathbf{Z}}(w)^{\top})^{\top}$ . In this section,  $\widetilde{\mathbf{Z}} : \mathbb{R}^{k'} \to \mathbb{R}^m$  is regarded as a general basis expansion that does not need to satisfy the local support assumptions in the previous section. Estimation is performed in the following form

$$\widehat{\boldsymbol{\gamma}}^{\dagger}(\tau) = (\widehat{\boldsymbol{\alpha}}(\tau)^{\top}, \widehat{\boldsymbol{\beta}}^{\dagger}(\tau)^{\top})^{\top} := \underset{\mathbf{a} \in \mathbb{R}^{k}, \mathbf{b} \in \mathbb{R}^{m}}{\operatorname{argmin}} \sum_{i} \rho_{\tau} (Y_{i} - \mathbf{a}^{\top} V_{i} - \mathbf{b}^{\top} \widetilde{\mathbf{Z}}(W_{i})). \tag{3.2}$$

For a theoretical analysis of  $\hat{\gamma}^{\dagger}$ , define population coefficients  $\gamma_n^{\dagger}(\tau) := (\boldsymbol{\alpha}(\tau)^{\top}, \boldsymbol{\beta}_n^{\dagger}(\tau)^{\top})^{\top}$ , where

$$\boldsymbol{\beta}_{n}^{\dagger}(\tau) := \operatorname*{argmin}_{\boldsymbol{\beta} \in \mathbb{R}^{m}} \mathbb{E}[f_{Y|X}(Q(X;\tau)|X)(h(W;\tau) - \boldsymbol{\beta}^{\top} \widetilde{\mathbf{Z}}(W))^{2}]$$
(3.3)

similar to (2.7); see Remark 3.3 for additional explanations.

To state our main result, we need to define a class of functions

$$\mathcal{U}_{\tau} := \left\{ w \mapsto g(w) \middle| g \text{ measurable and } \mathbb{E}[g^2(W) f_{Y|X}(Q(X;\tau)|X)] < \infty, w \in \mathbb{R}^{k'} \right\}.$$

For  $V \in \mathbb{R}^k$ , define for j = 1, ..., k,

$$h_{VW,j}(\cdot;\tau) := \underset{g \in \mathcal{U}_{\tau}}{\operatorname{argmin}} \mathbb{E}[(V_j - g(W))^2 f_{Y|X}(Q(X;\tau)|X)]$$
(3.4)

where  $V_j$  denotes the j-th entry of the random vector V. By the definition of  $h_{VW,j}$ , we have for all  $\tau \in \mathcal{T}$  and  $g \in \mathcal{U}_{\tau}$ ,

$$\mathbb{E}[(V - h_{VW}(W; \tau))g(W)f_{Y|X}(Q(X; \tau)|X)] = \mathbf{0}_k. \tag{3.5}$$

The matrix A is defined as coefficient matrix of the best series approximation of  $h_{VW}(W;\tau)$ :

$$A(\tau) := \underset{A}{\operatorname{argmin}} \mathbb{E}[f_{Y|X}(Q(X;\tau)|X) || h_{VW}(W;\tau) - A\widetilde{\mathbf{Z}}(W)||^{2}].$$
 (3.6)

The following two assumptions are needed in our main results.

(C1) Define  $c_n^{\dagger} := \sup_{\tau, w} |\widetilde{\mathbf{Z}}(w)^{\top} \boldsymbol{\beta}_n^{\dagger}(\tau) - h(w; \tau)|$  and assume that

$$\xi_m c_n^{\dagger} = o(1); \tag{3.7}$$

$$\sup_{\tau \in \mathcal{T}} \mathbb{E}[f_{Y|X}(Q(X;\tau)|X) \|h_{VW}(W;\tau) - A(\tau)\widetilde{\mathbf{Z}}(W))\|^2] = O(\lambda_n^2) \text{ with } \xi_m \lambda_n^2 = o(1); \quad (3.8)$$

(C2) We have  $\max_{j \le k} |V_j| < C$  almost surely for some constant C > 0.

Bounds on  $c_n^{\dagger}$  can be obtained under various assumptions on the basis expansion and smoothness of the function  $w \mapsto h(w;\tau)$ . Assume for instance that  $\mathcal{W} = [0,1]^{k'}$ , that  $h(\cdot;\cdot) \in \Lambda_c^{\eta}(\mathcal{W},\mathcal{T})$  and that  $\widetilde{\mathbf{Z}}$  corresponds to a tensor product B-spline basis of order q on  $\mathcal{W}$  with  $m^{1/k'}$  equidistant knots in each coordinate. Assuming that (V,W) has a density  $f_{V,W}$  such that  $0 < \inf_{v,w} f_{V,W}(v,w) \le \sup_{v,w} f_{V,W}(v,w) < \infty$  and  $q > \eta$ , we show in Remark B.1 that  $c_n^{\dagger} = O(m^{-\lfloor \eta \rfloor/k'})$ . Assumption (3.8) essentially states that  $h_{VW}$  can be approximated by a series estimator sufficiently well. This assumption is necessary to ensure that  $\alpha(\tau)$  is estimable at a parametric rate without undersmoothing when estimating  $h(\cdot;\tau)$ . In general, (3.8) is a non-trivial high-level assumption. It can be verified under smoothness conditions on the joint density of (X,Y) by applying arguments similar to those in Appendix S.1 of Cheng et al. (2014).

In addition to (C1)-(C2), we need the following condition.

#### (B1') Assume that

$$\left(\frac{m\xi_m^{2/3}\log n}{n}\right)^{3/4} + c_n^{\dagger 2}\xi_m = o(n^{-1/2}).$$

Moreover, assume that  $c_n^{\dagger} \lambda_n = o(n^{-1/2})$  and  $m c_n^{\dagger} \log n = o(1)$ .

We now are ready to state the main result of this section.

**Theorem 3.1.** Let Conditions (A1)-(A3) hold with  $\mathbf{Z} = (V^{\top}, \widetilde{\mathbf{Z}}(W)^{\top})^{\top}$ , (B1') and (C1)-(C2) hold for  $\beta_n^{\dagger}(\tau)$  defined in (3.3). For any sequence  $\mathbf{w}_n \in \mathbb{R}^m$  with  $\mathbb{E}[|\mathbf{w}_n^{\top} M_2(\tau_2)^{-1} \widetilde{\mathbf{Z}}(W)|] = o(||\mathbf{w}_n||)$  where  $M_2(\tau) := \mathbb{E}[\widetilde{\mathbf{Z}}(W)\widetilde{\mathbf{Z}}(W)^{\top} f_{Y|X}(Q(X;\tau)|X)]$ , if

$$\Gamma_{22}(\tau_1, \tau_2) = \lim_{n \to \infty} \|\mathbf{w}_n\|^{-2} \mathbf{w}_n^{\mathsf{T}} M_2(\tau_1)^{-1} \mathbb{E}[\widetilde{\mathbf{Z}}(W) \widetilde{\mathbf{Z}}(W)^{\mathsf{T}}] M_2(\tau_2)^{-1} \mathbf{w}_n$$
(3.9)

exists, then

$$\left(\frac{\sqrt{n}(\widehat{\boldsymbol{\alpha}}(\cdot) - \boldsymbol{\alpha}(\cdot))}{\frac{\sqrt{n}}{\|\mathbf{w}_n\|}} \mathbf{w}_n^{\top} (\widehat{\boldsymbol{\beta}}^{\dagger}(\cdot) - \boldsymbol{\beta}_n^{\dagger}(\cdot))\right) \rightsquigarrow (\mathbb{G}_1(\cdot), ..., \mathbb{G}_k(\cdot), \mathbb{G}_h(\cdot))^{\top} in (\ell^{\infty}(\mathcal{T}))^{k+1},$$
(3.10)

and the multivariate process  $(\mathbb{G}_1(\cdot),...,\mathbb{G}_k(\cdot),\mathbb{G}_h(\cdot))$  has the covariance function

$$\Gamma(\tau_1, \tau_2; \mathbf{w}_n) = (\tau_1 \wedge \tau_2 - \tau_1 \tau_2) \begin{pmatrix} \Gamma_{11}(\tau_1, \tau_2) & \mathbf{0}_k \\ \mathbf{0}_k^\top & \Gamma_{22}(\tau_1, \tau_2) \end{pmatrix}$$
(3.11)

with

$$\Gamma_{11}(\tau_1, \tau_2) = M_{1,h}(\tau_1)^{-1} \mathbb{E}\left[ (V - h_{VW}(W; \tau_1))(V - h_{VW}(W; \tau_2))^{\top} \right] M_{1,h}(\tau_2)^{-1}$$
(3.12)

where  $M_{1,h}(\tau) = \mathbb{E}[(V - h_{VW}(W;\tau))(V - h_{VW}(W;\tau))^{\top} f_{Y|X}(Q(X;\tau)|X)]$ . In addition, at any fixed  $w_0 \in \mathbb{R}^{k'}$ , let  $\mathbf{w}_n = \widetilde{\mathbf{Z}}(w_0)$  satisfy the above conditions,  $\widehat{h}(w_0;\tau) = \widetilde{\mathbf{Z}}(w_0)^{\top} \widehat{\boldsymbol{\beta}}^{\dagger}(\tau)$ ,  $c_n^{\dagger} = o(\|\widetilde{\mathbf{Z}}(w_0)\|n^{-1/2})$ , then

$$\left(\frac{\sqrt{n}\{\widehat{\boldsymbol{\alpha}}(\cdot) - \boldsymbol{\alpha}(\cdot)\}}{\frac{\sqrt{n}}{\|\widetilde{\mathbf{Z}}(w_0)\|}\{\widehat{h}(w_0; \cdot) - h(w_0; \cdot)\}}\right) \rightsquigarrow \left(\mathbb{G}_1(\cdot), ..., \mathbb{G}_k(\cdot), \mathbb{G}_h(w_0; \cdot)\right)^\top in \left(\ell^{\infty}(\mathcal{T})\right)^{k+1},$$
(3.13)

where  $(\mathbb{G}_1(\cdot),...,\mathbb{G}_k(\cdot),\mathbb{G}_h(w_0;\cdot))$  are centered Gaussian processes with joint covariance function  $\Gamma_{w_0}(\tau_1,\tau_2)$  of the form (3.11) where  $\Gamma_{22}(\tau_1,\tau_2)$  is defined through the limit in (3.9) with  $\mathbf{w}_n$  replaced by  $\widetilde{\mathbf{Z}}(w_0)$ . In particular, there exists a version of  $\mathbb{G}_h(w_0;\cdot)$  with almost surely continuous sample paths.

The proof of Theorem 3.1 is presented in Section A.3. The invertibility of the matrices  $M_{1,h}(\tau)$  and  $M_2(\tau)$  is discussed in Remark 5.5. In general,  $\widehat{\alpha}(\tau)$  is not semiparametric efficient, as its covariance matrix  $\tau(1-\tau)\Gamma_{11}$  does not achieve the efficiency bound given in Section 5 of Lee (2003).

The joint asymptotic process convergence result (in  $\ell^{\infty}(\mathcal{T})$ ) presented in Theorem 3.1 is new in the quantile regression literature. The block structure of covariance function  $\Gamma$  defined in (3.11) implies that  $\widehat{\alpha}(\tau)$  and  $\widehat{h}(w_0;\tau)$  are asymptotically independent for any fixed  $\tau$ . This effect was recently discovered by Cheng and Shang (2015) in the case of mean regression, named as *joint* asymptotics phenomenon.

Remark 3.2. We point out that  $\mathbb{E}[|\mathbf{w}_n^{\top} M_2(\tau_2)^{-1} \widetilde{\mathbf{Z}}(W)|] = o(||\mathbf{w}_n||)$  is a crucial sufficient condition for asymptotic independence between the parametric and nonparametric parts. We conjecture that this condition is also necessary. This condition holds, for example, for  $\mathbf{w}_n = \widetilde{\mathbf{Z}}(w_0)$  or  $\mathbf{w}_n = \partial_{w_j} \widetilde{\mathbf{Z}}(w_0)$  at a fixed  $w_0$ , j = 1, ..., k', where  $\widetilde{\mathbf{Z}}(w)$  is a vector of B-spline basis. However, this condition may not hold for other estimators. Consider for instance the case  $\mathcal{W} = [0, 1]$ , B-splines of order zero  $\widetilde{\mathbf{Z}}$  and the vector  $\mathbf{w}_n = \int_0^{\lambda} \widetilde{\mathbf{Z}}(w) dw$  for some  $\lambda > 0$ . In this case  $||\mathbf{w}_n|| \approx 1$ , and one can show that  $\mathbb{E}[|\mathbf{w}_n^{\top} M_2(\tau_2)^{-1} \widetilde{\mathbf{Z}}(W)|] \approx 1$  instead. A more detailed investigation of related questions is left to future research.

**Remark 3.3.** A seemingly more natural choice for the centering vector, which was also considered in Belloni et al. (2011), is

$$\boldsymbol{\gamma}_n^*(\tau) = (\boldsymbol{\alpha}_n^*(\tau), \boldsymbol{\beta}_n^*(\tau)) := \arg\min_{(\mathbf{a}, \mathbf{b})} \mathbb{E}[\rho_{\tau}(Y - \mathbf{a}^{\top}V - \mathbf{b}^{\top}\widetilde{\mathbf{Z}}(W))],$$

which gives  $g_n(\gamma_n^*(\tau)) = 0$ . However, a major drawback of centering with  $\gamma_n^*(\tau)$  is that in this representation, it is difficult to find bounds for the difference  $\alpha_n^*(\tau) - \alpha(\tau)$ .

# 4 Applications of Weak Convergence Results

In this section, we consider applications of the process convergence results to the estimation of conditional distribution functions and non-crossing quantile curves via rearrangement operators. For the former estimation, define the functional (see Dette and Volgushev (2008), Chernozhukov et al. (2010) or Volgushev (2013) for similar ideas)

$$\Phi: \begin{cases} \ell^{\infty}((\tau_L, \tau_U)) \to \ell^{\infty}(\mathbb{R}) \\ \Phi(f)(y) := \tau_L + \int_{\tau_L}^{\tau_U} \mathbf{1}\{f(\tau) < y\} d\tau. \end{cases}$$

A simple calculation shows that

$$\Phi(Q(x;\cdot))(y) = \begin{cases} \tau_L & \text{if } F_{Y|X}(y|x) < \tau_L \\ F_{Y|X}(y|x) & \text{if } \tau_L \le F_{Y|X}(y|x) \le \tau_U \\ \tau_U & \text{if } F_{Y|X}(y|x) > \tau_U. \end{cases}$$

The latter identity motivates the following estimator of the conditional distribution function

$$\widehat{F}_{Y|X}(y|x) := \tau_L + \int_{\tau_L}^{\tau_U} \mathbf{1}\{\widehat{Q}(x;\tau) < y\} d\tau,$$

where  $\widehat{Q}(x;\tau)$  denotes the estimator of the conditional quantile function in any of the three settings discussed in Sections 2 and 3. By following the arguments in Chernozhukov et al. (2010), one can easily show that under suitable assumptions the functional  $\Phi$  is compactly differentiable (see Section A.5 for more details). Hence, the general process convergence results in Sections 2 and 3 allow to easily establish the asymptotic properties of  $\widehat{F}_{Y|X}$  - see Theorem 4.1 at the end of this section.

The second functional of interest is the monotone rearrangement operator, defined as follows

$$\Psi: \left\{ \begin{array}{ll} \ell^{\infty}((\tau_L, \tau_U)) & \to & \ell^{\infty}((\tau_L, \tau_U)) \\ \Psi(f)(\tau) & := & \inf \left\{ y : \Phi(Q(x; \cdot))(y) \ge \tau \right\}. \end{array} \right.$$

The main motivation for considering  $\Psi$  is that the function  $\tau \mapsto \Psi(f)(\tau)$  is by construction nondecreasing. Thus for any initial estimator  $\widehat{Q}(x;\cdot)$ , its rearranged version  $\Psi(\widehat{Q}(x;\cdot))(\tau)$  is an estimator of the conditional quantile function which avoids the issue of quantile crossing. For more detailed discussions of rearrangement operators and their use in avoiding quantile crossing we refer the interested reader to (Dette and Volgushev, 2008) and (Chernozhukov et al., 2010).

**Theorem 4.1** (Convergence of  $\widehat{F}(y|x)$  and  $\Psi(\widehat{Q}(x;\tau))$ ). For any fixed  $x_0$  and an initial estimator  $\widehat{Q}(x_0,\cdot)$ , we have for any compact sets  $[\tau_U,\tau_L]\subset\mathcal{T},\mathcal{Y}\subset\mathcal{Y}_{0,\mathcal{T}}:=\{y:F_{Y|X}(y|x_0)\in\mathcal{T}\}$ 

$$a_n \big\{ \widehat{F}_{Y|X}(\cdot|x_0) - F_{Y|X}(\cdot|x_0) \big\} \leadsto -f_{Y|X}(F_{Y|X}(\cdot|x_0)|x_0) \mathbb{G} \big( x_0; F_{Y|X}(\cdot|x_0) \big) \text{ in } \ell^{\infty}(\mathcal{Y}),$$

$$a_n \big\{ \Psi(\widehat{Q}(x_0; \cdot))(\cdot) - Q(x_0; \cdot) \big\} \leadsto \mathbb{G}(x_0; \cdot) \text{ in } \ell^{\infty}((\tau_U, \tau_L)),$$

where  $\widehat{Q}(x_0,\cdot)$ , the normalization  $a_n$ , and the process  $\mathbb{G}(x_0;\cdot)$  are stated as follows

- 1. (Linear model with increasing dimension) Suppose  $\mathbf{Z}(X) = X$ ,  $\widehat{Q}(x_0, \cdot) = \widehat{\gamma}(\cdot)^{\top} x_0$  and the conditions in Corollary 2.2 hold. In this case, we have  $a_n = \sqrt{n}/\|x_0\|$ .  $\mathbb{G}(x_0; \cdot)$  is a centered Gaussian process with covariance function  $H_1(\tau_1, \tau_2; x_0)$  defined in (2.5).
- 2. (Nonparametric model) Suppose  $\widehat{Q}(x_0,\cdot) = \boldsymbol{\beta}(\cdot)^{\top} \mathbf{B}(x_0)$  and the conditions in Theorems 2.4 hold. In this case, we have  $a_n = \sqrt{n}/\|\mathbf{B}(x_0)\|$ .  $\mathbb{G}(x_0;\cdot)$  is a centered Gaussian process with covariance function  $\widetilde{H}(\tau_1,\tau_2;\mathbf{B}(x_0))$  defined in (2.8).
- 3. (Partial linear model) Suppose  $x_0^{\top} = (v_0^{\top}, w_0^{\top})$ ,  $\widehat{Q}(x_0, \cdot) = \widehat{\gamma}^{\dagger}(\tau)^{\top}(v_0^{\top}, \widetilde{\mathbf{Z}}(w_0)^{\top})^{\top}$  and the conditions in Theorem 3.1 hold. In this case, we have  $a_n = \sqrt{n}/\|\widetilde{\mathbf{Z}}(w_0)\|$ .  $\mathbb{G}(x_0; \cdot)$  is a centered Gaussian process with covariance function  $\Gamma_{22}(\tau_1, \tau_2; \widetilde{\mathbf{Z}}(w_0))$  defined in (3.9).

The proof of Theorem 4.1 is a direct consequence of the functional delta method. Details can be found in Section A.5.

# 5 Bahadur representations

In this section, we provide Bahadur representations for the estimators discussed in Sections 2 and 3. In Sections 5.1 and 5.2, we state Bahadur representations for general series estimators and a more specific choice of local basis function, respectively. In particular, the latter representation is developed with an improved remainder term. Section 5.3 contains a special case of the general theorem in Section 5.1 that is particularly tailored to partial linear models. The remainders in these representations are shown to have exponential tail probabilities (uniformly over  $\mathcal{T}$ ).

## 5.1 A Fundamental Bahadur Representation

Our first result gives a Bahadur representation for  $\widehat{\gamma}(\tau) - \gamma_n(\tau)$  for centering functions  $\gamma_n$  satisfying certain conditions. Recall the definition of  $\widehat{\gamma}(\tau)$  in (2.1). This kind of representation for quantile regression with an increasing number of covariates has previously been established in Theorem 2 of Belloni et al. (2011). Compared to their results, the Bahadur representation given below has several advantages. First, we allow for a more general centering. This is helpful for the analysis of partial linear models (see Sections 3 and S.1.2). Second, we provide exponential tail bounds on remainder terms, which is much more explicit and sharper than those in Belloni et al. (2011).

**Theorem 5.1.** Suppose Conditions (A1)-(A3) hold and that additionally  $m\xi_m^2 \log n = o(n)$ . Then, for any  $\gamma_n(\cdot)$  satisfying  $g_n(\gamma_n) = o(\xi_m^{-1})$  and  $c_n(\gamma_n) = o(1)$ , we have

$$\widehat{\gamma}(\tau) - \gamma_n(\tau) = -\frac{1}{n} J_m(\tau)^{-1} \sum_{i=1}^n \psi(Y_i, \mathbf{Z}_i; \gamma_n(\tau), \tau) + r_{n,1}(\tau) + r_{n,2}(\tau) + r_{n,3}(\tau) + r_{n,4}(\tau).$$

The remainder terms  $r_{n,j}$ 's can be bounded as follows:

$$\sup_{\tau \in \mathcal{T}} \|r_{n,1}(\tau)\| \le \frac{1}{\inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_m(\tau))} \frac{m\xi_m}{n} \quad a.s.$$
 (5.1)

Moreover, we have for any  $\kappa_n \ll n/\xi_m^2$ , sufficiently large n, and a constant C independent of n

$$P\left(\sup_{\tau \in \mathcal{T}} ||r_{n,j}(\tau)|| \le C\Re_j(\kappa_n)\right) \ge 1 - 2e^{-\kappa_n}, \quad j = 2, 3, 4,$$

where

$$\Re_2(\kappa_n) := \xi_m \left( \left( \frac{m}{n} \log n \right)^{1/2} + \left( \frac{\kappa_n}{n} \right)^{1/2} + g_n \right)^2, \tag{5.2}$$

$$\Re_3(\kappa_n) := \left( \left( \frac{m \log n}{n} \right)^{1/2} + \left( \frac{\kappa_n}{n} \right)^{1/2} + g_n \right)^{1/2} \left( \frac{m \xi_m \log n}{n} + \frac{\xi_m \kappa_n}{n} \right)^{1/2}, \tag{5.3}$$

$$\Re_4(\kappa_n) := c_n \left( \left( \frac{m}{n} \log n \right)^{1/2} + \left( \frac{\kappa_n}{n} \right)^{1/2} \right) + g_n. \tag{5.4}$$

The proof for Theorem 5.1 can be found in Section S.1.1.

#### 5.2 Bahadur Representation for Local Basis Series Estimator

In this section, we focus on basis expansions **B** satisfying (L) and derive a Bahadur representation for linear functionals of the form  $\mathbf{u}_n^{\top}(\widehat{\boldsymbol{\beta}}(\tau) - \boldsymbol{\beta}_n(\tau))$ , where the vector  $\mathbf{u}_n$  can have at most a finite

number of consecutive non-zero entries. Such linear functionals are of interest since the estimator of the quantile function itself as well as estimators of derivatives can be represented in exactly this form - see Remark 5.3 for additional details. The advantage of concentrating on vectors with this particular structure is that we can substantially improve the rates of remainder terms compared to the general setting in Theorem 5.1.

**Theorem 5.2.** Suppose Conditions (A1)-(A3) and (L) hold with  $\mathbf{Z}(x) = \mathbf{B}(x)$ . Assume additionally that  $m\xi_m^2(\log n)^2 = o(n)$  and that  $\widetilde{c}_n = o(1)$  and that  $\mathcal{I} \subset \{1, ..., m\}$  consists of at most L consecutive integers. Then, for  $\beta_n(\tau)$  defined as (2.7) and  $\mathbf{u}_n \in \mathcal{S}_{\mathcal{I}}^{m-1}$  we have

$$\mathbf{u}_n^{\top}(\widehat{\boldsymbol{\beta}}(\tau) - \boldsymbol{\beta}_n(\tau)) = -\mathbf{u}_n^{\top} \widetilde{J}_m(\tau)^{-1} n^{-1} \sum_{i=1}^n \mathbf{B}_i (\mathbf{1}\{Y_i \le \boldsymbol{\beta}_n(\tau)^{\top} \mathbf{B}_i\} - \tau) + \sum_{k=1}^4 r_{n,k}(\tau, \mathbf{u}_n), \quad (5.5)$$

where the remainder terms  $r_{n,j}$ 's can be bounded as follows:

$$\sup_{\mathbf{u}_n \in \mathcal{S}_{\tau}^{m-1}} \sup_{\tau \in \mathcal{T}} |r_{n,1}(\tau, \mathbf{u}_n)| \lesssim \frac{\xi_m \log n}{n} \quad a.s.$$
 (5.6)

$$\sup_{\mathbf{u}_n \in \mathcal{S}_T^{m-1}} \sup_{\tau \in \mathcal{T}} |r_{n,4}(\tau, \mathbf{u}_n)| \leq \frac{1}{n} + \frac{1}{2} \overline{f'} \widetilde{c}_n^2 \sup_{\mathbf{u}_n \in \mathcal{S}_T^{m-1}} \widetilde{\mathcal{E}}(\mathbf{u}_n, \mathbf{B}) \quad a.s.$$
 (5.7)

where  $\widetilde{\mathcal{E}}(\mathbf{u}_n, \mathbf{B}) := \sup_{\tau} \mathbb{E}|\mathbf{u}_n \widetilde{J}_m(\tau)^{-1}\mathbf{B}|$ . Moreover, we have for any  $\kappa_n \ll n/\xi_m^2$ , all sufficiently large n, and a constant C independent of n

$$P\Big(\sup_{\mathbf{u}_n \in \mathcal{S}_{\mathcal{I}}^{m-1}} \sup_{\tau \in \mathcal{T}} |r_{n,j}(\tau, \mathbf{u}_n)| \le C \widetilde{\Re}_j(\kappa_n)\Big) \ge 1 - n^2 e^{-\kappa_n}, \quad j = 2, 3$$

where

$$\widetilde{\Re}_{2}(\kappa_{n}) := C \sup_{\mathbf{u}_{n} \in \mathcal{S}_{\mathcal{I}}^{m-1}} \widetilde{\mathcal{E}}(\mathbf{u}_{n}, \mathbf{B}) \left( \frac{\xi_{m}(\log n + \kappa_{n}^{1/2})}{n^{1/2}} + \widetilde{c}_{n}^{2} \right)^{2}, \tag{5.8}$$

$$\widetilde{\Re}_3(\kappa_n) := C\left(\widetilde{c}_n \frac{\kappa_n^{1/2} \vee \log n}{n^{1/2}} + \frac{\xi_m^{1/2} (\kappa_n^{1/2} \vee \log n)^{3/2}}{n^{3/4}}\right). \tag{5.9}$$

Theorem 5.2 is proved in Section S.1.2. We note that by Hölder's inequality and assumptions (A1)-(A3), we have a simple bound for

$$\sup_{\mathbf{u}_n \in \mathcal{S}_{\mathcal{I}}^{m-1}} \widetilde{\mathcal{E}}(\mathbf{u}_n, \mathbf{B}) \leq \sup_{\mathbf{u}_n \in \mathcal{S}^{m-1}} \sup_{\tau} \left( \mathbf{u}_n^{\top} J_m^{-1}(\tau) \mathbb{E}[\mathbf{B} \mathbf{B}^{\top}] J_m^{-1}(\tau) \mathbf{u}_n \right)^{1/2} = O(1).$$

Remark 5.3. Theorem 5.2 enables us to study several quantities associated with the quantile function  $Q(x;\tau)$ . For instance, consider the spline setting of Example 2.3. Setting  $\mathbf{u}_n = \mathbf{B}(x)/\|\mathbf{B}(x)\|$  in the Theorem 5.2 yields a representation for  $\widehat{Q}(x;\tau)$ , while setting  $\mathbf{u}_n = \mathbf{B}'(x)/\|\mathbf{B}'(x)\|$  yields a representation for the estimator of the derivative  $\partial_x Q(x;\tau)$ . Uniformity in x follows once we observe that for different values of x, the support of the vector  $\mathbf{B}(x)$  is always consecutive so that there is at most  $n^l$ , l > 0, number of different sets  $\mathcal{I}$  that we need to consider.

#### 5.3 Bahadur Representation for Partial Linear Models

In this section, we provide a joint Bahadur representation for the parametric and non-parametric part of this model. Recall the partial linear model  $Q(X;\tau) = h(W;\tau) + \boldsymbol{\alpha}(\tau)^{\top}V$ .

