

Math 362: Mathematical Statistics II

Le Chen

le.chen@emory.edu
chenle02@gmail.com

Emory University
Atlanta, GA

Last updated on Spring 2021
Last compiled on January 15, 2023

2021 Spring

Creative Commons License
(CC By-NC-SA)

Chapter 10. Goodness-of-fit Tests

§ 10.1 Introduction

§ 10.2 The Multinomial Distribution

§ 10.3 Goodness-of-Fit Tests: All Parameters Known

§ 10.4 Goodness-of-Fit Tests: Parameters Unknown

§ 10.5 Contingency Tables

Plan

§ 10.1 Introduction

§ 10.2 The Multinomial Distribution

§ 10.3 Goodness-of-Fit Tests: All Parameters Known

§ 10.4 Goodness-of-Fit Tests: Parameters Unknown

§ 10.5 Contingency Tables

Chapter 10. Goodness-of-fit Tests

§ 10.1 Introduction

§ 10.2 The Multinomial Distribution

§ 10.3 Goodness-of-Fit Tests: All Parameters Known

§ 10.4 Goodness-of-Fit Tests: Parameters Unknown

§ 10.5 Contingency Tables

Rationale

! We want to test if the c.d.f. $F_Y(\cdot)$ is given by the true c.d.f. $F_0(\cdot)$, i.e.,

$$H_0 : F_Y(y) = F_0(y) \quad v.s. \quad H_1 : F_Y(y) \neq F_0(y)$$

~ By properly partitioning the domain, the random sample should follow
an induced multinomial distribution.

\Rightarrow Then testing $F_Y(\cdot) = F_0(\cdot)$ reduces to testing the induced multinomial distribution of the following form:

$$H_0 : p_1 = p'_1, \dots, p_n = p'_n$$

v.s.

$$H_1 : p_i \neq p'_i \quad \text{for at least one } i$$

Rationale

! We want to test if the c.d.f. $F_Y(\cdot)$ is given by the true c.d.f. $F_0(\cdot)$, i.e.,

$$H_0 : F_Y(y) = F_0(y) \quad \text{v.s.} \quad H_1 : F_Y(y) \neq F_0(y)$$

~ By properly partitioning the domain, the random sample should follow
an induced multinomial distribution.

\Rightarrow Then testing $F_Y(\cdot) = F_0(\cdot)$ reduces to testing the induced multinomial distribution of the following form:

$$\begin{aligned} H_0 : p_1 &= p'_1, \dots, p_n = p'_n \\ &\text{v.s.} \\ H_1 : p_i &\neq p'_i \quad \text{for at least one } i \end{aligned}$$

Rationale

! We want to test if the c.d.f. $F_Y(\cdot)$ is given by the true c.d.f. $F_0(\cdot)$, i.e.,

$$H_0 : F_Y(y) = F_0(y) \quad v.s. \quad H_1 : F_Y(y) \neq F_0(y)$$

~ By properly partitioning the domain, the random sample should follow
an induced multinomial distribution.

\Rightarrow Then testing $F_Y(\cdot) = F_0(\cdot)$ reduces to testing the induced multinomial distribution of the following form:

$$\begin{aligned} H_0 : p_1 &= p'_1, \dots, p_n = p'_n \\ &v.s. \\ H_1 : p_i &\neq p'_i \quad \text{for at least one } i \end{aligned}$$

Rationale

! We want to test if the c.d.f. $F_Y(\cdot)$ is given by the true c.d.f. $F_0(\cdot)$, i.e.,

$$H_0 : F_Y(y) = F_0(y) \quad v.s. \quad H_1 : F_Y(y) \neq F_0(y)$$

~ By properly partitioning the domain, the random sample should follow
an induced multinomial distribution.

\Rightarrow Then testing $F_Y(\cdot) = F_0(\cdot)$ reduces to testing the induced multinomial distribution of the following form:

$$H_0 : p_1 = p'_1, \dots, p_n = p'_n$$

v.s.

$$H_1 : p_i \neq p'_i \quad \text{for at least one } i$$

How

1. Suppose we are sampling from the c.d.f. $F(y)$
2. Divide the range of the distribution into k mutually exclusive and exhaustive intervals, say I_1, \dots, I_k .
3. Let $\pi_i = \mathbb{P}(X \in I_i)$, $i = 1, \dots, k$.
4. Let O_1, \dots, O_k be the respective observed numbers of the observations X_1, \dots, X_n in the intervals I_1, \dots, I_k .
5. Then $O = (O_1, \dots, O_k) \sim$ multinomial distribution with (π_1, \dots, π_k) , i.e.,

$$\mathbb{P}(O_1 = o_1, \dots, O_k = o_k) = \frac{n!}{\prod_{i=1}^k o_i!} \prod_{i=1}^k \pi_i^{o_i}$$

with $\sum_{i=1}^k \pi_i = 1$, $\sum_{i=1}^k o_i = n$, and

$$\mathbb{E}[O_i] = n\pi_i =: e_i, \quad \text{Var}(O_i) = n\pi_i(1 - \pi_i)$$

How

1. Suppose we are sampling from the c.d.f. $F(y)$
2. Divide the range of the distribution into k mutually exclusive and exhaustive intervals, say I_1, \dots, I_k .
3. Let $\pi_i = \mathbb{P}(X \in I_i)$, $i = 1, \dots, k$.
4. Let O_1, \dots, O_k be the respective observed numbers of the observations X_1, \dots, X_n in the intervals I_1, \dots, I_k .
5. Then $O = (O_1, \dots, O_k) \sim$ multinomial distribution with (π_1, \dots, π_k) , i.e.,

