

Tutorial 4

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Review

There are n random samples $\mathbf{X}_1, \mathbf{X}_2, \dots, \mathbf{X}_n$ from the $N_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$. Let $Z_j = \mathbf{a}'\mathbf{X}_j, j = 1, 2, \dots, n$. Then the sample mean and variance of the observed values z_1, z_2, \dots, z_n are

$$\bar{z} = \mathbf{a}'\bar{\mathbf{x}} \quad \text{and} \quad s_z^2 = \mathbf{a}'\mathbf{S}\mathbf{a}$$

where $\bar{\mathbf{x}}$ and \mathbf{S} are the sample mean vector and covariance matrix of the \mathbf{X}_j 's respectively.

Given a particular \mathbf{a} , the confidence interval is that set of $\mathbf{a}'\boldsymbol{\mu}$ values for which

$$|t| = \left| \frac{\sqrt{n}(\mathbf{a}'\bar{\mathbf{x}} - \mathbf{a}'\boldsymbol{\mu})}{\sqrt{\mathbf{a}'\mathbf{S}\mathbf{a}}} \right| \leq t_{n-1}(\alpha/2)$$

Simultaneously for all \mathbf{a} , the interval

$$\left(\mathbf{a}'\bar{\mathbf{x}} - \sqrt{\frac{p(n-1)}{n(n-p)} F_{p,n-p}(\alpha) \mathbf{a}'\mathbf{S}\mathbf{a}}, \quad \mathbf{a}'\bar{\mathbf{x}} + \sqrt{\frac{p(n-1)}{n(n-p)} F_{p,n-p}(\alpha) \mathbf{a}'\mathbf{S}\mathbf{a}} \right)$$

will contain $\mathbf{a}'\boldsymbol{\mu}$ with probability $1 - \alpha$

Example 1

Perspiration from 20 healthy females was analyzed. Three components, X_1 = sweat rate, X_2 = sodium content, and X_3 = potassium content, were measured. For the data the computer calculation provides the sample mean vector, the sample covariance matrix and its inverse as follows:

$$\bar{\mathbf{x}} = \begin{pmatrix} 4.640 \\ 45.400 \\ 9.965 \end{pmatrix}, \quad \mathbf{S} = \begin{pmatrix} 2.879 & 10.010 & -1.810 \\ 10.010 & 199.788 & -5.640 \\ -1.810 & -5.640 & 3.628 \end{pmatrix}$$

and

$$\mathbf{S}^{-1} = \begin{pmatrix} .586 & -.022 & -.258 \\ -.022 & .006 & -.002 \\ .258 & -.002 & .402 \end{pmatrix}$$

Solve the following two problems under the significant level $\alpha = 0.1$

- (a) Construct the simultaneous confidence interval for $2\mu_1 - \mu_3$ and μ_2 .
(b) For three variables, testing $H_0 : \boldsymbol{\mu} = (5, 45, 10)'$ against $H_1 : \boldsymbol{\mu} \neq (5, 45, 10)'$.
(Useful formula: the T^2 statistic for a test $H_0 : \boldsymbol{\mu} = \boldsymbol{\mu}_0$ against $H_1 : \boldsymbol{\mu} \neq \boldsymbol{\mu}_0$ is given by

$$T^2 = n(\bar{\mathbf{x}} - \boldsymbol{\mu}_0)' \mathbf{S}^{-1} (\bar{\mathbf{x}} - \boldsymbol{\mu}_0)$$

its corresponding F-statistic is $F = \frac{n-p}{(n-1)p} T^2$. The latter follows $F(p, n-p)$ distribution when the null hypothesis is true.
 $F_{3,17}(.10) = 2.44$, $F_{1,19}(0.10) = 2.9899$, $t_{19}(0.05) = 1.7921$.)