**Theorem 5.4.** Let conditions (A1)-(A3), (C1)-(C2) hold with  $\mathbf{Z} = (V^{\top}, \widetilde{\mathbf{Z}}(W)^{\top})^{\top}$  and assume  $m\xi_m^2(\log n)^2 = o(n)$ . Then

$$\begin{pmatrix} \widehat{\boldsymbol{\alpha}}(\tau) - \boldsymbol{\alpha}(\tau) \\ \widehat{\boldsymbol{\beta}}^{\dagger}(\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau) \end{pmatrix} = -J_{m}(\tau)^{-1}n^{-1}\sum_{i=1}^{n} \mathbf{Z}_{i}(\mathbf{1}\{Y_{i} \leq \{\boldsymbol{\gamma}_{n}(\tau)^{\dagger}\}^{\top}\mathbf{Z}_{i}\} - \tau) + \sum_{j=1}^{4} r_{n,j}(\tau),$$

where the remainder terms  $r_{n,j}$ 's satisfy the bounds stated in Theorem 5.1 with  $g_n = \xi_m c_n^{\dagger 2}$ . Additionally, the matrix  $J_m^{-1}(\tau)$  can be represented as

$$J_m^{-1}(\tau) = \begin{pmatrix} M_1(\tau)^{-1} & -M_1(\tau)^{-1}A(\tau) \\ -A(\tau)^{\top}M_1(\tau)^{-1} & M_2(\tau)^{-1} + A(\tau)^{\top}M_1(\tau)^{-1}A(\tau) \end{pmatrix}$$
(5.10)

where

$$M_1(\tau) := \mathbb{E}[f_{Y|X}(Q(X;\tau)|X)(V - A(\tau)\widetilde{\mathbf{Z}}(W))(V - A(\tau)\widetilde{\mathbf{Z}}(W))^{\top}],$$
  

$$M_2(\tau) = \mathbb{E}[\widetilde{\mathbf{Z}}(W)\widetilde{\mathbf{Z}}(W)^{\top}f_{Y|X}(Q(X;\tau)|X)],$$

and  $A(\tau)$  is defined in (3.6).

See Section S.1.3 for the proof of Theorem 5.4.

**Remark 5.5.** We discuss the positive definiteness of  $M_1(\tau)$  and  $M_2(\tau)$ . Following Condition (A1) with  $\mathbf{Z} = (V^\top, \widetilde{\mathbf{Z}}(W)^\top)^\top$ , we have

$$1/M \le \inf_{\tau \in \mathcal{T}} \lambda_{\min}(M_1(\tau)) \le \sup_{\tau \in \mathcal{T}} \lambda_{\max}(M_1(\tau)) \le M; \tag{5.11}$$

$$1/M \le \inf_{\tau \in \mathcal{T}} \lambda_{\min}(M_2(\tau)) \le \sup_{\tau \in \mathcal{T}} \lambda_{\max}(M_2(\tau)) \le M, \tag{5.12}$$

for all n. To see this, observe that  $M_1(\tau) = (I_k - A(\tau))J_m(\tau)(I_k - A(\tau))^{\top}$  where  $I_k$  is the k-dimensional identity matrix, [A|B] denotes the block matrix with A in the left block and B in the right block, and

$$J_m(\tau) = \begin{pmatrix} M_1(\tau) + A(\tau)M_2(\tau)A(\tau)^\top & A(\tau)M_2(\tau) \\ M_2(\tau)A(\tau)^\top & M_2(\tau) \end{pmatrix},$$

whose form follows from the definition and the condition (3.5) (see the proof for Theorem 5.4 for more details). Thus, for an arbitrary nonzero vector  $\mathbf{a} \in \mathbb{R}^k$ , by Condition (A1),

$$0 < 1/M \le \mathbf{a}^{\top} M_1(\tau) \mathbf{a} = \mathbf{a}^{\top} [I_k| - A(\tau)] J_m(\tau) [I_k| - A(\tau)]^{\top} \mathbf{a} \le M < \infty$$

by the strictly positive definiteness of  $J_m(\tau)$  for some M > 0.

The strictly positive definiteness of  $M_2(\tau)$  follows directly from the observation that

$$0 < 1/M \le \mathbf{b}^{\top} M_2(\tau) \mathbf{b} = (\mathbf{0}_k^{\top}, \mathbf{b}^{\top}) J_m(\tau) (\mathbf{0}_k^{\top}, \mathbf{b}^{\top})^{\top} \le M < \infty,$$

for all nonzero  $\mathbf{b} \in \mathbb{R}^m$  and some M > 0.

#### APPENDIX

This appendix gives technical details of the results shown in the main text. Appendix A contains all the proofs for weak convergence results in Theorems 2.1, 2.4 and 3.1. Appendix B discusses basis approximation errors with full technical details.

Additional Notations. Define for a function  $x \mapsto f(x)$  that  $\mathbb{G}_n(f) := n^{1/2} \int f(x) (d\mathbb{P}_n(x) - dP(x))$  and  $||f||_{L_p(P)} = (\int |f(x)|^p dP(x))^{1/p}$  for  $0 . For a class of functions <math>\mathcal{G}$ , let  $||\mathbb{P}_n - P||_{\mathcal{G}} := \sup_{f \in \mathcal{G}} |\mathbb{P}_n f - Pf|$ . For any  $\epsilon > 0$ , the covering number  $N(\epsilon, \mathcal{G}, L_p)$  is the minimal number of balls of radius  $\epsilon$  (under  $L_p$ -norm) that is needed to cover  $\mathcal{G}$ . The bracketing number  $N_{[\ ]}(\epsilon, \mathcal{G}, L_p)$  is the minimal number of  $\epsilon$ -brackets that is needed to cover  $\mathcal{G}$ . An  $\epsilon$ -bracket refers to a pair of functions within an  $\epsilon$  distance:  $||u - l||_{L_p} < \epsilon$ . Throughout the proofs,  $C, C_1, C_2$  etc. will denote constants which do not depend on n but may have different values in different lines.

# APPENDIX A: Proofs for Process Convergence

#### A.1 Proof of Theorem 2.1

#### **A.1.1** Proof of (2.2)

Under conditions (A1)-(A3) and those in Theorem 2.1 it follows from Theorem 5.1 applied with  $\kappa_n = c \log n$  for a suitable constant c (note that the conditions  $g_n = o(\xi_m^{-1})$  and  $c_n = o(1)$  in Theorem 5.1 follow under the assumptions of Theorem 2.1) that

$$\mathbf{u}_{n}^{\top}(\widehat{\gamma}(\tau) - \gamma(\tau)) + \frac{1}{n}\mathbf{u}_{n}^{\top}J_{m}(\tau)^{-1}\sum_{i=1}^{n}\mathbf{Z}_{i}(\mathbf{1}\{Y_{i} \leq Q(X_{i};\tau)\} - \tau) = I(\tau) + o_{P}(\|\mathbf{u}_{n}\|n^{-1/2}),$$

where the remainder term is uniform in  $\mathcal{T}$  and

$$I(\tau) := -n^{-1}\mathbf{u}_n^{\top} J_m(\tau)^{-1} \sum_{i=1}^n \mathbf{Z}_i (\mathbf{1}\{Y_i \le \mathbf{Z}_i^{\top} \gamma_n(\tau))\} - \mathbf{1}\{Y_i \le Q(X_i; \tau)\}).$$

Under the assumption  $\chi_{\gamma_n}(\mathbf{u}_n, \mathbf{Z}) = o(\|\mathbf{u}_n\|n^{-1/2})$ , we have  $\sup_{\tau \in \mathcal{T}} |\mathbb{E}[I(\tau)]| = o(\|\mathbf{u}_n\|n^{-1/2})$  and moreover

$$\sup_{\tau \in \mathcal{T}} |I(\tau) - \mathbb{E}[I(\tau)]| \le \|\mathbf{u}_n\| [\inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_m(\tau))]^{-1} \|\mathbb{P}_n - P\|_{\mathcal{G}_5},$$

where the class of functions  $\mathcal{G}_5$  is defined as

$$\mathcal{G}_5(\mathbf{Z}, \boldsymbol{\gamma}_n) = \{(X, Y) \mapsto \mathbf{a}^\top \mathbf{Z}(X) \mathbf{1}\{\|\mathbf{Z}(X)\| \le \xi_m\} (\mathbf{1}\{Y \le \mathbf{Z}(X)^\top \boldsymbol{\gamma}_n(\tau)\} - \mathbf{1}\{Y \le Q(X; \tau)\}) | \tau \in \mathcal{T}, \mathbf{a} \in \mathcal{S}^{m-1}\}.$$

It remains to bound  $\|\mathbb{P}_n - P\|_{\mathcal{G}_5}$ . For any  $f \in \mathcal{G}_5$  and a sufficiently large C, we obtain

$$|f| \le |\mathbf{a}^{\top} \mathbf{Z}| \le \xi_m,$$
  
$$||f||_{L_2(P)}^2 \le 2\bar{f}c_n \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}]) \le Cc_n.$$

By Lemma 21 of Belloni et al. (2011), the VC index of  $\mathcal{G}_5$  is of the order O(m). Therefore, we obtain from (S.2.2)

$$\mathbb{E}\|\mathbb{P}_n - P\|_{\mathcal{G}_5} \le \widetilde{C}\left[\left(\frac{mc_n}{n}\log\frac{\xi_m}{\sqrt{c_n}}\right)^{1/2} + \frac{m\xi_m}{n}\log\frac{\xi_m}{\sqrt{c_n}}\right]. \tag{A.1}$$

For any  $\kappa_n > 0$ , let

$$r'_{N,3}(\kappa_n) = C \left[ \left( \frac{mc_n}{n} \log \frac{\xi_m}{\sqrt{c_n}} \right)^{1/2} + \frac{m\xi_m}{n} \log \frac{\xi_m}{\sqrt{c_n}} + \left( \frac{c_n}{n} \kappa_n \right)^{1/2} + \frac{\xi_m}{n} \kappa_n \right]$$

for a sufficiently large constant C > 0. We obtain from (S.2.3) combined with (A.1)

$$P\left\{\sup_{\tau\in\mathcal{T}}|I(\tau)|\geq \|\mathbf{u}_n\|r'_{N,3}(\kappa_n)+\sup_{\tau\in\mathcal{T}}\left|\mathbb{E}\left[I(\tau)\right]\right|\right\}\leq e^{-\kappa_n}.$$

Finally, note that under condition  $mc_n \log n = o(1)$  and

$$m^3 \xi_m^2 (\log n)^3 = o(n)$$

we have that  $r'_{N,3}(\log n) = o(n^{-1/2})$ . This completes the proof of (2.2).

#### **A.1.2** Proof of (2.4)

Throughout this subsection assume without loss of generality that  $\|\mathbf{u}_n\| = 1$ . It suffices to prove finite dimensional convergence and asymptotic equicontinuity. Asymptotic equicontinuity follows from (A.33). The existence of a version of the limit with continuous sample paths is a consequence of Theorem 1.5.7 and Addendum 1.5.8 in van der Vaart and Wellner (1996).

So, we only need to focus on finite dimensional convergence.

Let

$$\mathbb{G}_n(\tau) := \frac{1}{\sqrt{n}} \mathbf{u}_n^{\top} J_m(\tau)^{-1} \sum_{i=1}^n \mathbf{Z}_i (\mathbf{1} \{ Y_i \le Q(X_i; \tau) \} - \tau),$$

and  $\mathbb{G}$  be the Gaussian process defined in (2.4). From Cramér-Wold theorem, the goal is to show for arbitrary set of  $\{\tau_1, ..., \tau_L\}$  and  $\{\lambda_1, ..., \lambda_L\} \in \mathbb{R}^L$ , we have

$$\sum_{l=1}^{L} \lambda_l \mathbb{G}_n(\tau_l) \stackrel{d}{\to} \sum_{l=1}^{L} \lambda_l \mathbb{G}(\tau_l).$$

Let the triangular array  $V_{n,i}(\tau) := n^{-1/2} \mathbf{u}_n^{\top} J_m(\tau)^{-1} \mathbf{Z}_i (\mathbf{1}\{Y_i \leq Q(X_i; \tau)\} - \tau)$ . Then for all  $\tau \in \mathcal{T}$ , we have  $\mathbb{E}[V_{n,i}(\tau)] = 0$ ,  $|V_{n,i}| \leq n^{-1/2} \xi_m$  and  $\operatorname{var}(V_{n,i}(\tau)) = n^{-1} \mathbf{u}_n^{\top} J_m(\tau)^{-1} \mathbb{E}[\mathbf{Z}_i \mathbf{Z}_i^{\top}] J_m(\tau)^{-1} \mathbf{u}_n \tau (1 - \tau) < \infty$  by Conditions (A1)-(A3). We can express  $\mathbb{G}_n(\tau) = \sum_{i=1}^n V_{n,i}(\tau)$  and  $\sum_{l=1}^L \lambda_l \mathbb{G}_n(\tau_l) = \sum_{i=1}^n \sum_{l=1}^L \lambda_l V_{n,i}(\tau_l)$ . Observe that  $\operatorname{var}(\sum_{i=1}^n \sum_{l=1}^L \lambda_l V_{n,i}(\tau_l)) =: \sigma_{n,L}^2$  where

$$\sigma_{n,L}^2 = \sum_{l,l'=1}^L \lambda_l \lambda_{l'} \mathbf{u}_n^\top J_m(\tau_l)^{-1} \mathbb{E}[\mathbf{Z}_i \mathbf{Z}_i^\top] J_m(\tau_{l'})^{-1} \mathbf{u}_n(\tau_l \wedge \tau_{l'} - \tau_l \tau_{l'}).$$

If  $0 = \lim_{n \to \infty} \sigma_{n,L}^2 = \sum_{l,l'=1}^L \lambda_l \lambda_{l'} H(\tau_l, \tau_{l'}; \mathbf{u}_n) = \operatorname{var}(\sum_{l=1}^L \lambda_l \mathbb{G}(\tau_l))$ , then by Markov's inequality  $\sum_{i=1}^n \sum_{l=1}^L \lambda_l V_{n,i}(\tau_l) \to 0$  in probability, which coincides with the distribution of  $\sum_{l=1}^L \lambda_l \mathbb{G}(\tau_l)$ ,

which is a single point mass at 0. Next, consider the case  $\sigma_{n,L}^2 \to \sigma_L^2 > 0$ . For sufficiently large n and arbitrary v > 0, Markov's inequality implies

$$\sigma_{n,L}^{-2} \sum_{i=1}^{n} \mathbb{E} \left[ \left( \sum_{l=1}^{L} \lambda_{l} V_{n,i}(\tau_{l}) \right)^{2} \mathbf{1} \left( \sum_{l=1}^{L} \lambda_{l} V_{n,i}(\tau_{l}) > v \right) \right]$$

$$\lesssim \xi_{m}^{2} n^{-1} \sigma_{n,L}^{-2} \sum_{i=1}^{n} \mathbb{E} \left[ \mathbf{1} \left( \sum_{l=1}^{L} \lambda_{l} V_{n,i}(\tau_{l}) > v \right) \right]$$

$$\lesssim \xi_{m}^{2} n^{-1} \sigma_{n,L}^{-2} v^{-2} \sum_{l,l'=1}^{L} \lambda_{l} \lambda_{l'} \mathbf{u}_{n}^{\top} J_{m}(\tau_{l})^{-1} \mathbb{E} \left[ \mathbf{Z}_{i} \mathbf{Z}_{i}^{\top} \right] J_{m}(\tau_{l'})^{-1} \mathbf{u}_{n}(\tau_{l} \wedge \tau_{l'} - \tau_{l} \tau_{l'})$$

$$= o(1)$$

since  $\xi_m^2 n^{-1} = o(1)$  by the assumption  $m\xi_m^2 \log n = o(n)$ . Hence the Lindeberg condition is verified. The finite dimensional convergence follows from Cramér-Wold devise. This completes the proof.

## A.2 Proofs of Theorem 2.4, Example 2.3 and Remark 2.6

We begin by introducing some notations and useful preliminary results. For a vector  $\mathbf{u} = (u_1, ..., u_m)^{\top}$  and a set  $\mathcal{I} \subset \{1, ..., m\}$ , let  $\mathbf{u}^{(\mathcal{I})} \in \mathbb{R}^m$  denote the vector that has entries  $u_i$  for  $i \in \mathcal{I}$  and zero otherwise. For a vector  $\mathbf{a} \in \mathbb{R}^m$ , let  $k_{\mathbf{a}}$  denote the position of the first non-zero entry of  $\mathbf{a}$  with  $\|\mathbf{a}\|_0$  non-zero consecutive entries

$$\mathcal{I}(\mathbf{a}, D) := \{ i : |i - k_{\mathbf{a}}| \le ||\mathbf{a}||_0 + D \}, \tag{A.2}$$

$$\mathcal{I}'(\mathbf{a}, D) := \{ 1 \le j \le m : \exists i \in \mathcal{I}(\mathbf{a}, D) \text{ such that } |j - i| \le \|\mathbf{a}\|_0 \}, \tag{A.3}$$

**Lemma A.1.** Under (L), for an arbitrary vector  $\mathbf{a} \in \mathbb{R}^m$  with at most  $\|\mathbf{a}\|_0$  non-zero consecutive entries we have for a constant  $\gamma \in (0,1)$  independent of  $n, \tau$ 

$$|(\mathbf{a}^{\top} \widetilde{J}_{m}^{-1}(\tau))_{j}| \le C_{1} \|\mathbf{a}\|_{\infty} \sum_{q=k_{\mathbf{a}}}^{k_{\mathbf{a}} + \|\mathbf{a}\|_{0}} \gamma^{|q-j|}.$$
 (A.4)

$$\|\mathbf{a}^{\top} \widetilde{J}_{m}^{-1}(\tau) - (\mathbf{a}^{\top} \widetilde{J}_{m}^{-1}(\tau))^{(\mathcal{I}(\mathbf{a},D))}\| \lesssim \|\mathbf{a}\|_{\infty} \|\mathbf{a}\|_{0} \gamma^{D}$$
(A.5)

$$\|\mathbf{a}^{\top} J_m^{-1}(\tau) - (\mathbf{a}^{\top} J_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{a},D))}\| \lesssim \|\mathbf{a}\|_{\infty} \|\mathbf{a}\|_{0} \gamma^{D}$$
 (A.6)

**Proof for Lemma A.1.** Under (L) the matrix  $\mathbf{Z}(x)\mathbf{Z}(x)^{\top}$  has no non-zero entries that are further than r away from the diagonal. Thus  $\widetilde{J}_m^{-1}$  is a band matrix with band width no larger that 2r. Apply similar arguments as in the proof of Lemma 6.3 in Zhou et al. (1998) to find that under (L) the entries of  $\widetilde{J}_m^{-1}(\tau)$  satisfy

$$\sup_{\tau,m} |(\widetilde{J}_m^{-1}(\tau))_{j,k}| \le C_1 \gamma^{|j-k|}$$

for some  $\gamma \in (0,1)$  and a constant  $C_1$  where both  $\gamma$  and  $C_1$  do not depend on  $n,\tau$ . It follows that

$$|(\mathbf{a}^{\top} \widetilde{J}_m^{-1}(\tau))_j| \le C_1 ||\mathbf{a}||_{\infty} \sum_{q=k_{\mathbf{a}}}^{k_{\mathbf{a}} + ||\mathbf{a}||_0} \gamma^{|q-j|}.$$

and thus (A.4) is established. For a proof of (A.5) note that by (A.4) we have

$$|(\mathbf{a}^{\top}\widetilde{J}_{m}^{-1}(\tau))_{j}| \leq C_{1} \|\mathbf{a}\|_{\infty} \sum_{q=k_{\mathbf{a}}}^{k_{\mathbf{a}} + \|\mathbf{a}\|_{0}} \gamma^{|q-j|} \leq C_{1} \|\mathbf{a}\|_{\infty} \|\mathbf{a}\|_{0} \gamma^{|k_{\mathbf{a}} - j| - \|\mathbf{a}\|_{0}}.$$

By the definition of  $\mathcal{I}(\mathbf{a}, D)$  we find

$$\|\mathbf{a}^{\top}\widetilde{J}_{m}^{-1}(\tau) - (\mathbf{a}^{\top}\widetilde{J}_{m}^{-1}(\tau))^{(\mathcal{I}(\mathbf{a},D))}\| \leq C\|\mathbf{a}\|_{\infty}\|\mathbf{a}\|_{0}\gamma^{D}$$

for a constant C independent of n. The proof of (A.6) is similar to the proof of (A.5).

**Proof for Theorem 2.4.** By Theorem 5.2 and Condition (B1), we first obtain

$$\mathbf{u}_n^{\top} (\widehat{\boldsymbol{\beta}}(\tau) - \boldsymbol{\beta}(\tau)) = -\frac{1}{n} \mathbf{u}_n^{\top} \widetilde{J}_m(\tau)^{-1} \sum_{i=1}^n \mathbf{B}_i (\mathbf{1} \{ Y_i \le \mathbf{B}(x)^{\top} \boldsymbol{\beta}_n(\tau) \} - \tau) + o_P(\|\mathbf{u}_n\| n^{-1/2}). \quad (A.7)$$

Let  $\widetilde{U}_{1,n}(\tau) := n^{-1}\mathbf{u}_n^{\top}\widetilde{J}_m(\tau)^{-1}\sum_{i=1}^n\mathbf{B}_i(\mathbf{1}\{Y_i \leq \mathbf{B}(x)^{\top}\boldsymbol{\beta}_n(\tau)\} - \tau)$ . We claim that

$$\mathbf{u}_n^{\top} (\widetilde{U}_{1,n}(\tau) - U_n(\tau)) = o_P(\|\mathbf{u}_n\| n^{-1/2}), \tag{A.8}$$

where  $U_n(\tau) := n^{-1}\mathbf{u}_n^{\top}J_m(\tau)^{-1}\sum_{i=1}^n\mathbf{B}_i(\mathbf{1}\{Y_i \leq Q(X_i;\tau)\} - \tau)$ . Given (A.8), the process convergence of  $\mathbf{u}_n^{\top}(\widehat{\boldsymbol{\beta}}(\tau) - \boldsymbol{\beta}(\tau))$  and continuity of the sample paths of the limiting process follows from process convergence of  $\mathbf{u}_n^{\top}U_n(\tau)$ , which can be shown via exactly the same steps as in Section A.1.2 by replacing  $\mathbf{Z}$  by  $\mathbf{B}$  given assumptions (A1)-(A3).

To show (A.8), we proceed in two steps. Given  $\mathbf{u}_n \in \mathcal{S}_{\mathcal{I}}^{m-1}$ , let  $U_{1,n}(\tau) := n^{-1}\mathbf{u}_n^{\top}\widetilde{J}_m(\tau)^{-1}\sum_{i=1}^n \mathbf{B}_i(\mathbf{1}\{Y_i \leq Q(X_i;\tau)\} - \tau)$ .

Step 1: 
$$\sup_{\tau \in \mathcal{T}} \left| \mathbf{u}_n^{\top} (\widetilde{U}_{1,n}(\tau) - U_{1,n}(\tau)) \right| = o_P(n^{-1/2})$$
, for all  $\mathbf{u}_n \in \mathcal{S}_{\mathcal{I}}^{m-1}$ .  
Let  $\widetilde{I}_0(\tau) := \mathbb{E} \left[ \mathbf{u}_n^{\top} (\widetilde{U}_{1,n}(\tau) - U_{1,n}(\tau)) \right]$  and observe the decomposition

$$\mathbf{u}_{n}^{\top} (\widetilde{\boldsymbol{U}}_{1,n}(\tau) - \boldsymbol{U}_{1,n}(\tau)) - \mathbb{E} \left[ \mathbf{u}_{n}^{\top} (\widetilde{\boldsymbol{U}}_{1,n}(\tau) - \boldsymbol{U}_{1,n}(\tau)) \right] + \mathbb{E} \left[ \mathbf{u}_{n}^{\top} (\widetilde{\boldsymbol{U}}_{1,n}(\tau) - \boldsymbol{U}_{1,n}(\tau)) \right]$$

$$= \left( \mathbf{u}_{n} \widetilde{J}_{m}^{-1}(\tau) - \left( \mathbf{u}_{n} \widetilde{J}_{m}^{-1}(\tau) \right)^{(\mathcal{I}(\mathbf{u}_{n},D))} \right) (\mathbb{P}_{n} - P) \left\{ \mathbf{B}_{i} (\mathbf{1} \{ Y_{i} \leq \mathbf{B}(x)^{\top} \boldsymbol{\beta}_{n}(\tau) \} - \mathbf{1} \{ Y_{i} \leq Q(X_{i};\tau) \}) \right\}$$

$$+ (\mathbb{P}_{n} - P) \left\{ \left( \mathbf{u}_{n} \widetilde{J}_{m}^{-1}(\tau) \right)^{(\mathcal{I}(\mathbf{u}_{n},D))} \mathbf{B}_{i} (\mathbf{1} \{ Y_{i} \leq \mathbf{B}(x)^{\top} \boldsymbol{\beta}_{n}(\tau) \} - \mathbf{1} \{ Y_{i} \leq Q(X_{i};\tau) \}) \right\} + \widetilde{I}_{0}(\tau)$$

$$=: \widetilde{I}_{1}(\tau) + \widetilde{I}_{2}(\tau) + \widetilde{I}_{0}(\tau).$$

For  $\sup_{\tau \in \mathcal{T}} |\widetilde{I}_0(\tau)|$ , by the construction of  $\beta_n(\tau)$  in (2.7),

$$\begin{split} \sup_{\tau \in \mathcal{T}} \left| \widetilde{I}_0(\tau) \right| &\leq \sup_{\tau \in \mathcal{T}} \left| \mathbf{u}_n^\top \widetilde{J}_m^{-1}(\tau) \mathbb{E} \Big[ \mathbf{B}_i \big( \mathbf{1} \{ Y_i \leq Q(X_i; \tau) \} - \mathbf{1} \{ Y_i \leq \mathbf{B}_i^\top \boldsymbol{\beta}_n(\tau) \} \big) \Big] \right| \\ &\leq \widetilde{c}_n^2 \overline{f}' \sup_{\mathbf{u} \in \mathcal{S}^{m-1}} \mathbb{E} |\mathbf{u}^\top \widetilde{J}_m^{-1}(\tau) \mathbf{B}| \leq \widetilde{c}_n^2 \overline{f}' \sup_{\mathbf{u} \in \mathcal{S}^{m-1}} \Big( \mathbf{u}^\top \widetilde{J}_m^{-1}(\tau) \mathbb{E} [\mathbf{B} \mathbf{B}^\top] \widetilde{J}_m^{-1}(\tau) \mathbf{u} \Big)^{1/2} = o(n^{-1/2}), \end{split}$$

where the final rate follows from assumptions (A2) and  $\tilde{c}_n^2 = o(n^{-1/2})$  in (B1).

