

# Math 362: Mathematical Statistics II

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# Chapter 6. Hypothesis Testing

## § 6.1 Introduction

## § 6.2 The Decision Rule

## § 6.3 Testing Binomial Data – $H_0 : p = p_0$

## § 6.4 Type I and Type II Errors

## § 6.5 A Notion of Optimality: The Generalized Likelihood Ratio

# Plan

§ 6.1 Introduction

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§ 6.4 Type I and Type II Errors

§ 6.5 A Notion of Optimality: The Generalized Likelihood Ratio

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## § 6.1 Introduction

## § 6.2 The Decision Rule

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## § 6.4 Type I and Type II Errors

## § 6.5 A Notion of Optimality: The Generalized Likelihood Ratio

	True State of Nature	
	$H_0$ is true	$H_1$ is true
Fail to reject $H_0$	Correct	Type II error
Reject $H_0$	Type I error	Correct

Table of error types		Null hypothesis ( $H_0$ ) is	
		True	False
Decision about null hypothesis ( $H_0$ )	Don't reject	Correct inference (true negative) (probability = $1 - \alpha$ )	Type II error (false negative) (probability = $\beta$ )
	Reject	Type I error (false positive) (probability = $\alpha$ )	Correct inference (true positive) (probability = $1 - \beta$ )

## Type I error $\sim \alpha$

$$\alpha := \mathbb{P}(\text{Type I error}) = \mathbb{P}(\text{Reject } H_0 | H_0 \text{ is true})$$

By convention,  $H_0$  is always of the form, e.g.,  $\mu = \mu_0$ . So this probability can be exactly determined. It is equal to the level of significance  $\alpha$ .

(Simple null test)

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## Type II error $\sim \beta$

$$\beta := \mathbb{P}(\text{Type II error}) = \mathbb{P}(\text{Fail to reject } H_0 | H_1 \text{ is true})$$

In order to compute Type II error, we need to specify a concrete alternative hypothesis.

Figure: One-sided inference  $H_1 : \mu > \mu_0$

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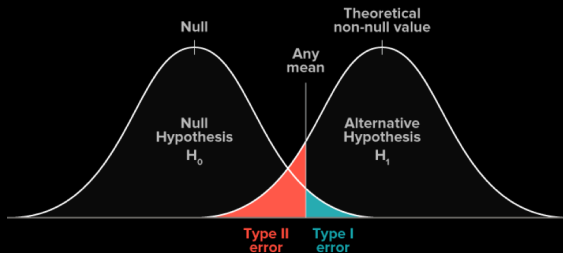


Figure: One-sided inference  $H_1 : \mu > \mu_0$

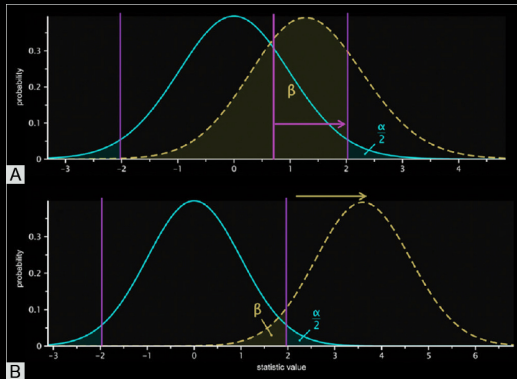
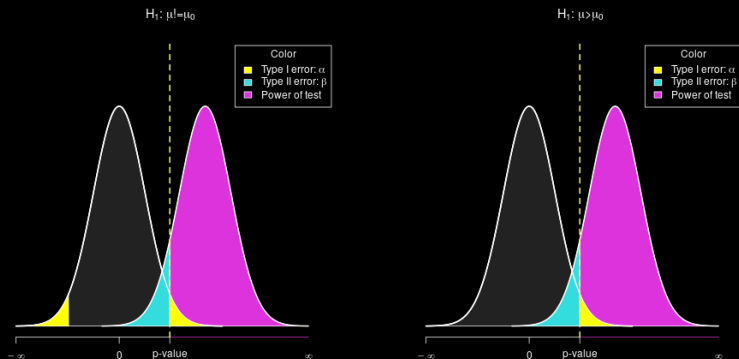


