

## 习题 4

$$1. E(X) = (-1) \times 0.2 + 0 \times 0.3 + 1 \times 0.5 = 0.3.$$

$$E(X^2) = (-1)^2 \times 0.2 + 0^2 \times 0.3 + 1^2 \times 0.5 = 0.7$$

$$E(-2X^2 + 3) = -2E(X^2) + 3 = -2 \times 0.7 + 3 = 1.6$$

### 2. 参习题 3 第 6 题:

数 1 出现的概率是 1/10, 它可被自己整除  $X=1$

数 2、3、5、7 出现的概率是 4/10, 它可被  $X=2$  个整数整除

数 4、9 出现的概率是 2/10, 它们可以被  $X=3$  个整数整除

数 6、8、10 出现的概率是 2/10, 它们可以被  $X=4$  个整数整除

于是给出分布率:

X	1	2	3	4
P	1/10	4/10	2/10	3/10

$$E(X) = 1 \times \frac{1}{10} + 2 \times \frac{4}{10} + 3 \times \frac{2}{10} + 4 \times \frac{3}{10} = \frac{27}{10} = 2.7.$$

$$3. E(X) = \int_{-\infty}^{+\infty} xf(x)dx = \int_{-\infty}^{+\infty} x \cdot \frac{1}{2} e^{-|x|} dx = \int_{-\infty}^0 \frac{x}{2} e^x dx + \int_0^{+\infty} \frac{x}{2} e^{-x} dx = 0$$

$$\begin{aligned} E(X^2) &= \int_{-\infty}^{+\infty} x^2 f(x) dx = \int_{-\infty}^{+\infty} x^2 \cdot \frac{1}{2} e^{-|x|} dx = \int_0^{+\infty} x^2 e^{-x} dx = - \int_0^{+\infty} x^2 de^{-x} \\ &= - \left[ 0 + 2 \int_0^{+\infty} x de^{-x} \right] = -2 \left[ xe^{-x} \Big|_0^{+\infty} - \int_0^{+\infty} e^{-x} dx \right] = -2 \left[ 0 + e^{-x} \Big|_0^{+\infty} \right] = 2. \end{aligned}$$

4.

$$10. E(X) = \int_{-\infty}^{+\infty} xf(x)dx = \int_0^1 x \cdot x dx + \int_1^2 x \cdot (2-x) dx = \frac{x^3}{3} \Big|_0^1 + \left[ x^2 - \frac{x^3}{3} \right]_1^2 = 1$$

$$\begin{aligned} D(X) &= \left( \int_{-\infty}^{+\infty} [x - E(X)]^2 f(x) dx \right) = E(X^2) - [E(X)]^2 = \int_{-\infty}^{+\infty} x^2 f(x) dx - 1^2 \\ &= \int_0^1 x^2 \cdot x dx + \int_1^2 x^2 \cdot (2-x) dx - 1 = \frac{x^4}{4} \Big|_0^1 + \left[ \frac{2}{3} x^3 - \frac{x^4}{4} \right]_1^2 - 1 = \frac{1}{6}. \end{aligned}$$

5.

$$E(Y) = E(X_1) - 2E(X_2) + 3E(X_3)$$

$$= \frac{0+6}{2} - 2 \cdot 1 + 3 \cdot \frac{1}{3}$$

$$= 2.$$

$$D(Y) = D(X_1) + (-2)^2 D(X_2) + 3^2 D(X_3)$$

$$= \frac{(6-0)^2}{12} + 4 \cdot 3 + 9 \cdot \frac{1}{3^2}$$

$$= 16.$$

$$E(Y^2) = D(Y) + (E(Y))^2 = 16 + 4 = 20.$$

12. 设  $(X, Y)$  的分布律为

$Y \backslash X$	-1	1	2
-1	$\frac{5}{20}$	$\frac{2}{20}$	$\frac{6}{20}$
1	$\frac{3}{20}$	$\frac{3}{20}$	$\frac{1}{20}$

求: (1)  $X+Y$ ; (2)  $X-Y$  的期望和方差.

