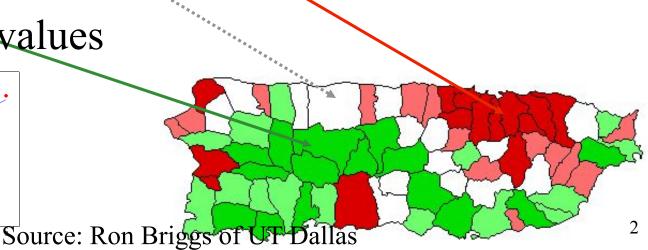
Spatial Autocorrelation of Areal Data

Positive spatial autocorrelation

- high values
 surrounded by nearby high values
- intermediate values surrounded by nearby intermediate values
- low values surrounded by nearby low values

} 0

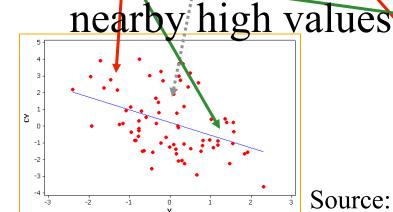


2002 population

density

Negative spatial autocorrelation

- high values surrounded by nearby low values
- intermediate values surrounded by nearby intermediate values
- low values surrounded by



Source: Ron Briggs of UT Dallas

competition for space

Grocery store density

Spatial Weight Matrix

- Core concept in statistical analysis of areal data
- Two steps involved:
 - define which relationships between observations are to be given a nonzero weight, i.e., define spatial neighbors
 - assign weights to the neighbors

Making the neighbors and weights is not easy as

it seems to be

– Which states are near Texas?

Spatial Neighbors

Contiguity-based neighbors

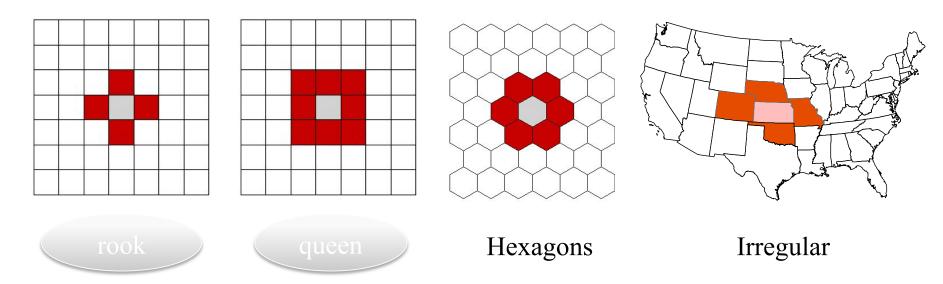
- Zone i and j are neighbors if zone i is contiguity or adjacent to zone j
- But what constitutes contiguity?

Distance-based neighbors

- Zone i and j are neighbors if the distance between them are less than the threshold distance
- But what distance do we use?

Contiguity-based Spatial Neighbors

- Sharing a border or boundary
 - Rook: sharing a border
 - Queen: sharing a border <u>or</u> a point



Which use?

Example

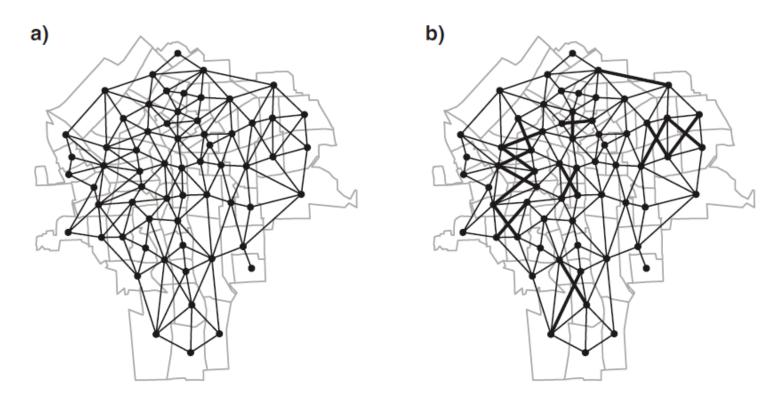
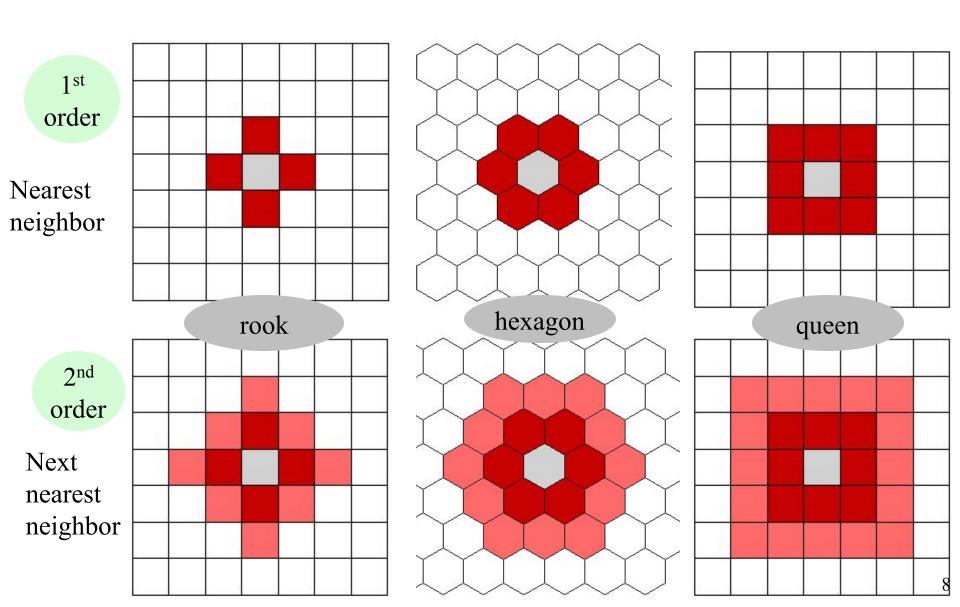


