

Example: The production manager of a bulb manufacturing company wishes to study the effect of new manufacturing process on the lifetimes of • bulbs produced through it.

Here the population under study is

\mathcal{P} : collection of lifetimes of all bulbs produced ~~by~~ using new process.

In most practical situation \mathcal{P} is generally large.

Probability theory: A mathematical tool for modeling uncertainty. (e.g. to describe the law according to which values of the life time a bulb vary across \mathcal{P}).

The only way to collect information about any random phenomenon is to perform experiment. As an example selecting a set of bulbs manufactured by the new process and putting them on test for measuring their lifetimes. Each experiment terminates in an outcome which can not be predicted in advance prior to the performance of experiment.

- Defⁿ (Random experiment): A random (or statistical) experiment is an experiment in which
- (a) All outcomes of the experiment are known in advance
 - (b) Outcome of a particular performance of the experiment can not be predicted in advance.
 - (c) the experiment can be repeated under identical conditions.

Sample space: The collection of all possible outcomes of a random experiment is called its sample space. Sample space is denoted by Ω .

Ex (I) Tossing a coin : $\Omega = \{H, T\}$

(II) Throwing a die $\Omega = \{1, 2, 3, 4, 5, 6\}$

(III) Birth of child, $\Omega = \{M, F\}$

If we consider weight $\Omega = (0, 7)$

(IV) Age at death of a person : $\Omega = (0, 120)$

(V) Life of a bulb : $\Omega = (0, \infty)$

(VI) Throwing two dice

$$\begin{aligned}\Omega &= \{ (1,1), (1,2), \dots, (1,6), (2,1), (2,2), \dots, (2,6) \\ &\quad \dots, (6,1), (6,2), \dots, (6,6) \} \\ &= \{ (i,j), i,j \in \{1,2,3,4,5,6\} \}\end{aligned}$$

Event: An event is any subset of the sample space. If the outcome of a random experiment is a member of the set $E \subseteq \Omega$ we say that event E has occurred.

Impossible event: ϕ , Sure event: Ω .

Let A and B are two event

- I) $A \cup B \rightarrow$ occurrence of at least one of A & B
- II) $\bigcup_{i=1}^n A_i \rightarrow$ occurrence of at least one $A_i, i=1, 2, \dots, n$
- III) $A \cap B \rightarrow$ simultaneous occurrence of A & B .
- IV) $\bigcap_{i=1}^n A_i \rightarrow$ simultaneous occurrence of $A_i, i=1, 2, \dots, n$.

Exhaustive events: If $\bigcup_{i=1}^n A_i = \Omega$ we call A_1, A_2, \dots, A_n to be exhaustive events

If $A \cap B = \phi$ then A & B are called mutually exclusive events i.e. happening or occurrence of one of them excludes the possibility of occurrence of other.

Let A_1, A_2, \dots are event then

$A_i \cap A_j = \phi \quad i \neq j$, then we say A_1, A_2, \dots are pair wise disjoint or mutually exclusive.

$A^c \rightarrow$ not happening of A

$A-B \rightarrow$ happening of A not $B = A \cap B^c$

~~In general~~ Generally we are interested in specific subsets of Ω which we will treat as event. So the event space (events under consideration) \mathcal{G} is a subset of power set of Ω .

So the event space is $\mathcal{G} \subseteq \mathcal{P}(\Omega)$, Here $\mathcal{P}(\Omega)$ is the power set of Ω .

The choice of \mathcal{G} is an important one

(I) If Ω contains at most a countable number of points we can always take \mathcal{G} to be the $\mathcal{P}(\Omega)$. (This is certainly a σ -field). In this case each point set is a member of \mathcal{G} and is the fundamental object of interest. Every subset of Ω is an event.

(II) If Ω has uncountably many points the class of all subsets of Ω is still a σ -field but it is much too large a class of sets to be of interest. If $\Omega = \mathbb{R}$ or any interval then Ω is uncountable. In this case we would like to consider all one point subsets of Ω all intervals (closed, open, or semiclosed) to be events. We consider the Borel σ -field \mathcal{B}_1 generated by

the class of all semiclosed intervals $(a, b]$, which is a σ -field in \mathbb{R} .

We say that the σ event space $\mathcal{G} \subseteq \mathcal{P}(\Omega)$ contains all subsets of Ω actually encountered in ordinary analysis and probability. It is large enough for all ~~pract~~ practical purposes.

The algebra of set theory is applicable in prob. theory. Probability is a measure of uncertainty. We are interested in quantifying uncertainty associated with various outcomes of a random experiment by assigning probability to these outcomes.