By (A.5) in Lemma A.1, let  $D = c \log n$  for large enough c > 0, we have almost surely

$$\sup_{\tau \in \mathcal{T}} \left| \widetilde{I}_1(\tau) \right| \leq \sup_{\tau \in \mathcal{T}} \left\| \mathbf{u}_n^{\mathsf{T}} \widetilde{J}_m^{-1}(\tau) - (\mathbf{u}_n^{\mathsf{T}} \widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{u}_n, D))} \right\| \xi_m \leq \|\mathbf{u}_n\|_{\infty} \|\mathbf{u}_n\|_0 n^{c \log \gamma} \xi_m = o(n^{-1/2}).$$

For bounding  $\sup_{\tau \in \mathcal{T}} |\widetilde{I}_2(\tau)|$ , observe that

$$\frac{1}{n} \sum_{i=1}^{n} (\mathbf{u}_{n}^{\top} \widetilde{J}_{m}(\tau))^{(\mathcal{I}(\mathbf{u}_{n},D))} \mathbf{B}_{i} (\mathbf{1} \{ Y_{i} \leq \mathbf{B}_{i}^{\top} \boldsymbol{\beta}_{n}(\tau) \} - \mathbf{1} \{ Y_{i} \leq Q(X_{i};\tau) \}) 
= \frac{1}{n} \sum_{i=1}^{n} \mathbf{u}_{n}^{\top} \widetilde{J}_{m}(\tau) \mathbf{B}_{i}^{(\mathcal{I}(\mathbf{u}_{n},D))} (\mathbf{1} \{ Y_{i} \leq \mathbf{B}_{i}^{\top} \boldsymbol{\beta}_{n}(\tau) \} - \mathbf{1} \{ Y_{i} \leq Q(X_{i};\tau) \}) 
= \frac{1}{n} \sum_{\{i: \text{supp}(\mathbf{B}_{i}) \cap \mathcal{I}(\mathbf{u}_{n},D) \neq \emptyset\}} \mathbf{u}_{n}^{\top} \widetilde{J}_{m}(\tau) \mathbf{B}_{i}^{(\mathcal{I}(\mathbf{u}_{n},D))} (\mathbf{1} \{ Y_{i} \leq \mathbf{B}_{i}^{\top} \boldsymbol{\beta}_{n}(\tau) \} - \mathbf{1} \{ Y_{i} \leq Q(X_{i};\tau) \}) 
= \frac{1}{n} \sum_{\{i: \mathbf{B}_{i}^{(\mathcal{I}'(\mathbf{u}_{n},D)^{c})} = \mathbf{0} \}} \mathbf{u}_{n}^{\top} \widetilde{J}_{m}(\tau) \mathbf{B}_{i}^{(\mathcal{I}(\mathbf{u}_{n},D))} (\mathbf{1} \{ Y_{i} \leq \mathbf{B}_{i}^{\top} \boldsymbol{\beta}_{n}(\tau) \} - \mathbf{1} \{ Y_{i} \leq Q(X_{i};\tau) \}) 
= \frac{1}{n} \sum_{i=1}^{n} \mathbf{u}_{n}^{\top} \widetilde{J}_{m}(\tau) \mathbf{B}_{i}^{(\mathcal{I}(\mathbf{u}_{n},D))} (\mathbf{1} \{ Y_{i} \leq \mathbf{B}_{i}^{\top} \boldsymbol{\beta}_{n}(\tau)^{(\mathcal{I}'(\mathbf{u}_{n},D))} \} - \mathbf{1} \{ Y_{i} \leq Q(X_{i};\tau) \}), 
= \frac{1}{n} \sum_{i=1}^{n} \mathbf{u}_{n}^{\top} \widetilde{J}_{m}(\tau) \mathbf{B}_{i}^{(\mathcal{I}(\mathbf{u}_{n},D))} (\mathbf{1} \{ Y_{i} \leq \mathbf{B}_{i}^{\top} \boldsymbol{\beta}_{n}(\tau)^{(\mathcal{I}'(\mathbf{u}_{n},D))} \} - \mathbf{1} \{ Y_{i} \leq Q(X_{i};\tau) \}),$$

where the third equality follows from the fact that  $\mathbf{B}_{i}^{(\mathcal{I}(\mathbf{u}_{n},D))} \neq 0$  can only happen for  $i \in \{i : \mathbf{B}_{i}^{(\mathcal{I}'(\mathbf{u}_{n},D)^{c})} = \mathbf{0}\}$ , because  $\mathbf{B}$  can only be nonzero in r consecutive entries by assumption (L), where  $\mathcal{I}'(\mathbf{u}_{n},D)^{c} = \{1,...,m\} - \mathcal{I}'(\mathbf{u}_{n},D)$  is the complement of  $\mathcal{I}'(\mathbf{u}_{n},D)$  in  $\{1,...,m\}$ . By restricting ourselves on set  $\{i : \mathbf{B}_{i}^{(\mathcal{I}'(\mathbf{u}_{n},D)^{c})} = \mathbf{0}\}$ , it is enough to look at the coefficient  $\boldsymbol{\beta}_{n}(\tau)^{(\mathcal{I}'(\mathbf{u}_{n},D))}$  in the last equality in (A.9). Hence,

$$\sup_{\tau \in \mathcal{T}} |\widetilde{I}_{2}(\tau)| \leq \|\mathbb{P}_{n} - P\|_{\widetilde{\mathcal{G}}_{5}(\mathcal{I}(\mathbf{u}_{n}, D), \mathcal{I}'(\mathbf{u}_{n}, D))}$$

where for any two index sets  $\mathcal{I}_1$  and  $\mathcal{I}'_1$ 

$$\widetilde{\mathcal{G}}_5(\mathcal{I}_1, \mathcal{I}_1') = \left\{ (X, Y) \mapsto \mathbf{a}^\top \mathbf{B}(X)^{(\mathcal{I}_1)} \left( \mathbf{1} \{ Y \leq \mathbf{B}(X)^\top \mathbf{b}^{(\mathcal{I}_1')} \} - \mathbf{1} \{ Y \leq Q(X; \tau) \} \right) \middle| \tau \in \mathcal{T}, \mathbf{b} \in \mathbb{R}^m, \mathbf{a} \in \mathcal{S}^{m-1} \right\}.$$

With the choice of  $D = c \log n$ , the cardinality of both  $\mathcal{I}(\mathbf{u}_n, D)$  and  $\mathcal{I}'(\mathbf{u}_n, D)$  is of order  $O(\log n)$ . Hence, the VC index of  $\widetilde{\mathcal{G}}_5(\mathcal{I}(\mathbf{u}_n, D), \mathcal{I}'(\mathbf{u}_n, D))$  is bounded by  $O(\log n)$ . Note that for any  $f \in \widetilde{\mathcal{G}}_5(\mathcal{I}(\mathbf{u}_n, D), \mathcal{I}'(\mathbf{u}_n, D))$ ,  $|f| \lesssim \xi_m$  and  $||f||_{L_2(P)} \lesssim \widetilde{c}_n$ . Applying (S.2.3) yields

$$P\bigg(\sup_{\tau \in \mathcal{T}} \big| \widetilde{I}_2(\tau) \big| \le C \bigg[ \bigg( \frac{\widetilde{c}_n(\log n)^2}{n} \bigg)^{1/2} + \frac{\xi_m(\log n)^2}{n} + \bigg( \frac{\widetilde{c}_n \kappa_n}{n} \bigg)^{1/2} + \frac{\kappa_n \xi_m}{n} \bigg] \bigg) \ge 1 - e^{\kappa_n}.$$

Taking  $\kappa_n = C \log n$ ,  $\widetilde{c}_n^2 = o(n^{-1/2})$  and  $\xi_m^4 (\log n)^6 = o(n)$  in (B1) implies that  $\sup_{\tau \in \mathcal{T}} \left| \widetilde{I}_2(\tau) \right| = o_P(n^{1/2})$ .

Step 2: 
$$\sup_{\tau \in \mathcal{T}} \left| \mathbf{u}_n^{\top} \left( U_{1,n}(\tau) - U_n(\tau) \right) \right| = o_P(n^{-1/2})$$
, for all  $\mathbf{u}_n \in \mathcal{S}_{\mathcal{T}}^{m-1}$ .

Observe that

$$\mathbf{u}_{n}^{\top} \left( \widetilde{U}_{1,n}(\tau) - U_{n}(\tau) \right)$$

$$= \frac{1}{n} \left( \left\{ \mathbf{u}_{n}^{\top} \left( \widetilde{J}_{m}(\tau)^{-1} - J_{m}(\tau)^{-1} \right) \right\} - \left\{ \mathbf{u}_{n}^{\top} \left( \widetilde{J}_{m}(\tau)^{-1} - J_{m}(\tau)^{-1} \right) \right\}^{(\mathcal{I}(\mathbf{u}_{n},D))} \right) \sum_{i=1}^{n} \mathbf{B}_{i} (\mathbf{1} \left\{ Y_{i} \leq Q(X_{i};\tau) \right\} - \tau)$$

$$+ \frac{1}{n} \left\{ \mathbf{u}_{n}^{\top} \left( \widetilde{J}_{m}(\tau)^{-1} - J_{m}(\tau)^{-1} \right) \right\}^{(\mathcal{I}(\mathbf{u}_{n},D))} \sum_{i=1}^{n} \mathbf{B}_{i} (\mathbf{1} \left\{ Y_{i} \leq Q(X_{i};\tau) \right\} - \tau)$$

$$=: \widetilde{I}_{3}(\tau) + \widetilde{I}_{4}(\tau).$$

Applying (A.5) and (A.6) in Lemma A.1 with  $D = c \log n$  where c > 0 is chosen sufficiently large, we have almost surely

$$\sup_{\tau \in \mathcal{T}} \left| \widetilde{I}_{3}(\tau) \right| \leq \left( \sup_{\tau \in \mathcal{T}} \left\| \mathbf{u}_{n}^{\mathsf{T}} \widetilde{J}_{m}^{-1}(\tau) - (\mathbf{u}_{n}^{\mathsf{T}} \widetilde{J}_{m}^{-1}(\tau))^{(\mathcal{I}(\mathbf{u}_{n},D))} \right\| + \sup_{\tau \in \mathcal{T}} \left\| \mathbf{u}_{n}^{\mathsf{T}} J_{m}^{-1}(\tau) - (\mathbf{u}_{n}^{\mathsf{T}} J_{m}^{-1}(\tau))^{(\mathcal{I}(\mathbf{u}_{n},D))} \right\| \right) \xi_{m}$$

$$\leq 2 \|\mathbf{u}_{n}\|_{\infty} \|\mathbf{u}_{n}\|_{0} n^{c \log \gamma} \xi_{m} = o(n^{-1/2}). \tag{A.10}$$

Now it is left to bound  $\sup_{\tau \in \mathcal{T}} |\widetilde{I}_4(\tau)|$ . We have

$$\widetilde{I}_{4}(\tau) = \frac{1}{n} \sum_{i=1}^{n} \left\{ \mathbf{u}_{n}^{\top} (\widetilde{J}_{m}(\tau)^{-1} - J_{m}(\tau)^{-1}) \right\}^{(\mathcal{I}(\mathbf{u}_{n},D))} \mathbf{B}_{i} (\mathbf{1} \{ Y_{i} \leq Q(X_{i};\tau) \} - \tau)$$

$$= \frac{1}{n} \sum_{i=1}^{n} \mathbf{u}_{n}^{\top} (\widetilde{J}_{m}(\tau)^{-1} - J_{m}(\tau)^{-1}) \mathbf{B}_{i}^{(\mathcal{I}(\mathbf{u}_{n},D))} (\mathbf{1} \{ Y_{i} \leq Q(X_{i};\tau) \} - \tau).$$

Hence,

$$\sup_{\tau \in \mathcal{T}} |\widetilde{I}_4(\tau)| \leq \sup_{\tau \in \mathcal{T}} \|\mathbf{u}_n^\top (\widetilde{J}_m(\tau)^{-1} - J_m(\tau)^{-1})\| \|\mathbb{P}_n - P\|_{\mathcal{G}_0(\mathcal{I}(\mathbf{u}_n, D)) \cdot \mathcal{G}_4}$$

where for any  $\mathcal{I}$ ,

$$\mathcal{G}_0(\mathcal{I}) := \{ (\mathbf{B}, Y) \mapsto \mathbf{a}^\top \mathbf{B}^{(\mathcal{I})} \mathbf{1} \{ || \mathbf{B}|| \le \xi_m \} | \mathbf{a} \in \mathcal{S}^{m-1} \},$$

$$\mathcal{G}_4 := \{ (X, Y) \mapsto \mathbf{1} \{ Y_i \le Q(X; \tau) \} - \tau | \tau \in \mathcal{T} \}.$$

The cardinality of the set  $\mathcal{I}(\mathbf{u}_n, c \log n)$  is of order  $O(\log n)$ . Thus, the VC index for  $\mathcal{G}_0(\mathcal{I}(\mathbf{u}_n, D))$  is of order  $O(\log n)$ . The VC index of  $\mathcal{G}_4$  is 2 (see Lemma S.2.4). By Lemma S.2.2,

$$N(\mathcal{G}_0(\mathcal{I}(\mathbf{u}_n, D)) \cdot \mathcal{G}_4, L_2(\mathbb{P}_n); \varepsilon) \le \left(\frac{A \|F\|_{L_2(\mathbb{P}_n)}}{\varepsilon}\right)^{v_0(n)},$$

where  $v_0(n) = O(\log n)$ . In addition, for any  $f \in \mathcal{G}_0(\mathcal{I}(\mathbf{u}_n, D)) \cdot \mathcal{G}_4$ ,  $|f| \lesssim \xi_m$  and  $||f||_{L_2(P)} = O(1)$  by (A1). Furthermore, by assumptions (A1)-(A2) and the definition of  $\tilde{c}_n$ ,

$$\|\mathbf{u}_n^{\top} (\widetilde{J}_m(\tau)^{-1} - J_m(\tau)^{-1})\| \le \widetilde{c}_n \lambda_{\max}(\mathbb{E}[\mathbf{B}(X)\mathbf{B}(X)^{\top}]) \overline{f}' \lesssim \widetilde{c}_n.$$
(A.11)

By (S.2.3), we have for some constant C > 0,

$$P\bigg(\sup_{\tau\in\mathcal{T}}\big|\widetilde{I}_2(\tau)\big|\leq C\widetilde{c}_n\bigg[\Big(\frac{(\log n)^2}{n}\Big)^{1/2}+\frac{\xi_m(\log n)^2}{n}+\Big(\frac{\kappa_n}{n}\Big)^{1/2}+\frac{\kappa_n\xi_m}{n}\bigg]\bigg)\geq 1-e^{\kappa_n}.$$

Taking  $\kappa_n = C \log n$ , an application of (B1) completes the proof.

**Proof for Example 2.3.** As  $\widetilde{J}_m(\tau)$  is a band matrix, applying similar arguments as in the proof of Lemma 6.3 in Zhou et al. (1998) gives

$$\sup_{\tau,m} |(\widetilde{J}_m^{-1}(\tau))_{j,j'}| \le C_1 \gamma^{|j-j'|},\tag{A.12}$$

for some  $\gamma \in (0,1)$  and  $C_1 > 0$ . Let  $k_{\mathbf{B}(x)}$  be the index of the first nonzero element of the vector  $\mathbf{B}(x)$ . Then by (A.12), we have

$$\sup_{\tau,m} |(\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau))_j| \le C_1 ||\mathbf{B}(x)||_{\infty} \sum_{j'=k_{\mathbf{u}_n}}^{k_{\mathbf{B}(x)} + ||\mathbf{B}(x)||_0} \gamma^{|j'-j|},$$

and also

$$\sup_{\tau,m} \mathbb{E}|\mathbf{B}(x)^{\top} \widetilde{J}_{m}^{-1}(\tau) \mathbf{B}(X)| \le C_{1} \|\mathbf{B}(x)\|_{\infty} \max_{l \le m} \mathbb{E}|B_{l}(X)| \sum_{j=1}^{m} \sum_{j'=k_{\mathbf{B}(x)}}^{k_{\mathbf{B}(x)} + \|\mathbf{B}(x)\|_{0}} \gamma^{|j'-j|}.$$
(A.13)

Since  $\|\mathbf{B}(x)\|_0$  is bounded by a constant, the sum in (A.13) is bounded uniformly. Moreover, in the present setting we have  $\|\mathbf{B}(x)\|_{\infty} = O(m^{1/2})$  and  $\max_{l \leq m} \mathbb{E}|B_l(X)| = O(m^{-1/2})$ . Therefore, for each m we have

$$\sup_{\tau \in \mathcal{T}, x \in \mathcal{X}} \mathbb{E}|\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau) \mathbf{B}(X)| = O(1).$$

**Proof of Remark 2.6.** Consider the product  $B_j(x)B_{j'}(x)$  of two B-spline functions. The fact that  $B_j(x)$  is locally supported on  $[t_j, t_{j+r}]$  implies that for all j' satisfying  $|j-j'| \geq r$ ,  $B_j(x)B_{j'}(x) = 0$  for all x, where  $r \in \mathbb{N}$  is the degree of spline. This also implies  $J_m(\tau)$  and  $\mathbb{E}[\mathbf{B}\mathbf{B}^\top]$  are a band matrices with each column having at most  $L_r := 2r+1$  nonzero elements and each non-zero element is at most r entries away from the main diagonal. Recall also the fact that  $\max_{j\leq m} \sup_{t\in\mathbb{R}} |B_j(t)| \lesssim m^{1/2}$  (by the discussion following assumption (A1)).

Define  $J_{m,D}(\tau) := D_m(\tau)\mathbb{E}[\mathbf{B}\mathbf{B}^{\top}]$ , where matrix  $D_m(\tau) := diag(f_{Y|X}(Q(t_j;\tau)|t_j), \ j=1,...,m)$ , and  $R_m(\tau) := J_m(\tau) - J_{m,D}(\tau)$ . Both  $J_{m,D}(\tau)$  and  $R_m(\tau)$  have the same band structure as  $J_m(\tau)$ . For arbitrary  $j, j' = 1, ..., m, \tau \in \mathcal{T}$  and a universal constant C > 0,

$$\begin{aligned} &|(R_{m}(\tau))_{j,j'}| \\ &= \left| \mathbb{E} \left[ B_{j}(X) B_{j'}(X) \left\{ f_{Y|X}(Q(X;\tau)|X) - f_{Y|X}(Q(t_{j};\tau)|t_{j}) \right\} \right] \right| \\ &\leq 2 \max_{j \leq m} \sup_{t \in \mathbb{R}} |B_{j}(t)|^{2} \int_{0}^{1} \mathbf{1} \left\{ |x - t_{j}| \leq C \frac{r}{m} \right\} \left| f_{Y|X}(Q(x;\tau)|x) - f_{Y|X}(Q(t_{j};\tau)|t_{j}) \right| f_{X}(x) dx \\ &\leq 2Cm \int_{0}^{1} \mathbf{1} \left\{ |x - t_{j}| \leq C \frac{r}{m} \right\} |x - t_{j}| dx \\ &= O(m^{-1}), \end{aligned} \tag{A.14}$$

where the second inequality is an application of the upper bound of  $\max_{j \leq m} \sup_{t \in \mathbb{R}} |B_j(t)| \lesssim m^{1/2}$  and the local support property of  $B_j$ ; the third inequality follows by the assumption (2.10) and bounded  $f_X(x)$ . This shows that  $\max_{j,j'=1,\ldots,m} \sup_{\tau \in \mathcal{T}} |(R_m(\tau))_{j,j'}| = O(m^{-1})$ .

Now we show a stronger result that  $\sup_{\tau \in \mathcal{T}} ||R_m(\tau)|| = O(m^{-1/2})$  for later use. Let  $\mathbf{v} = (v_1, ..., v_m)$ . Denote  $k_j$  the index with the first nonzero entry in the jth column of  $R_m(\tau)$ . By the band structure of  $R_m(\tau)$ ,

$$\sup_{\tau \in \mathcal{T}} \|R_{m}(\tau)\|^{2} = \sup_{\tau \in \mathcal{T}} \sup_{\mathbf{v} \in \mathcal{S}^{m-1}} \|R_{m}(\tau)\mathbf{v}\|_{2}^{2} = \sup_{\tau \in \mathcal{T}} \sup_{\mathbf{v} \in \mathcal{S}^{m-1}} \sum_{j=1}^{m} \left(\sum_{i=k_{j}}^{k_{j}+L_{r}-1} v_{i}(R_{m}(\tau))_{i,j}\right)^{2}$$

$$\lesssim \sup_{\tau \in \mathcal{T}} \max_{j,j'} |(R_{m}(\tau))_{j,j'}|^{2} m = O(m^{-1}), \tag{A.15}$$

where the last equality follows by (A.14). Note that from assumptions (A1)-(A3) that

$$f_{\min} M^{-1} < \lambda_{\min}(J_{m,D}(\tau)) \le \lambda_{\max}(J_{m,D}(\tau)) < \bar{f}M$$
(A.16)

uniformly in  $\tau \in \mathcal{T}$ , where the constant M > 0 is defined as in Assumption (A1). Using (A.16), assumptions (A1)-(A3) and (A.15),

$$||J_m^{-1}(\tau) - J_{m,D}^{-1}(\tau)|| \le ||J_{m,D}^{-1}(\tau)|| ||J_{m,D}(\tau) - J_m(\tau)|| ||J_m^{-1}(\tau)|| \lesssim \sup_{\tau \in \mathcal{T}} ||R_m(\tau)|| = O(m^{-1/2})$$

uniformly in  $\tau \in \mathcal{T}$ .

Without loss of generality, from now on we drop the term  $\tau_1 \wedge \tau_2 - \tau_1 \tau_2$  out of our discussion and focus on the matrix part in the covariance function  $\widetilde{H}(\tau_1, \tau_2; \mathbf{u}_n)$  defined in (2.8). From (A1) we have  $\|\mathbb{E}[\mathbf{B}\mathbf{B}^{\top}]\| < M$  for some constant M > 0 so for any  $\tau_1, \tau_2 \in \mathcal{T}$ ,

$$\|\mathbf{u}_{n}\|^{-2} \left| \mathbf{u}_{n}^{\top} J_{m}^{-1}(\tau_{1}) \mathbb{E}[\mathbf{B}\mathbf{B}^{\top}] J_{m}^{-1}(\tau_{2}) \mathbf{u}_{n} - \mathbf{u}_{n}^{\top} J_{m,D}(\tau_{1})^{-1} \mathbb{E}[\mathbf{B}\mathbf{B}^{\top}] J_{m,D}(\tau_{2})^{-1} \mathbf{u}_{n} \right|$$

$$\lesssim \sup_{\tau \in \mathcal{T}} \|R_{m}(\tau)\| \sup_{\tau \in \mathcal{T}} \|\mathbb{E}[\mathbf{B}\mathbf{B}^{\top}] J_{m,D}(\tau)^{-1}\| + \sup_{\tau \in \mathcal{T}} \|R_{m}(\tau)\| \sup_{\tau \in \mathcal{T}} \|\mathbb{E}[\mathbf{B}\mathbf{B}^{\top}] J_{m}(\tau)^{-1}\|$$

$$= O(m^{-1/2}). \tag{A.17}$$

Moreover, note that

$$\mathbf{u}_{n}^{\top} J_{m,D}(\tau_{1})^{-1} \mathbb{E}[\mathbf{B}\mathbf{B}^{\top}] J_{m,D}(\tau_{2})^{-1} \mathbf{u}_{n} = \mathbf{u}_{n}^{\top} D_{m}(\tau_{1})^{-1} \mathbb{E}[\mathbf{B}\mathbf{B}^{\top}]^{-1} D_{m}(\tau_{2})^{-1} \mathbf{u}_{n}$$
(A.18)

If  $\mathbf{u}_n = \mathbf{B}(x)$ , observe that for l = 1, ..., m, as suggested by the local support property, we only need to focus on the index l satisfying  $|x - t_l| \le Cr/m$ , for a universal constant C > 0. We have

$$(\mathbf{B}(x)^{\top} D_m(\tau)^{-1})_l = B_l(x) f_{Y|X}(Q(t_l;\tau)|t_l)^{-1} = B_l(x) f_{Y|X}(Q(x;\tau)|x)^{-1} + R'(t_l), \tag{A.19}$$

where by assumption (2.10),  $|R'(t_l)| \leq \max_{j \leq m} \sup_{t \in \mathbb{R}} |B_j(t)| C f_{\min}^{-2} |x - t_l| = O(m^{-1/2})$ . Therefore, the sparse vector  $\mathbf{B}(x)^\top D_m(\tau)^{-1} = f_{Y|X}(Q(x;\tau)|x)^{-1}\mathbf{B}(x)^\top + \mathbf{a}_{\mathbf{B}(x)}$ , where  $\mathbf{a}_{\mathbf{B}(x)} \in \mathbb{R}^m$  is a vector with the same support as  $\mathbf{B}(x)$  (only  $r < \infty$  nonzero components) and with nonzero components of order  $O(m^{-1/2})$ . Hence,  $\|\mathbf{a}_{\mathbf{B}(x)}\| = O(m^{-1/2})$ . Continued from (A.18), for any  $x \in [0,1]$ ,

$$\|\mathbf{B}(x)\|^{-2} \left| \mathbf{B}(x)^{\top} D_{m}(\tau_{1})^{-1} \mathbb{E}[\mathbf{B}\mathbf{B}^{\top}]^{-1} D_{m}(\tau_{2})^{-1} \mathbf{B}(x) - \frac{\mathbf{B}(x)^{\top} \mathbb{E}[\mathbf{B}\mathbf{B}^{\top}]^{-1} \mathbf{B}(x)}{f_{Y|X}(Q(x;\tau_{1})|x) f_{Y|X}(Q(x;\tau_{2})|x)} \right|$$

$$\leq \|\mathbf{B}(x)\|^{-1} \|\mathbf{a}_{\mathbf{B}(x)}\| \sup_{\tau \in \mathcal{T}} \|\mathbb{E}[\mathbf{B}\mathbf{B}^{\top}]^{-1} D_{m}(\tau)^{-1}\| + \|\mathbf{B}(x)\|^{-1} \|\mathbf{a}_{\mathbf{B}(x)}\| \|\mathbb{E}[\mathbf{B}\mathbf{B}^{\top}]^{-1}\| / f_{\min}$$

$$= O(m^{-1}).$$

We observe that  $\mathbf{B}(x)^{\top} \mathbb{E}[\mathbf{B}\mathbf{B}^{\top}]^{-1}\mathbf{B}(x)$  does not depend on  $\tau_1$  and  $\tau_2$  and can be treated as a scaling factor and shifted out of the covariance function as (2.11). Therefore, we finish the proof.

A.3 Proof of Theorem 3.1

Observe that

$$\widehat{m{lpha}}_j(\cdot) - m{lpha}_j(\cdot) = \mathbf{e}_j^ op(\widehat{m{\gamma}}^\dagger(\cdot) - m{\gamma}_n^\dagger(\cdot))$$

where  $\mathbf{e}_j$  denotes the j-th unit vector in  $\mathbb{R}^{m+k}$  for j=1,...,m+k, and

$$\mathbf{w}_n^\top \big( \widehat{\boldsymbol{\beta}}^\dagger(\tau) - \boldsymbol{\beta}_n^\dagger(\tau) \big) = (\mathbf{0}_k^\top, \mathbf{w}_n^\top) (\widehat{\boldsymbol{\gamma}}^\dagger(\cdot) - \boldsymbol{\gamma}_n^\dagger(\cdot)).$$

Let  $h_n^{\dagger}(w,\tau) = \widetilde{\mathbf{Z}}(w)^{\top} \boldsymbol{\beta}_n^{\dagger}(\tau)$ . The following results will be established at the end of the proof.

$$\sup_{\tau \in \mathcal{T}, j=1,\dots,k} \left| \mathbb{E} \left[ \mathbf{e}_j^\top J_m(\tau)^{-1} \mathbf{Z} \left( \mathbf{1} \{ Y \le Q(X; \tau) \} - \mathbf{1} \{ Y \le \boldsymbol{\alpha}(\tau)^\top V + h_n^{\dagger}(W, \tau) \} \right) \right] \right| = o(n^{-1/2}), \tag{A.20}$$

$$\sup_{\tau \in \mathcal{T}} \left| (\mathbf{0}_k^\top, \mathbf{w}_n^\top / \|\mathbf{w}_n\|) J_m(\tau)^{-1} \mathbb{E} \left[ \mathbf{Z} (\mathbf{1} \{ Y \le Q(X; \tau) \} - \mathbf{1} \{ Y \le \boldsymbol{\alpha}(\tau)^\top V + h_n^\dagger(W, \tau) \}) \right] \right| = o(n^{-1/2}).$$
(A.21)

From Theorem 5.4, we obtain that under Condition (B1')

$$\mathbf{e}_{j}^{\top}(\widehat{\boldsymbol{\gamma}}^{\dagger}(\tau) - \boldsymbol{\gamma}_{n}^{\dagger}(\tau)) = -n^{-1/2}\mathbf{e}_{j}^{\top}J_{m}(\tau)^{-1}\mathbb{G}_{n}(\psi(\cdot;\boldsymbol{\gamma}_{n}^{\dagger}(\tau),\tau)) + o_{P}(n^{-1/2}), \ j = 1,...,k; \ (A.22)$$

$$(\mathbf{0}_{k}^{\top},\mathbf{w}_{n}^{\top})(\widehat{\boldsymbol{\gamma}}^{\dagger}(\tau) - \boldsymbol{\gamma}_{n}^{\dagger}(\tau)) = -n^{-1/2}(\mathbf{0}_{k}^{\top},\mathbf{w}_{n}^{\top})J_{m}(\tau)^{-1}\mathbb{G}_{n}(\psi(\cdot;\boldsymbol{\gamma}_{n}^{\dagger}(\tau),\tau)) + o_{P}(n^{-1/2})$$

$$(A.23)$$

uniformly in  $\tau \in \mathcal{T}$ . Equation (A.20) implies that for j = 1, ..., k

$$\mathbf{e}_{i}^{\top} J_{m}(\tau)^{-1} \mathbb{E}[\psi(Y_{i}, \mathbf{Z}_{i}; \boldsymbol{\gamma}_{n}^{\dagger}(\tau), \tau)] = \mathbf{e}_{i}^{\top} J_{m}(\tau)^{-1} \mathbb{E}[\mathbf{1}\{Y_{i} \leq Q(X_{i}; \tau)\} - \tau] + o(n^{-1/2}) = o(n^{-1/2}).$$

Following similar arguments as given in the proof of (2.2) in Section A.1.1, (A.20) and (A.22) imply that

$$\mathbf{e}_{j}^{\top}(\widehat{\boldsymbol{\gamma}}^{\dagger}(\tau) - \boldsymbol{\gamma}_{n}^{\dagger}(\tau)) = -n^{-1}\mathbf{e}_{j}^{\top}J_{m}(\tau)^{-1}\sum_{i=1}^{n}\mathbf{Z}_{i}(\mathbf{1}\{Y_{i} \leq Q(X_{i};\tau)\} - \tau) + o_{P}(n^{-1/2}), \ j = 1, ..., k.$$

uniformly in  $\tau \in \mathcal{T}$ . Similarly, by (A.21) and (A.23) we have

$$\|\mathbf{w}_n\|^{-1}\mathbf{w}_n^{\top}(\widehat{\boldsymbol{\beta}}^{\dagger}(\tau)-\boldsymbol{\beta}_n^{\dagger}(\tau)) = -n^{-1}(\mathbf{0}_k^{\top}, \|\mathbf{w}_n\|^{-1}\mathbf{w}_n^{\top})J_m(\tau)^{-1}\sum_{i=1}^n \mathbf{Z}_i(\mathbf{1}\{Y_i \leq Q(X_i; \tau)\} - \tau) + o_P(n^{-1/2}).$$

Thus, the claim will follow once we prove

$$\mathbf{G}_n(\cdot) := (\mathbb{G}_{n,1}(\cdot), ..., \mathbb{G}_{n,k}(\cdot), \mathbb{G}_{n,h}(\cdot)) \leadsto \mathbf{G}(\cdot) \text{ in } (\ell^{\infty}(\mathcal{T}))^{k+1}$$

where

$$\mathbb{G}_{n,j}(\tau) := -n^{-1/2} \mathbf{e}_j^{\top} J_m(\tau)^{-1} \sum_{i=1}^n \mathbf{Z}_i (\mathbf{1} \{ Y_i \le Q(X_i; \tau) \} - \tau), \quad j = 1, ..., k$$

and

$$\mathbb{G}_{n,h}(\tau) := -\|\mathbf{w}_n\|^{-1} n^{-1/2} (\mathbf{0}_k^\top, \mathbf{w}_n^\top) J_m(\tau)^{-1} \sum_{i=1}^n \mathbf{Z}_i (\mathbf{1}\{Y_i \le Q(X_i; \tau)\} - \tau).$$

We need to establish tightness and finite dimensional convergence. By Lemma 1.4.3 of van der Vaart and Wellner (1996), it is enough to show the tightness of  $\mathbb{G}_{n,j}$ 's and  $\mathbb{G}_{n,h}$  individually. Tightness follows from asymptotic equicontinuity which can be proved by an application of Lemma A.3. More precisely, apply Lemma A.3 with  $\mathbf{u}_n = \mathbf{e}_j$  to prove tightness of  $\mathbb{G}_{n,j}(\cdot)$  for j = 1, ..., k, and Lemma A.3 with  $\mathbf{u}_n^{\top} = (\mathbf{0}_k^{\top}, \mathbf{w}_n^{\top})$  to prove tightness of  $\mathbb{G}_{n,h}(w_0; \cdot)$ . Continuity of the sample paths of  $\mathbb{G}_{n,h}(w_0; \cdot)$  follows by the same arguments as given at the beginning of Section A.1.2.