Figure: Two-sided inference  $H_1 : \mu \neq \mu_0$

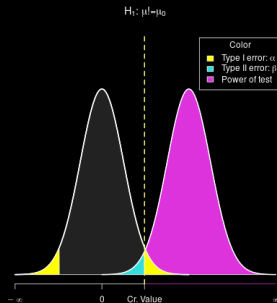
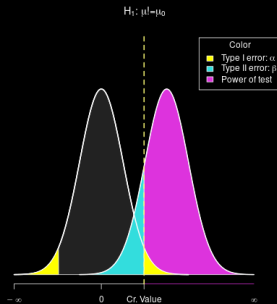
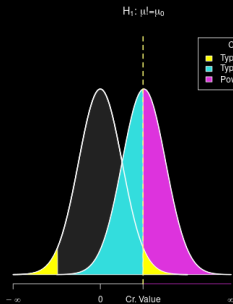
# Power of test $1 - \beta$

$$\text{Power of test} = \mathbb{P}(\text{Reject } H_0 | H_1 \text{ is true}) = 1 - \beta$$

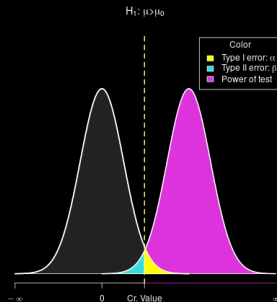
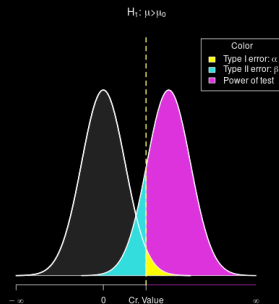
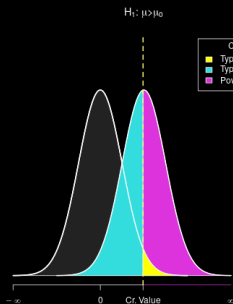


One online interactive show all  $\alpha$ ,  $\beta$  and  $1 - \beta$ :  
<https://rpsychologist.com/d3/NHST/>

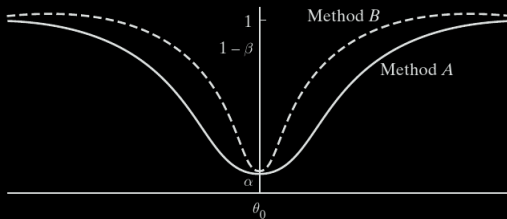
# Two-sided test



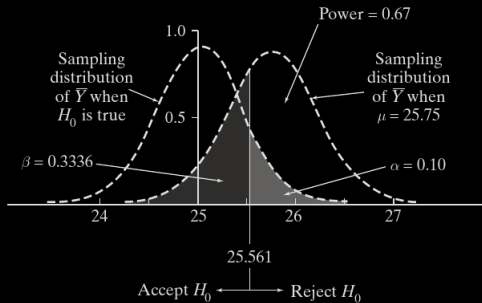
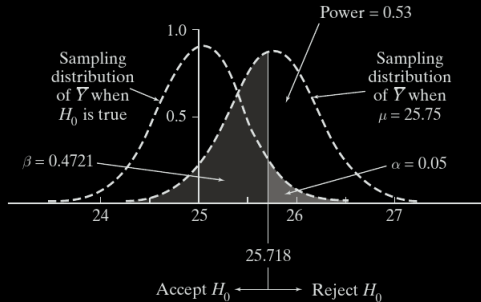
# One-sided test



Use the **power curves** to select methods  
(steepest one!)



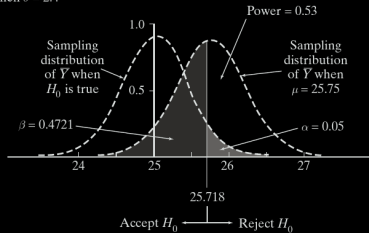
$$\alpha \uparrow \implies \beta \downarrow \text{ and } (1 - \beta) \uparrow$$



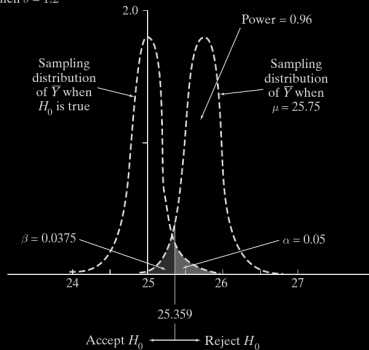


$$\sigma \downarrow \implies \beta \downarrow \text{ and } (1 - \beta) \uparrow$$

When  $\sigma = 2.4$



When  $\sigma = 1.2$



One usually cannot control the given parameter  $\sigma$ . But one can achieve the same power of test by increasing the sample size  $n$ .

