

# IS 604 Assignment 2

*David Stern*

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1. Suppose that  $X$  is a discrete random variable having probability function  $Pr(X = k) = ck^2$  for  $k = 1, 2, 3$ . Find  $c$ ,  $Pr(X \leq 2)$ ,  $E[X]$ , and  $Var(X)$ .

First, we will solve for  $c$  by summing the probabilities:

$$P(x = 1) = 1^2c$$

$$P(x = 2) = 2^2c$$

$$P(x = 3) = 3^2c$$

$$P(X = 1) + P(X = 2) + P(X = 3) = 1$$

$$c + 4c + 9c = 1$$

$$c = \frac{1}{14}$$

To find the probability that  $x$  is less than or equal to 2, we need to sum the probabilities that  $x$  is 1 or 2.

$$P(X \leq 2) = P(x = 1) + P(x = 2) = \frac{1}{14} + 4\left(\frac{1}{14}\right) = \frac{5}{14}$$

Now we will find  $E[X]$  by summing the product of each outcome by its probability.

$$E[X] = 1 \times \frac{1}{14} + 2 \times \frac{4}{14} + 3 \times \frac{9}{14} = \frac{36}{14} = 2.57$$

With  $E[X]$ , we can calculate  $Var[X]$ :

$$Var[X] = (1 - 2.57)^2 \times \frac{1}{14} + (2 - 2.57)^2 \times \frac{4}{14} + (3 - 2.57)^2 \times \frac{9}{14} = 0.3877571$$

2. Suppose that  $X$  is a continuous random variable having p.d.f.  $f(x) = cx^2$  for  $1 \leq x \leq 2$ . Find  $c$ ,  $Pr(X \geq 1)$ ,  $E[X]$ , and  $Var(X)$ .

$$\int_1^2 cx^2 dx = \left[ \frac{cx^3}{3} \right]_1^2$$

$$\frac{8c}{3} - \frac{c}{3} = 1$$

$$c = \frac{3}{7}$$

Since the range of value is  $1 \leq x \leq 2$ , the  $P(x \geq 1) = 1$ .

$$E[X] = \int_a^b xf(x)dx$$

$$E[X] = \int_1^2 x \frac{3x^2}{7} dx = \left[ \frac{3x^4}{28} + C \right]_1^2 = \left( \frac{48}{28} + C \right) - \left( \frac{3}{28} + C \right) = \frac{45}{28} = 1.607$$

$$Var[X] = E[X^2] - E[X]^2 = \int_1^2 x^2 \frac{3x^2}{7} dx - \left( \frac{45}{28} \right)^2 = \left[ \frac{3x^5}{35} + C \right]_1^2 - \left( \frac{45}{28} \right)^2 = \frac{96}{35} - \frac{3}{35} - \frac{2025}{784} = 0.0742$$

3. Suppose that X and Y are jointly continuous random variables with:  $y - x$  for  $0 < x < 1$  and  $1 < y < 2$ , otherwise 0

a. Compute and plot  $f_X(x)$  and  $f_Y(y)$ .

$$\text{for } 0 < x < 1, \quad f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x, y) dy$$

