Allocation Problems in Ride-sharing Platforms: Online Matching with Offline Reusable Resources

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Bipartite-matching markets pair agents on one side of a market with agents, items, or contracts on the opposing side. Prior work addresses online bipartite-matching markets, where agents arrive over time and are dynamically matched to a known set of disposable resources. In this article, we propose a new model, *Online Matching with (offline) Reusable Resources under Known Adversarial Distributions* (OM-RR-KAD), in which resources on the offline side are *reusable* instead of disposable; that is, once matched, resources become available again at some point in the future. We show that our model is tractable by presenting an LP-based non-adaptive algorithm that achieves an online competitive ratio of $\frac{1}{2} - \epsilon$ for any given constant $\epsilon > 0$. We also show that no adaptive algorithm can achieve a ratio of $\frac{1}{2} + o(1)$ based on the same benchmark LP. Through a data-driven analysis on a massive openly available dataset, we show our model is robust enough to capture the application of taxi dispatching services and ride-sharing systems. We also present heuristics that perform well in practice.

CCS Concepts: • **Theory of computation** → *Design and analysis of algorithms*; *Online algorithms*;

Additional Key Words and Phrases: Online-matching, rideshare, randomized algorithms, matching markets

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1 INTRODUCTION

In bipartite-matching problems, agents on one side of a market are paired with agents, contracts, or transactions on the other. Classical matching problems—assigning students to schools, papers

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to reviewers, or medical residents to hospitals—take place in a static setting, where all agents exist at the time of matching, are simultaneously matched, and then the market concludes. In contrast, many matching problems are dynamic, where one side of the market arrives in an *online* fashion and is matched sequentially to the other side.

Online bipartite-matching problems are primarily motivated by Internet advertising. In the basic version of the problem, we are given a bipartite graph G=(U,V,E) where U represents the offline vertices (advertisers) and V represents the online vertices (keywords or impressions). There is an edge e=(u,v) if advertiser u bids for a keyword v. When a keyword v arrives, a central clearinghouse must make an instant and irrevocable decision to either reject v or assign v to one of its "neighbors" v in v and obtain a profit v for the match v is matched, it is no longer available for matches with other keywords (in the most basic case) or its budget is reduced. The goal is to design an efficient online algorithm such that the expected total weight (profit) of the matching obtained is maximized. Following the seminal work of Karp et al. [29], there has been a large body of research on related variants (overviewed by Mehta [40]). One particular flavor of problems is **online-matching with known identical independent distributions** (OM-KIID) [14, 21, 26, 28, 36]. In this context, agents arrive over v rounds, and their arrival distributions are assumed to be *identical and independent* over all v rounds; additionally, this distribution is known to the algorithm beforehand.

Apart from the Internet-advertising application, online bipartite-matching models have been used to capture a wide range of online resource allocation and scheduling problems. Typically, we have an offline and an online party representing, respectively, the **service providers (SP)** and online users; once an online user arrives, we need to match it to an offline SP immediately. In many cases, the service is *reusable* in the sense that once an SP is matched to a user, it will be gone for some time, but will then rejoin the system afterwards. Besides that, in many real settings the arrival distributions of online users do change from time to time (i.e., they are not i.i.d.). Consider the following motivating examples.

Taxi Dispatching Services and Ride-sharing Systems. Traditional taxi services and ridesharing systems such as Uber and Didi Chuxing match drivers to would-be riders [30, 31, 50, 55]. Here, the offline SPs are different vehicle drivers. Once an online request (potential rider) arrives, the system matches it to a nearby driver instantly such that the rider's waiting time is minimized. In most cases, the driver will rejoin the system and can be matched again once she finishes the service. Additionally, the arrival rates of requests changes dramatically across the day. Consider the online arrivals during peak hours and off-peak hours for example: The arrival rates in the former case can be much larger than the latter. Though our model is primarily motivated by taxidispatching services and ride-sharing platforms, we acknowledge that it does not address the full generality of these applications. In particular, our work focuses on the temporal components of rideshare; spatial components such as match-specific driver movement are not modeled [17, 18, 33]. Since we make the simplifying assumption that the stochastic process that determines the parameters in the environment is independent of the algorithm's decisions, this may not fully capture the scenario where requests are heterogenous with vastly different ride times. One potential direction to extend this model is to consider correlated reusable rates, which makes the theory much more challenging.

Organ Allocation. Chronic kidney disease affects tens of millions of people worldwide at great societal and monetary cost [43, 49]. Organ donation—either via a deceased or living donor—is a lifesaving alternative to organ failure. In the case of kidneys, a donor organ can last up to 15 years in a patient before failing again. Various nationwide organ donation systems exist and operate under different ethical and logistical constraints [12, 20, 37], but all share a common online structure:

The offline party is the set of patients (who reappear every 5 to 15 years based on donor organ longevity), and the online party is the set of donors or donor organs, who arrive over time. Similarly, in some blood or plasma donation settings, donors might reappear after some number of weeks.

Similar scenarios can be seen in other areas such as wireless network connection management (SPs are different wireless access points) [59] and online cloud computing service scheduling [42, 60]. Inspired by the above applications, we generalize the model of OM-KIID in the following two ways.

Reusable Resources. Once we assign v to u, u will rejoin the system after C_e rounds with e = (u, v), where $C_e \in \{0, 1, ..., T\}$ is an integral random variable with known distribution. In this article, we call C_e the *occupation time* of u w.r.t. e. In fact, we show that our setting can directly be extended to the case when C_e is time-sensitive: When matching v to u at time t, u will rejoin the system after $C_{e,t}$ rounds. This extension makes our model adaptive to nuances in real-world settings. For example, consider the taxi dispatching or ride-sharing service: The occupation time of a driver u from a matching with an online user v does depend on both the user type of v (such as destination) and the time when the matching occurs (peak hours can differ significantly from off-peak hours).

Known Adversarial Distributions (KAD). Suppose we have T rounds and that for each round $t \in [T]$, a vertex v is sampled from V according to an arbitrary known distribution \mathcal{D} where the marginal for v is $\{p_{v,t}\}$ such that $\sum_{v \in V} p_{v,t} \leq 1$ for all t. Also, the arrivals at different times are independent (and according to these given distributions). The setting of KAD was introduced by References [4, 5] and is known as Prophet Inequality matching.

