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Ch. 11 More Topics in Linear Unobserved Effects Models

Example 11.2 lung cancer and sales of cigarette

$$y_{it} = \mathbf{z}_{it} \gamma + \delta w_{it} + c_i + u_{it}, \quad w_{it} = \mathbf{z}_{it} \xi + \rho_1 y_{i,t-1} + \psi c_i + r_{it}$$

 \mathbf{z}_n : strictly exogenous, \mathbf{w}_n : sequentially exogenous

 w_n : number of lung cancer patients per 1000, v_n : cigarette sales

If $\rho_1 \neq 0$, strict exogeneity is not satisfied $(E(w_{i,i+1}u_{it}) = \rho_1 E(u_{it}^2) \neq 0)$

Even with sequential moment restrictions, the FE or FD estimator of the model with feedback is not consistent without strict exogeneity

⇒ Estimate FD transformation model with strictly exogenous instruments

11.1 Unobserved Effects Models without the Strict Exogeneity Assumption Models in Ch. 10 assume strict exogeneity $E(u_{it} | \mathbf{x}_{it}, \dots, \mathbf{x}_{it}, c_i) = 0$. However, this is not satisfied if y affects future x (ex. y: output, x: capital)

$$y_{it} = \mathbf{x}_{it}\mathbf{\beta} + c_i + u_{it}, \quad t = 1, \dots, T$$

Sequential Moment Restrictions: \mathbf{x}_{tt} are sequentially exogenous conditional on the unobserved effect if $E(u_{tt} \mid \mathbf{x}_{tt}, \dots, \mathbf{x}_{tt}, c_t) = 0$, $t = 1, \dots, T$

 u_{ii} can correlate to future x_{ii} , but uncorrelated to past x_{ii}

Ex. u_n : technology shock. Future capital x_{it+1} depends on past GDP y_n and thus past technology shocks u_n through investment. But u_n does not depend on past capital

How to estimate the model with Sequentially exogenous regressors (Pooled 2SLS)

- 1. FD transformation: $\Delta y_{it} = \Delta \mathbf{x}_{it} \boldsymbol{\beta} + \Delta u_{it}, \quad t = 2, \dots, T \quad \Delta y_{it} = y_{it} y_{i,t-1}, \Delta \mathbf{x}_{it} = \mathbf{x}_{it} \mathbf{x}_{i,t-1}, \Delta u_{it} = u_{it} u_{i,t-1}$
- 2. Since $E(\mathbf{x}_{is}'\Delta u_{it}) = 0$ for $s = 1, \dots, t-1$, estimate the model by POLS using $\Delta \mathbf{x}_{i,t-1}$ (or variants of $\mathbf{x}_{i,t-1}^o = (\mathbf{x}_{i1}, \dots, \mathbf{x}_{i,t-1})$) as instruments for $\Delta \mathbf{x}_{it}$

Note that $E(\Delta x_{i+1}) \neq \emptyset$. For the IV to be valid, this correlation must be high enough

Example 11.3

11.1.2 Models with Strictly and Sequentially Exogenous Explanatory Variables

$$y_{it} = \mathbf{z}_{it} \boldsymbol{\gamma} + \mathbf{w}_{it} \boldsymbol{\delta} + c_i + u_{it}, \quad t = 1, \dots, T$$

 z_n : strictly exogenous, w_n : sequentially exogenous

How to estimate the model with sequentially exogenous (Pooled 2SLS)

- 1. FD transformation: $\Delta y_{it} = \Delta \mathbf{z}_{it} \boldsymbol{\gamma} + \Delta \mathbf{w}_{it} \boldsymbol{\delta} + \Delta u_{it}, \quad t = 2, \dots, T$
- 2. P2SLS estimation using for example $(\Delta \mathbf{z}_{i}, w_{i,t-1}, w_{i,t-2})$ as instruments for $(\Delta \mathbf{z}_{i}, \Delta \mathbf{w}_{i})$ (Note that any of $(\mathbf{z}_{i}, \mathbf{w}_{n-1}, \mathbf{L}, \mathbf{w}_{n})$ can be instruments)

Estimation Procedure

1. Remove unobservable effect by either FD or FE

$$\Delta \log(wage_{it}) = \Delta \mathbf{z}_{it} \mathbf{\gamma} + \delta_1 \Delta cigs_{it} + \Delta u_{it}$$

