

## STAT 610: Discussion 10

### 1 Summary

- Union-Intersection and Intersection Tests (UIT/IUT):

- UIT: Suppose we want to test  $H_0 : \theta \in \cap_{\gamma \in \Gamma} \Theta_\gamma$  v.s.  $H_1 : \theta \in \cup_{\gamma \in \Gamma} \Theta_\gamma^c$ . Suppose  $R_\gamma$  is the rejection region of test  $H_0 : \theta \in \Theta_\gamma$  v.s.  $H_1 : \theta \in \Theta_\gamma^c$ . Then, UIT is  $R = \cup_{\gamma \in \Gamma} R_\gamma$ . If  $R_\gamma$  is of the form  $\{x : T_\gamma(x) > c\}$ , then

$$R = \cup_{\gamma \in \Gamma} \{x : T_\gamma(x) > c\} = \{x : \sup_{\gamma \in \Gamma} T_\gamma(x) > c\}.$$

- IUT: Suppose we want to test  $H_0 : \theta \in \cup_{\gamma \in \Gamma} \Theta_\gamma$  v.s.  $H_1 : \theta \in \cap_{\gamma \in \Gamma} \Theta_\gamma^c$ . If each  $H_{0\gamma}$  has the rejection region  $R_\gamma = \{x : T_\gamma(x) > c\}$ , then IUT has the rejection region

$$R = \cap_{\gamma \in \Gamma} \{x : T_\gamma(x) > c\} = \{x : \inf_{\gamma \in \Gamma} T_\gamma(x) > c\}.$$

- Level of UIT: If  $R_\gamma$  has level  $\alpha_\gamma$ , then the overall level of UIT is at most  $\sum_{\gamma \in \Gamma} \alpha_\gamma$ .
- Level of IUT: If  $R_\gamma$  has level  $\alpha_\gamma$ , then the overall level of IUT is at most  $\min_{\gamma \in \Gamma} \alpha_\gamma$ .
- Relationship between LRT and UIT: Refer to *Theorem 8.3.21* in the textbook.

- Confidence Interval:

- The *coverage probability* is defined as  $\mathbb{P}_\theta(\theta \in [L(\mathbf{X}), U(\mathbf{X})])$  for an interval  $[L(\mathbf{X}), U(\mathbf{X})]$ .
- The *confidence coefficient* is defined as  $\inf \mathbb{P}_\theta(\theta \in [L(\mathbf{X}), U(\mathbf{X})])$ , which is the infimum of the coverage probability.

- Pivotal quantities: A random variable  $Q(\mathbf{X}, \theta) = Q(X_1, \dots, X_n, \theta)$  is a *pivotal quantity* if the distribution of  $Q(\mathbf{X}, \theta)$  is independent of all parameters. That is, if  $\mathbf{X} \sim F(\mathbf{x} | \theta)$ , then  $Q(\mathbf{x}, \theta)$  has the same distribution for all values of  $\theta$ .

- e.g. If  $X_i \stackrel{i.i.d.}{\sim} \text{Exp}(\theta)$ , then  $T = \sum_{i=1}^n X_i \sim \text{Gamma}(n, \theta)$ . Hence  $Q(T, \theta) = 2T/\theta \sim \chi_{2n}^2$  is a pivotal quantity.
- After figuring out  $a, b$  such that  $\mathbb{P}(a \leq Q(\mathbf{X}, \theta) \leq b) \geq 1 - \alpha$ , then  $C(\mathbf{x}) = \{\theta : a \leq Q(\mathbf{x}, \theta) \leq b\}$  is a  $1 - \alpha$  confidence set for  $\theta$ .
- If  $Q(\mathbf{x}, \theta)$  is a monotone function of  $\theta$ , then  $C(\mathbf{x})$  will be an interval.

- Pivoting a cdf: Let  $T$  be a statistic with cdf  $F_T(t | \theta)$ . Let  $\alpha_1 + \alpha_2 = \alpha$  be constants.

- If  $F_T(t | \theta)$  is a decreasing function of  $\theta$  for each  $t$ , define  $\theta_L(t)$  and  $\theta_U(t)$  by

$$F_T(t | \theta_U(t)) = \alpha_1, \quad F_T(t | \theta_L(t)) = 1 - \alpha_2.$$

- If  $F_T(t | \theta)$  is an increasing function of  $\theta$  for each  $t$ , define  $\theta_L(t)$  and  $\theta_U(t)$  by

$$F_T(t | \theta_U(t)) = 1 - \alpha_2, \quad F_T(t | \theta_L(t)) = \alpha_1.$$

Then  $[\theta_L(T), \theta_U(T)]$  is a  $1 - \alpha$  confidence interval for  $\theta$ .

## 2 Questions

1. Consider testing  $H_0 : \theta \in \cup_{j=1}^k \Theta_j$ . For each  $j = 1, \dots, k$ , let  $p_j(\mathbf{x})$  denote a valid p-value for testing  $H_{0j} : \theta \in \Theta_j$ . Let  $p(\mathbf{x}) = \max_{1 \leq j \leq k} p_j(\mathbf{x})$ .
  - (a) Show that  $p(\mathbf{X})$  is a valid p-value for testing  $H_0$ .
  - (b) Show that the  $\alpha$  level test defined by  $p(\mathbf{X})$  is the same as an  $\alpha$  level IUT defined in terms of individual tests based on the  $p_j(\mathbf{x})$ s.
  
2. Find a  $1 - \alpha$  confidence interval for  $\theta$  using pivots, given  $X$  with pdf
  - (a)  $f(x | \theta) = 1, \theta - \frac{1}{2} < x < \theta + \frac{1}{2}$ .
  - (b)  $f(x | \theta) = 2x/\theta^2, 0 < x < \theta$ .

3. Let  $X_1, \dots, X_n$  be i.i.d uniform $(0, \theta)$ . Let  $Y$  be the largest order statistic. Prove that  $Y/\theta$  is a pivotal quantity and show that  $[y, y \cdot \alpha^{-1/n}]$  is the shortest  $1 - \alpha$  pivotal interval.

4. If  $X_1, \dots, X_n$  are i.i.d from a location pdf  $f(x - \theta)$ . Show that the confidence set

$$C(x_1, \dots, x_n) = \{\theta : \bar{x} - k_1 \leq \theta \leq \bar{x} + k_2\},$$

where  $k_1, k_2$  are constants, has constant coverage probability.

5. Let  $X_1, \dots, X_n$  be a random variable with pdf  $f_X(x) = \theta a^\theta x^{-(\theta+1)} I_{(a, \infty)}(x)$ , where  $\theta > 0$  and  $a > 0$ .

- (a) When  $\theta$  is known, derive a confidence interval for  $a$  with confidence coefficient  $1 - \alpha$  by pivoting the cdf of the smallest order statistic  $X_{(1)}$ .
- (b) When both  $a$  and  $\theta$  are unknown and  $n \geq 2$ , derive a confidence interval for  $\theta$  with confidence coefficient  $1 - \alpha$  by pivoting the cdf of  $T = \prod_{i=1}^n (X_i/X_{(1)})$ .

*Hint: You can use the fact that  $2\theta \log T \sim \chi_{2(n-1)}^2$  and then write the cdf of  $T$  in terms of the cdf of  $\chi_{2(n-1)}^2$ .*

- (c) When both  $a$  and  $\theta$  are unknown, construct a confidence set for  $(a, \theta)$  with confidence coefficient  $1 - \alpha$  using a pivotal quantity.

*Hint: Notice that  $X_{(1)}/a$  is free of  $a$ , and  $X_{(1)}^\theta$  is free of  $\theta$ .*