6. 联合分布与边缘分布为

$Y \backslash X$	-1	1	2	$P\{X=i\}$
-1	$\frac{5}{20}$	$\frac{2}{20}$	$\frac{6}{20}$	$\frac{13}{20}$
1	$\frac{3}{20}$	$\frac{3}{20}$	$\frac{1}{20}$	$\frac{7}{20}$
$P\{Y=j\}$	$\frac{8}{20}$	$\frac{5}{20}$	$\frac{7}{20}$	

$$E(X) = (-1) \times \frac{13}{20} + 1 \times \frac{7}{20} = -\frac{6}{20},$$

$$E(X^2) = (-1)^2 \times \frac{13}{20} + 1^2 \times \frac{7}{20} = 1,$$

$$E(Y) = (-1) \times \frac{8}{20} + 1 \times \frac{5}{20} + 2 \times \frac{7}{20} = \frac{11}{20},$$

$$E(Y^2) = (-1)^2 \times \frac{8}{20} + 1^2 \times \frac{5}{20} + 2^2 \times \frac{7}{20} = \frac{41}{20}$$

$$(1) E(X+Y) = E(X) + E(Y) = -\frac{6}{20} + \frac{11}{20} = \frac{5}{20} = \frac{1}{4}.$$

因  $X$  与  $Y$  不独立, 因此不能用  $D(X+Y) = D(X) + D(Y)$  计算

$$\begin{aligned} D(X+Y) &= E\{(X+Y) - [E(X+Y)]^2\} \\ &= (-1-1-\frac{1}{4})^2 \times \frac{5}{20} + (-1+1-\frac{1}{4})^2 \times \frac{2}{20} + (-1+2-\frac{1}{4})^2 \times \frac{6}{20} \\ &\quad + (1-1-\frac{1}{4})^2 \times \frac{3}{20} + (1+1-\frac{1}{4})^2 \times \frac{3}{20} + (1+2-\frac{1}{4})^2 \times \frac{1}{20} \\ &= \frac{1}{4^2 \times 20} (81 \times 5 + 2 + 9 \times 6 + 3 + 49 \times 3 + 121) = \frac{183}{80} \end{aligned}$$

(也可先求  $E((X+Y)^2)$ , 再求  $D(X+Y)$ .)

$$(2) E(X-Y) = E(X) - E(Y) = -\frac{11}{20} - \frac{6}{20} = -\frac{17}{20}.$$

$$\begin{aligned}
 D(X-Y) &= E\{[(X-Y) - E(X-Y)]^2\} \\
 &= (-1+1+\frac{17}{20})^2 \times \frac{5}{20} + (-1-1+\frac{17}{20})^2 \times \frac{2}{20} + (-1-2+\frac{17}{20})^2 \times \frac{6}{20} \\
 &\quad + (1+1+\frac{17}{20})^2 \times \frac{3}{20} + (1-1+\frac{17}{20})^2 \times \frac{3}{20} + (1-2+\frac{17}{20})^2 \times \frac{1}{20} \\
 &= \frac{24220}{20^2 \times 20} = \frac{1211}{400}
 \end{aligned}$$

法 2:

P	5/20	2/20	6/20	3/20	3/20	1/20
(X,Y)	(-1,-1)	(-1,1)	(-1,2)	(1,-1)	(1,1)	(1,2)
X+Y	-2	0	1	0	2	3
X-Y	0	-2	-3	2	0	-1

(1)

$$E(X+Y)^2 = 4 \cdot \frac{5}{20} + 1 \cdot \frac{6}{20} + 4 \cdot \frac{3}{20} + 9 \cdot \frac{1}{20} = \frac{47}{20}$$

$$D(X+Y) = E(X+Y)^2 - (E(X+Y))^2 = \frac{47}{20} - (\frac{1}{4})^2 = \frac{183}{80}$$

(2)

$$E(X-Y)^2 = 4 \cdot \frac{2}{20} + 9 \cdot \frac{6}{20} + 4 \cdot \frac{3}{20} + 1 \cdot \frac{1}{20} = \frac{75}{20}$$

$$D(X-Y) = E(X-Y)^2 - (E(X-Y))^2 = \frac{75}{20} - (-\frac{17}{20})^2 = \frac{1211}{400}$$

$$\begin{aligned}
 7. \quad E(Z) &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} z f(x, y) dy dx = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} z(x, y) f_X(x) f_Y(y) dy dx \\
 &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} \sqrt{x^2 + y^2} \cdot \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dy dx = \int_0^{2\pi} \int_0^{+\infty} r \cdot \frac{1}{2\pi} e^{-\frac{r^2}{2}} r dr d\theta \\
 &= \frac{1}{2\pi} \cdot 2\pi \int_0^{+\infty} (-r) de^{-\frac{r^2}{2}} = -re^{-\frac{r^2}{2}} \Big|_0^{+\infty} + \int_0^{+\infty} e^{-\frac{r^2}{2}} dr = \sqrt{2} \int_0^{+\infty} e^{-\frac{r^2}{2}} d\frac{r}{\sqrt{2}} = \sqrt{\frac{\pi}{2}} \triangleq a.
 \end{aligned}$$

