

Linear Regression

1 Derived Distributions of OLS Estimates

1.1 General Matrix Expression

$$\begin{aligned} \mathbf{Y} &= \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}, \text{ where } \boldsymbol{\epsilon} \stackrel{\text{i.i.d.}}{\sim} \text{Dist}(\mathbf{0}_n, \sigma^2 \mathbf{I}_{n \times n}) \\ \boldsymbol{\beta} &= (\mathbf{X}'\mathbf{X})^{-1} \mathbf{X}'\mathbf{Y} \\ \text{cov}(\boldsymbol{\beta}) &= \sigma^2 (\mathbf{X}'\mathbf{X})^{-1} \\ \boldsymbol{\beta} &\sim \text{Dist}(\boldsymbol{\beta}, \sigma^2 (\mathbf{X}'\mathbf{X})^{-1}) \end{aligned}$$

1.2 When Errors are Normally Distributed

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i, \text{ where } \epsilon_i \stackrel{\text{i.i.d.}}{\sim} N(0, \sigma^2)$$

$$1. \quad \hat{\beta}_1 = \frac{\sum_{i=1}^n (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\sum_{i=1}^n (X_i - \bar{X})Y_i}{\sum_{i=1}^n (X_i - \bar{X})^2} = \sum_{i=1}^n W_i Y_i \text{ (i.e., weighted sum),}$$

where

$$W_i = \frac{(X_i - \bar{X})}{\sum_{i=1}^n (X_i - \bar{X})^2}, \text{ so } \sum_{i=1}^n W_i (X_i - \bar{X}) = \sum_{i=1}^n \frac{(X_i - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^2} = 1, \text{ then}$$

$$\begin{aligned} \sum_{i=1}^n W_i Y_i &= \sum_{i=1}^n W_i (\beta_0 + \beta_1 X_i + \epsilon_i) = \sum_{i=1}^n W_i (\bar{Y} - \beta_1 \bar{X} - \bar{\epsilon} + \beta_1 X_i + \epsilon_i) \\ &= \bar{Y} \sum_{i=1}^n W_i - \bar{\epsilon} \sum_{i=1}^n W_i + \beta_1 \sum_{i=1}^n W_i (X_i - \bar{X}) + \sum_{i=1}^n W_i \epsilon_i \\ &= \bar{Y} \sum_{i=1}^n W_i - 0 + \beta_1 + \sum_{i=1}^n W_i \epsilon_i = \hat{\beta}_1 \end{aligned}$$

$$\hat{\beta}_1 - \beta_1 = \text{constant} + \sum_{i=1}^n W_i \epsilon_i$$

$$\begin{aligned}
Var(\hat{\beta}_1) &= E[(\hat{\beta}_1 - \beta_1)^2] = E[(\sum_{i=1}^n W_i \epsilon_i)^2] = E[\sum_{i=1}^n W_i^2 \epsilon_i^2 + \sum_{i=1}^n \sum_{j \neq i}^n W_i W_j \epsilon_i \epsilon_j] \\
&= \sum_{i=1}^n W_i^2 E(\epsilon_i^2) + \sum_{i=1}^n \sum_{j \neq i}^n W_i W_j E(\epsilon_i \epsilon_j) = \sum_{i=1}^n W_i^2 E(\epsilon_i^2) \\
E(\epsilon_i \epsilon_j) &= 0 \text{ due to the uncorrelation assumption between error pair} \\
&= \sigma^2 \sum_{i=1}^n \frac{(X_i - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^4} = \frac{\sigma^2}{\sum_{i=1}^n (X_i - \bar{X})^2}
\end{aligned}$$

$$\boxed{\hat{\beta}_1 \sim N(\beta_1, \frac{\sigma^2}{\sum_{i=1}^n (X_i - \bar{X})^2})}$$

2. $\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}$

$$\begin{aligned}
Var(\hat{\beta}_0) &= Var(\bar{Y} - \hat{\beta}_1 \bar{X}) = Var(\frac{1}{n} \sum_{i=1}^n Y_i) + (\bar{X})^2 Var(\hat{\beta}_1) \\
&= \frac{1}{n^2} \sum_{i=1}^n Var(\epsilon_i) + \frac{\sigma^2 \bar{X}^2}{\sum_{i=1}^n (X_i - \bar{X})^2} = \frac{\sigma^2}{n} + \frac{\sigma^2 \bar{X}^2}{\sum_{i=1}^n (X_i - \bar{X})^2}
\end{aligned}$$

$$\boxed{\hat{\beta}_0 \sim N(\beta_0, \sigma^2(\frac{1}{n} + \frac{\bar{X}^2}{\sum_{i=1}^n (X_i - \bar{X})^2}))}$$

3. $\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 X$

$$\begin{aligned}
Var(\hat{Y}) &= Var(\hat{\beta}_0 + \hat{\beta}_1 X) = Var(\hat{\beta}_0) + X^2 Var(\hat{\beta}_1) \\
&= \sigma^2(\frac{1}{n} + \frac{\bar{X}^2}{\sum_{i=1}^n (X_i - \bar{X})^2}) + X^2 \frac{\sigma^2}{\sum_{i=1}^n (X_i - \bar{X})^2} \\
&= \sigma^2(\frac{1}{n} + \frac{(X - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^2})
\end{aligned}$$

$$\boxed{\hat{Y} \sim N(\beta_0 + \beta_1 X, \sigma^2(\frac{1}{n} + \frac{(X - \bar{X})^2}{\sum_{i=1}^n (X_i - \bar{X})^2}))}$$