

# Debiased Bayesian inference for average treatment effects

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## Motivation

Estimating a **causal effect** from observational data is a common problem, e.g. in healthcare.

If treatment assignment is **independent** of outcome (e.g. a randomized controlled trial), can use standard methods. However, often **not** the case.

**Goal:** estimate **population average treatment effect** from **observational data**, e.g. to decide whether to recommend a new treatment or policy.

Two major difficulties:

- **missing counterfactual** outcomes,
- **selection bias** in treatment assignment.

Bayesian (and other) methods can be **badly biased**, especially in complex models or with high-dimensional features (see picture).

## Causal inference model

In the **potential outcomes** model, every individual  $i$  has two 'potential outcomes':

$$\begin{aligned} Y_i^{(1)} & \text{ with treatment} \\ Y_i^{(0)} & \text{ without treatment.} \end{aligned}$$

Study (unobserved) **treatment effect**  $Y_i^{(1)} - Y_i^{(0)}$ .

Use **nonparametric causal regression** model:

$$Y_i = m(X_i, R_i) + \varepsilon_i, \quad \varepsilon_i \sim N(0, \sigma^2),$$

$i = 1, \dots, n$ , where

- $X_i \sim^{iid} F$  are **features** in  $\mathbb{R}^d$ ,
- $R_i = 1$  (treatment) or  $R_i = 0$  (no treatment),
- $Y_i = R_i Y_i^{(1)} + (1 - R_i) Y_i^{(0)} \in \mathbb{R}$  is the **observed outcome**.

We assume **unconfoundedness**:

$$R \perp\!\!\!\perp Y^{(1)}, Y^{(0)} | X,$$

i.e.  $R$  (treatment assignment) and  $Y^{(0)}, Y^{(1)}$  (outcomes) are **conditionally independent** given measured features  $X$ .

## Main idea

Study **marginal posterior** for **average treatment effect**

$$\psi = E[Y^{(1)} - Y^{(0)}] = \int_{\mathbb{R}^d} m(x, 1) - m(x, 0) dF(x).$$

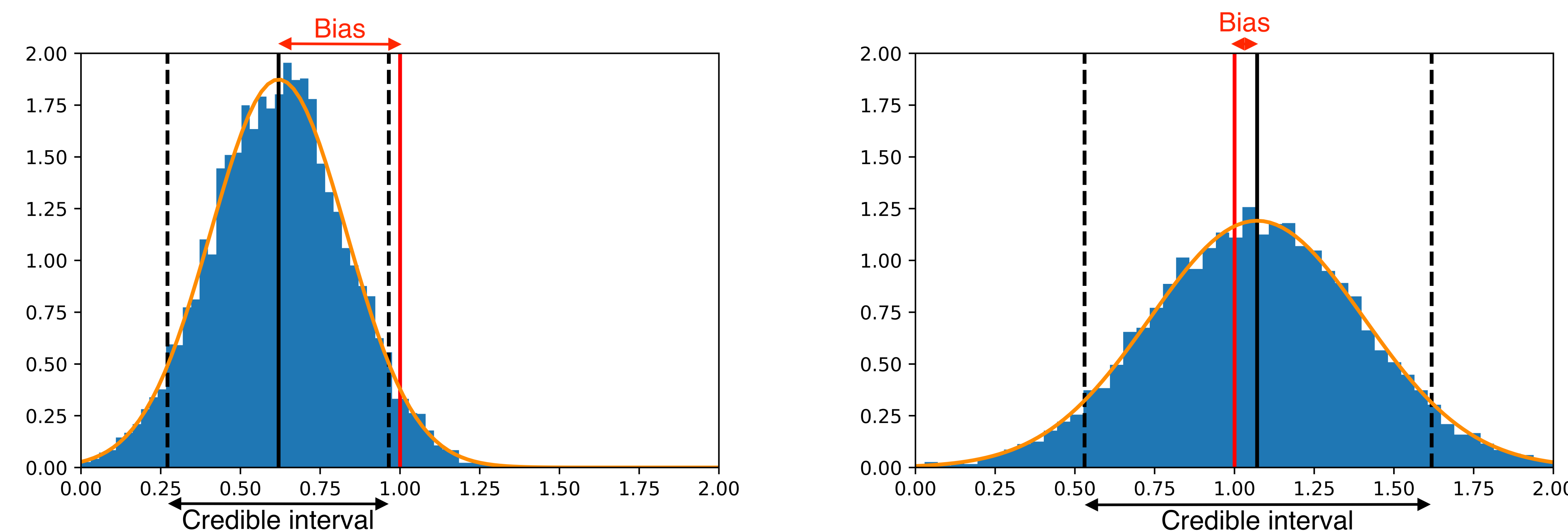
Carefully incorporate an estimate  $\hat{\pi}$  of the **propensity score**  $\pi(x) = P(R = 1 | X = x)$  (e.g. using **logistic regression**) into the prior:

$$m(x, r) = W(x, r) + \lambda \left( \frac{r}{\hat{\pi}(x)} + \frac{1-r}{1-\hat{\pi}(x)} \right), \quad F \sim \text{Dirichlet Process}, \quad \lambda \sim N(0, \nu_n^2) \text{ independent},$$

with  $W$  a standard nonparametric prior (e.g. **Gaussian process**) and  $\nu_n^2 \sim 1/n$ .

**Idea:** increase/decrease prior correlation within/across treatment groups in a heterogeneous way compared to a standard prior ( $\lambda = 0$ ).

**Reduces bias** (correct centering) at the cost of **moderately inflating** the posterior variance.



**Figure 1:** Posterior for average treatment effect  $\psi$  with **true  $\psi = 1.0$**  **Left:** unmodified GP prior ( $\lambda = 0$ ) with bias; **Right:** propensity score corrected GP prior with bias reduction

## Sample simulation results with $d = 100$ features and $n = 1000$

Features  $X_1, \dots, X_{1000} \sim^{iid} N(0, 1)$ , **non-linear** response surface  $m$ , **non-linear** propensity score  $\pi$  and **heterogeneous treatment effect**  $m(x, 1) - m(x, 0)$ .

Method	Abs. error $\pm$ sd	Width 95% CI $\pm$ sd	Coverage
GP (unmodified)	0.321 $\pm$ 0.027	0.613 $\pm$ 0.027	0.38
GP with debiasing (our method)	<b>0.063 <math>\pm</math> 0.042</b>	0.883 $\pm$ 0.040	<b>1.00</b>
BART	0.228 $\pm$ 0.186	1.723 $\pm$ 0.490	<b>1.00</b>
BART with propensity score	0.134 $\pm$ 0.092	0.741 $\pm$ 0.079	0.99
Bayesian Causal Forests	0.144 $\pm$ 0.109	0.535 $\pm$ 0.066	0.87
Causal Forests (AIPW)	0.138 $\pm$ 0.097	<b>0.695 <math>\pm</math> 0.103</b>	0.96
Causal Forests (TMLE)	0.136 $\pm$ 0.100	0.891 $\pm$ 0.152	0.99
Propensity Score Matching	0.234 $\pm$ 0.178	1.282 $\pm$ 0.158	0.97

## Discussion of results

- The unmodified GP **performs badly**, with **biased estimation** and **poor coverage**.
- Our method substantially improves both the **estimation accuracy** and **coverage** of a GP.
- Makes GPs competitive with state-of-the-art.
- Provides a **general route** to improving priors (e.g. BART).
- Can put prior on  $\pi$ , but much slower than using estimator  $\hat{\pi}$  ('empirical Bayes').

## Intuition and theory

- Should help most when  $m$  or  $\pi$  is **difficult to estimate**, especially with **high-dimensional features**.
- In such cases, bias  $\gg$  variance, so need **bias correction** (like we do).
- Idea theoretically investigated in idealized model with simple priors in **Ray & van der Vaart (2018)**.
- Show that for **large sample size**, the posterior for  $\psi$  is **asymptotically Gaussian** (see picture):

$$\psi | (X_i, R_i, Y_i)_{i=1}^n \approx^d N(\hat{\psi}_n, n^{-1} I_0^{-1}),$$

centred at a good (**efficient**) estimator  $\hat{\psi}_n$  with **best possible variance**.

## Future directions

- Scalability: sparse GP approximations, variational Bayes, distributed computing methods.
- Higher order bias corrections for Bayes (our method corrects 'first order' bias)
- Other causal models.

## References

Ray and van der Vaart (2018). Semiparametric Bayesian causal inference. *Annals of Statistics*, to appear.