$$\mathbb{P}(O_1 = o_1, \dots, O_k = o_k) = \frac{n!}{\prod_{i=1}^k o_i!} \prod_{i=1}^k \pi_i^{o_i}$$

with $\sum_{i=1}^k \pi_i = 1$, $\sum_{i=1}^k o_i = n$, and

$$\mathbb{E}[O_i] = n\pi_i =: e_i, \quad \text{Var}(O_i) = n\pi_i(1 - \pi_i)$$

How

1. Suppose we are sampling from the c.d.f. $F(y)$
2. Divide the range of the distribution into k mutually exclusive and exhaustive intervals, say I_1, \dots, I_k .
3. Let $\pi_i = \mathbb{P}(X \in I_i)$, $i = 1, \dots, k$.
4. Let O_1, \dots, O_k be the respective observed numbers of the observations X_1, \dots, X_n in the intervals I_1, \dots, I_k .
5. Then $O = (O_1, \dots, O_k) \sim$ multinomial distribution with (π_1, \dots, π_k) , i.e.,

$$\mathbb{P}(O_1 = o_1, \dots, O_k = o_k) = \frac{n!}{\prod_{i=1}^k o_i!} \prod_{i=1}^k \pi_i^{o_i}$$

with $\sum_{i=1}^k \pi_i = 1$, $\sum_{i=1}^k o_i = n$, and

$$\mathbb{E}[O_i] = n\pi_i =: e_i, \quad \text{Var}(O_i) = n\pi_i(1 - \pi_i)$$

How

1. Suppose we are sampling from the c.d.f. $F(y)$
2. Divide the range of the distribution into k mutually exclusive and exhaustive intervals, say I_1, \dots, I_k .
3. Let $\pi_i = \mathbb{P}(X \in I_i)$, $i = 1, \dots, k$.
4. Let O_1, \dots, O_k be the respective observed numbers of the observations X_1, \dots, X_n in the intervals I_1, \dots, I_k .
5. Then $O = (O_1, \dots, O_k) \sim$ multinomial distribution with (π_1, \dots, π_k) , i.e.,

$$\mathbb{P}(O_1 = o_1, \dots, O_k = o_k) = \frac{n!}{\prod_{i=1}^k o_i!} \prod_{i=1}^k \pi_i^{o_i}$$

with $\sum_{i=1}^k \pi_i = 1$, $\sum_{i=1}^k o_i = n$, and

$$\mathbb{E}[O_i] = n\pi_i =: e_i, \quad \text{Var}(O_i) = n\pi_i(1 - \pi_i)$$

How

1. Suppose we are sampling from the c.d.f. $F(y)$
2. Divide the range of the distribution into k mutually exclusive and exhaustive intervals, say I_1, \dots, I_k .
3. Let $\pi_i = \mathbb{P}(X \in I_i)$, $i = 1, \dots, k$.
4. Let O_1, \dots, O_k be the respective observed numbers of the observations X_1, \dots, X_n in the intervals I_1, \dots, I_k .
5. Then $O = (O_1, \dots, O_k) \sim$ multinomial distribution with (π_1, \dots, π_k) , i.e.,

$$\mathbb{P}(O_1 = o_1, \dots, O_k = o_k) = \frac{n!}{\prod_{i=1}^k o_i!} \prod_{i=1}^k \pi_i^{o_i}$$

with $\sum_{i=1}^k \pi_i = 1$, $\sum_{i=1}^k o_i = n$, and

$$\mathbb{E}[O_i] = n\pi_i =: e_i, \quad \text{Var}(O_i) = n\pi_i(1 - \pi_i)$$

How

1. Suppose we are sampling from the c.d.f. $F(y)$
2. Divide the range of the distribution into k mutually exclusive and exhaustive intervals, say I_1, \dots, I_k .
3. Let $\pi_i = \mathbb{P}(X \in I_i)$, $i = 1, \dots, k$.
4. Let O_1, \dots, O_k be the respective observed numbers of the observations X_1, \dots, X_n in the intervals I_1, \dots, I_k .
5. Then $O = (O_1, \dots, O_k) \sim$ multinomial distribution with (π_1, \dots, π_k) , i.e.,

$$\mathbb{P}(O_1 = o_1, \dots, O_k = o_k) = \frac{n!}{\prod_{i=1}^k o_i!} \prod_{i=1}^k \pi_i^{o_i}$$

with $\sum_{i=1}^k \pi_i = 1$, $\sum_{i=1}^k o_i = n$, and

$$\mathbb{E}[O_i] = n\pi_i =: e_i, \quad \text{Var}(O_i) = n\pi_i(1 - \pi_i)$$

6. When $k = 2$, by CLT, as $n \rightarrow \infty$,

$$\begin{aligned} \frac{O_1 - n\pi_1}{\sqrt{n\pi_1(1 - \pi_1)}} &\xrightarrow{d} N(0, 1) \quad \implies \quad \frac{(O_1 - n\pi_1)^2}{n\pi_1(1 - \pi_1)} \xrightarrow{d} \chi_1^2 \\ &\quad \parallel \\ &\frac{(O_1 - n\pi_1)^2}{n\pi_1} + \frac{(O_2 - n\pi_2)^2}{n\pi_2} \\ &\quad \parallel \\ &\frac{(O_1 - e_1)^2}{e_1} + \frac{(O_2 - e_2)^2}{e_2} \end{aligned}$$

