

Surrogate Model Development using Arbitrary Polynomial Chaos Expansion : Independent Variables

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Outline

- 1 Introduction
- 2 Reasons for Slower Convergence
- 3 Examples
 - $Y = x^2$
 - Rosenbrock Function
- 4 Data-driven APCE for Practical Use
 - Ishigami function
- 5 Summary

Use of Traditional PCE

Estimations of structural vibration responses are studied using traditional PCE:

- Estimation of VIV-induced fatigue damage
Inputs: ΔA_{\max} and $\Delta\omega$ —Shifted Generalized Log-normal Distribution (SGLD)
- Long-term extreme responses of a moored floating structure
Inputs: H_s and T_p —Lognormal and Weibull
- Wood floor vibrations
Inputs: seven ρ_i —Gaussians

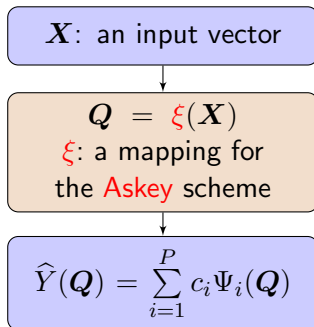
Hermite polynomial family has been applied with a mapping from \mathbf{X} (physical) to \mathbf{Q} (standard normal); also with Nataf and Rosenblatt transformation if needed.

Outside Askey Scheme

- For certain distributions, there exist a best representation which guarantees fast convergence rate: **Askey scheme** - Xiu, 2003.
- But, for inputs outside the Askey scheme, the Askey polynomial families converge, but not exponentially - Witteveen, 2006.
Also, the other reasons are:
 - correlations between inputs - Eldred, 2009.
 - combinations of standard inputs - Oladyshkin, 2012.
- These problems cannot be handled efficiently by using the Askey polynomial families.

Transformation of Input Domain

Flowchart of Askey PCE



$$\underset{\substack{\text{truth model} \\ \text{in } \mathbf{X} \text{ domain}}}{Y(\mathbf{X})} = \underset{\substack{\text{truth model} \\ \text{in } \mathbf{Q} \text{ domain}}}{Y(\mathbf{Q})} \equiv \underset{\substack{\text{surrogate} \\ \text{in } \mathbf{Q} \text{ domain}}}{\hat{Y}(\mathbf{Q})}$$

- ξ can be a non-linear
- Y and \mathbf{X} vs. Y and \mathbf{Q}
- $\hat{Y}(\mathbf{Q})$ aims to fit $Y(\mathbf{Q})$, not $Y(\mathbf{X})$

Ex. 1: Legendre PCE for a Gaussian input

A stupid example:

Truth model: $Y_X(x) = x^2$, $X \sim N(0, 1)$, $Y \sim \chi^2_{\text{dof}=1}$

What if we try **Legendre PCE** (instead of Hermite) to the truth model?

Q : Unif[-1,1] X : N(0,1)

$Q = \xi(\mathbf{X})$:
Askey scheme

$$\begin{aligned} q &= F_Q^{-1}(F_X(x)) \\ &= F_Q^{-1}\left(\frac{1}{2}\left(1 + \operatorname{erf}\left(\frac{x}{\sqrt{2}}\right)\right)\right) \\ &= \operatorname{erf}\left(\frac{x}{\sqrt{2}}\right) \end{aligned}$$

Q	\Rightarrow	Ψ
Normal	\Rightarrow	Hermite
Uniform	\Rightarrow	Legendre

$$Y_X(x) = x^2 = Y_Q(q) = 2\left(\operatorname{erf}^{-1}(q)\right)^2$$

Ex. 1: Legendre PCE for a Gaussian input (cont'd)

$$Y_X(x) = x^2$$

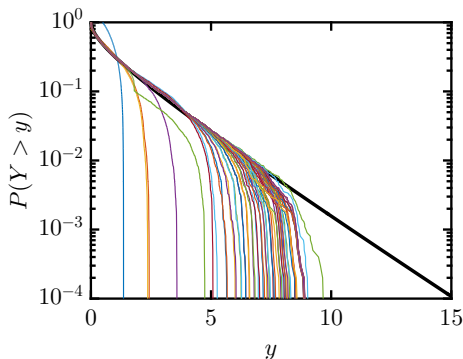
$$Y_Q(q) = 2\left(\operatorname{erf}^{-1}(q)\right)^2$$