E.g. Test  $H_0: \mu = \mu_0$  vs.  $H_1: \mu > \mu_0$  at  $\alpha = 0.05$  with  $\sigma = 1$  known.  
Requirement:  $\beta \leq 0.05$  if  $\mu = \mu_1$ .  
Find smallest sample size  $n$ .

Remark: Two conditions:  $\alpha = 0.05$  and  $\beta \leq 0.05$ .  
Two unknowns: Critical value  $c$  and sample size  $n$ .

Sol.

$$0 = \left| z - z_{1-\alpha} = \frac{\bar{x} - \mu_0}{\frac{1}{\sqrt{n}}} = \frac{\bar{x} - \mu_1}{\frac{1}{\sqrt{n}}} = z \right|$$

One usually cannot control the given parameter  $\sigma$ . But one can achieve the same power of test by increasing the sample size  $n$ .

E.g. Test  $H_0 : \mu = 100$  v.s.  $H_1 : \mu > 100$  at  $\alpha = 0.05$  with  $\sigma = 14$  known.  
Requirement:  $1 - \beta = 0.60$  when  $\mu = 103$ .  
Find smallest sample size  $n$ .

Remark: Two conditions:  $\alpha = 0.05$  and  $1 - \beta = 0.60$

Two unknowns: Critical value  $y^*$  and sample size  $n$

Sol.

$$0 = \left[ z - z_{\alpha} + \frac{y^* - \mu_0}{\sigma/\sqrt{n}} - z_{\beta} \right]$$

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Requirement:  $1 - \beta = 0.60$  when  $\mu = 103$ .  
Find smallest sample size  $n$ .

Remark: Two conditions:  $\alpha = 0.05$  and  $1 - \beta = 0.60$   
Two unknowns: Critical value  $y^*$  and sample size  $n$

Sol.

$$C = \left\{ z : z = \frac{\bar{y} - \mu_0}{\sigma/\sqrt{n}} \geq z_\alpha \right\}.$$

$$\begin{aligned}
1 - \beta &= \mathbb{P} \left( \frac{\bar{Y} - \mu_0}{\sigma/\sqrt{n}} \geq z_\alpha \mid \mu_1 \right) \\
&= \mathbb{P} \left( \frac{\bar{Y} - \mu_1}{\sigma/\sqrt{n}} + \frac{\mu_1 - \mu_0}{\sigma/\sqrt{n}} \geq z_\alpha \mid \mu_1 \right) \\
&= \mathbb{P} \left( Z \geq -\frac{\mu_1 - \mu_0}{\sigma/\sqrt{n}} + z_\alpha \mid \mu_1 \right) \\
&= \Phi \left( \frac{\mu_1 - \mu_0}{\sigma/\sqrt{n}} - z_\alpha \right)
\end{aligned}$$

$$\frac{\mu_1 - \mu_0}{\sigma/\sqrt{n}} - z_\alpha = \Phi^{-1}(1 - \beta) \quad \Longleftrightarrow \quad n = \left( \sigma \times \frac{\Phi^{-1}(1 - \beta) + z_\alpha}{\mu_1 - \mu_0} \right)^2$$

$$n = \left\lceil \left( 14 \times \frac{0.2533 + 1.645}{103 - 100} \right)^2 \right\rceil = \lceil 78.48 \rceil = 79.$$

□

R  
 $z_\alpha = \text{qnorm}(1 - \alpha)$   
 $\Phi^{-1}(1 - \beta) = \text{qnorm}(1 - \beta)$

Python  
 $z_\alpha = \text{scipy.stats.norm.ppf}(1 - \alpha)$   
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$$\begin{aligned}
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R

```

z_alpha = qnorm(1 - alpha)
Phi_inv(1 - beta) = qnorm(1 - beta)

```

Python

```

z_alpha = scipy.stats.norm.ppf(1 - alpha)
Phi_inv(1 - beta) = scipy.stats.norm.ppf(1 - beta)

```

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# Nonnormal data

Test  $H_0 : \theta = \theta_0$ , with  $f_Y(y; \theta)$  is not normal distribution.

