

Problem 1

(a) First, we compute $P(A)$, $P(B)$, $P(A \cap B)$

$$P(A) = 1 - p$$

$$P(B) = p\varepsilon_0 + (1 - p)\varepsilon_1$$

$$P(A \cap B) = (1 - p)\varepsilon_1$$

Next, we check when $P(A \cap B) = P(A) \cdot P(B)$

$$\begin{aligned} (1 - p)\varepsilon_1 &= (1 - p)p\varepsilon_0 + (1 - p)^2\varepsilon_1 \\ \Rightarrow p(1 - p)(\varepsilon_0 - \varepsilon_1) &= 0 \end{aligned}$$

Therefore, A and B are independent if $p = 0$, $p = 1$, or $\varepsilon_0 = \varepsilon_1$.

(b) First, we compute $P(A|C)$, $P(B|C)$, $P(A \cap B|C)$

$$P(A|C) = \frac{P(A \cap C)}{P(C)} = \frac{(1 - p)(\alpha_1 + \varepsilon_1)}{p(\alpha_0 + \varepsilon_0) + (1 - p)(\alpha_1 + \varepsilon_1)}$$

$$P(B|C) = \frac{P(B \cap C)}{P(C)} = \frac{p\varepsilon_0 + (1 - p)\varepsilon_1}{p(\alpha_0 + \varepsilon_0) + (1 - p)(\alpha_1 + \varepsilon_1)}$$

$$P(A \cap B|C) = \frac{P((A \cap B) \cap C)}{P(C)} = \frac{(1 - p)\varepsilon_1}{p(\alpha_0 + \varepsilon_0) + (1 - p)(\alpha_1 + \varepsilon_1)}$$

Next, we check if $P(A \cap B|C) = P(A|C) \cdot P(B|C)$

$$\begin{aligned} P(A \cap B|C) &= \frac{(1 - p)\varepsilon_1}{p(\alpha_0 + \varepsilon_0) + (1 - p)(\alpha_1 + \varepsilon_1)} \\ P(A|C) \cdot P(B|C) &= \frac{(1 - p)(\alpha_1 + \varepsilon_1)}{p(\alpha_0 + \varepsilon_0) + (1 - p)(\alpha_1 + \varepsilon_1)} \cdot \frac{p\varepsilon_0 + (1 - p)\varepsilon_1}{p(\alpha_0 + \varepsilon_0) + (1 - p)(\alpha_1 + \varepsilon_1)} \end{aligned}$$

We can simplify the check to

$$\begin{aligned} &P((A \cap B) \cap C) \cdot P(C) \\ &= (1 - p)\varepsilon_1 \cdot [p(\alpha_0 + \varepsilon_0) + (1 - p)(\alpha_1 + \varepsilon_1)] \\ &= (1 - p)[p\varepsilon_1\alpha_0 + p\varepsilon_0\varepsilon_1 + (1 - p)\varepsilon_1\alpha_1 + (1 - p)\varepsilon_1^2] \\ &P(A \cap C) \cdot P(B \cap C) \\ &= (1 - p)[p\varepsilon_0\alpha_1 + p\varepsilon_0\varepsilon_1 + (1 - p)\varepsilon_1\alpha_1 + (1 - p)\varepsilon_1^2] \\ &= (1 - p)(\alpha_1 + \varepsilon_1)[p\varepsilon_0 + (1 - p)\varepsilon_1] \end{aligned}$$

Therefore, events A and B are conditionally independent given the event C if $p = 0$, $p = 1$, or $\varepsilon_0\alpha_1 = \varepsilon_1\alpha_0$.

(c) By total probability theorem, we have

$$P(\text{Bit delivered correctly}) = p(a - \varepsilon_0 - \alpha_0) + (1 - p)(1 - \varepsilon_1 - \alpha_1)$$

Problem 2

- (a) Let X be the number of "1"s transmitted in the time interval. Given $Z = n$, X follows a binomial distribution

$$X|(Z = n) \sim \text{Binomial}(n, p)$$

Then, the probability that $X = k$ is

$$\begin{aligned} P(X = k) &= \sum_{n=k}^{\infty} P(X = k|Z = n)P(Z = n) \\ &= \sum_{n=k}^{\infty} \binom{n}{k} p^k (1-p)^{n-k} \frac{e^{-\lambda} \lambda^n}{n!} \\ &= \frac{e^{-\lambda} p^k \lambda^k}{k!} \sum_{n=k}^{\infty} \frac{(1-p)^{n-k} \lambda^{n-k}}{(n-k)!} \\ &= \frac{e^{-\lambda} p^k \lambda^k}{k!} \sum_{n=0}^{\infty} \frac{(1-p)^n \lambda^n}{n!} \\ &= \frac{e^{-\lambda} p^k \lambda^k}{k!} e^{\lambda(1-p)} \\ &= \frac{e^{-\lambda p} (\lambda p)^k}{k!} \end{aligned}$$

Therefore, $X \sim \text{Poisson}(\lambda p)$.