- $X \sim N(0, 1)$
- $X : -\infty < x < \infty$
- A quadratic polynomial
- $Q \sim \operatorname{Unif}[-1, 1]$
- $Q : -1 < q < 1$
- Became complicated

$$\hat{Y}_Q(q) \approx \sum_{i=0}^M c_i \Psi_i(q): M \text{ should be larger than } 2$$

Ex. 1: Legendre PCE for a Gaussian input (cont'd)

Comparison of PoEs by $1 - F_Y(y)$ and MCS of **surrogate** models



- Black solid line: $Y_X(x)$
- Others: $\hat{Y}_Q(q)$ of orders from 1 to 60 with $3 \times \#$ LR samples
- $\hat{Y}_Q(q)$ show slow convergence to estimate low PoEs (quite inefficient)

But Hermite PCE with an order 2 shows the convergence to $Y_X(x) = x^2$

Ex. 1: Comparison of Coefficients and Basis Functions

Hermite PCE ($M = 2$), $X \sim N(0, 1)$, $c_i = \frac{\mathbb{E}[Y_X(x)H_i(x)]}{\mathbb{E}[H_i^2(x)]}$

$$\hat{Y}_X(x) = \sum_{i=0}^M c_i H_i(x) = \underset{c_0}{1} \cdot \underset{H_0}{1} + \underset{c_1}{0} \cdot \underset{H_1}{x} + \underset{c_2}{1} \cdot \underset{H_2}{(x^2 - 1)}$$

Legendre PCE ($M = 2$), $Q \sim \text{Unif}[-1, 1]$, $c_i = \frac{\mathbb{E}[Y_Q(q)L_i(q)]}{\mathbb{E}[L_i^2(q)]}$

$$\begin{aligned} \hat{Y}_Q(q) &= \sum_{i=0}^M c_i L_i(q) = \underset{c_0}{1} \cdot \underset{L_0}{1} + \underset{c_1}{\left[3 \int_{-1}^1 \{\text{erf}^{-1}(q)\}^2 q \, dq \right]} \cdot \underset{L_1}{q} \\ &\quad + \underset{c_2}{\left[\frac{5}{2} \int_{-1}^1 \{\text{erf}^{-1}(q)\}^2 (3q^2 - 1) \, dq \right]} \cdot \underset{L_2}{\left\{ \frac{1}{2} (3q^2 - 1) \right\}} \end{aligned}$$

Ex. 2: Rosenbrock Function with Lognormal Variables

Truth model: $Y_{\mathbf{X}}(x_1, x_2) = 100(x_2 - x_1^2)^2 + (1 - x_1)^2$, $X_{1,2} \sim LN(1, 0.5)$

No **Askey** scheme for lognormal variables \rightarrow try Hermite

$$Q_1: N(0, 1) \quad X_1: LN(1, 0.5)$$

$$\begin{aligned} q_1 &= F_Q^{-1}(F_X(x)) \\ &= \frac{\ln(x_1) - 1}{0.5} \end{aligned}$$

$Q = \xi(\mathbf{X})$:
Askey scheme

$$\begin{aligned} Y_Q(q_1, q_2) &= 100(e^{0.5q_2+1} - e^{q_1+2})^2 \\ &\quad + (1 - e^{0.5q_1+1})^2 \end{aligned}$$