1. Identify a sufficient estimator  $\hat{\theta}$  for  $\theta$ .
2. Find the critical region  $C$ : Least compatible with  $H_0$ , but still admissible under  $H_1$ .
3. Three types of questions:
  - Given  $\alpha \rightarrow$  find  $C \rightarrow 1 - \beta$
  - From  $C \rightarrow$  determine  $\alpha$
  - From  $\alpha \rightarrow$  and  $F$ -value

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From  $\theta_e \rightarrow$  find  $P$ -value

## Examples for nonnormal data

E.g. 1. A random sample of size  $n$  from uniform distr.  $f_Y(y; \theta) = 1/\theta, y \in [0, \theta]$ .

To test

$$H_0 : \theta = 2.0 \quad \text{v.s.} \quad H_1 : \theta < 2.0$$

at the level  $\alpha = 0.10$  of significance, one can use the decision rule based on  $Y_{\max}$ . Find the probability of committing a Type II error when  $\theta = 1.7$ .

Remark:  $Y_{\max}$  is a sufficient estimator for  $\theta$ . Why?

Sol: 1) The critical region should has the form:  $C = \{Y_{\max} > c\}$  for some  $c$ .

2) We need to use the condition  $\alpha = 0.10$  to find  $c$ .

3) Find the prob. of Type II error.

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Sol: 1) The critical region should has the form:  $C = \{Y_{\max} > c\}$  for some  $c$ .

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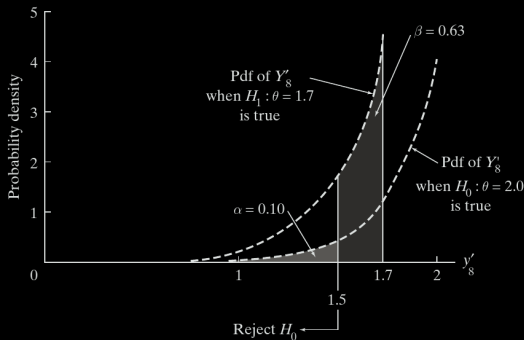
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$$f_{Y_{max}}(y) = \dots = n \frac{y^{n-1}}{\theta^n} \quad y \in [0, \theta].$$

$$\alpha = \int_0^c n \frac{y^{n-1}}{\theta_0^n} dy = \left( \frac{c}{\theta_0} \right)^n \implies c = \theta_0 \alpha^{1/n} \quad (\text{Under } H_0 : \theta = \theta_0)$$

$$\beta = \int_{\theta_0 \alpha^{1/n}}^{\theta_1} n \frac{y^{n-1}}{\theta_1^n} dy = 1 - \left( \frac{\theta_0}{\theta_1} \right)^n \alpha \quad (\text{Under } \theta = \theta_1)$$

Finally, we need only plug in the values  $\theta_0 = 2$ ,  $\theta_1 = 1.7$  and  $\alpha = 0.10$ . □

E.g. 2. A random sample of size 4 from  $\text{Poisson}(\lambda)$ :  $p_X(k; \lambda) = e^{-\lambda} \lambda^k / k!$ ,  $k = 0, 1, \dots$ . One wants to test

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$$\bar{X} \sim \text{Poisson}(3.2)$$

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# Finding critical region

k	P(X=k)	P(X≤ k)	P(X>k)	P(X≥k)
0	0.0408	0.0408	0.9592	1
1	0.1304	0.1712	0.8288	0.9592
2	0.2087	0.3799	0.6201	0.8288
3	0.2226	0.6025	0.3975	0.6201
4	0.1781	0.7806	0.2194	0.3975
5	0.114	0.8946	0.1054	0.2194
6	0.0608	0.9554	0.0446	0.1054
7	0.0278	0.9832	0.0168	0.0446
8	0.0111	0.9943	0.0057	0.0168
9	0.004	0.9982	0.0018	0.0057
10	0.0013	0.9995	0.0005	0.0018
11	0.0004	0.9999	0.0001	0.0005
12	0.0001	1	0	0.0001
13	0	1	0	0
14	0	1	0	0

Poisson lambda= 3.2

```
1 > qpois(1-0.10,3.2)
2 [1] 6
```

```
1 > scipy.stats.poisson.ppf(1-0.10,3.2)
2 [1] 6
```