(a) Suppose the 20 random samples X_1, X_2, \dots, X_{20} came from the $N_3(\mu, \Sigma)$. \bar{x} and S are the sample mean vector and covariance matrix of the X_j 's respectively. Let $\mathbf{a}_1' = (2, 0, -1)$, $\mathbf{a}_2' = (0, 1, 0)$
The simultaneous confidence interval for $2\mu_1 - \mu_3$ and μ_2 .

$$\left(\mathbf{a}_1' \bar{x} - \sqrt{\frac{57}{340} F_{3,17}(0.1) \mathbf{a}_1' \mathbf{S} \mathbf{a}_1}, \quad \mathbf{a}_1' \bar{x} + \sqrt{\frac{57}{340} F_{3,17}(0.1) \mathbf{a}_1' \mathbf{S} \mathbf{a}_1} \right)$$

and

$$\left(\mathbf{a}_2' \bar{x} - \sqrt{\frac{57}{340} F_{3,17}(0.1) \mathbf{a}_2' \mathbf{S} \mathbf{a}_2}, \quad \mathbf{a}_2' \bar{x} + \sqrt{\frac{57}{340} F_{3,17}(0.1) \mathbf{a}_2' \mathbf{S} \mathbf{a}_2} \right)$$

$$(b)\mu_0 = (5, 45, 10)'$$

$$T^2 = 20 (\bar{\mathbf{x}} - \mu_0)' \mathbf{S}^{-1} (\bar{\mathbf{x}} - \mu_0) = 1.6758$$

$$F = \frac{n-p}{(n-1)p} T^2 = \frac{17}{19 \times 3} T^2 = 0.4998 < F_{3,17}(.10) = 2.44$$

Thus H_0 can't be rejected.

Comparing Mean Vectors from Two Populations

Result 4.7 If $X_{11}, X_{12}, \dots, X_{1n_1}$ is a random sample of size n_1 from $N_p(\mu_1, \Sigma)$ and $\mathbf{X}_{21}, \mathbf{X}_{22}, \dots, \mathbf{X}_{2n_2}$ is an independent random sample size n_2 from $N_p(\mu_2, \Sigma)$, then

$T^2 = [\bar{X}_1 - \bar{X}_2 - (\mu_1 - \mu_2)]' \left[\frac{1}{n_1} + \frac{1}{n_2} S_{pooled} \right]^{-1} [\bar{X}_1 - \bar{X}_2 - (\mu_1 - \mu_2)]$ is distributed as

$$\frac{(n_1 + n_2 - 2)p}{n_1 + n_2 - p - 1} F_{p, n_1 + n_2 - p - 1}$$

Consequently,

$$P(T^2 \leq c^2) = 1 - \alpha$$

where

$$c^2 = \frac{(n_1 + n_2 - 2)p}{n_1 + n_2 - p - 1} F_{p, n_1 + n_2 - p - 1}(\alpha)$$

Testing for Equality of Covariance Matrices

With g populations, the null hypothesis is

$$H_0 : \Sigma_1 = \Sigma_2 = \cdots = \Sigma_g = \Sigma$$

where Σ_l is the covariance matrix for the l th population, $l = 1, 2, \dots, g$, and Σ is the presumed common covariance matrix. The alternative hypothesis is that at least two of the covariance matrices are not equal. Here n_l is the sample size for the l th group, \mathbf{S}_l is the l th group sample covariance matrix and \mathbf{S}_{pool} is the pooled sample covariance matrix given by

$$\mathbf{S}_{\text{pooled}} = \frac{1}{\sum_l (n_l - 1)} \{ (n_1 - 1) \mathbf{S}_1 + (n_2 - 1) \mathbf{S}_2 + \cdots + (n_g - 1) \mathbf{S}_g \}$$