Ex. 2: Comparison of Coefficients and Basis Functions

Truth model: $Y_{\mathbf{X}}(x_1, x_2) = 100(x_2 - x_1^2)^2 + (1 - x_1)^2$, $X_{1,2} \sim LN(1, 0.5)$

multi-index, α_i

An illustrative case: $p = 2$

	X_1	X_2	P_i
$i = 0$	0	0	$P_{(0,0)}$
1	1	0	$P_{(1,0)}$
2	0	1	$P_{(0,1)}$
3	2	0	$P_{(2,0)}$
4	1	1	$P_{(1,1)}$
5	0	2	$P_{(0,2)}$

$$\text{PCE : } \hat{Y}_{\mathbf{Q}}(q_1, q_2) = \sum_{i=0}^5 \overbrace{c_i}^{\text{LR}} \underbrace{H_i(q_1, q_2)}_{\text{pre-defined}}$$

$$\text{aPCE : } \hat{Y}_{\mathbf{X}}(x_1, x_2) = \sum_{i=0}^5 \overbrace{c_i}^{\text{LR}} \underbrace{\Phi_i(x_1, x_2)}_{\text{Gram-Schmidt}}$$

E.g.)

$$H_4(q_1, q_2) = \underbrace{\hat{H}_1(q_1)}_{Q_1} \underbrace{\hat{H}_1(q_2)}_{Q_2}$$

$$L^2 \equiv \int_{\mathbf{X}} \Phi_i \Phi_j \underbrace{f_{\mathbf{X}}}_{\text{LN jpdf}} d\mathbf{x} = 0, \quad \text{for } i \neq j$$

Ex. 2: Univariate Gram-Schmidt Polynomials: $\hat{\Phi}$

make $\{1, x_1, x_1^2\}$ orthogonal each other $\rightarrow \{1, \hat{\Phi}_1, \hat{\Phi}_2\}$

$$\hat{\Phi}_0(x_1) = 1$$

$$\hat{\Phi}_1(x_1) = x_1 - \underbrace{\frac{\langle x_1, \hat{\Phi}_0 \rangle}{\langle \hat{\Phi}_0, \hat{\Phi}_0 \rangle} \hat{\Phi}_0(x_1)}_{=1}$$

$$\hat{\Phi}_2(x_1) = x_1^2 - \underbrace{\frac{\langle x_1^2, \hat{\Phi}_0 \rangle}{\langle \hat{\Phi}_0, \hat{\Phi}_0 \rangle} \hat{\Phi}_0(x_1)}_{=1} - \frac{\langle x_1^2, \hat{\Phi}_1 \rangle}{\langle \hat{\Phi}_1, \hat{\Phi}_1 \rangle} \hat{\Phi}_1(x_1)$$

$$\langle \hat{\Phi}_i, \hat{\Phi}_i \rangle \equiv \int_{\mathbf{X}} (\hat{\Phi}_i)^2 f_{\mathbf{X}} d\mathbf{x}$$

$$m_p = \int x_1^p f_{X_1} dx_1$$

$$\hat{\Phi}_1(x_1) = x_1 - \frac{\langle x_1, 1 \rangle}{\langle 1, 1 \rangle} \cdot 1 = x_1 - \int_{X_1} x_1 f_{X_1} dx_1 = x_1 - m_1$$

$$\hat{\Phi}_2(x_1) = x_1^2 - \frac{\langle x_1^2, 1 \rangle}{\langle 1, 1 \rangle} \cdot 1 - \frac{\langle x_1^2, x_1 - m_1 \rangle}{\langle x_1 - m_1, x_1 - m_1 \rangle} (x_1 - m_1)$$

$$= \underbrace{1}_{a_2^{(2)}} \cdot x_1^2 + \underbrace{\left(\frac{m_1 m_2 - m_3}{m_2 - m_1^2} \right)}_{a_1^{(2)}} x_1 + \underbrace{\left(\frac{m_1 m_3 - m_2^2}{m_2 - m_1^2} \right)}_{a_0^{(2)}}$$

$$\hat{\Phi}_i(x_1) = \sum_{k=0}^i a_k^{(i)} x_1^k$$

for $a_k^{(i)}$
a matrix form
is compact

Ex. 2: Comparison of Multi-indices

$$\begin{aligned}
 Y_{\mathbf{X}}(x_1, x_2) &= 100(x_2 - x_1^2)^2 + (1 - x_1)^2 \\
 &= \underbrace{1}_{(0,0)} - \underbrace{2x_1}_{(1,0)} + \underbrace{x_1^2}_{(2,0)} + \underbrace{100x_2^2}_{(0,2)} - \underbrace{200x_1^2x_2}_{(2,1)} + \underbrace{100x_1^4}_{(4,0)}
 \end{aligned}$$