Next, we prove finite-dimensional convergence. Observe the decomposition

$$\mathbf{G}_n(\tau) = -n^{-1/2} \sum_{i=1}^n \left\{ \begin{pmatrix} M_1(\tau)^{-1} (V_i - A(\tau) \widetilde{\mathbf{Z}}(W_i)) \\ \|\mathbf{w}_n\|^{-1} \mathbf{w}_n^\top M_2(\tau)^{-1} \widetilde{\mathbf{Z}}(W_i) \end{pmatrix} \left( \mathbf{1} \{ Y_i \le Q(X_i; \tau) \} - \tau \right) + \begin{pmatrix} 0 \\ \varphi_i(\tau) \end{pmatrix} \right\}$$

where

$$\varphi_i(\tau) := -\|\mathbf{w}_n\|^{-1}\mathbf{w}_n^{\mathsf{T}}A(\tau)^{\mathsf{T}}M_1(\tau)^{-1}(V_i - A(\tau)\widetilde{\mathbf{Z}}(W_i))(\mathbf{1}\{Y_i \le Q(X_i; \tau)\} - \tau).$$

By definition, we have  $\mathbb{E}[\varphi_i(\tau)] = 0$  and moreover

$$\mathbb{E}[\varphi_i^2(\tau)] \lesssim \|\mathbf{w}_n\|^{-2} \mathbf{w}_n^{\mathsf{T}} A(\tau)^{\mathsf{T}} M_1(\tau)^{-1} \mathbb{E}[(V_i - A(\tau)\widetilde{\mathbf{Z}}(W_i))(V_i - A(\tau)\widetilde{\mathbf{Z}}(W_i))^{\mathsf{T}}] M_1(\tau)^{-1} A(\tau) \mathbf{w}_n.$$

Since  $f_{Y|X}(Q(X_i;\tau)|X)$  is bounded away from zero uniformly, it follows that

$$\|\mathbb{E}[(V_i - A(\tau)\widetilde{\mathbf{Z}}(W_i))(V_i - A(\tau)\widetilde{\mathbf{Z}}(W_i))^{\top}]\| \le f_{\min}\lambda_{\max}(M_1(\tau)) < \infty,$$

by Remark 5.5. Moreover, by Lemma A.2 proven later,  $||A(\tau)\mathbf{w}_n|| = O(1)$  uniformly in  $\tau$ , and thus by  $||\mathbf{w}_n|| \to \infty$ ,  $\sup_{\tau \in \mathcal{T}} \mathbb{E}[\varphi_i^2(\tau)] = o(1)$ . This implies that  $n^{-1/2} \sum_{i=1}^n \varphi_i(\tau) = o_P(1)$  for every fixed  $\tau \in \mathcal{T}$ . Hence it suffices to prove finite dimensional convergence of

$$-n^{-1/2} \sum_{i=1}^{n} \left( \frac{M_1(\tau)^{-1} (V_i - A(\tau) \widetilde{\mathbf{Z}}(W_i))}{\|\mathbf{w}_n\|^{-1} \mathbf{w}_n^{\top} M_2(\tau)^{-1} \widetilde{\mathbf{Z}}(W_i)} \right) \left( \mathbf{1} \{ Y_i \le Q(X_i; \tau) \} - \tau \right).$$

Observe that  $\mathbb{E}[M_1(\tau)^{-1}(h_{VW}(W_i;\tau)-A(\tau)\widetilde{\mathbf{Z}}(W_i))(\mathbf{1}\{Y_i\leq Q(X_i;\tau)\}-\tau)]=0$  and by assumptions (A1)-(A3), (C1)

$$\sup_{\tau \in \mathcal{T}} \mathbb{E}[\|M_1(\tau)^{-1}(h_{VW}(W;\tau) - A(\tau)\widetilde{\mathbf{Z}}(W))\|^2] 
\leq \sup_{\tau \in \mathcal{T}} \frac{1}{f_{\min}\lambda_{\min}(M_1(\tau))} \mathbb{E}[\|f_{Y|X}(Q(X;\tau)|X)(h_{VW}(w;\tau) - A(\tau)\widetilde{\mathbf{Z}}(W))\|^2] = o(1).$$

Thus,  $n^{-1/2} \sum_{i=1}^{n} M_1(\tau)^{-1} (h_{VW}(W_i; \tau) - A(\tau) \widetilde{\mathbf{Z}}(W_i)) (\mathbf{1}\{Y_i \leq Q(X_i; \tau)\} - \tau) = o_P(1)$  for every fixed  $\tau \in \mathcal{T}$ . So, now we only need to consider finite dimensional convergence of

$$\sum_{i=1}^{n} \psi_i(\tau) := -n^{-1/2} \sum_{i=1}^{n} \left( \frac{M_1(\tau)^{-1} (V_i - h_{VW}(W_i; \tau))}{\|\mathbf{w}_n\|^{-1} \mathbf{w}_n^{\top} M_2(\tau)^{-1} \widetilde{\mathbf{Z}}(W_i)} \right) \left( \mathbf{1} \{ Y_i \le Q(X_i; \tau) \} - \tau \right). \tag{A.24}$$

Note that

$$\mathbb{E}[\psi_i(\tau_1)\psi_i(\tau_2)^{\top}] = (\tau_1 \wedge \tau_2 - \tau_1 \tau_2) \begin{pmatrix} \Gamma_{11}(\tau_1, \tau_2) + o(1) & \Gamma_{12}(\tau_1, \tau_2) \\ \Gamma_{12}(\tau_1, \tau_2)^{\top} & \Gamma_{22}(\tau_1, \tau_2) + o(1) \end{pmatrix},$$

where  $\Gamma_{12}(\tau_1, \tau_2) := \|\mathbf{w}_n\|^{-1} \mathbb{E}[M_1(\tau_2)^{-1}(V_i - h_{VW}(W_i; \tau_1))\mathbf{w}_n^{\top} M_2(\tau_2)^{-1} \widetilde{\mathbf{Z}}(W_i)]$ . We shall now show that  $\Gamma_{12}(\tau_1, \tau_2) = o(1)$  uniformly in  $\tau_1, \tau_2$ . Note that from the definition of  $h_{VW}(W; \tau)$  in (3.4), by standard argument we can write

$$h_{VW}(W;\tau) = \mathbb{E}[f(Q(X;\tau)|X)|W]^{-1}\mathbb{E}[Vf(Q(X;\tau)|X)|W].$$
 (A.25)

From (A3) we obtain  $\mathbb{E}[f(Q(X;\tau)|X)|W] \geq f_{\min} > 0$ , and from (C1), (A2) it follows that  $|\mathbb{E}[V^{(j)}f(Q(X;\tau)|X)|W]| \leq C\bar{f} = O(1)$ , i.e. the components of  $h_{VW}(W;\tau)$  are bounded by a constant almost surely. Hence,

$$\|\Gamma_{12}(\tau_{1}, \tau_{2})\| = \|\mathbf{w}_{n}\|^{-1} \|M_{1}(\tau_{1})^{-1} \mathbb{E}[(V_{i} - h_{VW}(W_{i}; \tau)) \mathbf{w}_{n}^{\top} M_{2}(\tau_{2})^{-1} \widetilde{\mathbf{Z}}(W_{i})]\|$$

$$\leq \|\mathbf{w}_{n}\|^{-1} \|M_{1}(\tau_{1})^{-1}\| \|\mathbb{E}[(V_{i} - h_{VW}(W_{i}; \tau_{1})) \mathbf{w}_{n}^{\top} M_{2}(\tau_{2})^{-1} \widetilde{\mathbf{Z}}(W_{i})]\|$$

$$\lesssim \|\mathbf{w}_{n}\|^{-1} \mathbb{E}[\|V_{i} - h_{VW}(W_{i}; \tau_{1})\| |\mathbf{w}_{n}^{\top} M_{2}(\tau_{2})^{-1} \widetilde{\mathbf{Z}}(W_{i})|]$$

$$\lesssim \|\mathbf{w}_{n}\|^{-1} \mathbb{E}[|\mathbf{w}_{n}^{\top} M_{2}(\tau_{2})^{-1} \widetilde{\mathbf{Z}}(W_{i})|]$$

$$= o(1),$$

where the third inequality applies the lower bound for  $\inf_{\tau \in \mathcal{T}} \lambda_{\min}(M_1(\tau))$  in Remark 5.5; the fourth inequality follows from  $\sup_{1 \leq j \leq k, \tau} |V^{(j)}| + |h_{VW}^{(j)}(W;\tau)| < \infty$  a.s., while the last equality follows by the assumptions of the Theorem.

Now we prove the finite dimensional convergence (A.24). Taking arbitrary collections  $\{\tau_1, ..., \tau_J\} \subset \mathcal{T}$ ,  $\mathbf{c}_1, ..., \mathbf{c}_J \in \mathbb{R}^{k+1}$ , we need to show that

$$\sum_{j=1}^J \mathbf{c}_j^\top \mathbf{G}_n(\tau_j) = \sum_{j=1}^J \sum_{i=1}^n \mathbf{c}_j^\top \boldsymbol{\psi}_i(\tau_j) \overset{d}{\to} \sum_{j=1}^J \mathbf{c}_j^\top \mathbf{G}(\tau_j).$$

Define  $V_{i,J} = \sum_{j=1}^{J} \mathbf{c}_j^{\top} \psi_i(\tau_j)$ . Note that  $\mathbb{E}[V_{i,J}] = 0$  and  $|V_{i,J}| \lesssim n^{-1/2} \xi_m$ . Using the results derived above, we have

$$\operatorname{var}(V_{i,J}) = o(n^{-1}) + n^{-1} \sum_{j,j'=1}^{J} (\tau_j \wedge \tau_{j'} - \tau_j \tau_{j'}) \mathbf{c}_j^{\top} \Gamma(\tau_j, \tau_{j'}) \mathbf{c}_{j'},$$

where  $\Gamma(\tau_j, \tau_{j'})$  is defined as (3.11). If  $\sum_{j,j'=1}^{J} (\tau_j \wedge \tau_{j'} - \tau_j \tau_{j'}) \mathbf{c}_j^{\top} \Gamma(\tau_j, \tau_{j'}) \mathbf{c}_{j'} = 0$ , then the distribution of  $\sum_{j=1}^{J} \mathbf{c}_j^{\top} \mathbf{G}(\tau_j)$  is a single point mass at 0, and  $\sum_{j=1}^{J} \mathbf{c}_j^{\top} \mathbf{G}_n(\tau_j) = \sum_{j=1}^{J} \sum_{i=1}^{n} \mathbf{c}_j^{\top} \psi_i(\tau_j)$  converges to 0 in probability by Markov's inequality.

If  $n^{-1} \sum_{j,j'=1}^{J} (\tau_j \wedge \tau_{j'} - \tau_j \tau_{j'}) \mathbf{c}_j^{\top} \Gamma(\tau_j, \tau_{j'}) \mathbf{c}_{j'} > 0$  define  $s_{n,J}^2 = \sum_{i=1}^n \text{var}(V_{i,J})$ . We will now verify that the triangular array of random variables  $(V_{i,J})_{i=1,\dots,n}$  satisfies the Lindeberg condition. For any v > 0 and sufficiently large n, Markov's inequality gives

$$s_{n,J}^{-2} \sum_{i=1}^{n} \mathbb{E} \big[ V_{i,J}^2 I(V_{i,J} \ge v) \big] \lesssim \xi_m^2 s_{n,J}^{-2} \mathbb{E} \big[ I(V_{1,J} \ge v) \big] \lesssim \xi_m^2 s_{n,J}^{-2} v^{-2} n^{-1} s_{n,J}^2,$$

where  $\xi_m^2 n^{-1} = o(1)$  by (B1'). Thus the Lindeberg condition holds and it follows that

$$\frac{\sum_{i=1}^{n} V_{i,J}}{s_{n,J}} \stackrel{d}{\to} \mathcal{N}(0,1).$$

Finally, it remains to prove (A.20) and (A.21). Begin by observing that

$$\begin{split} & \left\| \mathbb{E} \left[ (V - A(\tau) \widetilde{\mathbf{Z}}(W)) \left( F_{Y|X}(Q(X;\tau)|X) - F_{Y|X}(\boldsymbol{\alpha}(\tau)^{\top}V + h_{n}^{\dagger}(W,\tau)|X) \right) \right] \right\| \\ & \lesssim \left\| \mathbb{E} \left[ (V - A(\tau) \widetilde{\mathbf{Z}}(W)) f_{Y|X}(Q(X_{i};\tau)|X) \left( h(W,\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top} \widetilde{\mathbf{Z}}(W) \right) \right] \right\| \\ & + \left\| \mathbb{E} \left[ (V - A(\tau) \widetilde{\mathbf{Z}}(W)) f_{Y|X}^{\prime} (\bar{Q}^{\dagger}(X_{i},\tau)|X) \left( h(W,\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top} \widetilde{\mathbf{Z}}(W) \right)^{2} \right] \right\| \\ & \lesssim \left\| \mathbb{E} \left[ (h_{VW}(W,\tau) - A(\tau) \widetilde{\mathbf{Z}}(W)) f_{Y|X}(Q(X_{i};\tau)|X) \left( h(W,\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top} \widetilde{\mathbf{Z}}(W) \right) \right] \right\| + c_{n}^{\dagger 2} \\ & \lesssim c_{n}^{\dagger} \mathbb{E} \left[ \left\| h_{VW}(W,\tau) - A(\tau) \widetilde{\mathbf{Z}}(W) \right\| f_{Y|X}(Q(X_{i};\tau)|X) \right] + c_{n}^{\dagger 2} \\ & = O(c_{n}^{\dagger}\lambda_{n} + c_{n}^{\dagger 2}), \end{split} \tag{A.26}$$

where the first inequality is an application of Taylor expansion, with  $\bar{Q}^{\dagger}(X_i, \tau)$  lying on the line segment of  $Q(X;\tau)$  and  $\boldsymbol{\alpha}(\tau)^{\top}V_i + h_n^{\dagger}(W_i,\tau)$ ; the second inequality is the result of the orthogonality condition (3.5), Condition (A2), Conditions (C1)-(C2); the third inequality follows from Conditions (A2), (C1), and the last line follows from condition (C1) and the Hölder inequality. For a proof of (A.20) observe that by (5.10)  $\mathbf{e}_j^{\top} J_m(\tau)^{-1} \mathbf{Z} = \mathbf{e}_j^{\top} M_1(\tau)^{-1} (V - A(\tau) \widetilde{\mathbf{Z}}(W))$  for j = 1, ..., k. Thus we obtain from Remark 5.5

$$\begin{split} & \left| \mathbb{E}[\mathbf{e}_{j}^{\top} J_{m}(\tau)^{-1} \mathbf{Z} \left( \mathbf{1} \{ Y \leq Q(X; \tau) \} - \mathbf{1} \{ Y \leq \boldsymbol{\alpha}(\tau)^{\top} V + h_{n}^{\dagger}(W, \tau) \} \right) \right] \right| \\ & = \left| \mathbf{e}_{j}^{\top} M_{1}(\tau)^{-1} \mathbb{E} \left[ (V - A(\tau) \widetilde{\mathbf{Z}}(W)) \left( F_{Y|X}(Q(X; \tau)|X) - F_{Y|X}(\boldsymbol{\alpha}(\tau)^{\top} V + h_{n}^{\dagger}(W, \tau)|X) \right) \right] \right| \\ & = O(c_{n}^{\dagger} \lambda_{n} + c_{n}^{\dagger 2}), \end{split}$$

To prove (A.21), without loss of generality, let  $\|\mathbf{w}_n\| = 1$ . We note that by (5.10)

$$(\mathbf{0}_k^\top, \mathbf{w}_n^\top) J_m(\tau)^{-1} \mathbf{Z}_i = -\mathbf{w}_n^\top A(\tau)^\top M_1(\tau)^{-1} (V_i - A(\tau) \widetilde{\mathbf{Z}}(W_i)) + \mathbf{w}_n^\top M_2(\tau)^{-1} \widetilde{\mathbf{Z}}(W_i).$$

From (A.26), (5.11) in Remark 5.5 and Lemma A.2 we obtain

$$\left| \mathbb{E} \left[ \mathbf{w}_n^{\top} A(\tau)^{\top} M_1(\tau)^{-1} (V_i - A(\tau) \widetilde{\mathbf{Z}}(W_i)) \left( \mathbf{1} \{ Y_i \le Q(X_i; \tau) \} - \mathbf{1} \{ Y_i \le \boldsymbol{\alpha}(\tau)^{\top} V_i + h_n^{\dagger}(W_i, \tau) \} \right) \right] \right|$$

$$= O(c_n^{\dagger} \lambda_n + c_n^{\dagger 2}). \tag{A.27}$$

Moreover,

$$\begin{aligned}
&\left| \mathbb{E} \left[ \mathbf{w}_{n}^{\top} M_{2}(\tau)^{-1} \widetilde{\mathbf{Z}}(W_{i}) \left( \mathbf{1} \left\{ Y_{i} \leq Q(X_{i}; \tau) \right\} - \mathbf{1} \left\{ Y_{i} \leq \boldsymbol{\alpha}(\tau)^{\top} V_{i} + h_{n}^{\dagger}(W_{i}, \tau) \right\} \right) \right] \right| \\
&\lesssim \left| \mathbf{w}_{n}^{\top} M_{2}(\tau)^{-1} \mathbb{E} \left[ \widetilde{\mathbf{Z}}(W_{i}) f_{Y|X}(Q(X_{i}; \tau)|X) \left( h(W, \tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top} \widetilde{\mathbf{Z}}(W) \right) \right] \right| \\
&+ \left| \mathbf{w}_{n}^{\top} M_{2}(\tau)^{-1} \mathbb{E} \left[ \widetilde{\mathbf{Z}}(W_{i}) f_{Y|X}^{\prime} (\bar{Q}^{\dagger}(X_{i}, \tau)|X) \left( h(W, \tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top} \widetilde{\mathbf{Z}}(W) \right)^{2} \right\} \right] \right| \\
&= \left| \mathbb{E} \left[ \mathbf{w}_{n}^{\top} M_{2}(\tau)^{-1} \widetilde{\mathbf{Z}}(W_{i}) f_{Y|X}^{\prime} (\bar{Q}^{\dagger}(X_{i}, \tau)|X) \left( h(W, \tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top} \widetilde{\mathbf{Z}}(W) \right)^{2} \right\} \right] \right| \\
&\lesssim \mathbb{E} \left[ \left| \mathbf{w}_{n}^{\top} M_{2}(\tau_{2})^{-1} \widetilde{\mathbf{Z}}(W_{i}) \right| \right] c_{n}^{\dagger 2} = o(\|\mathbf{w}_{n}\| c_{n}^{\dagger 2}), \tag{A.28}
\end{aligned}$$

where the first inequality follows from the Taylor expansion with  $\bar{Q}^{\dagger}(X_i, \tau)$  lying on the line segment of  $Q(X;\tau)$  and  $\alpha(\tau)^{\top}V_i + h_n^{\dagger}(W_i, \tau)$ ; the second equality follows from the first order condition of (3.3), and the last line follows Conditions (A2), (C1) and the conditions of the theorem. Combining (A.27) and (A.28), and (B1') we obtain (A.21).

**Lemma A.2.** Under Assumptions (A1)-(A3), (C2) and  $\sup_{\tau \in \mathcal{T}} \mathbb{E}[|\widetilde{\mathbf{Z}}(W)^{\top} M_2(\tau)^{-1} \mathbf{w}_n|] = o(||\mathbf{w}_n||)$ , we have

$$\sup_{\tau} ||A(\tau)\mathbf{w}_n||^2 = o(||\mathbf{w}_n||)$$
(A.29)

**Proof for Lemma A.2.** By the first order condition for obtaining  $A(\tau)$ ,

$$A(\tau) = \mathbb{E}[h_{VW}(W;\tau)\widetilde{\mathbf{Z}}(W)^{\top}f_{Y|X}(Q(X;\tau)|X)]M_2(\tau)^{-1}.$$

By the orthogonality condition (3.5),

$$||A(\tau)\mathbf{w}_n|| = ||\mathbb{E}[h_{VW}(W;\tau)\widetilde{\mathbf{Z}}(W)^{\top}f_{Y|X}(Q(X;\tau)|X)]M_2(\tau)^{-1}\mathbf{w}_n||$$

$$= ||\mathbb{E}[(h_{VW}(W;\tau) - V + V)\widetilde{\mathbf{Z}}(W)^{\top}f_{Y|X}(Q(X;\tau)|X)]M_2(\tau)^{-1}\mathbf{w}_n||$$

$$= ||\mathbb{E}[V\widetilde{\mathbf{Z}}(W)^{\top}f_{Y|X}(Q(X;\tau)|X)]M_2(\tau)^{-1}\mathbf{w}_n||.$$

By the assumption that at fixed j,  $|V_j| \leq C$ , the uniform boundedness of the conditional density in (A2), and the hypothesis  $\sup_{\tau \in \mathcal{T}} \mathbb{E}[|\widetilde{\mathbf{Z}}(W)^{\top} M_2(\tau)^{-1} \mathbf{w}_n|] = o(\|\mathbf{w}_n\|)$ ,

$$\sup_{\tau \in \mathcal{T}} \mathbb{E}[|V_j \widetilde{\mathbf{Z}}(W)^\top M_2(\tau)^{-1} \mathbf{w}_n | f_{Y|X}(Q(X;\tau)|X)] \le \bar{f} C \sup_{\tau \in \mathcal{T}} \mathbb{E}[|\widetilde{\mathbf{Z}}(W)^\top M_2(\tau)^{-1} \mathbf{w}_n|] = o(\|\mathbf{w}_n\|). \tag{A.30}$$

This completes the proof of (A.29) by noting that  $\sup_{\tau \in \mathcal{T}} \|M_2(\tau)^{-1}\| = O(1)$ .

# A.4 Asymptotic Tightness of Quantile Process

In this section we establish the asymptotic tightness of the process  $n^{1/2}\mathbf{u}_n^{\mathsf{T}}U_n(\tau)$  in  $\ell^{\infty}(\mathcal{T})$  with  $\mathbf{u}_n \in \mathbb{R}^m$  being an arbitrary vector, where

$$U_n(\tau) = n^{-1} J_m^{-1}(\tau) \sum_{i=1}^n \mathbf{Z}_i (\mathbf{1}\{Y_i \le Q(X_i; \tau)\} - \tau).$$
(A.31)

Note that the results obtained in this section, in particular Lemma A.3, apply to any series expansion  $\mathbf{Z} = \mathbf{Z}(X_i)$  satisfying Assumptions (A1).

The following definition is only used in this subsection: For any non-decreasing, convex function  $\Psi: \mathbb{R}^+ \to \mathbb{R}^+$  with  $\Phi(0) = 0$ , the *Orlicz norm* of a real-valued random variable Z is defined as (see e.g. Chapter 2.2 of van der Vaart and Wellner (1996))

$$||Z||_{\Psi} = \inf \{C > 0 : \mathbb{E}\Phi(|Z|/C) \le 1\}.$$
 (A.32)

**Lemma A.3** (Asymptotic Equicontinuity of Quantile Process). Under (A1)-(A3) and  $\xi_m^2(\log n)^2 = o(n)$ , we have for any  $\varepsilon > 0$  and vector  $\mathbf{u}_n \in \mathbb{R}^m$ ,

$$\lim_{\delta \to 0} \limsup_{n \to \infty} P\left( \|\mathbf{u}_n\|^{-1} n^{1/2} \sup_{\tau_1, \tau_2 \in \mathcal{T}, |\tau_1 - \tau_2| \le \delta} \left| \mathbf{u}_n^\top \mathbf{U}_n(\tau_1) - \mathbf{u}_n^\top \mathbf{U}_n(\tau_2) \right| > \varepsilon \right) = 0, \tag{A.33}$$

where  $U_n(\tau)$  is defined in (A.31).

