

## Probability Inequalities

**Thm 1 (Gaussian Tail Inequality):** Let  $X \sim \mathcal{N}(0, 1)$ . Then

$$\mathbb{P}(|X| > \epsilon) \leq \frac{2}{\epsilon} e^{-\epsilon^2/2} \quad (1)$$

Additionally:

$$\mathbb{P}(|\bar{X}_n| > \epsilon) \leq \frac{1}{\sqrt{n}\epsilon} e^{-n\epsilon^2/2} \quad (2)$$

**Thm 2 (Markov Inequality):** Let  $X$  be a non-negative random variable s.t.  $\mathbb{E}(X)$  exists. Then  $\forall t > 0$

$$\mathbb{P}(X > t) \leq \frac{\mathbb{E}(X)}{t} \quad (3)$$

**Thm 3 (Chebyshev's Inequality):** Let  $\mu = \mathbb{E}(X)$  and  $\sigma^2 = \text{Var}(X)$ . Then

$$\begin{aligned} \mathbb{P}(|X - \mu| \geq t) &\leq \frac{\sigma^2}{t^2} \\ \mathbb{P}(|(X - \mu)/\sigma| \geq t) &\leq \frac{1}{t^2} \end{aligned} \quad (4) \quad (5)$$

**Lemma 4:** Let  $\mathbb{E}(X) = 0$  and  $a \leq X \leq b$ . Then

$$\mathbb{E}(e^{tX}) \leq e^{t^2(b-a)^2/8} \quad (6)$$

**Lemma 5:** Let  $X$  be any random variable. Then

$$\mathbb{P}(X > \epsilon) \leq \inf_{t \geq 0} e^{-t\epsilon} \mathbb{E}(e^{tX}) \quad (7)$$

**Thm 6 (Hoeffding's Inequality):**  $X_1, \dots, X_n$  iid,  $\mathbb{E}(X_i) = \mu$ ,  $a \leq X_i \leq b$ . Then  $\forall \epsilon > 0$

$$\mathbb{P}(|\bar{X} - \mu| \geq \epsilon) \leq 2e^{-2n\epsilon^2/(b-a)^2} \quad (8)$$

**Thm 9 (McDiarmid):**

**Thm 12 (Cauchy-Schwartz inequality):**

**Thm 13 (Jensen's inequality):**

**Ex 15 (Kullback Leibler distance):**

**Thm 18?:**

$O_p$  and  $o_p$ :

## Shattering

Note: remember uniform bounds and union bound.

$\mathcal{A}$  picks out  $G \subset F$ .

$S(\mathcal{A}, F)$ .

$F$  shattered by  $\mathcal{A}$  if  $S(\mathcal{A}, F) = 2^{|F|}$  (ie if  $\mathcal{A}$  picks out all  $G \subset F$ ).

The shatter coefficient  $s_n(\mathcal{A}) = \sup_{F \in \mathcal{F}_n} s(\mathcal{A}, F)$ . Note  $n = |F|$  and  $s_n(\mathcal{A}) \leq 2^n$ .

**Thm 5:**

The VC dimension  $d(\mathcal{A}) = \text{largest } n \text{ s.t. } s_n(\mathcal{A}) = 2^n$ .

## Random Samples

For  $X_1, \dots, X_n \sim F$  a statistic is any  $T = g(X_1, \dots, X_n)$ .

E.g.  $\bar{X}_n, S_n = \sum_i (X_i - \bar{X}_n)^2 / (n-1), (X_{(1)}, \dots, X_{(n)})$

**Note:**  $\mathbb{E}(\bar{X}_n) = \mathbb{E}(X_i)$ ,  $\text{Var}(\bar{X}_n) = \text{Var}(X_i)/n$ ,  $\mathbb{E}(S_n)^2 = \text{Var}(X_i)$ .

**Note:** sum of bernoulli is binomial(n,p), sum of exp(beta) is gamma(n,beta), sum of standard normal is chi-squared(n-dof).

**Thm. 1:**  $X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2) \implies \bar{X}_n \sim \mathcal{N}(\mu, \sigma^2/n)$ .

## Convergence

$X, X_1, X_2, \dots$  random variables.