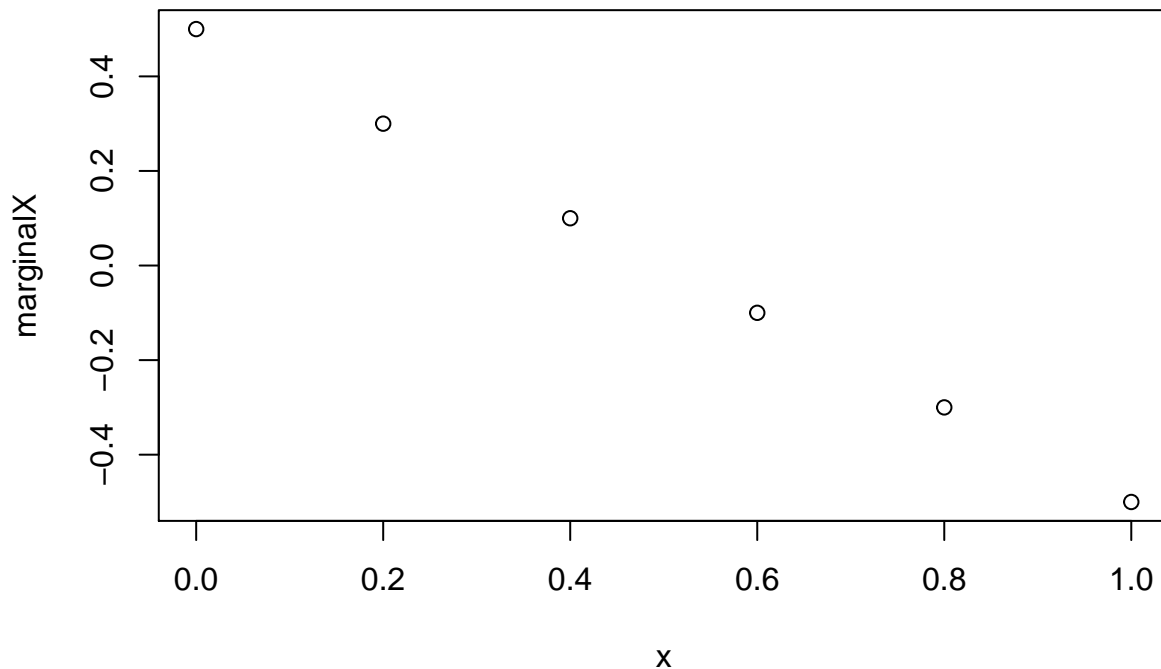
$$f_X(x) = \int_0^1 y - x dy = \left[ \frac{y^2}{2} - xy \right]_0^1$$

$$f_X(x) = \frac{1}{2} - x$$

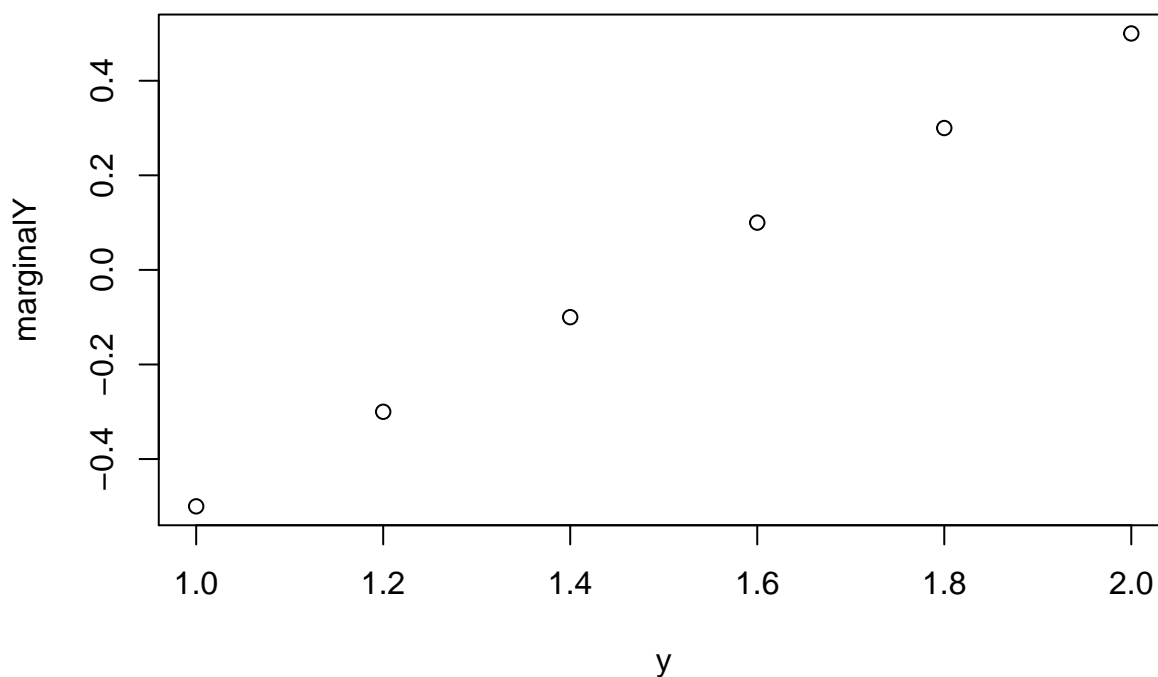
$$\text{for } 1 < y < 2, \quad f_Y(y) = \int_1^2 y - x dx = \left[ xy - \frac{x^2}{2} \right]_1^2$$

$$f_Y(y) = (2y - 2) - \left( y - \frac{1}{2} \right) = y - \frac{3}{2}$$

```
x <- seq(0,1,by=0.2)
fxx <- function(x){0.5-x}
marginalX <- fxx(x)
y <- seq(1,2,by=0.2)
fyy <- function(y){y-1.5}
marginalY <- fyy(y)
plot(x,marginalX)
```



```
plot(y,marginalY)
```



b. Are X and Y independent?

