Generative Models for Classification and Missclassification Big Data y Machine Learning para Economía Aplicada

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Agenda

- 1 Recap
- 2 Generative Models for Classification
 - Discriminant Analysis
 - Naive Bayes
- 3 Misclassification Rates
 - ROC curve
- 4 Multiple Classes
 - KNN
 - Multinomial Logit

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Recap

- ▶ We observe (y_i, X_i) i = 1, ..., n
- ► Estimate Probabilities
 - Logit

$$p_i = rac{e^{X_ieta}}{1+e^{X_ieta}}$$

- ightharpoonup get β
 - Logit, with the $\hat{\beta}$

$$\hat{p}_i = \frac{e^{X_i \beta}}{1 + e^{X_i \hat{\beta}}}$$

Classification

$$\hat{Y}_i = 1[\hat{p}_i > c]$$

(1)

(2)

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Linear Discriminant Analysis

Reverend Bayes to the rescue: Bayes Theorem

$$Pr(Y=1|X) (4)$$

Linear Discriminant Analysis, k=1

Extensions

- ▶ If we have k > 1 predictors?
- ▶ then $X|Y \sim NM(\mu, \Sigma)$

$$f(X|Y=j) = \frac{1}{(2\pi)^{k/2} |\Sigma|^{1/2}} exp(-\frac{1}{2}(x-\mu_j)'\Sigma_j(x-\mu_j)$$
 (5)

- \blacktriangleright μ_i is the vector of the sample means in each partition j=0,1
- \triangleright Σ_i is the matrix of variance and covariances of each partition j=0,1

Linear Discriminant Analysis

- ► Why is it called linear?
- ► Note

$$p > \frac{1}{2} \iff ln(\frac{p}{(1-p)}) \tag{6}$$

Logit with one predictor

$$\beta_1 + \beta_2 X \tag{7}$$

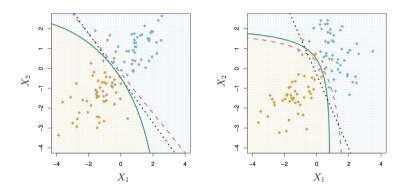
- ► Classification: in the probability of space
- ▶ Discrimination: in the space of X
- \triangleright $\beta_1 + \beta_2 X$ is the discrimination function for logit (it is lineal)

Linear Discriminant Analysis

- ► LDA?
- ▶ One predictor with $\sigma_0 = \sigma_1$ (equal variance)

Quadratic Discriminant Analysis

▶ QDA assumes diferent variances for the components



Naive Bayes

$$Pr(Y=1|X) = \frac{f(X|Y=1)\pi(Y=1)}{f(X|Y=1)\pi(Y=1) + f(X|Y=0)(1-\pi(Y=1))}$$
(8)

- \blacktriangleright $\pi(Y=1)$
- ightharpoonup f(X|Y=1)



Naive Bayes

► NB assumes independence

$$f(X|Y=1) = f(x_1|Y=1) \times \dots \times f(x_k|Y=1)$$
 (9)

Example: Default

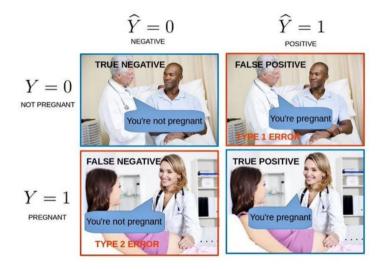


photo from https://www.dailydot.com/parsec/batman-1966-labels-tumblr-twitter-vine/

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Misclassification Rates



Misclassification Rates

$$\begin{array}{cccc}
 & \hat{y}_i \\
 & 0 & 1 \\
 & 0 & \text{TN FP} \\
 & y_i & 1 & \text{FN TP}
\end{array}$$

▶ We have several types of error associated with this that we can use as a measure of performance

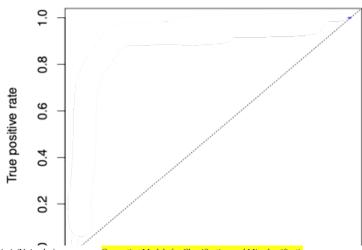
$$\begin{array}{cccc}
 & \hat{y}_i \\
 & 0 & 1 \\
 & 0 & \text{TN FP} \\
 & y_i & 1 & \text{FN TP}
\end{array}$$

- ► A classification rule, or cutoff, is the probability *p* at which you predict
 - $ightharpoonup \hat{y}_i = 0 \text{ if } p_i < c$
 - $\hat{y}_i = 1 \text{ if } p_i > c$
- ▶ Bayes classifier c = 0.5
- ► Changing *c* changes predictions, changes FP and FN
- ▶ There is a trade-off: reducing one error increases the other

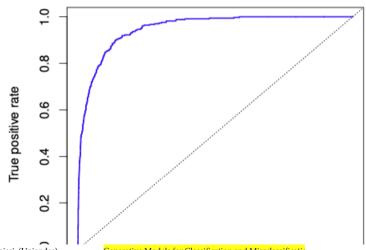


- ▶ ROC curve: Receiver operating characteristic curve
- ▶ ROC curve illustrates the trade-off of the classification rule
- ► Gives us the ability
 - Measure the predictive capacity of our model
 - Compare between models

ROC Curve







Example: Default

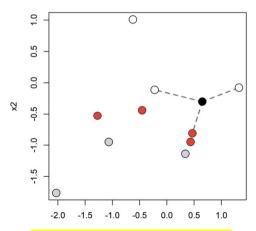


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- ▶ What happens when we have to predict multiple outcomes?
- ► K nearest neighbor (K-NN) algorithm predicts class \hat{y} for x by asking What is the most common class for observations around x?