Hence, as $n \rightarrow \infty$,

$$\sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \xrightarrow{d} \chi_{k-1}^2$$

6. When $k = 2$, by CLT, as $n \rightarrow \infty$,

$$\begin{aligned} \frac{O_1 - n\pi_1}{\sqrt{n\pi_1(1 - \pi_1)}} &\xrightarrow{d} N(0, 1) \implies \frac{(O_1 - n\pi_1)^2}{n\pi_1(1 - \pi_1)} \xrightarrow{d} \chi_1^2 \\ &\parallel \\ \frac{(O_1 - n\pi_1)^2}{n\pi_1} + \frac{(O_2 - n\pi_2)^2}{n\pi_2} & \\ &\parallel \\ \frac{(O_1 - e_1)^2}{e_1} + \frac{(O_2 - e_2)^2}{e_2} & \end{aligned}$$

Hence, as $n \rightarrow \infty$,

$$\sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \xrightarrow{d} \chi_{k-1}^2$$

6. When $k = 2$, by CLT, as $n \rightarrow \infty$,

$$\begin{aligned} \frac{O_1 - n\pi_1}{\sqrt{n\pi_1(1 - \pi_1)}} &\xrightarrow{d} N(0, 1) \implies \frac{(O_1 - n\pi_1)^2}{n\pi_1(1 - \pi_1)} \xrightarrow{d} \chi_1^2 \\ &\parallel \\ \frac{(O_1 - n\pi_1)^2}{n\pi_1} + \frac{(O_2 - n\pi_2)^2}{n\pi_2} & \\ &\parallel \\ \frac{(O_1 - e_1)^2}{e_1} + \frac{(O_2 - e_2)^2}{e_2} & \end{aligned}$$

Hence, as $n \rightarrow \infty$,

$$\sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \xrightarrow{d} \chi_{k-1}^2$$

6. When $k = 2$, by CLT, as $n \rightarrow \infty$,

$$\begin{aligned} \frac{O_1 - n\pi_1}{\sqrt{n\pi_1(1 - \pi_1)}} &\xrightarrow{d} N(0, 1) \implies \frac{(O_1 - n\pi_1)^2}{n\pi_1(1 - \pi_1)} \xrightarrow{d} \chi_1^2 \\ &\parallel \\ \frac{(O_1 - n\pi_1)^2}{n\pi_1} + \frac{(O_2 - n\pi_2)^2}{n\pi_2} \\ &\parallel \\ \frac{(O_1 - e_1)^2}{e_1} + \frac{(O_2 - e_2)^2}{e_2} \end{aligned}$$

Hence, as $n \rightarrow \infty$,

$$\sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \xrightarrow{d} \chi_{k-1}^2$$

7. For general k ,

$$\sum_{i=1}^k \frac{(O_i - n\pi_i)^2}{n\pi_i} = \sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i}$$

follows a complicated, but exact, distribution, from which, one can show

$$\sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \xrightarrow{d} \chi_{k-1}^2$$

↓

Thm. When n is large enough, namely, when $n\pi_i \geq 5$ for all i ,

$$D = \sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \underset{\text{appr.}}{\sim} \chi_{k-1}^2.$$

Rmk: The above is called *Pearson's chi-square test*. It is asymptotically equivalent to the generalized likelihood ratio test.

7. For general k ,

$$\sum_{i=1}^k \frac{(O_i - n\pi_i)^2}{n\pi_i} = \sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i}$$

follows a complicated, but exact, distribution, from which, one can show

$$\sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \xrightarrow{d} \chi_{k-1}^2$$

\Downarrow

Thm. When n is large enough, namely, when $n\pi_i \geq 5$ for all i ,

$$D = \sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \underset{\text{appr.}}{\sim} \chi_{k-1}^2.$$

Rmk: The above is called *Pearson's chi-square test*. It is asymptotically equivalent to the generalized likelihood ratio test.

7. For general k ,

$$\sum_{i=1}^k \frac{(O_i - n\pi_i)^2}{n\pi_i} = \sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i}$$

follows a complicated, but exact, distribution, from which, one can show

$$\sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \xrightarrow{d} \chi_{k-1}^2$$

\Downarrow

Thm. When n is large enough, namely, when $n\pi_i \geq 5$ for all i ,

$$D = \sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \underset{\text{appr.}}{\sim} \chi_{k-1}^2.$$

Rmk: The above is called *Pearson's chi-square test*. It is asymptotically equivalent to the generalized likelihood ratio test.

7. For general k ,

$$\sum_{i=1}^k \frac{(O_i - n\pi_i)^2}{n\pi_i} = \sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i}$$

follows a complicated, but exact, distribution, from which, one can show

$$\sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \xrightarrow{d} \chi_{k-1}^2$$

\Downarrow

Thm. When n is large enough, namely, when $n\pi_i \geq 5$ for all i ,

$$D = \sum_{i=1}^k \frac{(O_i - e_i)^2}{e_i} \underset{\text{appr.}}{\sim} \chi_{k-1}^2.$$

Rmk: The above is called *Pearson's chi-square test*. It is asymptotically equivalent to the generalized likelihood ratio test.