We call our new model **Online Matching with (offline) Reusable Resources under Known Adversarial Distributions (**OM-RR-KAD). Note that the OM-KIID model can be viewed as a special case when C_e is a constant (with respect to T) and $\{p_{v,t}|v \in V\}$ are the same for all $t \in [T]$.

Competitive Ratio. Let $E[ALG(I,\mathcal{D})]$ denote the expected value obtained by an algorithm ALG on an input I and arrival distribution \mathcal{D} . Let E[OPT(I)] denote the expected *offline optimal value*, which refers to the optimal solution when we are allowed to make choices after observing the entire sequence of online arrival vertices. Then, the competitive ratio is defined as $\min_{I,\mathcal{D}} \frac{E[ALG(I,\mathcal{D})]}{E[OPT(I)]}$. It is a common technique to use an LP optimal value to upper bound E[OPT(I)] (called the benchmark LP) and hence get a valid lower bound on the resulting competitive ratio.

Adaptive vs. non-adaptive algorithms. An online algorithm ALG is called *non-adaptive* if the strategy for all time-steps t, denoted by ALG(t), is pre-computed *before* the realizations of the online process. In contrast, adaptive algorithms can choose the strategy for time-step t *after* seeing the random realizations of all the random processes (e.g., random seeds used in the algorithm, arrivals of online requests and occupation times of drivers) in steps $1, 2, \ldots, t-1$.

1.1 Our Contributions

First, we propose the new model of OM-RR-KAD to capture a wide range of real-world applications related to online scheduling, organ allocation, rideshare dispatch, among others. We claim that this model is tractable enough to obtain good algorithms with theoretically provable guarantees and general enough to capture many real-life instances. Our model assumptions take a significant step forward from the usual assumptions in the online-matching literature where the offline side is assumed to be <code>single-use</code> or <code>disposable</code>. This leads to a larger range of potential applications that can be modeled by online-matching. The first part of the article focuses on the abstract model, and

¹Throughout this article, we use [N] to denote the set $\{1, 2, \ldots, N\}$, for any positive integer N.

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though the experiments are motivated using ride-share, we do not get into the application-specific assumptions/issues one may run into. For specific discussion on issues that arise when applied to ride-share, we defer to Section 5.

Second, we show how this model can be *cleanly* analyzed under a theoretical framework. We first construct a **linear program** (LP) LP (1), which we show is a valid upper bound on the expected offline optimal value (note that the latter is hard to characterize). Next, we propose an efficient *non-adaptive* algorithm that achieves a competitive ratio of $\frac{1}{2} - \epsilon$ for any given constant $\epsilon > 0$. This algorithm solves the LP and obtains an optimal fractional solution. It uses this optimal solution as a guide in the online phase. Using Monte-Carlo simulations (called simulations henceforth), and combining with this optimal solution, our algorithm makes the online decisions. In particular, Theorem 1 describes our first theoretical results formally.

Theorem 1. LP (1) is a valid benchmark for OM-RR-KAD. There exists a non-adaptive online algorithm that achieves an online competitive ratio of $\frac{1}{2} - \epsilon$ for any given $\epsilon > 0$ against the benchmark LP (1).

Third, we show that the online-competitive analysis of the non-adaptive algorithm is tight with respect to the choice of our benchmark LP. Specifically, we show that when restricted to LP (1), no adaptive algorithm can achieve a competitive ratio better than $\frac{1}{2}$ (Theorem 2), and no non-adaptive algorithm can beat $\frac{1}{2}$ even when all C_e are deterministic and equal (Theorem 3).

Theorem 2. No adaptive algorithm can achieve a competitive ratio better than $\frac{1}{2} + o(1)$ against the benchmark LP (1). Here, o(1) is a vanishing term when T is sufficiently large.

Theorem 3. No non-adaptive algorithm can achieve a competitive ratio better than $\frac{1}{2} + o(1)$ against the benchmark LP (1) even when all C_e are deterministic and equal. Here, o(1) is a vanishing term when both of C_e and T/C_e are sufficiently large.

Finally, through a data-driven analysis on a massive openly available dataset, we show that our model is robust enough to capture the setting of taxi hailing/sharing at least. Additionally, we provide certain *simpler* heuristics that also give good performance. Hence, we can combine these theoretically grounded algorithms with such heuristics to obtain further improved ratios in practice. Section 5 provides a detailed qualitative and quantitative discussion.

1.2 Other Related Work

In addition to the arrival assumptions of KIID and KAD, there are several other important, well-studied variants of online-matching problems. Under *adversarial ordering*, an adversary can arrange the arrival order of all items in an arbitrary way (e.g., online-matching [29, 53] and Ad-Words [16, 41]). Under a *random arrival order*, all items arrive in a random permutation order (e.g., online-matching [35] and AdWords [24]). Finally, under *unknown distributions*, in each round, an item is sampled from a fixed but unknown distribution. (e.g., Reference [19]). For each of the categories above, we list only a few examples considered under that setting. For a more complete list, please refer to the book by Mehta [40].

Despite the fact that our model is inspired by online bipartite-matching, it also overlaps with **stochastic online scheduling problems** (SOS) [38, 39, 52]. We first restate our model in the language of SOS: We have |U| nonidentical parallel machines and |V| jobs; at every time-step a single job v is sampled from V with probability $p_{v,t}$; the jobs have to be assigned immediately after its arrival (or rejected right away); additionally, each job v can be processed non-preemptively on a specific subset of machines; once we assign v to v, we get a profit of v and v will be occupied for v counds with v and v with v and

the key difference between our model and SOS is in the objective: The former is to maximize the expected profit from the completed jobs, while the latter is to minimize the total or the maximum completion time of all jobs. In a recent concurrent work [22], they consider the dynamic assortment of reusable resources. In their context, the offline side is the set of reusable resources, while the online side is the set of consumers. They consider the arrival setting of KAD (which they call the Bayesian model with non-identical distributions). The critical difference is that they assume that in each round, on the arrival of an online customer, the algorithm should assign her a set of offline resources, where each offline resource has a given budget. The benchmark LP has an exponential number of variables, which is not solvable in polynomial time. They overcome this by assuming an offline oracle, which returns an optimal solution to the benchmark LP for any given arrival sequence of online customers. They design a similar simulation-based online policy that achieves a competitive ratio of 1/2.