**Proof of Lemma A.3.** Without loss of generality, we will assume that  $\mathbf{u}_n$  is a sequence of vectors with  $\|\mathbf{u}_n\| = 1$ , which can always be achieved by rescaling. Define

$$\mathbb{G}_n(\tau) := n^{1/2} \mathbf{u}_n^{\top} \mathbf{U}_n(\tau).$$

Consider the decomposition

$$\mathbf{u}_{n}^{\top} \mathbf{U}_{n}(\tau_{1}) - \mathbf{u}_{n}^{\top} \mathbf{U}_{n}(\tau_{2}) = n^{-1} \mathbf{u}_{n}^{\top} (J_{m}^{-1}(\tau_{1}) - J_{m}^{-1}(\tau_{2})) \sum_{i} \mathbf{Z}_{i} (\mathbf{1} \{ Y_{i} \leq Q(X_{i}; \tau_{1}) \} - \tau_{1})$$
$$+ n^{-1} \mathbf{u}_{n}^{\top} J_{m}^{-1}(\tau_{2}) \sum_{i} \mathbf{Z}_{i} \Delta_{i}(\tau_{1}, \tau_{2}),$$

where

$$\Delta_i(\tau_1, \tau_2) := \mathbf{1}\{Y_i \le Q(X_i; \tau_1)\} - \tau_1 - (\mathbf{1}\{Y_i \le Q(X_i; \tau_2)\} - \tau_2).$$

Note that for any  $L \geq 2$ ,

$$\mathbb{E}\left[\left|\mathbf{u}_{n}^{\top}J_{m}^{-1}(\tau_{2})\mathbf{Z}_{i}\Delta_{i}(\tau_{1},\tau_{2})\right|^{L}\right] \lesssim \xi_{m}^{L-2}\mathbb{E}\left[\left|\mathbf{u}_{n}^{\top}J_{m}^{-1}(\tau_{2})\mathbf{Z}_{i}\Delta_{i}(\tau_{1},\tau_{2})\right|^{2}\right] \\
= \xi_{m}^{L-2}\mathbf{u}_{n}^{\top}J_{m}^{-1}(\tau_{2})\mathbb{E}\left[\mathbf{Z}_{i}\mathbf{Z}_{i}^{\top}\Delta_{i}(\tau_{1},\tau_{2})^{2}\right]J_{m}^{-1}(\tau_{2})\mathbf{u}_{n} \\
\lesssim \xi_{m}^{L-2}|\tau_{1}-\tau_{2}|. \tag{A.34}$$

By the Lipschitz continuity of  $\tau \mapsto J_m^{-1}(\tau)$  (cf. Lemma 13 of Belloni et al. (2011)) and positive definiteness of  $J_m^{-1}(\tau)$ , we have

$$||J_{m}^{-1}(\tau_{1}) - J_{m}^{-1}(\tau_{2})|| = ||J_{m}^{-1}(\tau_{2})\{J_{m}(\tau_{1}) - J_{m}(\tau_{2})\}J_{m}^{-1}(\tau_{1})||$$

$$\leq \frac{\bar{f}'}{f_{\min}}|\tau_{1} - \tau_{2}|\Big(\inf_{\tau \in \mathcal{T}}\lambda_{\min}(J_{m}(\tau))\Big)^{-2}\lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}]),$$

where  $\|\cdot\|$  denotes the operator norm of a matrix. Thus, we have for  $L \geq 2$ ,

$$\mathbb{E}\Big[\Big|\mathbf{u}_{n}^{\top}\Big\{J_{m}^{-1}(\tau_{1}) - J_{m}^{-1}(\tau_{2})\Big\}\mathbf{Z}_{i}(\mathbf{1}(Y_{i} \leq Q(X_{i};\tau)) - \tau)\Big|^{L}\Big] 
\lesssim \xi_{m}^{L-2}\mathbb{E}\Big[\Big|\mathbf{u}_{n}^{\top}\Big\{J_{m}^{-1}(\tau_{1}) - J_{m}^{-1}(\tau_{2})\Big\}\mathbf{Z}_{i}(\mathbf{1}(Y_{i} \leq Q(X_{i};\tau)) - \tau)\Big|^{2}\Big] 
\leq \xi_{m}^{L-2}\mathbb{E}\Big[\Big|\mathbf{u}_{n}^{\top}\Big\{J_{m}^{-1}(\tau_{1}) - J_{m}^{-1}(\tau_{2})\Big\}\mathbf{Z}_{i}\Big|^{2}\Big] 
= \xi_{m}^{L-2}\mathbf{u}_{n}^{\top}\Big\{J_{m}^{-1}(\tau_{1}) - J_{m}^{-1}(\tau_{2})\Big\}\mathbb{E}\Big[\mathbf{Z}_{i}\mathbf{Z}_{i}^{\top}\Big]\Big\{J_{m}^{-1}(\tau_{1}) - J_{m}^{-1}(\tau_{2})\Big\}\mathbf{u}_{n} 
\leq \xi_{m}^{L-2}\Big\|\Big\{J_{m}^{-1}(\tau_{1}) - J_{m}^{-1}(\tau_{2})\Big\}\mathbb{E}\Big[\mathbf{Z}_{i}\mathbf{Z}_{i}^{\top}\Big]\Big\{J_{m}^{-1}(\tau_{1}) - J_{m}^{-1}(\tau_{2})\Big\}\Big\| 
\lesssim \xi_{m}^{L-2}|\tau_{1} - \tau_{2}|^{2}.$$
(A.35)

To simplify notations, define

$$\widetilde{V}_{n,i}(\tau_1,\tau_2) := \mathbf{u}_n^{\top} \big\{ J_m^{-1}(\tau_1) - J_m^{-1}(\tau_2) \big\} \mathbf{Z}_i(\mathbf{1}(Y_i \leq Q(X_i;\tau)) - \tau) + \mathbf{u}_n^{\top} J_m^{-1}(\tau_2) \mathbf{Z}_i \Delta_i(\tau_1,\tau_2).$$

Combining the bounds (A.34) and (A.35) yields

$$\mathbb{E}\left[|\widetilde{V}_{n,i}(\tau_{1},\tau_{2})|^{2L}\right]^{1/2L} \\
\leq \mathbb{E}\left[|\mathbf{u}_{n}^{\top}J_{m}^{-1}(\tau_{2})\mathbf{Z}_{i}\Delta_{i}(\tau_{1},\tau_{2})|^{2L}\right]^{1/2L} + \mathbb{E}\left[|\mathbf{u}_{n}^{\top}\left\{J_{m}^{-1}(\tau_{1}) - J_{m}^{-1}(\tau_{2})\right\}\mathbf{Z}_{i}(\mathbf{1}(Y_{i} \leq Q(X_{i};\tau)) - \tau)|^{2L}\right]^{1/2L} \\
\lesssim \left(\xi_{m}^{2(L-1)}|\tau_{1} - \tau_{2}|\right)^{1/2L}. \tag{A.36}$$

Note that (A.36) holds for all positive integers  $L \geq 1$ . By the fact that  $\mathbb{G}_n(\tau_1) - \mathbb{G}_n(\tau_2) = n^{-1/2} \sum_{i=1}^n \widetilde{V}_{n,i}(\tau_1,\tau_2)$  and  $\mathbb{E}\widetilde{V}_{n,i}(\tau_1,\tau_2) = 0$ , we obtain from (A.36) that

$$\begin{split} &\mathbb{E}[|\mathbb{G}_{n}(\tau_{1}) - \mathbb{G}_{n}(\tau_{2})|^{2L}] \\ &= n^{-L}\mathbb{E}\left[\left(\sum_{i=1}^{n} \widetilde{V}_{n,i}(\tau_{1},\tau_{2})\right)^{2L}\right] \\ &= n^{-L}\mathbb{E}\left[\left(\sum_{i=1}^{n} \widetilde{V}_{n,i}(\tau_{1},\tau_{2})^{2L}\right) + \sum_{l=1}^{L-1} \sum_{\substack{1 \leq i_{1},i_{2} \leq n \\ i_{1} \neq i_{2}}} \mathbb{E}\left[\widetilde{V}_{n,i_{1}}(\tau_{1},\tau_{2})^{2L-2l}\right] \mathbb{E}\left[\widetilde{V}_{n,i_{2}}(\tau_{1},\tau_{2})^{2l}\right] \\ &\quad + \sum_{\substack{l_{1}+l_{2} < L \\ l_{1}=1,l_{2}=1}} \sum_{\substack{1 \leq i_{1},i_{2},i_{3} \leq n \\ i_{1} \neq i_{2} \neq i_{3}}} \mathbb{E}\left[\widetilde{V}_{n,i_{1}}(\tau_{1},\tau_{2})^{2L-2(l_{1}+l_{2})}\right] \mathbb{E}\left[\widetilde{V}_{n,i_{2}}(\tau_{1},\tau_{2})^{2l_{1}}\right] \mathbb{E}\left[\widetilde{V}_{n,i_{3}}(\tau_{1},\tau_{2})^{2l_{2}}\right] \\ &\quad + \ldots + \sum_{\substack{1 \leq i_{1},\ldots,i_{L} \leq n \\ i_{1} \neq \ldots \neq i_{L}}} \prod_{j=1}^{L} \mathbb{E}\left[\widetilde{V}_{n,i_{j}}(\tau_{1},\tau_{2})^{2}\right]\right) \\ &\leq C_{L}n^{-L}\left(n\xi_{m}^{2(L-1)}|\tau_{1}-\tau_{2}| + \binom{n}{2}\xi_{m}^{2(L-1-1)}|\tau_{1}-\tau_{2}|^{2} + \binom{n}{3}\xi_{m}^{2(L-1-2)}|\tau_{1}-\tau_{2}|^{3} + \ldots + \binom{n}{L}|\tau_{1}-\tau_{2}|^{L}\right) \\ &\lesssim \sum_{k=0}^{L-1}\frac{\xi_{m}^{2(L-k-1)}}{n(L-k-1)}|\tau_{1}-\tau_{2}|^{k+1}. \end{split}$$

In particular we obtain for  $|\tau_1 - \tau_2| \ge \xi_m^2/n$ ,

$$\mathbb{E}[|\mathbb{G}_n(\tau_1) - \mathbb{G}_n(\tau_2)|^{2L}] \lesssim |\tau_1 - \tau_2|^L. \tag{A.37}$$

For  $\Psi(z)=z^{2L}$ , the above equation implies that the Orlicz norm (defined in (A.32)) of  $\mathbb{G}_n(\tau_1)$  –  $\mathbb{G}(\tau_2)$  satisfies

$$\|\mathbb{G}_n(\tau_1) - \mathbb{G}(\tau_2)\|_{\Psi} \lesssim |\tau_1 - \tau_2|^{1/2}$$
.

Let  $d(\tau, \tau') = \sqrt{|\tau - \tau'|}$ , which is a metric on  $\mathcal{T}$ . The packing number  $D(\epsilon, d)$  of  $\mathcal{T}$  with respect to d satisfies  $D(\epsilon, d) \lesssim 1/\epsilon^2$ . Let  $\bar{\omega}_n = 2\xi_m/\sqrt{n} \to 0$  as  $n \to \infty$ . We have

$$\int_{\bar{\omega}_n/2}^{\omega} \Psi^{-1}(D(\epsilon, d)) d\epsilon \lesssim \int_{\bar{\omega}_n/2}^{\omega} \epsilon^{-1/L} d\epsilon = \frac{\omega^{1 - L^{-1}}}{1 - L^{-1}} - \frac{(\bar{\omega}_n/2)^{1 - L^{-1}}}{1 - L^{-1}}.$$
(A.38)

For  $\omega > 0$ ,

$$\Psi^{-1}(D^2(\omega, d)) \lesssim \Psi^{-1}\left(\frac{1}{\omega^4}\right) = \omega^{-2/L}.$$
(A.39)

Therefore, applying Lemma S.2.1 yields for any  $\delta > 0$ 

$$\sup_{|\tau_1 - \tau_2| \le \delta} |\mathbb{G}_n(\tau_1) - \mathbb{G}_n(\tau_2)| = \sup_{|\tau_1 - \tau_2|^{1/2} \le \delta^{1/2}} |\mathbb{G}_n(\tau_1) - \mathbb{G}_n(\tau_2)|$$

$$\le S_{1,n}(\delta) + 2 \sup_{|\tau' - \tau|^{1/2} \le \bar{\omega}_n, \tau \in \widetilde{\mathcal{T}}} |\mathbb{G}_n(\tau') - \mathbb{G}_n(\tau)|, \tag{A.40}$$

where  $\widetilde{\mathcal{T}} \subset \mathcal{T}$  has at most  $D(\bar{\omega}_n, d) \lesssim \bar{\omega}_n^{-2}$  points and  $S_{1,n}(\delta)$  is a random variable that satisfies

$$P(|S_{1,n}(\delta)| > z) \le \left(z \left[8K \left(\int_{\bar{\omega}_n/2}^{\omega} \Psi^{-1} \left(D(\epsilon, d)\right) d\epsilon + (\delta^{1/2} + 2\bar{\omega}_n) \Psi^{-1} \left(D^2(\omega, d)\right)\right)\right]^{-1}\right)^{-2L}$$

$$\lesssim \left(\frac{\omega^{1-L-1}}{\frac{1-L-1}{1-L-1}} - \frac{(\bar{\omega}_n/2)^{1-L-1}}{z} + (\delta^{1/2} + 2\bar{\omega}_n)\omega^{-2/L}}{z}\right)^{2L}.$$
(A.41)

for a constant K > 0. Let  $\omega = \delta$  and L = 6. As  $n \to \infty$ ,  $\omega > \bar{\omega}_n$ . We obtain  $\lim_{\delta \to 0} \limsup_{n \to \infty} P(|S_{1,n}(\delta)| > z) = 0$  for any z > 0.

To bound the remaining term in (A.40), observe that

$$\sup_{d(\tau,\tau')\leq \bar{\omega}_n,\tau\in\tilde{\mathcal{T}}} |\mathbb{G}_n(\tau') - \mathbb{G}_n(\tau)| = \sup_{|\tau-\tau'|\leq \bar{\omega}_n^2,\tau\in\tilde{\mathcal{T}}} |\mathbb{G}_n(\tau') - \mathbb{G}_n(\tau)| \leq \sup_{|\tau-\tau'|\leq \bar{\omega}_n^2,\tau,\tau'\in\mathcal{T}} |\mathbb{G}_n(\tau') - \mathbb{G}_n(\tau)|.$$

Now by Lemma A.4 we have

$$P\Big(\sup_{|\tau'-\tau|\leq \bar{\omega}_n^2, \tau, \tau'\in \mathcal{T}} |\mathbb{G}_n(\tau') - \mathbb{G}_n(\tau)| > r_n(\kappa_n)\Big) < e^{-\kappa_n},$$

where

$$r_n(\kappa_n) := C \left[ \bar{\omega}_n \log^{1/2} \frac{\xi_m}{\bar{\omega}_n} + \frac{\xi_m}{\sqrt{n}} \log \frac{\xi_m}{\bar{\omega}_n} + \kappa_n^{1/2} \bar{\omega}_n + \frac{\xi_m}{\sqrt{n}} \kappa_n \right], \tag{A.42}$$

for a sufficiently large constant C > 0. Take  $\kappa_n = \log n$ . Since  $\bar{\omega}_n(\log n)^{1/2} = 2\xi_m(\log n)^{1/2}/\sqrt{n} = o(1)$  and  $\xi_m \log(n)/\sqrt{n} = o(1)$  by assumption, it follows that  $r_n(\log n) \to 0$ . Therefore, we conclude from Lemma A.4 that

$$\sup_{d(\tau,\tau')\leq \bar{\omega}_n, \tau\in\widetilde{\mathcal{T}}} \left| \mathbb{G}_n(\tau') - \mathbb{G}_n(\tau) \right| = o_P(1). \tag{A.43}$$

Applying bounds (A.41) and (A.43) to (A.40) verifies the asymptotic equicontinuity

$$\lim_{\delta \to 0} \limsup_{n \to \infty} P \left( \sup_{|\tau_1 - \tau_2| \le \delta} |\mathbb{G}_n(\tau_1) - \mathbb{G}_n(\tau_2)| > z \right) = 0$$

for all 
$$z > 0$$
.

The following result is applied in the proof of Lemma A.3.

**Lemma A.4.** Under (A1)-(A3), we have for any  $\kappa_n > 0$ ,  $1/n \ll \delta < 1$ ,

$$P\left(\sup_{0 \le h \le \delta} \sup_{\tau \in [\epsilon, 1 - \epsilon - h]} |\mathbf{u}_n^\top \mathbf{U}_n(\tau + h) - \mathbf{u}_n^\top \mathbf{U}_n(\tau)| \ge Cr_n(\delta, \kappa_n)\right) \le 3e^{-\kappa_n}, \tag{A.44}$$

where  $\kappa_n > 0$ ,  $U_n(\tau)$  is defined in (A.31) and  $\mathbf{u}_n \in \mathbb{R}^m$  is arbitrary, and

$$r_n(\delta, \kappa_n) = \|\mathbf{u}_n\| \left(\frac{\delta}{n} \log \frac{\xi_m}{\sqrt{\delta}}\right)^{1/2} + \frac{\|\mathbf{u}_n\| \xi_m}{n} \log \frac{\xi_m}{\sqrt{\delta}} + \|\mathbf{u}_n\| \left(\frac{\kappa_n \delta}{n}\right)^{1/2} + \frac{\|\mathbf{u}_n\| \xi_m}{n} \kappa_n.$$

To prove Lemma A.4, we need to establish some preliminary results. For any fixed vector  $\mathbf{u} \in \mathbb{R}^m$  and  $\delta > 0$ , define the function classes

$$\mathcal{G}_{3}(\mathbf{u}) := \{ (\mathbf{Z}, Y) \mapsto \mathbf{u}^{\top} J_{m}(\tau)^{-1} \mathbf{Z} \mathbf{1} \{ \|\mathbf{Z}\| \leq \xi_{m} \} \big| \tau \in \mathcal{T} \}, 
\mathcal{G}_{4} := \{ (X, Y) \mapsto \mathbf{1} \{ Y_{i} \leq Q(X; \tau) \} - \tau \big| \tau \in \mathcal{T} \}, 
\mathcal{G}_{6}(\mathbf{u}, \delta) := \{ (\mathbf{Z}, Y) \mapsto \mathbf{u}^{\top} \{ J_{m}(\tau_{1})^{-1} - J_{m}(\tau_{2})^{-1} \} \mathbf{Z} \mathbf{1} \{ \|\mathbf{Z}\| \leq \xi_{m} \} \big| \tau_{1}, \tau_{2} \in \mathcal{T}, |\tau_{1} - \tau_{2}| \leq \delta \}, 
\mathcal{G}_{7}(\delta) := \{ (X, Y) \mapsto \mathbf{1} \{ Y_{i} \leq Q(X, \tau_{1}) \} - \mathbf{1} \{ Y_{i} \leq Q(X, \tau_{2}) \} - (\tau_{1} - \tau_{2}) \big| \tau_{1}, \tau_{2} \in \mathcal{T}, |\tau_{1} - \tau_{2}| \leq \delta \}.$$

Denote  $G_3$ ,  $G_6$  and  $G_7$  as the envelope functions of  $G_3$ ,  $G_6$  and  $G_7$ , respectively. The following covering number results will be shown in Section S.2.2: for any probability measure Q,

$$N(\epsilon \|G_3\|_{L_2(Q)}, \mathcal{G}_3(\mathbf{u}), L_2(Q)) \le \frac{C_0}{\epsilon}, \tag{A.45}$$

$$N(\epsilon \|G_6\|_{L_2(Q)}, \mathcal{G}_6(\mathbf{u}, \delta), L_2(Q)) \le 2\left(\frac{C_0}{\epsilon}\right)^2, \tag{A.46}$$

$$N(\epsilon ||G_7||_{L_2(Q)}, \mathcal{G}_7(\delta), L_2(Q)) \le \left(\frac{A_7}{\epsilon}\right)^2, \tag{A.47}$$

where  $C_0 := \frac{\bar{f}'}{f_{\min}} \frac{\lambda_{\max}(\mathbb{E}[\mathbf{ZZ}^{\top}])}{\inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_m(\tau))} < \infty$  given Assumptions (A1)-(A3), and  $A_7 > 0$  is a constant. Also,  $\mathcal{G}_4$  has VC index 2 according to Lemma S.2.4.

**Proof of Lemma A.4.** Observe the decomposition

$$\mathbf{u}_n^{\mathsf{T}} \boldsymbol{U}_n(\tau_1) - \mathbf{u}_n^{\mathsf{T}} \boldsymbol{U}_n(\tau_2) = I_1(\tau_1, \tau_2) + I_2(\tau_1, \tau_2),$$

where

$$I_1(\tau_1, \tau_2) := n^{-1} \mathbf{u}_n^{\top} \left\{ J_m(\tau_1)^{-1} - J_m(\tau_2)^{-1} \right\} \sum_{i=1}^n \mathbf{Z}_i \left( \mathbf{1} \left\{ Y_i \le Q(X_i, \tau_1) \right\} - \tau_1 \right),$$

$$I_2(\tau_1, \tau_2) := n^{-1} \mathbf{u}_n^{\top} J_m(\tau_2)^{-1} \sum_{i=1}^n \mathbf{Z}_i \left( \mathbf{1} \left\{ Y_i \le Q(X_i, \tau_1) \right\} - \mathbf{1} \left\{ Y_i \le Q(X_i, \tau_2) \right\} - (\tau_1 - \tau_2) \right).$$

Step 1: bounding  $I_1(\tau_1, \tau_2)$ .

Note that  $\sup_{\tau_1,\tau_2\in\mathcal{T}, |\tau_1-\tau_2|<\delta} |I_1(\tau_1,\tau_2)| \leq \|\mathbb{P}_n - P\|_{\mathcal{G}_6(\mathbf{u}_n,\delta)\cdot\mathcal{G}_4}$ , where

$$\mathcal{G}_{6}(\mathbf{u}_{n},\delta)\cdot\mathcal{G}_{4} = \left\{\mathbf{u}_{n}^{\top}\left\{J_{m}(\tau_{1})^{-1} - J_{m}(\tau_{2})^{-1}\right\}\mathbf{Z}_{i}\left(\mathbf{1}\left\{Y_{i} \leq Q(X_{i},\tau_{3})\right\} - \tau_{3}\right) \middle| \tau_{1},\tau_{2},\tau_{3} \in \mathcal{T}, |\tau_{1} - \tau_{2}| \leq \delta\right\}.$$

Theorem 2.6.7 of van der Vaart and Wellner (1996) and Part 1 of Lemma S.2.4 give

$$N(\epsilon \|G_4\|_{L_2(\mathbb{P}_n)}, \mathcal{G}_4, L_2(\mathbb{P}_n)) \le \frac{A_4}{\epsilon}$$

where the envelope for  $\mathcal{G}_4$  is  $G_4 = 2$  and  $A_4$  is a universal constant. Part 2 of Lemma S.2.4 and Part 2 of Lemma S.2.2 imply that

$$N(\epsilon \| G_6 G_4 \|_{L_2(\mathbb{P}_n)}, \mathcal{G}_6(\mathbf{u}_n, \delta) \cdot \mathcal{G}_4, L_2(\mathbb{P}_n)) \le \frac{2A_4}{\epsilon} \left(\frac{2C_0}{\epsilon}\right)^2 \le \left(\frac{2A_4^{1/3}C_0^{2/3}}{\epsilon}\right)^3, \tag{A.48}$$

where  $2A_4^{1/3}C_0^{2/3} \leq C$  for a large enough universal constant C > 0. To bound  $\sup_{f \in \mathcal{G}_6(\mathbf{u}_n, \delta) \cdot \mathcal{G}_4} ||f||_{L^2(P)}$ , note that

$$\mathbb{E}\left(\mathbf{u}_n^{\top}\left\{J_m(\tau_1)^{-1} - J_m(\tau_2)^{-1}\right\}\mathbf{Z}_i\left(\mathbf{1}\left\{Y_i \leq Q(X_i, \tau_3)\right\} - \tau_3\right)\right)^2 \\ \leq 4\|\mathbf{u}_n\|^2 \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}])\left[\inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_m(\tau))\right]^{-2} C_0^2 \delta^2 \leq C\|\mathbf{u}_n\|^2 \delta^2.$$

for a large enough constant C. In addition, an upper bound for the functions in  $\mathcal{G}_6(\mathbf{u}_n, \delta) \cdot \mathcal{G}_4$  is

$$2\xi_m \|\mathbf{u}_n\| [\inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_m(\tau))]^{-1} C_0 \delta \le C \delta \xi_m \|\mathbf{u}_n\|,$$

and we can take this upper bound as envelope.

Applying the bounds (S.2.2) and (S.2.3) and taking into account (A.48), for any  $\mathbf{u}_n$  and  $\delta > 0$ ,

$$\mathbb{E}\|\mathbb{P}_n - P\|_{\mathcal{G}_6(\mathbf{u}_n,\delta)\cdot\mathcal{G}_4} \le c_1 \left[ \|\mathbf{u}_n\| \delta \left(\frac{\log(\xi_m)}{n}\right)^{1/2} + \frac{\delta \|\mathbf{u}_n\| \xi_m}{n} \log(\xi_m) \right]. \tag{A.49}$$

Finally, for any  $\kappa_n > 0$ , let

$$r_{n,1}(\delta,\kappa_n) = \widetilde{C}\left[\|\mathbf{u}_n\|\delta\left(\frac{1}{n}\log(\xi_m)\right)^{1/2} + \frac{\delta\|\mathbf{u}_n\|\xi_m}{n}\log(\xi_m) + \left(\frac{\kappa_n}{n}\right)^{1/2}\|\mathbf{u}_n\|\delta + \frac{\delta\|\mathbf{u}_n\|\xi_m}{n}\kappa_n\right]$$

for a sufficiently large constant  $\tilde{C} > 0$ . From this, we obtain

$$P\bigg\{\sup_{\tau_1,\tau_2\in\mathcal{T},|\tau_1-\tau_2|<\delta}|I_1(\tau_1,\tau_2)|\geq r_{n,1}(\delta,\kappa_n)\bigg\}\leq P\big\{\|\mathbb{P}_n-P\|_{\mathcal{G}_6(\mathbf{u}_n,\delta)\cdot\mathcal{G}_4}\geq r_{n,1}(\delta,\kappa_n)\big\}\leq e^{-\kappa_n}.$$

### Step 2: bounding $I_2(\tau_1, \tau_2)$ .

Note that  $\sup_{\tau_1,\tau_2\in\mathcal{T},|\tau_1-\tau_2|<\delta}|I_2(\tau_1,\tau_2)|\leq \|\mathbb{P}_n-P\|_{\mathcal{G}_3(\mathbf{u}_n)\cdot\mathcal{G}_7(\delta)}$ , where

$$\mathcal{G}_{3}(\mathbf{u}_{n}) \cdot \mathcal{G}_{7}(\delta) = \left\{ \mathbf{u}_{n}^{\top} J_{m}(\tau_{3})^{-1} \mathbf{Z}_{i} \left( \mathbf{1} \{ Y_{i} \leq Q(X_{i}, \tau_{1}) \} - \mathbf{1} \{ Y_{i} \leq Q(X_{i}, \tau_{2}) \} - (\tau_{1} - \tau_{2}) \right) \middle| \right.$$

$$\tau_{1}, \tau_{2}, \tau_{3} \in \mathcal{T}, |\tau_{1} - \tau_{2}| \leq \delta \right\}.$$

Lemma S.2.3, Part 2 of Lemma S.2.4 and Part 2 of Lemma S.2.2 imply that

$$N(\epsilon \| G_3 G_7 \|_{L_2(\mathbb{P}_n)}, \mathcal{G}_3(\mathbf{u}_n) \cdot \mathcal{G}_7(\delta), L_2(\mathbb{P}_n)) \le \frac{2C_0}{\epsilon} \left(\frac{2A_7}{\epsilon}\right)^2 \le \left(\frac{2C_0^{1/3} A_7^{2/3}}{\epsilon}\right)^3, \tag{A.50}$$

where  $2C_0^{1/3}A_7^{2/3} \leq C$  for a large enough constant C > 0. To bound  $\sup_{f \in \mathcal{G}_3(\mathbf{u}_n) \cdot \mathcal{G}_7(\delta)} ||f||_{L^2(P)}$  note that

$$\mathbb{E}\left(\mathbf{u}_{n}^{\top}J_{m}(\tau_{3})^{-1}\mathbf{Z}_{i}\left(\mathbf{1}\left\{Y_{i} \leq Q(X_{i}, \tau_{1})\right\} - \mathbf{1}\left\{Y_{i} \leq Q(X_{i}, \tau_{2})\right\} - (\tau_{1} - \tau_{2})\right)\right)^{2}$$

$$\leq 3\|\mathbf{u}_{n}\|^{2}\lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}])\left[\inf_{\tau \in \mathcal{T}}\lambda_{\min}(J_{m}(\tau))\right]^{-2}\delta \leq C\|\mathbf{u}_{n}\|^{2}\delta.$$

Moreover

$$\sup_{f \in \mathcal{G}_3(\mathbf{u}_n) \cdot \mathcal{G}_7(\delta)} \sup \|f\|_{\infty} \le 2 \|\mathbf{u}_n\| \xi_m [\inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_m(\tau))]^{-1} \le C \|\mathbf{u}_n\| \xi_m$$

for some constant C. Applying the bounds (S.2.2) and (S.2.3) and taking into account (A.50)

$$\mathbb{E}\|\mathbb{P}_n - P\|_{\mathcal{G}_3(\mathbf{u}_n) \cdot \mathcal{G}_7(\delta)} \le c_1 \left[\|\mathbf{u}_n\| \left(\frac{\delta}{n} \log \frac{\xi_m}{\sqrt{\delta}}\right)^{1/2} + \frac{\|\mathbf{u}_n\| \xi_m}{n} \log \frac{\xi_m}{\sqrt{\delta}}\right]. \tag{A.51}$$

For any  $\kappa_n > 0$ , let

$$r_{n,2}(\delta, \kappa_n) = C \left[ \|\mathbf{u}_n\| \left( \frac{\delta}{n} \log \frac{\xi_m}{\sqrt{\delta}} \right)^{1/2} + \frac{\|\mathbf{u}_n\| \xi_m}{n} \log \frac{\xi_m}{\sqrt{\delta}} + \|\mathbf{u}_n\| \left( \frac{\kappa_n \delta}{n} \right)^{1/2} + \frac{\|\mathbf{u}_n\| \xi_m}{n} \kappa_n \right] \right]$$

for a constant C > 0 sufficiently large, we obtain

$$P\bigg\{\sup_{\tau_1,\tau_2\in\mathcal{T},|\tau_1-\tau_2|<\delta}|I_2(\tau_1,\tau_2)|\geq r_{n,2}(\delta,\kappa_n)\bigg\}\leq P\big\{\|\mathbb{P}_n-P\|_{\mathcal{G}_3(\mathbf{u}_n)\cdot\mathcal{G}_7(\delta)}\geq r_{n,2}(\delta,\kappa_n)\big\}\leq e^{-\kappa_n}.$$

Finally,  $r_{n,1}(\delta, \kappa_n) \leq r_{n,2}(\delta, \kappa_n)$  when  $\delta < 1$ . Hence, we conclude (A.44).

#### A.5 Proof of Theorem 4.1

As the argument  $x_0$  in  $Q(x_0; \tau)$  and  $F_{Y|X}(y|x_0)$  is fixed, simplify notations by writing  $Q(x_0; \tau) = Q(\tau)$ ,  $\widehat{Q}(x_0; \tau) = \widehat{Q}(\tau)$  and  $F_{Y|X}(y|x_0) = F(y)$ ,  $\widehat{F}_{Y|X}(y|x_0) = \widehat{F}(y)$  as functions of the single arguments in  $\tau$  and y, respectively. From Theorems 2.4, 3.1 or Corollary 2.2, we have

$$a_n(\widehat{Q}(\cdot) - Q(\cdot)) \leadsto \mathbb{G}(\cdot) \text{ in } \ell^{\infty}([\tau_L, \tau_U]),$$
 (A.52)

where  $a_n$  and  $\mathbb{G}$  depend on the model for  $Q(x;\tau)$  and  $\mathbb{G}$  has continuous sample paths almost surely. Next, note that for  $y \in \mathcal{Y}$ 

$$a_n(\widehat{F}(y) - F(y)) = a_n(\Phi(\widehat{Q})(y) - \Phi(Q)(y)).$$

Finally, observe that  $\Phi(f)(y) = \tau_L + (\tau_U - \tau_L)(\Phi^* \circ R)(f)(y)$  where  $\Phi^*(f)(y) := \int_0^1 \mathbf{1}\{f(u) < y\}du$  and  $R(f)(y) := f(\tau_L + y(\tau_U - \tau_L))$ . The map  $R : \ell^{\infty}((\tau_L, \tau_U)) \to \ell^{\infty}((0, 1))$  is linear and continuous, hence compactly differentiable with derivative R. The map  $\Phi^*$  is compactly differentiable tangentially to C(0, 1) at any strictly increasing, differentiable function  $f_0$  and the derivative of  $\Phi$  at  $f_0$  is given by  $d\Phi_{f_0}^*(h)(y) = -h(f_0^{-1}(y))/f_0'(f_0^{-1}(y))$  - see Corollary 1 in Chernozhukov et al. (2010). Hence the map  $\Phi^* \circ R$  is compactly differentiable at any strictly increasing function  $f_0 \in \ell^{\infty}((\tau_L, \tau_U))$  tangentially to  $C(\tau_L, \tau_U)$ . Combining this with the representation  $\Phi(f)(y) = \tau_L + (\tau_U - \tau_L)(\Phi^* \circ R)(f)(y)$  it follows that  $\Phi$  is compactly differentiable at any strictly increasing function  $f_0 \in \ell^{\infty}((\tau_L, \tau_U))$  with derivative  $d\Phi_{f_0}(h)(y) = -f'_0(f_0^{-1}(y))h(f_0^{-1}(y))$ . Thus weak convergence of  $a_n(\widehat{F}(y) - F(y))$  follows from the functional delta method.