$$\begin{aligned}
 Y_{\mathbf{Q}}(q_1, q_2) &= 100(e^{0.5q_2+1} - e^{q_1+2})^2 + (1 - e^{0.5q_1+1})^2 \\
 &= 1 - 2e^{0.5q_1+2} + e^{q_1+2} + 100e^{q_1+2} - 200e^{q_1+0.5q_2+3} + 100e^{2q_1+4}
 \end{aligned}$$

multi-index inconveniently applicable

$$e^{q_1} = 1 + q_1 + \frac{q_1^2}{2!} + \frac{q_1^3}{3!} + \dots$$

Ex. 2: Parameter Estimation

Same 45 random samples are used as a part of LR; $3 \times \binom{4+2}{2} = 45$.

$$Y_{\mathbf{X}}(x_1, x_2) = 100(x_2 - x_1^2)^2 + (1 - x_1)^2 \approx \hat{Y}_{\mathbf{X}}(x_1, x_2) = \sum_{i=0}^{14} c_i \Phi_i(x_1, x_2)$$

$$Y_{\mathbf{Q}}(q_1, q_2) = 100(e^{0.5q_2+1} - e^{q_1+2})^2 + (1 - e^{0.5q_1+1})^2 \approx \hat{Y}_{\mathbf{Q}}(q_1, q_2) = \sum_{i=0}^{14} c_i H_i(q_1, q_2)$$

c_i	H_i (pre-defined)	c_i	Φ_i (Gram-Schmidt)
24323	1	34063	1
41547	q_1	73678	$x_1 - 3.08$
-1985	q_2	-1533	$x_2 - 3.08$
36366	$q_1^2 - 1$	33880	$x_1^2 - 9.03x_1 + 15.64$
-2488	$q_1 q_2$	-1807	$(x_1 - 3.08)(x_2 - 3.08)$
-2170	$q_2^2 - 1$	100	$x_2^2 - 9.03x_2 + 15.64$
13018	$q_1^3 - 3q_1$	3945	$x_1^3 - 19.97x_1^2 + 101.43x_1 - 130.97$
-29	$q_2(q_1^2 - 1)$	-200	$(x_1^2 - 9.03x_1 + 15.64)(x_2 - 3.08)$
-942	$q_1(q_2^2 - 1)$	0	$(x_1 - 3.08)(x_2^2 - 9.03x_2 + 15.64)$
-47	$q_2^3 - 3q_2$	0	$x_2^3 - 19.97x_2^2 + 101.43x_2 - 130.97$
4411	$q_1^4 - 6q_1^2 + 3$	100	$x_1^4 - 39.4x_1^3 + 442.9x_1^2 - 1677.4x_1 + 1808$
10	$q_2(q_1^3 - 3q_1)$	0	$(x_1^3 - 19.97x_1^2 + 101.43x_1 - 130.97)(x_2 - 3.08)$
-117	$(q_1^2 - 1)(q_2^2 - 1)$	0	$(x_1^2 - 9.03x_1 + 15.64)(x_2^2 - 9.03x_2 + 15.64)$
-298	$q_1(q_2^3 - 3q_2)$	0	$(x_1 - 3.08)(x_2^3 - 19.97x_2^2 + 101.43x_2 - 130.97)$
-560	$q_2^4 - 6q_2^2 + 3$	0	$x_2^4 - 39.4x_2^3 + 442.9x_2^2 - 1677.4x_2 + 1808$

Ex. 2: Coefficients by LR: OLS

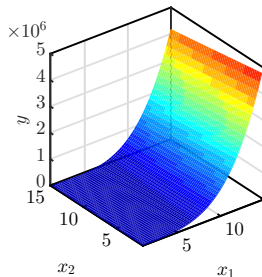
For c_i , the same samples are drawn for LR:



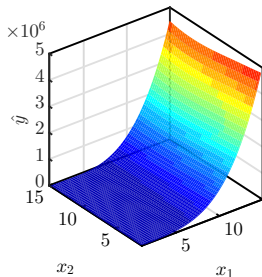
$$\begin{bmatrix} H_0(\mathbf{q}^{(1)}) & \cdots & H_{14}(\mathbf{q}^{(1)}) \\ \vdots & \ddots & \vdots \\ H_0(\mathbf{q}^{(45)}) & \cdots & H_{14}(\mathbf{q}^{(45)}) \end{bmatrix} \begin{Bmatrix} c_0 \\ \vdots \\ c_{14} \end{Bmatrix} = \overbrace{\begin{Bmatrix} Y_{\mathbf{X}}(\mathbf{x}^{(1)}) \\ \vdots \\ Y_{\mathbf{X}}(\mathbf{x}^{(45)}) \end{Bmatrix}}^{\text{4th order polynomial}} = \overbrace{\begin{Bmatrix} Y_{\mathbf{Q}}(\mathbf{q}^{(1)}) \\ \vdots \\ Y_{\mathbf{Q}}(\mathbf{q}^{(45)}) \end{Bmatrix}}^{\text{complicated than } Y_{\mathbf{X}}}$$

$$\begin{bmatrix} \Phi_0(\mathbf{x}^{(1)}) & \cdots & \Phi_{14}(\mathbf{x}^{(1)}) \\ \vdots & \ddots & \vdots \\ \Phi_0(\mathbf{x}^{(45)}) & \cdots & \Phi_{14}(\mathbf{x}^{(45)}) \end{bmatrix} \begin{Bmatrix} c_0 \\ \vdots \\ c_{14} \end{Bmatrix} = \begin{Bmatrix} Y_{\mathbf{X}}(\mathbf{x}^{(1)}) \\ \vdots \\ Y_{\mathbf{X}}(\mathbf{x}^{(45)}) \end{Bmatrix}$$

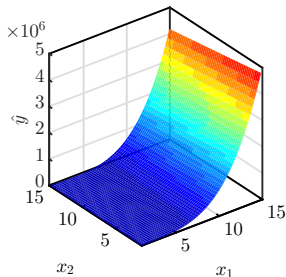
Ex. 2: Response Surfaces



Truth



PCE
Order= 13



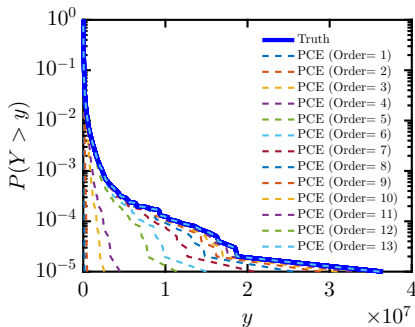
aPCE
Order= 4

- An order-4 aPCE shows convergence to the truth.
- An order-13 PCE still has inaccuracy at some regions.

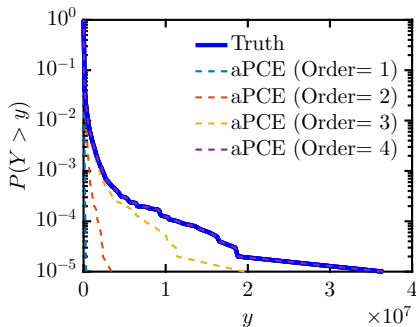
Ex. 2: Exceedance Probabilities

Truth: $Y_{\mathbf{X}}(x_1, x_2)$ PCE: $\hat{Y}_{\mathbf{Q}}(q_1, q_2)$ aPCE: $\hat{Y}_{\mathbf{X}}(x_1, x_2)$

$P(Y > y)$ obtained empirically by using MCS of **truth** and **surrogates**
Identical random seeds are used for the comparison