$$\begin{aligned}
 D(Z) &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} [z - E(Z)]^2 f(x, y) dy dx = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} (\sqrt{x^2 + y^2} - a)^2 \cdot \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}} dy dx \\
 &= \frac{1}{2\pi} \int_0^{2\pi} \int_0^{+\infty} (r - a)^2 e^{-\frac{r^2}{2}} r dr d\theta = \int_0^{+\infty} (r - a)^2 e^{-\frac{r^2}{2}} r dr = -\int_0^{+\infty} (r - a)^2 de^{-\frac{r^2}{2}} \\
 &= -(r - a)^2 e^{-\frac{r^2}{2}} \Big|_0^{+\infty} + \int_0^{+\infty} e^{-\frac{r^2}{2}} \cdot 2(r - a) dr = a^2 + 2 \int_0^{+\infty} e^{-\frac{r^2}{2}} r dr - 2a \int_0^{+\infty} e^{-\frac{r^2}{2}} dr \\
 &= a^2 - 2e^{-\frac{r^2}{2}} \Big|_0^{+\infty} - 2a\sqrt{2} \int_0^{+\infty} e^{-\frac{r^2}{2}} d\frac{r}{\sqrt{2}} = a^2 + 2 - 2a\sqrt{2} \cdot \frac{\sqrt{\pi}}{2} = \frac{\pi}{2} + 2 - \pi = 2 - \frac{\pi}{2}.
 \end{aligned}$$

8. 设  $(X, Y)$  的密度函数为

$$f(x, y) = \begin{cases} 12y^2 & 0 \leq y \leq x \leq 1 \\ 0 & \text{其它} \end{cases}$$

求  $E(X), E(Y), E(XY), E(X^2 + Y^2)$ .

$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_0^x 12y^2 dy, & 0 < x \leq 1 \\ 0, & \text{其它} \end{cases} = \begin{cases} 4x^3, & 0 < x \leq 1 \\ 0, & \text{其它} \end{cases}$$

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_y^1 12y^2 dx, & 0 < y < 1 \\ 0, & \text{其它} \end{cases} = \begin{cases} 12y^2(1 - y), & 0 < y < 1 \\ 0, & \text{其它} \end{cases}$$

$$E(X) = \int_{-\infty}^{+\infty} x f_X(x) dx = \int_0^1 x \cdot 4x^3 dx = \frac{4}{5}.$$

$$E(Y) = \int_{-\infty}^{+\infty} y f_Y(y) dy = \int_0^1 y \cdot 12y^2(1 - y) dy = 3y^4 \Big|_0^1 - \frac{12}{5} y^5 \Big|_0^1 = \frac{3}{5}.$$

$$E(XY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f(x, y) dy dx = \int_0^1 \int_0^x xy \cdot 12y^2 dy dx = 3 \int_0^1 x^5 dx = \frac{3}{6} = \frac{1}{2}.$$

$$\begin{aligned}
 E(X^2 + Y^2) &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} (x^2 + y^2) f(x, y) dy dx = \int_0^1 \int_0^x (x^2 + y^2) \cdot 12y^2 dy dx \\
 &= 3 \int_0^1 (4x^5 + \frac{12}{5} x^5) dx = \frac{16}{15}.
 \end{aligned}$$

$$9. \quad 1 = F(+\infty, +\infty) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x, y) dy dx = \int_0^1 \int_0^x k dy dx = k \int_0^1 x dx = \frac{k}{2}$$

$$(\text{或由分布均匀知}) k = \frac{1}{f \text{非零区域面积}} = \frac{1}{\frac{1}{2} \cdot 1 \cdot 1} = 2.)$$

$$\therefore f(x, y) = \begin{cases} 2, & 0 < y < x, 0 < x < 1 \\ 0, & \text{其它} \end{cases}$$

$$E(XY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f(x, y) dy dx = \int_0^1 \int_0^x xy \cdot 2 dy dx = \frac{1}{4}.$$