- ► K nearest neighbor (K-NN) algorithm predicts class \hat{y} for x by asking What is the most common class for observations around x?
- ightharpoonup Algorithm: given an input vector x_f where you would like to predict the class label
 - ▶ Find the K nearest neighbors in the dataset of labeled observations, $\{x_i, y_i\}_{i=1}^n$, the most common distance is the Euclidean distance:

$$d(x_i, x_f) = \sqrt{\sum_{j=1}^{p} (x_{ij} - x_{fj})^2}$$
 (10)

► This yields a set of the *K* nearest observations with labels:

$$[x_{i1}, y_{i1}], \dots, [x_{iK}, y_{iK}]$$
 (11)

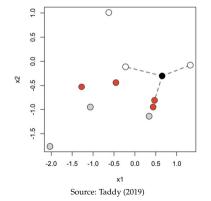
ightharpoonup The predicted class of x_f is the most common class in this set

$$\hat{y}_f = mode\{y_{i1}, \dots, y_{iK}\} \tag{12}$$

- ► There are some major problems with practical implications
 - ► Knn predictions are unstable as a function of *K*

$$K = 1 \implies \hat{p}(white) = 0$$

 $K = 2 \implies \hat{p}(white) = 1/2$
 $K = 3 \implies \hat{p}(white) = 2/3$
 $K = 4 \implies \hat{p}(white) = 1/2$



- ▶ There are some major problems with practical implications
 - ► Knn predictions are unstable as a function of *K*
 - ► This instability of prediction makes it hard to choose the optimal K and cross validation doesn't work well for KNN
 - ▶ Since prediction for each new *x* requires a computationally intensive counting, KNN is too expensive to be useful in most big data settings.
 - ► KNN is a good idea, but too crude to be useful in practice

- ► The MNLM can be thought of as simultaneously fitting binary logits for all comparisons among the alternatives.
- ► For example,
 - We have a categorical variable with the outcomes for Democrat, for Independent, and for Republican.
 - ► Assume that there is one independent variable measuring income in 1,000s.
 - We can examine the effect of income on party by fitting three binary logits,

► The MNLM can be thought of as simultaneously fitting binary logits for all comparisons among the alternatives.

$$\ln \frac{Pr(D|X)}{Pr(I/X)} = \beta_{0,D|I} + \beta_{1,D|I}Income$$

$$\ln \frac{Pr(R|X)}{Pr(I/X)} = \beta_{0,R|I} + \beta_{1,R|I}Income$$

$$\ln \frac{Pr(D|X)}{Pr(R/X)} = \beta_{0,D|R} + \beta_{1,D|R}Income$$
(13)

 \blacktriangleright where the subscripts to the $\beta's$ indicate which comparison is being made.

► These logits include redundant info

$$\ln \frac{Pr(D|X)}{Pr(I/X)} - \ln \frac{Pr(R|X)}{Pr(I/X)} = \ln \frac{Pr(D|X)}{Pr(R/X)}$$
(14)

which implies

$$\beta_{0,D|I} - \beta_{0,R|I} = \beta_{0,D|R} \tag{15}$$

$$\beta_{1,D|I} - \beta_{1,R|I} = \beta_{1,D|R} \tag{16}$$

▶ In general, with J alternatives, only J - 1 binary logits need to be fit (minimal set)

		Binary		
	(1)	(2)	(3)	
VARIABLES	dem_ind	rep_ind	dem_rep	
income	-0.00249	0.0157***	-0.0184***	
	(0.00355)	(0.00374)	(0.00230)	
Constant	1.605***	0.659***	0.953***	
	(0.149)	(0.162)	(0.105)	
Observations	844	689	1,231	

▶ Fitting the MNLM by fitting a series of binary logits is not optimal

	(1)	Binary (2)	(3)	
VARIABLES	dem_ind	rep_ind	dem_rep	
income	-0.00249 (0.00355)	0.0157*** (0.00374)	-0.0184*** (0.00230)	
Constant	1.605*** (0.149)	0.659*** (0.162)	0.953*** (0.105)	
Observations	844	689	1,231	

- Fitting the MNLM by fitting a series of binary logits is not optimal
 - ▶ Binary logit is based on a different sample.
 - ▶ It ignores the restricctions that are implicit in the definition of the MNLM

The multinomial logit model: formal statement

► Formally

$$\ln \Omega_{m|b}(X) = \ln \frac{Pr(y=m|X)}{Pr(y=b|X)} = X\beta_{m|b} \text{ for } m=1,\dots,J$$
(17)

- ▶ were *b* is the base outcome (reference category)
- ▶ These J equations can be solved to compute the probabilities for each outcome

$$Pr(y = m|X) = \frac{exp(X\beta_{m|b})}{\sum_{i=1}^{J} exp(X\beta_{i|b})}$$
(18)

	Binary			Multinomial Logit		
	(1)	(2)	(3)			
VARIABLES	dem_ind	rep_ind	dem_rep	Democrat	Independent	Republican
income	-0.00249	0.0157***	-0.0184***	-0.00272		0.0152***
	(0.00355)	(0.00374)	(0.00230)	(0.00372)		(0.00366)
Constant	1.605***	0.659***	0.953***	1.613***		0.678***
	(0.149)	(0.162)	(0.105)	(0.153)		(0.160)
Observations	844	689	1,231	1,382	1,382	1,382