Next, observe that  $\Psi(f) = \Theta \circ \Phi(f)$  where  $\Theta(f)(\tau) = \inf\{y : f(y) \ge \tau\}$  denotes the generalized inverse. Compact differentiability of  $\Theta$  at differentiable, strictly increasing functions  $f_0$  tangentially to the space of contunuous functions is established in Lemma 3.9.23 of van der Vaart and Wellner (1996), and the derivative of  $\Theta$  at  $f_0$  is given by  $d\Theta_{f_0}(h)(y) = -h(f_0^{-1}(y))/f'_0(f_0^{-1}(y))$ . By the chain rule for Hadamard derivatives this implies compact differentiability of  $\Psi$  tangentially to  $\mathcal{C}(\tau_L, \tau_U)$ . Thus the second weak convergence result again follows by the functional delta method.

## **APPENDIX B: Technical Remarks on Estimation Bias**

**Remark B.1.** In this remark we show the bound  $\tilde{c}_n = o(m^{-\lfloor \eta \rfloor})$  for univariate spline models discussed in Example 2.3, as well as  $c_n^{\dagger} = O(m^{-\lfloor \eta \rfloor/k'})$  for partial linear model in Section 3. We first show the latter below. We note that

Assume that  $W = [0, 1]^{k'}$ , that  $h(\cdot; \cdot) \in \Lambda_c^{\eta}(W, \mathcal{T})$  and that  $\widetilde{\mathbf{Z}}$  corresponds to a tensor product B-spline basis of order q on W with  $m^{1/k'}$  equidistant knots in each coordinate. Moreover, assume that (V, W) has a density  $f_{V,W}$  such that  $0 < \inf_{v,w} f_{V,W}(v,w) \le \sup_{v,w} f_{V,W}(v,w) < \infty$ . We shall show that in this case  $c_n^{\dagger} = O(m^{-\lfloor \eta \rfloor/k'})$  where  $c_n^{\dagger}$  is defined in Assumption (C1). Define

$$\boldsymbol{\beta}_{n,g}(\tau) := \underset{\boldsymbol{\beta} \in \mathbb{R}^m}{\operatorname{argmin}} \int \left( \widetilde{\mathbf{Z}}(w)^{\top} \boldsymbol{\beta} - h(w;\tau) \right)^2 \int f_{Y|X}(Q(v,w;\tau)|(v,w)) f_{V,W}(v,w) dv dw. \tag{B.1}$$

Note that  $w \mapsto \widetilde{\mathbf{Z}}(w)^{\top} \boldsymbol{\beta}_{n,g}(\tau)$  can be viewed as a projection of a function  $g : \mathcal{W} \to \mathbb{R}$  onto the spline space  $\mathcal{B}_m(\mathcal{W}) := \{w \mapsto \widetilde{\mathbf{Z}}(w)^{\top} \mathbf{b} : \mathbf{b} \in \mathbb{R}^m\}$ , with respect to the inner product  $\langle g_1, g_2 \rangle = \int g_1(w)g_2(w)d\nu(w)$ , where  $d\nu(w) := (\int_v f_{Y|X}(Q(v,w;\tau)|v,w)f_{V,W}(v,w)dv)dw$ .

We first apply Theorem A.1 on p.1630 of Huang (2003). To do so, we need to verify Condition A.1-A.3 of Huang (2003). Condition A.1 can be verified by invoking (A2)-(A3) in our paper and using the bounds on  $f_W$ . The choice of basis functions and knots ensures that Conditions A.2 and A.3 hold (see the discussion on p.1630 of Huang (2003)). Thus, Theorem A.1 on p.1630 of Huang (2003) implies there exists a constant C independent of n such that for any function on W,

$$\sup_{w \in \mathcal{W}} \left| \widetilde{\mathbf{Z}}(w)^{\top} \boldsymbol{\beta}_{n,g}(\tau) \right| \leq C \sup_{w \in \mathcal{W}} \left| g(w) \right|.$$

Recall that W is a compact subset of  $\mathbb{R}^d$  and  $h(w;\tau) \in \Lambda_c^{\eta}(W,\mathcal{T})$ . Since  $\mathcal{B}_m(W)$  is a finite dimensional vector space of functions, by a compactness argument there exists  $g^*(\cdot;\tau) \in \mathcal{B}_m(W)$  such that  $\sup_{w \in \mathcal{W}} |h(w;\tau) - g^*(w;\tau)| = \inf_{g \in \mathcal{B}_m(W)} \sup_{w \in \mathcal{W}} |h(w;\tau) - g(w)|$  for each fixed  $\tau$ . With  $m > \eta$ , the inequality in the proof for Theorem 12.8 in Schumaker (1981), with their " $m_i$ " being our  $\eta$  and  $\overline{\Delta}_i \approx m^{-1/k'}$  yields

$$\begin{split} \widetilde{c}_{n} &= \sup_{w,\tau} \left| \widetilde{\mathbf{Z}}(w)^{\top} \boldsymbol{\beta}_{n,h(w;\tau)}(\tau) - h(w;\tau) \right| \\ &= \sup_{w,\tau} \left| \widetilde{\mathbf{Z}}(w)^{\top} \boldsymbol{\beta}_{n,h(w;\tau)}(\tau) - g^{*}(w;\tau) + g^{*}(w;\tau) - h(w;\tau) \right| \\ &\leq \sup_{w,\tau} \left| \widetilde{\mathbf{Z}}(w)^{\top} \boldsymbol{\beta}_{n,h(w;\tau)-g^{*}(w;\tau)}(\tau) \right| + \sup_{w,\tau} \left| g^{*}(w;\tau) - h(w;\tau) \right| \\ &\leq (C+1) \sup_{\tau \in \mathcal{T}} \inf_{g \in \mathcal{B}(\mathcal{W})} \sup_{x} \left| h(w;\tau) - g(w) \right| \\ &\lesssim m^{-\lfloor \eta \rfloor/k'} k' \max_{|j| \leq \eta} \sup_{\tau \in \mathcal{T}} \sup_{w \in \mathcal{W}} |D^{j}h(w;\tau)|, \end{split}$$

where  $\lfloor \eta \rfloor$  is the greatest integer less than  $\eta$ ,  $\max_{|j| \leq \eta} \sup_{\tau \in \mathcal{T}} \sup_{x \in \mathcal{W}} |D^j h(w;\tau)| = O(1)$  by the assumption that  $h(w;\tau) \in \Lambda_c^{\eta}(\mathcal{W},\mathcal{T})$  and fixed k'. An extension of Theorem 12.8 of Schumaker (1981) to Besov spaces (see Example 6.29 of Schumaker (1981) or p.5573 of Chen (2007)) in similar manner as Theorem 6.31 of Schumaker (1981) could refine the rate to  $\tilde{c}_n \lesssim m^{-\eta/k'}$ , but we do not pursue this direction here.

Next we show the bound  $\tilde{c}_n = o(m^{-\lfloor \eta \rfloor})$  in the setting of Example 2.3. Assume the density  $f_X(x)$  of X exists and  $0 < \inf_{x \in \mathcal{X}} f_X(x) \le \sup_{x \in \mathcal{X}} f_X(x) < \infty$ . Define the measure  $\nu(u)$  by  $d\nu(u) = f(Q(u;\tau)|u)f_X(u)du$ . Thus,  $x \mapsto \mathbf{B}(x)^{\top} \boldsymbol{\beta}_{n,g}(\tau)$  with  $\boldsymbol{\beta}_{n,g}$  defined similarly to (B.1) is now viewed as a projection of a function  $g: \mathcal{X} \to \mathbb{R}$  onto the space  $\mathcal{B}(\mathcal{X})$  with respect to the inner product  $\langle g_1, g_2 \rangle = \int g_1(u)g_2(u)d\nu(u)$ . The remaining proof is similar to the partial linear model, with  $h(w;\tau)$  being replaced by  $Q(x;\tau)$ , and we omit the details.

# References

- Angrist, J., Chernozhukov, V., and Fernández-Val, I. (2006). Quantile regression under misspecification, with an application to the U.S. wage structure. *Econometrica*, 74(2):539–563.
- Bassett, Jr., G. and Koenker, R. (1982). An empirical quantile function for linear models with iid errors. *Journal of American Statistical Association*, 77(378):407–415.
- Belloni, A., Chernozhukov, V., and Fernández-Val, I. (2011). Conditional quantile processes based on series or many regressors. arXiv preprint arXiv:1105.6154.
- Briollais, L. and Durrieu, G. (2014). Application of quantile regression to recent genetic and -omic studies. *Human Genetics*, 133:951–966.
- Cade, B. S. and Noon, B. R. (2003). A gentle introduction to quantile regression for ecologists. Frontiers in Ecology and the Environment, 1(8):412–420.
- Chen, X. (2007). Large sample sieve estimation of semi-nonparametric models. In Heckman, J. J. and Leamer, E., editors, *Handbook of Econometrics*, chapter 76. North-Holland.
- Cheng, G. and Shang, Z. (2015). Joint asymptotics for semi-nonparametric regression models with partially linear structure. *Annals of Statistics*, 43(3):1351–1390.
- Cheng, G., Zhou, L., and Huang, J. Z. (2014). Efficient semiparametric estimation in generalized partially linear additive models for longitudinal/clustered data. *Bernoulli*, 20(1):141–163.
- Chernozhukov, V., Fernández-Val, I., and Galichon, A. (2010). Quantile and probability curves without crossing. *Econometrica*, 78(3):1093–1125.
- Delgado, M. A. and Escanciano, J. C. (2013). Conditional stochastic dominance testing. *Journal of Business & Economic Statistics*, 31(1):16–28.
- Demko, S., Moss, W. F., and Smith, P. W. (1984). Decay rates for inverses of band matrices. *Mathematics of Computation*, 43(168):491–499.
- Dette, H. and Volgushev, S. (2008). Non-crossing non-parametric estimates of quantile curves. Journal of the Royal Statistical Society: Series B, 70(3):609–627.
- He, X. and Shi, P. (1994). Convergence rate of B-spline estimators of nonparametric conditional quantile functions. *Journal title of Nonparametric Statistics*, 3(3-4):299–308.

- He, X. and Shi, P. (1996). Bivariate tensor-product B-splines in a partly linear model. *Journal of Multivariate Analysis*, 58(2):162–181.
- Huang, J. Z. (2003). Local asymptotics for polynomial spline regression. *Annals of Statistics*, 31(5):1600–1635.
- Kley, T., Volgushev, S., Dette, H., and Hallin, M. (2015+). Quantile spectral processes: Asymptotic analysis and inference. *Bernoulli*.
- Knight, K. (2008). Asymptotics of the regression quantile basic solution under misspecification. *Applications of Mathematics*, 53(3):223–234.
- Koenker, R. (2005). Quantile Regression. Cambridge University Press, New York.
- Koenker, R. and Bassett, Jr., G. (1978). Regression quantiles. *Econometrica*, 46(1):33–50.
- Koenker, R. and Hallock, K. F. (2001). Quantile regression. *Journal of Economic Perspectives*, 15(4):143–156.
- Koenker, R. and Xiao, Z. (2002). Inference on the quantile regression process. *Econometrica*, 70(4):1583–1612.
- Koltchinskii, V. (2006). Local Rademacher complexities and oracle inequalities in risk minimization. Annals of Statistics, 34(6):2593–2656.
- Lee, S. (2003). Efficient semiparametric estimation of a partially linear quantile regression model. *Econometric Theory*, 19:1–31.
- Massart, P. (2000). About the constants in Talagrand's concentration inequalities for empirical processes. *Annals of Probability*, pages 863–884.
- Puntanen, S. and Styan, G. P. H. (2005). Schur complements in statistics and probabiliy. In Zhang, F., editor, *The Schur Complement and Its Applications*, volume 4 of *Numerical Methods and Algorithms*. Springer.
- Qu, Z. and Yoon, J. (2015). Nonparametric estimation and inference on conditional quantile processes. *Journal of Econometrics*, 185:1–19.
- Schumaker, L. (1981). Spline Functions: Basic Theory. Wiley, New York.
- Shang, Z. and Cheng, G. (2015). A Bayesian splitotic theory for nonparametric models. arXiv preprint arXiv:1508.04175.
- van der Vaart, A. and Wellner, J. (1996). Weak Convergence and Empirical Processes: With Applications to Statistics. Springer.
- Volgushev, S. (2013). Smoothed quantile regression processes for binary response models. arXiv preprint arXiv:1302.5644.

- Volgushev, S., Chao, S.-K., and Cheng, G. (2016). Data compression for extraordinarily large data: A quantile regression approach. Technical report, Purdue University.
- Zhao, T., Cheng, G., and Liu, H. (2016). A partially linear framework for massive heterogeneous data. *Annals of Statistics*.
- Zhou, S., Shen, X., and Wolfe, D. (1998). Local asymptotics for regression splines and confidence regions. *Annals of Statistics*, 26(5):1760–1782.

# SUPPLEMENTARY MATERIAL: QUANTILE PROCESS FOR SEMI-NONPARAMETRIC REGRESSION MODELS

In this supplemental material, we provide the auxiliary proofs needed in the appendices. Section S.1 develops the technicalities for Bahadur representations. Section S.2 presents some empirical process results and computes the covering number of some function classes encountered in the proofs of asymptotic tightness of quantile process.

# S.1: Proofs for Bahadur Representations

# S.1.1 Proof of Theorem 5.1

Some rearranging of terms yields

$$\mathbb{P}_{n}\psi(\cdot;\widehat{\boldsymbol{\gamma}}(\tau),\tau) = n^{-1/2}\mathbb{G}_{n}(\psi(\cdot;\widehat{\boldsymbol{\gamma}}(\tau),\tau)) - n^{-1/2}\mathbb{G}_{n}(\psi(\cdot;\boldsymbol{\gamma}_{n}(\tau),\tau)) 
+ \widetilde{J}_{m}(\tau)(\widehat{\boldsymbol{\gamma}}(\tau)-\boldsymbol{\gamma}_{n}(\tau)) + n^{-1/2}\mathbb{G}_{n}(\psi(\cdot;\boldsymbol{\gamma}_{n}(\tau),\tau)) 
+ \mu(\boldsymbol{\gamma}_{n}(\tau),\tau) + \left(\mu(\widehat{\boldsymbol{\gamma}}(\tau),\tau) - \mu(\boldsymbol{\gamma}_{n}(\tau),\tau) - \widetilde{J}_{m}(\tau)(\widehat{\boldsymbol{\gamma}}(\tau)-\boldsymbol{\gamma}_{n}(\tau))\right).$$

In other words

$$\widehat{\gamma}(\tau) - \gamma_n(\tau) = -n^{-1/2} J_m(\tau)^{-1} \mathbb{G}_n(\psi(\cdot; \gamma_n(\tau), \tau)) + r_{n,1}(\tau) + r_{n,2}(\tau) + r_{n,3}(\tau) + r_{n,4}(\tau) \quad (S.1.1)$$

where

$$r_{n,1}(\tau) := \widetilde{J}_m(\tau)^{-1} \mathbb{P}_n \psi(\cdot; \widehat{\gamma}(\tau), \tau),$$

$$r_{n,2}(\tau) := -\widetilde{J}_m(\tau)^{-1} \Big( \mu(\widehat{\gamma}(\tau), \tau) - \mu(\gamma_n(\tau), \tau) - \widetilde{J}_m(\tau)(\widehat{\gamma}(\tau) - \gamma_n(\tau)) \Big),$$

$$r_{n,3}(\tau) := -n^{-1/2} \widetilde{J}_m(\tau)^{-1} \Big( \mathbb{G}_n(\psi(\cdot; \widehat{\gamma}(\tau), \tau)) - \mathbb{G}_n(\psi(\cdot; \gamma_n(\tau), \tau)) \Big),$$

$$r_{n,4}(\tau) := -n^{-1/2} (J_m(\tau)^{-1} - \widetilde{J}_m(\tau)^{-1}) \mathbb{G}_n(\psi(\cdot; \gamma_n(\tau), \tau)) - \widetilde{J}_m(\tau)^{-1} \mu(\gamma_n(\tau), \tau)).$$

The remaining proof consists in bounding the individual remainder terms.

The bound on  $r_{n,1}$  follows from results on duality theory for convex optimization, see Lemma 26 on page 66 in Belloni et al. (2011) for a proof.

To bound  $r_{n,2}$  and  $r_{n,3}$ , define the class of functions

$$\mathcal{G}_1 := \left\{ (\mathbf{Z}, Y) \mapsto \mathbf{a}^\top \mathbf{Z} (\mathbf{1} \{ Y \le \mathbf{Z}^\top \mathbf{b} \} - \tau) \mathbf{1} \{ \|\mathbf{Z}\| \le \xi_m \} \middle| \tau \in \mathcal{T}, \mathbf{b} \in \mathbb{R}^m, \mathbf{a} \in \mathcal{S}^{m-1} \right\}.$$
 (S.1.2)

Moreover, let

$$s_{n,1} := \|\mathbb{P}_n - P\|_{\mathcal{G}_1}.$$

Observe that by Lemma S.1.2 with t = 2 we have

$$\Omega_{1,n} := \left\{ \sup_{\tau \in \mathcal{T}} \|\widehat{\boldsymbol{\gamma}}(\tau) - \boldsymbol{\gamma}_n(\tau)\| \le \frac{4(s_{n,1} + g_n)}{\inf_{\tau \in \mathcal{T}} \lambda_{\min}(\widetilde{J}_m(\tau))} \right\} \supseteq \left\{ s_{n,1} + g_n < \frac{\inf_{\tau \in \mathcal{T}} \lambda_{\min}^2(\widetilde{J}_m(\tau))}{8\xi_m \overline{f'} \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^\top])} \right\} =: \Omega_{2,n}.$$

Define the event

$$\Omega_{3,n} := \left\{ s_{n,1} \le C \left[ \left( \frac{m}{n} \log n \right)^{1/2} + \frac{m\xi_m}{n} \log n + \left( \frac{\kappa_n}{n} \right)^{1/2} + \frac{\xi_m \kappa_n}{n} \right] \right\}.$$

Now it follows from Lemma S.1.3 that  $P(\Omega_{3,n}) \geq 1 - e^{-\kappa_n}$  [note that  $\xi_m = O(n^b)$  yields  $\log \xi_m = O(\log n)$ ]. Moreover, the assumption  $m\xi_m^2 \log n = o(n), \xi_m = O(n^b), \xi_m g_n = o(1)$  implies that for  $\kappa_n \ll n/\xi_m^2$  and large enough n,

$$C\left[\left(\frac{m}{n}\log n\right)^{1/2} + \frac{m\xi_m}{n}\log n + \left(\frac{\kappa_n}{n}\right)^{1/2} + \frac{\xi_m\kappa_n}{n}\right] + g_n \le \frac{\inf_{\tau \in \mathcal{T}} \lambda_{\min}^2(\widetilde{J}_m(\tau))}{8\xi_m \overline{f'}\lambda_{\max}(\mathbb{E}[\mathbf{ZZ}^\top])}$$
(S.1.3)

for n large enough. Thus, for all n for which (S.1.3) holds,  $\Omega_{3,n} \subseteq \Omega_{2,n} \subseteq \Omega_{1,n}$ . From this we obtain that on  $\Omega_{3,n}$ , for a constant  $C_2$  which is independent of n, we have

$$\sup_{\tau \in \mathcal{T}} \|\widehat{\gamma}(\tau) - \gamma_n(\tau)\| \le C_2 \left[ \left( \frac{m}{n} \log n \right)^{1/2} + \frac{m\xi_m}{n} \log n + \left( \frac{\kappa_n}{n} \right)^{1/2} + \frac{\xi_m \kappa_n}{n} + g_n \right].$$

In particular, for all n for which (S.1.3) holds,

$$P\left(\sup_{\tau \in \mathcal{T}} \|\widehat{\gamma}(\tau) - \gamma_n(\tau)\| \le C_2 \left[ \left( \frac{m}{n} \log n \right)^{1/2} + \frac{m\xi_m}{n} \log n + \left( \frac{\kappa_n}{n} \right)^{1/2} + \frac{\xi_m \kappa_n}{n} + g_n \right] \right) \ge 1 - e^{-\kappa_n}.$$
 (S.1.4)

The bound on  $r_{n,2}$  is now a direct consequence of Lemma S.1.1 and the fact that  $n^{-1}m\xi_m \log n = o((n^{-1}m\log n)^{1/2})$  and  $\xi_m\kappa_n n^{-1} = o(\kappa_n^{1/2}n^{-1/2})$ . The bound on  $r_{n,4}$  follows once we observe that for any  $\mathbf{a} \in \mathcal{S}^{m-1}$ 

$$|\mathbf{a}^{\top}(J_m(\tau) - \widetilde{J}_m(\tau))\mathbf{a}| \le \overline{f'}c_n\lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}]).$$
 (S.1.5)

Together with the identity  $A^{-1} - B^{-1} = B^{-1}(B - A)A^{-1}$  this implies that for sufficiently large n we have  $\sup_{\tau \in \mathcal{T}} ||r_{n,4}(\tau)|| \le C_1(c_n s_{n,1} + g_n)$  for a constant  $C_1$  which does not depend on n.

Thus, it remains to bound  $r_{n,3}$ . Observe that on the set  $\{\sup_{\tau \in \mathcal{T}} \|\widehat{\gamma}(\tau) - \gamma_n(\tau)\| \leq \delta\}$  we have the bound

$$\sup_{\tau \in \mathcal{T}} \|r_{n,3}(\tau)\| \le \frac{1}{\inf_{\tau \in \mathcal{T}} \lambda_{\min}(\widetilde{J}_m(\tau))} \|\mathbb{P}_n - P\|_{\mathcal{G}_2(\delta)},$$

where the class of functions  $\mathcal{G}_2(\delta)$  is defined as follows

$$\mathcal{G}_2(\delta) := \left\{ (\mathbf{Z}, Y) \mapsto \mathbf{a}^\top \mathbf{Z} (\mathbf{1} \{ Y \le \mathbf{Z}^\top \mathbf{b}_1 \} - \mathbf{1} \{ Y \le \mathbf{Z}^\top \mathbf{b}_2 \}) \mathbf{1} \{ \| \mathbf{Z} \| \le \xi_m \} \right|$$

$$\mathbf{b}_1, \mathbf{b}_2 \in \mathbb{R}^m, \| \mathbf{b}_1 - \mathbf{b}_2 \| \le \delta, \mathbf{a} \in \mathcal{S}^{m-1} \}. \quad (S.1.6)$$

It thus follows that for any  $\delta, \alpha > 0$ 

$$P\Big(\sup_{\tau \in \mathcal{T}} \|r_{n,3}(\tau)\| \ge \alpha\Big) \le P\Big(\sup_{\tau \in \mathcal{T}} \|\widehat{\gamma}(\tau) - \gamma_n(\tau)\| \ge \delta\Big) + P\Big(\frac{\|\mathbb{P}_n - P\|_{\mathcal{G}_2(\delta)}}{\inf_{\tau \in \mathcal{T}} \lambda_{\min}(\widetilde{J}_m(\tau))} \ge \alpha\Big).$$

Letting  $\delta := C((n^{-1}m\log n)^{1/2} + (\kappa_n/n)^{1/2} + g_n)$  and  $\alpha := C\zeta_n(\delta_n, \kappa_n)$  (here,  $\zeta_n$  is defined in (S.1.11) in the statement of Lemma S.1.3) with a suitable constant C, the bounds in Lemma S.1.3 and (S.1.4) yield the desired bound.

## S.1.1.1 Technical details for the proof of Theorem 5.1

**Lemma S.1.1.** Under assumptions (A1)-(A3) we have for any  $\delta > 0$ ,

$$\sup_{\tau \in \mathcal{T}} \sup_{\|\mathbf{b} - \boldsymbol{\gamma}_n(\tau)\| \le \delta} \|\mu(\mathbf{b}, \tau) - \mu(\boldsymbol{\gamma}_n(\tau), \tau) - \widetilde{J}_m(\tau)(\mathbf{b} - \boldsymbol{\gamma}_n(\tau))\| \le \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^\top])\overline{f'}\delta^2 \xi_m,$$

**Proof of Lemma S.1.1.** Note that  $\mu'(\gamma_n(\tau), \tau) = \mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}f_{Y|X}(\mathbf{Z}^{\top}\gamma_n(\tau)|X)] = \widetilde{J}_m(\tau)$  where we use the notation  $\mu'(\mathbf{b}, \tau) := \partial_{\mathbf{b}}\mu(\mathbf{b}, \tau)$ . Additionally, we have

$$\mu(\mathbf{b}, \tau) = \mu(\gamma_n(\tau), \tau) + \mu'(\bar{\gamma}_n, \tau)(\mathbf{b} - \gamma_n(\tau)),$$

where  $\bar{\gamma}_n = \mathbf{b} + \lambda_{\mathbf{b},\tau}(\gamma_n(\tau) - \mathbf{b})$  for some  $\lambda_{\mathbf{b},\tau} \in [0,1]$ . Moreover, for any  $\mathbf{a} \in \mathbb{R}^m$ ,

$$\mathbf{a}^{\top}[\mu(\mathbf{b},\tau) - \mu(\boldsymbol{\gamma}_n(\tau),\tau) - \mu'(\boldsymbol{\gamma}_n(\tau),\tau)(\mathbf{b} - \boldsymbol{\gamma}_n(\tau))] = \mathbf{a}^{\top}[(\mu'(\bar{\boldsymbol{\gamma}}_n,\tau) - \mu'(\boldsymbol{\gamma}_n(\tau),\tau))(\mathbf{b} - \boldsymbol{\gamma}_n(\tau))]$$

and thus we have for any  $\|\mathbf{b} - \boldsymbol{\gamma}_n(\tau)\| \leq \delta$ 

$$\begin{aligned} & \| \mu(\mathbf{b}, \tau) - \mu(\boldsymbol{\gamma}_{n}(\tau), \tau) - \mu'(\boldsymbol{\gamma}_{n}(\tau), \tau)(\mathbf{b} - \boldsymbol{\gamma}_{n}(\tau)) \| \\ & \leq \sup_{\|\mathbf{a}\|=1} \left| \mathbb{E} \left[ (\mathbf{a}^{\top} \mathbf{Z}) \mathbf{Z}^{\top} (\mathbf{b} - \boldsymbol{\gamma}_{n}(\tau)) (f_{Y|X} (\mathbf{Z}^{\top} \bar{\boldsymbol{\gamma}}_{n}|X) - f_{Y|X} (\mathbf{Z}^{\top} \boldsymbol{\gamma}_{n}(\tau)|X)) \right] \right| \\ & \leq \overline{f'} \sup_{\|\mathbf{a}\|=1} \mathbb{E} \left[ \left| \mathbf{a}^{\top} \mathbf{Z} \right| \left| \mathbf{Z}^{\top} (\bar{\boldsymbol{\gamma}}_{n} - \boldsymbol{\gamma}_{n}(\tau)) \right| \left| \mathbf{Z}^{\top} (\mathbf{b} - \boldsymbol{\gamma}_{n}(\tau)) \right| \right] \\ & \leq \overline{f'} \xi_{m} \mathbb{E} \left[ \left| \mathbf{Z}^{\top} (\bar{\boldsymbol{\gamma}}_{n} - \boldsymbol{\gamma}_{n}(\tau)) \right| \left| \mathbf{Z}^{\top} (\mathbf{b} - \boldsymbol{\gamma}_{n}(\tau)) \right| \right] \\ & \leq \xi_{m} \delta^{2} \overline{f'} \sup_{\|\mathbf{a}\|=1} \mathbb{E} \left[ \left| \mathbf{a}^{\top} \mathbf{Z} \right|^{2} \right], \end{aligned}$$

here the last inequality follows by Chauchy-Schwarz. Since the last line does not depend on  $\tau$ , **b**, this completes the proof.

**Lemma S.1.2.** Let assumptions (A1)-(A3) hold. Then, for any t > 1

$$\left\{\sup_{\tau \in \mathcal{T}} \|\widehat{\boldsymbol{\gamma}}(\tau) - \boldsymbol{\gamma}_n(\tau)\| \le \frac{2t(s_{n,1} + g_n)}{\inf_{\tau \in \mathcal{T}} \lambda_{\min}(\widetilde{J}_m(\tau))}\right\} \supseteq \left\{(s_{n,1} + g_n) < \frac{\inf_{\tau \in \mathcal{T}} \lambda_{\min}^2(\widetilde{J}_m(\tau))}{4t\xi_m \overline{f'} \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^\top])}\right\},$$

where  $s_{n,1} := \|\mathbb{P}_n - P\|_{\mathcal{G}_1}$  and  $\mathcal{G}_1$  is defined in (S.1.2).