(1)  $X_n$  converges **almost surely**  $X_n \xrightarrow{a.s.} X$  if  $\forall \epsilon > 0$

$$\mathbb{P}(\lim_{n \rightarrow \infty} |X_n - X| < \epsilon) = 1 \quad (9)$$

(2)  $X_n$  converges **in probability**  $X_n \xrightarrow{p} X$  if  $\forall \epsilon > 0$

$$\lim_{n \rightarrow \infty} \mathbb{P}(|X_n - X| > \epsilon) = 0 \quad (10)$$

(3)  $X_n$  converges **in quadratic mean**  $X_n \xrightarrow{qm} X$  if

$$\lim_{n \rightarrow \infty} \mathbb{E}[(X_n - X)^2] = 0 \quad (11)$$

(4)  $X_n$  converges **in distribution**  $X_n \rightsquigarrow X$  if

$$\lim_{n \rightarrow \infty} F_{X_n}(t) = F_X(t) \quad (12)$$

$\forall t$  on which  $F_X$  is continuous.

**Thm 7:** Conv. a.s. and in q.m. imply conv. in prob. All three imply conv. in distribution. Conv. in distribution to a point-mass also implies conv. in prob.

Ex from class: Showed conv. in prob  $\not\Rightarrow$  conv. a.s.. Showed conv. in prob  $\not\Rightarrow$  conv. in q.m.. Showed conv. in distro  $\not\Rightarrow$  conv. in prob.

**Thm 10a:**  $X, X_n, Y, Y_n$  random variables. Then

$$(a) \quad X_n \xrightarrow{p} X, Y_n \xrightarrow{p} Y \implies X_n + Y_n \xrightarrow{p} X + Y \quad (13)$$

$$(b) \quad X_n \xrightarrow{p} X, Y_n \xrightarrow{p} Y \implies X_n Y_n \xrightarrow{p} XY \quad (14)$$

$$(c) \quad X_n \xrightarrow{qm} X, Y_n \xrightarrow{qm} Y \implies X_n + Y_n \xrightarrow{qm} X + Y \quad (15)$$

**Thm 10b (Slutzky's Thm):**  $X, X_n, Y_n$  random variables. Then

$$(a) \quad X_n \rightsquigarrow X, Y_n \rightsquigarrow c \implies X_n + Y_n \rightsquigarrow X + c \quad (16)$$

$$(b) \quad X_n \rightsquigarrow X, Y_n \rightsquigarrow c \implies X_n Y_n \rightsquigarrow cX \quad (17)$$

**Thm 12 (Law of Large Numbers):**  $X_1, \dots, X_n$  iid,  $\mathbb{E}(X_i) = \mu \implies \bar{X}_n \xrightarrow{qm} \mu$ .

**Thm 14 (CLT):**  $X_1, \dots, X_n$  iid,  $\mathbb{E}(X_i) = \mu$ ,  $\text{Var}(X_i) = \sigma^2$

$\implies \sqrt{n}(\bar{X}_n - \mu)/\sigma \rightsquigarrow \mathcal{N}(0, 1)$

$\implies \bar{X}_n \rightsquigarrow \mathcal{N}(\mu, \sigma^2/n)$

$\implies \sqrt{n}(\bar{X}_n - \mu)/S_n \rightsquigarrow \mathcal{N}(0, 1)$

**Thm 18 (delta method):** If  $\sqrt{n}(Y_n - \mu)/\sigma \rightsquigarrow \mathcal{N}(0, 1)$ ,  $g'(\mu) \neq 0$

$\implies \sqrt{n}(g(Y_n) - g(\mu))/|g'(\mu)|\sigma \rightsquigarrow \mathcal{N}(0, 1)$

ie  $Y_n \approx \mathcal{N}(\mu, \sigma^2/n) \implies g(Y_n) \approx \mathcal{N}(g(\mu), g'(\mu)^2 \sigma^2/n)$

**Thm 18b (2nd order delta method):??** Should I include this?

## Sufficiency

If  $X_1, \dots, X_n \sim p(x; \theta)$ ,  $T$  **sufficient** for  $\theta$  if  $p(x^n | t; \theta) = p(x^n | t)$ .

**Thm 9 (factorization):** for  $X^n \sim p(x; \theta)$ ,  $T(X^n)$  sufficient for  $\theta$  if the joint probability can be factorized as.

$$p(x^n; \theta) = h(x^n) \times g(t; \theta) \quad (18)$$

$T$  is a **minimal sufficient statistic (MSS)** if  $T$  is sufficient and  $T = g(U)$  for all other sufficient stats  $U$ .