Research in ridesharing platforms and similar allocation problems is an active area of research within multiple fields, including computer science, operations research, and transportation engineering. State-independent policies were studied previously using theory from control and queuing systems [11, 13, 46]. The role of pricing in the dynamics of drivers in ridesharing platforms is also an active area of research in computational economics and AI/ML (e.g., References [2, 10, 17, 33, 44, 47, 58]). Our problem is a form of *online-matching in dynamic environments*, which is an active area of research within the AI/ML community. In particular, References [20, 31, 54, 55] have studied algorithms for matching in various dynamic bipartite markets such as kidney exchange, spatial crowdsourcing, labor markets, and so on. A similar line of work on general graphs is also prominent in the literature (e.g., References [3, 6–8]).

$$\text{maximize } \sum_{t \in [T]} \sum_{e \in E} w_e x_{e,t} \tag{1}$$

subject to
$$\sum_{e \in E_v} x_{e,t} \le p_{v,t}$$
 $\forall v \in V, t \in [T]$ (2)

$$\sum_{t' < t} \sum_{e \in E_u} x_{e,t'} \Pr[C_e > t - t'] + \sum_{e \in E_u} x_{e,t} \le 1 \qquad \forall u \in U, t \in [T]$$
(3)

$$0 \le x_{e,t} \le 1 \qquad \qquad \forall e \in E, t \in [T] \tag{4}$$

2 MAIN MODEL

In this section, we present a formal statement of our main model. Suppose we have a bipartite graph G=(U,V,E) where U and V represent the offline and online parties, respectively. We have a finite time horizon T (known beforehand) and for each time $t\in [T]$, a vertex v will be sampled (we use the term v arrives) from a known probability distribution $\{p_{v,t}\}$ such that $\sum_{v\in V}p_{v,t}\leq 1^2$ (noting that such a choice is made independently for each round t). The expected number of times v arrives across the T rounds, $\sum_{t\in [T]}p_{v,t}$, is called the arrival rate for vertex v. Once a vertex v arrives, we need to make an irrevocable decision immediately: either to reject v or assign v to one of its neighbors in v. For each v, once it is assigned to some v, it becomes unavailable for v0 rounds with v0, and subsequently rejoins the system. Here, v0 is an integral random variable taking values from v0, 1, . . . , v1 and the distribution is known in advance. Each assignment v1 is associated with a weight v2 and our goal is to design an online assignment policy such that the

²Thus, with probability $1 - \sum_{v \in V} p_{v,t}$, none of the vertices from V will arrive at t.

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total expected weights of all assignments made is maximized. Following prior work, we assume $|V| \gg |U|$ and $T \gg 1$. Throughout this article, we use edge e = (u, v) and assignment of v to u interchangeably.

For an assignment e, let $x_{e,t}$ be the probability that e is chosen at t in any offline optimal algorithm. For each u (likewise for v), let E_u (E_v) be the set of neighboring edges incident to u (v). We use the LP (1) as a benchmark to upper bound the offline optimal. We now interpret the constraints. For each round t, once an online vertex v arrives, we can assign it to at most one of its neighbors. Thus, we have: If v arrives at t, then the total number of assignments for v at t is at most 1; if v does not arrive, then the total is 0. The LHS of (2) is exactly the expected number of assignments made at t for v. It should be no more than the probability that v arrives at v, which is the RHS of (2). Constraint (3) is the v ovel part of our problem formulation. Consider a given v and v is not available at v while the second term (summation over v and v is the probability that v is assigned to some driver at v, which is no larger than probability v is available at v. Thus, the sum of the first term and second term on LHS is no larger than 1. This argument implies that the LP forms a valid upper-bound on the offline optimal solution and hence, we have the first part of Lemma 4.

Lemma 4. The optimal value to LP (1) is a valid upper bound for the offline optimal. Moreover, suppose for some $\delta \geq 0$, we have an estimate f(e,y) of $\Pr[C_e > y]$ for all edges e and $y \geq 0$, where $f(e,y)/\Pr[C_e > y]$ always lies in $[1/(1+\delta), 1+\delta]$. Then, by using f(e,t-t') in the LP instead of $\Pr[C_e > t-t']$ and scaling down the resultant vector \mathbf{x} by $(1+\delta)$, we only get a further loss of $(1+\delta)$ in the competitive ratio.

PROOF. Fix any offline optimal algorithm OPT. For each assignment e = (u, v) and t, let $X_{e,t}$ denote an indicator random variable for the event that e is matched at t in OPT, which includes the event that v arrives at time t. Let $\mathbb{E}[X_{e,t}] = x_{e,t}$. Therefore, by linearity of expectation, the expected performance of OPT is $\mathbb{E}[OPT] = \sum_t \sum_e w_e x_{e,t}$. Now, we justify that the solution $\{x_{e,t}\}$ is feasible to all constraints in LP (1).

In constraint (2), the LHS denotes the probability that v is matched at t, which should be no larger than the probability that vertex v arrives at time t. In constraint (3), the first part $\sum_{t' < t} \sum_{e \in E_u} x_{e,t'} \Pr[C_e > t - t']$ denotes the probability that u is occupied due to assignments made prior to t while the second part $\sum_{e \in E_u} x_{e,t}$ is the probability that an assignment incident to u is made at time t. Thus, the sum of these two parts should be no larger than 1. Constraint (4) is satisfied, since $\{x_{e,t}\}$ are all probability values. Therefore, we have shown that the values $\{x_{e,t}\}$ is feasible to all constraints in LP (1), which implies that the optimal value of LP (1) is a valid upper bound for the performance of any offline optimal.

The second part follows directly from the fact that **x** is scaled down by a factor $(1+\delta)$ and hence the objective is scaled down by a factor $(1+\delta)$.

3 SIMULATION-BASED ALGORITHM

In this section, we present a simulation-based algorithm. We will first give a gentle introduction to simulation-based algorithms that has been developed and used in prior works (References [1] and [15]).

