## STAT 610: Discussion 13

## 1 Summary

- Asymptotic tests:
  - **LRT statistic**: Let  $X_1, \ldots, X_n$  from  $f_{\theta}(x)$ . Under regularity conditions in §10.6.2, if  $\theta \in \Theta_0$ , then the distribution of the statistic  $-2 \log \lambda(\mathbf{X}) \sim \chi_r^2$  as  $n \to \infty$ , where r is the difference between the dimension of  $\theta$  and number of free parameters in  $\Theta_0$ .
  - Based on the above asymptotic result above, we can construct an asymptotic α-level test as  $R = {\mathbf{X} : -2 \log \lambda(\mathbf{X}) \ge \chi_{r,\alpha}^2}$ .
  - Wald test: Suppose  $H_0$  is equivalent to a set of equations  $R(\theta) = 0$  where R is a continuous function from  $\mathbb{R}^k \to \mathbb{R}^l$ . Wald (1943) introduces the following statistic:

$$W_n = R(\hat{\theta})^\mathsf{T} \{ c(\hat{\theta})^\mathsf{T} I_n(\hat{\theta})^{-1} c(\hat{\theta}) \}^{-1} R(\hat{\theta}),$$

where  $C(\theta) = \partial R(\theta)/\partial \theta$ ,  $I_n(\hat{\theta})$  is the fisher information matrix based on  $X_1, \ldots, X_n$  and  $\hat{\theta}$  is the MLE of  $\theta$ . Then, under regularity conditions,  $W_n \xrightarrow{d} \chi_r^2$ , where r equals the difference of number of all free parameters and free parameters in  $\Theta_0$ .

- For testing  $H_0: \theta = \theta_0$  with  $\theta_0$  known,  $R(\theta) = \theta - \theta_0$  and  $W_n$  simplifies to

$$W_n = (\hat{\theta} - \theta_0)^{\mathsf{T}} I_n(\hat{\theta}) (\hat{\theta} - \theta_0).$$

- Score test: Rao (1947) proposed the following statistic:

$$R_n = s_n(\hat{\theta}_0)^{\mathsf{T}} I_n(\hat{\theta}_0)^{-1} s_n(\hat{\theta}_0),$$

where  $\hat{\theta}_0$  is the MLE under  $H_0$  and  $s_n(\theta) = \partial \log f_{\theta}(x)/\partial \theta$  is called the score function. Under regularity conditions,  $W_n \xrightarrow{d} \chi_r^2$ .

- LRT, Wald test and score test are asymptotic equivalent.
- Asymptotic confidence sets:
  - **Asymptotic pivots**: A known function  $q_n(\mathbf{X}, \theta)$  is a *Asymptotic pivot* if the asymptotic distribution of  $q_n(\mathbf{X}, \theta)$  is free of the unknown parameter  $\theta$ .
  - We can also obtain asymptotic confidence sets by inverting asymptotic tests.

## 2 Questions

1. Let  $X_1, \ldots, X_n$  be i.i.d. samples from  $\mathcal{N}(\mu, \sigma^2)$ . Derive a score test statistic for testing  $H_0: \sigma = \sigma_0$  if  $\mu$  is known.

2. Let  $X_1, \ldots, X_n$  be a random sample from  $\mathcal{N}(\mu, \varphi)$  with unknown  $\theta = (\mu, \varphi)$ . Obtain  $1 - \alpha$  asymptotically correct confidence sets for  $\mu$  by inverting acceptance regions of LR tests, Wald's tests and score test.

- 3. Let  $X_1, \ldots, X_n$  be i.i.d. negative binomial(r, p). Consider  $Y = \sum_{i=1}^n X_i \sim \text{nb}(nr, p)$ .
  - (a) Prove that  $2pY \xrightarrow{d} \chi^2_{2nr}$  as  $p \to 0$ .
  - (b) Show that for small p, the interval

$$\left\{ p : \frac{\chi_{2nr,1-\alpha/2}^2}{2Y} \le p \le \frac{\chi_{2nr,\alpha/2}^2}{2Y} \right\}$$

is an approximate  $1-\alpha$  confidence interval.

(c) Obtain a minimum length  $1 - \alpha$  CI based on the asymptotic pivot 2pY.

 $\textit{Hint: mgf of $\chi^2_r$ is $(1-2t)^{-k/2}$ for $t<\frac{1}{2}$, and mgf of $nb(r,p)$ is $[\frac{p}{1-(1-p)e^t}]^r$ for $t<\log(1-p)$.}$