**Proof of Lemma S.1.2.** Observe that  $f: \mathbf{b} \mapsto \mathbb{P}_n \rho_{\tau}(Y_i - \mathbf{Z}_i^{\top} \mathbf{b})$  is convex, and the vector  $\mathbb{P}_n \psi(\cdot; \mathbf{b}, \tau)$  is a subgradient of f at the point  $\mathbf{b}$ . Recalling that  $\widehat{\gamma}(\tau)$  is a minimizer of  $\mathbb{P}_n \rho_{\tau}(Y_i - \mathbf{Z}_i^{\top} \mathbf{b})$ , it follows that for any a > 0,

$$\{\sup_{\tau \in \mathcal{T}} \|\widehat{\boldsymbol{\gamma}}(\tau) - \boldsymbol{\gamma}_n(\tau)\| \le a(s_{n,1} + g_n)\} \supseteq \{\inf_{\tau} \inf_{\|\boldsymbol{\delta}\| = 1} \boldsymbol{\delta}^{\top} \mathbb{P}_n \psi(\cdot; \boldsymbol{\gamma}_n(\tau) + a(s_{n,1} + g_n)\boldsymbol{\delta}, \tau) > 0\}.$$
 (S.1.7)

To see this, define  $\boldsymbol{\delta} := (\widehat{\boldsymbol{\gamma}}(\tau) - \boldsymbol{\gamma}_n(\tau)) / \|\widehat{\boldsymbol{\gamma}}(\tau) - \boldsymbol{\gamma}_n(\tau)\|$  and note that by definition of the subgradient we have for any  $\widetilde{\zeta}_n > 0$ 

$$\mathbb{P}_n \rho_{\tau}(Y_i - \mathbf{Z}_i^{\top} \widehat{\boldsymbol{\gamma}}(\tau)) \geq \mathbb{P}_n \rho_{\tau}(Y_i - \mathbf{Z}_i^{\top} (\boldsymbol{\gamma}_n(\tau) + \widetilde{\boldsymbol{\zeta}}_n \boldsymbol{\delta})) + (\|\widehat{\boldsymbol{\gamma}}(\tau) - \boldsymbol{\gamma}_n(\tau)\| - \widetilde{\boldsymbol{\zeta}}_n) \boldsymbol{\delta}^{\top} \mathbb{P}_n \psi(\cdot; \boldsymbol{\gamma}_n(\tau) + \widetilde{\boldsymbol{\zeta}}_n \boldsymbol{\delta}, \tau).$$

Set  $\widetilde{\zeta}_n = a(s_{n,1} + g_n)$ . By the definition of  $\widehat{\gamma}(\tau)$  as minimizer, the inequality above can only be true if  $(\|\widehat{\gamma}(\tau) - \gamma_n(\tau)\| - \widetilde{\zeta}_n) \delta^{\top} \mathbb{P}_n \psi(\cdot; \gamma_n(\tau) + \widetilde{\zeta}_n \delta, \tau) \leq 0$ , which yields (S.1.7).

The proof is finished once we minorize the empirical score  $\boldsymbol{\delta}^{\top} \mathbb{P}_n \psi(\cdot; \boldsymbol{\gamma}_n(\tau) + a(s_{n,1} + g_n) \boldsymbol{\delta}, \tau)$  in (S.1.7) in terms of  $s_{n,1} + g_n$ . To proceed, observe that under assumptions (A1)-(A3) we have by Lemma S.1.1

$$\sup_{\|\boldsymbol{\delta}\|=1} \left| \mathbb{E}[\boldsymbol{\delta}^{\top} \{ \psi(Y, \mathbf{Z}; \mathbf{b}, \tau) - \psi(Y, \mathbf{Z}; \boldsymbol{\gamma}_{n}(\tau), \tau) - \mathbf{Z}\mathbf{Z}^{\top} f_{Y|X} (\boldsymbol{\gamma}_{n}(\tau)^{\top} \mathbf{Z}|X) (\mathbf{b} - \boldsymbol{\gamma}_{n}(\tau)) \}] \right|$$

$$\leq \xi_{m} \overline{f'} \lambda_{\max} (\mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}]) \|\mathbf{b} - \boldsymbol{\gamma}_{n}(\tau)\|^{2}.$$
(S.1.8)

Therefore, we have for arbitrary  $\|\boldsymbol{\delta}\| = 1, \tau \in \mathcal{T}$  that

$$\boldsymbol{\delta}^{\top} \mathbb{P}_{n} \psi(\cdot; \boldsymbol{\gamma}_{n}(\tau) + a(s_{n,1} + g_{n}) \boldsymbol{\delta}, \tau)$$

$$\geq -s_{n,1} - g_{n} + \boldsymbol{\delta}^{\top} \left( \mathbb{E} \left[ \psi(Y, \mathbf{Z}; \boldsymbol{\gamma}_{n}(\tau) + a(s_{n,1} + g_{n}) \boldsymbol{\delta}, \tau) \right] - \mathbb{E} \left[ \psi(Y, \mathbf{Z}; \boldsymbol{\gamma}_{n}(\tau), \tau) \right] \right)$$

$$\geq a(s_{n,1} + g_{n}) \inf_{\tau \in \mathcal{T}} \lambda_{\min}(\widetilde{J}_{m}(\tau)) - s_{n,1} - g_{n} - \xi_{m} \overline{f'} \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}]) a^{2}(s_{n,1} + g_{n})^{2},$$

where for the first inequality we recall the definition  $g_n = \sup_{\tau \in \mathcal{T}} \|\mathbb{E}[\psi(Y, \mathbf{Z}; \gamma_n(\tau), \tau)]\|$  and  $s_{n,1} = \|\mathbb{P}_n - P\|_{\mathcal{G}_1}$ ; the second inequality follows by (S.1.8). Setting  $a = 2t/\inf_{\tau \in \mathcal{T}} \lambda_{\min}(\widetilde{J}_m(\tau))$  in  $\widetilde{\zeta}_n$ , we see that the right-hand side of the display above is positive when

$$s_{n,1} + g_n < \frac{(2t-1)\inf_{\tau \in \mathcal{T}} \lambda_{\min}^2(\widetilde{J}_m(\tau))}{4t^2 \xi_m \overline{f'} \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^\top])}.$$

Observing that for t > 1 we have  $(2t - 1)/t^2 \ge 1/t$  and plugging  $a = 2t/\inf_{\tau \in \mathcal{T}} \lambda_{\min}(\widetilde{J}_m(\tau))$  in equation (S.1.7) completes the proof.

**Lemma S.1.3.** Consider the classes of functions  $\mathcal{G}_1, \mathcal{G}_2(\delta)$  defined in (S.1.2) and (S.1.6), respectively. Under assumptions (A1)-(A3) we have for some constant C independent of n and all  $\kappa_n > 0$  provided that  $\xi_m = O(n^b)$  for some fixed b

$$P\left(\|\mathbb{P}_n - P\|_{\mathcal{G}_1} \ge C\left[\left(\frac{m}{n}\log\xi_m\right)^{1/2} + \frac{m\xi_m}{n}\log\xi_m + \left(\frac{\kappa_n}{n}\right)^{1/2} + \frac{\xi_m\kappa_n}{n}\right]\right) \le e^{-\kappa_n}.$$
 (S.1.9)

For any  $\delta_n$  satisfying  $\xi_m \delta_n \gg n^{-1}$ , we have for sufficiently large n and arbitrary  $\kappa_n > 0$ 

$$P(\|\mathbb{P}_n - P\|_{\mathcal{G}_2(\delta_n)} \ge C\zeta_n(\delta_n, \kappa_n)) \le e^{-\kappa_n},$$
(S.1.10)

where

$$\zeta_n(t,\kappa_n) := \xi_m^{1/2} t^{1/2} \left(\frac{m}{n} \log(\xi_m \vee n)\right)^{1/2} + \frac{m\xi_m}{n} \log(\xi_m \vee n) + n^{-1/2} (\xi_m t \kappa_n)^{1/2} + n^{-1} \xi_m \kappa_n.$$
(S.1.11)

**Proof of Lemma S.1.3.** Observe that for each  $f \in \mathcal{G}_1$  we have  $|f(x,y)| \leq \xi_m$ , and the same holds for  $\mathcal{G}_2(\delta)$  for any value of  $\delta$ . Additionally, similar arguments as those in the proof of Lemma 18 in Belloni et al. (2011) imply together with Theorem 2.6.7 in van der Vaart and Wellner (1996) that, almost surely,

$$N(\mathcal{G}_2(\delta), L_2(\mathbb{P}_n); \varepsilon) \le \left(\frac{A\|F\|_{L^2(\mathbb{P}_n)}}{\varepsilon}\right)^{v_1(m)}, \quad N(\mathcal{G}_1, L_2(\mathbb{P}_n); \varepsilon) \le \left(\frac{A\|F\|_{L^2(\mathbb{P}_n)}}{\varepsilon}\right)^{v_2(m)},$$

where A is some constant and  $v_1(m) = O(m), v_2(m) = O(m)$ . Finally, for each  $f \in \mathcal{G}_1$  we have

$$\mathbb{E}[f^2] \leq \sup_{\|\mathbf{a}\|=1} \mathbf{a}^\top \mathbb{E}[\mathbf{Z}\mathbf{Z}^\top] \mathbf{a} = \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^\top]).$$

On the other hand, each  $f \in \mathcal{G}_2(\delta_n)$  satisfies

$$\mathbb{E}[f^{2}] \leq \sup_{\|\mathbf{a}\|=1} \sup_{\|\mathbf{b}_{1}-\mathbf{b}_{2}\| \leq \delta_{n}} \mathbb{E}[(\mathbf{a}^{\top}\mathbf{Z})^{2}\mathbf{1}\{|Y-\mathbf{Z}^{\top}\mathbf{b}_{1}| \leq |\mathbf{Z}^{\top}(\mathbf{b}_{1}-\mathbf{b}_{2})|\}]$$

$$\leq \sup_{\mathbf{b}\in\mathbb{R}^{m}} \sup_{\|\mathbf{a}\|=1} \mathbb{E}[(\mathbf{a}^{\top}\mathbf{Z})^{2}\mathbf{1}\{|Y-\mathbf{Z}^{\top}\mathbf{b}| \leq \xi_{m}\delta_{n}\}]$$

$$\leq 2\overline{f}\xi_{m}\delta_{n}\lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}]).$$

Note that under assumptions (A1)-(A3) the right-hand side is bounded by  $c\xi_m\delta_n$  where c is a constant that does not depend on n. Thus the bound in (S.2.2) implies that for  $\xi_m\delta_n\gg n^{-1}$  we have for some constant C which is independent of n,

$$\mathbb{E}\|\mathbb{P}_n - P\|_{\mathcal{G}_1} \le C\left[\left(\frac{m}{n}\log(\xi_m \vee n)\right)^{1/2} + \frac{m\xi_m}{n}\log(\xi_m \vee n)\right],\tag{S.1.12}$$

$$\mathbb{E}\|\mathbb{P}_n - P\|_{\mathcal{G}_2(\delta_n)} \le C \left[ \xi_m^{1/2} \delta_n^{1/2} \left( \frac{m}{n} \log(\xi_m \vee n) \right)^{1/2} + \frac{m\xi_m}{n} \log(\xi_m \vee n) \right].$$
 (S.1.13)

Thus (S.1.9) and (S.1.10) follow from the bound in (S.2.3) by setting  $t = \kappa_n$ .

# S.1.2 Proof of Theorem 5.2

We begin with the following useful decomposition

$$\widehat{\boldsymbol{\beta}}(\tau) - \boldsymbol{\beta}_n(\tau) = -n^{-1/2} \widetilde{J}_m^{-1}(\tau) \mathbb{G}_n(\psi(\cdot; \boldsymbol{\beta}_n(\tau), \tau)) + \widetilde{J}_m^{-1}(\tau) \sum_{k=1}^4 R_{n,k}(\tau)$$
 (S.1.14)

where

$$R_{n,1}(\tau) := \mathbb{P}_n \psi(\cdot; \widehat{\boldsymbol{\beta}}(\tau), \tau),$$

$$R_{n,2}(\tau) := -\left(\mu(\widehat{\boldsymbol{\beta}}(\tau), \tau) - \mu(\boldsymbol{\beta}_n(\tau), \tau) - \widetilde{J}_m(\tau)(\widehat{\boldsymbol{\beta}}(\tau) - \boldsymbol{\beta}_n(\tau))\right),$$

$$R_{n,3}(\tau) := -n^{-1/2} \Big( \mathbb{G}_n(\psi(\cdot; \widehat{\boldsymbol{\beta}}(\tau), \tau)) - \mathbb{G}_n(\psi(\cdot; \boldsymbol{\beta}_n(\tau), \tau)) \Big),$$

$$R_{n,4}(\tau) := -\mu(\boldsymbol{\beta}_n(\tau), \tau).$$

Define  $r_{n,2}(\tau, \mathbf{u}_n) := \mathbf{u}_n^{\top} \widetilde{J}_m^{-1}(\tau) R_{n,2}(\tau), \ r_{n,k}(\tau, \mathbf{u}_n) := (\mathbf{u}_n^{\top} \widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{u}_n, D))} R_{n,k}(\tau) \text{ for } k = 1, 3 \text{ and } r_{n,k}(\tau)$ 

$$r_{n,4}(\tau,\mathbf{u}_n) := \mathbf{u}_n^{\top} \widetilde{J}_m^{-1}(\tau) R_{n,4}(\tau) + \Big( \mathbf{u}_n^{\top} \widetilde{J}_m^{-1}(\tau) - (\mathbf{u}_n^{\top} \widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{u}_n,D))} \Big) \Big( R_{n,1}(\tau) + R_{n,3}(\tau) \Big).$$

With those definitions we obtain

$$\mathbf{u}_n^{\top} (\widehat{\boldsymbol{\beta}}(\tau) - \boldsymbol{\beta}_n(\tau)) = -n^{-1/2} \mathbf{u}_n^{\top} \widetilde{J}_m^{-1}(\tau) \mathbb{G}_n(\psi(\cdot; \boldsymbol{\beta}_n(\tau), \tau)) + \sum_{k=1}^4 r_{n,k}(\tau, \mathbf{u}_n).$$

We will now show that the terms  $r_{n,k}(\tau, \mathbf{u}_n)$  defined above satisfy the bounds given in the statement of Theorem 5.2 if we let  $D = c \log n$  for a sufficiently large constant c.

The bound on  $r_{n,1}$  follows from Lemma S.1.7. To bound  $r_{n,2}$  apply Lemma S.1.4 and Lemma S.1.6. To bound  $r_{n,3}$  observe that by Lemma S.1.4 the probability of the event

$$\Omega_1 := \Big\{ \sup_{\tau, x} |\mathbf{B}(x)^\top \widehat{\boldsymbol{\beta}}(\tau) - \mathbf{B}(x)^\top \boldsymbol{\beta}_n(\tau)| \le C \Big( \widehat{c}_n^2 + \frac{\xi_m (\log n + \kappa_n^{1/2})}{n^{1/2}} \Big) \Big\}.$$

is at least  $1-(m+1)e^{-\kappa_n}$ . Letting  $\delta_n:=C\Big(\widetilde{c}_n^2+\frac{\xi_m(\log n+\kappa_n^{1/2})}{n^{1/2}}\Big)$  we find that on  $\Omega_1$ 

$$\sup_{\mathbf{u}_n \in \mathcal{S}_{\mathcal{I}}^{m-1}} \sup_{\tau} |r_{n,3}(\tau, \mathbf{u}_n)| \lesssim \|\mathbb{P}_n - P\|_{\mathcal{G}_2(\delta_n, \mathcal{I}(\mathbf{u}_n, D), \mathcal{I}'(\mathbf{u}_n, D))},$$

this follows since

$$n^{-1/2}(\mathbf{u}_n^{\top}\widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{u}_n,D))}\mathbb{G}_n(\psi(\cdot;\widehat{\boldsymbol{\beta}}(\tau),\tau)) = n^{-1/2}(\mathbf{u}_n^{\top}\widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{u}_n,D))}\mathbb{G}_n(\psi^{(\mathcal{I}'(\mathbf{u}_n,D))}(\cdot;\widehat{\boldsymbol{\beta}}(\tau),\tau))$$

and a similar identity holds with  $\beta_n$  instead of  $\widehat{\beta}$ . The bound on  $r_{n,3}$  now follows form Lemma S.1.5. To bound the first part of  $r_{n,4}(\tau, \mathbf{u}_n)$ , we proceed as in the proof of equation (S.1.18) to obtain

$$\left| (\mathbf{u}_n^{\top} \widetilde{J}_m^{-1}(\tau)) \mu(\boldsymbol{\beta}_n(\tau), \tau) \right| \leq \frac{1}{2} \overline{f'} \widetilde{c}_n^2 \widetilde{\mathcal{E}}(\mathbf{u}_n, \mathbf{B})$$

where the last line follows after a Taylor expansion. To bound the second part of  $r_{n,4}(\tau, \mathbf{u}_n)$  note that  $\sup_{\tau} \|R_{n,1}(\tau)\| + \|R_{n,3}(\tau)\| \le 3\xi_m$  almost surely and thus choosing  $D = c \log n$  with c sufficiently large yields  $\|(\mathbf{u}_n^{\top} \widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{u}_n,D))} - \mathbf{u}_n^{\top} \widetilde{J}_m^{-1}(\tau)\| \le n^{-1} 3^{-1} \xi_m^{-1}$  where we used (A.5). This completes the proof.

### S.1.2.1 Technical details for the proof of Theorem 5.2

**Lemma S.1.4.** Under the assumptions of Theorem 5.2 we have for sufficiently large n and any  $\kappa_n \ll n/\xi_m^2$ ,

$$P\Big(\sup_{\tau,x}|\mathbf{B}(x)^{\top}(\widehat{\boldsymbol{\beta}}(\tau)-\boldsymbol{\beta}_n(\tau))| \ge C\Big(\frac{\xi_m\log n + \xi_m\kappa_n^{1/2}}{n^{1/2}} + \widetilde{c}_n^2\Big)\Big) \le (m+1)e^{-\kappa_n}.$$

where the constant C does not depend on n.

**Proof of Lemma S.1.4.** Apply (A.5) with  $\mathbf{a} = \mathbf{B}(x)$  to obtain

$$\|\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau) - (\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{B}(x),D))}\| \lesssim m \xi_m \gamma^D$$

where  $\mathcal{I}(\mathbf{B}(x), D)$  is defined as (A.2), and  $\gamma \in (0, 1)$  is a constant independent of n. Next observe the decomposition

$$\mathbf{B}(x)^{\top}(\widehat{\boldsymbol{\beta}}(\tau) - \boldsymbol{\beta}_n(\tau)) = -n^{-1/2}(\mathbf{B}(x)^{\top}\widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{B}(x),D))}\mathbb{G}_n(\psi(\cdot;\widehat{\boldsymbol{\beta}}(\tau),\tau)) + \sum_{k=1}^4 r_{n,k}(\tau,x)$$

where

$$r_{n,1}(\tau,x) := (\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{B}(x),D))} \mathbb{P}_n \psi(\cdot; \widehat{\boldsymbol{\beta}}(\tau),\tau),$$

$$r_{n,2}(\tau,x) := -(\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau)) \Big( \mu(\widehat{\boldsymbol{\beta}}(\tau),\tau) - \mu(\boldsymbol{\beta}_n(\tau),\tau) - \widetilde{J}_m(\tau)(\widehat{\boldsymbol{\beta}}(\tau) - \boldsymbol{\beta}_n(\tau)) \Big),$$

$$r_{n,3}(\tau,x) := -(\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau)) \mu(\boldsymbol{\beta}_n(\tau),\tau).$$

and

$$r_{n,4}(\tau,x) := \left(\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau) - (\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{B}(x),D))}\right) \left(\mathbb{P}_n \psi(\cdot; \widehat{\boldsymbol{\beta}}(\tau),\tau) - n^{-1/2} \mathbb{G}_n(\psi(\cdot; \widehat{\boldsymbol{\beta}}(\tau),\tau))\right).$$

Letting  $D = c \log n$  for a sufficiently large constant c, (S.1.14) yields  $\sup_{\tau,x} |r_{n,4}(\tau,x)| \le n^{-1}$  almost surely. Lemma S.1.7 yields the bound

$$\sup_{x,\tau} |r_{n,1}(\tau,x)| \lesssim \frac{\xi_m^2 \log n}{n} \quad a.s. \tag{S.1.15}$$

Let  $\delta_n := \sup_{x,\tau} |\mathbf{B}(x)^{\top} \widehat{\boldsymbol{\beta}}(\tau) - \mathbf{B}(x)^{\top} \boldsymbol{\beta}_n(\tau)|$ . From Lemma S.1.6 we obtain under (L)

$$\sup_{x,\tau} |r_{n,2}(\tau,x)| \le \delta_n^2 \sup_{x,\tau} \mathbb{E}\left[|\mathbf{B}(x)^\top \widetilde{J}_m^{-1}(\tau)\mathbf{B}|\right] \times \delta_n^2, \tag{S.1.16}$$

where  $\sup_{x,\tau} \mathbb{E}[|\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau) \mathbf{B}|] = O(1)$  by assumption (L). Finally, note that

$$n^{-1/2}(\mathbf{B}(x)^{\top}\widetilde{J}_{m}^{-1}(\tau))^{(\mathcal{I}(\mathbf{B}(x),D))}\mathbb{G}_{n}(\psi(\cdot;\widehat{\boldsymbol{\beta}}(\tau),\tau))$$

$$= n^{-1/2}(\mathbf{B}(x)^{\top}\widetilde{J}_{m}^{-1}(\tau))^{(\mathcal{I}(\mathbf{B}(x),D))}\mathbb{G}_{n}(\psi^{(\mathcal{I}'(\mathbf{B}(x),D))}(\cdot;\widehat{\boldsymbol{\beta}}(\tau),\tau))$$

where  $\mathcal{I}'(\mathbf{B}(x), D)$  is defined as (A.3). This yields

$$\sup_{\tau,x} \left| n^{-1/2} (\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{B}(x),D))} \mathbb{G}_n(\psi(\cdot;\widehat{\boldsymbol{\beta}}(\tau),\tau)) \right| \lesssim \xi_m \sup_{x \in \mathcal{X}} \|\mathbb{P}_n - P\|_{\mathcal{G}_1(\mathcal{I}(\mathbf{B}(x),D),\mathcal{I}'(\mathbf{B}(x),D))}$$

By the definition of  $\mathcal{I}(\mathbf{B}(x), D)$ ,  $\mathcal{I}'(\mathbf{B}(x), D)$ , the supremum above ranges over at most m distinct terms. Additionally,  $\sup_{x \in \mathcal{X}} |\mathcal{I}(\mathbf{B}(x), c \log n)| + |\mathcal{I}'(\mathbf{B}(x), c \log n)| \lesssim \log n$ . Thus Lemma S.1.5 yields

$$P\Big(\sup_{\tau,x} \Big| n^{-1/2} (\mathbf{B}(x)^{\top} \widetilde{J}_m^{-1}(\tau))^{(\mathcal{I}(\mathbf{B}(x),c\log n))} \mathbb{G}_n(\psi(\cdot;\widehat{\boldsymbol{\beta}}(\tau),\tau)) \Big| \ge C\Big(\frac{\xi_m^2 (\log n)^2}{n}\Big)^{1/2} + C\frac{\xi_m \kappa_n^{1/2}}{n^{1/2}}\Big) \le me^{-\kappa_n}.$$
(S.1.17)

Finally, from the definition of  $\beta_n$  as minimizer we obtain

$$|r_{n,3}(\tau,x)| = \left| (\mathbf{B}(x)^{\top} \widetilde{J}_{m}^{-1}(\tau)) \mu(\boldsymbol{\beta}_{n}(\tau),\tau) \right|$$

$$= \left| (\mathbf{B}(x)^{\top} \widetilde{J}_{m}^{-1}(\tau)) \mathbb{E} [\mathbf{B} (\mathbf{1} \{ Y \leq \boldsymbol{\beta}_{n}(\tau)^{\top} \mathbf{B} \} - \tau)] \right|$$

$$= \left| (\mathbf{B}(x)^{\top} \widetilde{J}_{m}^{-1}(\tau)) \mathbb{E} [\mathbf{B} (F_{Y|X}(\boldsymbol{\beta}_{n}(\tau)^{\top} \mathbf{B} | X) - F_{Y|X}(Q(X;\tau) | X))] \right|$$

$$\leq \frac{1}{2} \overline{f'} \widetilde{c}_{n}^{2} O(1)$$
(S.1.18)

where the last line follows after a Taylor expansion and the fact that  $\mathbb{E}[\mathbf{B}f_{Y|X}(Q(X;\tau)|X)(\boldsymbol{\beta}_n^{\top}\mathbf{B} - Q(X;\tau))] = 0$  from the definition of  $\boldsymbol{\beta}_n$  and making use of (L). Combining this with (S.1.15) - (S.1.17) yields

$$\delta_n \le C \left[ \left( \frac{\xi_m^2 (\log n)^2}{n} \right)^{1/2} + \frac{\xi_m \kappa_n^{1/2}}{n^{1/2}} + \frac{\xi_m^2 \log n}{n} + \delta_n^2 + \widetilde{c}_n^2 \right]$$

with probability at least  $me^{-\kappa_n}$ . By Lemma S.1.2 we have  $P(\delta_n \ge 1/(2C)) \le e^{-\kappa_n}$  for any  $\kappa_n$  satisfying  $\xi_m \kappa_n \gg n^{-1}$  This yields the assertion.

**Lemma S.1.5.** Let  $\mathcal{Z} := \{\mathbf{B}(x) | x \in \mathcal{X}\}$  where  $\mathcal{X}$  is the support of X. For  $\mathcal{I}_1, \mathcal{I}'_1 \subset \{1, ..., m\}$ , define the classes of functions

$$\widetilde{\mathcal{G}}_{1}(\mathcal{I}_{1}, \mathcal{I}'_{1}) := \left\{ (\mathbf{Z}, Y) \mapsto \mathbf{a}^{\top} \mathbf{Z}^{(\mathcal{I}_{1})} (\mathbf{1} \{ Y \leq \mathbf{Z}^{\top} \mathbf{b}^{(\mathcal{I}'_{1})} \} - \tau) \mathbf{1} \{ \| \mathbf{Z} \| \leq \xi_{m} \} \middle| \tau \in \mathcal{T}, \mathbf{b} \in \mathbb{R}^{m}, \mathbf{a} \in \mathcal{S}^{m-1} \right\},$$
(S.1.19)

$$\widetilde{\mathcal{G}}_{2}(\delta, \mathcal{I}_{1}, \mathcal{I}'_{1}) := \left\{ (\mathbf{Z}, Y) \mapsto \mathbf{a}^{\top} \mathbf{Z}^{(\mathcal{I}_{1})} (\mathbf{1} \{ Y \leq \mathbf{Z}^{\top} \mathbf{b}_{1}^{(\mathcal{I}'_{1})} \} - \mathbf{1} \{ Y \leq \mathbf{Z}^{\top} \mathbf{b}_{2}^{(\mathcal{I}'_{1})} \}) \mathbf{1} \{ \mathbf{Z} \in \mathcal{Z} \} \right|$$

$$\mathbf{b}_{1}, \mathbf{b}_{2} \in \mathbb{R}^{m}, \sup_{\mathbf{v} \in \mathcal{Z}} \| \mathbf{v}^{\top} \mathbf{b}_{1} - \mathbf{v}^{\top} \mathbf{b}_{2} \| \leq \delta, \mathbf{a} \in \mathcal{S}^{m-1} \right\}. \quad (S.1.20)$$

Under assumptions (A1)-(A3) we have

$$P\Big(\|\mathbb{P}_n - P\|_{\widetilde{\mathcal{G}}_1(\mathcal{I}_1, \mathcal{I}_1')} \ge C\Big[\Big(\frac{\max(|\mathcal{I}_1|, |\mathcal{I}_1'|)}{n}\log \xi_m\Big)^{1/2} + \frac{\max(|\mathcal{I}_1|, |\mathcal{I}_1'|)\xi_m}{n}\log \xi_m + \Big(\frac{\kappa_n}{n}\Big)^{1/2} + \frac{\xi_m \kappa_n}{n}\Big]\Big) \le e^{-\kappa_n}$$
(S.1.21)

and for any  $\delta_n$  satisfying  $\xi_m \delta_n \gg n^{-1}$  we have for sufficiently large n and arbitrary  $\kappa_n > 0$ 

$$P\Big(\|\mathbb{P}_n - P\|_{\widetilde{\mathcal{G}}_2(\delta, \mathcal{I}_1, \mathcal{I}_1')} \ge C\zeta_n(\delta, \mathcal{I}_1, \mathcal{I}_1', \kappa_n)\Big) \le e^{-\kappa_n},\tag{S.1.22}$$

where

$$\zeta_n(t, \mathcal{I}_1, \mathcal{I}'_1, \kappa_n) := t^{1/2} \left( \frac{\max(|\mathcal{I}_1|, |\mathcal{I}'_1|)}{n} \log(\xi_m \vee n) \right)^{1/2} + \frac{\max(|\mathcal{I}_1|, |\mathcal{I}'_1|) \xi_m}{n} \log(\xi_m \vee n) + n^{-1/2} (t \kappa_n)^{1/2} + n^{-1} \xi_m \kappa_n.$$

**Proof of Lemma S.1.5.** We begin by observing that

$$N(\widetilde{\mathcal{G}}_2(\delta, \mathcal{I}_1, \mathcal{I}'_1), L_2(\mathbb{P}_n); \varepsilon) \leq \left(\frac{A\|F\|_{L^2(\mathbb{P}_n)}}{\varepsilon}\right)^{v_1(m)}, \quad N(\widetilde{\mathcal{G}}_1(\mathcal{I}_1, \mathcal{I}'_1), L_2(\mathbb{P}_n); \varepsilon) \leq \left(\frac{A\|F\|_{L^2(\mathbb{P}_n)}}{\varepsilon}\right)^{v_2(m)},$$

where  $v_1(m) = O(\max(|\mathcal{I}_1|, |\mathcal{I}'_1|)), v_2(m) = O(\max(|\mathcal{I}_1|, |\mathcal{I}'_1|))$ . The proof of the bound for  $\widetilde{\mathcal{G}}_1(\mathcal{I}_1, \mathcal{I}'_1)$  now follows by similar arguments as the proof of Lemma S.1.3. For a proof of the second part, note that for  $f \in \widetilde{\mathcal{G}}_2$  we have

$$\begin{split} \mathbb{E}[f^2] &\leq \sup_{\|\mathbf{a}\|=1} \sup_{\mathbf{b}_1, \mathbf{b}_2 \in R(\delta_n)} \mathbb{E}\Big[ (\mathbf{a}^{\top} \mathbf{B}^{(\mathcal{I}_1)})^2 \mathbf{1} \Big\{ |Y - \mathbf{B}^{\top} \mathbf{b}_1^{(\mathcal{I}_1')}| \leq |\mathbf{B}^{\top} (\mathbf{b}_1^{(\mathcal{I}_1')} - \mathbf{b}_2^{(\mathcal{I}_1')})| \Big\} \Big] \\ &\leq \sup_{\mathbf{b} \in \mathbb{R}^m} \sup_{\|\mathbf{a}\|=1} \mathbb{E}\Big[ \Big( (\mathbf{a}^{(\mathcal{I}_1')})^{\top} \mathbf{B} \mathbf{B}^{\top} \mathbf{a}^{(\mathcal{I}_1')} \Big)^2 \mathbf{1} \{ |Y - \mathbf{B}^{\top} \mathbf{b}^{(\mathcal{I}_1')}| \leq \delta_n \} \Big] \\ &\leq 2\overline{f} \delta_n \lambda_{\max}(\mathbb{E}[\mathbf{B} \mathbf{B}^{\top}]) \end{split}$$

where we defined  $R(\delta) := \left\{ \mathbf{b}_1, \mathbf{b}_2 \in \mathbb{R}^m, \sup_{\mathbf{v} \in \mathcal{Z}} \|\mathbf{v}^\top \mathbf{b}_1 - \mathbf{v}^\top \mathbf{b}_2\| \le \delta \right\}$ . The rest of the proof follows by similar arguments as the proof of Lemma S.1.3.