Computing power of test

k	P(X=k)	P(X≤k)	P(X>k)	P(X≥k)
0	0.0082	0.0082	0.9918	1
1	0.0395	0.0477	0.9523	0.9918
2	0.0948	0.1425	0.8575	0.9523
3	0.1517	0.2942	0.7058	0.8575
4	0.182	0.4763	0.5237	0.7058
5	0.1747	0.651	0.349	0.5237
6	0.1398	0.7908	0.2092	0.349
7	0.0959	0.8867	0.1133	0.2092
8	0.0575	0.9442	0.0558	0.1133
9	0.0307	0.9749	0.0251	0.0558
10	0.0147	0.9896	0.0104	0.0251
11	0.0064	0.996	0.004	0.0104
12	0.0026	0.9986	0.0014	0.004
13	0.0009	0.9995	0.0005	0.0014
14	0.0003	0.9999	0.0001	0.0005
15	0.0001	1	0	0.0001
16	0	1	0	0
17	0	1	0	0
18	0	1	0	0
19	0	1	0	0
20	0	1	0	0

Poisson lambda= 4.8

$$1 - \beta = \mathbb{P}(\text{Reject } H_0 \mid H_1 \text{ is true}) = \mathbb{P}(\bar{X} \geq 6 \mid \bar{X} \sim \text{Poisson}(4.8))$$

□

```
1 > 1-ppois(6-1,4.8)
2 [1] 0.3489936
```

```
1 > 1-sciipy.stats.poisson.cdf(6-1,4.8)
2 [1] 0.3489935627305083
```

```

1 PlotPoissonTable <- function(n=14,lambda=3.2,png_filename,TableTitle) {
2   library (gridExtra)
3   library (grid)
4   library (gtable)
5   x = seq(1,n,1)
6   # qpois(0.90,lambda)
7   tb = cbind(x,
8             round(dpois(x,lambda),4),
9             round(ppois(x,lambda),4),
10            round(1-ppois(x,lambda),4),
11            round(c(1,(1-ppois(x,lambda))[1:n]),4))
12   colnames(tb) <- c("k", "P(X=k)", "P(X<= k)", "P(X>k)", "P(X>=k)")
13   rownames(tb) <- x
14   table <- tableGrob(tb,rows = NULL)
15   title <- textGrob(TableTitle,gp=gpar(fontsize=12))
16   footnote <- textGrob(paste("Poisson lambda=",lambda),
17                       x=0, hjust=0, gp=gpar( fontface=" italic "))
18   padding <- unit(0.2,"line")
19   table <- gtable_add_rows(table, heights = grobHeight(title) + padding,pos = 0)
20   table <- gtable_add_rows(table, heights = grobHeight(footnote)+ padding)
21   table <- gtable_add_grob(table, list ( title , footnote),
22                             t=c(1, nrow(table)), l=c(1,2),r=ncol(table))
23   png(png_filename)
24   grid.draw(table)
25   dev.off ()
26 }
27
28 PlotPoissonTable(14,3.2,"Example_6-4-3_1.png","Finding critical region")
29 PlotPoissonTable(20,4.8,"Example_6-4-3_2.png","Computing power of test")

```

The *R* code to produce the previous two Poisson tables.

E.g. 3. A random sample of size 7 from  $f_Y(y; \theta) = (\theta + 1)y^\theta$ ,  $y \in [0, 1]$ . Test

$$H_0 : \theta = 2.0 \quad \text{v.s.} \quad H_1 : \theta > 2.0$$

Decision rule: Let  $X$  be the number of  $y_i$ 's that exceed 0.9;

Reject  $H_0$  if  $X \geq 4$ .

Find  $\alpha$ .

Sol. 1)  $X \sim \text{binomial}(7, p)$

2) Find  $p$

$$p = P(Y > 0.9 | H_0 \text{ is true})$$

$$= \int_{0.9}^1 3y^2 dy = 0.271$$

3) Compute  $\alpha$

$$\alpha = P(X \geq 4 | n=7) = \sum_{k=4}^7 \binom{7}{k} (0.271)^k (0.729)^{7-k} = 0.092$$

```
1 > 1-pbinom(3,7,0.271)
2 [1] 0.09157663
```

```
1 > 1-sciPy.stats.binom.cdf(3, 7, 0.271)
2 [1] 0.09157663095582469
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