Here we will not discuss how probabilities are assigned (which is a part of prob. modeling) rather we will discuss properties of a probability measure.

Defⁿ (Probability function or Probability measure)

A probability funⁿ (or prob measure) is a real valued set function, defined on the event space \mathcal{G} satisfying the following axioms

(a) $P(A) \geq 0 \quad \forall A \in \mathcal{G}$

(b) $P(\Omega) = 1$

(c) If $\{A_n\}$ is a sequence of mutually exclusive (disjoint) sets in \mathcal{G} i.e. $A_i \cap A_j = \phi$ $i \neq j$

then

$$P\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} P(A_i) \quad (\text{countable additivity}).$$

We call $P(A)$ the probability of event A .

The triplet (Ω, \mathcal{G}, P) is called probability space.

Properties of Probability Measure:

$$(P_1) \quad P(\phi) = 0$$

Proof: Let $A_1 = \Omega$ and $\emptyset A_i = \phi, i=2, 3, \dots$

Then $P(A_1) = 1$. Also we have $A_1 = \bigcup_{i=1}^{\infty} A_i, A_i \cap A_j = \phi$

$i \neq j$. Therefore

$$1 = P(A_1) = P\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} P(A_i) = 1 + \sum_{i=2}^{\infty} P(A_i)$$

$$\Rightarrow \sum_{i=2}^{\infty} P(\phi) = 0 \Rightarrow P(\phi) = 0$$

(P2) Let $A_1, A_2, \dots, A_n \in \mathcal{G}, \exists A_i \cap A_j = \phi, i \neq j \Rightarrow$

$$P\left(\bigcup_{i=1}^n A_i\right) = \sum_{i=1}^n P(A_i)$$

Proof: Take $A_i = \phi, i = n+1, n+2, \dots$

$$\text{Then } P\left(\bigcup_{i=1}^n A_i\right) = P\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} P(A_i) = \sum_{i=1}^n P(A_i)$$

(P3) $\forall A \in \mathcal{S}, 0 \leq P(A) \leq 1$ and $P(A^c) = 1 - P(A)$

Proof:

$$1 = P(\Omega) = P(A \cup A^c) = P(A) + P(A^c)$$

$$\Rightarrow P(A^c) = 1 - P(A)$$

Also $P(A) \leq 1$

So $0 \leq P(A) \leq 1$ & $P(A^c) = 1 - P(A)$

(P4) $A_1, A_2 \in \mathcal{S}$ and $A_1 \subseteq A_2 \Rightarrow P(A_2 - A_1) = P(A_2) - P(A_1)$
and $P(A_1) \leq P(A_2)$.

Proof:

$$A_2 = A_1 \cup (A_2 - A_1)$$

$$A_1 \cap (A_2 - A_1) = \emptyset$$

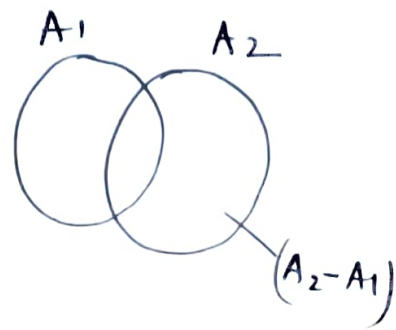
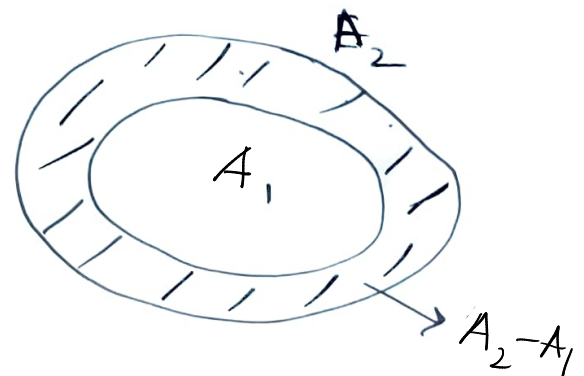
So $P(A_2) = P(A_1) + P(A_2 - A_1)$

$$\Rightarrow P(A_2 - A_1) = P(A_2) - P(A_1)$$

$\therefore P(A_2 - A_1) \geq 0 \Rightarrow P(A_2) \geq P(A_1)$ (Monotonicity)

(P5) $A_1, A_2 \in \mathcal{S}, P(A_1 \cup A_2) = P(A_1) + P(A_2) - P(A_1 \cap A_2)$

$$\begin{aligned} P(A_1 \cup A_2) &= P(A_1 \cup (A_2 - A_1)) \\ &= P(A_1) + P(A_2 - A_1) \quad \text{--- (i)} \end{aligned}$$