If  $E[XY] = E[X]E[Y]$ , then X and Y are independent.

From the solutions below, we know that  $E[X] = \frac{-1}{12}$ ,  $E[Y] = \frac{1}{12}$ , and  $E[XY] = \frac{2}{3}$ . Since  $\frac{2}{3} \neq \frac{-1}{144}$ , we know that X and Y are not independent.

c. Compute  $F_X(x)$  and  $F_Y(y)$ .

$$\begin{aligned}
 f_X(x) &= \frac{dF_X x}{dx} \\
 f_X(x) &= \frac{1}{2} - x \\
 F_X x &= \int_{-\infty}^{\infty} \frac{1}{2} - x = \frac{x - x^2}{2} + C \\
 f_Y(y) &= y - \frac{3}{2} \\
 F_Y(y) &= \int_{-\infty}^{\infty} y - \frac{3}{2} = \frac{y^2 - 3y}{2} + C
 \end{aligned}$$

d. Compute  $E[X]$ ,  $\text{Var}(X)$ ,  $E[Y]$ ,  $\text{Var}(Y)$ ,  $\text{Cov}(X,Y)$ , and  $\text{Corr}(X,Y)$ .

$$E[X] = \int_0^1 u f_x(u) du = \int_0^1 u \left( \frac{1}{2} - u \right) du = \int_0^1 \frac{u}{2} - u^2 = \left[ \frac{u^2}{4} - \frac{u^3}{3} \right]_0^1 = \frac{1}{4} - \frac{1}{3} = -\frac{1}{12}$$

$$E[X] = -\frac{1}{12}$$

$$Var[X] = \int_0^1 u^2 f_x(u) du - \left(\frac{-1}{12}\right)^2$$

$$Var[X] = \int_0^1 \frac{u^2}{2} - u^3 - \frac{1}{144}$$

$$Var[X] = \left[ \frac{u^3}{6} - \frac{u^4}{4} \right]_0^1 - \frac{1}{144}$$

$$Var[X] = \frac{1}{6} - \frac{1}{4} - \frac{1}{144} = \frac{24}{144} - \frac{36}{144} - \frac{1}{144} = \frac{-13}{144}$$

$$E[Y] = \int_1^2 u f_y(u) du = \int_1^2 u(u - \frac{3}{2}) du = \int_1^2 u^2 - \frac{3u}{2} du$$

$$E[Y] = \left[ \frac{u^3}{3} - \frac{3u^2}{4} \right]_1^2$$

$$E[Y] = \left(\frac{8}{3} - \frac{12}{4}\right) - \left(\frac{1}{3} - \frac{3}{4}\right)$$

$$E[Y] = \frac{1}{12}$$

$$Var[Y] = \int_1^2 u^2 f_y(u) du - \left(\frac{1}{12}\right)^2$$

$$Var[Y] = \int_1^2 u^2(u - \frac{3}{2}) du - \frac{1}{144}$$

$$Var[Y] = \int_1^2 u^3 - \frac{3u^2}{2} du - \frac{1}{144}$$

$$Var[Y] = \left[ \frac{u^4}{4} - \frac{u^3}{2} \right]_1^2 - \frac{1}{144}$$

$$Var[Y] = \left(\frac{16}{4} - \frac{8}{2}\right) - \left(\frac{1}{4} - \frac{1}{2}\right) - \frac{1}{144}$$

$$Var[Y] = \frac{35}{144}$$

$$Cov(X, Y) = E[XY] - E[X]E[Y]$$

$$E[XY] = \int_0^1 \int_1^2 xy(y - x) dy dx = \int_0^1 \int_1^2 xy^2 - x^2y dy dx$$

$$E[XY] = \int_1^2 \left[ \frac{x^2y^2}{2} - \frac{x^3y}{3} \right]_0^1 dy$$

$$E[XY] = \int_1^2 \frac{y^2}{2} - \frac{y}{3} dy = \left[ \frac{y^3}{6} - \frac{y^2}{6} \right]_1^2$$

$$E[XY] = \left(\frac{8}{6} - \frac{4}{6}\right) - \left(\frac{1}{6} - \frac{1}{6}\right) = \frac{2}{3}$$

$$Cov(X, Y) = \frac{2}{3} - \frac{-1}{12} \times \frac{1}{12}$$

$$Cov(X, Y) = \frac{49}{144} = 0.34$$

$$Corr(X, Y) = \frac{Cov(X, Y)}{\sqrt{Var(X)Var(Y)}} = \frac{\frac{49}{144}}{\sqrt{\frac{-13}{144} \times \frac{35}{144}}}$$