**Thm 15:**  $T$  is a MSS if:

$$\frac{p(y^n; \theta)}{p(x^n; \theta)} \text{ constant in } \theta \iff T(y^n) = T(x^n) \quad (19)$$

## Parametric Point Estimation

make sure i've defined:  $\mathbb{E}_\theta(\hat{\theta})$ , bias, sampling distro, standard error,  $\hat{\theta}_n$  consistent.

**Method of Moments:** Define equations

- (a)  $(\sum_i X_i)/n = \mathbb{E}_{\hat{\theta}}(X_i)$
- (b)  $(\sum_i X_i^2)/n = \mathbb{E}_{\hat{\theta}}(X_i^2)$
- (c) ...

And solve for  $\hat{\theta}$ .

**Maximum Likelihood (MLE):** The MLE is

$$\hat{\theta} = \arg \min_{\theta} L(\theta) = \arg \min_{\theta} l(\theta) \quad (20)$$

Often suffices to solve for  $\theta$  in  $\frac{\partial l(\theta)}{\partial \theta} = 0$ . The MLE is **equivariant**  $\implies$  if  $\eta = g(\theta)$  then  $\hat{\eta} = g(\hat{\theta})$ .

**Bayes Estimation:** For prior  $\pi(\theta)$ , choose

$$\hat{\theta} = \mathbb{E}(\theta|x^n) = \int \theta \pi(\theta|x^n) d\pi \quad (21)$$

**Mean Squared Error (MSE):** The MSE is

$$\text{MSE} = \mathbb{E}(\hat{\theta} - \theta)^2 = \int (\hat{\theta} - \theta)^2 p(x^n; \theta) dx^n = \text{bias}(\hat{\theta})^2 + \text{Var}(\hat{\theta}) \quad (22)$$

Notes: **bias**( $\hat{\theta}$ ) =  $\mathbb{E}(\hat{\theta}) - \theta$ . We say  $\hat{\theta}$  is **consistent** if  $\hat{\theta} = \hat{\theta}_n \xrightarrow{P} \theta$ . The **standard error** of  $\hat{\theta}$ ,  $\text{se}(\hat{\theta})$ , is the standard deviation of  $\hat{\theta}$ . Ex (in class): MSE for normal.

## Risks and Estimators

$L(\theta, \hat{\theta})$  is the **loss** of an estimator  $\hat{\theta} = \hat{\theta}(x^n)$  for  $x^n \sim p(x^n; \theta)$ .

The **risk** of  $\hat{\theta}$  is

$$R(\theta, \hat{\theta}) = \mathbb{E}[L(\theta, \hat{\theta})] = \int L(\theta, \hat{\theta}) p(x^n; \theta) dx^n \quad (23)$$

When  $L(\theta, \hat{\theta}) = (\theta - \hat{\theta})^2$ , the risk is the MSE.

The **max risk** of  $\hat{\theta}$  over a set  $\theta \in \Theta$  is

$$\overline{R}(\hat{\theta}) = \sup_{\theta \in \Theta} R(\theta, \hat{\theta}) \quad (24)$$

The **minimax estimator** is

$$\hat{\theta} = \arg \inf_{\hat{\theta}} \overline{R}(\hat{\theta}) \quad (25)$$

The **Bayes risk** of  $\hat{\theta}$  given a prior  $\pi(\theta)$  is

$$B_\pi(\hat{\theta}) = \int R(\theta, \hat{\theta}) \pi(\theta) d\theta \quad (26)$$

The **posterior risk** of  $\hat{\theta}$  given a prior  $\pi(\theta)$  is

$$r(\hat{\theta}|x^n) = \int L(\theta, \hat{\theta}) \pi(\theta|x_1, \dots, x_n) d\theta \quad (27)$$

where  $\pi(\theta|x^n) = \frac{\mathbb{P}(x^n; \theta) \pi(\theta)}{m(x^n)}$  is the posterior over  $\theta$ .

The **Bayes estimator** is

$$\hat{\theta} = \arg \inf_{\hat{\theta}} B_\pi(\hat{\theta}) = \arg \inf_{\hat{\theta}} r(\hat{\theta}|x^n) \quad (28)$$

which equals  $\mathbb{E}(\theta|x^n)$  when  $L(\theta, \hat{\theta}) = (\theta - \hat{\theta})^2$ .

**Thm 10:** If  $\hat{\theta}$  is a Bayes estimator for some prior  $\pi$  and  $R(\theta, \hat{\theta})$  is constant, then  $\hat{\theta}$  is a minimax estimator.

**Note:** The MLE is approximately minimax (as n increases, if dimension of the parameter is fixed).