$$\begin{aligned} 10. \quad E(X) &= \int_{-\infty}^{+\infty} x f(x) dx = \int_0^{+\infty} x \frac{x}{\sigma^2} e^{-\frac{x^2}{2\sigma^2}} dx = -\int_0^{+\infty} x de^{-\frac{x^2}{2\sigma^2}} \\ &= -xe^{-\frac{x^2}{2\sigma^2}} \Big|_0^{+\infty} + \int_0^{+\infty} e^{-\frac{x^2}{2\sigma^2}} dx = 0 + \frac{\sqrt{2\pi}\sigma}{2} \cdot 2 \int_0^{+\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{x^2}{2\sigma^2}} dx \\ &= \sqrt{\frac{\pi}{2}} \sigma \cdot 1 \end{aligned}$$

$$\begin{aligned} E(X^2) &= \int_0^{+\infty} x^2 \frac{x}{\sigma^2} e^{-\frac{x^2}{2\sigma^2}} dx = -\int_0^{+\infty} x^2 de^{-\frac{x^2}{2\sigma^2}} \\ &= -x^2 e^{-\frac{x^2}{2\sigma^2}} \Big|_0^{+\infty} + 2 \int_0^{+\infty} x e^{-\frac{x^2}{2\sigma^2}} dx = -2\sigma^2 e^{-\frac{x^2}{2\sigma^2}} \Big|_0^{+\infty} \\ &= 2\sigma^2 \end{aligned}$$

$$DX = E(X^2) - (EX)^2 = 2\sigma^2 - \frac{\pi}{2}\sigma^2 = \frac{4-\pi}{4}\sigma^2.$$

$$\begin{aligned} 11. \quad E(X) &= \sum_{k=1}^{\infty} k P\{X=k\} = \sum_{k=1}^{\infty} kp(1-p)^{k-1} = -p \sum_{k=1}^{\infty} [(1-p)^k]' = -p \left( \sum_{k=1}^{\infty} (1-p)^k \right)' \\ &= -p \left( \sum_{k=0}^{\infty} (1-p)^k - 1 \right)' = -p \left( \frac{1}{1-(1-p)} - 1 \right)' = \frac{1}{p}. \\ D(X) &= \sum_{k=1}^{\infty} \left( k - \frac{1}{p} \right)^2 \cdot p(1-p)^{k-1} = \sum_{k=1}^{\infty} k^2 p(1-p)^{k-1} - 2 \sum_{k=1}^{\infty} k(1-p)^{k-1} + \frac{1}{p} \sum_{k=1}^{\infty} (1-p)^{k-1} \\ &= \sum_{k=1}^{\infty} k^2 p(1-p)^{k-1} - \frac{2}{p} E(x) + \frac{1}{p} \cdot \frac{1}{1-(1-p)} = \sum_{k=1}^{\infty} k^2 p(1-p)^{k-1} - \frac{1}{p^2} \\ \text{而 } \sum_{k=1}^{\infty} k^2 p(1-p)^{k-1} &= p \sum_{k=1}^{\infty} (k+1)k(1-p)^{k-1} - \sum_{k=1}^{\infty} kp(1-p)^{k-1} = p \sum_{k=1}^{\infty} [(1-p)^{k+1}]'' - E(X) \end{aligned}$$

$$\begin{aligned}
 &= p \left[ \sum_{k=1}^{\infty} (1-p)^{k+1} \right]^n - \frac{1}{p} = p \left[ \frac{1}{1-(1-p)} - (1-p) - 1 \right]^n - \frac{1}{p} = \frac{2}{p^2} - \frac{1}{p} \\
 \therefore D(X) &= \frac{2}{p^2} - \frac{1}{p} - \frac{1}{p^2} = \frac{1}{p^2} - \frac{1}{p}.
 \end{aligned}$$

12. (1) 若  $X$  为离散型随机变量, 则

$$\begin{aligned}
 \varphi(k) &= E(X-k)^2 = \sum_{i=1}^{\infty} (x_i - k)^2 p_i = \sum_{i=1}^{\infty} x_i^2 p_i - 2k \sum_{i=1}^{\infty} x_i p_i + k^2 \sum_{i=1}^{\infty} p_i \\
 &= \sum_{i=1}^{\infty} x_i^2 p_i - 2kE(X) + k^2 \cdot 1
 \end{aligned}$$

令  $\varphi'(k) = 0 - 2E(X) + 2k = 0$  得  $k = E(X)$ , 又  $\varphi''(k) = 2 > 0$ ,

故  $\varphi(k)$  在  $k = E(X)$  取得最小值.

