

Assignment 3

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Given the following regression model:

$$RPmsoft_t = \beta_1 + \beta_2 RP sandp_t + \beta_3 Dprod_t + \beta_4 Dinflation_t + \beta_5 Dterm_t + \beta_6 m1_t + \epsilon_t$$

where

- $RPmsoft_t$ is the excess return of the Microsoft stock,
- $RP sandp_t$ is the risk premium of the S&P 500 index,
- $Dprod_t$ is the change in production,
- $Dinflation_t$ is the change in inflation,
- $Dterm_t$ is the change in term structure,
- $m1_t$ is the money supply growth,
- ϵ_t is the error term.

2.a

Using the data microsoft.csv, the resulting regression model is as follows:

$$\widehat{RPmsoft}_t = -0.9291 + 1.3232 RP sandp_t - 1.5216 Dprod_t + 0.4716 Dinflation_t + 4.1588 Dterm_t + 5.4352 m1_t$$

. reg rpmsoft rpsandp dprod dinflation dterm m1						
Source	SS	df	MS	Number of obs	=	324
Model	13628.1699	5	2725.63398	F(5, 318)	=	17.26
Residual	50211.9204	318	157.899121	Prob > F	=	0.0000
				R-squared	=	0.2135
				Adj R-squared	=	0.2011
Total	63840.0903	323	197.647338	Root MSE	=	12.566

rpmsoft	Coefficient	Std. err.	t	P> t	[95% conf. interval]	
rpsandp	1.323168	.1524728	8.68	0.000	1.023185	1.623151
dprod	-1.521569	1.283074	-1.19	0.237	-4.045955	1.002817
dinflation	.4715841	2.351165	0.20	0.841	-4.15422	5.097388
dterm	4.158737	2.487168	1.67	0.095	-.7346462	9.05212
m1	5.435169	2.869448	1.89	0.059	-.2103327	11.08067
_cons	-.9290695	.7598421	-1.22	0.222	-2.424022	.5658834

2.b

To test the *January effect* which is that on average, every else equal, the returns (or excess returns) are larger in the month of January than the rest of the months, we set-up the following hypothesis test:

$$H_0 : \beta_6 = 0 \quad (\text{No January effect}) \quad \text{vs} \quad H_a : \beta_6 > 0 \quad (\text{January effect exists and is positive})$$

where β_6 is the estimated coefficient for the regressor $m1$ which is a 1 for January and 0 otherwise.

From there OLS regression results, we have $t - \text{value} = 1.89$. At $\alpha = 1\%$ and degrees of freedom $n - K = 324 - 6 = 298$, the $t - \text{critical}_{0.01}(298) = 2.339$. Because $t - \text{value} < t - \text{critical}$, we fail to reject the null hypothesis. That is, the data does not provide evidence to reject the claim that there is no January effect in the excess returns of Microsoft stock.

2.c

The starting point for use of the $t - \text{test}$ statistic is the (conditional) sampling distribution of the $\hat{\beta}_k$ which is derived from the classical assumptions plus normality. Thus, the assumptions other than normality are:

- (A1) Linearity: The regression model has been correctly specified such that $y = X\beta + \epsilon$
- (A2) Strict exogeneity: The error term has an expected value of zero given any values of the regressors in all time periods, i.e. $E(\epsilon_i|X) = 0$ for all i .
- (A3) Homoskedasticity: The variance of the error term is constant across all levels of the regressors, i.e. $\text{Var}(\epsilon_i|X) = \sigma^2$ for all i .
- (A4) Disturbances are uncorrelated: The error terms are uncorrelated across observations, i.e. $\text{Cov}(\epsilon_i, \epsilon_j|X) = 0$ for all $i \neq j$.

In addition, to derive the distribution of the $t - \text{test}$ statistic, we used

$$\frac{\hat{\beta}_k - r}{\sqrt{\sigma^2(X'X)^{-1}_{kk}}} \underset{\text{under } H_0}{\sim} N(0, 1)$$