4. Suppose that the following 10 observations come from some distribution (not highly skewed) with unknown mean  $\mu$ .

```
sample <- c(7.3, 6.1, 3.8, 8.4, 6.9, 7.1, 5.3, 8.2, 4.9, 5.8)
mean(sample)
```

```
## [1] 6.38
```

```
var(sample)
```

```
## [1] 2.161778
```

The sample mean is  $\bar{X} = 6.38$  and the variance is  $S^2 = 2.161778$ .

We know that 95% of the values in a normally distributed population will fall within 2 standard deviations of the mean. We can construct the 95% confidence interval for  $\mu$  as follows:

$$\begin{aligned} \bar{x} \pm 2\sigma_x \\ 6.38 \pm 2\sqrt{\left(\frac{2.161778}{10}\right)} \\ 6.38 \pm 0.93 = (5.45, 7.31) \end{aligned}$$

5. A random variable  $X$  has the memoryless property if, for all  $s, t > 0$ ,  $P(x > t + s \mid x > t) = P(x > s)$ .

Show that the exponential distribution has the memoryless property.

We can rearrange the above equivalency as follows:

$$\begin{aligned} \frac{P(x > t + s, x > t)}{P(x > t)} &= P(x > s) \\ P(x > s + t) &= P(x > s)P(x > t) \end{aligned}$$

If we take the exponential distribution as:  $f(x, \lambda) = \lambda e^{-\lambda x}$ , then we can see that the equation holds true for all values of  $s$  as long as  $t \neq 0$ . This equality makes sense when we consider a scenario with waiting times where  $x$  is the time it takes for an event to occur. If an event fails to occur within the initial time  $s$ , then the probability of the event occurring in the initial time  $t$  should be equal to the unconditioned probability (the right-hand side).

$$\begin{aligned} P(x > t) &= e^{-\lambda t} \\ P(x > s) &= e^{-\lambda s} \\ P(x > s + t) &= e^{-\lambda s} e^{-\lambda t} = e^{-\lambda(s+t)} \end{aligned}$$

6. Suppose  $X_1, X_2, \dots, X_n$  are i.i.d.  $Exp(\lambda = 1)$ . Use the Central Limit Theorem to find the approximate value of  $P(100 \leq \sum_{i=1}^{100} X_i \leq 110)$ .

First, we know that the mean and the variance of an exponentially distributed random with  $\lambda = 1$  are:

$$E[X] = \frac{1}{\lambda} = 1$$

$$Var[X] = \frac{1}{\lambda^2} = 1$$

From the Central Limit Theorem, we know that the sample mean,  $\bar{X} = \frac{\sum_{i=1}^n X_i}{n}$ . Since  $n = 100$ , we can rewrite  $P(100 \leq \sum_{i=1}^{100} X_i \leq 110)$  as:

$$P(1 \leq \bar{X} \leq 1.1) = F(1.1) - F(1)$$

We determine the probability by finding the z-scores that correspond to the standard normal distribution.

$$z = \frac{X - \mu}{\frac{\sigma}{\sqrt{n}}} = \frac{1.1 - 1}{\frac{1}{\sqrt{100}}} = 1$$

The area under the curve to the left of  $z=1$  is 0.8413.

$$F(1.1)z = \frac{X - \mu}{\frac{\sigma}{\sqrt{n}}} = \frac{1 - 1}{\frac{1}{\sqrt{100}}} = 0$$

The area under the curve to the left of  $z=0$  is 0.5000.

$$P(1 \leq \bar{X} \leq 1.1) = 0.8413 - 0.5000 = 0.3413$$

Exercises from Discrete-Event System Simulation:

5.13

A random variavle  $X$  that has a pmf given by  $p(x) = 1/(n + 1)$  over the range  $R_X = 0, 1, 2, \dots, n$  is said to have a discrete uniform distribution.

- a. Find the mean and variance of this distribution.