2. 2SLS using strictly exogenous \mathbf{z}_i or (under the assumption that u is not serially correlated) 2 lagged w or y, or other exogenous variable as IV

P2SLS or GMM estimate the model using cigarette price or cigarette tax (in the U.S., there is a variation across shops /states) as instruments for $\Delta cigs_{it}$

It is possible to use FE transformation to remove unobserved effect, but with FEtransformation, we cannot use lagged w as IV (FD is better)

11.1.3 Models with Contemporaneous Correlation between Some Explanatory Variables and the Idiosyncratic Error

$$y_{it} = \mathbf{z}_{it} \mathbf{\gamma} + \mathbf{w}_{it} \mathbf{\delta} + c_i + u_{it}, \quad t = 1, \dots, T$$

 z_n : strictly exogenous, w_n : contemporaneously correlated with u_n (due to **omitted** variables, measurement errors, simultaneity)

Example 11.5 Effect of smoking on wage

$$\log(wage_{ij}) = \mathbf{z}_{ij}\gamma + \delta_1 cigs_{ij} + c_i + u_{ij}$$
, $cigs_{ij}$: smokes per day

If cigarette is a normal good, high income (wage) leads to high consumption of cigarette (δ_i is upward biased)

⇒ Correlation between regressors and u (OLS is inconsistent, and IVs neede)

With measurement error (attenuation bias)

$$y_{tt} = \beta x_{tt}^{*} + c_{t} + u_{tt}, \quad E(u_{tt} | \mathbf{x}_{t}^{*}, \mathbf{x}_{t}, c_{t}) = 0, \quad t = 1, \dots, T,$$

$$x_{tt}^{*} \text{:true value}, \quad x_{tt} \text{:observed} \quad x_{tt} = x_{tt}^{*} + r_{tt}$$

$$p \lim_{N \to \infty} \hat{\beta}_{POLS} = \beta + \frac{Cov(x_{tt}, c_{t}) - \beta \sigma_{r}^{2}}{Var(x_{tt})}, \quad \sigma_{r}^{2} = Var(r_{tt}) = Cov(x_{tt}, r_{tt})$$

$$p \lim_{N \to \infty} \hat{\beta}_{FD} = \beta + \left(1 - \frac{\sigma_r^2(1 - \rho_r)}{\sigma_{x^*}^2(1 - \rho_{x^*}) + \sigma_r^2(1 - \rho_r)}\right), \quad \rho_{x^*} = Corr(x^*_{it}, x^*_{i,t-1}), \\ \rho_r = Corr(r_{it}, r_{i,t-1})$$

=> do not know which generates smaller bias, POLS or FD

FD + IV for the consistent estimation

11.1.4 Summary of Models without Strictly Exogenous Explanatory Variables

In general

- 1) FD or FE transformation to remove unobserved effect
- 2) IV estimation

11.2.2 General Models with Individual-Specific Slopes

$$y_{it} = \mathbf{z}_{it}\mathbf{a}_i + \mathbf{x}_{it}\boldsymbol{\beta} + u_{it}, \quad t = 1,\dots,T, \quad \text{a_i: Jx1 (individual specific slope)}$$

Assumption FE.1' (strict exogeneity): $E(u_{ii} | \mathbf{x}_i, \mathbf{z}_i, \mathbf{a}_i) = 0$

Stack across time and write with the vector $\mathbf{y}_i = \mathbf{Z}_i \mathbf{a}_i + \mathbf{X}_i \mathbf{\beta} + \mathbf{u}_i$,

$$\begin{aligned} & \mathbf{Z}_i = (\mathbf{z}_{i1}, \cdots, \mathbf{z}_{iT})', \quad \mathbf{X}_i = (\mathbf{x}_{i1}, \cdots, \mathbf{x}_{iT})' \text{ here } \mathbf{M}_i = \mathbf{I}_T - \mathbf{Z}_i (\mathbf{Z}_i' \mathbf{Z}_i)^{-1} \mathbf{Z}_i' \text{ and } \ddot{\mathbf{X}}_i = \mathbf{M}_i \mathbf{X}_i, \quad \ddot{\mathbf{y}}_i = \mathbf{M}_i \mathbf{y}_i, \\ & \ddot{\mathbf{u}}_i = \mathbf{M}_i \mathbf{u}_i, \text{ then } \ddot{\mathbf{y}}_i = \ddot{\mathbf{X}}_i \boldsymbol{\beta} + \ddot{\mathbf{u}}_i \end{aligned}$$

<u>Assumption FE.2'</u> (rank condition): $rankE(\ddot{\mathbf{X}}_i'\ddot{\mathbf{X}}_i) = K$, where $\ddot{\mathbf{X}}_i = \mathbf{M}_i\mathbf{X}_i$, Under AFE1' and AFE2', SOLS is consistent.

