## Financial Mathematics

MATH 5870/6870<sup>1</sup> Fall 2021

Le Chen

lzc0090@auburn.edu

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Auburn University
Auburn AL

<sup>&</sup>lt;sup>1</sup>Based on Robert L. McDonald's *Derivatives Markets*. 3rd Ed. Pearson. 2013.

Chapter 18. The Lognormal Distribution

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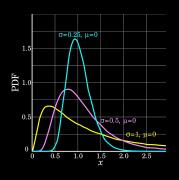
## Chapter 18. The Lognormal Distribution

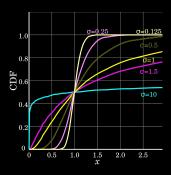
- § 18.1 The normal distribution
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- § 18.4 Lognormal probability calculations
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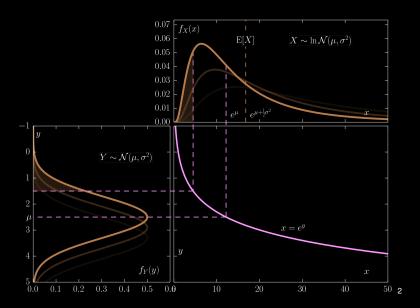
## Chapter 18. The Lognormal Distribution

- § 18.1 The normal distribution
- § 18.2 The lognormal distribution
- § 18.3 A lognormal model of stock prices
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- § 18.5 A. The expectation of a lognormal variable
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- § 18.7 Problems

Definition 18.2-1 A random variable Y is *lognormally distributed* with parameters  $\mu$  and  $\sigma > 0$  if  $\ln(Y) \sim N(\mu, \sigma^2)$ .







<sup>&</sup>lt;sup>2</sup>Image from Wikipedia.

Theorem 18.2-1 If  $Y_1$  and  $Y_2$  are lognormally distributed, so is  $Y_1 Y_2$ .

## Proof.

Since  $Y_1$  and  $Y_2$  are lognormally distributed,  $\ln(Y_1)$  and  $\ln(Y_2)$  are normally distributed. Hence,

$$\ln(\mathbf{Y}_1) + \ln(\mathbf{Y}_2) = \ln(\mathbf{Y}_1 \mathbf{Y}_2)$$

is normally distributed too. Therefore,  $Y_1 Y_2$  is lognormally distributed.

Theorem 18.2-2 If  $\ln(Y) \sim N(\mu, \sigma^2)$ , then the density of Y is given by

$$f_Y(y) = \frac{1}{y\sigma\sqrt{2\pi}}e^{-\frac{1}{2}\left(\frac{\ln(y)-\mu}{\sigma}\right)^2}.$$

Proof. For y > 0,

$$\begin{split} \mathbb{P}(Y \leq y) &= \mathbb{P}\left(\ln(Y) \leq \ln(y)\right) \\ &= \mathbb{P}\left(\frac{\ln(Y) - \mu}{\sigma} \leq \frac{\ln(y) - \mu}{\sigma}\right) \\ &= \Phi\left(\frac{\ln(y) - \mu}{\sigma}\right). \end{split}$$

Hence,

$$f_{Y}(y) = \frac{d}{dy} \Phi\left(\frac{\ln(y) - \mu}{\sigma}\right)$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{\ln(y) - \mu}{\sigma}\right)^{2}} \frac{d}{dy} \frac{\ln(y) - \mu}{\sigma}$$

$$= \frac{1}{v\sigma\sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{\ln(y) - \mu}{\sigma}\right)^{2}}.$$

Recall that the moment generating function of the normal random variable  $X \sim N(\mu, \sigma^2)$  is

$$\mathbb{E}\left(\mathbf{e}^{t\,X}\right) = \mathbf{e}^{\mu t + \sigma^2 t^2/2}, \quad \text{for all } t \in \mathbb{R}.$$

Remark 18.2-1 If Y is lognormally distributed with parameters  $\mu$  and  $\sigma$ , then

$$\mathbb{E}\left(Y^{t}\right)=\pmb{e}^{\mu t+\sigma^{2}t^{2}/2}, \qquad ext{for all } t\in\mathbb{R}.$$

Remark 18.2-2 By Jensen's inequality, if g is a convex function, then

$$\mathbb{E}\left(g(X)\right) \leq g\left(\mathbb{E}(X)\right).$$

Hence, for  $g(x) = e^x$ , we see that

$$\mathbb{E}\left(e^{X}\right) \leq e^{\mathbb{E}(X)}.$$

This is consistent with our computations above because

$$LHS = e^{\mu + \sigma^2/2} \ge e^{\mu} = RHS.$$

Theorem 18.2-3 If Y is lognormally distributed such that  $\ln(Y) \sim N(\mu, \sigma^2)$ , then

$$\mathbb{E}(Y) = e^{\mu + \frac{1}{2}\sigma^2}$$
 and  $\operatorname{Var}(Y) = e^{2\mu + \sigma^2} \left( e^{\sigma^2} - 1 \right)$ .

Proof. Let  $X = \ln(Y)$ . By (1),

$$\mathbb{E}(Y) = \mathbb{E}\left(e^{X}\right) = e^{\mu + \frac{1}{2}\sigma^{2}}$$

and

$$\mathbb{E}(\mathbf{Y}^2) = \mathbb{E}\left(\mathbf{e}^{2X}\right) = \mathbf{e}^{2\mu + 2\sigma^2}.$$

Therefore,

$$\operatorname{Var}(Y) = \mathbb{E}(Y^2) - \mathbb{E}(Y)^2 = e^{2\mu + 2\sigma^2} - e^{2\left(\mu + \frac{1}{2}\sigma^2\right)} = e^{2\mu + \sigma^2} \left(e^{\sigma^2} - 1\right).$$