(2) 若  $X$  为连续型随机变量, 则

$$\begin{aligned}
 \varphi(k) &= E(X-k)^2 = \int_{-\infty}^{+\infty} (x-k)^2 f(x) dx = \int_{-\infty}^{+\infty} x^2 f(x) dx - 2k \int_{-\infty}^{+\infty} x f(x) dx + k^2 \int_{-\infty}^{+\infty} f(x) dx \\
 &= E(X^2) - 2kE(X) + k^2 \cdot 1
 \end{aligned}$$

令  $\varphi'(k) = 0 - 2E(X) + 2k = 0$  得  $k = E(X)$ , 又  $\varphi''(k) = 2 > 0$ ,

故  $\varphi(k)$  在  $k = E(X)$  取得最小值.

13. (1) 对概率密度为  $f(x) = \frac{1}{\pi} \frac{1}{1+x^2}$  的柯西分布, 有

$$\begin{aligned}
 E(X) &= \int_{-\infty}^{+\infty} x f(x) dx = \int_{-\infty}^{+\infty} x \cdot \frac{1}{\pi} \frac{1}{1+x^2} dx = \frac{1}{\pi} \lim_{\substack{a \rightarrow -\infty \\ b \rightarrow +\infty}} [\ln(1+x^2)]_a^b \\
 &= \frac{1}{\pi} \lim_{\substack{a \rightarrow -\infty \\ b \rightarrow +\infty}} [\ln(1+b^2) - \ln(1+a^2)] = \frac{1}{\pi} \lim_{\substack{a \rightarrow -\infty \\ b \rightarrow +\infty}} \ln \frac{1+b^2}{1+a^2} \text{ 不存在}
 \end{aligned}$$

(2) 对概率分布为  $P\{X = (-1)^{j+1} \frac{3^j}{j}\} = \frac{j}{3^j}$  的离散型随机变量  $X$ , 有

$$E(X) = \sum_{j=1}^{\infty} x_j p_j = \sum_{j=1}^{\infty} \left( (-1)^{j+1} \frac{3^j}{j} \cdot \frac{j}{3^j} \right) = \sum_{j=1}^{\infty} (-1)^{j+1} \text{ 不存在.}$$

14. 因  $E(X), D(X)$  为常数, 所以由均值的性质得

$$\begin{aligned}
 E(X^*) &= E\left(\frac{X - E(X)}{\sqrt{D(X)}}\right) = \frac{1}{\sqrt{D(X)}} E(X - E(X)) \\
 &= \frac{1}{\sqrt{D(X)}} (E(X) - E(X)) = 0 \\
 D(X^*) &= \left(\frac{1}{\sqrt{D(X)}}\right)^2 D(X + (-E(X))) = \frac{1}{D(X)} D(X) = 1 \text{ (性质3的推论)}
 \end{aligned}$$

$$15. (1) \quad E(\bar{X}) = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} E\left(\sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n E(X_i) = \frac{1}{n} \sum_{i=1}^n \mu = \mu.$$

由  $X_1, X_2, \dots, X_n$  相互独立, 得

$$D(\bar{X}) = \left(\frac{1}{n}\right)^2 D\left(\sum_{i=1}^n X_i\right) = \frac{1}{n^2} \sum_{i=1}^n D(X_i) = \frac{1}{n^2} \sum_{i=1}^n \sigma^2 = \frac{\sigma^2}{n}.$$

$$\begin{aligned} (2) \quad E(S^2) &= E\left(\frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2\right) = \frac{1}{n-1} E\left(\sum_{i=1}^n (X_i^2 - 2X_i\bar{X} + \bar{X}^2)\right) \\ &= \frac{1}{n-1} E\left(\sum_{i=1}^n X_i^2 - 2n\bar{X} \cdot \bar{X} + n\bar{X}^2\right) = \frac{1}{n-1} \left[ \sum_{i=1}^n E(X_i^2) - nE(\bar{X}^2) \right] \\ &= \frac{1}{n-1} \left\{ \sum_{i=1}^n [DX_i + (EX_i)^2] - n[D(\bar{X}) + (E\bar{X})^2] \right\} \\ &= \frac{1}{n-1} [n\sigma^2 + n\mu^2 - n\left(\frac{\sigma^2}{n} + \mu^2\right)] = \frac{1}{n-1} (n-1)\sigma^2 = \sigma^2. \end{aligned}$$