- (b) (i) This event occurs when all X_i are less than or equal to k .
 CDF of X is $F_X(k) = \prod_{i=1}^n P(X_i \leq k) = \prod_{i=1}^n 1 - P(X_i > k) = [1 - (1-p)^k]^n$
 PMF of X is $P(X = k) = F_X(k) - F_X(k-1) = [1 - (1-p)^k]^n - [1 - (1-p)^{k-1}]^n$
- (ii) This event occurs when at least one X_i is less than or equal to k .
 CDF of Y is $F_Y(k) = 1 - P(\min(X_1, \dots, X_n) > k) = 1 - [(1-p)^k]^n = 1 - (1-p)^{nk}$
 PMF of Y is $P(Y = k) = F_Y(k) - F_Y(k-1) = (1-p)^{n(k-1)} - (1-p)^{nk}$
- (iii) $P(Y = k) = (1-p)^{n(k-1)} - (1-p)^{nk} = (1-p)^{n(k-1)}[1 - (1-p)^n]$
 Let $p' = 1 - (1-p)^n$, then $P(Y = k) = (1-p')^{(k-1)} p'$, this is the PMF of a Geometric random variable with success probability p' .
- (c) According to the Binomial distribution, we have

$$P(X_{S_2} = k) = \binom{123}{k} p_T^k (1-p_T)^{123-k} \text{ for } k = 0, 1, \dots, 123$$

Problem 3

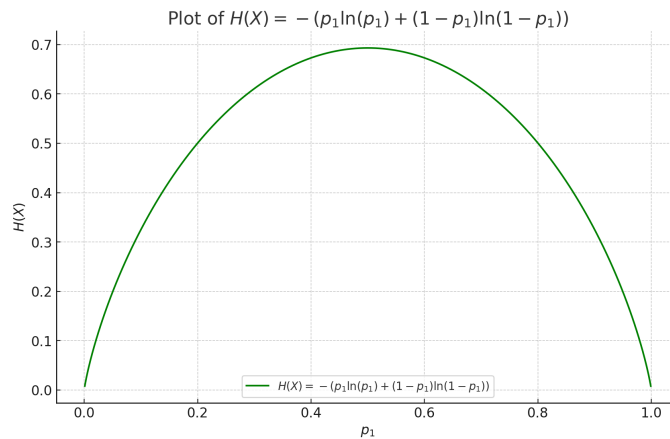
(a)

$$\begin{aligned}
 H\left(\frac{1}{2}, \frac{1}{3}, \frac{1}{6}\right) &= -\left(\frac{1}{2} \ln \frac{1}{2} + \frac{1}{3} \ln \frac{1}{3} + \frac{1}{6} \ln \frac{1}{6}\right) \\
 &= \left(\frac{1}{2} + \frac{1}{6}\right) \ln 2 + \left(\frac{1}{3} + \frac{1}{6} \ln 3\right) \\
 &= \frac{2}{3} \ln 2 + \frac{1}{2} \ln 3 \\
 H\left(\frac{1}{2}, \frac{1}{2}\right) + \frac{1}{2} \cdot H\left(\frac{2}{3}, \frac{1}{3}\right) &= -\left(\frac{1}{2} \ln \frac{1}{2} + \frac{1}{2} \ln \frac{1}{2}\right) - \frac{1}{2} \cdot \left(\frac{2}{3} \ln \frac{2}{3} + \frac{1}{3} \ln \frac{1}{3}\right) \\
 &= \ln 2 + \frac{1}{2} \left(\frac{2}{3} (\ln 3 - \ln 2) + \frac{1}{3} \ln 3\right) \\
 &= \frac{2}{3} \ln 2 + \frac{1}{2} \ln 3
 \end{aligned}$$

Therefore, $H\left(\frac{1}{2}, \frac{1}{3}, \frac{1}{6}\right) = H\left(\frac{1}{2}, \frac{1}{2}\right) + \frac{1}{2} \cdot H\left(\frac{2}{3}, \frac{1}{3}\right)$.

(b) Maximum entropy: $H_{\max} = \ln 2$ at $p_1 = 0.5$

Minimum entropy: $H_{\min} = 0$ at $p_1 = 1$ or $p_1 = 0$



(c) Let $w_i = p_i$ and $x_i = \frac{1}{p_i}$.

$$\begin{aligned}
 \frac{p_1 \frac{1}{p_1} + \cdots + p_n \frac{1}{p_n}}{1} &= n \geq \prod_{i=1}^n \left(\frac{1}{p_i}\right)^{p_i} \\
 \Rightarrow \ln n &\geq \ln\left(\prod_{i=1}^n \left(\frac{1}{p_i}\right)^{p_i}\right) = -\sum_{i=1}^n p_i \ln p_i = H(X)
 \end{aligned}$$

The equality holds when $p_1 = \cdots = p_n = \frac{1}{n}$. Therefore, $H(X)_{\max} = \ln n$ when $\{p_i = \frac{1}{n}\}_{i=1}^n$.