Alternative: G-test

– the likelihood ratio test for multinomial model

1. Under $H_0 : \pi_i = p_i, i = 1, \dots, k$, the MLE of π_i are

$$\tilde{\pi}_i = p_i = \frac{np_i}{n} = \frac{e_i}{n}, \quad \forall i.$$

2. When there are no constraints, for $i = 1, \dots, k - 1$,

$$\frac{\partial}{\partial \pi_i} \ln L(\pi_1, \dots, \pi_{k-1} | o_1, \dots, o_k) = 0, \quad 1 \leq i \leq k - 1$$

$$\Updownarrow$$

$$\frac{o_i}{\hat{\pi}_i} = \frac{o_k}{1 - \hat{\pi}_1 - \dots - \hat{\pi}_{k-1}}, \quad 1 \leq i \leq k - 1$$

$$\Updownarrow$$

$$\hat{\pi}_i = \frac{o_i}{n}, \quad 1 \leq i \leq k.$$

Alternative: G-test

– the likelihood ratio test for multinomial model

1. Under $H_0 : \pi_i = p_i, i = 1, \dots, k$, the MLE of π_i are

$$\tilde{\pi}_i = p_i = \frac{np_i}{n} = \frac{e_i}{n}, \quad \forall i.$$

2. When there are no constraints, for $i = 1, \dots, k - 1$,

$$\frac{\partial}{\partial \pi_i} \ln L(\pi_1, \dots, \pi_{k-1} | o_1, \dots, o_k) = 0, \quad 1 \leq i \leq k - 1$$

$$\Updownarrow$$

$$\frac{o_i}{\hat{\pi}_i} = \frac{o_k}{1 - \hat{\pi}_1 - \dots - \hat{\pi}_{k-1}}, \quad 1 \leq i \leq k - 1$$

$$\Updownarrow$$

$$\hat{\pi}_i = \frac{o_i}{n}, \quad 1 \leq i \leq k.$$

\Rightarrow

$$\begin{aligned}\lambda &:= \ln \left(\frac{L(\tilde{\pi}_1, \dots, \tilde{\pi}_{k-1} | \mathbf{o}_1, \dots, \mathbf{o}_k)}{L(\hat{\pi}_1, \dots, \hat{\pi}_{k-1} | \mathbf{o}_1, \dots, \mathbf{o}_k)} \right) = \log \left(\frac{\prod_{i=1}^k \tilde{\pi}_i^{o_i}}{\prod_{i=1}^k \hat{\pi}_i^{o_i}} \right) \\ &= \sum_{i=1}^k o_i \ln \left(\frac{\tilde{\pi}_i}{\hat{\pi}_i} \right) \\ &= \sum_{i=1}^k o_i \ln \left(\frac{\mathbf{e}_i}{\mathbf{o}_i} \right)\end{aligned}$$

Critical region: $\lambda < \lambda_* < 0$.

Def.

$$G := -2\lambda = -2 \sum_{i=1}^k o_i \ln \left(\frac{\mathbf{e}_i}{\mathbf{o}_i} \right) = 2 \sum_{i=1}^k o_i \ln \left(\frac{\mathbf{o}_i}{\mathbf{e}_i} \right)$$

$G \overset{\text{approx.}}{\sim} \chi_{k-1}^2$ for large n .

Critical region: $G \geq G_* = \chi_{1-\alpha, k-1}^2$.

\Rightarrow

$$\begin{aligned}\lambda &:= \ln \left(\frac{L(\tilde{\pi}_1, \dots, \tilde{\pi}_{k-1} | \mathbf{o}_1, \dots, \mathbf{o}_k)}{L(\hat{\pi}_1, \dots, \hat{\pi}_{k-1} | \mathbf{o}_1, \dots, \mathbf{o}_k)} \right) = \log \left(\frac{\prod_{i=1}^k \tilde{\pi}_i^{o_i}}{\prod_{i=1}^k \hat{\pi}_i^{o_i}} \right) \\ &= \sum_{i=1}^k o_i \ln \left(\frac{\tilde{\pi}_i}{\hat{\pi}_i} \right) \\ &= \sum_{i=1}^k o_i \ln \left(\frac{\mathbf{e}_i}{\mathbf{o}_i} \right)\end{aligned}$$

Critical region: $\lambda < \lambda_* < 0$.

Def.

$$G := -2\lambda = -2 \sum_{i=1}^k o_i \ln \left(\frac{\mathbf{e}_i}{\mathbf{o}_i} \right) = 2 \sum_{i=1}^k o_i \ln \left(\frac{\mathbf{o}_i}{\mathbf{e}_i} \right)$$

$G \overset{\text{approx.}}{\sim} \chi_{k-1}^2$ for large n .

Critical region: $G \geq G_* = \chi_{1-\alpha, k-1}^2$.

\Rightarrow

$$\begin{aligned}\lambda &:= \ln \left(\frac{L(\tilde{\pi}_1, \dots, \tilde{\pi}_{k-1} | o_1, \dots, o_k)}{L(\hat{\pi}_1, \dots, \hat{\pi}_{k-1} | o_1, \dots, o_k)} \right) = \log \left(\frac{\prod_{i=1}^k \tilde{\pi}_i^{o_i}}{\prod_{i=1}^k \hat{\pi}_i^{o_i}} \right) \\ &= \sum_{i=1}^k o_i \ln \left(\frac{\tilde{\pi}_i}{\hat{\pi}_i} \right) \\ &= \sum_{i=1}^k o_i \ln \left(\frac{e_i}{o_i} \right)\end{aligned}$$

Critical region: $\lambda < \lambda_* < 0$.

Def.

$$G := -2\lambda = -2 \sum_{i=1}^k o_i \ln \left(\frac{e_i}{o_i} \right) = 2 \sum_{i=1}^k o_i \ln \left(\frac{o_i}{e_i} \right)$$

$G \overset{\text{approx.}}{\sim} \chi_{k-1}^2$ for large n .