Simulation-based algorithms. We use the term *simulation* throughout this article to refer to Monte Carlo simulation and the term *simulation-based attenuation* to refer to the simulation and

 $^{^{3}}$ We would like to point out that our LP constraint (3) on u is inspired by Ma [34]. The proof is similar to that by Alaei et al. [4] and Alaei et al. [5].

attenuation techniques as shown in References [1] and [15].⁴ At a high level, suppose we have a randomized algorithm such that for some event E (i.e., u is available at t), we have $\Pr[E] \geq c$, then we modify the algorithm as follows: (i) We first use simulation to estimate a value \hat{E} that lies in the range $\Pr[E]$, $(1+\epsilon)\Pr[E]$ with probability at least $1-\delta$. (ii) By "ignoring" E (i.e., attenuation, in a problem-specific manner) with probability $\sim 1-c/\hat{E}$, we can ensure that the final effective value of $\Pr[E]$ is arbitrarily close to c, i.e., in the range $[c/(1+\epsilon),c]$ with probability at least $1-\delta$. This simple idea of attenuating the probability of an event to come down approximately to a certain value c is what we term simulation-based attenuation. The number of samples needed to obtain the estimate \hat{E} is $\Theta(\frac{1}{c\epsilon^2} \cdot \log(\frac{1}{\delta}))$ via a standard Chernoff-bound argument. In our applications, we will take $\epsilon = 1/\text{poly}(N)$ where N is the problem-size, and the error ϵ will only impact lower-order terms in our approximations.

We adapt the above simulation-based attenuation technique to our setting. Let \mathbf{x}^* denote an optimal solution to LP (1). Suppose we aim to develop an online algorithm achieving a ratio of $\gamma \in [0,1]$. In particular, for every edge $e \in E$, we want to ensure that the $\Pr[e]$ is matched $] \geq \gamma$. Consider an assignment e = (u,v) when some v arrived at time t. Let $\mathrm{SF}_{e,t}$ be the event that e is safe at t, i.e., u is available at t. Using Monte-Carlo simulations, we obtain an approximate estimate $\beta_{e,t}$ of the quantity $\Pr[\mathrm{SF}_{e,t}]$. Note that to obtain $\beta_{e,t}$, we need the estimates $\beta_{e,t'}$ for every t' < t. With the estimated quantity $\beta_{e,t}$, we match a safe edge e at time t with probability $\frac{x_{e,t}^*}{p_{v,t}} \frac{\gamma}{\beta_{e,t}}$. This is a valid probability (i.e., not exceeding 1), if and only if $\gamma \leq \beta_{e,t}$. We denote an algorithm $\operatorname{ADAP}(\gamma)$ to be valid if and only if for every $e \in E$ and every $t \in [T]$ we have $\gamma \leq \beta_{e,t}$.

When the condition $\gamma \leq \beta_{e,t}$ is satisfied for all $e \in E$ and $t \in [T]$, we have that the probability an edge e = (u, v) is matched conditioned on the event that v arrives and u is available is at least $\gamma x_{e,t}^*$. Thus, using linearity of expectation this immediately implies that the expected reward obtained by the algorithm is at least $\gamma \sum_{t \in [T]} \sum_{e \in E} w_e x_{e,t}^*$. Therefore, the competitive ratio is at least γ .

At the outset, this looks similar to the **Inverse Propensity Scoring (IPS)** used in the multiarmed bandit literature [9]. However, there is a key difference between IPS estimates and our estimates. In the bandit literature, one usually scales the value by the probability of playing an action, since this is the *cost* of observing only bandit feedback. However, here we scale by a quantity that depends on the probability of a certain event happening during the *run* of the algorithm, because of playing other actions. The linear program gives a distribution over the edges assuming that all the neighbors are available. Hence, this scaling can be interpreted as the *cost* the algorithm needs to incur when some neighbors are already matched.

The simulation-based attenuation technique has been used previously for other problems, such as stochastic knapsack [34] and stochastic matching [1]. Throughout the analysis, we assume that we know the exact value of $\beta_{e,t} := \Pr[SF_{e,t}]$ for all t and e. (It is easy to see that the sampling error can be folded into a multiplicative factor of $(1-\epsilon)$ in the competitive ratio by standard Chernoff bounds and hence, ignoring it leads to a cleaner presentation.) The formal statement of our algorithm, denoted by $ADAP(\gamma)$, is as follows: For each v and t, let $E_{v,t}$ be the set of safe assignments for v at t.

ALGORITHM 1: A simulation-based adaptive algorithm ADAP(γ)

For each time t, let v denote the request arriving at time t.

If $E_{v,t} = \emptyset$, then reject v; otherwise, choose $e \in E_{v,t}$ with probability $\frac{x_{e,t}^*}{p_{v,t}} \frac{\gamma}{\beta_{e,t}}$ where e = (u,v).

⁴This is called "dumping factor" in Reference [1]. See Appendix B in Reference [1] for a formal treatment.

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Remarks. Our simulation-based adaptive algorithm described above is *non-adaptive* according to our definition. Assume ADAP is valid with respect to a given $\gamma \in (0,1)$. We can first solve the benchmark LP (1) and get an optimal solution $\{x_{e,t}^*\}_{e \in E, t \in [T]}$. Then, by simulating the random arrivals of online agents according to known distributions $\{p_{v,t}\}_{v \in V, t \in [T]}$ and ADAP(γ) itself sequentially from $t = 1, 2, \ldots, T$, we can get an arbitrarily accurate estimate of $\beta_{e,t}$ for each e and t. All these procedures can be done in an offline manner (i.e., before the online process).

LEMMA 5. ADAP(γ) is a valid algorithm (i.e., $\gamma \leq \beta_{e,t}$ for every $e \in E$ and $t \in [T]$) when $\gamma = \frac{1}{2}$.