Fig. 9.3. (a) Queen-style census tract contiguities, Syracuse; (b) Rook-style contiguity differences shown as thicker lines

Source: Bivand and Pebesma and Gomez-Rubio

Higher-Order Contiguity



Distance-based Neighbors

- How to measure distance between polygons?
- Distance metrics
 - 2D Cartesian distance (projected data)
 - 3D spherical distance/great-circle distance (lat/long data)
 - Haversine formula

```
Haversine a = \sin^2(\Delta \phi/2) + \cos(\phi_1).\cos(\phi_2).\sin^2(\Delta \lambda/2)
formula: c = 2.a \tan 2(\sqrt{a}, \sqrt{(1-a)})
d = R.c
```

Distance-based Neighbors

• k-nearest neighbors

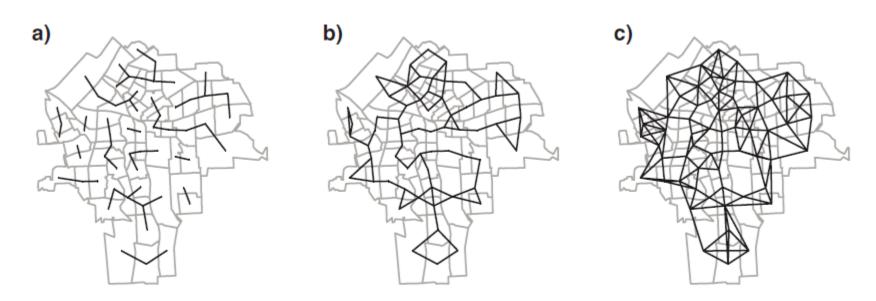


Fig. 9.5. (a) k = 1 neighbours; (b) k = 2 neighbours; (c) k = 4 neighbours

Source: Bivand and Pebesma and Gomez-Rubio

Distance-based Neighbors

thresh-hold distance (buffer)

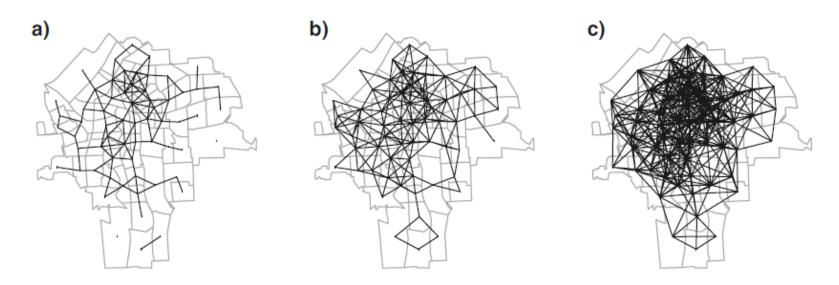
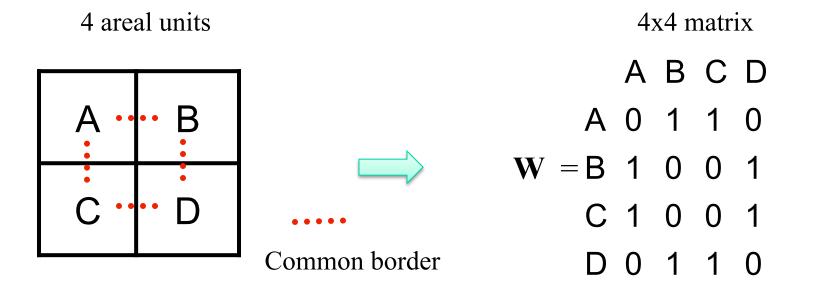