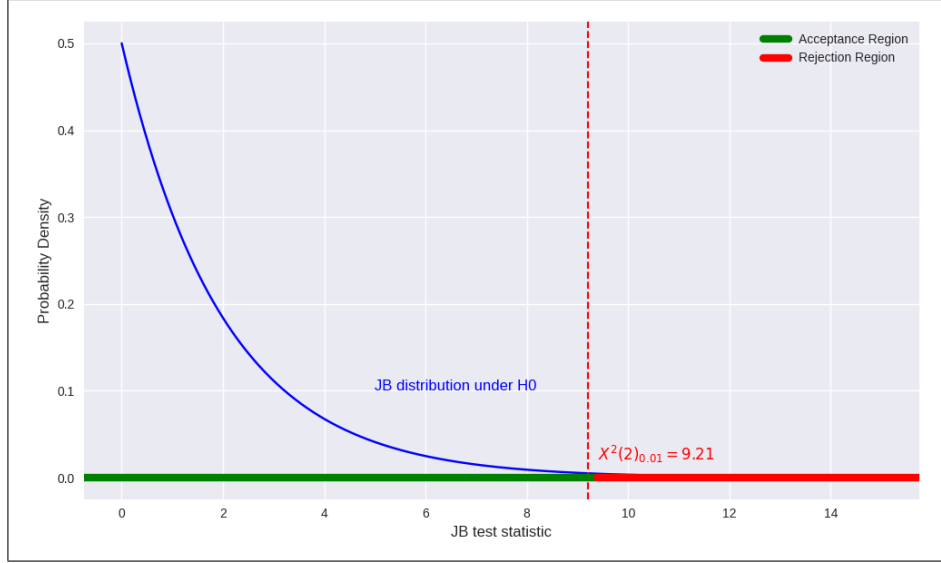
where r is the value of β_k under the null hypothesis. Thus, the distribution of the $t - \text{test}$ statistic is derived under the assumption that H_0 is true.

2.d

The *Jarque-Bera* (JB) test can be used to check for normality of the disturbances. The JB-test statistic is given by:

$$JB = \frac{n}{6} \left(sk^2 + \frac{(kur - 3)^2}{4} \right) \underset{\text{under } H_0}{\overset{a}{\sim}} \chi^2(2)$$

where sk is the sample coefficient of the skewness of the variable, kur is its sample coefficient of kurtosis, and n is the sample size. Using significance of level of $\alpha = 0.01$ and the critical value $\chi^2_{0.01}(2) = 9.21$. The acceptance and rejection regions are in the illustration below:



A normal distribution has skewness of 0 and kurtosis of 3. Thus, a deviation from these values can indicate a departure from normality. The corresponding hypothesis test is set-up as follows:

$$H_0 : SK = 0 \text{ and } KUR = 3 \quad vs \quad H_A : \text{not } H_0$$

where SK and KUR are the population coefficients of skewness and kurtosis for the disturbances, respectively.

Under normality, the mean and variance of skewness is 0 and $\frac{6}{n}$ and 3 and $\frac{24}{n}$ for kurtosis, respectively. Thus, we can see the $JB - test$ statistic is composed of mahalanobis distances for SK and KUR . If the disturbances are normally distributed, then the $JB - test$ statistic should be small, otherwise it will have a large positive value. Therefore, the acceptance region is from 0 up to the critical value determined by the significance level.

2.e

If the assumptions regarding the dgp for consistency of OLS estimator are met, then $\hat{\beta} \xrightarrow{d} \beta$. Note that $\hat{\epsilon} = y - \hat{y} = X\beta + \epsilon - X\hat{\beta} = X(\beta - \hat{\beta}) + \epsilon$, by *Continuous Mapping Theorem*, we have $\hat{\epsilon} \xrightarrow{d} X(\beta - \beta) + \epsilon = \epsilon$. Thus, $\hat{\epsilon} \xrightarrow{d} \epsilon$.

2.f

Using Stata, the $JB - test$ value is 1809. At $\alpha = 0.01$ and degrees of freedom = 2, the critical value is $\chi^2_{0.01}(2) = 9.21$. Since $1809 > 9.21$, we reject the null hypothesis that the disturbances are normally distributed. Thus, there is evidence to suggest that the disturbances are not normally distributed.

2.g

Testing again the *January effect* using asymptotic t-test, we have critical value = $z_{0.01} = 2.326$. Since the $t - value = 1.89 < 2.326$, we fail to reject the null hypothesis. Thus, the conclusion remains the same as in part (b) that there is no evidence to suggest the existence of January effect in the excess returns of Microsoft stock.

2.h

Yes. Because the normality assumption for the exact $t - test$ is not satisfied

2.i

The $t - test$ statistic distribution has fatter tails than the standard normal distribution or the z statistic. This means that at the tails, using the same significance level, the value of the $t - test$ statistic will be smaller than the z statistic.