Set

$$u = \left[\sum_I \frac{1}{(n_I - 1)} - \frac{1}{\sum_I (n_I - 1)} \right] \left[\frac{2p^2 + 3p - 1}{6(p + 1)(g - 1)} \right]$$

$$M = \left\{ \left[\sum_{\ell} (n_{\ell} - 1) \right] \ln |\mathbf{S}_{\text{pooled}}| - \sum_{\ell} [(n_{\ell} - 1) \ln |\mathbf{S}_{\ell}|] \right\}$$

where p is the number of variables and g is the number of groups. Then Box's test statistic

$$C = (1 - u)M = (1 - u) \left\{ \left[\sum_I (n_I - 1) \right] \ln |\mathbf{S}_{\text{pooled}}| - \sum_I [(n_I - 1) \ln |\mathbf{S}_I|] \right\}$$

has an approximate χ^2 distribution with

$$\nu = g \frac{1}{2} p(p + 1) - \frac{1}{2} p(p + 1) = \frac{1}{2} p(p + 1)(g - 1)$$

degrees of freedom. At significance level α , reject H_0 if

$$C > \chi_{p(p+1)(g-1)/2}^2(\alpha)$$

Example 2

Fifty bars of soap are manufactured in each of two ways. Two characteristics, X_1 = lather and X_2 = mildness, are measured. The summary statistics for the bars produced by methods 1 and 2 are

$$\bar{\mathbf{x}}_1 = \begin{bmatrix} 8.3 \\ 4.1 \end{bmatrix}, \quad \mathbf{S}_1 = \begin{bmatrix} 2 & 1 \\ 1 & 6 \end{bmatrix}, \quad \bar{\mathbf{x}}_2 = \begin{bmatrix} 10.2 \\ 3.9 \end{bmatrix}, \quad \mathbf{S}_2 = \begin{bmatrix} 2 & 1 \\ 1 & 4 \end{bmatrix}$$

(a) Testing $H_0 : \Sigma_1 = \Sigma_2$ against $\Sigma_1 \neq \Sigma_2$ at the significant level $\alpha = 0.10$.

(b) Testing $H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2$ against $\boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2$ at the significant level $\alpha = 0.10$. where $\boldsymbol{\mu}_1, \boldsymbol{\mu}_2$ are mean vectors for the bars produced by method 1 and 2. Σ_1 and Σ_2 are their covariance matrices. (Box's test statistic $C = (1 - u)M$ where

$$u = \left[\sum_{\ell} \frac{1}{(n_{\ell} - 1)} - \frac{1}{\sum_{\ell} (n_{\ell} - 1)} \right] \left[\frac{2p^2 + 3p - 1}{6(p + 1)(g - 1)} \right]$$

$$M = \left\{ \left[\sum_{\ell} (n_{\ell} - 1) \right] \ln |\mathbf{S}_{\text{pooled}}| - \sum_{\ell} [(n_{\ell} - 1) \ln |\mathbf{S}_{\ell}|] \right\}$$

$$F_{2,97}(0.10) = 2.3581, \chi_3^2(0.10) = 6.2514)$$

$$(a) n_1 = n_2 = 50, p = 2, g = 2,$$

$$S_{pooled} = \frac{1}{n_1 + n_2 - 2} [(n_1 - 1)S_1 + (n_2 - 1)S_2] = \begin{bmatrix} 2 & 1 \\ 1 & 5 \end{bmatrix}$$

$$u = \left[\frac{1}{n_1 - 1} + \frac{1}{n_2 - 1} - \frac{1}{n_1 + n_2 - 2} \right] \left[\frac{2p^2 + 3p - 1}{6(p+1)(g-1)} \right] = 0.0221$$

$$M = (n_1 + n_2 - 2) \ln |\mathbf{S}_{pooled}| - (n_1 - 1) \ln |\mathbf{S}_1| - (n_2 - 1) \ln |\mathbf{S}_2| = 2.4815$$