$$\begin{aligned} 16. \text{法1} \quad E(X) &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xf(x, y) dy dx = \int_0^{\frac{\pi}{2}} \int_0^{\frac{\pi}{2}} x \cdot \frac{1}{2} \sin(x+y) dy dx \\ &= -\frac{1}{2} \int_0^{\frac{\pi}{2}} \left[ \int_0^{\frac{\pi}{2}} x d \cos(x+y) \right] dx = -\frac{1}{2} \int_0^{\frac{\pi}{2}} \left[ x \cos(x+y) - \sin(x+y) \right]_0^{\frac{\pi}{2}} dx \\ &= -\frac{1}{2} \int_0^{\frac{\pi}{2}} \left[ \frac{\pi}{2} \cos\left(x + \frac{\pi}{2}\right) - \sin\left(x + \frac{\pi}{2}\right) + \sin x \right] dx \\ &= -\frac{1}{2} \left[ \frac{\pi}{2} \sin\left(x + \frac{\pi}{2}\right) + \cos\left(x + \frac{\pi}{2}\right) - \cos x \right]_0^{\frac{\pi}{2}} \\ &= -\frac{1}{2} \left[ -1 - \left(\frac{\pi}{2} - 1\right) \right] = \frac{\pi}{4}. \end{aligned}$$

$$\begin{aligned} E(X^2) &= \int_0^{\frac{\pi}{2}} \int_0^{\frac{\pi}{2}} x^2 \cdot \frac{1}{2} \sin(x+y) dy dx \\ &= -\frac{1}{2} \int_0^{\frac{\pi}{2}} \left[ \int_0^{\frac{\pi}{2}} x^2 d \cos(x+y) \right] dx = -\frac{1}{2} \int_0^{\frac{\pi}{2}} \left[ \left[ x^2 \cos(x+y) \right]_0^{\frac{\pi}{2}} - 2 \int_0^{\frac{\pi}{2}} x \cos(x+y) dy \right] dx \\ &= -\frac{1}{2} \int_0^{\frac{\pi}{2}} \left[ \frac{\pi^2}{4} \cos\left(x + \frac{\pi}{2}\right) - 2 \left[ x \sin(x+y) + \cos(x+y) \right]_0^{\frac{\pi}{2}} \right] dx \\ &= -\frac{1}{2} \int_0^{\frac{\pi}{2}} \left[ \frac{\pi^2}{4} \cos\left(x + \frac{\pi}{2}\right) - 2 \left[ \frac{\pi}{2} \sin\left(x + \frac{\pi}{2}\right) + \cos\left(x + \frac{\pi}{2}\right) - \cos x \right] \right] dx \\ &= -\frac{1}{2} \left[ \frac{\pi^2}{4} \sin\left(x + \frac{\pi}{2}\right) + \pi \cos\left(x + \frac{\pi}{2}\right) - 2 \sin\left(x + \frac{\pi}{2}\right) + 2 \sin x \right]_0^{\frac{\pi}{2}} \\ &= -\frac{1}{2} \left[ -\pi + 2 - \left(\frac{\pi^2}{4} - 2\right) \right] = \frac{\pi}{2} + \frac{\pi^2}{8} - 2 \end{aligned}$$

$$16. \quad f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_0^{\frac{\pi}{2}} \frac{1}{2} \sin(x+y) dy, & 0 \leq x \leq \frac{\pi}{2} \\ 0, & \text{其它} \end{cases} = \begin{cases} \frac{1}{2}(\sin x + \cos x), & 0 \leq x \leq \frac{\pi}{2} \\ 0, & \text{其它} \end{cases}$$

$$f_Y(y) = \begin{cases} \frac{1}{2}(\sin y + \cos y), & 0 \leq y \leq \frac{\pi}{2} \\ 0, & \text{其它} \end{cases}$$

$$\begin{aligned} E(Y) = E(X) &= \int_{-\infty}^{+\infty} x f_X(x) dx = \int_0^{\frac{\pi}{2}} x \cdot \frac{1}{2}(\sin x + \cos x) dx = \frac{1}{2} \int_0^{\frac{\pi}{2}} x d(-\cos x + \sin x) \\ &= \frac{1}{2} [x(-\cos x + \sin x)]_0^{\frac{\pi}{2}} - \int_0^{\frac{\pi}{2}} (-\cos x + \sin x) dx = \frac{1}{2} \left[ \frac{\pi}{2} + [\sin x + \cos x]_0^{\frac{\pi}{2}} \right] = \frac{\pi}{4}. \end{aligned}$$