Critical region: $G \geq G_* = \chi_{1-\alpha, k-1}^2$.

\Rightarrow

$$\begin{aligned}\lambda &:= \ln \left(\frac{L(\tilde{\pi}_1, \dots, \tilde{\pi}_{k-1} | o_1, \dots, o_k)}{L(\hat{\pi}_1, \dots, \hat{\pi}_{k-1} | o_1, \dots, o_k)} \right) = \log \left(\frac{\prod_{i=1}^k \tilde{\pi}_i^{o_i}}{\prod_{i=1}^k \hat{\pi}_i^{o_i}} \right) \\ &= \sum_{i=1}^k o_i \ln \left(\frac{\tilde{\pi}_i}{\hat{\pi}_i} \right) \\ &= \sum_{i=1}^k o_i \ln \left(\frac{e_i}{o_i} \right)\end{aligned}$$

Critical region: $\lambda < \lambda_* < 0$.

Def.

$$G := -2\lambda = -2 \sum_{i=1}^k o_i \ln \left(\frac{e_i}{o_i} \right) = 2 \sum_{i=1}^k o_i \ln \left(\frac{o_i}{e_i} \right)$$

$G \overset{\text{approx.}}{\sim} \chi_{k-1}^2$ for large n .

Critical region: $G \geq G_* = \chi_{1-\alpha, k-1}^2$.

\Rightarrow

$$\begin{aligned}\lambda &:= \ln \left(\frac{L(\tilde{\pi}_1, \dots, \tilde{\pi}_{k-1} | \mathbf{o}_1, \dots, \mathbf{o}_k)}{L(\hat{\pi}_1, \dots, \hat{\pi}_{k-1} | \mathbf{o}_1, \dots, \mathbf{o}_k)} \right) = \log \left(\frac{\prod_{i=1}^k \tilde{\pi}_i^{o_i}}{\prod_{i=1}^k \hat{\pi}_i^{o_i}} \right) \\ &= \sum_{i=1}^k o_i \ln \left(\frac{\tilde{\pi}_i}{\hat{\pi}_i} \right) \\ &= \sum_{i=1}^k o_i \ln \left(\frac{e_i}{o_i} \right)\end{aligned}$$

Critical region: $\lambda < \lambda_* < 0$.

Def.

$$G := -2\lambda = -2 \sum_{i=1}^k o_i \ln \left(\frac{e_i}{o_i} \right) = 2 \sum_{i=1}^k o_i \ln \left(\frac{o_i}{e_i} \right)$$

$G \overset{\text{approx.}}{\sim} \chi_{k-1}^2$ for large n .

Critical region: $G \geq G_* = \chi_{1-\alpha, k-1}^2$.

Relation G-test and Pearson's Chi square test

By second order Taylor expansion around 1,

$$\begin{aligned} G &= -2 \sum_{i=1}^k o_i \ln \left(\frac{e_i}{o_i} \right) \\ &\approx -2 \sum_{i=1}^k o_i \left[\left(\frac{e_i}{o_i} - 1 \right) - \frac{1}{2} \left(\frac{e_i}{o_i} - 1 \right)^2 \right] \\ &= -2 \sum_{i=1}^k (e_i - o_i) + \sum_{i=1}^k o_i \left(\left(1 - \frac{o_i}{e_i} \right) + \frac{o_i}{e_i} \right) \left(\frac{e_i}{o_i} - 1 \right)^2 \\ &= 0 + \sum_{i=1}^n \frac{o_i^2}{e_i} \left(1 - \frac{o_i}{e_i} \right)^3 + \sum_{i=1}^k \frac{(e_i - o_i)^2}{e_i} \\ &\approx \sum_{i=1}^k \frac{(e_i - o_i)^2}{e_i} \\ &\quad \parallel \\ &\quad D \end{aligned}$$

∴ Pearson's Chi-square test is an approximation of G-test.

Relation G-test and Pearson's Chi square test

By second order Taylor expansion around 1,

$$\begin{aligned} G &= -2 \sum_{i=1}^k o_i \ln \left(\frac{e_i}{o_i} \right) \\ &\approx -2 \sum_{i=1}^k o_i \left[\left(\frac{e_i}{o_i} - 1 \right) - \frac{1}{2} \left(\frac{e_i}{o_i} - 1 \right)^2 \right] \\ &= -2 \sum_{i=1}^k (e_i - o_i) + \sum_{i=1}^k o_i \left(\left(1 - \frac{o_i}{e_i} \right) + \frac{o_i}{e_i} \right) \left(\frac{e_i}{o_i} - 1 \right)^2 \\ &= 0 + \sum_{i=1}^n \frac{o_i^2}{e_i} \left(1 - \frac{o_i}{e_i} \right)^3 + \sum_{i=1}^k \frac{(e_i - o_i)^2}{e_i} \\ &\approx \sum_{i=1}^k \frac{(e_i - o_i)^2}{e_i} \\ &\quad \parallel \\ &\quad D \end{aligned}$$

∴ Pearson's Chi-square test is an approximation of G-test.