**Lemma S.1.6.** Under assumptions (A1)-(A3) we have for any  $\mathbf{a}, \mathbf{b} \in \mathbb{R}^m$ 

$$\left| \mathbf{a}^{\top} \widetilde{J}_{m}(\tau)^{-1} \mu(\mathbf{b}, \tau) - \mathbf{a}^{\top} \widetilde{J}_{m}(\tau)^{-1} \mu(\boldsymbol{\beta}_{n}(\tau), \tau) - \mathbf{a}^{\top} (\mathbf{b} - \boldsymbol{\beta}_{n}(\tau)) \right|$$

$$\leq \overline{f'} \sup_{x} |\mathbf{B}(x)^{\top} \mathbf{b} - \mathbf{B}(x)^{\top} \boldsymbol{\beta}_{n}(\tau)|^{2} \mathbb{E}[|\mathbf{a}^{\top} \widetilde{J}_{m}(\tau)^{-1} \mathbf{B}|].$$

**Proof of Lemma S.1.6.** Note that  $\mu'(\beta_n(\tau), \tau) = \mathbb{E}[\mathbf{B}\mathbf{B}^{\top} f_{Y|X}(\mathbf{B}^{\top} \beta_n(\tau)|X)] = \widetilde{J}_m(\tau)$ . Additionally, we have

$$\mu(\mathbf{b}, \tau) = \mu(\beta_n(\tau), \tau) + \mu'(\bar{\mathbf{b}}, \tau)(\mathbf{b} - \beta_n(\tau)),$$

where  $\bar{\mathbf{b}} = \mathbf{b} + \lambda_{\mathbf{b},\tau}(\boldsymbol{\beta}_n(\tau) - \mathbf{b})$  for some  $\lambda_{\mathbf{b},\tau} \in [0,1]$ . Moreover,

$$\mathbf{a}^{\top}[\widetilde{J}_m(\tau)^{-1}\mu(\mathbf{b},\tau)-\widetilde{J}_m(\tau)^{-1}\mu(\boldsymbol{\beta}_n(\tau),\tau)-(\mathbf{b}-\boldsymbol{\beta}_n(\tau))]=\mathbf{a}^{\top}\widetilde{J}_m(\tau)^{-1}[\mu'(\bar{\mathbf{b}},\tau)-\widetilde{J}_m(\tau)](\mathbf{b}-\boldsymbol{\beta}_n(\tau))$$

and thus

$$\begin{aligned} & \left| \mathbf{a}^{\top} \widetilde{J}_{m}(\tau)^{-1} \mu(\mathbf{b}, \tau) - \mathbf{a}^{\top} \widetilde{J}_{m}(\tau)^{-1} \mu(\boldsymbol{\beta}_{n}(\tau), \tau) - \mathbf{a}^{\top} (\mathbf{b} - \boldsymbol{\beta}_{n}(\tau)) \right| \\ &= \left| \mathbb{E} \left[ \mathbf{a}^{\top} \widetilde{J}_{m}(\tau)^{-1} \mathbf{B} \mathbf{B}^{\top} \left( f_{Y|X} (\mathbf{B}^{\top} \bar{\mathbf{b}} | X) - f_{Y|X} (\mathbf{B}^{\top} \boldsymbol{\beta}_{n}(\tau) | X) \right) \left( \mathbf{b} - \boldsymbol{\beta}_{n}(\tau) \right) \right] \right| \\ &\leq \overline{f'} \mathbb{E} \left[ \left| \mathbf{a}^{\top} \widetilde{J}_{m}(\tau)^{-1} \mathbf{B} \right| \left\{ \mathbf{B}^{\top} (\mathbf{b} - \boldsymbol{\beta}_{n}(\tau)) \right\}^{2} \right] \\ &\leq \overline{f'} \sup_{x} |\mathbf{B}(x)^{\top} \mathbf{b} - \mathbf{B}(x)^{\top} \boldsymbol{\beta}_{n}(\tau) |^{2} \mathbb{E} [|\mathbf{a}^{\top} \widetilde{J}_{m}(\tau)^{-1} \mathbf{B}|]. \end{aligned}$$

**Lemma S.1.7.** Under assumptions (A1)-(A3) and (L) we have for any  $\mathbf{a} \in \mathbb{R}^m$  having zero entries everywhere except at L consecutive positions:

$$\left|\mathbf{a}^{\top} \mathbb{P}_n \psi(\cdot; \widehat{\boldsymbol{\beta}}(\tau), \tau)\right| \leq \frac{(L+2r) \|\mathbf{a}\| \xi_m}{n}.$$

**Proof of Lemma S.1.7.** From standard arguments of the optimization condition of quantile regression (p.35 of Koenker (2005), also see equation (2.2) on p.224 of Knight (2008)), we know that for any  $\tau \in \mathcal{T}$ ,

$$\mathbb{P}_n \psi(\cdot; \widehat{\boldsymbol{\beta}}(\tau), \tau) = \frac{1}{n} \sum_{i=1}^n \mathbf{B}_i (\mathbf{1}\{Y_i \leq \mathbf{B}_i^{\top} \widehat{\boldsymbol{\beta}}(\tau)\} - \tau) = \frac{1}{n} \sum_{i \in H_{\tau}} v_i \mathbf{B}_i$$

where  $v_i \in [-1,1]$  and  $H_{\tau} = \{i : Y_i = \mathbf{B}_i^{\top} \widehat{\boldsymbol{\beta}}(\tau)\}$ . Since  $\mathbf{a}$  has at most L non-zero entries, the dimension of the subspace spanned by  $\{\mathbf{B}_i : \mathbf{a}^{\top} \mathbf{B}_i \neq 0\}$  is at most L + 2r [each vector  $\mathbf{B}_i$  by construction has at most r nonzero entries and all of those entries are consecutive]. Since the conditional distribution of Y given covariates has a density, the data are in general position almost surely, i.e. no more than k of the points  $(\mathbf{B}_i, Y_i)$  lie in any k-dimensional linear space, it follows that the cardinality of the set  $H_{\tau} \cap \{i : \mathbf{a}^{\top} \mathbf{B}_i \neq 0\}$  is bounded by L + 2r. The assertion follows after an elementary calculation.

#### S.1.3 Proof of Theorem 5.4

The statement follows from Theorem 5.1 if we prove that the vector  $\boldsymbol{\gamma}_n^{\dagger}(\tau)$  satisfies

$$\sup_{\tau \in \mathcal{T}} \|\mu(\gamma_n^{\dagger}(\tau); \tau)\| = O(\xi_m c_n^{\dagger 2})$$
(S.1.23)

as  $c_n^{\dagger} = o(\xi_m^{-1})$  in Condition (C1), and establish the identity in (5.10). For the identity (5.10), we first observe the representation

$$J_m(\tau) = \begin{pmatrix} M_1(\tau) + A(\tau)M_2(\tau)A(\tau)^\top & A(\tau)M_2(\tau) \\ M_2(\tau)A(\tau)^\top & M_2(\tau) \end{pmatrix},$$
 (S.1.24)

which follows from (3.5) and

$$\mathbb{E}[(V - A(\tau)\widetilde{\mathbf{Z}}(W))\widetilde{\mathbf{Z}}(W)^{\top} f_{Y|X}(Q(X;\tau)|X)] = 0, \text{ for all } \tau \in \mathcal{T}.$$

To simplify the notations, we suppress the argument in  $\tau$  in the following matrix calculations. Recall the following identity for the inverse of  $2 \times 2$  block matrix (see equation (6.0.8) on p.165 of Puntanen and Styan (2005))

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix}^{-1} = \begin{pmatrix} (A - BD^{-1}C)^{-1} & -(A - BD^{-1}C)^{-1}BD^{-1} \\ -D^{-1}C(A - BD^{-1}C)^{-1} & D^{-1} + D^{-1}C(A - BD^{-1}C)^{-1}BD^{-1} \end{pmatrix}.$$

Identifying the blocks in the representation (S.1.24) with the blocks in te above representation yields the result after some simple calculations. For a proof of (S.1.23) observe that

$$\mu(\boldsymbol{\gamma}_n^{\dagger}(\tau);\tau) = \mathbb{E}[(\boldsymbol{V}^{\top}, \widetilde{\mathbf{Z}}(\boldsymbol{W})^{\top})^{\top}(F_{Y|X}(\boldsymbol{\gamma}_n^{\dagger}(\tau)^{\top}\mathbf{Z}|\boldsymbol{X}) - \tau)]$$

Now on one hand we have, uniformly in  $\tau \in \mathcal{T}$ ,

$$\begin{split} & \left\| \mathbb{E}[V(F_{Y|X}(\boldsymbol{\alpha}(\tau)^{\top}V + \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W)|X) - \tau)] \right\| \\ &= \left\| \mathbb{E}[Vf_{Y|X}(Q(X;\tau)|X)(Q(X;\tau) - \boldsymbol{\alpha}(\tau)^{\top}V - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W))] \right\| + O(c_{n}^{\dagger 2}) \\ &= \left\| \mathbb{E}[Vf_{Y|X}(Q(X;\tau)|X)(h(W;\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W))] \right\| + O(c_{n}^{\dagger 2}) \\ &= \left\| \mathbb{E}[(V - h_{VW}(W;\tau) + h_{VW}(W;\tau))f_{Y|X}(Q(X;\tau)|X)(h(W;\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W))] \right\| + O(c_{n}^{\dagger 2}) \\ &= \left\| \mathbb{E}[h_{VW}(W;\tau)f_{Y|X}(Q(X;\tau)|X)(h(W;\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W))] \right\| + O(c_{n}^{\dagger 2}) \\ &= \left\| \mathbb{E}[(h_{VW}(W;\tau) - A(\tau)\widetilde{\mathbf{Z}}(W) + A(\tau)\widetilde{\mathbf{Z}}(W))f_{Y|X}(Q(X;\tau)|X)(h(W;\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W))] \right\| + O(c_{n}^{\dagger 2}) \\ &= \left\| \mathbb{E}[(h_{VW}(W;\tau) - A(\tau)\widetilde{\mathbf{Z}}(W))f_{Y|X}(Q(X;\tau)|X)(h(W;\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W))] \right\| + O(c_{n}^{\dagger 2}) \\ &= \left\| \mathbb{E}[(h_{VW}(W;\tau) - A(\tau)\widetilde{\mathbf{Z}}(W))f_{Y|X}(Q(X;\tau)|X)(h(W;\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W))] \right\| + O(c_{n}^{\dagger 2}) \\ &= O(c_{n}^{\dagger 2} + \lambda_{n}c_{n}^{\dagger}). \end{split}$$

Here, the first equation follows after a Taylor expansion taking into account that, by the definition of the conditional quantile function,  $F_{Y|X}(Q(X;\tau)|X) \equiv \tau$ . The fourth equality is a consequence of (3.5), the sixth equality follows since

$$\mathbb{E}[\widetilde{\mathbf{Z}}(W)f_{Y|X}(Q(X;\tau)|X)(h(W;\tau) - \boldsymbol{\beta}_n^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W))] = 0$$
 (S.1.25)

by the definition of  $\beta_n^{\dagger}(\tau)$  as minimizer. The last line follows by the Cauchy-Schwarz inequality. On the other hand

$$\mathbb{E}[\widetilde{\mathbf{Z}}(W)(F_{Y|X}(\boldsymbol{\alpha}(\tau)^{\top}V + \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W)|X) - \tau)]$$

$$= \mathbb{E}[\widetilde{\mathbf{Z}}(W)f_{Y|X}(Q(X;\tau)|X)(h(W;\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W))]$$

$$+ \frac{1}{2}\mathbb{E}[\widetilde{\mathbf{Z}}(W)f_{Y|X}'(\zeta(X;\tau)|X)(h(W;\tau) - \boldsymbol{\beta}_{n}^{\dagger}(\tau)^{\top}\widetilde{\mathbf{Z}}(W))^{2}].$$

By (S.1.25), the first term in the representation above is zero, and the norm of the second term is of the order  $O(\xi_m c_n^{\dagger 2})$ . This completes the proof.

# APPENDIX S.2: Auxiliary Results

# S.2.1 Results on empirical process theory

In this section, we collect some basic results from empirical process theory needed in our proofs. Denote by  $\mathcal{G}$  a class of functions that satisfies  $|f(x)| \leq F(x) \leq U$  for every  $f \in \mathcal{G}$  and let  $\sigma^2 \geq \sup_{f \in \mathcal{G}} Pf^2$ . Additionally, let for some A > 0, V > 0 and all  $\varepsilon > 0$ ,

$$N(\varepsilon, \mathcal{G}, L_2(\mathbb{P}_n)) \le \left(\frac{A\|F\|_{L^2(\mathbb{P}_n)}}{\varepsilon}\right)^V.$$
 (S.2.1)

Note that if  $\mathcal{G}$  is a VC-class, then V is the VC-index of the set of subgraphs of functions in  $\mathcal{G}$ . In that case, the symmetrization inequality and inequality (2.2) from Koltchinskii (2006) yield

$$\mathbb{E}\|\mathbb{P}_n - P\|_{\mathcal{G}} \le c_0 \left[\sigma \left(\frac{V}{n} \log \frac{A\|F\|_{L^2(P)}}{\sigma}\right)^{1/2} + \frac{VU}{n} \log \frac{A\|F\|_{L^2(P)}}{\sigma}\right]$$
(S.2.2)

for a universal constant  $c_0 > 0$  provided that  $1 \ge \sigma^2 > \text{const} \times n^{-1}$  [in fact, the inequality in Koltchinskii (2006) is for  $\sigma^2 = \sup_{f \in \mathcal{G}} Pf^2$ . However, this is not a problem since we can replace  $\mathcal{G}$  by  $\mathcal{G}\sigma/(\sup_{f \in \mathcal{G}} Pf^2)^{1/2}$ ]. The second inequality (a refined version of Talagrand's concentration inequality) states that for any countable class of measurable functions  $\mathcal{F}$  with elements mapping into [-M, M]

$$P\Big\{\|\mathbb{P}_n - P\|_{\mathcal{F}} \ge 2\mathbb{E}\|\mathbb{P}_n - P\|_{\mathcal{F}} + c_1 n^{-1/2} \Big(\sup_{f \in \mathcal{F}} Pf^2\Big)^{1/2} \sqrt{t} + n^{-1} c_2 Mt\Big\} \le e^{-t},$$
 (S.2.3)

for all t > 0 and universal constants  $c_1, c_2 > 0$ . This is a special case of Theorem 3 in Massart (2000) [in the notation of that paper, set  $\varepsilon = 1$ ].

**Lemma S.2.1** (Lemma 7.1 of Kley et al. (2015)). Let  $\{\mathbb{G}_t : t \in T\}$  be a separable stochastic process with  $\|\mathbb{G}_s - \mathbb{G}_t\|_{\Psi} \leq Cd(s,t)$  ( $\|\cdot\|_{\Psi}$  is defined in (A.32)) for all s,t satisfying  $d(s,t) \geq \bar{\omega}/2 \geq 0$ . Denote by  $D(\epsilon,d)$  the packing number of the metric space (T,d). Then, for any  $\delta > 0$ ,  $\omega \geq \bar{\omega}$ , there exists a random variable  $S_1$  and a constant  $K < \infty$  such that

$$\sup_{d(s,t)\leq\delta} |\mathbb{G}_s - \mathbb{G}_t| \leq S_1 + 2 \sup_{d(s,t)\leq\bar{\omega}, t\in\widetilde{T}} |\mathbb{G}_s - \mathbb{G}_t|, \tag{S.2.4}$$

where the set  $\widetilde{T}$  contains at most  $D(\bar{\omega}, d)$  points, and  $S_1$  satisfies

$$||S_1||_{\Psi} \le K \left[ \int_{\bar{\omega}/2}^{\omega} \Psi^{-1} \left( D(\epsilon, d) \right) d\epsilon + (\delta + 2\bar{\omega}) \Psi^{-1} \left( D^2(\omega, d) \right) \right]$$
(S.2.5)

$$P(|S_1| > x) \le \left(\Psi \left\{ x \left[ 8K \left( \int_{\bar{\omega}/2}^{\omega} \Psi^{-1} \left( D(\epsilon, d) \right) d\epsilon + (\delta + 2\bar{\omega}) \Psi^{-1} \left( D^2(\omega, d) \right) \right) \right]^{-1} \right\} \right)^{-1}.$$
 (S.2.6)

# S.2.2 Covering number calculation

A few useful lemmas on covering number are given in this section.

**Lemma S.2.2.** Suppose  $\mathcal{F}$  and  $\mathcal{G}$  are two function classes with envelopes F and G.

1. The class  $\mathcal{F} - \mathcal{G} := \{f - g | f \in \mathcal{F}, g \in \mathcal{G}\}$  with envelope F + G and

$$\sup_{Q} N(\epsilon \| F + G \|_{Q,2}, \mathcal{F} - \mathcal{G}, L_2(Q)) \le \sup_{Q} N\left(\epsilon \frac{\| F \|_{Q,2}}{\sqrt{2}}, \mathcal{F}, L_2(Q)\right) \sup_{Q} N\left(\epsilon \frac{\| G \|_{Q,2}}{\sqrt{2}}, \mathcal{G}, L_2(Q)\right), \tag{S.2.7}$$

2. The class  $\mathcal{F} \cdot \mathcal{G} := \{fg : f \in \mathcal{F}, g \in \mathcal{G}\}$  with envelope FG and

$$\sup_{Q} N\left(\epsilon \|FG\|_{Q,2}, \mathcal{F} \cdot \mathcal{G}, L_2(Q)\right) \le \sup_{Q} N\left(\frac{\epsilon \|F\|_{Q,2}}{2}, \mathcal{F}, L_2(Q)\right) \sup_{Q} N\left(\frac{\epsilon \|G\|_{Q,2}}{2}, \mathcal{G}, L_2(Q)\right), \tag{S.2.8}$$

where the suprema are taken over the appropriate subsets of all finitely discrete probability measures Q.

# Proof of Lemma S.2.2.

1. It is obvious that  $|f - g| \le |f| + |g| \le F + G$  for any  $f \in \mathcal{F}$  and  $g \in \mathcal{G}$ . Hence, F + G is an envelop for the function class  $\mathcal{F} - \mathcal{G}$ . Suppose that  $A = \{f_1, ..., f_J\}$  and  $B = \{g_1, ..., g_K\}$  are the centers of  $\frac{\epsilon ||F||_{Q,2}}{\sqrt{2}}$ -net for  $\mathcal{F}$  and  $\frac{\epsilon ||G||_{Q,2}}{\sqrt{2}}$ -net for  $\mathcal{F}$  and  $\mathcal{G}$  respectively. For any f - g, there exists  $f_j$  and  $g_k$  such that

$$||(f_j - g_k) - (f - g)||_{Q,2}^2 = ||(f_j - f) - (g_k - g)||_{Q,2}^2 \le 2(||f_j - f||_{Q,2}^2 + ||g_k - g||_{Q,2}^2)$$
  
$$\le \epsilon^2 (||F||_{Q,2}^2 + ||G||_{Q,2}^2) \le \epsilon^2 ||F + G||_{Q,2}^2,$$

where the last inequality follows from the fact that both F and G are nonnegative. Hence,  $\{f_j+g_k: 1 \leq j \leq J, 1 \leq k \leq K\}$  forms an  $\epsilon ||F+G||_{Q,2}$ -net for the class  $\mathcal{F}-\mathcal{G}$ , with cardinality JK.

2. See Lemma 6 of Belloni et al. (2011).

For any fixed vector  $\mathbf{u} \in \mathbb{R}^m$  and  $\delta > 0$ , recall the function classes

$$\mathcal{G}_{3}(\mathbf{u}) = \left\{ (\mathbf{Z}, Y) \mapsto \mathbf{u}^{\top} J_{m}(\tau)^{-1} \mathbf{Z} \middle| \tau \in \mathcal{T} \right\},$$

$$\mathcal{G}_{4} = \left\{ (X, Y) \mapsto \mathbf{1} \left\{ Y_{i} \leq Q(X; \tau) \right\} - \tau \middle| \tau \in \mathcal{T} \right\},$$

$$\mathcal{G}_{6}(\mathbf{u}, \delta) = \left\{ (\mathbf{Z}, Y) \mapsto \mathbf{u}^{\top} \left\{ J_{m}(\tau_{1})^{-1} - J_{m}(\tau_{2})^{-1} \right\} \mathbf{Z} \middle| \tau_{1}, \tau_{2} \in \mathcal{T}, |\tau_{1} - \tau_{2}| \leq \delta \right\},$$

$$\mathcal{G}_{7}(\delta) = \left\{ (X, Y) \mapsto \mathbf{1} \left\{ Y_{i} \leq Q(X, \tau_{1}) \right\} - \mathbf{1} \left\{ Y_{i} \leq Q(X, \tau_{2}) \right\} - (\tau_{1} - \tau_{2}) \middle| \tau_{1}, \tau_{2} \in \mathcal{T}, |\tau_{1} - \tau_{2}| \leq \delta \right\}.$$

Recall the following Lipschitz continuity property of  $I^{-1}(\tau)$  by Lemma 13 of Belloni et al. (2011)

Recall the following Lipschitz continuity property of  $J_m^{-1}(\tau)$  by Lemma 13 of Belloni et al. (2011): for  $\tau_1, \tau_2 \in \mathcal{T}$ ,

$$||J_{m}^{-1}(\tau_{1}) - J_{m}^{-1}(\tau_{2})|| \leq \frac{\bar{f}'}{f_{\min}} |\tau_{1} - \tau_{2}| \Big( \inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_{m}(\tau)) \Big)^{-2} \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}])$$

$$:= C_{0} \inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_{m}(\tau))^{-1} |\tau_{1} - \tau_{2}|.$$
(S.2.9)

where  $C_0 = \frac{\bar{f}'}{f_{\min}} \frac{\lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}])}{\inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_m(\tau)}$ .

**Lemma S.2.3.**  $\mathcal{G}_3(\mathbf{u})$  has an envelope  $G_3(\mathbf{Z}) = \|\mathbf{u}\| \xi_m [\inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_m(\tau))]^{-1}$  and

$$N(\epsilon ||G_3||_{L_2(Q)}, \mathcal{G}_3(\mathbf{u}), L_2(Q)) \le \frac{C_0}{\epsilon},$$

where  $C_0 = \frac{\bar{f}'}{f_{\min}} \frac{\lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^\top])}{\inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_m(\tau))} < \infty$ , for any probability measure Q and x.

**Proof of Lemma S.2.3.** By (S.2.9), for any  $\tau_1, \tau_2 \in \mathcal{T}$  and  $\mathbf{u}$ ,

$$\begin{aligned} \left| \mathbf{u}^{\top} J_{m}(\tau_{1})^{-1} \mathbf{Z} - \mathbf{u}^{\top} J_{m}(\tau_{2})^{-1} \mathbf{Z} \right| &\leq \|\mathbf{u}\| \xi_{m} \frac{\bar{f}'}{f_{\min}} \lambda_{\max}(\mathbb{E}[\mathbf{Z}\mathbf{Z}^{\top}]) [\inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_{m}(\tau))]^{-2} |\tau_{1} - \tau_{2}| \\ &= C_{0} \|G_{3}\|_{L_{2}(Q)} |\tau_{1} - \tau_{2}|. \end{aligned}$$

Applying the relation of the covering and bracketing number on p.84 and Theorem 2.7.11 of van der Vaart and Wellner (1996) yields for each  $\mathbf{u}$  and any probability measure Q,

$$N(\epsilon \|G_3\|_{L_2(Q)}, \mathcal{G}_3(\mathbf{u}), L_2(Q)) \leq N_{[\ ]}(2\epsilon \|G_3\|_{L_2(Q)}, \mathcal{G}_3(\mathbf{u}), L_2(Q)) \leq N(\epsilon, \mathcal{T}, |\cdot|) \leq \frac{C_0}{\epsilon}.$$

**Lemma S.2.4.** 1.  $\mathcal{G}_4$  is a VC-class with VC index 2.

2. The envelopes for  $\mathcal{G}_6(\mathbf{u}, \delta)$  and  $\mathcal{G}_7$  are  $G_6 = \xi_m^2 [\inf_{\tau \in \mathcal{T}} \lambda_{\min}(J_m(\tau))]^{-1} C_0 \delta$  and  $G_7 = 2$ . Furthermore, it holds for any fixed x and  $\delta \leq |\mathcal{T}|$  that

$$N(\epsilon ||G_6||_{L_2(Q)}, \mathcal{G}_6(\mathbf{u}, \delta), L_2(Q)) \le 2\left(\frac{C_0}{\epsilon}\right)^2, \tag{S.2.10}$$

$$N(\epsilon \|G_7\|_{L_2(Q)}, \mathcal{G}_7(\delta), L_2(Q)) \le \left(\frac{A_7}{\epsilon}\right)^2,$$
 (S.2.11)

where  $A_7$  is a universal constant and Q is an arbitrary probability measure.

# Proof of Lemma S.2.4.

- 1. Due to the fact that  $Q(X;\tau)$  is monotone in  $\tau$ , it can be argued with basic VC subgraph argument that  $\mathcal{G}_4$  has VC index 2, under the definition given in p.135 of van der Vaart and Wellner (1996).
- 2. By (S.2.9), the envelope for  $\mathcal{G}_6(\mathbf{u}_n, \delta)$  is  $\|\mathbf{u}\| \xi_m \lambda_{\min}(J_m(\tau))^{-1} C_0 \delta$ . The envelope for  $\mathcal{G}_7(\delta)$  is obvious. By the fact that  $\mathcal{G}_7(\delta) \subset \mathcal{G}_4 \mathcal{G}_4$  and the covering number of  $\mathcal{G}_4$  (implied by Theorem 2.6.7 of van der Vaart and Wellner (1996)), (S.2.11) thus follows by (S.2.7) of Lemma S.2.2. As for (S.2.10), we note that  $\mathcal{G}_6(\mathbf{u}, \delta) \subset \mathcal{G}_3(\mathbf{u}) \mathcal{G}_3(\mathbf{u})$ . Then, (S.2.7) of Lemma S.2.2 and Lemma S.2.3 imply

$$N(\epsilon \|G_6\|_{L_2(Q)}, \mathcal{G}_6(\mathbf{u}, \delta), L_2(Q)) \le N\left(\frac{\epsilon}{\sqrt{2}} \|G_3\|_{L_2(Q)}, \mathcal{G}_3(\mathbf{u}), L_2(Q)\right)^2 \le 2\left(\frac{C_0}{\epsilon}\right)^2,$$

where Q is an arbitrary probability measure.