PROOF. Essentially, we need to show that $\beta_{e,t} \ge \gamma = \frac{1}{2}$ for all e and t. We prove it by induction on t as follows:

When t = 1, $\beta_{e,t} = 1$ for all e = (u,*). Therefore, we are done. Assume for all t' < t, $\beta_{e,t'} \ge 1/2$ and ADAP(γ) is valid for all rounds t'. In other words, we assume each e is assigned with probability *exactly* equal to $x_{e,t'}^* \cdot \frac{1}{2}$ for all t' < t. Now consider a given e = (u,v). Observe that e is unsafe at t iff u is assigned with some v' at t' < t such that the assignment e' = (u,v') makes u unavailable at t. Therefore,

$$1 - \beta_{e,t} = 1 - \Pr[SF_{e,t}] = \sum_{t' < t} \sum_{e \in E_t} \frac{x_{e,t'}^*}{2} \Pr[C_e > t - t'] \le \frac{1}{2}.$$

The last inequality above is due to Constraint (3) in the benchmark LP (1). Therefore, we have $\beta_{e,t} \ge 1/2$ and we are done.

The main Theorem 1 follows directly from Lemmas 4 and 5.

Extension from C_e **to** $C_{e,t}$. Consider the case when the occupation time of u from e is sensitive to t. In other words, each u will be unavailable for $C_{e,t}$ rounds from the assignment e = (u, v) at t. We can accommodate the extension by simply updating the constraints (3) on u in the benchmark LP (1) to the following. We have that $\forall u \in U, t \in [T]$,

$$\sum_{t' < t} \sum_{e \in E_u} x_{e,t'} \Pr[C_{e,t'} > t - t'] + \sum_{e \in E_u} x_{e,t} \le 1.$$
 (5)

The rest of our algorithm remains the same as before. We can verify that (1) LP (1) with constraints (3) replaced by Equation (5) is a valid benchmark; (2) ADAP achieves a competitive ratio of $\frac{1}{2} - \epsilon$ for any given $\epsilon > 0$ for the new model based on the new valid benchmark LP. The modifications to the analysis transfer through in a straightforward way and for brevity we omit the details here.

4 TIGHTNESS OF ONLINE ANALYSIS AGAINST THE BENCHMARK LP

In this section, we prove the main result as stated in Theorem 2. Consider the following example:

EXAMPLE 1. Consider a star graph G = (U, V, E), where U consists of one single node u and $V = \{v_1, v_2\}$. Set T = N + 1. For j = 1, 2 and $t \in [T]$, let $p_{j,t}$ denote the arrival probability of the vertex of type v_j at time t. For t = 1, $p_{11} = 1$ and $p_{21} = 0$. For $2 \le t \le T$, $p_{1,t} = 0$ and $p_{2,t} = 1/N$. In other words, with probability 1 - 1/N, no vertex will arrive (or we can assume a dummy node will arrive) during round $t \ge 2$. Let $w_1 \doteq w_{(u,v_1)} = 1 - 1/N$ and $w_2 \doteq w_{(u,v_2)} = 1$. For $C_1 \doteq C_{(u,v_1)}$, it takes value of T and T with respective probabilities T = T and T and T with probability T.

LEMMA 6. The benchmark LP (1) has an optimal value of 2 - 1/N on Example 1.

PROOF. For ease of exposition, let x_t and y_t denote the probabilities that $e = (u, v_1)$ and $e = v_1$ (u, v_2) get chosen at time t in any offline optimal, respectively. Thus, the updated benchmark LP is as follows:

maximize
$$\sum_{t \in [T]} (1 - 1/N)x_t + \sum_{t \in [T]} y_t$$
 (6)

s.t.
$$x_1 \le 1, y_1 \le 0$$
 (7)

$$x_t \le 0, y_t \le 1/N \qquad \forall t \ge 2 \tag{8}$$

$$x_t \le 0, y_t \le 1/N \qquad \forall t \ge 2$$

$$x_1(1 - 1/N) + y_t \le 1 \qquad \forall t \ge 2$$
(8)

$$0 \le x_t, y_t \le 1 \qquad \forall t \in [T]. \tag{10}$$

We can verfiy that the optimal solution to the above LP is as follows: $x_1 = 1$ and $y_t = 1/N$ for all $t \ge 2$, all the rest are zeros. Therefore, the optimal value is $(1-1/N) + (1/N) \cdot (T-1) = 2-1/N$.

Lemma 7. The optimal (adaptive) online algorithm has an expected performance of 1 + 1/N on Example 1.

PROOF. Let ALG be an optimal online algorithm and suppose that it matches the edge $e = (u, v_1)$ with probability $\alpha \in [0, 1]$ at t = 1. Let E[ALG] be the expected total weight of all matches achieved by ALG. Thus,

$$E[ALG] = \alpha (1 + (1/N) \cdot (1/N) \cdot N) + (1 - \alpha) \cdot (1/N) \cdot N = 1 + \alpha/N.$$

Thus, the optimal online algorithm will choose $\alpha = 1$ and the resultant expected performance is 1 + 1/N.

Proof of Theorem 2.

PROOF. Combining the two lemmas above, we see that the optimal algorithm can achieve a competitive ratio of (1 + 1/N)/(2 - 1/N) on Example 1 w.r.t. to benchmark LP (1). This completes the proof.

The Special Case of Deterministic C_e Being a Constant

Consider a complete bipartite graph G = (U, V, E) where |U| = K, $|V| = n^2$. Suppose we have T=n rounds and $p_{v,t}=\frac{1}{n^2}$ for each v and t. In other words, in each round t, each v is sampled uniformly from V. For each e, let C_e be deterministically equal to K, which implies that each u will be unavailable for a constant K rounds after each assignment. Assume all assignments have a uniform weight (i.e., $w_e = 1$ for all e). Split the whole online process of n rounds into n - K + 1consecutive windows $W = \{W_\ell\}$ such that $W_\ell = \{\ell, \ell+1, \dots, \ell+K-1\}$ for each $1 \le \ell \le n-K+1$. The benchmark LP (1) then reduces to the following:

$$\max \sum_{t \in [T]} \sum_{e \in E} x_{e,t} \tag{11}$$

s.t.
$$\sum_{e \in E_{v}} x_{e,t} \le \frac{1}{n^2} \qquad \forall v \in V, t \in [T]$$
 (12)

$$\sum_{t \in W_{\ell}} \sum_{e \in E_u} x_{e,t} \le 1 \quad \forall u \in U, 1 \le \ell \le n - K + 1$$