$$\begin{aligned} E(X^2) &= \int_{-\infty}^{+\infty} x^2 \cdot \frac{1}{2}(\sin x + \cos x) dx = \frac{1}{2} \int_0^{\frac{\pi}{2}} x^2 d(-\cos x + \sin x) \\ &= \frac{1}{2} [x^2(-\cos x + \sin x)]_0^{\frac{\pi}{2}} - \int_0^{\frac{\pi}{2}} 2x(-\cos x + \sin x) dx = \frac{1}{2} \left[ \frac{\pi^2}{4} + 2 \int_0^{\frac{\pi}{2}} x d(\sin x + \cos x) \right] \\ &= \frac{\pi^2}{8} + x(\sin x + \cos x) \Big|_0^{\frac{\pi}{2}} - \int_0^{\frac{\pi}{2}} (\sin x + \cos x) dx = \frac{\pi^2}{8} + \frac{\pi}{2} + [\cos x - \sin x]_0^{\frac{\pi}{2}} \\ &= \frac{\pi^2}{8} + \frac{\pi}{2} - 2 \end{aligned}$$

$$D(Y) = D(X) = E(X^2) - (E(X))^2 = \frac{\pi^2}{8} + \frac{\pi}{2} - 2 - \left(\frac{\pi}{4}\right)^2 = \frac{\pi^2 + 8\pi - 32}{16}.$$

17. 由习题册第10题(书第4题)知  $E(X) = 1$ ,  $D(X) = \frac{1}{6}$ . 由  $Y \sim \pi(1)$  知

$$E(Y) = \frac{1}{\lambda} = 1, \quad D(Y) = \frac{1}{\lambda^2} = 1$$

由  $X$  与  $Y$  相互独立得  $E(XY) = E(X)E(Y) = 1$

$$E((XY)^2) = E(X^2)E(Y^2) = [D(X) + (E(X))^2][D(Y) + (E(Y))^2] = \frac{7}{6} \cdot 2 = \frac{7}{3}$$

$$D(XY) = E((XY)^2) - (E(XY))^2 = \frac{7}{3} - 1 = \frac{4}{3} \quad (\text{书上答案为 } \frac{7}{6} = E(X^2)).$$

18. 利用书上 P237 答案中给的离散型随机变量的例子:

写出联合及边缘分布律如下

X \ Y	Y				
		-1	0	1	$P\{X=i\}$
	X				



-1	1/8	1/8	1/8	3/8
0	1/8	0	1/8	2/8
1	1/8	1/8	1/8	3/8
$P\{Y=j\}$	3/8	2/8	3/8	

由表可知 $X$ 与 $Y$ 不是相互独立的,即

$$P(X=x_i, Y=y_j) \neq P(X=x_i) \cdot P(Y=y_j)$$

再由

$$E(XY) = \sum_{j=1}^3 \sum_{i=1}^3 x_i y_j p_{ij} = (-1)^2 \times \frac{1}{8} + 0 + (-1) \times 1 \times \frac{1}{8} + 0 + 0 + 0 + 1 \times (-1) \times \frac{1}{8} + 1^2 \times \frac{1}{8} = 0$$

$$E(Y) = E(X) = (-1) \times \frac{3}{8} + 0 \times \frac{2}{8} + 1 \times \frac{3}{8} = 0$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = 0,$$

$\therefore \rho_{XY} = 0$ , 故 $X$ 与 $Y$ 也不相关.

$$19. (1) 1 = F(+\infty, +\infty) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} f(x, y) dy dx = \int_0^1 \int_0^1 k(x+y) dy dx = k$$

$$(2) f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_0^1 (x+y) dy, & 0 \leq x \leq 1 \\ 0, & \text{其它} \end{cases} = \begin{cases} x+1/2, & 0 \leq x \leq 1 \\ 0, & \text{其它} \end{cases},$$