$$p(p+1)(g-1)/2 = 3$$

$$C = (1 - u)M = 2.4267 < \chi_3^2(0.10) = 6.2514$$

(b) $\Sigma_1 = \Sigma_2$ If $\mu_1 = \mu_2$, then

$$T^2 = [\bar{x}_1 - \bar{x}_2]' \left[\left(\frac{1}{n_1} + \frac{1}{n_2} \right) S_{\text{pooled}} \right]^{-1} [\bar{x}_1 - \bar{x}_2] \text{ is}$$

distributed as $\frac{(n_1+n_2-2)p}{n_1+n_2-p-1} F_{p, n_1+n_2-p-1}$

$$T^2 = [\bar{x}_1 - \bar{x}_2]' \left[\left(\frac{1}{n_1} + \frac{1}{n_2} \right) S_{\text{pooled}} \right]^{-1} [\bar{x}_1 - \bar{x}_2] = 52.4722$$

$$F = \frac{n_1 + n_2 - p - 1}{(n_1 + n_2 - 2)p} T^2 = 25.9684 > F_{2,97}(0.10) = 2.3581$$

Then H_0 can be rejected.

Exmple 3

Using Moody's bond ratings, samples of 20 Aa (middle-high quality) corporate bonds and 20 Baa(top-medium quality) corporate were selected. For each of the corresponding companies, the ratio X_1 = current ratio (a measure of short-term liquidity), X_2 = debt to equity ratio (a measure of financial risk or leverage) were recorded. The summary statistics are as follows:

Aa bound companies:

$$\bar{\mathbf{x}}_1 = \begin{bmatrix} 2.287 \\ .347 \end{bmatrix}, \quad \mathbf{S}_1 = \begin{bmatrix} .459 & -.026 \\ -.026 & .030 \end{bmatrix}, \quad n_1 = 20$$

Baa bond companies:

$$\bar{\mathbf{x}}_2 = \begin{bmatrix} 2.404 \\ .524 \end{bmatrix}, \quad \mathbf{S}_2 = \begin{bmatrix} .944 & .002 \\ .002 & .024 \end{bmatrix}, \quad n_2 = 20$$

(a) Test $H_0 : \Sigma_1 = \Sigma_2$ vs $H_1 : \Sigma_1 \neq \Sigma_2$ at the significant level $\alpha = 0.05$.

(b) Test $H_0 : \mu_1 - \mu_2 = 0$ vs $H_1 : \mu_1 - \mu_2 \neq 0$ at the significant level $\alpha = 0.05$

(Useful formula: the T^2 statistics for two populations is given by

$$T^2 = [\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2 - (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2)]' \left[\left(\frac{1}{n_1} + \frac{1}{n_2} \right) \mathbf{S}_{\text{pooled}} \right]^{-1} [\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2 - (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2)]$$

when $\Sigma_1 = \Sigma_2$, and when $\Sigma_1 \neq \Sigma_2$

$$T^2 = [\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2 - (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2)]' \left[\frac{1}{n_1} \mathbf{S}_1 + \frac{1}{n_2} \mathbf{S}_2 \right]^{-1} [\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2 - (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2)]$$
$$\mathbf{S}_{\text{pooled}} = \frac{n_1-1}{n_1+n_2-2} \mathbf{S}_1 + \frac{n_2-1}{n_1+n_2-2} \mathbf{S}_2$$

$$F_{2,37}(0.05) = 3.25, \chi_2^2(0.05) = 5.99, \chi_3^2(0.05) = 7.8147)$$

$$(a) S_{pooled} = \frac{1}{n_1+n_2-2}[(n_1-1)S_1 + (n_2-1)S_2] = \begin{bmatrix} 0.7015 & -0.012 \\ -0.0120 & 0.027 \end{bmatrix}$$

$$u = \left[\frac{1}{n_1-1} + \frac{1}{n_2-1} - \frac{1}{n_1+n_2-2} \right] \left[\frac{2p^2+3p-1}{6(p+1)(g-1)} \right] = 0.0570$$