$$\tag{13}$$

$$0 \le x_{e,t} \le 1 \qquad \forall e \in E, t \in [T]. \tag{14}$$

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We can verify that an optimal solution to the above LP is as follows: $x_{e,t}^* = 1/(n^2K)$ for all e and t with the optimal objective value of n. We investigate the performance of any optimal non-adaptive algorithm. Notice that the expected arrivals of any v in the full sequence of online arrivals is 1/n. Thus, for any non-adaptive algorithm NADAP, it needs to specify the allocation distribution \mathcal{D}_v for each v during the first arrival. Consider a given NADAP parameterized by $\{\alpha_{u,v} \in [0,1]\}$ for each v and $v \in E_v$ such that $v \in E_v$ and $v \in E_v$ such that $v \in E_v$ and $v \in E_v$ are the first arrival as $v \in E_v$. In other words, NADAP will assign $v \in E_v$ with probability $v \in E_v$ and $v \in E_v$ when $v \in E_v$ are for the first time and $v \in E_v$ and $v \in E_v$ are the first time and $v \in E_v$ are all $v \in E_v$ and $v \in E_v$ are the first time and $v \in E_v$ and $v \in E_v$ are the first time and $v \in E_v$ are the first time and $v \in E_v$ are the first time and $v \in E_v$ and $v \in E_v$ are the first time and $v \in E_v$ and $v \in E_v$ are the first time and $v \in E_v$ are the first t

Let $\beta_u = \sum_{v \in E_u} \alpha_{u,v} * \frac{1}{n^2}$, which is the probability that u is matched in each round if it is safe at the beginning of that round, when running NADAP. Hence,

$$\sum_{u \in U} \beta_u = \sum_{u \in U} \sum_{v \in E_u} \alpha_{u,v} \cdot \frac{1}{n^2} = \sum_{v \in V} \sum_{u \in E_v} \alpha_{u,v} \cdot \frac{1}{n^2} \le 1.$$

Consider a given u with β_u and let $\gamma_{u,t}$ be the probability that u is available at t. Then, the expected number of matches of u after the n rounds is $\sum_t \beta_u \gamma_{u,t}$. We have the recursive inequalities on $\gamma_{u,t}$ as in Lemma 8, with $\gamma_{u,t} = 1$, t = 1.

Lemma 8. $\forall 1 < t \leq n$, we have

$$\gamma_{u,t} + \beta_u \sum_{t-K+1 \le t' < t} \gamma_{u,t'} = 1.$$

PROOF. The inequality for t=1 is due to the fact that u is safe at t=1. For each time t>1, let $SF_{u,t}$ be the event that u is safe at t and $A_{u,t}$ be the event that u is matched at t. Observe that for each window of K time slots, $\{SF_{u,t}, A_{u,t'}, t-K+1 \le t' < t\}$ are mutually exclusive and collectively exhaustive events. Therefore,

$$1 = \Pr[SF_{u,t}] + \sum_{t-K+1 \le t' < t} \Pr[A_{u,t'}]$$

$$= \gamma_{u,t} + \beta_u \sum_{t-K+1 \le t' < t} \gamma_{u,t'}.$$

Note that the OPT of our benchmark LP is n while the performance of NADAP is $\sum_{u} \sum_{t} \beta_{u} \gamma_{u,t}$. The resulting competitive ratio achieved by an optimal NADAP is captured by the following maximization problem:

$$\max \frac{\sum_{u} \sum_{t} \beta_{u} \gamma_{u,t}}{n} \tag{15}$$

s.t.
$$\sum_{u \in U} \beta_u \le 1 \tag{16}$$

$$\gamma_{u,t} + \beta_u \sum_{t-K+1 \le t' < t} \gamma_{u,t'} = 1$$
 $\forall 1 < t \le n, u \in U$ (17)

$$\beta_u \ge 0, \gamma_{u,1} = 1$$
 $\forall u \in U.$ (18)

We prove the following Lemma that implies Theorem 3:

Lemma 9. The optimal value of the program (15) is at most $\frac{1}{2-1/K} + K/n$.

PROOF. Focus on a given vertex $u \in U$. Notice that $\gamma_{u,t} + \beta_u \sum_{t-K+1 \le t' < t} \gamma_{u,t'} = 1$ for all $1 \le t \le n$. Summing both sides over $t \in [n]$, we have the following:

$$\left(1 + \beta_u(K - 1) \right) \sum_{t \in [n]} \gamma_{u,t} = n + \beta_u(K - 1) \gamma_{u,n} + \beta_u(K - 2) \gamma_{u,n-1} + \dots + \beta_u \gamma_{u,n-K+2}$$

$$\leq n + K - 1.$$

Therefore, we have,

$$\sum_{t \in [n]} \gamma_{u,\,t} \leq \frac{n}{1+\beta_u(K-1)} + \frac{K-1}{1+\beta_u(K-1)} \leq \frac{n}{1+\beta_u(K-1)} + \frac{1}{\beta_u}.$$

Define $H_u \doteq \sum_t \beta_u \gamma_{u,t}$. From the above analysis, we have that $H_u \leq \frac{n\beta_u}{1+\beta_u(K-1)} + 1$. Thus, the objective value in the program (15) can be upper-bounded as follows:

$$\frac{\sum_{u}\sum_{t}\beta_{u}\gamma_{u,t}}{n}=\sum_{u\in U}\frac{H_{u}}{n}\leq\sum_{u\in U}\frac{\beta_{u}}{1+\beta_{u}(K-1)}+\frac{K}{n}.$$

We claim that the optimal value to the program (15) can be upper bounded by the following maximization program:

$$\left\{ \max \sum_{u \in [U]} \frac{\beta_u}{1 + \beta_u(K - 1)} + \frac{K}{n} : \sum_{u \in U} \beta_u = 1, \beta_u \ge 0, \forall u \in U \right\}.$$

According to our assumption K = o(n), the second term can be ignored. Let g(x) = x/(1 + x(K-1)). For any $K \ge 2$, it is a concave function, which implies that maximization of g subject to $\sum_{u} \beta_{u} = 1$ will be achieved when all $\beta_{u} = 1/K$. The resultant value is $\frac{1}{2-1/K} + o(1)$. Thus, we are done.