Relation G-test and Pearson's Chi square test

By second order Taylor expansion around 1,

$$\begin{aligned} G &= -2 \sum_{i=1}^k o_i \ln \left(\frac{e_i}{o_i} \right) \\ &\approx -2 \sum_{i=1}^k o_i \left[\left(\frac{e_i}{o_i} - 1 \right) - \frac{1}{2} \left(\frac{e_i}{o_i} - 1 \right)^2 \right] \\ &= -2 \sum_{i=1}^k (e_i - o_i) + \sum_{i=1}^k o_i \left(\left(1 - \frac{o_i}{e_i} \right) + \frac{o_i}{e_i} \right) \left(\frac{e_i}{o_i} - 1 \right)^2 \\ &= 0 + \sum_{i=1}^n \frac{o_i^2}{e_i} \left(1 - \frac{o_i}{e_i} \right)^3 + \sum_{i=1}^k \frac{(e_i - o_i)^2}{e_i} \\ &\approx \sum_{i=1}^k \frac{(e_i - o_i)^2}{e_i} \\ &\quad \parallel \\ &\quad D \end{aligned}$$

∴ Pearson's Chi-square test is an approximation of G-test.

Relation G-test and Pearson's Chi square test

By second order Taylor expansion around 1,

$$\begin{aligned} G &= -2 \sum_{i=1}^k o_i \ln \left(\frac{e_i}{o_i} \right) \\ &\approx -2 \sum_{i=1}^k o_i \left[\left(\frac{e_i}{o_i} - 1 \right) - \frac{1}{2} \left(\frac{e_i}{o_i} - 1 \right)^2 \right] \\ &= -2 \sum_{i=1}^k (e_i - o_i) + \sum_{i=1}^k o_i \left(\left(1 - \frac{o_i}{e_i} \right) + \frac{o_i}{e_i} \right) \left(\frac{e_i}{o_i} - 1 \right)^2 \\ &= 0 + \sum_{i=1}^n \frac{o_i^2}{e_i} \left(1 - \frac{o_i}{e_i} \right)^3 + \sum_{i=1}^k \frac{(e_i - o_i)^2}{e_i} \\ &\approx \sum_{i=1}^k \frac{(e_i - o_i)^2}{e_i} \\ &\quad || \\ &\quad D \end{aligned}$$

∴ Pearson's Chi-square test is an approximation of G-test.

E.g. 1 *Benford's law*:

Initial digits

Use this law to check whether the bookkeepers have made up entries.

Assume that bookkeepers are not aware of Benford's law.

E.g. 1 *Benford's law*:

Table 10.3.1	
Digit, i	$\log_{10}(i + 1) - \log_{10}(i)$
1	0.301
2	0.176
3	0.125
4	0.097
5	0.079
6	0.067
7	0.058
8	0.051
9	0.046

Initial digits

Digit	Observed, k_i
1	111
2	60
3	46
4	29
5	26
6	22
7	21
8	20
9	20
	<hr/> 355

Use this law to check whether the bookkeepers have made up entries.

Assume that bookkeepers are not aware of Benford's law.

Sol. The test should be

$$H_0 : p_1 = p_{10}, \dots, p_9 = p_{90}$$

v.s.

$$H_1 : p_i \neq p_{i0} \quad \text{for at least one } i = 1, \dots, 9.$$

Critical region: $(\chi^2_{95,8}, \infty) = (15.507, \infty)$.

Sol. The test should be

$$H_0 : p_1 = p_{10}, \dots, p_9 = p_{90}$$

v.s.

$$H_1 : p_i \neq p_{i0} \quad \text{for at least one } i = 1, \dots, 9.$$

Critical region: $(\chi^2_{.95,8}, \infty) = (15.507, \infty)$.

Compute the D and G scores:

Digit	o_i	p_i	e_i	$(o_i - e_i)^2 / e_i$	$2o_i \ln(e_i / o_i)$
1	111	0.301			
2	60	0.176			
3	46	0.125			
4	29	0.097			
5	26	0.079			
6	22	0.067			
7	21	0.058			
8	20	0.051			
9	20	0.046			
sum	355	1	355	$d = \underline{\hspace{2cm}}$	$g = \underline{\hspace{2cm}}$

Digit	o_i	p_i	e_i	$(o_i - e_i)^2 / e_i$	$2o_i \ln(e_i / o_i)$
1	111	0.301	106.9	0.16	8.449
2	60	0.176	62.5	0.10	-4.860
3	46	0.125	44.4	0.06	3.309
4	29	0.097	34.4	0.86	-9.963
5	26	0.079	28.0	0.15	-3.937
6	22	0.067	23.8	0.13	-3.433
7	21	0.058	20.6	0.01	0.828
8	20	0.051	18.1	0.20	3.982
9	20	0.046	16.3	0.82	8.109
sum	355	1	355	$d = \underline{2.49}$	$g = \underline{2.48}$

Conclusion: Fail to reject.

```

1 > # EX 10.3.2
2 > library(data.table)
3 > mydat <- fread('http://math.emory.edu/~lchen41/teaching/2020_Spring/Case_10-3-2.data')
4 trying URL 'http://math.emory.edu/~lchen41/teaching/2020_Spring/Case_10-3-2.data'
5 Content type 'unknown' length 153 bytes
6 =====
7 downloaded 153 bytes
8
9 > head(mydat)
10      Digit  Oi   Pi
11 1:      1 111 0.301
12 2:      2  60 0.176
13 3:      3  46 0.125
14 4:      4  29 0.097
15 > pi = mydat[,3]
16 > oi = mydat[,2]
17 > n = sum(oi)
18 > ei = n*pi
19 > di = (ei-oi)^2/ei
20 > gi = 2*oi*log(oi/ei)
21 > print(paste("Using Pearson's test, D value is equal to ", round(sum(di),3)))
22 [1] "Using Pearson's test, D value is equal to  2.491"
23 > print(paste("Using the G-test, G value is equal to ", round(sum(gi),3)))
24 [1] "Using the G-test, G value is equal to  2.484"

```

Codes available

http://math.emory.edu/~lchen41/teaching/2020_Spring/Case_10-3-2.R

E.g. 2 Test for randomness

Is the following sample of size 40 from $f_Y(y) = 6y(1 - y)$, $y \in [0, 1]$?