$$M = (n_1 + n_2 - 2) \ln |\mathbf{S}_{pooled}| - (n_1 - 1) \ln |\mathbf{S}_1| - (n_2 - 1) \ln |\mathbf{S}_2| = 3.3238$$

$$p(p+1)(g-1)/2 = 3$$

$$C = (1 - u)M = 3.1344 < \chi_3^2(0.05) = 7.8147$$

(b) $\Sigma_1 = \Sigma_2$ If $\mu_1 - \mu_2 = 0$, then

$$T^2 = [\bar{x}_1 - \bar{x}_2]' \left[\left(\frac{1}{n_1} + \frac{1}{n_2} \right) S_{\text{pooled}} \right]^{-1} [\bar{x}_1 - \bar{x}_2] \text{ is}$$

distributed as $\frac{(n_1+n_2-2)p}{n_1+n_2-p-1} F_{p, n_1+n_2-p-1}$

$$T^2 = [\bar{x}_1 - \bar{x}_2]' \left[\left(\frac{1}{n_1} + \frac{1}{n_2} \right) S_{\text{pooled}} \right]^{-1} [\bar{x}_1 - \bar{x}_2] = 12.15328$$

$F = \frac{n_1+n_2-p-1}{(n_1+n_2-2)p} T^2 = 5.9167 > F_{2,37}(0.05) = 3.25$ Then H_0 can be rejected.

Review

The "distance" of the point $[x_1, x_2, \dots, x_p]'$ to origin

$$\begin{aligned}(\text{distance})^2 &= a_{11}x_1^2 + a_{22}x_2^2 + \dots + a_{pp}x_p^2 \\ &\quad + 2(a_{12}x_1x_2 + a_{13}x_1x_3 + \dots + a_{p-1,p}x_{p-1}x_p)\end{aligned}$$

A geometric interpretation based on the eigenvalues and eigenvectors of the matrix A .

For example, suppose $p = 2$, Then the points $\mathbf{x}' = [x_1, x_2]$ of constant distance c from the origin satisfy

$$\mathbf{x}'\mathbf{A}\mathbf{x} = a_{11}x_1^2 + a_{22}x_2^2 + 2a_{12}x_1x_2 = c^2$$

By the spectral decomposition,

$$\mathbf{A} = \lambda_1\mathbf{e}_1\mathbf{e}_1' + \lambda_2\mathbf{e}_2\mathbf{e}_2'$$

so

$$\mathbf{x}'\mathbf{A}\mathbf{x} = \lambda_1 (\mathbf{x}'\mathbf{e}_1)^2 + \lambda_2 (\mathbf{x}'\mathbf{e}_2)^2$$

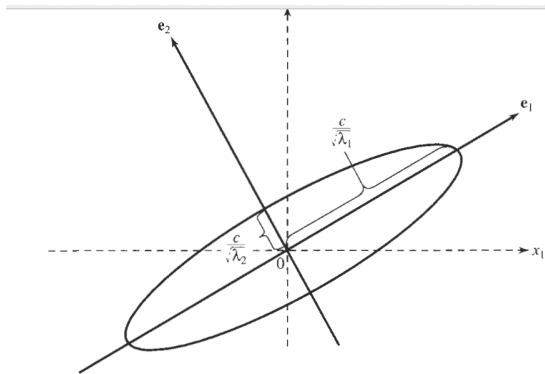


Figure 2.6 Points a constant distance c from the origin ($p = 2, 1 \leq \lambda_1 < \lambda_2$).

Assignment 1 5

Consider the sets of points (x_1, x_2) whose "distance" from the origin are given by

$$c^2 = 4x_1^2 + 3x_2^2 + 2\sqrt{2}x_1x_2$$

for $c^2 = 1$ and for $c^2 = 4$. Determine the major and minor axes of the ellipse of constant distances and their associated length. Sketch the ellipse of constant distances and comment on their positions.