Fig. 9.6. (a) Neighbours within 1,158 m; (b) neighbours within 1,545 m; (c) neighbours within 2,317 m

Source: Bivand and Pebesma and Gomez-Rubio

A Simple Example for Rook case

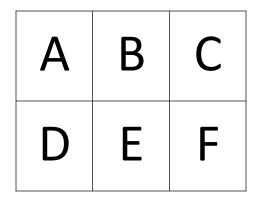
- Matrix contains a:
 - 1 if share a border
 - 0 if do not share a border



Style of Spatial Weight Matrix

- Row
 - a weight of unity for each neighbor relationship
- Row standardization
 - Symmetry not guaranteed
 - can be interpreted as allowing the calculation of average values across neighbors
- General spatial weights based on distances

Row vs. Row standardization



Divide each number by the **row sum**

Total number of neighbors
--some have more than others



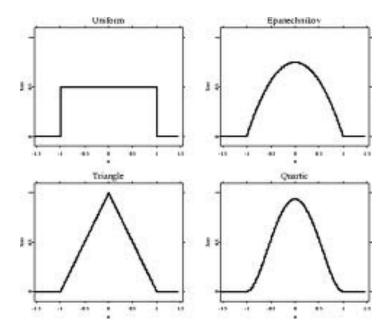
	A	В	С	D	E	F	Row Sum
Α	0	1	0	1	0	0	2
В	1	0	1	0	1	0	3
С	0	1	0	0	0	1	2
D	1	0	0	0	1	0	2
Ε	0	1	0	1	0	1	3
F	0	0	1	0	1	0	2

Row standardized --usually use this

	A	В	С	D	E	F	Row Sum
Α	0.0	0.5	0.0	0.5	0.0	0.0	1
В	0.3	0.0	0.3	0.0	0.3	0.0	1
С	0.0	0.5	0.0	0.0	0.0	0.5	1
D	0.5	0.0	0.0	0.0	0.5	0.0	1
E	0.0	0.3	0.0	0.3	0.0	0.3	1
F	0.0	0.0	0.5	0.0	0.5	0.0	1

General Spatial Weights Based on Distance

- Decay functions of distance
 - Most common choice is the inverse (reciprocal) of the distance between locations i and j $(w_{ij} = 1/d_{ij})$
 - Other functions also used
 - inverse of <u>squared</u> distance $(w_{ij} = 1/d_{ij}^2)$, or
 - negative exponential $(w_{ij} = e^{-d} \ or \ w_{ij} = e^{-d^2})$



Measure of Spatial Autocorrelation

Global Measures and Local Measures

Global Measures

- A single value which applies to the entire data set
 - The same pattern or process occurs over the entire geographic area
 - An average for the entire area

Local Measures

- A value calculated for <u>each</u> observation unit
 - Different patterns or processes may occur in different parts of the region
 - A unique number for each location
- Global measures usually can be decomposed into a combination of local measures

Global Measures and Local Measures

- Global Measures
 - Moran's I, Getis-Ord's G
- Local Measures
 - Local Moran's I, Getis-Ord's G

Formula for Moran's I

$$I = \frac{N \sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij} (x_i - \overline{x}) (x_j - \overline{x})}{(\sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij}) \sum_{i=1}^{n} (x_i - \overline{x})^2}$$

Where:

 $\frac{N}{\overline{X}}$ is the number of observations (points or polygons) is the mean of the variable X_i is the variable value at a particular location X_i is the variable value at another location W_{ij} is a weight indexing location of i relative to j

Moran's I

• Expectation of Moran's I under no spatial autocorrelation

$$E(I) = -1/(N-1)$$

- Variance of Moran's is complex and exact equation is given at textbook d&G&L
- [-1, 1]

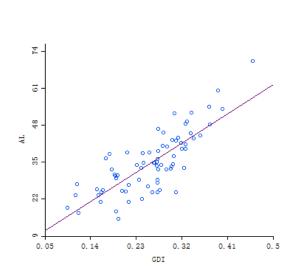
Moran's I and Correlation Coefficient

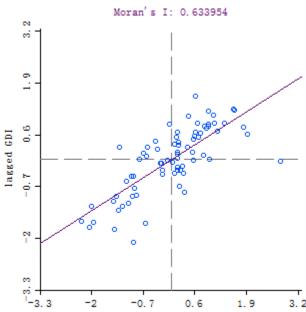
Correlation Coefficient [-1, 1]