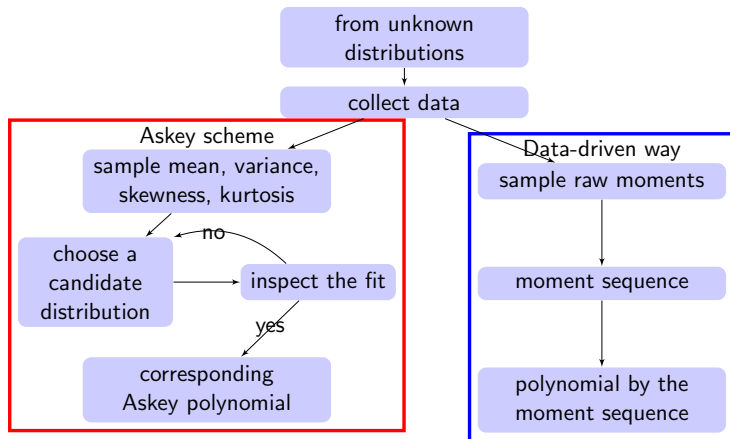


PCE



aPCE

Framework: APCE



Ishigami function

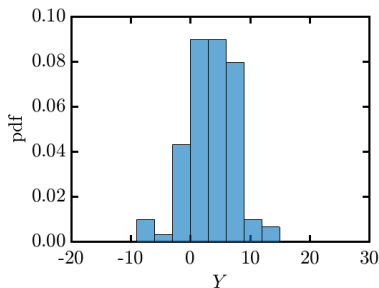
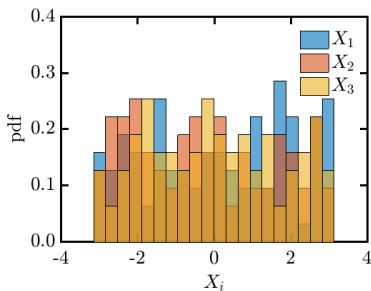
The Ishigami function:

$$Y_{\mathbf{X}}(\mathbf{x}) = \sin(x_1) + a(\sin(x_2))^2 + bx_3^4 \sin(x_1).$$

X_1 , X_2 , and X_3 follow uniform distributions over $[-\pi, \pi]^3$.

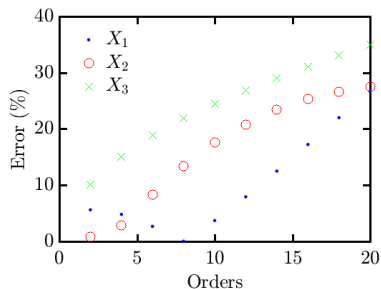
a and b are 7 and 0.1, respectively.

Data is given: $\mathcal{D} = \{\mathbf{x}_i, y_i\}_{i=1}^N$, $N = 100$



Sample Raw Moments

$$\text{Error (\%)} = \frac{|\text{Exact} - \text{SRM}|}{\text{Exact}} \times 100$$



APCE

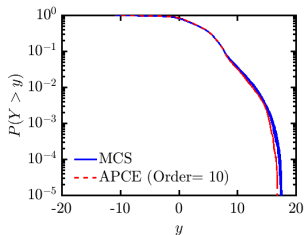
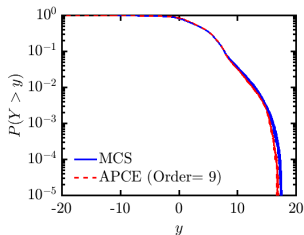
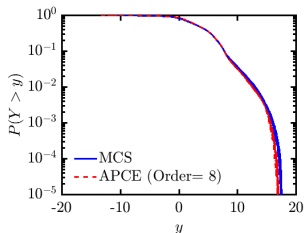
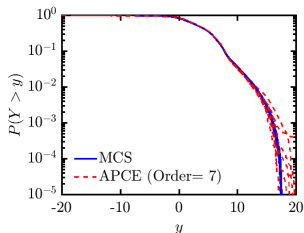
$$\hat{Y}_{\mathbf{X}}(\mathbf{x}) = \sum_{i=0}^P c_i \Psi_i(\mathbf{x})$$

Ψ is constructed based on \mathcal{D} ;

c_i is also constructed based on \mathcal{D} with resampling (if needed);

Uncertainty propagation is also done by resampling in \mathcal{D} .