11.2 Models with Individual-Specific Slopes

11.2.1 A Random Trend Model

$$y_{it} = c_i + g_i t + \mathbf{x}_{it} \mathbf{\beta} + u_{it}, \quad t = 1, \dots, T, \quad E(u_{it} \mid \mathbf{x}_i, c_i, g_i) = 0$$
: strict exogeneity

Estimation Method If taking FD, the model is just one with unobserved effect model! $\Delta y_n = g_n + \Delta x_n \beta + \Delta u_n$

FD or FE estimation (need T > 2)

Same estimation method can be used if the coefficients on explanatory variables varies across time

Assumption FE.3' (homoskedasticity): $E(\mathbf{u}_i \mathbf{u}_i' | \mathbf{z}_i, \mathbf{x}_i, \mathbf{a}_i) = \sigma_u^2 \mathbf{I}_T$

Asymptotic variance
$$A \hat{\text{var}}(\hat{\boldsymbol{\beta}}_{FE}) = \hat{\sigma}_u^2 \left(\sum_{i=1}^N \ddot{\mathbf{X}}_i ' \ddot{\mathbf{X}}_i \right)^{-1}$$

$$\hat{\sigma}_u^2 = [N(T-J) - K]^{-1} \sum_{i=1}^N \sum_{i=1}^T \hat{u}_{ii}^2 = SSR/[N(T-J) - K], \quad \hat{u}_{it} = \ddot{y}_{it} - \ddot{\mathbf{x}}_{it} \hat{\boldsymbol{\beta}}_{FE}$$

F-test is valid. If AFE3' is not satisfied, use robust variance matrix

How to obtain $a = E(a_i)$

$$\mathbf{a} = E[(\mathbf{Z}_i'\mathbf{Z}_i)^{-1}\mathbf{Z}_i'(\mathbf{y}_i - \mathbf{X}_i\boldsymbol{\beta})], \quad \hat{\mathbf{a}} = N^{-1}\sum_{i=1}^{N}(\mathbf{Z}_i'\mathbf{Z}_i)^{-1}\mathbf{Z}_i'(\mathbf{y}_i - \mathbf{X}_i\hat{\boldsymbol{\beta}}_{FE})$$

Expanding this equation and using asymptotic normality of beta, we can calculate asymptotic variance of \hat{a}

11.3 GMM	Approaches to	Linear	Unobserved	Effects	Models
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11.3.1 Equivalence between 3SLS and Standard Panel Data Estimation

11.3.2 Chamberlain's Approach to Unobserved Effects Models

11.4 Hausman and Taylor-Type Models

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- I. Example 11.3
- II. Example 11.6
- **III.** 11.15

11.5 Applying Panel Data Methods to Matched Pairs and Cluster Samples

Matched pairs sample

Ex. sibling
$$y_{i1} = \mathbf{x}_{i1}\mathbf{\beta} + f_i + u_{i1}$$
, $y_{i2} = \mathbf{x}_{i2}\mathbf{\beta} + f_i + u_{i2}$
=> estimate by **RE**, **FE**

Cluster sample

Ex. Cluster by school
$$y_{is} = \mathbf{x}_{is}\mathbf{\beta} + c_i + u_{is}$$

Even if the cluster size is not same, demean and FE estimation

Peer effects

Ex. School average traits
$$y_{is} = \mathbf{x}_{is} \boldsymbol{\beta} + \overline{\mathbf{w}}_{i(s)} \boldsymbol{\delta} + u_{is}$$

$$\overline{\mathbf{w}}_{i(s)}$$
: cluster average of \mathbf{x}_{is} except for s

Under strict exogeneity, POLS is consistent