- Relationship between <u>two</u> different variables

Moran's I [-1, 1]

- Spatial autocorrelation and often involves <u>one</u> (spatially indexed) variable only
- Correlation between observations of a spatial variable at location X and "spatial lag" of X formed by averaging all the observation at neighbors of X





$$\frac{\sum_{i=1}^{n} 1(y_i - \overline{y})(x_i - \overline{x})/n}{\sqrt{\sum_{i=1}^{n} (y_i - \overline{y})^2} \sqrt{\sum_{i=1}^{n} (x_i - \overline{x})^2}}$$

Correlation Coefficient

Note the similarity of the numerator (top) to the measures of spatial association discussed earlier if we view Yi as being the Xi for the neighboring polygon

(see next slide)

$$\frac{N \sum_{i=1}^{n} \sum_{j=1}^{n} W_{ij}(X_{i} - \overline{X})(X_{j} - \overline{X})}{(\sum_{i=1}^{n} \sum_{j=1}^{n} W_{ij}) \sum_{i=1}^{n} (X_{i} - \overline{X})^{2}}$$

Spatial auto-correlation

$$\frac{\sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij}(x_{i} - \overline{x})(x_{j} - \overline{x}) / \sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij}}{\sqrt{\sum_{i=1}^{n} (x_{i} - \overline{x})^{2}} \sqrt{\sum_{i=1}^{n} (x_{i} - \overline{x})^{2}}} \sqrt{\frac{\sum_{i=1}^{n} (x_{i} - \overline{x})^{2}}{n}}$$

Source: Ron Briggs of UT Dallas

$$\frac{\sum_{i=1}^{n} 1(y_i - \overline{y})(x_i - \overline{x})/n}{\sqrt{\sum_{i=1}^{n} (y_i - \overline{y})^2} \sqrt{\sum_{i=1}^{n} (x_i - \overline{x})^2}}$$

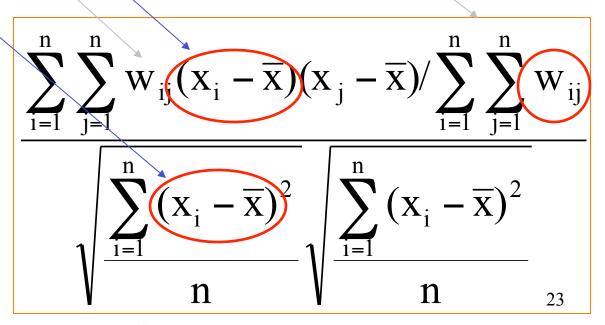
Correlation Coefficient

Spatial weights

Yi is the Xi for the neighboring polygon

$$\frac{N \sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij} (x_i - \overline{x}) (x_j - \overline{x})}{(\sum_{i=1}^{n} \sum_{j=1}^{n} w_{ij}) \sum_{i=1}^{n} (x_i - \overline{x})^2}$$

Moran's I



Source: Ron Briggs of UT Dallas

Statistical Significance Tests for Moran's I

• Based on the normal frequency distribution with

$$Z = \frac{I - E(I)}{S_{error(I)}}$$

Where: I is the calculated value for Moran's I from the sample

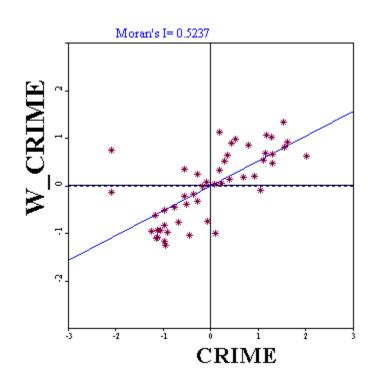
E(I) is the expected value if random

S is the standard error

- Statistical significance test
 - Monte Carlo test, as we did for spatial pattern analysis
 - Permutation test
 - Non-parametric
 - Data-driven, no assumption of the data
 - Implemented in GeoDa

Moran Scatter Plots

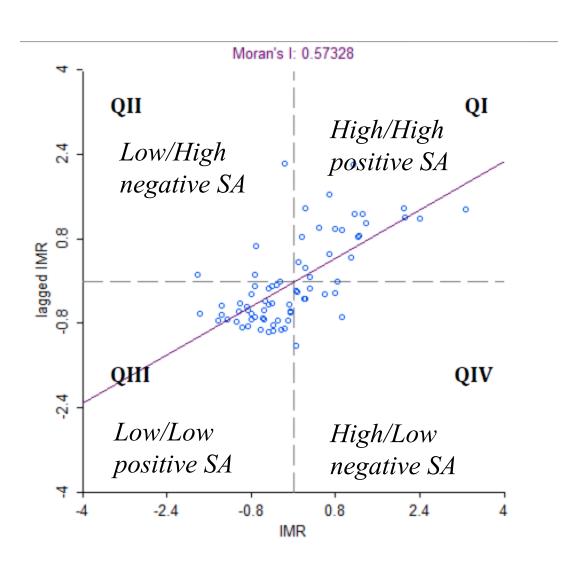
We can draw a scatter diagram between these two variables (in standardized form): **X** and **lag-X** (or W_X)