Solutions

$$c^2 = 4x_1^2 + 3x_2^2 + 2\sqrt{2}x_1x_2 = \begin{pmatrix} x_1 & x_2 \end{pmatrix} \begin{pmatrix} 4 & \sqrt{2} \\ \sqrt{2} & 3 \end{pmatrix} \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}$$

$$\text{Let } A = \begin{pmatrix} 4 & \sqrt{2} \\ \sqrt{2} & 3 \end{pmatrix}, \text{ then } |A - \lambda I| = (\lambda - 2)(\lambda - 5), \quad \lambda_1 = 2, \lambda_2 = 5$$

$$\text{When } \lambda = 2, \quad e_1 = \begin{bmatrix} -\frac{1}{\sqrt{3}} \\ \frac{\sqrt{2}}{\sqrt{3}} \end{bmatrix}$$

$$\text{When } \lambda = 5, \quad e_2 = \begin{bmatrix} \frac{\sqrt{2}}{\sqrt{3}} \\ \frac{1}{\sqrt{3}} \end{bmatrix}$$

$$\text{For } c^2 = 1, \text{ major axis is } \frac{1}{\sqrt{2}} \begin{bmatrix} -\frac{1}{\sqrt{3}} \\ \frac{\sqrt{2}}{\sqrt{3}} \end{bmatrix}, \quad \text{minor axis is } \frac{1}{\sqrt{5}} \begin{bmatrix} \frac{\sqrt{2}}{\sqrt{3}} \\ \frac{1}{\sqrt{3}} \end{bmatrix}$$

For $c^2 = 4$, major axis is $\sqrt{2} \begin{bmatrix} -\frac{1}{\sqrt{3}} \\ \frac{\sqrt{2}}{\sqrt{3}} \end{bmatrix}$, minor axis is $\frac{2}{\sqrt{5}} \begin{bmatrix} \frac{\sqrt{2}}{\sqrt{3}} \\ \frac{1}{\sqrt{3}} \end{bmatrix}$ Let

$$\mathbf{x} = (x_1, x_2)'$$

$$\mathbf{x}'\mathbf{A}\mathbf{x} = \lambda_1 (\mathbf{x}'\mathbf{e}_1)^2 + \lambda_2 (\mathbf{x}'\mathbf{e}_2)^2 = c^2$$

Assume that the original coordinate axes are rotated through an angle of θ

$$x_\theta = x_1 \cos \theta + x_2 \sin \theta$$

$$y_\theta = -x_1 \sin \theta + x_2 \cos \theta$$

$$\text{Let } x_\theta = x_1 \cos \theta + x_2 \sin \theta = \mathbf{x}'\mathbf{e}_1 = x_1 \cos \frac{1}{\sqrt{3}} + x_2 \sin \frac{\sqrt{2}}{\sqrt{3}}$$

$$\theta = \arccos\left(-\frac{1}{\sqrt{3}}\right) = 0.6959$$

Thus, the ellipse rotates an angle between 90 and 180 from the original axes.

Assignment 1 3

Are the following distance functions valid for the distance from the origin ?
Explain. (a) $x_1^2 + 4x_2^2 + x_1x_2 = (\text{distance})^2$ (b) $x_1^2 - 2x_2^2 = (\text{distance})^2$

The points $\mathbf{x}' = [x_1, x_2]$ of constant distance c from the origin satisfy

$$\mathbf{x}'\mathbf{A}\mathbf{x} = a_{11}x_1^2 + a_{22}x_2^2 + 2a_{12}x_1x_2 = c^2$$

If \mathbf{A} is a positive definite matrix, the distant function is valid.

(a) Because $\begin{pmatrix} 1 & 0.5 \\ 0.5 & 4 \end{pmatrix}$ is a positive definite matrix, and the origin is $(0, 0)$, so the distant function is valid.

(b) Because $\begin{pmatrix} 1 & 0 \\ 0 & -2 \end{pmatrix}$ is not a positive definite matrix, so the distant function is not valid.