E.g. 2 Test for randomness

Is the following sample of size 40 from $f_Y(y) = 6y(1 - y)$, $y \in [0, 1]$?

Table 10.3.4				
0.18	0.06	0.27	0.58	0.98
0.55	0.24	0.58	0.97	0.36
0.48	0.11	0.59	0.15	0.53
0.29	0.46	0.21	0.39	0.89
0.34	0.09	0.64	0.52	0.64
0.71	0.56	0.48	0.44	0.40
0.80	0.83	0.02	0.10	0.51
0.43	0.14	0.74	0.75	0.22

Sol. Test continuous pdf \rightarrow reduce to a set of classes:

$$d = \dots = 1.84.$$

Critical region: $\{x : \chi^2 = 15.992, \alpha = 0.05\}$

Conclusion: Fail to reject.

Sol. Test continuous pdf \rightarrow reduce to a set of classes:

Table 10.3.5			
Class	Observed Frequency, k_i	P_{i_0}	$40 p_{i_0}$
$0 \leq y < 0.20$	8	0.104	4.16
$0.20 \leq y < 0.40$	8	0.248	9.92
$0.40 \leq y < 0.60$	14	0.296	11.84
$0.60 \leq y < 0.80$	5	0.248	9.92
$0.80 \leq y < 1.00$	5	0.104	4.16

$$d = \cdots = 1.84.$$

Critical region: $\chi^2_{0.05, 4} = 9.488$ (see Table 10.3.6)

Conclusion: Fail to reject.

Sol. Test continuous pdf \rightarrow reduce to a set of classes:

Table 10.3.5			
Class	Observed Frequency, k_i	P_{i_o}	$40 p_{i_o}$
$0 \leq y < 0.20$	8	0.104	4.16
$0.20 \leq y < 0.40$	8	0.248	9.92
$0.40 \leq y < 0.60$	14	0.296	11.84
$0.60 \leq y < 0.80$	5	0.248	9.92
$0.80 \leq y < 1.00$	5	0.104	4.16

Table 10.3.6			
Class	Observed Frequency, k_i	P_{i_o}	$40 p_{i_o}$
$0 \leq y < 0.40$	16	0.352	14.08
$0.40 \leq y < 0.60$	14	0.296	11.84
$0.60 \leq y \leq 1.00$	10	0.352	14.08

$$d = \dots = 1.84.$$

Critical region: $\chi^2 \geq \chi^2_{0.05, 2} = 5.991$

Conclusion: Fail to reject

Sol. Test continuous pdf \rightarrow reduce to a set of classes:

Table 10.3.5			
Class	Observed Frequency, k_i	P_{i_o}	$40 p_{i_o}$
$0 \leq y < 0.20$	8	0.104	4.16
$0.20 \leq y < 0.40$	8	0.248	9.92
$0.40 \leq y < 0.60$	14	0.296	11.84
$0.60 \leq y < 0.80$	5	0.248	9.92
$0.80 \leq y < 1.00$	5	0.104	4.16

Table 10.3.6			
Class	Observed Frequency, k_i	P_{i_o}	$40 p_{i_o}$
$0 \leq y < 0.40$	16	0.352	14.08
$0.40 \leq y < 0.60$	14	0.296	11.84
$0.60 \leq y \leq 1.00$	10	0.352	14.08

$$d = \dots = 1.84.$$

Critical region: $(\chi^2_{.95,2}, \infty) = (5.992, \infty)$.

Conclusion: Fail to reject.

Sol. Test continuous pdf \rightarrow reduce to a set of classes:

Table 10.3.5			
Class	Observed Frequency, k_i	P_{i_o}	$40 p_{i_o}$
$0 \leq y < 0.20$	8	0.104	4.16
$0.20 \leq y < 0.40$	8	0.248	9.92
$0.40 \leq y < 0.60$	14	0.296	11.84
$0.60 \leq y < 0.80$	5	0.248	9.92
$0.80 \leq y < 1.00$	5	0.104	4.16

Table 10.3.6			
Class	Observed Frequency, k_i	P_{i_o}	$40 p_{i_o}$
$0 \leq y < 0.40$	16	0.352	14.08
$0.40 \leq y < 0.60$	14	0.296	11.84
$0.60 \leq y \leq 1.00$	10	0.352	14.08

$$d = \dots = 1.84.$$

Critical region: $(\chi^2_{.95,2}, \infty) = (5.992, \infty)$.