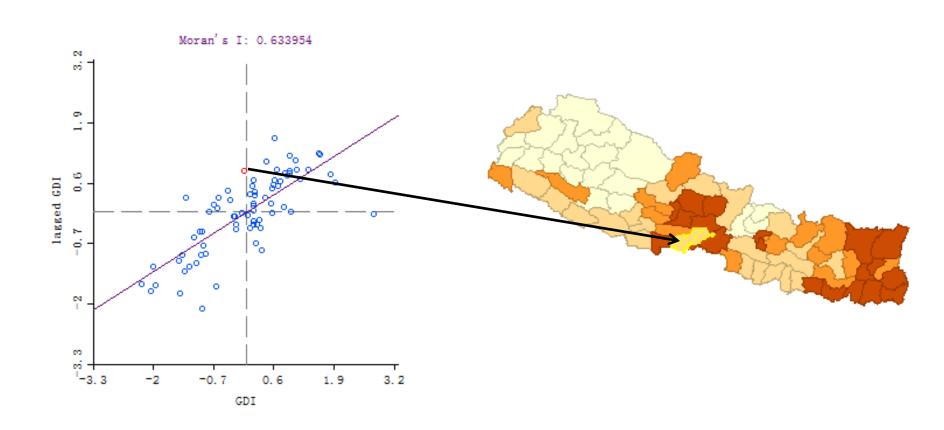


The <u>slope</u> of this *regression line* is Moran's I

Moran Scatter Plots



Moran Scatterplot: Example



Local Measures of Spatial Autocorrelation

Local Indicators of Spatial Association (LISA)

- Local versions of Moran's I, Geary's C, and the Getis-Ord G statistic
- Moran's I is most commonly used, and the local version is often called Anselin's LISA, or just LISA

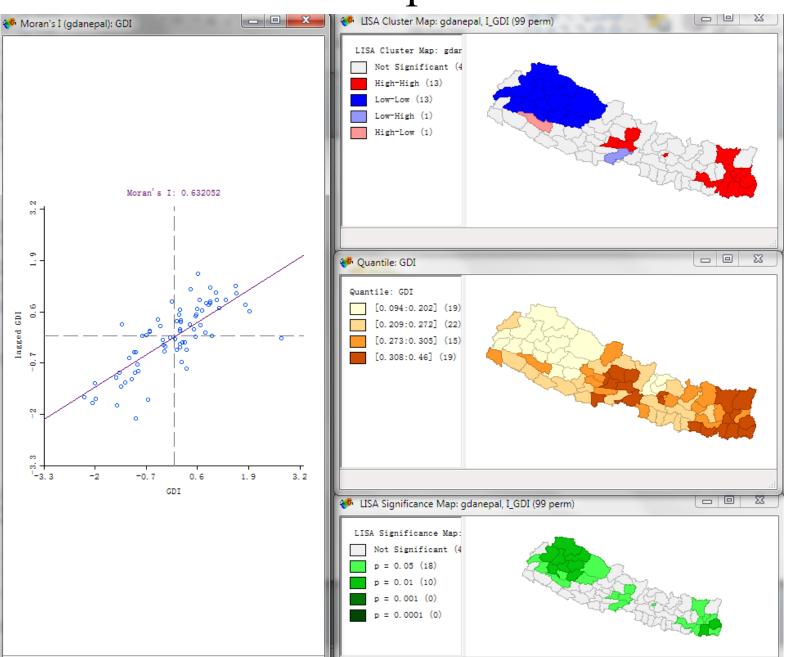
See:

Luc Anselin 1995 Local Indicators of Spatial Association-LISA Geographical Analysis 27: 93-115

Local Indicators of Spatial Association (LISA)

- The statistic is calculated for **each** areal unit in the data
- For each polygon, the index is calculated <u>based on neighboring</u> polygons with which it shares a border
- A measure is available for <u>each</u> polygon, these can be mapped to indicate how <u>spatial autocorrelation varies</u> over the study region
- Each index has an associated test statistic, we can also map which of the polygons has a <u>statistically significant relationship</u> with its neighbors, and show <u>type</u> of relationship

Example:



Calculating Anselin's LISA

• The local Moran statistic for areal unit *i* is:

$$I_i = z_i \sum_j w_{ij} z_j$$

where z_i is the original variable x_i in "standardized form"

$$z_i = \frac{x_i - x}{SD_x}$$

or it can be in "deviation form"

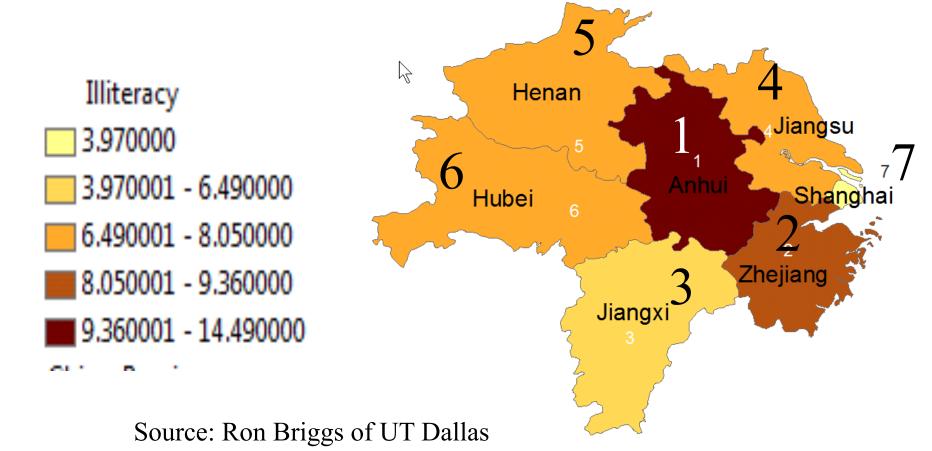
$$x_i - \overline{x}$$

and w_{ij} is the spatial weight

The summation \sum_{j}^{∞} is across each <u>row</u> i of the spatial weights matrix.

An example follows

Contiguit	y Matrix	1	2	3	4	5	6	7			
	Code	Anhui	Zhejiang	Jiangxi	Jiangsu	Henan	Hubei	Shanghai	Sum	Neighbors	Illiteracy
Anhui	1	0	1	1	1	1	1	0	5	65432	14.49
Zhejiang	2	1	0	1	1	0	0	1	4	7 4 3 1	9.36
Jiangxi	3	1	1	0	0	0	1	0	3	621	6.49
Jiangsu	4	1	1	0	0	0	0	1	3	7 2 1	8.05
Henan	5	1	0	0	0	0	1	0	2	6 1	7.36
Hubei	6	1	0	1	0	1	0	0	3	1 3 5	7.69
Shanghai	7	0	1	0	1	0	0	0	2	2 4	3.97



Contiguity Matrix and Row Standardized Spatial Weights Matrix

Contiguity	Matrix Code	1 Anhui	2 Zhejiang	3 Jiangxi	4 Jiangsu	5 Henan	6 Hubei	7 Shanghai	Sum
Anhui	1	0	1	1	1	1	1	0	5
Zhejiang	2	1	0	1	1	0	0	1	4
Jiangxi	3	1	1	0	0	0	1	0	3
Jiangsu	4	1	1	0	0	0	0	\bigcirc	3
Henan	5	1	0	0	0	0	1	0	2
Hubei	6	1	0	1	0	1	0	0	3
Shanghai	7	0	1	0	1	0	0	0	₂ (1/3)
Row Stand	dardized : Code	Spatial Weigh Anhui	its Matrix Zhejiang	Jiangxi	Jiangsu	Henan	Hubei	Shanghai	Sum
Anhui	1	0.00	0.20	0.20	0.20	0.20	0.20	0.00	1
Zhejiang	2	0.25	0.00	0.25	0.25	0.00	0.00	0.25	1
Jiangxi	3	0.33	0.33	0.00	0.00	0.00	0.33	0.00	1
Jiangsu	4	0.33	0.33	0.00	0.00	0.00	0.00	0.33	1
Henan	5	0.50	0.00	0.00	0.00	0.00	0.50	0.00	1
Hubei	6	0.33	0.00	0.33	0.00	0.33	0.00	0.00	1
Shanghai	7	0.00	0.50	0.00	0.50	0.00	0.00	0.00	1