The mean of the distribution is:

$$E[X] = \frac{1}{n} \sum_{i=1}^n x_i$$

$$\sum_{i=1}^n x_i = \frac{n(n+1)}{2}$$

$$E[X] = \frac{1}{n} \cdot \frac{n(n+1)}{2} = \frac{n+1}{2}$$

$$\begin{aligned}
\text{Var}[X] &= E[X^2] - E[X]^2 \\
E[X^2] &= \frac{1}{n} \sum_{i=1}^n x_{i^2} \\
\sum_{i=1}^n x_{i^2} &= \frac{n(n+1)(2n+1)}{6} \\
E[X^2] &= \frac{1}{n} \cdot \frac{n(n+1)(2n+1)}{6} = \frac{(n+1)(2n+1)}{6} \\
\text{Var}[X] &= \frac{(n+1)(2n+1)}{6} - \frac{(n+1)^2}{4} \\
\text{Var}[X] &= \frac{n^2 - 1}{12}
\end{aligned}$$

b. If  $R_X = a, a+1, a+2, \dots, b$ , compute the mean and variance of  $X$ .

If the discrete uniform distribution has parameters  $a$  and  $b$  where  $a < b$ , then the probability mass function is  $f(x) = \frac{1}{b-a+1}$ . The solution for  $E[X]$  above shows use that we should expect the mean to fall in the middle of a discrete uniform distribution. If our interval is  $(a, b)$ , we can intuit that the mean should be:

$$E[X] = \frac{a+b}{2}$$

We also know that the variance will behave the same over the interval  $(a, b)$ , so we can plug in  $n = b - a + 1$  in  $\text{Var}[X] = \frac{n^2-1}{12}$ :

$$\text{Var}[X] = \frac{(b-a+1)^2 - 1}{12}$$

5.14

The lifetime, in years, of a satellite placed in orbit is given by the following pdf:

$$f(x) = 0.4e^{-0.4x} \text{ when } x \geq 0, \text{ otherwise } 0$$

$$P(x > t) = \int_t^\infty \lambda e^{-\lambda x} dx = \left[ \lambda \frac{-e^{-\lambda x}}{\lambda} \right]_t^\infty = e^{-\lambda t}$$

a. What is the probability that this satellite is still “alive” after 5 years?

$$P(x > 5) = e^{-(0.4)(5)} = 0.135$$

b. What is the probability that the satellite dies between 3 and 6 years from the time it is placed in orbit?

$$P(3 \leq x \leq 6) = F(6) - F(3) = (1 - e^{-(0.4)(6)}) - (1 - e^{-(0.4)(3)})$$

$$P(3 \leq x \leq 6) = 0.210$$

5.39

Three shafts are made and assembled into a linkage. The length of each shaft, in centimeters, is distributed as follows:

Shaft 1:  $N(60, 0.09)$  Shaft 2:  $N(40, 0.05)$  Shaft 3:  $N(50, 0.11)$

- a. What is the distribution of the length of the linkage?

The sum of multiple mutually independent normal random variables also has a normal distribution. The mean and variance of the distribution can be summed from the individual distributions.

$$E[X] = 60 + 40 + 50 = 150$$

$$Var[X] = 0.09 + 0.05 + 0.11 = 0.25$$

The distribution of the linkage is  $N(150, 0.25)$ .

- b. What is the probability that the linkage will be longer than 150.2 centimeters?

Here we will use a z-table to find  $P(X > 150.2)$ .

$$z = \frac{150.2 - 150}{\sqrt{0.25}} = 0.4$$

$$P(X > 150.2) = 1 - P(X \leq 150.2)$$

$$= 1 - 0.6554$$

$$= 0.3446$$

The probability that the linkage is longer than 150.2 cm is 0.3446.

- c. The tolerance limits for the assembly are (149.83, 150.21). What proportion of assemblies are within the tolerance limits?

We will use z-scores again the portion of the assemblies that fall within the tolerance limits.

$$P(149.83 \leq X \leq 150.21) = P(X \leq 150.21) - P(X \leq 149.83)$$

$$z = \frac{150.21 - 150}{\sqrt{0.25}} = 0.42$$

The area under the curve to the left of  $z = 0.42$  is 0.6628.

$$z = \frac{149.83 - 150}{\sqrt{0.25}} = -0.34$$

The area under the curve to the left of  $z = -0.34$  is 0.3669.

$$0.6628 - 0.3669 = 0.2959$$

The proportion of assemblies that fall within the tolerance limits is 0.2959.