Exceedance Plots



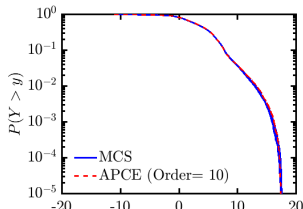
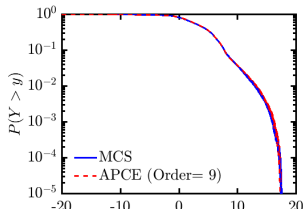
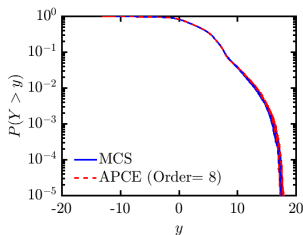
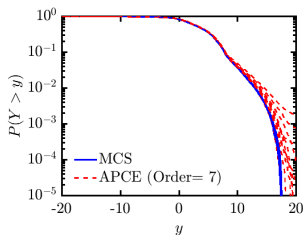
An order-8 APCE models look convergent to the truth.
The estimations at low $P(Y > y)$ levels exhibit inaccuracy.

Reason for Inaccuracy

Inaccurate estimation of high-order raw moments \Rightarrow inaccurate estimation of high-order basis functions

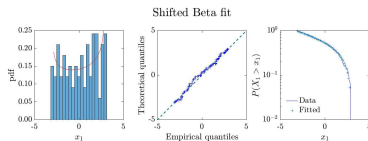
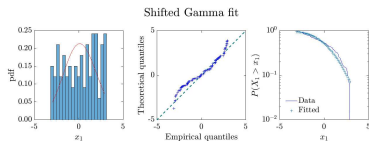
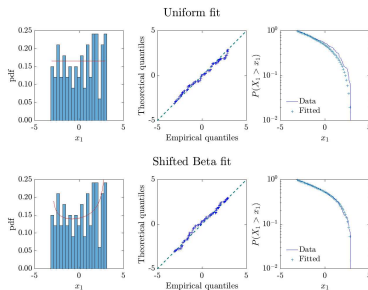
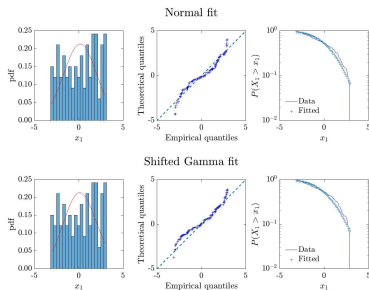
Accuracy is improved when $N = 1000$.

High variability of order-7 APCE models might be due to high sampling variability when $N = 1000$.



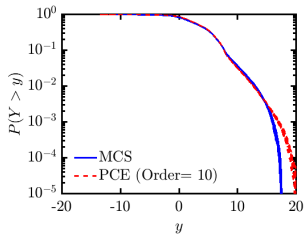
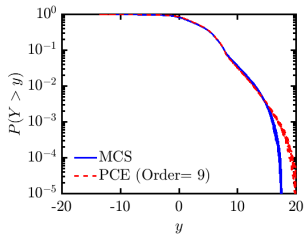
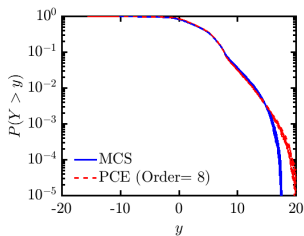
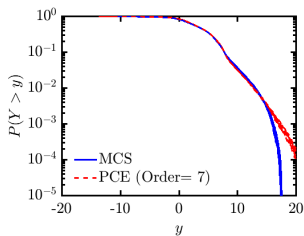
Classic PCE: Askey Scheme

Candidate distributions are investigated for the Askey scheme PCE.
Normal (Hermite), Uniform (Legendre), Gamma (Laguerre), Beta (Jacobi)



A Shifted Beta fit is selected; Jacobi polynomials are used.

Exceedance Plots



Pros and Cons: APCE

Pros:

- Can make a surrogate without fitting input variables to certain distribution types.
- Fast convergence to the truth model (regarding response surface).

Cons:

- Surrogate model accuracy depends on the estimation of raw moments.
- High-order sample raw moments can be inaccurate with a small number of samples.
- A high-order basis function can be inaccurate, too.

Future work:

- To improve accuracy of high-order raw moment approximation.

Thank you very much for your attention!