Source: Ron Briggs of UT Dallas

Calculating standardized (z) scores

Deviations from Mean and z scores. $ x_i - x_i = x_i - x_i - x_i = x_i - x_i - x_i = x_i $							
	X	X-Xmean	X-Mean2	$z \sim z_i$	$=\frac{}{SD_x}$		
Anhui	14.49	6.29	39.55	2.101			
Zhejiang	9.36	1.16	1.34	0.387			
Jiangxi	6.49	(1.71)	2.93	(0.572)			
Jiangsu	8.05	(0.15)	0.02	(0.051)			
Henan	7.36	(0.84)	0.71	(0.281)			
Hubei	7.69	(0.51)	0.26	(0.171)			
Shanghai	3.97	(4.23)	17.90	(1.414)			
Mean and Standard I	Deviation						
Sum	57.41	0.00	62.71				
Mean	57.41	/ 7 =	8.20				
Variance	62.71	/ 7 =	8.96				
SD	√ 8.96	=	2.99		2.5		

Source: Ron Briggs of UT Dallas

Row Standardized Spatial Weights Matrix

Calculating LISA

Zhejiang 2 0.25 0.00 0.25 0.25 0.00 0.00 0.25 Jiangxi 3 0.33 0.33 0.00 0.00 0.00 0.33 0.00		Code	Anhui	Zhejiang	Jiangxi	Jiangsu	Henan	Hubei	Shanghai
Jiangxi 3 0.33 0.33 0.00 0.00 0.00 0.33 0.00	Anhui	1	0.00	0.20	0.20	0.20	0.20	0.20	0.00
	Zhejiang	2	0.25	0.00	0.25	0.25	0.00	0.00	0.25
Jiangsu 4 0.33 0.33 0.00 0.00 0.00 0.00 0.33	Jiangxi	3	0.33	0.33	0.00	0.00	0.00	0.33	0.00
	Jiangsu	4	0.33	0.33	0.00	0.00	0.00	0.00	0.33
Henan 5 0.50 0.00 0.00 0.00 0.00 0.50 0.00	Henan	5	0.50	0.00	0.00	0.00	0.00	0.50	0.00
Hubei 6 0.33 0.00 0.33 0.00 0.33 0.00 0.00	Hubei	6	0.33	0.00	0.33	0.00	0.33	0.00	0.00
Shanghai 7 0.00 0.50 0.00 0.50 0.00 0.00 0.00	Shanghai	7	0.00	0.50	0.00	0.50	0.00	0.00	0.00

W	•	
* *	1	1
	•	J

Z-Scores for row Province and its potential neighbors

		Anhui	Zhejiang	Jiangxi	Jiangsu	Henan	Hubei	Shanghai
	Zi							
Anhui	2.101	2.101	0.387	(0.572)	(0.051)	(0.281)	(0.171)	(1.414)
Zhejiang	0.387	2.101	0.387	(0.572)	(0.051)	(0.281)	(0.171)	(1.414)
Jiangxi	(0.572)	2.101	0.387	(0.572)	(0.051)	(0.281)	(0.171)	(1.414)
Jiangsu	(0.051)	2.101	0.387	(0.572)	(0.051)	(0.281)	(0.171)	(1.414)
Henan	(0.281)	2.101	0.387	(0.572)	(0.051)	(0.281)	(0.171)	(1.414)
Hubei	(0.171)	2.101	0.387	(0.572)	(0.051)	(0.281)	(0.171)	(1.414)
Shanghai	(1.414)	2.101	0.387	(0.572)	(0.051)	(0.281)	(0.171)	(1.414)

$I_i = z_i \sum_j w_{ij} z_j$

Spatial Weight Matrix multiplied by Z-Score Matrix (cell by cell multiplication)

		Anhui	Zhejiang	Jiangxi	Jiangsu	Henan	Hubei	Shanghai	SumW
	Zi								0.0
Anhui	2.101	-	0.077	(0.114)	(0.010)	(0.056)	(0.034)	-	(0.1
Zhejiang	0.387	0.525	-	(0.143)	(0.013)	-	-	(0.353)	0.0
Jiangxi	(0.572)	0.700	0.129	-	-	-	(0.057)	-	0.7
Jiangsu	(0.051)	0.700	0.129	-	-	-	-	(0.471)	0.3
Henan	(0.281)	1.050	-	-	-	-	(0.085)	-	0.9
Hubei	(0.171)	0.700	-	(0.191)	-	(0.094)	-	-	0.4
Shanghai	(1.414)	_	0.194	-	(0.025)	-	-	-	0.1

		<u> </u>
WijZj .000	LISA	Lisa from GeoDA
137)	-0.289	-0.248
.016	0.006	0.005
.772	-0.442	-0.379
.358	-0.018	-0.016
.965	-0.271	-0.233
.416	-0.071	-0.061
.168	-0.238	-0.204