Hardness against the ex-post optimal solution. Here, we show a hardness result against the ex-post optimal solution. Manshadi et al. [36] prove that for the online-matching problem under known IID distributions (but disposable offline vertices), no algorithm can achieve a ratio better than 0.823. Since our setting generalizes this, their hardness result directly applies to our problem as well. It is worth noting that the gap between the LP relaxation (1) and the ex-post optimal solution is currently unknown. Thus, either the hardness result can be improved or the algorithm can use a tighter relaxation of the offline optimal solution.

5 EXPERIMENTS

To validate the approaches presented in this article, we use the New York City Yellow Cabs dataset,⁵ which contains the trip records for trips in Manhattan, Brooklyn, and Queens for the year 2013. The dataset is split into 12 months. For each month, we have numerous records each corresponding to a single trip. Each record has the following structure: We have an anonymized license number that is the primary key corresponding to a car. For privacy purposes, a long string is used as opposed to the actual license number. We then have the time at which the trip was initiated, the time at which the trip ended, and the total time of the trip in seconds. This is followed by the starting coordinates (i.e., latitude and longitude) of the trip and the destination coordinates of the trip.

Assumptions. We make two assumptions specific to our experimental setup: First, we assume that every car starts and ends at the same location for *all* trips that it makes. Initially, we assign every car a location (potentially the same) that corresponds to its *docking* position. On receiving a request, the car leaves from this docking position to the point of pick-up, executes the trip, and returns to this docking position. Second, we assume that **occupation time distributions (OTD)** associated with all matches are identically (and independently) distributed, i.e., $\{C_e\}$ follow the same distribution. Note that this is a much stronger assumption than what we made in the model

⁵http://www.andresmh.com/nyctaxitrips/.

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and is completely inspired by the dataset (see Section 5.2). We test our model on two specific distributions, namely, a *normal* distribution and a *power-law* distribution (see Figure 5). The docking position of each car and parameters associated with each distribution are all learned from the training dataset (described below in the **Training** discussion).

5.1 Experimental Setup

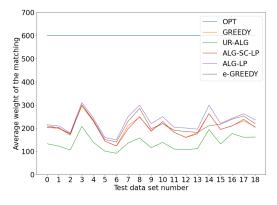
For our experimental setup, we randomly select 30 cabs (each cab is denoted by u). We discretize the Manhattan map into cells such that each cell is approximately 4 miles (increments of 0.15 degrees in latitude and longitude). For each pair of locations, say, (a,b), we create a request type v, which represents all trips with starting and ending locations falling into a and b, respectively. In our model, we have |U|=30 and $|V|\approx550$ (variations depending on day to day requests with low variance). We focus on the month of January 2013. We split the records into 31 parts, each corresponding to a day of January. We choose a random set of 12 parts for training purposes and use the remaining for testing purposes.

The edge weight w_e on e = (u, v) (i.e., edge from a car u to type v) is set as a function of two distances in our setup. The first is the trip distance (i.e., the distance from the starting location to the ending location of v, denoted L_1) while the second is the docking distance (i.e., the distance from the docking position of u to the starting/ending location of v, denoted L_2). We set $w_e = \max(L_1 - \alpha L_2, 0)$, where α is a parameter capturing the subtle balance between the positive contribution from the trip distance and negative contribution from the docking distance to the final profit. We set $\alpha = 0.5$ for the experiments. We consider each single day as the time horizon and set the total number of rounds $T = \frac{24*60}{5} = 288$ by discretizing the 24-hour period into a time-step of 5 minutes. Throughout this section, we use time-step and round interchangeably.

Training. We use the training dataset of 12 days to learn various parameters. As for the arrival rates $\{p_{v,t}\}$, we count the total number of appearances of each request type v at time-step t in the 12 parts (denote it by $c_{v,t}$) and set $p_{v,t} = c_{v,t}/12$ under KAD (note that $c_{v,t}$ is at most 12 and hence this value is always less than 1). When assuming KIID, we set $p_v = p_{v,t} = (c_v/12)/T$ where we have $c_v = \sum_{t \in [T]} c_{v,t}$ (i.e., the arrival distributions are assumed the same across all the time-steps for each v). The estimation of parameters for the two different occupation time distributions are processed as follows: We first compute the average number of seconds between two requests in the dataset (note this was 5 minutes in the experimental setup). We then assume that each time-step of our online process corresponds to a time difference of this average in seconds. We then compute the sample mean and sample variance of the trip lengths (as number of seconds taken by the trip divided by 5 minutes) in the 12 parts. Hence, we use the normal distribution obtained by this sample mean and standard deviation as the distribution with which a car is unavailable. We assign the docking position of each car to the location (in the discretized space) in which the majority of the requests were initiated (i.e., starting location of a request) and matched to this car.

5.2 Justifying the Two Important Model Assumptions

Known Adversarial Distributions. Figure 4 plots the number of arrivals of a particular type at various times during the day. Notice the significant increase in the number of requests in the middle of the day as opposed to the mornings and nights. This justified our arrival assumption of KAD that assumes different arrival distributions at different time-steps. Hence, the LP (and the corresponding algorithm) can exploit this vast difference in the arrival rates and potentially obtain improved results compared to the assumption of **Known Identical Independent Distributions** (**KIID**). This is confirmed by our experimental results shown in Figures 1 and 2.



700

GREEDY

UR-ALG

ALG-SC-LP

ALG-LP

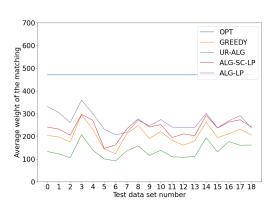
100

0 1 2 3 4 5 6 7 8 9 10 11 12 13 14 15 16 17 18

Test data set number

Fig. 1. OTD is normal distribution under KIID.

Fig. 2. OTD is normal distribution under KAD.



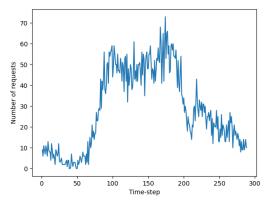


Fig. 3. OTD is power law distribution under KAD.

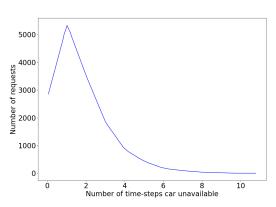
Fig. 4. The number of requests of a given type at various time-steps. x-axis: time-step, y-aixs: number of requests.