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} y+1/2, & 0 \leq y \leq 1 \\ 0, & \text{其它} \end{cases}.$$

$$E(Y) = E(X) = \int_{-\infty}^{+\infty} x f_X(x) dx = \int_0^1 x(x+\frac{1}{2}) dx = \frac{7}{12}.$$

$$E(Y^2) = E(X^2) = \int_{-\infty}^{+\infty} x^2 f_X(x) dx = \int_0^1 x^2(x+\frac{1}{2}) dx = \frac{5}{12}.$$

$$D(Y) = D(X) = E(X^2) - [E(X)]^2 = \frac{5}{12} - \frac{49}{144} = \frac{11}{144}.$$

$$E(XY) = \int_0^1 \int_0^1 xy(x+y) dy dx = \int_0^1 \left[ x^2 \frac{y^2}{2} + x \frac{y^3}{3} \right]_0^1 dx = \int_0^1 (\frac{x^2}{2} + \frac{x}{3}) dx = \frac{1}{3}.$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{1}{3} - \frac{49}{144} = -\frac{1}{144}.$$

$$\rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)D(Y)}} = \frac{-1/144}{11/144} = -\frac{1}{11}.$$

$$D(X+Y) = D(X) + D(Y) + 2\text{Cov}(X, Y) = \frac{11}{144} + \frac{11}{144} + 2 \times (-\frac{1}{144}) = \frac{5}{36}.$$

$$20. X+Y=n, \quad X \sim B(n, \frac{1}{2}), \quad Y \sim B(n, \frac{1}{2}).$$

$$\begin{aligned}
 \text{Cov}(X, Y) &= E(X(n-X)) - EX \cdot E(n-X) \\
 &= nEX - E(X^2) - nEX + (EX)^2 \\
 &= -DX = -npq = -\frac{n}{4} \\
 \rho_{XY} &= \frac{-DX}{\sqrt{DX} \cdot \sqrt{DY}} = \frac{-DX}{\sqrt{DX} \cdot \sqrt{DX}} = -1
 \end{aligned}$$

$$\begin{aligned}
 21. (1) Y_i &= X_i - \bar{X} = \frac{n-1}{n} X_i - \frac{n-1}{n} \sum_{\substack{k=1 \\ k \neq i}}^n X_k \\
 D(Y_i) &= \left(\frac{n-1}{n}\right)^2 D(X_i) + \frac{1}{n^2} \sum_{\substack{k=1 \\ k \neq i}}^n D(X_k) = \frac{(n-1)^2}{n^2} \cdot 1 + \frac{1}{n^2} \sum_{\substack{k=1 \\ k \neq i}}^n 1 \\
 &= \frac{(n-1)^2}{n^2} + \frac{n-1}{n^2} = \frac{n-1}{n}
 \end{aligned}$$

$$\begin{aligned}
 (2) \quad E(Y_1) &= E(X_1) - E(\bar{X}) = 0 - 0 = 0 = E(Y_n) \\
 E(Y_1 Y_n) &= E((X_1 - \bar{X})(X_n - \bar{X})) \\
 &= E(X_1 X_n - X_1 \bar{X} - X_n \bar{X} + \bar{X}^2) \\
 &= E(X_1 X_n) - 2E(X_1 \bar{X}) + E(\bar{X}^2) \\
 &= E(X_1)E(X_n) - \frac{2}{n} E(X_1^2 + \sum_{k=2}^n X_1 X_k) + D\bar{X} + (E\bar{X})^2 \\
 &= 0 \cdot 0 - \frac{2}{n} [(DX_1 + (EX_1)^2) + \sum_{k=2}^n 0] + \frac{1^2}{n} + 0^2 \\
 &= 0 \cdot 0 - \frac{2}{n} (1 + 0^2 + 0) + \frac{1}{n} = -\frac{1}{n} \\
 \text{Cov}(Y_1, Y_n) &= E(Y_1 Y_n) - E(Y_1)E(Y_n) = -\frac{1}{n}
 \end{aligned}$$

22. ♦任意♦♦λ有  $E[(\lambda X + Y)^2] \geq 0$ , 即

$$E(\lambda^2 X^2 + 2\lambda XY + Y^2) \geq 0$$

或  $E(X^2) \cdot \lambda^2 + 2E(XY) \cdot \lambda + E(Y^2) \geq 0$

$$\because E(X^2) > 0$$

$$\therefore \Delta = (2E(XY))^2 - 4E(X^2)E(Y^2) \leq 0$$

$$\text{即 } (E(XY))^2 \leq E(X^2)E(Y^2).$$