Now we have

$$(A_1 \cap A_2) \cap (A_2 - A_1) = \phi \quad \text{and} \quad A_2 = (A_1 \cap A_2) \cup (A_2 - A_1)$$

\Rightarrow

$$P(A_2) = P(A_1 \cap A_2) + P(A_2 - A_1)$$

$$\Rightarrow P(A_2 - A_1) = P(A_2) - P(A_1 \cap A_2) \quad \text{--- (11)}$$

Using (11) from (1) we get

$$P(A_1 \cup A_2) = P(A_1) + P(A_2) - P(A_1 \cap A_2)$$

Theorem

Let $A_1, A_2, A_3, \dots, A_n \in \mathcal{G}$, $n \geq 2$. Define

$$p_{1,n} = P(A_1) + P(A_2) + \dots + P(A_n) = \sum_{i=1}^n P(A_i)$$

$$p_{2,n} = \sum_{1 \leq i < j \leq n} P(A_i \cap A_j) \quad \left(\begin{array}{l} \text{sum of probabilities of all} \\ \text{possible intersections involving 2} \\ \text{events out of } n \text{ events } A_1, \dots, A_n \end{array} \right)$$

$$p_{i,n} = \sum_{1 \leq j_1 < j_2 < \dots < j_i \leq n} P(A_{j_1} \cap A_{j_2} \cap \dots \cap A_{j_i})$$

$\left(\begin{array}{l} \text{sum of probabilities of all possible intersections involving} \\ i \text{ events out of } n \text{ events } A_1, A_2, \dots, A_n \end{array} \right)$

Then

$$P\left(\bigcup_{i=1}^n A_i\right) = p_{1,n} - p_{2,n} + p_{3,n} - p_{4,n} + \dots + (-1)^{n-1} p_{n,n}$$

Proof: Do yourself. Use induction.

Remark (i) Let $A_1, A_2, A_3 \in \mathcal{G}$, then

$$\begin{aligned} P(A_1 \cup A_2 \cup A_3) &= P_{1,3} - P_{2,3} + P_{3,3} \\ &= P(A_1) + P(A_2) + P(A_3) - P(A_1 \cap A_2) - P(A_1 \cap A_3) - \\ &\quad P(A_2 \cap A_3) + P(A_1 \cap A_2 \cap A_3) \end{aligned}$$

(ii) We have $P(A_1 \cup A_2) \leq 1$

$$\Rightarrow P(A_1) + P(A_2) - P(A_1 \cap A_2) \leq 1$$

$$\Rightarrow P(A_1 \cap A_2) \geq P(A_1) + P(A_2) - 1$$

This inequality is known as Bonferroni's inequality.

Theorem: Let (Ω, \mathcal{G}, P) be a probability space and let $A_1, A_2, \dots, A_n \in \mathcal{G}$ ($n \in \mathbb{N}, n \geq 2$). Then

(I) Boole's inequality

$$P_{1,n} - P_{2,n} \leq P\left(\bigcup_{i=1}^n A_i\right) \leq P_{1,n}$$

(II) $P\left(\bigcap_{i=1}^n A_i\right) \geq P_{1,n} - (n-1)$. (Bonferroni's Inequality)

Example: Consider a random experiment throwing two dice one is red and other one white

Sample space: $\Omega = \{(i, j) : i = 1, 2, \dots, 6, j = 1, 2, \dots, 6\}$

For $(i, j) \in \Omega$

i : number of spots up on the red die

j : number of spots up on the white die.

Event space $\mathcal{S} = \text{power set of } \Omega = 2^\Omega$

For $E \in \mathcal{S}$ define $Q: \mathcal{S} \rightarrow \mathbb{R}$ as

$$Q(E) = \frac{|E|}{36}, \text{ where } |E| = \# \text{ of elements in } E$$

Then (i) $Q(\Omega) = \frac{|\Omega|}{36} = 1$

(ii) $Q(E) = \frac{|E|}{36} \geq 0 \quad \forall E \in \mathcal{S}.$

(iii) For mutually disjoint events E_1, E_2, \dots

$$Q\left(\bigcup_{i=1}^{\infty} E_i\right) = \frac{\left|\bigcup_{i=1}^{\infty} E_i\right|}{36} = \frac{\sum_{i=1}^{\infty} |E_i|}{36} = \sum_{i=1}^{\infty} Q(E_i)$$

Thus (Ω, \mathcal{S}, Q) is a prob. space.