(d) The entropy is minimized when the random variable X is deterministic. For a specific $i \in \{1, 2, \dots, n\}$

$$H(X) = -(1 \cdot \ln 1 + \sum_{j \neq i}^n 0 \cdot \ln 0) = 0$$

Therefore, $H(X)_{\min} = 0$ is achieved by any PMF where exactly one $i \in \{1, 2, \dots, n\}$, $p_i = 1$ and $p_j = 0$ for all $j \neq i$.

Problem 4

(a) (i)

$$E[X] = \sum_{k=1}^{\infty} k \cdot (1-p)^{k-1} p = p \sum_{k=1}^{\infty} k \cdot (1-p)^{k-1} = p \cdot \frac{1}{(1-(1-p))^2} = \frac{1}{p}$$

(ii)

$$\begin{aligned} E[e^{tX}] &= \sum_{k=1}^{\infty} e^{tk} \cdot (1-p)^{k-1} p \\ &= p \sum_{k=1}^{\infty} e^{tk} \cdot (1-p)^{k-1} \\ &= pe^t \sum_{k=1}^{\infty} [e^t(1-p)]^{k-1} \\ &= pe^t \cdot \frac{1}{1-(1-p)e^t} = \frac{pe^t}{1-(1-p)e^t} \end{aligned}$$

Additionally, we should verify the convergence condition

$$|e^t(1-p)| < 1 \Rightarrow e^t < \frac{1}{1-p} \Rightarrow t < \ln \frac{1}{1-p} = -\ln(1-p)$$

(b) (i) To find $Var[Z] = E[Z^2] - (E[Z])^2$, we first compute $E[Z]$ and $E[Z^2]$

$$E[Z^2] = \sum_{n=1}^{\infty} z_n^2 \cdot p_Z(z_n) = \sum_{n=1}^{\infty} n \cdot \frac{6}{(\pi n)^2} = \frac{6}{\pi^2} \sum_{n=1}^{\infty} \frac{1}{n} = \infty$$

Therefore, since $E[Z^2] = \infty$, the variance $Var[Z]$ does not exist.

(ii)

$$\sum_{n=1}^{\infty} z_n^3 \cdot p_Z(z_n) = \sum_{n=1}^{\infty} (-1)^n (\sqrt{n})^3 \cdot \frac{6}{(\pi n)^2} = \frac{6}{\pi^2} \sum_{n=1}^{\infty} (-1)^n \frac{1}{\sqrt{n}} = -\frac{6}{\pi^2} \eta\left(\frac{1}{2}\right) \approx -\frac{6}{\pi^2} \cdot 0.6049$$

(iii)

$$E[|Z^3|] = \sum_{n=1}^{\infty} |z_n^3| \cdot p_Z(z_n) = \sum_{n=1}^{\infty} n^{\frac{3}{2}} \cdot \frac{6}{(\pi n)^2} = \frac{6}{\pi^2} \sum_{n=1}^{\infty} \frac{1}{\sqrt{n}} = \infty$$

Since $E[|Z^3|]$ is infinite, $E[Z^3]$ does not exist.

Problem 5

(a) Partition: $[0.7, 0.3]$ / Prior: Uniform

- Test accuracy: 98.21%
- Observations:
 - The classifier performs exceptionally well for the ham class, with a precision and recall close to 100%.
 - The spam detection is slightly weaker, with a recall of 89%, indicating that some spam messages were missed.

(b) • Partition: $[0.7, 0.3]$ / Prior $[0.5, 0.5]$

- Test accuracy: 97.67%
- Observations:
 - * This prior slightly lowered the overall accuracy compared to the uniform prior.
 - * The spam detection was more balanced, achieving equal precision and recall, which implies the model's confidence in detecting both classes became more stable.
- Partition: $[0.7, 0.3]$ / Prior $[0.9, 0.1]$
 - Test accuracy: 98.27%
 - Observations:
 - * Setting a heavier prior towards ham improved the model's confidence in predicting spam but didn't significantly change the recall.
 - * The accuracy remains similar to the uniform prior but with slightly better precision for spam.
- Partition: $[0.8, 0.2]$ / Prior: Uniform
 - Test accuracy: 98.39%
 - Observations:
 - * Increasing the training data led to a slight improvement in accuracy.
 - * Spam detection precision increased to 99%, showing that more training data helps the model learn the nuances better.
- Partition: $[0.6, 0.4]$ / Prior: Uniform
 - Test accuracy: 98.34%
 - Observations:
 - * A larger test set still resulted in high accuracy.
 - * Spam detection improved slightly, demonstrating the model's robustness even with less training data.