Identical-occupation-time Distribution. We assume each car will be available again via an independent and identical random process regardless of the matches it received. The validity of our assumptions can be seen in Figures 5 and 6, where the x-axis represents the different occupation time and the y-axis represents the corresponding number of requests in the dataset responsible for each occupation time. It is clear that for most requests the occupation time is around two to three time-steps and dropping drastically beyond that with a long tail. Figure 6 displays occupation times for two representative (we chose two out of the many cars we plotted, at random) cars in the dataset; we see that the distributions roughly coincide with each other, suggesting that such distributions can be learned from historical data and used as a guide for future matches.

5.3 Results

Inspired by the experimental setup by References [55, 56], we run five different algorithms on our dataset. The first algorithm is the ALG-LP. In this algorithm, when a request v arrives, we choose a neighbor u with probability $x_{e,t}^*/p_{v,t}$ with e=(u,v) if u is available. Here, $x_{e,t}^*$ is an optimal solution to our benchmark LP and $p_{v,t}$ is the arrival rate of type v at time-step t. The second algorithm is called ALG-SC-LP. Recall that $E_{v,t}$ is the set of "safe" or available assignments with

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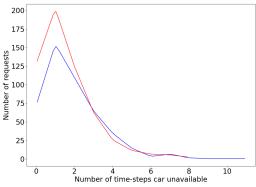


Fig. 5. Occupation time distribution of all cars. x-axis: number of time-steps, y-axis: number of requests.

Fig. 6. Occupation time distribution of two different cars. x-axis: number of time-steps, y-axis: number of requests.

respect to v when the type v arrives at t. Let $x_{v,t} = \sum_{e \in E_{v,t}} x_{e,t}^*$. In ALG-SC-LP, we sample a safe assignment for v with probability $x_{e,t}^*/x_{v,t}$. The next two algorithms are heuristics oblivious to the underlying LP. Our third algorithm is called GREEDY, which is as follows: When a request v comes, match it to the safe neighbor u with the highest edge weight. Our fourth algorithm is called UR-ALG, which chooses one of the safe neighbors uniformly at random. Finally, we use a combination of LP-oblivious algorithm and LP-based algorithm called ϵ -GREEDY. In this algorithm when a type v comes, with probability ϵ , we use the greedy choice and with probability $1 - \epsilon$, we use the optimal LP choice. In our algorithm, we optimized the value of ϵ and set it to $\epsilon = 0.1$. We summarize our results in the following plots: Figures 1, 2, and 3 show the performance of the five algorithms and OPT (optimal value of the benchmark LP) under the different assumptions of the OTD (normal or power law) and online arrives (KIID or KAD). In all three figures, the x-axis represents test dataset number and the y-axis represents average weight of matching.

Discussion. From the figures, it is clear that both the LP-based solutions, namely, ALG-LP and ALG-SC-LP, do better than choosing a free neighbor uniformly at random. Additionally, with distributional assumptions the LP-based solutions outperform greedy algorithm as well. We would like to draw attention to a few interesting details in these results. First, compared to the LP optimal solution, our LP-based algorithms have a competitive ratio in the range of 0.5 to 0.7. We believe this is because of our experimental setup. In particular, we have that the rates are high (> 0.1) only in a few time-steps, while in all other time-steps the rates are very close to 0. This means that it resembles the structure of the *theoretical* worst-case example we showed in Section 4. In future experiments, running our algorithms during *peak* periods (where the request rates are significantly larger than 0) may show that competitive ratios in those cases approach 1. Second, it is surprising that our algorithm is fairly robust to the *actual* distributional assumption we made. In particular, from Figures 2 and 3 it is clear that the difference between the assumption of normal distribution versus power-law distribution for the unavailability of cars is *negligible*. This is important, since it might not be easy to learn the *exact* distribution in many cases (e.g., cases where the sample complexity is high), and this shows that a close approximation will still be as good.

Simulation-based algorithm. We omit the results of the simulation-based algorithm, since the performance was similar to the algorithm without the scaling (i.e., ALG-LP). Here, we briefly describe the implementation details on performing the simulations efficiently in practice. The

estimates are computed even before the start of the algorithm. We first simulate the entire sequence of T requests, δ times. Using these δ samples, we first compute the estimates for the first time-step. We now re-use the same δ samples and the computed estimates in the first time-step to obtain the estimates for the second time-step. Hence in a sequential manner, we compute estimates at time t using the samples from time-steps $1, 2, \ldots, t-1$. The overall runtime of this implementation is $O(\delta T + \delta T \kappa)$, where κ denotes the running time of ADAP in every time-step. Hence during the online phase, the running time of ADAP is same as that of ALG-LP.

6 CONCLUSION AND FUTURE DIRECTIONS

In this work, we provide a model that captures the application of assignment in ride-sharing platforms. One key aspect in our model is to consider the reusable aspect of the offline resources. This helps in modeling many other important applications where agents enter and leave the system multiple times (e.g., organ allocation, crowdsourcing markets [27]). Our work opens several important research directions. The first direction is to generalize the online model to the batch setting (e.g., this subsequent work). In other words, in each round, we assume multiple arrivals from V. This assumption is useful in crowdsourcing markets (for example) where multiple tasks—but not all—become available at some time. The second direction is to consider a Markov model on the driver starting position. In this work, we assumed that each driver returns to her docking position. However, in many ride-sharing systems, drivers start a new trip from the position of the last dropoff. This leads to a Markovian system on the offline types, as opposed to the assumed static types in the present work. Finally, pairing our current work with more-applied stochastic-optimization and reinforcement-learning approaches would be of practical interest to policymakers running taxi and bikeshare services [23, 31, 45, 51, 57]. Following the initial conference publication of this article, subsequent work has appeared that addresses some of the aforementioned directions. The multi-capacitated version of the problem in the batch setting was studied in Reference [32], where the authors devise an algorithm that has a competitive ratio of 0.317. Driver-rider matching in rideshare was modeled as a Markovian system [18]; they show that under a practical condition this system convergences to the stationary distribution very fast. The unweighted multi-capacity version of the problem considered in this article under the adversarial arrival model was studied in Reference [25]. They proved an optimal competitive ratio 1 - 1/e improving over the ratio of 1/2 proved in References [22, 48].

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