Local Getis-Ord G and G* Statistics

Local Getis-Ord G

- It is the proportion of all x values in the study area accounted for by the neighbors of location *I*
- G* will include the self value

$$G_i(d) = \frac{\sum_{j} w_{ij} x_j}{\sum_{j} x_j}$$

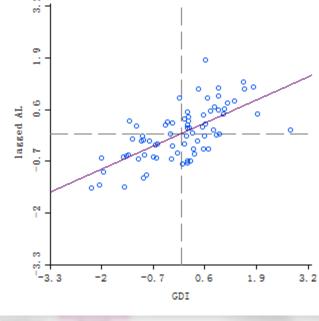
G will be <u>high</u> where <u>high</u> values cluster G will be <u>low</u> where <u>low</u> values cluster Interpreted relative to expected value if randomly distributed.

$$E(G_i(d)) = \frac{\sum_{j} w_{ij}(d)}{n-1}$$

Bivariate LISA

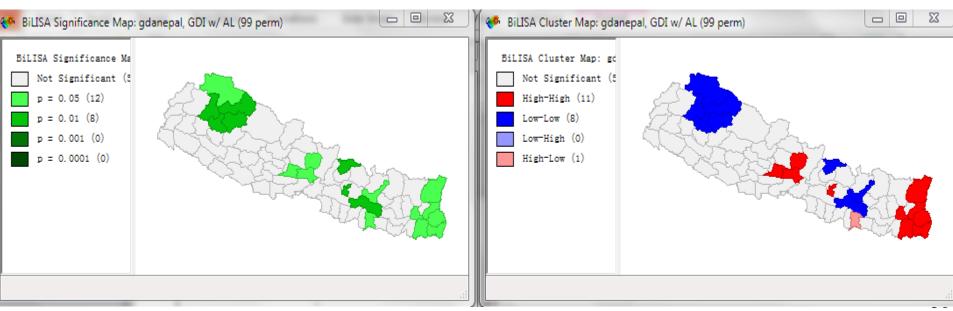
- Moran's I is the correlation between X and Lag-X--the <u>same</u> variable but in <u>nearby</u> areas
 - Univariate Moran's I
- Bivariate Moran's I is a correlation between X and a <u>different</u> variable in <u>nearby</u> areas.

Moran Significance Map for GDI vs. AL



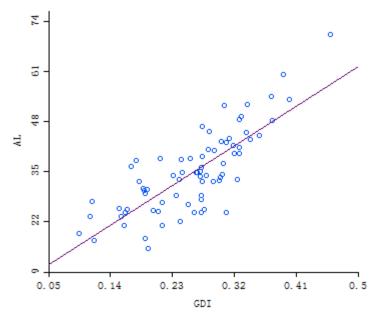
Moran Scatter Plot for GDI vs AL

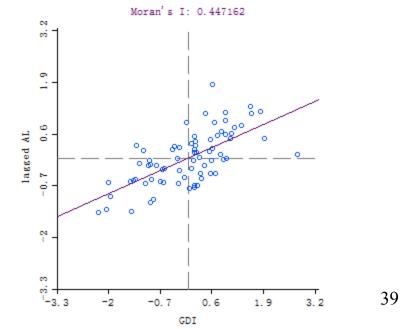
Moran's I: 0.447162



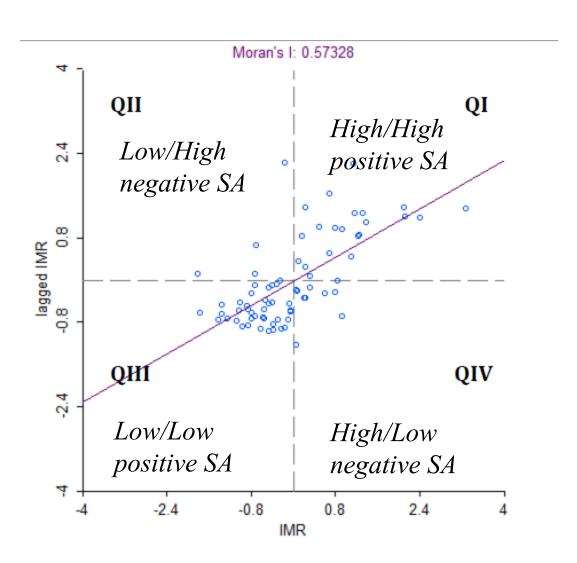
Bivariate LISA and the Correlation Coefficie

- Correlation Coefficient is the relationship between two <u>different</u> variables in the <u>same</u> area
- Bivariate LISA is a correlation between two <u>different</u> variables in an area and in <u>nearby</u> areas.





Bivariate Moran Scatter Plot



Summary

- Spatial autocorrelation of areal data
- Spatial weight matrix
- Measures of spatial autocorrelation
- Global Measure
 - Moran's I
- Local
 - LISA: Moran's I
 - Bivariate LISA
 - Significance test

• End of this topic