Conclusion: Fail to reject.

```

1 > # Case Study 10.3.2
2 > # Read data from the URL link
3 > library (data.table)
4 > mydat <- fread('http://math.emory.edu/~lchen41/teaching/2020_Spring/EX_10-3-1.data')
5 trying URL 'http://math.emory.edu/~lchen41/teaching/2020_Spring/EX_10-3-1.data'
6 Content type 'unknown' length 234 bytes
7 =====
8 downloaded 234 bytes
9
10 >d(mydat)
11   Col1 Col2 Col3 Col4 Col5
12 1: 0.18 0.06 0.27 0.58 0.98
13 2: 0.55 0.24 0.58 0.97 0.36
14 3: 0.48 0.11 0.59 0.15 0.53
15 4: 0.29 0.46 0.21 0.39 0.89
16 5: 0.34 0.09 0.64 0.52 0.64
17 6: 0.71 0.56 0.48 0.44 0.40
18 # Conditions for lower bounds
19 > lb=c(0,0.40,0.60)
20 > # Conditions for upper bounds
21 > up=c(0.40,0.60,1.00)
22 > # Store the results in d
23 > oi <- seq(1:length(lb))
24 > pi <- seq(1:length(lb))
25 > integrand <- function(y) {6*y*(1-y)}
26 > for (i in c(1:length(lb))) {
27 +   oi[i] <- table(mydat>=lb[i] & mydat<up[i])[2]
28 +   pi[i] <- integrate(integrand, lb[i], up[i])$value[1]
29 +   print(paste("the", i,"th bin has", oi[i],
30 +     "entries and pi is equal to", pi[i]))
31 + }

```



```

1 [1] "the 1 th bin has 16 entries and pi is equal to 0.352"
2 [1] "the 2 th bin has 14 entries and pi is equal to 0.296"
3 [1] "the 3 th bin has 10 entries and pi is equal to 0.352"
4 > pi <- unlist(pi)
5 > n <- sum(oi)
6 > ei <- n*pi
7 > di <- (ei-oi)^2/ei
8 > gi <- 2*oi*log(oi/ei)
9 > rbind(oi,pi,ei,di,gi)
10      [,1]      [,2]      [,3]
11 oi 16.0000000 14.0000000 10.0000000
12 pi  0.3520000 0.2960000 0.3520000
13 ei 14.0800000 11.8400000 14.0800000
14 di  0.2618182 0.3940541 1.182273
15 gi  4.0906679 4.6920636 -6.843405
16 > print(paste("Using Pearson's test, D value is equal to ",round(sum(di),3)))
17 [1] "Using Pearson's test, D value is equal to  1.838"
18 > print(paste("Using the G-test, G value is equal to ", round(sum(gi),3)))
19 [1] "Using the G-test, G value is equal to  1.939"<Paste>

```

http://math.emory.edu/~lchen41/teaching/2020_Spring/EX_10-3-1.R

E.g. 3 Fisher's suspicion on Mendel's experiments on 1866:

$$d = \dots = 0.47$$

$$P\text{-value} = \mathbb{P}(\chi_3^2 \leq 0.47) = 0.0746.$$

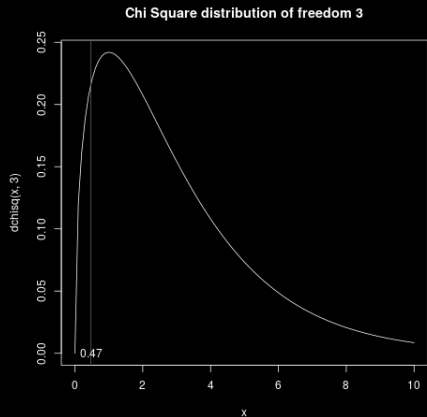
E.g. 3 Fisher's suspicion on Mendel's experiments on 1866:

Table 10.3.7			
Phenotype	Obs. Freq.	Mendel's Model	Exp. Freq.
(round, yellow)	315	9/16	312.75
(round, green)	108	3/16	104.25
(angular, yellow)	101	3/16	104.25
(angular, green)	32	1/16	34.75

$$d = \dots = 0.47$$

$$P\text{-value} = \mathbb{P}(\chi_3^2 \leq 0.47) = 0.0746.$$

```
1 > # Case Study 10.3.3
2 > x=seq(0,10,0.1)
3 > plot(x,dchisq(x,3),type = "l")
4 > abline(v=0.47,col = "gray60")
5 > text(0.47,0,"0.47")
6 > title("Chi Square distribution
7 +       of freedom 3")
8 > pchisq(0.47,3)
9 [1] 0.07456892
```



E.g. 2' A second look at the random generator in E.g. 2.

Does it fit the model too well? Find the P -value.

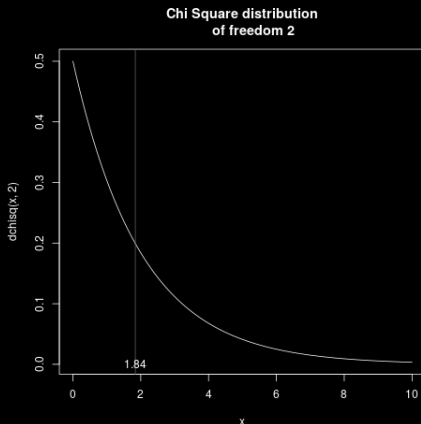
```
1 > # Example 10.3.1
2 > x=seq(0,10,0.1)
3 > plot(x,dchisq(x,2),type="l")
4 > abline(v=1.84,col="gray60")
5 > text(1.84,0,"1.84")
6 > title("Chi Square distribution
7 +      of freedom 2")
8 > pchisq(1.84,2)
9 [1] 0.601481
```

$P\text{-value} = 0.601 \implies \text{No.}$

E.g. 2' A second look at the random generator in E.g. 2.

Does it fit the model too well? Find the P -value.

```
1 > # Example 10.3.1
2 > x=seq(0,10,0.1)
3 > plot(x,dchisq(x,2),type="l")
4 > abline(v=1.84,col="gray60")
5 > text(1.84,0,"1.84")
6 > title("Chi Square distribution
7 +       of freedom 2")
8 > pchisq(1.84,2)
9 [1] 0.601481
```



$P\text{-value} = 0.601 \implies \text{No.}$